

3rd way of estimating spectrum

Predictors, Yule Walker equations, Wold theorem

lunedì 25 maggio 2020 16:24

$$x \in \mathbb{R}^n \quad x \in \mathbb{C}^n$$

$x \xrightarrow{\text{predict}} w \in \mathbb{R}$
 $\in \mathbb{C}$

$$\hat{w} = \alpha^+ x$$

$$\alpha = \begin{bmatrix} \alpha_0 \\ \vdots \\ \alpha_{n-1} \end{bmatrix}$$

$$\hat{w} = \sum_{j=0}^{n-1} \alpha_j^* x_j$$

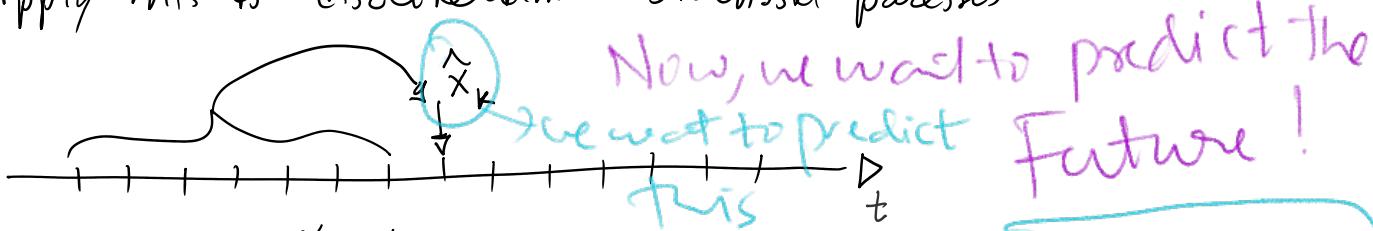
$$\mathcal{E} = E[|\varepsilon|^2] \quad \varepsilon = w - \hat{w}$$

Orthogonality principle

$$E[\varepsilon^* x] = \begin{bmatrix} E[\varepsilon^* x_0] \\ E[\varepsilon^* x_1] \\ \vdots \\ E[\varepsilon^* x_{n-1}] \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

$$E[\varepsilon^* w] = \mathcal{E}_{\min}$$

→ Apply this to discrete-time stochastic processes

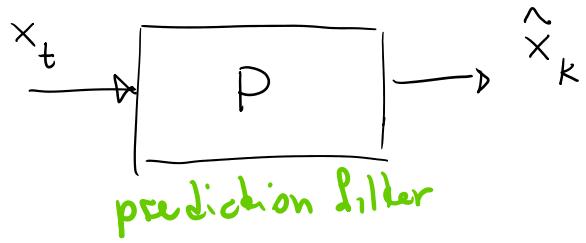


$x_{k-M}, \dots, x_{k-2}, x_{k-1}$
 memory of predictor

$$\hat{x}_k = \sum_{j=1}^M (-\alpha_j^*) x_{k-j}$$

We define the coefficient like this!

This is the equation of a filter



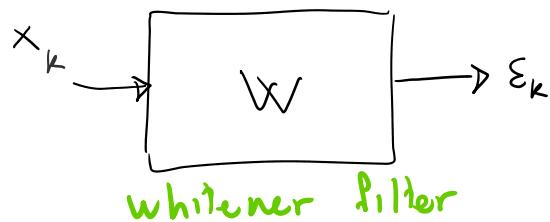
$\varepsilon_k = x_k - \hat{x}_k =$

we calculate error
after 1 CLK time =

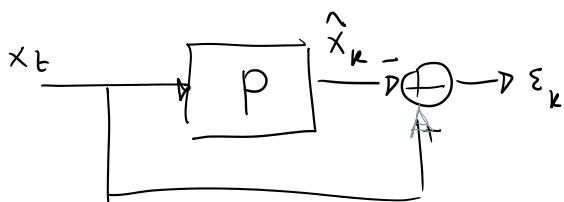
$$x_k - \sum_{j=1}^M (-\alpha_j^*) x_{k-j}$$

$$\sum_{j=0}^M \alpha_j^* x_{k-j}$$

$\alpha_0 = 1$



reproduce
a sequence of
ERRORS if
we continue
with the process
4 The sequence
of ERRORS is a
random process



$x_k \rightarrow \hat{x}_k, \varepsilon_k$ are stationary

$$\mathcal{E} = E[|\varepsilon_k|^2]$$

error is orthogonal to previous values
of \hat{x}_k

$$E[\varepsilon_k^* x_{k-l}] = 0$$

$$l = 1, 2, \dots, M$$

data considered in the prediction ✓

$$E[\varepsilon_k^* x_k] = \mathcal{E}_{\min}$$

$$M * - M$$

$$E[x_0^* x_{k-l-k+s}]$$

$$E\left[\left(\sum_{j=0}^M \alpha_j^* x_{k-j}\right)^* x_{k-e}\right] = \sum_{j=0}^M \alpha_j E[x_{k-j}^* x_{k-e}] = 0$$

$\underbrace{C_x(k-e-(k-j))}_{\parallel}$

$$\sum_{j=0}^M \alpha_j C_x(j-e) = 0 \quad \ell = 1, 2, \dots, M$$

$$E\left[\left(\sum_{j=0}^M \alpha_j^* x_{k-j}\right)^* x_k\right] = \sum_{j=0}^M \alpha_j C_x(j-e) = \varepsilon_{\min}$$

$\ell = 0$

Correlation matrix

$$\begin{bmatrix} C_x(0) & C_x(1) & \cdots & C_x(M) \\ C_x(-1) & C_x(0) & \cdots & C_x(M-1) \\ \vdots & \vdots & \ddots & \vdots \\ C_x(-M) & & & C_x(0) \end{bmatrix} = \begin{bmatrix} \alpha_0 \\ \alpha_1 \\ \vdots \\ \vdots \\ \alpha_M \end{bmatrix} = \begin{bmatrix} \varepsilon_{\min} \\ 0 \\ \vdots \\ \vdots \\ 0 \end{bmatrix}$$

R_x

Yule-Walker equations

$$R_x \begin{bmatrix} 1 \\ \alpha_1 \\ \vdots \\ \vdots \\ \alpha_M \end{bmatrix} = \begin{bmatrix} \varepsilon_{\min} \\ 0 \\ \vdots \\ \vdots \\ 0 \end{bmatrix}$$

For perfect prediction
the gis lies in the
NULLSPACE

Case 1

$\det R_x \neq 0$

Non singular

R_x is non-singular

only $0 \in \ker R_x$

$\nexists p \neq 0 \quad R_x p = 0$

no perfect prediction exists

all the unknowns are here

$\epsilon_{min} > 0$

$$R_x \begin{bmatrix} \frac{1}{\epsilon_{min}} \\ \frac{\alpha_1}{\epsilon_{min}} \\ \vdots \\ \frac{\alpha_n}{\epsilon_{min}} \end{bmatrix} = \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 1 \\ 0 \end{bmatrix}$$

$$b_0 = \frac{\alpha_0}{\epsilon_{min}} \quad \left(b_0 = \frac{1}{\epsilon_{min}} \right)$$

$$R_x \begin{bmatrix} b_0 \\ b_1 \\ \vdots \\ b_m \end{bmatrix} = \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 1 \\ 0 \end{bmatrix} \Rightarrow \begin{bmatrix} b_0 \\ b_1 \\ \vdots \\ b_m \end{bmatrix} = R_x^{-1} \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 1 \\ 0 \end{bmatrix}$$

Case 2

$\det R_x = 0$

$\exists p \quad R_x p = 0 \quad p \neq 0$

there is the chance of a perfect prediction

now ...
of a perfect prediction

$$p = \begin{bmatrix} p_0 \\ \vdots \\ p_M \end{bmatrix}$$

$$p \in \ker R_x$$

$$p_0 \neq 0$$

$$\alpha = \begin{bmatrix} 1 \\ \alpha_1 \\ \vdots \\ \alpha_M \end{bmatrix} = \underbrace{\begin{bmatrix} p \\ p_0 \end{bmatrix}}_{p_0 \neq 0}$$

$$p_0 = 0$$

$$\frac{p}{p_0} \text{ does not exist}$$

try again

$$R_x \begin{bmatrix} 1 \\ \alpha_1 \\ \vdots \\ \alpha_M \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

cannot
be zero

perfect prediction

$$p \in \ker R_x \Rightarrow p_0 = 0$$

$$\begin{bmatrix} C_x(0) & C_x(1) & C_x(2) & \dots & C_x(M) \\ C_x(-1) & C_x(0) & C_x(1) & \dots & \\ C_x(-2) & C_x(-1) & C_x(0) & \dots & \\ \vdots & \vdots & \vdots & \ddots & \\ C_x(-M) & & & & \end{bmatrix} \begin{bmatrix} C_x(0) \\ C_x(1) \\ \vdots \\ C_x(M) \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

coefficient count

discard first column of R_x
 first row p
 first row of R_x
 first row of the result

$$\begin{array}{c|ccc}
 & C(0) & C(1) & C(M) \\
 \hline
 & \cancel{C(0)} & C_x(0) & C_x(1) \dots C_x(M) \\
 & C_x(-1) & C_x(0) & \vdots \\
 & \vdots & \vdots & \vdots \\
 & C_x(-M+1) & \dots & C_x(0)
 \end{array} = \begin{bmatrix} 0 \\ \textcircled{0} \\ 0 \\ \vdots \\ 0 \end{bmatrix} \quad \mathcal{E}_{\min}$$

The same R_x

of order M



Yule-Walker equation

for a filter of order dimension reduced by '1'

M-1

— look-up vectors in kernel reduced R_x

— if every vector of the kernel starts with ϕ



discard smaller row

discard smaller column

↓

~~Yule-Walker equations for z
filter with $M-2$ lags~~

proceed until you find
a vector in \ker reduced R_x
starting with $\neq 0$



Whitener filter

(filter that produces sequence of errors)

$$\hat{x}_k = \sum_{j=1}^{\infty} (-\alpha_j^*) x_{k-j}$$

$$\varepsilon_k = \sum_{j=0}^{\infty} \alpha_j^* x_{k-j} \quad \alpha_0 = 1$$

The process of errors
in a Best predictors
is a WHITE PROCESS

"whitener" why?

$$x_t \rightarrow [w] \rightarrow \varepsilon_k$$

$$C_\varepsilon(z) = E[\varepsilon_0^* \varepsilon_z] \quad z > 0$$

$$= E\left[\left(\sum_{j=0}^{\infty} \alpha_j^* x_{0-j}\right)^* \varepsilon_z\right]$$

$$= \sum_{j=0}^{\infty} \alpha_j E[x_{0-j}^* \varepsilon_z] = \sum_{j=0}^{\infty} \alpha_j \cdot 0 = 0$$

Non Random

$$= \sum_{j=0}^{\infty} \beta_j \mathbb{E}[x_{0-j} \varepsilon_t] = \sum_{j=0}^{\infty} \beta_j \cdot 0 = 0$$

Orthogonal

$\varepsilon_t = x_t - \hat{x}_t$

$\hat{x}_t = \sum_{l=1}^{\infty} (-\beta_l^*) x_{t-l}$

$\varepsilon_t > 0 \quad t > -J$

$\varepsilon_t > 0 \quad t > 0$

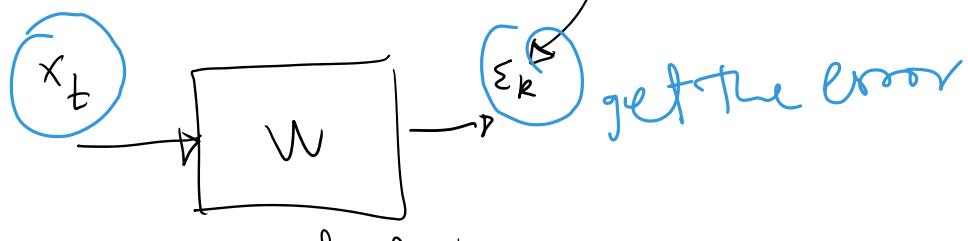
Correlation is zero from $-\infty$ to t for $t > 0$

stationarity

$$C_{\varepsilon}(z) = \begin{cases} 0 & z > 0 \\ E[|\varepsilon_0|^2] = \varepsilon_{\min} & z = 0 \\ C_{\varepsilon}^*(-z) = 0 & z < 0 \end{cases}$$

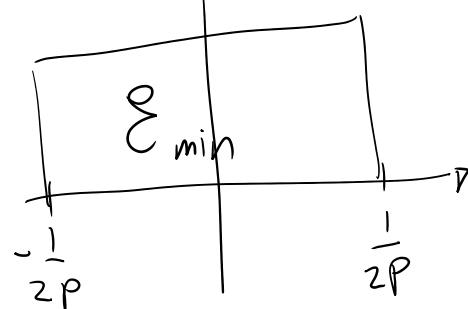
correlation of a white process

give input



transfer function

$$S_x^w(f) |W(f)|^2 = S_{\varepsilon}^w(f)$$



$$\underline{S_x^W(f)} \left| W(f) \right|^2 = P \underline{\mathcal{E}_{\min}}$$

original Power Spectrum

→ squared T/F of whitener

All pole estimation ~ maximum entropy

$$S_x^W(f) = \frac{P \mathcal{E}_{\min}}{|W(f)|^2}$$

This is one way of making Estimation

Recipe (take a large M)

- 1). collect samples
- 2). estimate $\hat{C}_x(z) \rightarrow \text{full } \hat{R}_x$

- 3). solve Yule-Walker equations → coefficients

$$\underline{\alpha_j}$$

$$\underline{\mathcal{E}_{\min}}$$

$$y) \quad \underline{\epsilon_k} = \sum_{j=0}^M \underline{\alpha_j^*} \underline{x_{k-j}}$$

improve response of whitener

$$W(f) = \sum_{j=0}^{M-1} \underline{\alpha_j^*} e^{-2\pi j f j}$$

$$\cdot \quad \underline{S_x^W(f)} = \frac{P \mathcal{E}_{\min}}{|W(f)|^2}$$

completely different way of
to estimate power spectrum

$$S_x^w(f) |W(f)|^2 = P \mathcal{E}_{\min}$$

$$\int_{-\frac{1}{2P}}^{\frac{1}{2P}} S_x^w(f) |W(f)|^2 df = \int_{-\frac{1}{2P}}^{\frac{1}{2P}} P \mathcal{E}_{\min} df$$

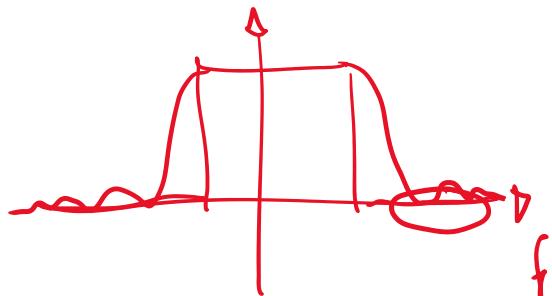
$$\int_{-\frac{1}{2P}}^{\frac{1}{2P}} S_x^w(f) |W(f)|^2 df = \mathcal{E}_{\min}$$

\rightarrow perfect predictability $\Rightarrow \mathcal{E}_{\min} = 0$

This is theoretically fundamental formulation of stochastic processes

$$\int_{-\frac{1}{2P}}^{\frac{1}{2P}} S_x^w(f) |W(f)|^2 df = 0$$

≥ 0



↓

either $S_x^w(f) = 0$

or $W(f) = 0$

possible only for some isolated frequencies

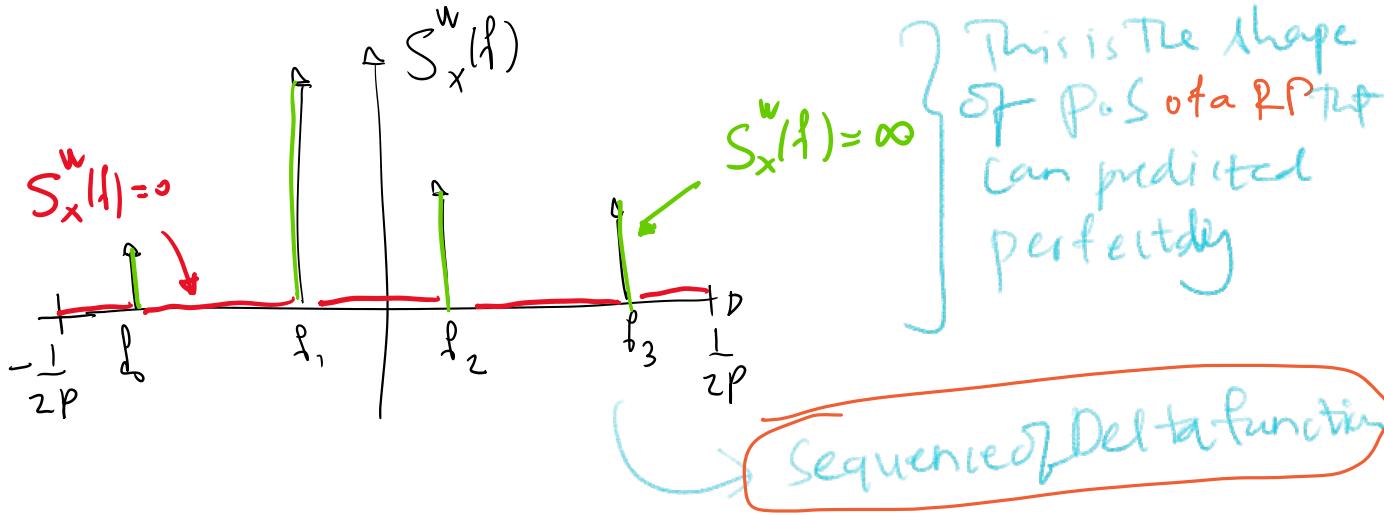
Theorem 3
Postoy -
Wiener

\rightarrow Predictable processes

$$\mathcal{E}_{\min} = 0$$

To be perfectly predictable we must multiply our process by whitener if it should be equal to zero!

$$S_x^W(f) = \sum_j \alpha_j \delta(f - f_j) \quad \sum_j \alpha_j = E[X_0^2]$$

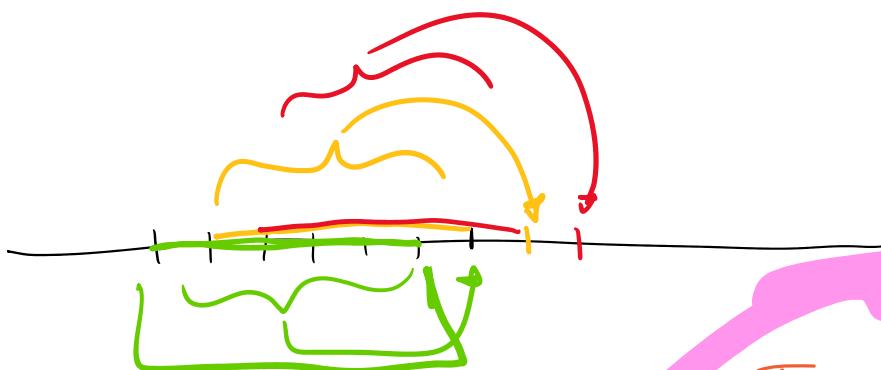


predictability = way of generating the process



$$\epsilon_{min} = 0$$

$$x_k = \hat{x}_k = \sum_{j=1}^M (-\beta_j^*) x_{k-j}$$



Texas Instrument
Frequency Synthesizers

look in the INTERNET

* Frequency Synthesis by
recursion

Regular processes
complementary of
predictable Processes

i.e They rarely touch Zero & Almost Never goto Infinity

$$S_x^w(f)$$

stays away from

ϕ

and

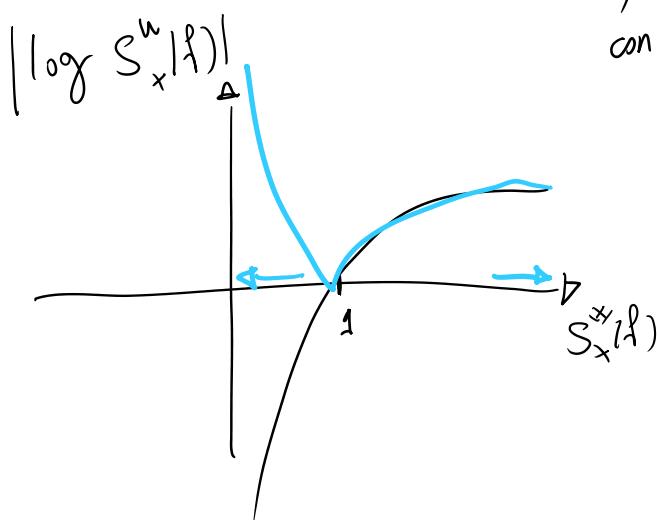
∞

by

$$\int_{-\frac{1}{2P}}^{\frac{1}{2P}} |\log S_x^w(f)| df < +\infty$$



Paley-Wiener
condition



* generate a regular process by differencing white noise

w_t
White
Noise



regular

Regular process
&
predictable process
are enough to
Generate all the
Stationary
processes

Wold decomposition theorem

f

x_k

is discrete time and stationary

then

$$1) \boxed{x_k = x_k^P + x_k^R}$$

predictable regular

$$2) \boxed{x_k^P \quad x_k^R} \text{ are orthogonal} \quad \cancel{\nabla_{j,k}} \quad E[(x_j^P)^* x_k^R] = 0$$

Consequence on power spectra

$$S_x(f) = \mathcal{F}[C_x](f)$$

$$C_x(\tau) = E[x_0^* x_\tau] \quad \underline{x_k = x_k^P + x_k^R}$$

$$= E[(x_0^P + x_0^R)^* (x_\tau^P + x_\tau^R)]$$

$$= E[(x_0^P)^* x_\tau^P] + E[(x_0^R)^* x_\tau^R] +$$

$$\cancel{E[(x_0^P)^* x_\tau^R]} + \cancel{E[(x_0^R)^* x_\tau^P]}$$

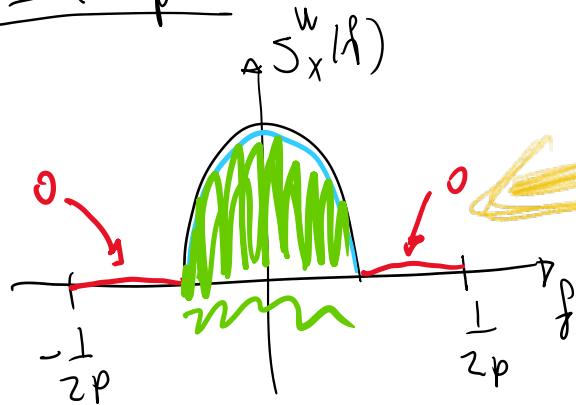
$$C_x(\tau) = C_{xP}(\tau) + C_{xR}(\tau)$$

$$\mathcal{F}[\quad] = \mathcal{F}[\quad] + \mathcal{F}[\quad]$$

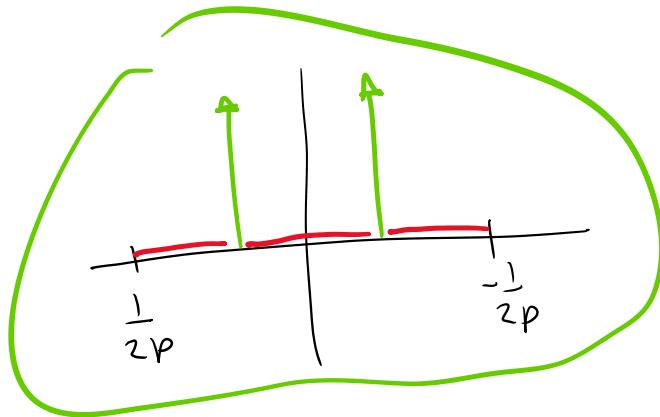
$$\underline{S_x(f) = S_{xP}(f) + S_{xR}(f)}$$

$$S_x^w(f) = S_{xp}^w(f) + S_{xr}^w(f)$$

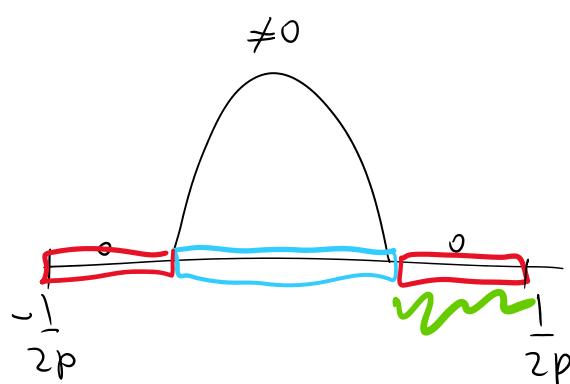
Examples



not a spectrum of
a stationary process

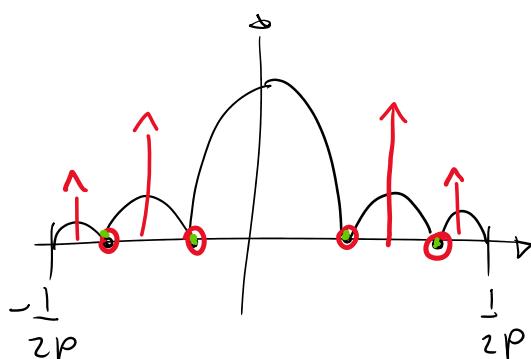


good spectrum of
a stationary process



$$\int_{-\frac{1}{2P}}^{\frac{1}{2P}} |\log S_x^w(f)| df = +\infty$$

||
0
+∞



$$\int_{-\frac{1}{2P}}^{\frac{1}{2P}} |\log S_x^w(f)| df < \infty$$

