Q42***: AR Model

Part I

Consider the following model:

$$A \sim \mathcal{N}(A; 0, 1.2)$$

 $R \sim \mathcal{IG}(R; 0.4, 250)$
 $x_k | x_{k-1}, A, R \sim \mathcal{N}(x_k; Ax_{k-1}, R)$
 $x_0 = 1$ $x_1 = -6$

- 1. Draw the directed graphical model and the factor graph
- 2. Write the expression for the full joint distribution and assign terms to the individual factors on the factor graph
- 3. Derive the full conditional distributions $p(A|R, x_0, x_1)$ and $p(R|A, x_0, x_1)$
- 4. Derive the joint distribution $p(A, R, x_0 = 1, x_1 = -6)$ and create a contour plot.

Part II

Consider the following model discussed in detail during the lectures.

$$A \sim \mathcal{N}(A; 0, P)$$
 $R \sim \mathcal{IG}(R; \nu, \nu/\beta)$
 $x_k | x_{k-1}, A, R \sim \mathcal{N}(x_k; Ax_{k-1}, R)$

where \mathcal{N} is a Gaussian and

$$\mathcal{IG}(R; a, b) = \exp\left(-(a+1)\log R - \frac{b}{R} - \log \Gamma(a) + a\log b\right)$$

Caution: (This definition is different from the definition of \mathcal{IG} given in some of the earlier lectures.) We are given the hyperparameters $\theta = (\nu, \beta, P)$

$$\nu = 0.4 \qquad \beta = 100 \qquad P = 1.2$$
 $x_0 = 1 \qquad x_1 = -6$

1. Derive and implement an EM algorithm to find the MAP estimate

$$R^* = \operatorname*{argmax}_{R} p(R|x_0, x_1, \theta)$$

2. Derive and implement an EM algorithm to find the MAP estimate

$$A^* = \operatorname*{argmax}_{A} p(A|x_0, x_1, \theta)$$

3. Derive and implement an ICM (Iterative conditional modes) algorithm to find

$$(R^*, A^*) = \underset{A,R}{\operatorname{argmax}} p(A, R|x_0, x_1, \theta)$$

4. In the lectures, we have shown that the unnormalised posterior is

$$\phi = p(A, R, x_1 = \hat{x}_1 | x_0 = \hat{x}_0, \theta) = \mathcal{N}(x_1; Ax_0, R) \mathcal{N}(A; 0, P) \mathcal{IG}(R; \nu, \nu/\beta)$$

$$\propto \exp\left(-\frac{1}{2}\frac{x_1^2}{R} + x_0 x_1 \frac{A}{R} - \frac{1}{2}\frac{x_0^2 A^2}{R} - \frac{1}{2}\log 2\pi R\right)$$

$$\exp\left(-\frac{1}{2}\frac{A^2}{P} - \frac{1}{2}\log|2\pi P|\right)$$

$$\exp\left(-(\nu + 1)\log R - \frac{\nu}{\beta}\frac{1}{R} - \log\Gamma(\nu) + \nu\log(\nu/\beta)\right)$$

We know also that the marginal log-likelihood

$$\log Z = \log p(x_1 = \hat{x}_1 | x_0 = \hat{x}_0, \theta)$$

is lower bounded by

$$\mathcal{B}_{VB} = \langle \log \phi \rangle_Q + H[Q]$$

where

$$Q = q(A)q(R)$$

$$q(A) = \mathcal{N}(A; m, \Sigma)$$

$$q(R) = \mathcal{IG}(R; a, b)$$

Extend the VB algorithm given in the slides so that you compute this bound at every iteration and plot the bound \mathcal{B} as a function of iterations. You should observe that the VB fixed point **monotonically** increases this lower bound. Restart your algorithm several times and compare the largest bound you find with the bound you find with importance sampling.