Homework 2

Solutions

1. (Geometric random variable) By definition of conditional probability, we have

$$P(\tilde{a} = a \mid \tilde{a} > 5) = \frac{P(\tilde{a} = a, \tilde{a} > a)}{P(\tilde{a} > a)}.$$
 (1)

The numerator is zero unless a > 5, in that case

$$P(\tilde{a} = a \mid \tilde{a} > 5) = \frac{(1 - \alpha)^{a-1} \alpha}{\sum_{b=6}^{\infty} (1 - \alpha)^{b-1} \alpha}.$$
 (2)

The denominator contains the geometric sum

$$\sum_{b=6}^{\infty} (1-\alpha)^{b-1} = \frac{(1-\alpha)^6}{\alpha}.$$
 (3)

We conclude

$$P(\tilde{a} = a \mid \tilde{a} > 5) = (1 - \alpha)^{a-6} \alpha.$$
 (4)

This is the same as the geometric pmf if we plug in a-5. In terms of the coin example, if we have obtained 5 tails, now the probability that we have to flip b=a-5 more times is distributed like a geometric random variable. This makes sense because the flips are independent, so there is no difference between this and just considering another sequence of flips starting from the beginning.

- 2. (Chess games)
 - (a) Under the independence assumption, we have

$$\mathcal{L}_X(\theta) = \theta^4 \alpha^2 \left(1 - \theta - \alpha \right)^4, \tag{5}$$

$$\log \mathcal{L}_X(\theta) = 4\log \theta + 2\log \alpha + 4\log (1 - \theta - \alpha). \tag{6}$$

The plot is shown in Figure 1.

(b) From the plot we can see that the function has a single maximum. To find it, we set the partial derivatives to zero. We have

$$\frac{\mathrm{d}\log\mathcal{L}_X(\theta)}{\mathrm{d}\theta} = \frac{4}{\theta} - \frac{4}{1 - \theta - \alpha},\tag{7}$$

$$\frac{\mathrm{d}\log\mathcal{L}_X(\alpha)}{\mathrm{d}\theta} = \frac{2}{\alpha} - \frac{4}{1 - \theta - \alpha}.$$
 (8)

(9)

Setting the first expression equal to zero yields $\alpha_{\rm ML} = 1 - 2\theta_{\rm ML}$. Plugging into the second and solving the equation, we conclude $\theta_{\rm ML} = 0.4$ and $\alpha_{\rm ML} = 0.2$.

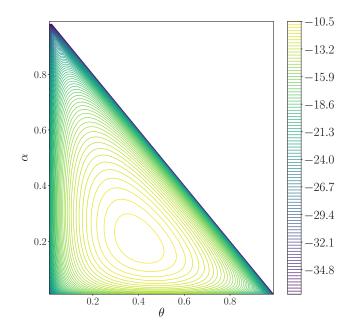


Figure 1: Log-likelihood for Problem 2.

- (c) The empirical pmf would assign 4/10 = 0.4 to the probability of Garry winning, 2/10 = 0.2 to the probability of Anish winning, and 4/10 = 0.4 to the probability of a draw. This is exactly equivalent to the parametric model.
- 3. (Darts) Let \tilde{a} denote the random variable. Note that the last attempt must always be a success. We can therefore decompose the event a attempts required into the intersection of k-1 successes over first a-1 attempts and ath attempt is a success. Since the attempts are all independent, we have

$$p_{\tilde{a}}(a) = P(k-1 \text{ successes over first } a-1 \text{ attempts})P(a\text{th attempt is a success}).$$
 (10)

By exactly the same reasoning we used to derive the binomial distribution, we have

$$P(k-1 \text{ successes over first } a-1 \text{ attempts}) = \binom{a-1}{k-1} \theta^{k-1} (1-\theta)^{a-1-(k-1)}, \qquad (11)$$

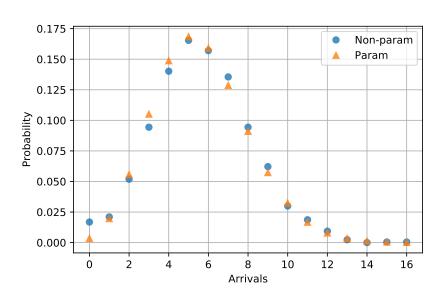
as long as $a \geq k$. We conclude

$$p_{\tilde{a}}(a) = {\binom{a-1}{k-1}} \theta^k (1-\theta)^{a-k}, \quad \text{for } a \ge k,$$
(12)

and zero otherwise.

- 4. (Air Traffic)
 - (a) We assume that flights arrive with a constant rate λ during the evening peak. We model the numbers of arrival with a Poisson distribution.

```
train_data = pd.read_csv("./train.csv", header=0, names=['time', 'arrivals']).set
test_data = pd.read_csv("./test.csv", header=0, names=['time', 'arrivals']).set_i
x_range = np.arange(train_data.max()+1)
# non-parametric model
def non_param_model(data):
    pmf = np.zeros(data.max()+1)
    counts = data.value_counts()
    emprical_prob = counts/counts.sum()
    for i in emprical_prob.index:
        pmf[i] = emprical_prob[i]
    return pmf
pmf_nonparam = non_param_model(train_data)
# parametric model
pmf_param = poisson.pmf(x_range, train_data.mean())
plt.plot(x_range, pmf_nonparam, 'o', label="Non-param", alpha=0.8)
plt.plot(x_range, pmf_param, '^', label="Param", alpha=0.8)
plt.legend()
plt.ylabel('Probability')
plt.xlabel('Arrivals')
plt.grid("on")
```



(b) The test RMSE is 0.016 for non-parametric model and 0.021 for parametric model.