**Data info:** Inflation rate from 1965 to 2021

View/Download Data: csvdata

# TIME SERIES ANALYSIS

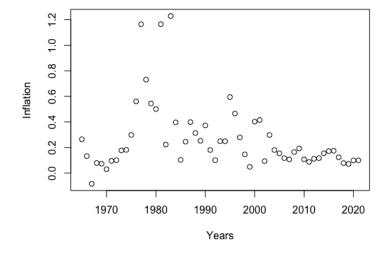
library(ggplot2) library(forecast) library(tseries) library(TSA) library(Kendall)

#### importing data

infla =read.csv("csvinflation.csv") View(infla)

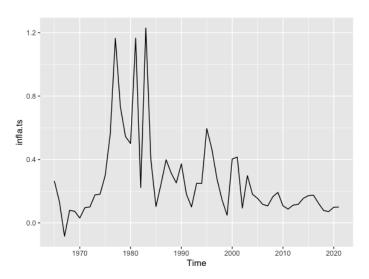
## Graphical visualization of data

plot(infla)



## converting original inflation data to time series data

infla.ts = ts(Inflation, start = c(1965,1), frequency = 1)autoplot(infla.ts)



test for trend, seasonality, normality and correlation

MannKendall(infla.ts) # Test for trend

**Output:** tau = -0.182, 2-sided pvalue = 0.045894

Conclusion: Inflation time series has a monotonic trend since p-value is less than the significance level (0.05)

shapiro.test(infla.ts) # Test for normality

#### **Output:**

Shapiro-Wilk normality test

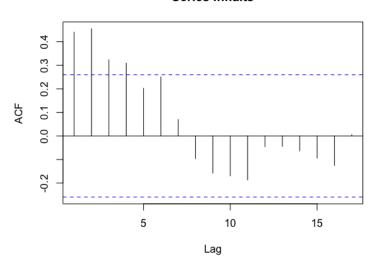
data: infla.ts

W = 0.74354, p-value = 1.243e-08

Conclusion: Inflation time series is not normally distributed since p-value is less than the significance level (0.05)

#### acf(infla.ts) # Autocorrelation Plot

#### Series infla.ts



## stationarity test

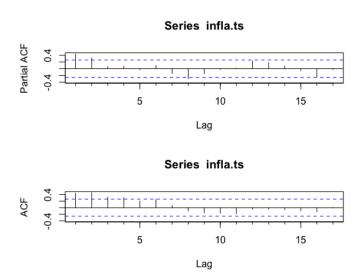
adf <- adf.test(infla.ts)
pp <- pp.test(infla.ts)
kpss <- kpss.test(infla.ts)
overall <- cbind(adf\$p.value, pp\$p.value, kpss\$p.value)
colnames(overall) <- c("ADF", "PP", "KPSS")
overall

**Output:** ADF PP KPSS [1,] 0.3554373 0.01 0.09596811

Conclusion: PP and KPSS says it's stationary since p-vaue < 0.05 and p-value > 0.05 respectively, while ADF says it is stationary.

## Acf & Pacf for model selection (Graphical)

Pacf(rain.ts) Acf(rain.ts) par(mfrow=c(2,1)) Pacf(rain.ts) Acf(rain.ts)



NB: Since we're getting different conclusions from our unit root test, we use train/test. But we first find the best or most suitable model for the two scenarios; treating it as stationary and differencing it

# model selection for UNDIFFERENCED data (automated)

m1 = auto.arima(infla.ts, stepwise = FALSE, approximation = FALSE, stationary = TRUE) m1

Output: Series: infla.ts

#### ARIMA(2,0,0) with non-zero mean

#### Coefficients:

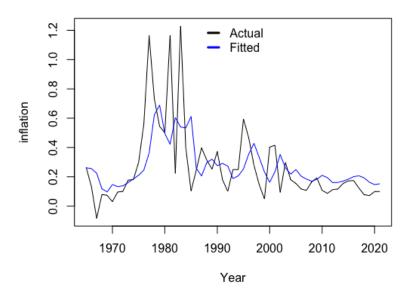
ar1 ar2 mean 0.2937 0.3211 0.2612 s.e. 0.1231 0.1234 0.0749

sigma^2 = 0.05414: log likelihood = 3.56 AIC=0.88 AICc=1.65 BIC=9.05

## Graphical comparison of Actual and Selected model

plot(infla.ts, main="ACTUAL VS FITTED", xlab = "Year", ylab="inflafall Amount", lwd=1) lines(fitted(m1), col="blue", lwd=1) legend("top", legend = c("Actual", "Fitted"), col = c("black", "blue"), lwd = 3,seg.len = 1, bty = "n", xpd = TRUE, inset = c(-0.1, 0))

#### **ACTUAL VS FITTED**



# Diagnostic of the selected model

checkresiduals(m1)

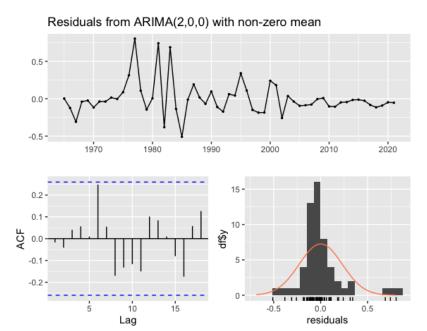
## Output:

Ljung-Box test

data: Residuals from ARIMA(2,0,0) with non-zero mean  $Q^*=8.8877,\,df=8,\,p\text{-value}=0.3519$ 

Model df: 2. Total lags used: 10

Conclusion: We conclude that Arima(2,0,0) is a suitable model since p-value > 0.05, thus we fail to reject the null hypothesis.



Mean of the residual of Arima(2,0,0)

resid1 = m1\$residuals mean(resid1)

Output: [1] 0.0007906026

Conclusion: Arima(2,0,0) is a suitable model since the mean of the residuals is very close to 0. Thus, it agrees with the Ljung-Box test conclusion.

## model selection for **DIFFERENCED** data

m2 = auto.arima(infla.ts, stepwise = FALSE, approximation = FALSE, stationary = FALSE) m2

Output:
Series: infla.ts
ARIMA(0,1,1)

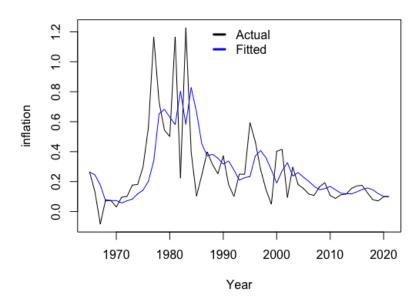
Coefficients: ma1 -0.6192 s.e. 0.1034

 $sigma^2 = 0.05658$ : log likelihood = 1.22AIC=1.56 AICc=1.78 BIC=5.61

Graphical comparison of Actual and Selected model plot(infla.ts, main="ACTUAL VS FITTED", xlab = "Year", ylab="inflafall Amount", lwd=1) lines(fitted(m2), col="blue", lwd=1)

legend("top", legend = c("Actual", "Fitted"), col = c("black", "blue"), lwd = 3, seg.len = 1, bty = "n", xpd = TRUE, inset = c(-0.1, 0))

#### **ACTUAL VS FITTED**



## Diagnostic of the selected model

checkresiduals(m2)

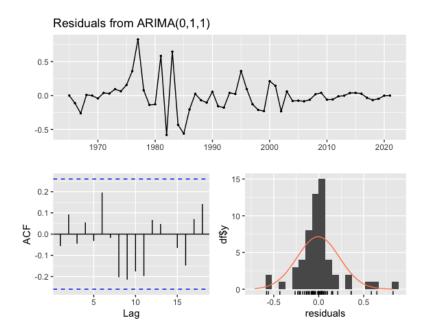
#### **Output:**

Ljung-Box test

data: Residuals from ARIMA(0,1,1) Q\* = 11.903, df = 9, p-value = 0.2188

Model df: 1. Total lags used: 10

**Conclusion:** We conclude that Arima(0,1,1) is a suitable model since p-value > 0.05, thus we fail to reject the null hypothesis.



Mean of the residual of Arima(0,1,1)

resid2 = m1\$residuals mean(resid2)

Output: [1] -0.005192656

**Conclusion:** Arima(0,1,1) is a suitable model since the mean of the residuals is very close to 0. Thus, it agrees with the Ljung-Box test conclusion.

## Using Train/Test to determine the best model between Arima(2,0,0) and Arima(0,1,1)

#### Train Set

mtrain = window(infla.ts, start = 1965, end = 2011, frequency = 1) mtrain

**Output:** 

Time Series: Start = 1965 End = 2011 Frequency = 1

 $[1] \quad 0.2644 \quad 0.1324 \quad -0.0842 \quad 0.0789 \quad 0.0732 \quad 0.0303 \quad 0.0956 \quad 0.1007 \quad 0.1768 \quad 0.1813 \quad 0.2982 \quad 0.5608 \quad 1.1645 \quad 0.0813 \quad 0.0982 \quad 0.0$ 

 $[14] \ 0.7309 \ 0.5444 \ 0.5007 \ 1.1650 \ 0.2230 \ 1.2287 \ 0.3967 \ 0.1031 \ 0.2457 \ 0.3982 \ 0.3136 \ 0.2522 \ 0.3726$ 

 $[27] \ \ 0.1803 \ \ 0.1006 \ \ 0.2496 \ \ 0.2487 \ \ 0.5946 \ \ 0.4656 \ \ 0.2789 \ \ 0.1462 \ \ 0.0487 \ \ 0.4024 \ \ 0.4151 \ \ 0.0936 \ \ 0.2977$ 

 $[40] \ \ 0.1804 \ \ 0.1544 \ \ 0.1168 \ \ 0.1073 \ \ 0.1649 \ \ 0.1925 \ \ 0.1073 \ \ 0.0873$ 

#### Test Set

mtest = window(infla.ts, start = 2012, frequency = 1) mtest

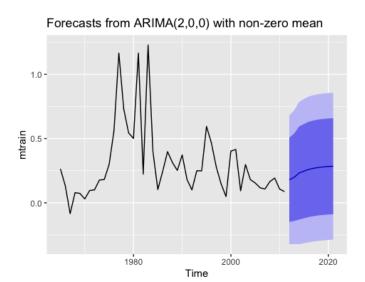
**Output:** 

Time Series: Start = 2012 End = 2021 Frequency = 1

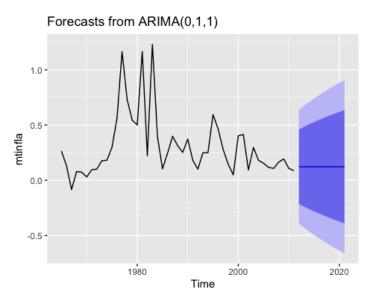
 $[1]\ 0.1119\ 0.1167\ 0.1549\ 0.1715\ 0.1745\ 0.1237\ 0.0781\ 0.0714\ 0.0989\ 0.0997$ 

tm1 = Arima(mtrain, order = c(2,0,0))

fm1 = forecast(tm1, h=length(mtest))
fm1
autoplot(fm1)



tm2 = Arima(mtrain, order = c(0,1,1)) fm2 = forecast(tm2, h=length(mtest)) autoplot(fm2)



Selection of best model by comparing accuracy output for the train/test

accuracy(fm1, mtest) accuracy(fm2, mtest)

### Output(tabular):

Model	RMSE	MAE	ME
Arima(2,0,0)	0.1418473	0.1305226	-0.13052265
Arima(0,1,1)	0.03460538	0.02934484	-0.00264422

Conclusion: We therefore conclude that Arima(0,1,1) is the better or more suitable model for our inflation data.

©Bentum Welson