

HOMEWORK 11

[Wei] 4.1. Notice that $Q(\mathbf{X}, \mu) = \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \sim N(0, 1)$ is a pivot. Then a 0.95 confidence interval is $\{-z_{0.05/2} \leq Q \leq z_{0.05/2}\}$. With $\bar{X} = 4.7832$, $\sigma = 0.01$, $n = 5$, the confidence interval is

$$[\bar{X} - \frac{\sigma}{\sqrt{n}}z_{0.025}, \bar{X} + \frac{\sigma}{\sqrt{n}}z_{0.025}] = [4.774, 4.792].$$

□

[Wei] 4.2. Notice that $Q(\mathbf{X}, \mu) = \frac{\sqrt{n}(\bar{X} - \mu)}{S} \sim t_{n-1}$ is a pivot. Then a $1 - \alpha$ confidence interval is $\{t_{n-1}(\alpha/2) \leq Q \leq t_{n-1}(\alpha/2)\}$, that is,

$$[\bar{X} - \frac{S}{\sqrt{n}}t_{n-1}(\alpha/2), \bar{X} + \frac{S}{\sqrt{n}}t_{n-1}(\alpha/2)].$$

With $\bar{X} = 4.7832$, $S = 0.0105$, $n = 5$, the confidence interval is $[4.770, 4.796]$ for $\alpha = 0.05$ and $[4.762, 4.805]$ for $\alpha = 0.01$. □

[Wei] 4.3. Similar to Exercise 4.2, a 0.05 confidence interval is $[\bar{X} - \frac{S}{\sqrt{n}}t_{n-1}(\alpha/2), \bar{X} + \frac{S}{\sqrt{n}}t_{n-1}(\alpha/2)] = [1784, 2116]$. □

[Wei] 4.4. Here we have

$$\frac{\sqrt{n}(\bar{X} - \mu)}{4} \sim N(0, 1).$$

and we may transform the probability as

$$P(\bar{X} - 1 < \mu < \bar{X} + 1) = P(-\frac{\sqrt{n}}{4} < Z < \frac{\sqrt{n}}{4}) \geq 0.9.$$

where $Z \sim N(0, 1)$. We also note that $z_{\alpha/2} = 1.65$, which implies that $\frac{\sqrt{n}}{4} \geq 1.65$, thus we have $n \geq 44$.

[Wei] 4.11. Notice that $Q = \frac{\bar{A} - \bar{B} - (a_1 - a_2)}{S} \sqrt{\frac{mn}{m+n}} \sim t_{m+n-2}$ is a pivot, where $S^2 = \frac{(m-1)S_A^2 + (n-1)S_B^2}{m+n-2}$ (See Page 137 for more details). Then a $1 - \alpha$ confidence interval is

$$[\bar{A} - \bar{B} - S\sqrt{\frac{m+n}{mn}}t_{n-1}(\alpha/2), \bar{A} - \bar{B} + S\sqrt{\frac{m+n}{mn}}t_{n-1}(\alpha/2)].$$

With $\bar{A} = 0.14125$, $\bar{B} = 0.1392$, $S = 0.00255$, $m = 4$, $n = 5$, the 0.05 confidence interval is $[-0.002, 0.006]$ for $a_1 - a_2$. □

[Wei] 4.12. From Page 138, $Q = \frac{\bar{Y} - \bar{X} - (b - a)}{\sqrt{S_1^2/m + S_2^2/n}} \xrightarrow{\mathcal{L}} N(0, 1)$ is a pivot. Then an asymptotic $1 - \alpha$ confidence interval for $b - a$ is

$$[\bar{Y} - \bar{X} - \sqrt{S_1^2/m + S_2^2/n}z_{\alpha/2}, \bar{Y} - \bar{X} + \sqrt{S_1^2/m + S_2^2/n}z_{\alpha/2}].$$

With the given data, the confidence interval is $[-155, -95]$. □

[Wei] 4.14. Facts:

- (1) $\frac{\bar{Y} - \bar{X} - (b-a)}{\sqrt{\sigma_1^2/m + \sigma_2^2/n}} \sim N(0, 1).$
- (2) $\frac{(m-1)S_1^2}{\sigma_1^2} + \frac{(n-1)S_2^2}{\sigma_2^2} \sim \chi_{m+n-2}^2.$

Since $\sigma_2^2 = \lambda\sigma_1^2$, $Q = \frac{\frac{\bar{Y} - \bar{X} - (b-a)}{\sqrt{\sigma_1^2/m + \sigma_2^2/n}}}{\sqrt{\frac{1}{m+n-2} \left(\frac{(m-1)S_1^2}{\sigma_1^2} + \frac{(n-1)S_2^2}{\sigma_2^2} \right)}} = \frac{\frac{\bar{Y} - \bar{X} - (b-a)}{\sqrt{1/m + \lambda/n}}}{\sqrt{\frac{1}{m+n-2} \left((m-1)S_1^2 + (n-1)S_2^2/\lambda \right)}} \sim t_{m+n-2}$

is a pivot. Then a $1 - \alpha$ confidence interval for $b - a$ is

$$\bar{Y} - \bar{X} \pm \sqrt{\frac{(1/m + \lambda/n)((m-1)S_1^2 + (n-1)S_2^2/\lambda)}{(m+n-2)}} t_{m+n-2}(\alpha/2).$$

(I abbreviate the interval expression since it's too long.)

[Wei] 4.10. The problem should be revised as “how many should n be at most?” Notice that $\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$, and from the definition of lower confidence limit,

$$0.95 \leq P\left(\sigma \geq \frac{\sqrt{(n-1)S^2}}{4}\right) = P(\chi_{n-1}^2 \leq 16).$$

Thus we have $n \leq 9$. □

[Wei] 4.15.

- (1) According to [Wei] Page 138, in the case that m and n are not sufficiently large, an approximate $1 - \alpha$ confidence interval for $\mu_2 - \mu_1$ is

$$[\bar{Y} - \bar{X} - S_* t_r(\alpha/2), \bar{Y} - \bar{X} + S_* t_r(\alpha/2)],$$

where $S_*^2 = \frac{S_1^2}{m} + \frac{S_2^2}{n} = 177.73$, $r = S_*^4 / \left[\frac{S_1^4}{m^2(m-1)} + \frac{S_2^4}{n^2(n-1)} \right] = 7.77 \approx 8$. Therefore, the confidence interval should be $[-39.72, 21.76]$. □

- (2) According to [Wei] Page 139 (4.2.13), a $1 - \alpha$ confidence interval for σ_1^2/σ_2^2 is

$$\left[\frac{S_1^2}{S_2^2} \frac{1}{F_{m-1, n-1}(\alpha/2)}, \frac{S_1^2}{S_2^2} \frac{1}{F_{m-1, n-1}(1 - \alpha/2)} \right].$$

With the given data, it should be $[4.88, 79.36]$. □

[Wei] 4.19. We only consider the case $n > 1$.

- (1) Let $Y_i = \frac{X_i - \theta_1}{\theta_2 - \theta_1} \sim U(0, 1)$, $Z_i = \frac{\theta_2 - X_i}{\theta_2 - \theta_1} \sim U(0, 1)$, $Q = \frac{(\theta_2 - \theta_1) - (X_{(n)} - X_{(1)})}{\theta_2 - \theta_1}$ then $Q = Y_{(1)} + Z_{(1)}$ and

$$f_{Y_{(1)}, Z_{(1)}}(y, z) = f_{Y_{(1)}, Y_{(n)}}(y, 1 - z) = n(n-1)(1 - z - y)^{n-2} I_{(0,1)}(y, z, y + z).$$

Therefore, Q is a pivot with *p.d.f.*

$$f_Q(x) = \int f_{Y_{(1)}, Z_{(1)}}(y, x - y) dy = n(n-1)x(1-x)^{n-2} I_{(0,1)}(x).$$

We can observe that $Q \sim \text{Beta}(2, n-1)$.

Suppose $P(Q \in (1-c, 1)) = 1-\alpha$ (this assumption is convenient for following analysis, you can generally suppose a confidence interval (a, b) instead), then $c = \text{Beta}_{n-1,2}(\alpha)$. Also notice that $Q = 1 - \frac{(X_{(n)} - X_{(1)})}{\theta_2 - \theta_1}$,

$$Q \in (1-c, 1) \Leftrightarrow \theta_2 - \theta_1 \in ((X_{(n)} - X_{(1)})/c, \infty).$$

As a result, a $1-\alpha$ confidence interval for $\theta_2 - \theta_1$ is

$$((X_{(n)} - X_{(1)})/\text{Beta}_{n-1,2}(\alpha), \infty).$$

□

- (2) Hint: the objective statistic can be rewritten as $R := \frac{Y_{(1)} + Y_{(n)} - 1}{Y_{(n)} - Y_{(1)}}$, which is auxiliary. Notice that $(X_{(1)}, X_{(n)})$ is sufficient and complete for (θ_1, θ_2) . Basu's theorem tells us $S := Y_{(n)} - Y_{(1)}$ is independent with R . Now write the *p.d.f* of $(Y_{(1)}, Y_{(n)})$. With the relationship

$$\begin{cases} Y_{(1)} = \frac{1}{2}(RS - S + 1) \\ Y_{(n)} = \frac{1}{2}(RS + S + 1) \end{cases},$$

write the *p.d.f* of (R, S) . Consequently, $f_R = f_{(R,S)}/f_S$. The explicit form is very complicated, and should be discussed with nodes 0, ±1. □

- (3) Using R as the pivot. □

[Wei] 4.23.

- (1) $X_{(1)} - \theta \sim \text{Exp}(n)$. □
 (2) Using $Q := X_{(1)} - \theta$ in (1) as a pivot, suppose the interval as $(0, c)$ for Q . After calculation, the interval for θ is $(X_{(1)} + \log \alpha/n, X_{(1)})$. □

[Wei] 5.56, 5.57. See [Wei] Page 135 (4.2.4), Page 141 (4.2.13) for the answer. But take exercise on using inverting hypothesis method!

	5.56	5.57
CI	$\left[\frac{(n-1)S^2}{\chi_{n-1}^2(\alpha/2)}, \frac{(n-1)S^2}{\chi_{n-1}^2(1-\alpha/2)} \right]$	$\left[\frac{S_1^2}{S_2^2} \frac{1}{F_{m-1, n-1}(\alpha/2)}, \frac{S_1^2}{S_2^2} \frac{1}{F_{m-1, n-1}(1-\alpha/2)} \right]$

□

CL	$\frac{(n-1)S^2}{\chi_{n-1}^2(\alpha)}$	$\frac{S_1^2}{S_2^2} \frac{1}{F_{m-1, n-1}(\alpha)}$
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CU	$\frac{(n-1)S^2}{\chi_{n-1}^2(1-\alpha)}$	$\frac{S_1^2}{S_2^2} F_{n-1, m-1}(\alpha)$
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