Logistic Regression and MaxEnt

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Generative vs. Discriminative Learning

• Generative models:

$$Pr[y \mid x] = \frac{Pr[x \mid y]Pr[y]}{Pr[x]}$$

$$\propto Pr[x \mid y]Pr[y] = Pr[x, y]$$

- The key is to model the generative probability: Pr[x | y].
- Example: Naive Bayes.
- Discriminative models:
 - models $Pr[y \mid x]$ directly as $g(x; \theta)$.
 - Example: Decision tree, Logistic Regression.
- Instance-based Learning.
 - Example: kNN classifier.

Linear Regression

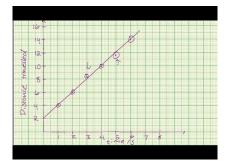


Figure: Linear Regression

Task

- Input: $(x^{(i)}, y^{(i)})$ pairs $(1 \le i \le n)$
- Preprocess: let $\mathbf{x}^{(i)} = \begin{bmatrix} 1 & x^{(i)} \end{bmatrix}^{\top}$
- Output: The best $\mathbf{w} = \begin{bmatrix} w_0 & w_1 \end{bmatrix}^\top$ such that $\hat{y} = \mathbf{w}^\top \mathbf{x}$ best explains the observations

Least Square Fit

The criterion for "best":

- Individual error: $\epsilon_i = \hat{y}^{(i)} y^{(i)}$
- Sum squared error: $\ell = \sum_{i=1}^n \epsilon_i^2$

Find **w** such that ℓ is minimized.

Minimizing a Function

Taylor Series of f(x) at point a

$$f(x) = \sum_{n=0}^{+\infty} \frac{f^{(i)}(a)}{n!} (x - a)^n$$
 (1)

$$= f(a) + f'(a) \cdot (x-a) + \frac{f''(a)}{2}(x-a)^2 + o((x-a)^2) \quad (2)$$

- Intuitively, f(x) is almost $f(a) + f'(a) \cdot (x a)$ for all a if it is close to x.
- If f(x) has local minimum x^* , then
 - $f'(x^*) = 0$, and
 - $f''(x^*) > 0$.

Minimum of the local minima is the global minimum if it is smaller than the function values at all the boundary points.

• Intuitively, f(x) is almost $f(a) + \frac{f''(a)}{2}(x-a)^2$ if a is close to x^* .

Find the Least Square Fit for Linear Regression

$$\frac{\partial \ell}{\partial w_j} = \sum_{i=1}^n 2\epsilon_i \frac{\partial \epsilon_i}{\partial w_j} = \sum_{i=1}^n 2\epsilon_i \frac{\partial \mathbf{w}^\top \mathbf{x}^{(i)}}{\partial w_j}$$
$$= \sum_{i=1}^n 2\epsilon_i x_j^{(i)} = 2\sum_{i=1}^n (\hat{y}^{(i)} - y^{(i)}) x_j^{(i)}$$

By setting the above to 0, this essentially requires, for all j

$$\sum_{i=1}^{n} \hat{y}^{(i)} x_{j}^{(i)} = \sum_{i=1}^{n} y^{(i)} x_{j}^{(i)}$$

what the model predicts

what the data says

Find the Least Square Fit for Linear Regression

In the simple 1D case, we have only two parameters in $\mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$

$$\sum_{i=1}^{n} (w_0 + w_1 x_1^{(i)}) x_0^{(i)} = \sum_{i=1}^{n} y^{(i)} x_0^{(i)}$$
$$\sum_{i=1}^{n} (w_0 + w_1 x_1^{(i)}) x_1^{(i)} = \sum_{i=1}^{n} y^{(i)} x_1^{(i)}$$

Since $x_0^{(i)} = 1$, they are essentially

$$\sum_{i=1}^{n} (w_0 + w_1 x_1^{(i)}) \cdot 1 = \sum_{i=1}^{n} y^{(i)} \cdot 1$$
$$\sum_{i=1}^{n} (w_0 + w_1 x_1^{(i)}) \cdot x_1^{(i)} = \sum_{i=1}^{n} y^{(i)} \cdot x_1^{(i)}$$

Example

Using the same example in https://en.wikipedia.org/wiki/Linear_least_squares_(mathematics)

$$\mathbf{X} = \begin{bmatrix} - & (x^{(1)})^{\top} & - \\ - & (x^{(2)})^{\top} & - \\ - & (x^{(3)})^{\top} & - \\ - & (x^{(4)})^{\top} & - \end{bmatrix} = \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \end{bmatrix} \quad \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix} \quad \mathbf{y} = \begin{bmatrix} 6 \\ 5 \\ 7 \\ 10 \end{bmatrix}$$

$$\begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \end{bmatrix} \qquad \begin{bmatrix} 6 \\ 5 \\ 7 \\ 10 \end{bmatrix} \qquad = \qquad \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \end{bmatrix} \qquad \begin{bmatrix} \hat{y}_1 \\ \hat{y}_2 \\ \hat{y}_3 \\ \hat{y}_4 \end{bmatrix}$$

Generalization to *m*-dim

• Easily generalizes to more than 2-dim:

$$\mathbf{X} = \begin{bmatrix} 1 & x_1^{(1)} & \dots & x_m^{(1)} \\ 1 & \dots & \dots & \dots \\ 1 & x_1^{(i)} & \dots & x_m^{(i)} \\ 1 & \dots & \dots & \dots \\ 1 & x_1^{(n)} & \dots & x_m^{(n)} \end{bmatrix} \quad \mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \dots \\ w_m \end{bmatrix} \quad \mathbf{y} = \begin{bmatrix} y^{(1)} \\ \dots \\ y^{(i)} \\ \dots \\ y^{(n)} \end{bmatrix}$$

- How to perform polynomial regression for one dimensional x?
 - $\hat{y} = w_0 + w_1 x + w_2 x^2 \ldots + w_m x^m.$
 - Let $x_j^{(i)} = (x_1^{(i)})^j \Longrightarrow \text{Polynomial least square fitting } (\text{http://mathworld.wolfram.com/} \\ \text{LeastSquaresFittingPolynomial.html})$

Probablistic Interpretation

High-level idea:

Any w is possible, but some w is most likely.

•
$$P(y^{(i)} | \hat{y}^{(i)}) = f_i(\mathbf{w})$$

- Assuming independence of training examples, the likelihood of the training dataset is $\prod_i f_i(\mathbf{w})$.
- We shall choose the **w*** that maximizes the likelihood.
 - Maximum likelihood estimation (MLE)
 - If we also incorporate some prior on w, this becomes Maximum Posterior Estimation (MAP)
 - If we assume some Gaussian prior on ${\bf w}$, this will add a ℓ_2 regularization term to the objective function.
- Many models and their variants can be deemed as different ways of estimating $P(y^{(i)} | \hat{y}^{(i)})$



Geometric Interpretation and the Closed Form Solution

Find **w** such that $\|\mathbf{y} - \mathbf{X}\mathbf{w}\|_2$ is minimized.

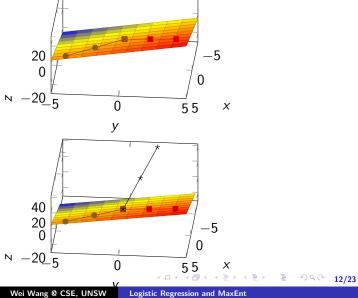
- What is Xw when X is fixed?
 - It is the hyperplane spanned by the d column vectors of \mathbf{X} .
- **y** in general is a vector outside the hyperplane. So the minimum distance is achieved when $\mathbf{X}\mathbf{w}^*$ is exactly the projection of \mathbf{y} on the hyperplane. This means (denote *i*-th column of \mathbf{X} as X_i)

$$\begin{array}{ll} X_1^\top(\mathbf{y} - \mathbf{X}\mathbf{w}) &= 0 \\ X_2^\top(\mathbf{y} - \mathbf{X}\mathbf{w}) &= 0 \\ \dots \dots &= 0 \\ X_d^\top(\mathbf{y} - \mathbf{X}\mathbf{w}) &= 0 \end{array} \right\} \Longrightarrow \mathbf{X}^\top(\mathbf{y} - \mathbf{X}\mathbf{w}) = \mathbf{0}$$

$$\bullet \ \mathbf{w} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y} = \mathbf{X}^{+}\mathbf{y}$$

(X⁺: pseudo inverse of X)

Illustration



Logistic Regression

Special case: $y^{(i)} \in \{0, 1\}.$

- Not appropriate to directly regress $y^{(i)}$.
- Rather, model $y^{(i)}$ as the observed outcome of a Bernoulli trial with an unknown parameter p_i
- How to model p_i
 - We assume that p_i depends on $\mathbf{x} \triangleq \mathbf{X}_{i\bullet} \Longrightarrow$ rename p_i to $p_{\mathbf{x}}$.
 - Still hard to estimate p_x reliably.
 - MLE: $p_x = E[y = 1 \mid x]$
 - What can we say about $p_{x+\epsilon}$ when p_x is given?
- ullet Answer: we impose a linear relationship between $p_{f x}$ and ${f x}$
 - What about a simple linear model $p_{\mathbf{x}} = \mathbf{w}^{\top} \mathbf{x}$ for some \mathbf{w} ? (Note: all points share the same parameter \mathbf{w})
 - Problem: mismatch of the domains: vs
 - Solution: mean function / inverse of link function: $g^{-1}: \Re \to \mathrm{params}$

Solution

• Solution: Link function $g(parameters) \rightarrow \Re$

$$g(p) = \operatorname{logit}(p) \triangleq \log \frac{p}{1-p} = \mathbf{w}^{\mathsf{T}} \mathbf{x}$$
 (3)

Equivalently, solve for p.

$$p = \frac{e^{\mathbf{w}^{\top}\mathbf{x}}}{1 + e^{\mathbf{w}^{\top}\mathbf{x}}} = \frac{1}{1 + e^{-\mathbf{w}^{\top}\mathbf{x}}} = \sigma(\mathbf{w}^{\top}\mathbf{x})$$
(4)

Where $\sigma(z) = \frac{1}{1 + \exp(-z)}$.

Recall that $p_{\mathbf{x}} = \mathbf{E}[y = 1 \mid \mathbf{x}].$

- Decision boundary is $p \ge 0.5$.
 - Equivalent to whether $\mathbf{w}^{\top}\mathbf{x} \geq 0$. Hence, LR is a linear classifier.



Learning the Parameter w

- Consider a training data point $\mathbf{x}^{(i)}$.
 - Recall that the conditional probability $(\Pr[y^{(i)} = 1 \mid \mathbf{x^{(i)}}])$ computed by the model is denoted by the shorthand notation p (which is a function of \mathbf{w} and $\mathbf{x^{(i)}}$).
 - The likelihood of $\mathbf{x^{(i)}}$ is $\begin{cases} p & \text{, if } y^{(i)} = 1 \\ 1-p & \text{, otherwise} \end{cases}$, or equivalently, $p^{y^{(i)}}(1-p)^{1-y^{(i)}}$.
- Hence, the likelihood of the whole training dataset is

$$L(\mathbf{w}) = \prod_{i=1}^{n} p(\mathbf{x}^{(i)})^{y^{(i)}} (1 - p(\mathbf{x}^{(i)}))^{1 - y^{(i)}}.$$

• Log-likelihood is (assume $\log \triangleq \ln$)

$$\ell(\mathbf{w}) = \sum_{i=1}^{n} y^{(i)} \log p(\mathbf{x}^{(i)}) + (1 - y^{(i)}) \log (1 - p(\mathbf{x}^{(i)}))$$
 (5)

Learning the Parameter w

ullet To maximize ℓ , notice that it is concave. So take its partial derivatives

$$\frac{\partial \ell(\mathbf{w})}{\partial \mathbf{w}_j} = \sum_{i=1}^n \left(y^{(i)} \frac{1}{p(\mathbf{x}^{(i)})} \frac{\partial p(\mathbf{x}^{(i)})}{\partial \mathbf{w}_j} + (1 - y^{(i)}) \frac{1}{1 - p(\mathbf{x}^{(i)})} \frac{\partial (1 - p(\mathbf{x}^{(i)}))}{\partial \mathbf{w}_j} \right)$$

$$= \sum_{i=1}^n \left(\mathbf{x}^{(i)}_j y^{(i)} - \mathbf{x}^{(i)}_j p(\mathbf{x}^{(i)}) \right)$$

ullet and set them to 0 essentially means, for all j

$$\sum_{i=1}^{n} \hat{y}^{(i)} \cdot \mathbf{x^{(i)}}_{j} = \sum_{i=1}^{n} p(\mathbf{x^{(i)}}) \mathbf{x^{(i)}}_{j} = \sum_{i=1}^{n} y^{(i)} \cdot \mathbf{x^{(i)}}_{j}$$

what the model predicts

what the data says

Understand the Equilibrium

 Consider one dimensional x. The above condition is simplified to

$$\sum_{i=1}^{n} p^{(i)} x^{(i)} = \sum_{i=1}^{n} y^{(i)} x^{(i)}$$

- The RHS is essentially the sum of x values **only** for the training data in class Y = 1.
- The LHS says: if we use our learned model to assign a probability (of belonging to the class Y=1) for **every** training data, the LHS is the expected sum of x values.
- If this is still abstract, think of an example.

Numeric Solution

- There is no closed-form solution to maximize ℓ .
- Use the *Gradient Ascent* algorithm to maximize ℓ .
- There are faster algorithms.

(Stochastic) Gradient Ascent

- w is intialized to some random value (e.g., 0).
- Since the gradient gives the steepest direction to increase a function's value, we move a small step towards that direction, i.e.,

$$w_j \leftarrow w_j + \alpha \frac{\partial \ell(\mathbf{w})}{\partial \mathbf{w}_j}$$
, or $w_j \leftarrow w_j + \alpha \sum_{i=1}^n (y^{(i)} - p(\mathbf{x^{(i)}})) \mathbf{x^{(i)}}_j$

where α (learning rate) is usually a small constant, or decreasing over the epochs.

 Stochastic version: using the gradient on a randomly selected training instance, i.e.,

$$w_j \leftarrow w_j + \alpha(y^{(i)} - p(\mathbf{x^{(i)}}))\mathbf{x^{(i)}}_j$$

Newton's Method

- Gradient Ascent moves to the "right" direction a tiny step a time. Can we find a good step size?
- Consider 1D case: **minimize** f(x) and the current point is a.

•
$$f(x) = f(a) + f'(a)(x - a) + \frac{f''(a)}{2}(x - a)^2$$
 for x near a.

• To minimize f(x), take $\frac{\partial f(x)}{\partial x} = 0$, i.e.,

$$\frac{\partial f(x)}{\partial x} = 0$$

$$\Leftrightarrow f'(a) \cdot 1 + \frac{f''(a)}{2} \cdot 2(x - a) \cdot 1 = f'(a) + f''(a)(x - a) = 0$$

$$\Leftrightarrow x = a - \frac{f'(a)}{f''(a)}$$

• Can be applied to multiple dimension cases too \Rightarrow need to use ∇ (gradient) and Hess (Hessian).



Regularization

- Regularization is another method to deal with overfitting.
 - It is designed to penalize large values of the model parameters.
 - Hence it encourages simpler models, which are less likely to overfit.
- Instead of optimizing for $\ell(\mathbf{w})$, we optimize $\ell(\mathbf{w}) + \lambda R(\mathbf{w})$.
 - \bullet λ is a hyper-parameter that controls the strength of regularization.
 - It is usually determined by cross validating with a list of possible values (e.g., 0.001, 0.01, 0.1, 1, 10, ...)
 - Grid search: http: //scikit-learn.org/stable/modules/grid_search.html
 There are alternative methods.
 - R(w) quantifies the "size" of the model parameters. Popular choices are:
 - L_2 regularization (Ridge LR) $R(\mathbf{w}) = \|w\|_2$
 - L_1 regularization (Lasso LR) $R(\mathbf{w}) = ||w||_1$
 - ullet L₁ regularization is more likely to result in sparse models.

Generalizing LR to Multiple Classes

■ LR can be generalized to multiple classes ⇒ MaxEnt.

$$\Pr[c \mid \mathbf{x}] \propto \exp\left(\mathbf{w}_c^{\top}\mathbf{x}\right) \implies \Pr[c \mid \mathbf{x}] = \frac{\exp\left(\mathbf{w}_c^{\top}\mathbf{x}\right)}{Z}$$

- Z is the normalization constant..
- Let \mathbf{c}^* be the last class in C, then $\mathbf{w}_{\mathbf{c}^*} = \mathbf{0}$.
- Derive LR from MaxEnt How?
- Both belong to *exponential* or *log-linear* classifiers.

Further Reading

- Andrew Ng's note: http://cs229.stanford.edu/notes/cs229-notes1.pdf
- Cosma Shalizi's note: http://www.stat.cmu.edu/ ~cshalizi/uADA/12/lectures/ch12.pdf
- Tom Mitchell's book chapter: https: //www.cs.cmu.edu/~tom/mlbook/NBayesLogReg.pdf