# Math-UA.233: Theory of Probability Lecture 14

#### Tim Austin

tim@cims.nyu.edu
cims.nyu.edu/~tim

#### From last time: continuous RVs

If *X* is a continuous RV, then:

(Definition) It has a **probability density function** ('**PDF**') *f*:

$$P(a \leqslant X \leqslant b) = \int_a^b f(x) \, dx$$
 (c.f. discrete:  $\sum_{i \text{ such that } a \leqslant x_i \leqslant b} p(x_i)$ )

(Definition) Expectation:

$$E[X] = \int_{-\infty}^{\infty} xf(x) dx$$
 (c.f. discrete:  $\sum_{i} x_{i}p(x_{i})$ )

(Theorem: LOTUS)

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x) dx$$
 (c.f. discrete:  $\sum_{i} g(x_{i})p(x_{i})$ )

A special case of LOTUS:

$$E[aX + b] = aE[X] + b$$

for any continuous RV X and real numbers a and b.

Next we turn to the variance.

#### Definition (Ross p183)

If X is a continuous RV with PDF f, and if we let  $\mu = E[X]$ , then its **variance** is the number

$$Var(X) = E[(X - \mu)^2].$$

(Identical to what we wrote for discrete RVs)

Letting  $g(x) = (x - \mu)^2$  and applying LOTUS, we get

$$Var(X) = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) \, dx,$$

whenever this integral is well-defined.

Just as for discrete RVs, we can now deduce the alternative formula

$$Var(X) = E[X^2] - (E[X])^2,$$

which is sometimes simpler to use.

#### Example (Ross 5.2e)

Find Var(X) when the PDF of X is

$$f(x) = \begin{cases} 2x & \text{if } 0 \leqslant x \leqslant 1 \\ 0 & \text{otherwise.} \end{cases}$$

Example (Ross E.g. 5.3a, part (b))

Find Var(X) when X is Unif(c, d).

Sometimes we need to know how variance changes when we apply a linear function to a RV.

#### Proposition (Ross p183)

For any constants a and b,

$$Var(aX + b) = a^2 Var(X).$$

## Finding the <u>distribution</u> of g(X)

We have seen a simple formula for E[g(X)]. But sometimes we need more information about g(X), such as its CDF.

That is, we want to determine the function

$$F_{g(X)}(a) = P(g(X) \leqslant a)$$

in terms of the PDF of X itself.

In general, we can do this by expressing the *event*  $\{g(X) \le a\}$  in terms of events of the form  $\{c \le X \le d\}$ , and then compute using the probabilities of those (e.g. using  $F_X$ .)

## Example (Ross E.g. 5.1d)

Suppose X is continuous with PDF  $f_X$ . Show that Y = 2X is also continuous, and find its PDF.

#### Example (Warning!)

Let X be Unif(-1,1), and let

$$g(x) = \begin{cases} 0 & \text{if } x \leq 0 \\ x & \text{if } x \geq 0. \end{cases}$$

Then g(X) is not discrete or continuous!

In the last example, g(X) is a **hybrid**: it behaves like a discrete RV on the event  $\{X < 0\}$ , and like a continuous one on the event  $\{X \geqslant 0\}$ . But there are even more complex examples which are actually neither. (Try looking up 'Cantor distribution' on Wikipedia.)



## Example (Ross E.g. 5.7a)

Let X be Unif(0,1) and let n be a positive integer. Find the CDF of  $Y = X^n$ . Is Y a continuous RV? If so, what is its PDF?

ANS: 
$$F_Y(a) = a^{1/n}$$
; Yes;  $f_Y(y) = (1/n)y^{1/n-1}$ .



The new RV g(X) can be hard to understand in general. But there are conditions on g which guarantee good properties.

#### Theorem (See Ross Theorem 5.7.1)

Suppose X is continuous with PDF  $f_X$ , and that it always takes values in the interval [a, b].

Assume g(x) is (i) strictly increasing and (ii) differentiable on [a,b].

Then Y = g(X) is also continuous,  $g(a) \leqslant Y \leqslant g(b)$ , and

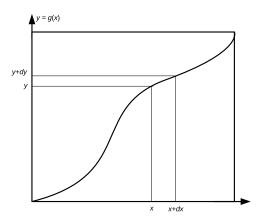
$$f_Y(y) = f_X(g^{-1}(y)) \cdot (g^{-1})'(y) \quad \text{for } g(a) \leqslant y \leqslant g(b)$$

This is all still OK if  $a = -\infty$  or  $b = +\infty$ .

IDEA: Find  $F_Y$  in terms of  $F_X$ , then differentiate to find  $f_Y$ .



#### Accompanying picture:



We have  $x = g^{-1}(y)$ , so  $dx = (g^{-1})'(y)dy$ , and

$$f_Y(y)dy \approx P(Y \in [y, y + dy]) = P(X \in [x, x + dx]) \approx f_X(x)dx.$$

If you want to apply the previous theorem, it's very important to *check the conditions*.

#### Example

Let  $n \ge 1$  be an integer and let X be continuous and non-negative, with PDF f. Let  $Y = X^n$ . Find  $f_Y$ .

NOTE: If n is even, then the function  $g(x) = x^n$  isn't increasing (or decreasing) on the whole real line, but it is increasing on  $[0,\infty)$ , which is where X takes its values.

Something we did last time...

Example (Ross E.g. 5.2b)

Let X be Unif(0,1). Find  $E[e^X]$ .

If the conditions aren't met, then that theorem can fail!

#### Example (Ross E.g. 5.7c)

Suppose X is a continuous RV with PDF  $f_X$ . Then Y = |X| is also continuous, and its PDF is 0 for  $y \le 0$  and

$$f_Y(y) = f_X(y) + f_X(-y)$$
 for  $y \geqslant 0$ .

#### Example (Ross E.g. 5.7b)

Suppose X is a continuous RV with PDF  $f_X$ . Then  $Y = X^2$  is also continuous, and its PDF is 0 for  $y \le 0$  and

$$f_Y(y) = rac{f_X(\sqrt{y}) + f_X(-\sqrt{y})}{2\sqrt{y}}$$
 for  $y \geqslant 0$ .

Here's an interesting example whose full understanding involves transformations of RVs.

Example (Bertrand's paradox; Ross E.g. 5.3d)

Consider a random chord of a circle. What is the probability that the length of the chord will be greater than the side of the equilateral triangle inscribed in that circle?

THE PROBLEM: We can interpret 'random chord of a circle' in two different ways.

## Normal RVs (Ross Sec 5.4)

It's time to meet the next important family of RVs.

#### **Definition**

Let  $\mu$  be a real value and  $\sigma > 0$ . A RV X is **normal with parameters**  $\mu$  and  $\sigma^2$ , or just  $\mathbf{N}(\mu, \sigma^2)$ , if it is continuous with PDF

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma}e^{-(x-\mu)^2/2\sigma^2}$$
 for  $-\infty < x < \infty$ .

It is **standard normal** if it is N(0, 1).

These are important because they appear in another powerful approximation theorem for binomial RVs — more on that later. For now we study their basic properties.

First, why the constant  $\frac{1}{\sqrt{2\pi}\sigma}$ ?

Since f is a PDF it needs to satisfy  $\int_{-\infty}^{\infty} f(x) dx = 1$ . It turns out (!) that

$$\int_{-\infty}^{\infty} e^{-(x-\mu)^2/2\sigma^2} dx = \sigma \int_{-\infty}^{\infty} e^{-y^2/2} dy = \sigma \sqrt{2\pi}.$$

(PROOF: Curious trick using polar coordinates.)