# Linear regression

Frank Edwards

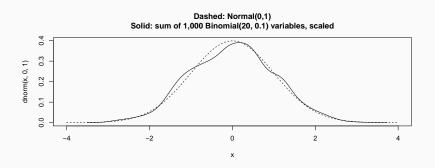
2/7/2020

### Is reality linear?

- · Linear regression is the dominant model in social science
- $\cdot$  It is (obviously) an inadequate scientific model
- · But still useful

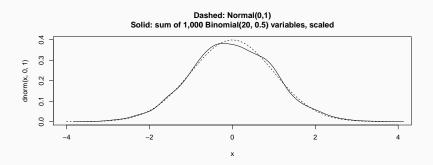
### Why Normal distributions are so common

```
output<-list()
for(i in 1:1000){output[[i]]<-rbinom(1000, 20, 0.1)}
combined<-bind_cols(output)
combined_sum<-apply(combined, 1, sum)
x<-seq(-4, 4, length.out=100)
plot(x, dnorm(x, 0, 1),
    lty = 2, type = "l",
    main = "Dashed: Normal(0,1)\nSolid: sum of 1,000 Binomial(20, 0.1) variables, scaled")
lines(density(scale(combined_sum)))</pre>
```



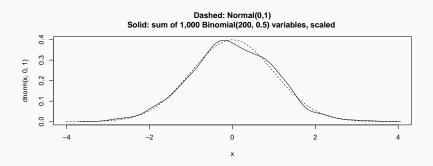
#### Why Normal distributions are so common

```
output<-list()
for(i in 1:1000){output[[i]]<-rbinom(1000, 20, 0.5)}
combined<-bind_cols(output)
combined_sum<-apply(combined, 1, sum)
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plot(x, dnorm(x, 0, 1),
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#### Why Normal distributions are so common

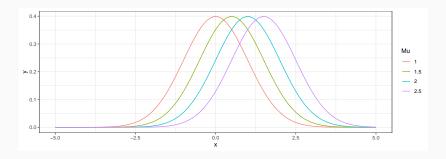
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```



### Getting to know the Normal

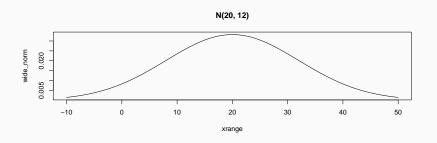
- Normal (Gaussian) have two parameters: mean  $(\mu)$  and standard deviation  $(\sigma)$
- These two parameters describe the location and spread of the distribution



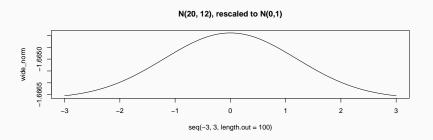


### Fun fact: every normal can be a N(0,1)

```
xrange<-seq(-10,50,length.out=100)
wide_norm<-dnorm(xrange, mean = 20, sd = 12)
plot(xrange, wide_norm, type ="l",
    main = "N(20, 12)")</pre>
```



### Rescaling N(20,12)



### Model notation: a primer

**Likelihood**:  $W \sim Binomial(N, p)$ 

 $\textbf{Prior}: p \sim \textit{Uniform}(0,1)$ 

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Note the relationship between W and p.

```
library(rethinking)
data("Howell1")
d<-Howell1
head(d)</pre>
```

```
## height weight age male
## 1 151.765 47.82561 63 1
## 2 139.700 36.48581 63 0
## 3 136.525 31.86484 65 0
## 4 156.845 53.04191 41 1
## 5 145.415 41.27687 51 0
## 6 163.830 62.99259 35 1
```

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What is our next step?

### Set priors!

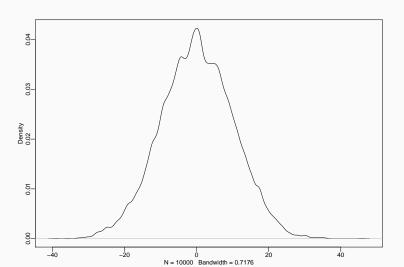
What should our prior for  $\mu$  be?

How about Normal(0, 10)?

$$\mu \sim \mathit{N}(\mathrm{0,10})$$

### What does our prior imply?

```
prior_sims<-rnorm(1e4, 0, 10)
dens(prior_sims)</pre>
```



The average adult male is 176.5 cm, the average adult female is 163 cm. Maybe we should set the mean as  $\frac{176.5+163}{2}=169.75$ 

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How certain are we that the true mean for this sample will be equal to 169.75?

- Completely certain?  $\mu \sim N(169.75, 0.01)$
- Pretty darn sure?  $\mu \sim N(169.75, 2)$
- Probably in the ballpark  $\mu \sim N(169.75, 10)$
- OK, not too sure, but not a bad guess  $\mu \sim \textit{N} (\text{169.75}, \text{25})$

### Simulate the priors for $\mu$

```
prior_1<-rnorm(1e4, 169.75, 1)
prior_2<-rnorm(1e4, 169.75, 2)
prior_3<-rnorm(1e4, 169.75, 10)
prior_4<-rnorm(1e4, 168.75, 25)</pre>
```

### Check out the simulations for $\mu$

```
quantile(prior 1, c(0.05, 0.5, 0.9))
         5%
##
                 50%
                          90%
## 168.1063 169.7659 171.0110
quantile(prior 2, c(0.05, 0.5, 0.9))
        5%
                50%
                          90%
##
## 166.5096 169.7547 172.3133
quantile(prior_3, c(0.05, 0.5, 0.9))
##
        5%
                 50%
                          90%
## 153.5469 169.7145 182.6973
quantile(prior_4, c(0.05, 0.5, 0.9))
##
         5%
                 50%
                          90%
## 128.0699 168.8352 201.1157
```

#### Now for $\sigma$

How much do individuals vary on average from the population average? How variable is height?

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Recall what a standard deviation is:

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Note:  $\sigma$  is constrained to be positive by definition

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Our prior for  $\sigma$  captures our beliefs about how much variance there is within the population.

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Always be sure to think through the scale of the outcome we are modeling (e.g. height)

- · No idea how variable the population is:  $\sigma \sim \text{Uniform}(0,50)$
- · We know there's not much variation:  $\sigma \sim \text{Uniform}(0,10)$

## Understanding the prior: prior predictive simulation

- 1. Simulate  $\mu$
- 2. Simulate  $\sigma$
- 3. Draw predictions from  $\mathit{N}(\mu,\sigma)$

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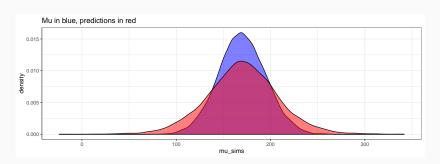
- 1. Simulate  $\mu$
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- 3. Draw predictions from  $\mathit{N}(\mu,\sigma)$

```
### simulate mu prior: mu ~ N(169.75, 25) prior
mu_sims<-rnorm(1e4, 169.75, 25)
### simulate sigma prior: sigma ~ U(0,50)
sigma_sims<-runif(1e4, 0, 50)
### simulate heights: N(mu, sigma)
height_preds<-rnorm(1e4, mu_sims, sigma_sims)
summary(height_preds)</pre>
```

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## -24.17 146.21 170.00 169.91 193.95 341.76
```

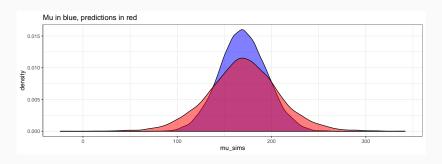
#### Visualize these predictions

```
plot_dat<-tibble(height_preds, mu_sims)
ggplot(plot_dat, aes(x = mu_sims)) +
   geom_density(fill = "blue", alpha = 0.5) +
   geom_density(aes(x = height_preds), fill = "red", alpha = 0.5) +
   labs(title = "Mu in blue, predictions in red")</pre>
```



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```



Predictions include both parameter uncertainty and sampling uncertainty

#### Model notation

**Likelihood:**  $h_i \sim Normal(\mu, \sigma)$ 

 $\mu$  Prior:  $\mu \sim \mathit{Normal}(\text{169.75}, \text{25})$ 

 $\sigma$  Prior:  $\sigma \sim \textit{Uniform}(0,50)$ 

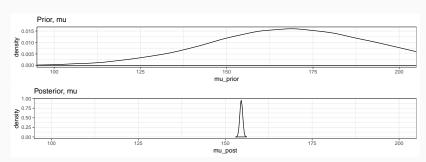
#### Now translate into an R formula

Likelihood:  $h_i \sim Normal(\mu, \sigma)$ 

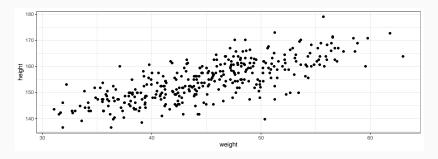
 $\mu$  Prior:  $\mu \sim Normal(169.75, 25)$ 

```
\sigma Prior: \sigma \sim Uniform(0,50)
## remove children with filter()
d2<- d %>%
  filter(age>=18)
## define model
model formula<-alist(
  height ~ dnorm(mu, sigma),
  mu ~ dnorm(169.75, 25),
  sigma ~ dunif(0,50)
## estimate with quap
m0<-quap(model_formula, data = d2)
summary(m0)
##
                           sd
                                    5.5%
                                              94.5%
               mean
         154.601182 0.4120254 153.942686 155.259678
## sigma
          7.731329 0.2913855 7.265638
                                           8.197019
```

#### Compare results to the prior



# Predicting height with other variables



What does it mean to assume a linear relationship between height and weight?

## The anatomy of a linear model

Begin with the Normal model of height

$$h_i \sim Normal(\mu, \sigma)$$
  
 $\mu \sim Normal(169.75, 25)$   
 $\sigma \sim Uniform(0, 50)$ 

Let  $x_i$  be the weight of the person in row i

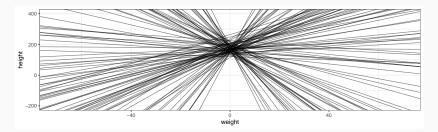
$$h_i \sim \text{Normal}(\mu_i, \sigma)$$
  
 $\mu_i = \alpha + \beta(x_i - \bar{x})$   
 $\alpha \sim \text{Normal}(169.75, 25)$   
 $\beta \sim \text{Uniform}(0, 10)$ 

# What does our prior imply?

Our prior now describes a line:  $mu = \alpha + \beta x$ 

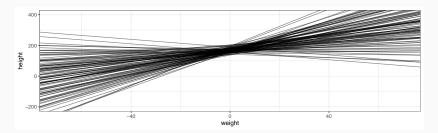
#### What does our prior imply?

### Our prior now describes a line: $mu = \alpha + \beta x$



#### Let's try again

Let's shrink the variance on  $\alpha$ , and set  $\beta$  to be more likely positive



## Estimating the model

Likelihood for height:  $h_i \sim \textit{Normal}(\mu_i, \sigma)$ 

Linear model for mean:  $\mu_i = \alpha + \beta x_i$ 

**Prior for intercept:**  $\alpha \sim Normal(169.75, 15)$ 

Prior for slope:  $\beta \sim Normal(2, 2)$ 

Prior for standard deviation:  $\sigma \sim \textit{Uniform}(0,50)$ 

```
data(Howell1)
d<-Howell1 %>%
  filter(age>=18)
xbar<-mean(d$weight)

m1<-quap(
  alist(
    height ~ dnorm( mu , sigma) ,
    mu <- a + b * weight,
    a ~ dnorm(178 , 20),
    b ~ dlnorm(0 , 1),
    sigma ~ dunif(0 , 50)
),
    data = d
)</pre>
```

#### Interpreting the posterior: table

#### summary(m1)

```
## mean sd 5.5% 94.5%

## a 114.534318 1.89774726 111.5013512 117.5672846

## b 0.890730 0.04175799 0.8239927 0.9574674

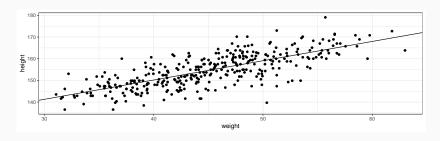
## sigma 5.072719 0.19124893 4.7670660 5.3783715
```

## Examining posterior samples

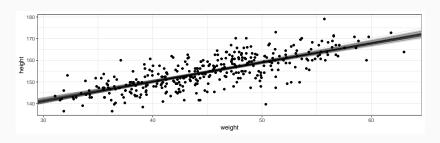
```
post_m1<-extract.samples(m1)
head(post_m1)</pre>
```

```
## a b sigma
## 1 114.0870 0.8958745 4.786056
## 2 117.2360 0.8368378 4.925024
## 3 117.3856 0.8397231 5.240202
## 4 119.2280 0.7849737 5.154901
## 5 115.1533 0.8658036 4.928546
## 6 116.7038 0.8448853 4.835857
```

## Visualizing the estimated posterior: posterior mean



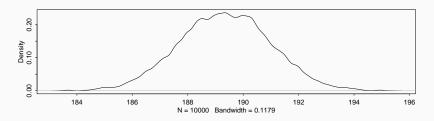
## But the posterior is more than one line



## At a single value of x

I weight about 84 kg. Let's solve for h

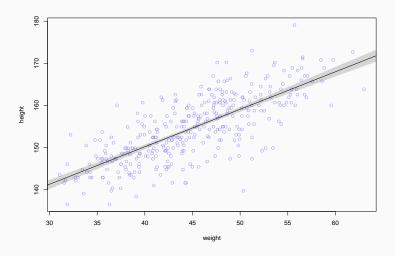
$$E(h_i) = \mu_i = \alpha + \beta x_i$$



These are the values of  $\mu$  for individuals of weight 84kg that are compatible with the data

## Summarizing the posterior over the full range of the data

# Visualizing the posterior for mu: 89 percent PI



## Predicting new data; validating the model

Recall the method for posterior prediction:

- 1. Draw samples of the parameters from the posterior
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```
## already got the samples in post_m1
head(post_m1)
```

```
## a b sigma
## 1 114.0870 0.8958745 4.786056
## 2 117.2360 0.8368378 4.925024
## 3 117.3856 0.8397231 5.240202
## 4 119.2280 0.7849737 5.154901
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## 6 116.7038 0.8448853 4.835857
```

We haven't used  $\sigma$  yet. It will tell our model how much sampling uncertainty to produce

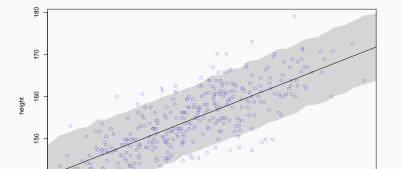
### Method for predictions

We could do it by hand, or we could just use sim(). It will give us 1,000 predictions for each of 43 unique height values we provided

```
height_preds<-sim(m1, data = sim_dat)
height.PI<-apply(height_preds, 2, PI, prob = 0.89)
```

## Prediction uncertainty

```
# plot raw data
plot( height ~ weight , d2 , col=col.alpha(rangi2,0.7) )
# draw MAP line
lines( weights , mu.mean )
# draw PI region for simulated heights
shade( height.PI , weights )
```



#### Next time

- · Polynomials and splines
- $\cdot$  Regression with multiple predictors

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- $\cdot$  I'll post the new homework on Friday

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- · Regression with multiple predictors
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- · A note on Q5:

boys\_after\_girls<-birth2[birth1==0]</pre>