## STA 522 Sample Exam 1 Solutions

## Problem 1

 $X_1, X_2, \dots, X_n$  are iid with common cdf

$$F(x \mid \alpha, \beta) = \begin{cases} 0, & x < 0 \\ (x/\beta)^{\alpha}, & 0 \le x \le \beta \\ 1, & x > \beta \end{cases}$$

Therefore, the common pdf of  $X_1, X_2, \dots, X_n$  is given by:

$$f(x \mid \alpha, \beta) = \frac{d}{dx} F(x \mid \alpha, \beta) = \frac{\alpha}{\beta^{\alpha}} x^{\alpha - 1} I(0 \le x \le \beta)$$

Part (a): The joint pdf of  $X_1, X_2, \ldots, X_n$  is:

$$f(x_1, \dots, x_n \mid \alpha, \beta) \stackrel{\text{iid}}{=} \prod_{i=1}^n f(x_i \mid \alpha, \beta)$$

$$= \frac{\alpha^n}{\beta^{n\alpha}} \left( \prod_{i=1}^n x_i \right)^{\alpha - 1} \prod_{i=1}^n I(0 \le x_i \le \beta)$$

$$= \underbrace{\frac{\alpha^n}{\beta^{n\alpha}} \left( \prod_{i=1}^n x_i \right)^{\alpha - 1} I(x_{(n)} \le \beta)}_{=g(T(\underline{x} \mid \alpha, \beta))} \underbrace{I(x_{(1)} > 0)}_{=h(\underline{x})}$$

where  $T(\underline{x}) = (\prod_{i=1}^n x_i, x_{(n)})$ . Therefore, by the Factorization theorem,  $T(\underline{X}) = (\prod_{i=1}^n X_i, X_{(n)})$  is sufficient for  $\alpha, \beta$ .

**Part** (b): The joint likelihood for  $\alpha, \beta$  is:

$$L(\alpha, \beta \mid \underline{x}) = \frac{\alpha^n}{\beta^{n\alpha}} \left( \prod_{i=1}^n x_i \right)^{\alpha - 1} I(x_{(n)} \le \beta) \ I(x_{(1)} > 0)$$

For any  $\alpha$ , the likelihood function is decreasing in  $\beta$  and is non-zero when  $\beta \geq x_{(n)}$ . Hence,  $\hat{\beta} = X_{(n)}$  is the MLE of  $\beta$  for all  $\alpha$ .

The profile likelihood for  $\alpha$  is:

$$\log \tilde{L}(\alpha \mid \underline{x}) = \frac{\alpha^n}{x_{(n)}^{n\alpha}} \left( \prod_{i=1}^n x_i \right)^{\alpha - 1} \implies \log \tilde{L}(\alpha \mid \underline{x}) = n \log \alpha - n\alpha \log x_{(n)} + (\alpha - 1) \sum_{i=1}^n \log x_i$$

Therefore

$$\frac{\partial}{\partial \alpha} \log \tilde{L}(\alpha \mid \underline{x}) = \frac{n}{\alpha} - n \log x_{(n)} + \sum_{i=1}^{n} \log x_{i} \gtrsim 0$$

$$\iff \frac{n}{\alpha} \gtrsim n \log x_{(n)} - \sum_{i=1}^{n} \log x_{i} = \log \left( \frac{x_{(n)}^{n}}{\prod_{i=1}^{n} x_{i}} \right)$$

$$\iff \alpha \lesssim \frac{n}{\log \left( \frac{x_{(n)}^{n}}{\prod_{i=1}^{n} x_{i}} \right)}$$

Hence, the MLE of  $\alpha$  is  $\hat{\alpha} = \frac{n}{\log\left(\frac{X_{(n)}^n}{\prod_{i=1}^n X_i}\right)}$  and the MLE of  $\beta$  is  $\hat{\beta} = X_{(n)}$ .

## Problem 2

**Part** (a): First find the MLE of  $\theta$ . This is from Lecture 6:

The likelihood of  $\theta$  is

$$L(\theta \mid \underline{x}) = \prod_{i=1}^{n} \exp(-\theta) \frac{\theta^{x_i}}{x_i!} = \exp(-n\theta) \frac{\theta^{\sum_{i=1}^{n} x_i}}{\prod_{i=1}^{n} x_i!}$$

The log likelihood is:

$$l(\theta \mid \underline{x}) = \log L(\theta \mid \underline{x}) = -n\theta + \left(\sum_{i=1}^{n} x_i\right) \log \theta - \log \left(\prod_{i=1}^{n} x_i!\right)$$

Therefore,

$$\frac{d \log L(\theta \mid \underline{x})}{d\theta} = -n + \left(\sum_{i=1}^{n} x_i\right) \frac{1}{\theta} \gtrsim 0 \text{ according as } \theta \lesssim \frac{1}{n} \sum_{i=1}^{n} x_i = \overline{x}$$

Therefore,  $\hat{\theta} = \overline{x}$  is the MLE of  $\theta$ . Hence, by the invariance property of the MLE  $e^{-\overline{X}}$  is the MLE of  $P(X_1 = 0) = e^{-\theta}$ .

Part (b): No, the MLE is <u>not</u> unbiased. To see this, first note that  $g(x) = -e^{-x}$  is a convex function. Therefore as suggested in the hint,

$$\begin{split} & \mathrm{E}[g(\overline{X})] \overset{\mathrm{Jensen}}{>} g(\mathrm{E}(\overline{X})) \quad \text{(strict inequality as $g$ is strictly convex)} \\ & \mathrm{i.e., E}[-e^{-\overline{X}}] > -e^{-\mathrm{E}(\overline{X})} = -e^{-\theta} \\ & \iff \mathrm{E}[e^{-\overline{X}}] < e^{-\theta} \end{split}$$