STA 522/Solutions to Homework 3

Problem 6.20 (Part a)

We start by finding a sufficient statistic. The joint pdf of X_1, X_2, \ldots, X_n is given by

$$f(x_1, \dots, x_n \mid \theta) = \prod_{i=1}^n \left\{ \frac{2x_i}{\theta^2} I(0 < x_i < \theta) \right\}$$
$$= \underbrace{\left(2^n \prod_{i=1}^n x_i \right)}_{=h(x)} \underbrace{\left(\frac{1}{\theta^{2n}} I\left(x_{(n)} < \theta \right) \right)}_{g(T(\underline{x}) \mid \theta)}$$

Therefore, by Factorization theorem, $T(\underline{X}) = X_{(n)} = \max_{1 \le i \le n} X_i$ is sufficient for θ . To obtain the pdf of $T = X_{(n)}$ first note that the common cdf of X_i is

$$F(x \mid \theta) = \int_0^x \frac{2y}{\theta^2} dy = \frac{x^2}{\theta^2}; \quad \text{for } 0 < x < \theta.$$

Therefore, it follows from a result discussed in class that the pdf of $T = X_{(n)}$ is given by:

$$f_T(t \mid \theta) = n \ f(t \mid \theta) \ \{F(t \mid \theta)\}^{n-1} = \frac{2n}{\theta^{2n}} t^{2n-1} \ I(0 < y < \theta)$$

To prove completeness start with a function g(t) with

$$u(\theta) := \mathcal{E}_{\theta}[g(T)] = \int_{0}^{\theta} g(t) \frac{2n}{\theta^{2n}} t^{2n-1} dt = 0 \text{ for all } \theta > 0.$$

Then $u'(\theta) = 0$ (derivative of a constant function) for all θ , which implies

$$0 = \frac{d}{d\theta} \left(\frac{2n}{\theta^{2n}} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)$$

$$= \left(\frac{d}{d\theta} \frac{2n}{\theta^{2n}} \right) \int_0^\theta g(t) \ t^{2n-1} \ dt + \frac{2n}{\theta^{2n}} \qquad \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)} \underbrace{\left(\frac{d}{d\theta} \int_0^\theta g(t) \ t^{2n-1} \ dt \right)}_{=g(\theta)$$

which implies $P_{\theta}(g(T) = 0) = 1$ for all $\theta \in (0, \theta)$. This means $T = X_{(n)}$ is complete.

Problem 6.22

Part (a): The joint pdf of X_1, X_2, \ldots, X_n is given by:

$$f(\underline{x} \mid \theta) = \underbrace{\theta^n \exp\left[(\theta - 1) \sum_{i=1}^n \log x_i\right]}_{=g(T(\underline{x})|\theta)} \underbrace{\prod_{i=1}^n I(0 < x_i < 1)}_{=h(\underline{x})}$$

Therefore by the Factorization theorem, $T(\underline{X}) = \sum_{i=1}^{n} \log X_i$ is a sufficient statistics for θ . Since $\sum_{i=1}^{n} X_i$ is not a one-to-one function of $T(\underline{X}) = \sum_{i=1}^{n} \log X_i$, therefore $\sum_{i=1}^{n} X_i$ is NOT a sufficient statistic for θ (from the reverse implication of the Factorization theorem).

Part (b): We note that $f(x \mid \theta)$ is a member of the exponential family:

$$f(x \mid \theta) = \underbrace{\theta^n}_{=c(\theta)} \underbrace{I(0 < x < 1)}_{=h(x)} \exp \left[\underbrace{(\theta - 1)}_{=w_1(\theta)} \underbrace{\log x}_{=T_1(x)} \right]$$

and the parameter space $\Theta = \{\theta : \theta > 0\}$ contains the open interval, e.g., (1,2) in \mathbb{R}^1 . Therefore, using a theorem on exponential family discussed in class, $\sum_{i=1}^n T_1(X_i) = \sum_{i=1}^n \log X_i$ is complete sufficient for θ .

Problem 6.40

Say $X_1, X_2, ..., X_n$ be iid from the location scale family $\frac{1}{\sigma} f\left(\frac{x-\mu}{\sigma}\right)$. Then we can write $X_i = \mu + \sigma Z_i$ where $Z_i \sim \text{iid } f(x)$.

Part (a): We have

$$\frac{T_1(X_1, X_2, \dots, X_n)}{T_2(X_1, X_2, \dots, X_n)} = \frac{T_1(\mu + \sigma Z_1, \dots, \mu + \sigma Z_n)}{T_2(\mu + \sigma Z_1, \dots, \mu + \sigma Z_n)} = \frac{\sigma}{\sigma} \frac{T_1(Z_1, \dots, Z_n)}{T_2(Z_1, \dots, Z_n)} = \frac{T_1(Z_1, \dots, Z_n)}{T_2(Z_1, \dots, Z_n)}$$

The right hand side involves random variables whose pdf does not involve the parameters μ and σ . Hence T_1/T_2 is ancillary.

Part (b): We have for a > 0, $b \in \mathbb{R}$ and any $\underline{x} = (x_1, \dots, x_n)$

$$R(aX_1 + b, ..., aX_n + b) = \max_i (aX_i + b) - \min_i (aX_i + b) = a\left(\max_i X_i - \min_i X_i\right) = aR(X_1, ..., X_n)$$

and

$$S(aX_1 + b, ..., aX_n + b) = \left(\frac{1}{n-1} \sum_{i=1}^n \left((aX_i + b) - \frac{1}{n} \sum_{i=1}^n (aX_i + b) \right)^2 \right)^{1/2}$$
$$= a \left(\frac{1}{n-1} \sum_{i=1}^n \left(X_i - \overline{X} \right)^2 \right)^{1/2}$$
$$= a S(X_1, ..., X_n)$$