STA 522, Spring 2021
Introduction to Theoretical Statistics II

Lecture 7

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AGENDA

- Wrap up discussions of the Bayesian approach to Statistics
- Methods of evaluating estimators
- Cramér-Rao Lower Bound

Review: Bayesian approach to statistics

- The parameter θ is considered a random variable. Consider a **prior distribution** for θ before observing any data.

- After drawing a sample find the likelihood function for θ, and use Bayes' Rule to update the prior with the likelihood function to get the **posterior distribution**:

$$\pi(\theta \mid \underline{x}) = \frac{f(\underline{x} \mid \theta)\pi(\theta)}{m(\underline{x})} = \frac{f(\underline{x}, \theta)}{m(\underline{x})},$$

  where $m(\underline{x})$ is the marginal distribution of $\underline{X}$:

$$m(\underline{x}) = \int f(\underline{x} \mid \theta)\pi(\theta)\, d\theta.$$

- The posterior distribution combines information from prior and likelihood.
- One Bayesian point estimator of θ is given by the posterior mean $E(\theta \mid \underline{X})$ (can also use posterior median, posterior mode, etc.)

Example (Binomial Bayes estimation): Let $X_1, X_2, \ldots, X_n \sim$ iid Bernoulli (p), and let $Y = \sum_{i=1}^{n} X_i$. Then $Y \sim$ binomial(n, p). Assume the prior distribution on p to be beta(α, β).

The posterior distribution of p is

$$p \mid Y \sim \text{Beta}(y + \alpha, n - y + \beta)$$

The posterior mean is:

$$\begin{aligned}
\hat{p}_B &= \text{E}(p \mid Y) \\
&= \frac{y + \alpha}{n + \alpha + \beta} \\
&= \left(\frac{n}{n + \alpha + \beta}\right) \underbrace{\left(\frac{y}{n}\right)}_{=\text{sample mean}} + \left(\frac{\alpha + \beta}{n + \alpha + \beta}\right) \underbrace{\left(\frac{\alpha}{\alpha + \beta}\right)}_{=\text{prior mean}}
\end{aligned}$$

Conjugate Family

Definition: Let $\mathcal{F}$ denote the class of pdfs or pmfs $f(x \mid \theta)$, indexed by θ. A class Π of prior distributions is a **conjugate family** for $\mathcal{F}$ if the posterior distribution is in the class Π for all $f \in \mathcal{F}$, all priors in Π, and all $x \in \mathcal{X}$.

- The beta family is conjugate for the binomial family, which is why it was chosen as the prior distribution in the previous example.
- the gamma family is conjugate for the Poisson family.
- the normal family is its own conjugate.

Example (Normal Bayes Estimator): Let $X \sim \mathsf{N}(\theta, \sigma^2)$, and suppose that the prior distribution on θ is $\mathsf{N}(\mu, \tau^2)$ where σ^2, μ and τ^2 are all known.

The posterior distribution of θ is also normal (Exercise 7.22; Homework) with

$$\mathsf{E}(\theta \mid x) = \frac{\tau^2}{\sigma^2 + \tau^2} x + \frac{\sigma^2}{\sigma^2 + \tau^2} \mu = \frac{1/\sigma^2}{1/\sigma^2 + 1/\tau^2} x + \frac{1/\tau^2}{1/\sigma^2 + 1/\tau^2} \mu$$

and

$$\mathsf{Var}(\theta \mid x) = \frac{\sigma^2 \tau^2}{\sigma^2 + \tau^2} = \frac{1}{1/\sigma^2 + 1/\tau^2}$$

Using the posterior mean, a Bayes point estimator is given by $\mathsf{E}(\theta \mid X)$.

Note that the Bayes estimator is again a linear combination of prior and sample means.

Method of Evaluating Estimators

- There may exist multiple estimators for the same problem, obtained from different approaches, e.g., method of moments, maximum likelihood, Bayesian approach etc.
- We want to compare these estimators and possibly obtain the "best" estimator.

Definition: The **mean squared error (MSE)** of an estimator W of a parameter θ is the function of θ defined by

$$\mathsf{MSE} = \mathsf{MSE}_\theta(W) = \mathsf{E}_\theta\left[(W-\theta)^2\right].$$

Note: **mean absolute error**, defined as

$$\mathsf{E}_\theta\left[|W-\theta|\right],$$

is an alternative for measuring the performance of an estimator.

Definition: The **bias** of a point estimator W of a parameter θ is the difference between the expected value of W and θ; that is,

$$\mathrm{Bias}_\theta(W) = \mathrm{E}_\theta(W) - \theta.$$

An estimator whose bias is identically equal to 0 *as a function of* θ is called **unbiased** and satisfies $\mathrm{E}_\theta(W) = \theta$ for all θ.

Note that $\mathrm{MSE} = \mathrm{E}_\theta\left[(W-\theta)^2\right] = \mathrm{Var}_\theta(W) + [\mathrm{Bias}_\theta(W)]^2$

- For an unbiased estimator, we have

$$\mathrm{MSE} = \mathrm{E}_\theta\left[(W-\theta)^2\right] = \mathrm{Var}_\theta(W)$$

- Unbiased is a good property for an estimator to have, but it can be misleading.

Example: Let $X_1, X_2, \ldots, X_n \sim$ iid Bernoulli(p). Let $W = X_1$. Since $E(W) = p$, W is unbiased, but doesn't use all the data. Note that for W,

$$\text{MSE}(W) = \text{Var}(W) = p(1-p).$$

Example: Let $X_1, X_2, \ldots, X_n \sim$ iid $\text{N}(\mu, \sigma^2)$. Since $\text{E}(\overline{X}) = \mu$ and $\text{E}(S^2) = \sigma^2$, $\overline{X}$ and S^2 are unbiased for μ and σ^2.

Thus, the mean squared errors (see Thms. 5.2.6 and 5.3.1) are

$$\text{MSE}(\overline{X}) = \text{Var}(\overline{X}) = \frac{\sigma^2}{n}$$

$$\text{MSE}(S^2) = \text{Var}(S^2) = \frac{2\sigma^4}{n-1}$$

Recall the MLE for σ^2 is $\hat{\sigma}^2 = \frac{n-1}{n} S^2$.

Note that

$$\mathrm{E}\left(\hat{\sigma}^2\right) = \mathrm{E}\left(\frac{n-1}{n} S^2\right) = \frac{n-1}{n}\sigma^2 \implies \left[\mathrm{Bias}\left(\hat{\sigma}^2\right)\right]^2 = \frac{\sigma^4}{n^2}$$

$$\mathrm{Var}\left(\hat{\sigma}^2\right) = \mathrm{Var}\left(\frac{n-1}{n} S^2\right) = \frac{2(n-1)\sigma^4}{n^2}$$

Therefore

$$\mathrm{MSE}\left(\hat{\sigma}^2\right) = \mathrm{Var}\left(\hat{\sigma}^2\right) + \left[\mathrm{Bias}\left(\hat{\sigma}^2\right)\right]^2 = \frac{2n-1}{n^2}\sigma^4$$

Note that

$$\frac{2n-1}{n^2} = \frac{2}{n} - \frac{1}{n^2} < \frac{2}{n-1}$$

which implies $\mathrm{MSE}\left(\hat{\sigma}^2\right) < \mathrm{MSE}\left(S^2\right)$.

Example (7.3.5; Contd.): Let $X_1, X_2, \ldots, X_n \sim$ iid Bernoulli(p). Then $Y = \sum_{i=1}^n X_i \sim$ binomial(n, p). Recall that we have developed two estimators for p, the MLE and the Bayes estimator:

$$\hat{p} = \frac{Y}{n} = \overline{X}$$

$$\hat{p}_B = \frac{Y + \alpha}{\alpha + \beta + n}$$

We have

$$\begin{aligned}
\mathrm{MSE}_p(\hat{p}) &= \frac{p(1-p)}{n} \\
\mathrm{MSE}_p(\hat{p}_B) &= \mathrm{Var}_p(\hat{p}_B) + (\mathrm{Bias}_p(\hat{p}_B))^2 \\
&= \mathrm{Var}_p\left(\frac{Y + \alpha}{\alpha + \beta + n}\right) + \left[\mathrm{E}_p\left(\frac{Y + \alpha}{\alpha + \beta + n}\right) - p\right]^2 \\
&= \frac{np(1-p)}{(\alpha + \beta + n)^2} + \left(\frac{np + \alpha}{\alpha + \beta + n} - p\right)^2
\end{aligned}$$

Choosing $\alpha = \beta = \frac{\sqrt{n}}{2}$ makes the MSE of $\hat{p}_B$ constant as a function of p. Under this choice the MSEs are as follows:

$$\text{MSE}(\hat{p}) = \frac{p(1-p)}{n}$$

$$\text{MSE}(\hat{p}_B) = \frac{n}{4(n+\sqrt{n})^2}$$

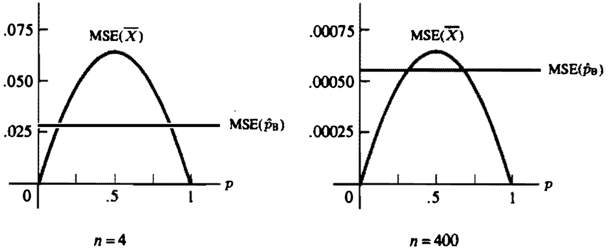

Figure 7.3.1. *Comparison of MSE of $\hat{p}$ and $\hat{p}_B$ for sample sizes $n = 4$ and $n = 400$ in Example 7.3.5*

Finding the "Best" Estimator

- Depends on what "best" means.
- Depends on the value of the parameter.
- There may be situations where a biased estimator (like $\hat{p}_B$) may be better.
- We first define "best" in relation to the variance of an unbiased estimator.

Definition: An estimator W^* is a **best unbiased estimator** of $\tau(\theta)$ if it satisfies the following:

(a) W^* is unbiased, i.e., $\mathrm{E}_\theta(W^*) = \tau(\theta)$ for all θ; and

(b) among all unbiased estimators, the variance (or MSE) of W^* is a minimum, i.e., for any other estimator W with $\mathrm{E}_\theta(W) = \tau(\theta)$, we have

$$\mathrm{MSE}(W^*) = \mathrm{Var}_\theta(W^*) \leq \mathrm{Var}_\theta(W) = \mathrm{MSE}(W)$$

for all θ. W^* may also be called a **uniform minimum variance unbiased estimator (UMVUE)** of $\tau(\theta)$.

Cramér-Rao Inequality

- It is usually hard to determine if a UMVUE exists.
- However, there is a lower bound on the variance of any unbiased estimator.
- So if we can find an unbiased estimator that achieves this lower bound, we know it must be UMVUE.

Theorem (7.3.9; Cramér-Rao Lower Bound)

Let $\underline{X} = (X_1, X_2, \ldots, X_n)$ have pdf $f(\underline{x} \mid \theta)$, and let $W(\underline{X})$ be any estimator satisfying

(a) $\dfrac{d}{d\theta} \mathrm{E}_\theta \left[W(\underline{X})\right] = \displaystyle\int_{\mathcal{X}} \frac{\partial}{\partial\theta} \left[W(\underline{x}) f(\underline{x} \mid \theta)\right] d\underline{x}$; and

(b) $\mathrm{Var}_\theta \left[W(\underline{X})\right] < \infty$.

Then

$$\mathrm{Var}_\theta(W(\underline{X})) \geq \frac{\left[\frac{d}{d\theta} \mathrm{E}_\theta \left[W(\underline{X})\right]\right]^2}{\mathrm{E}_\theta \left[\left[\frac{\partial}{\partial\theta} \log f(\underline{X} \mid \theta)\right]^2\right]}.$$

Proof: Note that for any two random variables U and V

$$[\text{Cov}(U,V)]^2 \overset{\text{Cauchy-Schwarz}}{\leq} \text{Var}(U)\,\text{Var}(V) \implies \text{Var}(U) \geq \frac{[\text{Cov}(U,V)]^2}{\text{Var}(V)}$$

Take $U \equiv W(\underline{X})$ and $V \equiv \frac{\partial}{\partial\theta} \log f(\underline{X} \mid \theta)$. Note that

$$\text{E}(V) = E_\theta\left(\frac{\partial}{\partial\theta} \log f(\underline{X} \mid \theta)\right) = \int_{\mathcal{X}} \frac{\partial}{\partial\theta} \log f(\underline{x} \mid \theta)\, f(\underline{x} \mid \theta)\, d\underline{x} = \frac{\partial}{\partial\theta} E(1) = 0$$

So, $\text{Cov}(U,V) = E(UV)$ and $\text{Var}(V) = \text{E}(V^2)$. Note that

$$\text{E}(UV) = \int_{\mathcal{X}} W(\underline{x})\, \frac{\partial}{\partial\theta} f(\underline{x} \mid \theta)\, d\underline{x} = \frac{d}{d\theta}\, \text{E}_\theta\left(W(\underline{X})\right)$$

Therefore, combining we get

$$\text{Var}_\theta(W(\underline{X})) = \text{Var}(U) \geq \frac{[\text{Cov}(U,V)]^2}{\text{Var}(V)} = \frac{[\text{E}(UV)]^2}{\text{E}(V^2)} = \frac{\left[\frac{d}{d\theta}\, \text{E}_\theta\left[W(\underline{X})\right]\right]^2}{\text{E}_\theta\left[\left[\frac{\partial}{\partial\theta} \log f(\underline{X} \mid \theta)\right]^2\right]}$$

Corollary: Cramér-Rao Lower Bound, iid case

If the assumptions of Theorem 7.3.9 are satisfied, and, additionally, if $X_1, X_2, \ldots, X_n$ are iid with pdf $f(x \mid \theta)$, then

$$\mathrm{Var}_\theta(W(\underline{X})) \geq \frac{\left[\frac{d}{d\theta}\,\mathrm{E}_\theta\left[W(\underline{X})\right]\right]^2}{n\,\mathrm{E}_\theta\left[\left[\frac{\partial}{\partial\theta}\log f(X_1 \mid \theta)\right]^2\right]}.$$

Proof: Homework. See p. 337 in the textbook.

Notes

- If $W(\underline{X})$ is unbiased for θ, then the numerator is

$$\left[\frac{d}{d\theta}\,\mathrm{E}_\theta\left[W(\underline{X})\right]\right]^2 = 1.$$

- The denominator is a function of the density, not the data.

Fisher Information

Definition: $\mathrm{E}_\theta\left[\left[\frac{\partial}{\partial\theta}\log f(\underline{X}\mid\theta)\right]^2\right]$ is called the **Fisher information** of the sample.

Lemma 7.3.11: Calculating the Fisher Information

If $f(x\mid\theta)$ satisfies

$$\frac{d}{d\theta}\mathrm{E}_\theta\left[\frac{\partial}{\partial\theta}\log f(X\mid\theta)\right]=\int_{\mathcal{X}}\frac{\partial}{\partial\theta}\left[\left[\frac{\partial}{\partial\theta}\log f(x\mid\theta)\right]f(x\mid\theta)\right]dx$$

(which is always true for an exponential family), then

$$\mathrm{E}_\theta\left[\left[\frac{\partial}{\partial\theta}\log f(X\mid\theta)\right]^2\right]=-\mathrm{E}_\theta\left[\frac{\partial^2}{\partial\theta^2}\log f(X\mid\theta)\right].$$

Example: Let $X_1, X_2, \ldots, X_n \sim$ iid gamma(α, β), and assume α is known. Show that $W(\underline{X}) = \frac{1}{\alpha}\overline{X}$ attains the Cramér-Rao lower bound for an unbiased estimator of β, and hence is UMVUE.

First note that $\mathrm{Var}_\beta(W(\underline{X})) = \mathrm{Var}_\beta\left(\dfrac{1}{\alpha}\overline{X}\right) = \dfrac{1}{\alpha^2}\mathrm{Var}_\beta\left(\overline{X}\right) = \dfrac{\beta^2}{\alpha n}$

Now obtain the CR lower bound. We have $f(x \mid \beta) = \frac{1}{\Gamma(\alpha)\beta^\alpha}\; x^{\alpha-1}e^{-x/\beta}$ which implies $\log f(x \mid \beta) = \log\Gamma(\alpha) - \alpha\log\beta + (\alpha-1)\log x - \frac{x}{\beta}$. Then

$$\frac{\partial}{\partial\beta}\log f(x \mid \beta) = -\frac{\alpha}{\beta} + \frac{x}{\beta^2}; \quad \frac{\partial^2}{\partial\beta^2}\log f(x \mid \beta) = \frac{\alpha}{\beta^2} - \frac{2x}{\beta^3}$$

Therefore

$$\mathrm{E}_\beta\left[\left[\frac{\partial}{\partial\beta}\log f(X_1 \mid \beta)\right]^2\right] = -\,\mathrm{E}_\beta\left[\frac{\partial^2}{\partial\beta^2}\log f(X_1 \mid \beta)\right] = \mathrm{E}_\beta\left[\frac{2X_1}{\beta^3} - \frac{\alpha}{\beta^2}\right] = \frac{\alpha}{\beta^2}$$

Of course, $W(\underline{X}) = \frac{1}{\alpha}\overline{X}$ being unbiased for β means $\frac{d}{d\beta}\mathrm{E}_\beta\left[W(\underline{X})\right] = 1$. Hence, the CR lower bound is (iid):

$$\frac{\left[\frac{d}{d\beta}\mathrm{E}_\beta\left[W(\underline{X})\right]\right]^2}{n\,\mathrm{E}_\beta\left[\left[\frac{\partial}{\partial\beta}\log f(X_1 \mid \beta)\right]^2\right]} = \frac{1}{\alpha/\beta^2} = \frac{\beta^2}{\alpha} = \mathrm{Var}_\beta(W(\underline{X}))$$

Example Let $X_1, X_2, \ldots, X_n \sim$ iid uniform$(0, \theta)$. The assumptions in CR inequality does not hold (verify; see p. 340 in the textbook). We will show that there exists an unbiased estimator of θ whose variance is uniformly smaller than the CRLB.

Note that here $\frac{\partial}{\partial\theta} \log f(x \mid \theta) = -\frac{1}{\theta} \implies \mathsf{E}_\theta\left[\left[\frac{\partial}{\partial\theta} \log f(X_1 \mid \theta)\right]^2\right] = \frac{1}{\theta^2}$.

So, CR lower bound $= \dfrac{\theta^2}{n}$.

Consider $Y = X_{(n)}$. $f_Y(y) = ny^{n-1}/\theta^n$, $0 < y < \theta$ so that

$$E_\theta(Y) = \int_0^\theta \frac{ny^n}{\theta^n}\, dy = \frac{n}{n+1}\,\theta \implies E_\theta\left(\underbrace{\frac{n+1}{n}Y}_{=U}\right) = \theta$$

i.e., U is an unbiased estimator of θ. We have

$$\mathsf{Var}_\theta\,(U) = \mathsf{Var}_\theta\left(\frac{n+1}{n}Y\right) = \frac{1}{n(n+2)}\theta^2 < \frac{1}{n}\theta^2$$

NOTE: In general, if the range of the pdf depends on the parameter, the Cramér-Rao Theorem will not be applicable.

Attainment

- There is no guarantee that the bound given in the Cramér-Rao Inequality is sharp. That is, our best unbiased estimator may not achieve the CRLB.

- Problem: when do we stop searching?

Homework

- Read p. 330 – 342.
- Exercises: TBA.