Stat 443: Time Series and Forecasting

Assignment 4: Analysis in the Frequency Domain

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March 28, 2024

Question 1

Consider the following second-order AR process AR(2) process for $\{X_t\}_{t\in\mathbb{Z}}$, where $\{Z_t\}_{t\in\mathbb{Z}}\stackrel{\mathrm{iid}}{\sim}\mathrm{WN}(0,\sigma^2)$.

$$X_t = \frac{7}{10}X_{t-1} - \frac{1}{10}X_{t-2} + Z_t$$

We have previously shown that the autocorrelation function $\gamma(h)$ for $h \in \mathbb{Z}$ is given by:

$$\rho(h) = \frac{16}{11} \left(\frac{1}{2}\right)^{|h|} - \frac{5}{11} \left(\frac{1}{5}\right)^{|h|}, \quad h \in \mathbb{Z}$$

Part A

Derive the normalized spectral density function $f^*(\omega)$ for $\{X_t\}_{t\in\mathbb{Z}}$.

Solution

We begin by verifying that the Fourier Transform is well defined.

$$\sum_{h=-\infty}^{\infty} |\rho(h)| = \sum_{h=-\infty}^{\infty} \left| \frac{16}{11} \left(\frac{1}{2} \right)^{|h|} - \frac{5}{11} \left(\frac{1}{5} \right)^{|h|} \right|^{?} < \infty$$

$$\sum_{t=-\infty}^{\infty} |\rho(h)| = \left(\frac{16}{11} - \frac{5}{11} \right) + 2 \left(\frac{16}{11} \sum_{h=1}^{\infty} \left(\frac{1}{2} \right)^{h} - \frac{5}{11} \sum_{h=1}^{\infty} \left(\frac{1}{5} \right)^{h} \right)$$

$$\sum_{t=-\infty}^{\infty} |\rho(h)| = 1 + 2 \left(\frac{16}{11} \left(\frac{1/2}{1 - 1/2} \right) - \frac{5}{11} \left(\frac{1/5}{1 - 1/5} \right) \right)$$

$$\sum_{t=-\infty}^{\infty} |\rho(h)| = 1 + 2 \left(\frac{16}{11} - \frac{5}{11} \left(\frac{1}{4} \right) \right) = \boxed{\frac{81}{22} < \infty, \therefore \text{ well-defined.}}$$

Now, we evaluate given ρ , recalling that for $\omega \in (0,1)$ and even functions, the normalized spectral density is given by:

$$f^{\star}(\omega) = \frac{1}{\pi} \left(\rho(0) + 2 \sum_{h=1}^{\infty} \rho(h) \cos(\omega h) \right), \quad \omega \in (0, 1)$$

Where $\rho(0) = 1$.

We will evaluate the infinite sum and substitute the result into the equation above. We will re-instate coefficients A_1 and A_2 from the previous assignment during intermediate steps for simplicity. In addition, we will let $d_1 = 1/2$ and $d_2 = 1/5$, noting that the geometric series equation is usable here as $|d_1|$ and $|d_2|$ are both less than 1.

$$\sum_{h=1}^{\infty} \rho(h) \cos(\omega h) = \sum_{h=1}^{\infty} \left(\frac{16}{11} \left(\frac{1}{2} \right)^{|h|} - \frac{5}{11} \left(\frac{1}{5} \right)^{|h|} \right) \cos(\omega h)$$

$$\sum_{h=1}^{\infty} \rho(h) \cos(\omega h) = \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} - A_2(d_2)^{|h|} \right) \cos(\omega h), \quad \text{using variable form.}$$

$$\sum_{h=1}^{\infty} \rho(h) \cos(\omega h) = \underbrace{\sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right)}_{\text{Term 1}} - \underbrace{\sum_{h=1}^{\infty} \left(A_2(d_2)^{|h|} \cos(\omega h) \right)}_{\text{Term 2}}$$

We will evaluate Term 1 and Term 2 separately. We will use the following identities without proof:

$$\cos(\omega h) = \frac{1}{2} \left(e^{ih\omega} + e^{-ih\omega} \right), \quad i = \sqrt{-1}$$
 (1)

$$\sum_{n=1}^{\infty} a \cdot r^n = \frac{ar}{(1-r)}, \quad |r| < 1, \ a \in \mathbb{R}$$
 (2)

Evaluating Term 1, noting that |h| = h since the summation spans $h \in \mathbb{Z}^+$.

$$\begin{split} \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= A_1 \sum_{h=1}^{\infty} (d_1)^{|h|} \left(\frac{1}{2} \left(e^{ih\omega} + e^{-ih\omega} \right) \right), \quad \text{by (1)} \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1}{2} \sum_{h=1}^{\infty} (d_1)^h \left(e^{ih\omega} + e^{-ih\omega} \right) \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1}{2} \left(\sum_{h=1}^{\infty} (d_1)^h e^{ih\omega} + \sum_{h=1}^{\infty} (d_1)^h e^{-ih\omega} \right) \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1}{2} \left(\sum_{h=1}^{\infty} \left(d_1 e^{i\omega} \right)^h + \sum_{h=1}^{\infty} \left(d_1 e^{-i\omega} \right)^h \right) \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1}{2} \left(\frac{d_1 e^{i\omega}}{1 - d_1 e^{i\omega}} + \frac{d_1 e^{-i\omega}}{1 - d_1 e^{-i\omega}} \right), \quad \text{by (2)} \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1}{2} \left(\frac{d_1 e^{i\omega} (1 - d_1 e^{-i\omega}) + d_1 e^{-i\omega} (1 - d_1 e^{i\omega})}{(1 - d_1 e^{i\omega}) (1 - d_1 e^{-i\omega})} \right) \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1}{2} \left(\frac{d_1 (e^{i\omega} + e^{-i\omega}) - 2d_1^2}{1 - d_1 (e^{i\omega} + e^{-i\omega}) + d_1^2} \right) \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1}{2} \left(\frac{2d_1 \cos(\omega) - 2d_1^2}{1 - 2d_1 \cos(\omega) + d_1^2} \right) \\ \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) &= \frac{A_1(d_1 \cos(\omega) - d_1^2)}{1 - 2d_1 \cos(\omega) + d_1^2} \end{split}$$

Similarly, if we repeat this exact same process with A_2 and d_2 , noting that $|d_2| < 1$ and $A_2 \in \mathbb{R}$ also satisfy the requirements of (1) and (2), we arrive at Term 2:

$$\sum_{h=1}^{\infty} \left(A_2(d_2)^{|h|} \cos(\omega h) \right) = \frac{A_2(d_2 \cos(\omega) - d_2^2)}{1 - 2d_2 \cos(\omega) + d_2^2}$$

Then, we can recombine these into our original expression for the infinite sum:

$$\sum_{h=1}^{\infty} \rho(h) \cos(\omega h) = \sum_{h=1}^{\infty} \left(A_1(d_1)^{|h|} \cos(\omega h) \right) - \sum_{h=1}^{\infty} \left(A_2(d_2)^{|h|} \cos(\omega h) \right)$$

$$\sum_{h=1}^{\infty} \rho(h) \cos(\omega h) = \frac{A_1(d_1 \cos(\omega) - d_1^2)}{1 - 2d_1 \cos(\omega) + d_1^2} - \frac{A_2(d_2 \cos(\omega) - d_2^2)}{1 - 2d_2 \cos(\omega) + d_2^2}$$

$$\sum_{h=1}^{\infty} \rho(h) \cos(\omega h) = \left(\frac{16}{11} \right) \frac{\left(\frac{1}{2} \cos(\omega) - \left(\frac{1}{2} \right)^2 \right)}{1 - 2\left(\frac{1}{2} \right) \cos(\omega) + \left(\frac{1}{2} \right)^2} - \left(\frac{5}{11} \right) \frac{\left(\frac{1}{5} \cos(\omega) - \left(\frac{1}{5} \right)^2 \right)}{1 - 2\left(\frac{1}{5} \right) \cos(\omega) + \left(\frac{1}{5} \right)^2}$$

$$\sum_{h=1}^{\infty} \rho(h) \cos(\omega h) = \left(\frac{16}{11} \right) \frac{2 \cos(\omega) - 1}{5 - 4 \cos(\omega)} - \left(\frac{5}{11} \right) \frac{5 \cos(\omega) - 1}{2(13 - 5 \cos(\omega))}$$

Then, combining the expressions we can get the final expression for the normalized spectral density:

$$f^{\star}(\omega) = \frac{1}{\pi} \left(1 + 2 \sum_{h=1}^{\infty} \rho(h) \cos(\omega h) \right)$$

$$f^{\star}(\omega) = \frac{1}{\pi} \left(1 + 2 \left(\left(\frac{16}{11} \right) \frac{2 \cos(\omega) - 1}{5 - 4 \cos(\omega)} - \left(\frac{5}{11} \right) \frac{5 \cos(\omega) - 1}{2(13 - 5 \cos(\omega))} \right) \right)$$

$$f^{\star}(\omega) = \left[\frac{1}{\pi} + \frac{32}{11\pi} \left(\frac{2 \cos(\omega) - 1}{5 - 4 \cos(\omega)} \right) - \frac{5}{11\pi} \left(\frac{5 \cos(\omega) - 1}{13 - 5 \cos(\omega)} \right) \right]$$

We can verify these results below:

```
w = pi/4
# define acf
rho <- function(h){ (16/11)*((1/2)^h) - (5/11)*((1/5)^h) }
# define the "infinite sum"
sum_vals = sum(sapply(1:1000, function(h){
    rho(h)*cos(w*h)
}))
# from the equation...
eqnval = (1/pi)*(rho(0) + 2*sum_vals)

# from our simplification
fw1 = 1/pi
fw2 = (32/(11*pi)) * ( (2 * cos(w) - 1) / (5 - 4*cos(w)) )
fw3 = (5/(11*pi)) * ( (5 * cos(w) - 1) / (13 - 5*cos(w)) )
# comparison
c(eqnval, fw1 + fw2 - fw3)</pre>
```

[1] 0.4561755 0.4561755

We see that the values are identical for at least the first 10,000 lags at fixed $\omega = \pi/4$.

Part B

Write down the power spectral density function of $\{X_t\}_{t\in\mathbb{Z}}$.

Solution

We recall from the definition of normalized spectral density that

$$f^{\star}(\omega) = \frac{f(\omega)}{\sigma_X^2}$$

Where σ_X^2 is the variance of $\{X_t\}_{t\in\mathbb{Z}}$.

Directly, then, we can write $f(\omega)$ as:

$$f(\omega) = \sigma_X^2 f^*(\omega) = \gamma(0) f^*(\omega)$$

We re-establish the Yule-Walker equations, where $\alpha_1 = 7/10$ and $\alpha_2 = -1/10$

$$\mathbb{E}(X_t X_t) = \alpha_1 \mathbb{E}(X_t X_{t-1}) - \alpha_2 \mathbb{E}(X_t X_{t-2}) + \mathbb{E}(X_t Z_t)$$

$$\mathbb{E}(X_t X_{t-1}) = \alpha_1 \mathbb{E}(X_{t-1} X_{t-1}) - \alpha_2 \mathbb{E}(X_{t-1} X_{t-2}) + \mathbb{E}(Z_t X_{t-1})$$

$$\mathbb{E}(X_t X_{t-2}) = \alpha_1 \mathbb{E}(X_{t-1} X_{t-2}) - \alpha_2 \mathbb{E}(X_{t-2} X_{t-2}) + \mathbb{E}(Z_t X_{t-2})$$

Which becomes the following system of three equations:

$$\gamma(0) = \alpha_1 \gamma(1) + \alpha_2 \gamma(2) + \sigma^2$$
$$\gamma(1) = \alpha_1 \gamma(0) + \alpha_2 \gamma(1)$$
$$\gamma(2) = \alpha_1 \gamma(1) + \alpha_2 \gamma(0)$$

test area:

$$\gamma(1) = \alpha_1 \gamma(0) + \alpha_2 \gamma(1)$$
$$\gamma(1) = \frac{\alpha_1}{1 - \alpha_2} \gamma(0)$$

then

$$\begin{split} \gamma(2) &= \alpha_1 \gamma(1) + \alpha_2 \gamma(0) \\ \gamma(2) &= \alpha_1 \left(\frac{\alpha_1}{1 - \alpha_2} \gamma(0) \right) + \alpha_2 \gamma(0) \\ \gamma(2) &= \left(\frac{\alpha_1^2 + \alpha_2 - \alpha_2^2}{1 - \alpha_2} \right) \gamma(0) \end{split}$$

finally

$$\gamma(0) = \alpha_1 \gamma(1) + \alpha_2 \gamma(2) + \sigma^2$$

$$\gamma(0) = \alpha_1 \left(\frac{\alpha_1}{1 - \alpha_2}\right) \gamma(0) + \alpha_2 \left(\frac{\alpha_1^2 + \alpha_2 - \alpha_2^2}{1 - \alpha_2}\right) \gamma(0) + \sigma^2$$

$$\gamma(0) = \left(\frac{\alpha_1^2 + \alpha_1^2 \alpha_2 + \alpha_2^2 - \alpha_2^3}{1 - \alpha_2}\right) \gamma(0) + \sigma^2$$

$$\gamma(0) \left(1 - \frac{\alpha_1^2 + \alpha_1^2 \alpha_2 + \alpha_2^2 - \alpha_2^3}{1 - \alpha_2}\right) = \sigma^2$$

$$\gamma(0) \left(\frac{(1 - \alpha_2) - \alpha_1^2 - \alpha_1^2 \alpha_2 - \alpha_2^2 + \alpha_2^3}{1 - \alpha_2}\right) = \sigma^2$$

$$\gamma(0) = \frac{\sigma^2 (1 - \alpha_2)}{1 - \alpha_2 - \alpha_1^2 - \alpha_1^2 \alpha_2 - \alpha_2^2 + \alpha_2^3}$$

Then, we can evaluate at our given α_1 and α_2 , simplifying the fraction above using Python to avoid human error.

275*sigma_sq/162

Then, we can verify both our results and Python's simplification as follows, assuming $\sigma^2 = 1$.

Computed Variance Using Simplified Fraction: 1.697531

```
cat("Computed Variance Using Computation:", gamma_0, "\n")
```

Computed Variance Using Computation: 1.697531

```
cat("Variance of Simulated AR(2) process:", simulated_gamma_0, "\n")
```

Variance of Simulated AR(2) process: 1.697897

It seems the computation very closely approximates the truth. Hence, we conclude that:

$$f(\omega) = \gamma(0)f^{\star}(\omega) = \frac{275\sigma^{2}}{162} \left(\frac{1}{\pi} + \frac{32}{11\pi} \left(\frac{2\cos(\omega) - 1}{5 - 4\cos(\omega)} \right) - \frac{5}{11\pi} \left(\frac{5\cos(\omega) - 1}{13 - 5\cos(\omega)} \right) \right)$$

Part C

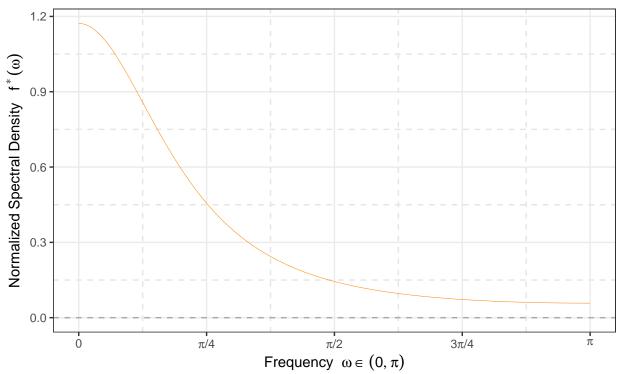
Plot the normalized spectral density and comment on its behaviour.

The normalized spectral density ended up being quite long, so we'll define each term one-by-one in the function below:

```
f = function(w){
  term_1 = 1
  term_2 = (32/11)*((2*cos(w) - 1) / (5 - 4*cos(w)))
  term_3 = (5 /11)*((5*cos(w) - 1) / (13 - 5*cos(w)))
  return ( (1/pi)*(term_1 + term_2 - term_3) )
omega = seq(from = 0, to = pi, length.out = 1e4)
# define data frame for values
p1df = data.frame(omega = omega,
           f_star_omega = f(omega))
# buld plot
p1 <- ggplot(p1df, aes(x = omega, y = f_star_omega)) +
  geom_line(color = "#ff8600", linewidth = 0.1) +
    title = "Normalized Spectral Density for AR(2) Process",
    subtitle = TeX(paste(
      "$X_t = \Lambda_1 X_{t-1} + \Lambda_2 X_{t-2} + Z_t$",
      "where",
      "$\{\,Z_t\,\}_{t \in Z}} \ \\sim \WN(0, \\sigma^2)$\",
      "and \alpha_1 = 7/10, \alpha_2 = -1/10),
    y = TeX("Normalized Spectral Density $f^{*}(\\omega)$"),
    x = TeX("Frequency $\\omega \\in (0, \\pi)$")
  ) + theme bw() +
  geom_hline(yintercept = 0, lty = 'dashed', col = "darkgrey")+
  scale_x_continuous(
    breaks = c(0, pi/4, pi/2, 3*pi/4, pi),
    labels = c(TeX("0"), TeX("$\pi$/4"), TeX("$\pi$/2"),
              TeX("$3\\pi$/4"), TeX("$\\pi$")))+
  theme(panel.grid.minor = element_line(
    color = "grey90",
    linetype = "dashed",
    linewidth = 0.5
  ))
print(p1)
```

Normalized Spectral Density for AR(2) Process

 $X_t = \alpha_1 X_{t-1} + \alpha_2 X_{t-2} + Z_t \text{ where } \{ \ Z_t \ \}_{t \in Z} \sim \text{WN} \big(0, \ \sigma^2 \big) \text{ and } \alpha_1 = 7/10, \ \alpha_2 = -1/10$



Comments: It appears that the normalized spectral density plot is largely dominated by low frequencies. We can tell this is the case due to the fact that the largest values of $f^*(\omega)$ is for small $\omega \in (0, \pi)$ specifically for $\omega \in (0, \pi/4)$ we see the majority of frequencies with observed normalized density greater than 0.5. This tells us that a greater proportion of the variance inherent in the process X_t can be attributed to the lower frequencies. As we would expect of a normalized spectral density dominated by low ω values, we see that $f^*(\omega)$ is strictly decreasing as $\omega \to \pi$.

Question 2