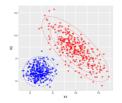
Introduction to Machine Learning

Classification Discriminant Analysis





Learning goals

- Understand the ideas of linear and quadratic discriminant analysis
- Understand how parameteres are estimated for LDA and QDA
- Understand how decision boundaries are computed for LDA and QDA

LINEAR DISCRIMINANT ANALYSIS (LDA)

LDA follows a generative approach

$$\pi_k(\mathbf{x}) = \mathbb{P}(y = k \mid \mathbf{x}) = \frac{\mathbb{P}(\mathbf{x}|y = k)\mathbb{P}(y = k)}{\mathbb{P}(\mathbf{x})} = \frac{p(\mathbf{x}|y = k)\pi_k}{\sum\limits_{j=1}^g p(\mathbf{x}|y = j)\pi_j},$$



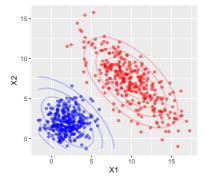
where we now have to pick a distributional form for $p(\mathbf{x}|y=k)$.

LINEAR DISCRIMINANT ANALYSIS (LDA)

LDA assumes that each class density is modeled as a *multivariate Gaussian*:

$$p(\mathbf{x}|y=k) = \frac{1}{(2\pi)^{\frac{\rho}{2}}|\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu_k})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu_k})\right)$$

with equal covariance, i. e. $\Sigma_k = \Sigma \quad \forall k$.





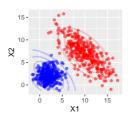
LINEAR DISCRIMINANT ANALYSIS (LDA)

Parameters heta are estimated in a straightforward manner by estimating

$$\hat{\pi_k} = \frac{n_k}{n}$$
, where n_k is the number of class- k observations

$$\hat{\mu_k} = \frac{1}{n_k} \sum_{i:y^{(i)}=k} \mathbf{x}^{(i)}$$

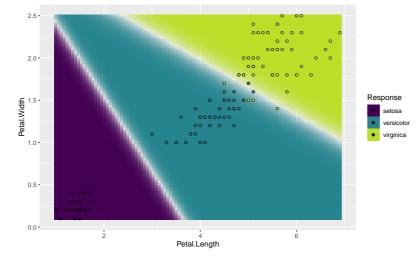
$$\hat{\Sigma} = \frac{1}{n-g} \sum_{k=1}^{g} \sum_{i:v^{(i)}=k} (\mathbf{x}^{(i)} - \hat{\boldsymbol{\mu}_k}) (\mathbf{x}^{(i)} - \hat{\boldsymbol{\mu}_k})^T$$





LDA AS LINEAR CLASSIFIER

Because of the equal covariance structure of all class-specific Gaussian, the decision boundaries of LDA are linear.





LDA AS LINEAR CLASSIFIER

We can formally show that LDA is a linear classifier, by showing that the posterior probabilities can be written as linear scoring functions - up to any isotonic / rank-preserving transformation.

$$\pi_k(\mathbf{x}) = \frac{\pi_k \cdot \rho(\mathbf{x}|y=k)}{\rho(\mathbf{x})} = \frac{\pi_k \cdot \rho(\mathbf{x}|y=k)}{\sum\limits_{j=1}^g \pi_j \cdot \rho(\mathbf{x}|y=j)}$$

As the denominator is the same for all classes we only need to consider

$$\pi_k \cdot p(\mathbf{x}|y=k)$$

and show that this can be written as a linear function of x.



LDA AS LINEAR CLASSIFIER

$$\pi_{k} \cdot p(\mathbf{x}|y = k)$$

$$\propto \qquad \pi_{k} \exp\left(-\frac{1}{2}\mathbf{x}^{T}\Sigma^{-1}\mathbf{x} - \frac{1}{2}\boldsymbol{\mu}_{k}^{T}\Sigma^{-1}\boldsymbol{\mu}_{k} + \mathbf{x}^{T}\Sigma^{-1}\boldsymbol{\mu}_{k}\right)$$

$$= \exp\left(\log \pi_{k} - \frac{1}{2}\boldsymbol{\mu}_{k}^{T}\Sigma^{-1}\boldsymbol{\mu}_{k} + \mathbf{x}^{T}\Sigma^{-1}\boldsymbol{\mu}_{k}\right) \exp\left(-\frac{1}{2}\mathbf{x}^{T}\Sigma^{-1}\mathbf{x}\right)$$

$$= \exp\left(\theta_{0k} + \mathbf{x}^{T}\theta_{k}\right) \exp\left(-\frac{1}{2}\mathbf{x}^{T}\Sigma^{-1}\mathbf{x}\right)$$

$$\propto \exp\left(\theta_{0k} + \mathbf{x}^{T}\theta_{k}\right)$$



by defining $\theta_{0k} := \log \pi_k - \frac{1}{2} \mu_k^T \Sigma^{-1} \mu_k$ and $\theta_k := \Sigma^{-1} \mu_k$.

We have again left out all constants which are the same for all classes k, so the normalizing constant of our Gaussians and $\exp\left(-\frac{1}{2}\mathbf{x}^T\Sigma^{-1}\mathbf{x}\right)$.

By finally taking the log, we can write our transformed scores as linear:

$$f_k(\mathbf{x}) = \boldsymbol{\theta}_{0k} + \mathbf{x}^T \boldsymbol{\theta}_k$$

QDA is a direct generalization of LDA, where the class densities are now Gaussians with unequal covariances Σ_k .

$$p(\mathbf{x}|y=k) = \frac{1}{(2\pi)^{\frac{\rho}{2}} |\Sigma_k|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu_k})^T \Sigma_k^{-1} (\mathbf{x} - \boldsymbol{\mu_k})\right)$$

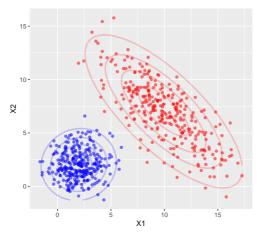


Parameters are estimated in a straightforward manner by:

$$\hat{\pi_k} = \frac{n_k}{n}$$
, where n_k is the number of class- k observations
$$\hat{\mu_k} = \frac{1}{n_k} \sum_{i:y^{(i)}=k} \mathbf{x}^{(i)}$$

$$\hat{\Sigma_k} = \frac{1}{n_k - 1} \sum_{i:y^{(i)}=k} (\mathbf{x}^{(i)} - \hat{\mu_k}) (\mathbf{x}^{(i)} - \hat{\mu_k})^T$$

- Covariance matrices can differ over classes.
- Yields better data fit but also requires estimation of more parameters.





$$\pi_k(\mathbf{x}) \propto \pi_k \cdot p(\mathbf{x}|y=k)$$

$$\propto \pi_k |\Sigma_k|^{-\frac{1}{2}} \exp(-\frac{1}{2}\mathbf{x}^T \Sigma_k^{-1} \mathbf{x} - \frac{1}{2}\boldsymbol{\mu}_k^T \Sigma_k^{-1} \boldsymbol{\mu}_k + \mathbf{x}^T \Sigma_k^{-1} \boldsymbol{\mu}_k)$$

Taking the log of the above, we can define a discriminant function that is quadratic in x.

$$\log \pi_k - \frac{1}{2} \log |\Sigma_k| - \frac{1}{2} \boldsymbol{\mu}_{\boldsymbol{k}}^T \boldsymbol{\Sigma}_k^{-1} \boldsymbol{\mu}_{\boldsymbol{k}} + \boldsymbol{x}^T \boldsymbol{\Sigma}_k^{-1} \boldsymbol{\mu}_{\boldsymbol{k}} - \frac{1}{2} \boldsymbol{x}^T \boldsymbol{\Sigma}_k^{-1} \boldsymbol{x}$$



