PREDICTING TRUCK DRIVERS' CRITICAL EVENTS: EFFICIENT

BAYESIAN HIERARCHICAL MODELS

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Dedication

I dedicate this dissertation to my parents, Zhimin Cai and Guizhen Xu, who believe in the power of higher education, hard work, and always support me.

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INTRODUCTION

1.1 Transportation safety

Traffic safety is a pressing public health issue that involves huge lives losses and financial burden across the world and in the United States (US). As reported by the World Health Organization (WHO 2018b), road injury was the eighth cause of death globally in 2016, killing approximately 1.4 million people, which consisted of about 2.5% of all deaths in the world. If no sustained action is taken, road injuries are predicted to be the seventh leading cause of death across the world by 2030 (WHO 2018a). Compared to the victims who were claimed lives by diseases, people killed in traffic are mostly early- or middle-aged, particularly those aged 4 to 44 years old (Litman 2013; Evans 2014). Without traffic accidents, these victims could have much longer lives with normal health.

Apart from fatal deaths, road traffic injuries were also reported to be the cause of 50 million non-fatal life injuries and approximately 75.5 million disability-adjusted life years globally (Staton et al. 2016). In high-income countries, the majority of non-death costs were attributable to non-fatal crashes, with 2% of non-fatal events leading to over 40% of life-time medical costs (Ameratunga, Hijar, and Norton 2006). Besides non-fatal injuries, traffic safety is a major economic burden. The global economic losses attributable to transportation safety were estimated to be 518 billion the United States Dollars (USD), which accounted for 1%

the gross domestic product (GDP) in low-income countries, 1.5% in middle-income countries, and 2% in high-income countries (Peden et al. 2004; Dalal et al. 2013).

Specifically in the United States, transportation contributed to the highest number of fatal occupational injuries, leading to 2,077 deaths and accounting for over 40% of all fatal occupational injuries in 2017 (The United States, Bureau of Labor Statistics 2017). The National Safety Council reported that the number of deaths attributable to car crashes will be at least 40,000 in 2018, which is the third straight year that this number is over 40,000 (The National Safety Council 2018). A comparison study of 26 developed countries revealed that 20 to 60 traffic deaths per billion kilometers were reduced from 2011 (Evans 2014). Even though fatality rates attributable to road traffic in the US were reduced by 40% during that period, the rates declined more rapidly in all other 25 countries. Given the large amounts of investments in roads, improved vehicle protection and traffic policy implementation, and advanced emergency and trauma care, the reduction in traffic associated fatality rates is nominal (Litman 2013). If the change of traffic fatality rates match those in other unremarkable countries, 20,000 traffic deaths could have been prevented each year (Evans 2014).

The impact of road injuries is even more impactful in developing countries than in developed countries (Goonewardene et al. 2010). Although low- and middle-income countries own only 48% of the registered vehicles in the world, 90% of road traffic fatalities and injuries were estimated to occur in these countries, which continue to be escalating due to rapid urbanization and motorization (Dalal et al. 2013; Staton et al. 2016). Ten developing countries, including Brazil, Cambodia, China, Egypt, India, Kenya, Mexico, Russia, Turkey, and Vietnam, account for almost half of all the road traffic in the world (Hyder et al. 2012). For example, China has around 100,000 traffic-related fatalities each year (Zhang et al. 2010), accounting for around 80% of all accidental deaths, with 87% of them were caused by motor vehicles in 2015 (Jiang et al. 2017). In comparison, 148,707 lives were claimed by road collisions in India in 2015, with the road fatality rate similar to the global average level of 17.4

deaths per 100,000 people (National Crime Records Bureau, Government of India 2015).

1.2   Truck safety

In the US, the large commercial truck industry is the backbone of the economy. Approximately 70% of freight is delivered via a truck at some point of their transportation, which account for 73.1% of value and 71.3% of volume of the domestic goods (Olson et al. 2016; Anderson et al. 2017). However, among all vehicles, large trucks are associated with more catastrophic accidents and therefore are the primary concern of traffic safety. In 2016, the Federal Motor Carrier Safety Administration (FMCSA) reported that 27% fatal crashes in work zones involved large trucks (FMCSA 2018a). Among all 4,079 crashes involving large trucks or buses in 2016, 4,564 people (1.12 people per crash) were killed in the accidents (FMCSA 2016). Large truck crashes approximately claim 5,000 lives and cause 120,000 injuries each year, but only 15% of these fatalities occur in the trucks, with a predominate 78% occurred in the other vehicles (Neeley and Richardson Jr 2009). Besides, the economic losses associated with large truck crashes are also higher than those with passenger vehicles, with an estimated average cost of 91,000 US dollars per crash (Zaloshnja, Miller, and others 2008).

The high risk of large trucks is attributed to two aspects of reasons (Huang et al. 2013). First, large truck drivers generally need to drive alone for long routes, under on-time demands, challenging weather and traffic conditions. Professional truck drivers usually need to work in shifts, and sometimes unavoidable late-night or early-morning shifts (Pylkkönen et al. 2015). These late-night or early-morning working shifts have been reported to be associated with sleep deprivation and disorders (Åkerstedt 1988; Mitler et al. 1997; Solomon et al. 2004; Sallinen et al. 2005). Besides, commercial truck drivers are exposed to long route, constant concentration, and overtime work, which intertwines with sleep deprivation and disorder, and induce the fatigue symptoms among truck drivers. It is estimated that fatigue among long distance truck drivers caused up to 31% of single vehicle fatal truck crashes (National

Transportation Safety Board 1990; Mitler et al. 1997).

On the other hand, trucks have huge weights, large physical dimensions, and potentially carry hazardous cargoes. Although these huge-size trucks boost the transportation efficiency by increasing cargo capacity and reducing fuel costs per trip, they also raise public safety concerns (Lemp, Kockelman, and Unnikrishnan 2011). Large trucks can weight up to 80,000 pounds by federal law, which are twenty times as much as a passenger vehicle (Department of Transportation, Utah 2019). If these trucks travel at the speed of 65 miles per hour on the highway, it will take them 525 feet to stop, which is about two times the length of a football field (Department of Transportation, Utah 2019). The large physical size also creates large blind spots on both sides of the truck, which poses more threat on smaller-sized vehicles. When a crash occurs between a large truck and a smaller vehicle, the sheer size and weight of the truck result in the tragedy that the victims are from the smaller vehicle instead of the trucks in around 80% of the cases (Neeley and Richardson Jr 2009). In even worse case, commercial trucks crashes can cause massive casualties and regional public health emergency when the carried hazardous materials (such as gasoline and sulfuric acid) are leaked.

The importance of truck industry and the potential catastrophic consequences of truck crashes underscore the need to reduce crash risk and improve the safety of truck transportation.

1.3 Crashes and critical events

To reduce the lives and economic losses associated with trucks, numerous studies attempted to screen the risk factors for truck-related traffic crashes and make accurate prediction. However, there are several limitations of studies using crash data. First, traffic crashes are characterized by rare events (dozens to thousands of times fewer crashes than non-crashes) (Theofilatos et al. 2016, 2018). To tackle this rare-event issue, the most common study design is a case-control study that matches a crash with one to up to ten non-crashes, and then use statistical models such as logistic regressions to explain the causes or predict the crashes (Braver et al.

1997; Chen Chen and Xie 2014; Meuleners et al. 2015; Née et al. 2019). Unfortunately, a case-control study is limited in estimating the incidence data or overall average treatment effect. It may be contentious in selecting the ratio of controls to cases and how to select these controls (Grimes and Schulz 2005; Sedgwick 2014). Second, due to the retrospective nature of crash data, it is unrealistic to trace back to the real-time traffic, weather, and other environmental factors that were associated with the crashes. Most of crash data reported by police and associated drivers were subject to recall and misinformation bias (Girotto et al. 2016). Third, crashes are underreported, especially those without injuries or economic losses, as well as those crashes with minor severity (Ye and Lord 2011). The National Highway Traffic Safety Adminitration estimated that 25% of minor-injury crashes and 50% of no-injury crashes were not reported, compared to 100% reporting rate for fatal crashes (Savolainen et al. 2011).

Past truck safety literature almost exclusively focused on crashes, while ignoring the precursors to crashes. A precursor to crashes, also known as a critical events, adverse events, or near-miss crashes, is an emerging pattern or signature associated with an increasing chance of truck crash (Saleh et al. 2013; Janakiraman, Matthews, and Oza 2016). Truck critical events deserve more attention since they occur more frequently than crashes, suggest fatigue and a lapse in performance, and they can lead to giant crashes (Dingus, Neale, et al. 2006).

With the rapid development of modern technology, real-time driving and critical events data are recorded more and more by truck companies (Janakiraman, Matthews, and Oza 2016). Although critical events do not always result in an accident, they could potentially be used as an early warning system to mitigate or prevent truck crashes (Kusano and Gabler 2012; Rome et al. 2018).

1.4 Proposal

I propose using real-time truck ping data provided by the J.B. Hunt, real-time weather and traffic data to

1) quantify the association between truck crashes and critical events;

2) construct predictive models for truck critical events;

3) establish scalable Bayesian hierarchical models for large truck data.

I believe that this work will contribute to statistical theories in constructing scalable Bayesian hierarchical statistical and reliability models using modern Markov chain Monte Carlo simulations. Realistically, these predictive models will inform policy-makers the functional relationship between driver characteristics, traffic, weather, and other real-time driving environment. These predictive models can be further used to provide data-driven justification to optimize trucking routes and minimize unsafe driving behaviors.

Chapter 2

LITERATURE REVIEW

2.1   Precursors to crashes

Truck crash data have several limitations. Firstly, truck crashes are extremely rare compared with non-crashes. According to the FMCSA (2018b), large truck and bus fatalities in 2017 were 0.156 per million travelled vehicle miles, which was a 6.8 percent increase from 2016. This rareness poses a challenge to infer unbiased estimates using traditional statistical models (Guo et al. 2010; Theofilatos et al. 2018). Secondly, most truck crash data reply on post hoc police reports. Although these data are generally accurate and detailed by police officers, they are limited in determining the information of the driver in the meaningful time period leading up to the crash (Dingus, Hanowski, and Klauer 2011). Some of the critical factors, such as distraction, are not reported or cannot be determined due to a variety of reasons (Dingus, Hanowski, and Klauer 2011). Thirdly, truck crashes are underreported, particularly for no-injury and minor-injury crashes (Ye and Lord 2011; Stern et al. 2019). It is estimated that 25% of minor-injury and 50% of non-injury crashes were not reported, while 100% of fatal crashes were reported (Savolainen et al. 2011).

In light of the aforementioned limitations, a growing number of literature is using natu-ralistic driving data to identify crash causation and develop effective crash countermeasures (Dingus, Klauer, et al. 2006; Klauer et al. 2009; Guo et al. 2010; Hickman, Hanowski,

and Bocanegra 2010). These studies use unobstrusive devices installed on vehicles to collect real-time driving behavior and and performance. Instead of collecting extremely rare vehicle crash data, these studies focus on near-crash events, defined as events that used last-second successful evasive maneuver that avoided crashes (Dingus, Hanowski, and Klauer 2011). Although these near-crashes were not real crashes, these studies suggested that they are predictive crashes (Dingus, Hanowski, and Klauer 2011; Guo et al. 2010).

A more formal definition of near-crash, or the more general accident precursor, in safety analysis and accident prevention field was proposed in Saleh et al. (2013). An accident precursor was defined as a chain of adverse events following an initial off-nominal event, which can result in an accident if compounded with additional adverse conditions (Saleh et al. 2013). A near-crash or near miss is a special case of accident precursor, with the feature of being close to a complete accident sequence. Accidence precursor has been widely studies in certain safety science in which the accidents are extremely rare, such as nuclear industry (Smith and Borgonovo 2007), chemistry (Phimister et al. 2003), and aerospace industry (Kirwan, Gibson, and Hickling 2008). However, there are limited number of studies in accident precursors in the truck safety field.

2.2 Naturalistic driving data

The first large-scale NDS was The 100-Car Naturalistic Driving study conducted in North Virginia, the United States (Neale et al. 2005).

Knipling (2015) challenges the validity of naturalistic driving data.

2.3 Risk factors

2.3.1 Fatigue

Among all driver-related safety events, fatigue has become the most pressing problem of truck accidents. It is estimated by National Sleep Foundation that approximately 32% of drivers drive with fatigue over twice a month (National Sleep Foundation 2008). The American Automobile Association Foundation for Traffic Safety claimed that 16.5% of fatal traffic accidents and 12.5% of injuries-related collisions were associated with driving with fatigue in 2010 (American Automobile Association Foundation for Traffic Safety 2010). The National Highway Traffic Safety Administration (NHTSA) estimated that 60% of fatal truck crashes were attributable to the driver falling asleep while driving (Craye et al. 2016; Cavuoto and Megahed 2017).

Fatigue is often defined as a multidimensional process that leads to diminished worker performance, which may be a result of prolonged work, psychological, socioeconomic, and environment factors (Yung 2016; Cavuoto and Megahed 2017). However, this definition has low specificity since there are other factors associated with a decreased worker performance, such as cell phone use, which does not result in driver fatigue. There is no uniform and succint definition on fatigue since it involves interactions between biological, behavior, and psychological process. A comprehensive review on fatigue definition and measurement is provided by Yung (2016).

The mechanism of fatigue leading to traffic safety events is that the driver's capability to stay alert to ambient traffic and pedestrians will be largely impaired. The reaction time is subsequently prolonged in that situation (Zhang et al. 2014). It is estimated that 17 hours of continuous working lead to a deterioration of driving performance equivalent to a blood alcohol level of 0.05% (MacLean, Davies, and Thiele 2003). What makes the outcomes worse is that fatigue driving is more likely to happen on expressways and major highways where the speed limit is over 55 miles per hour (Knipling and Wang 1994). This is especially concerning

because fatigue driving safety critical events are more likely to result in serious injuries and fatalities, compared with non-fatigue driving safety critical events.

Although fatigue has been recognized as the primary reason for traffic safety, little has been done about drowsy driving since there is no simple way to objectively measure fatigue driving (Dement 1997). In view of the difficulty of measuring fatigue, researchers have attempted to use different proxies of fatigue associated with truck drivers. Stern et al. (2019) reviewed the research related to the fatigue of commercial motor vehicle drivers.

Numerous studies used ocular metrics, such as eye closure and blinking patterns, to detect driver's fatigue (Golias and Mishra 2013). For example, Jackson et al. (2016) measured the eyelid closure of 22 healthy participants with 24 hours of sleep deprivation. They found a significant increase in proportions of time with closed eyelids after the sleep deprivation, which were significantly associated with crashes, Psychomotor Vigilance Task lapses, and subjective sleepiness scores (95% confidence interval of Spearman's correlation coefficient was 0.46 to 0.69).

Other studies used physiological measures, such as electroencephalogram (EEG) and heart rate, as indicators of fatigue (Golias and Mishra 2013). It has been demonstrated that awake but drowsy truck drivers during working have increased alpha and theta power densities in EEG (Kecklund and Åkerstedt 1993; Cajochen et al. 1995; Mitler et al. 1997). The variability of heart rates has been identified as passive proxy of fatigue among drivers (Mulder 1992; Patel et al. 2011).

Working load, such as cumulative driving time, is also measure of the fatigue of drivers. For example, Nakayama (2002) found that there was a significant increase in the fatigue of drivers after 12 hours of continuous driving. Jovanis, Wu, and Chen (2011) used cumulative hours of driving, time of the day, driving patterns over multiple days, rests after driving, and the 34-hour recovery policy as measures of driver fatigue. They found that more driving time was associated with increased odds of crashes among 686 less-than-truck-load drivers (224 crash-involved drivers and 464 randomly sampled non-crash drivers from the same terminal

in the same month), with the highest odds in the 11-th hour. From the fifth hour to the 11th hour, the odds of crashes were consistently increasing (Jovanis, Wu, and Chen 2012). In contrast, Soccolich et al. (2013) found no significant difference in safety outcomes between the 11th driving hour and driving hours 8, 9, or 10 using from the Naturalistic Truck Driving Study data. Desipte the fact that working hours were not significant, they suggested an interaction between driving hours and working hours: a work day that starts with several hours of non-driving work and then followed by 14 hours of driving was significantly associated with risk of safety-critical events.

Lack of sleep or specific sleep patterns have also been used as proxies of fatigue. For example, G. X. Chen et al. (2016) used negative binomial regression to identify the association between four sleep patterns and driving performance based on the Naturalistic Truck Driving Study data. They revealed that shorter sleep, early-stage sleep in a non-work period, and insufficient sleep between 1 a.m. and 5 a.m. were associated with increased safety-critical event rates. Sparrow et al. (2016) reported that truck drivers with a retart break of only one nighttime period (defined as 1 a.m. to 5 a.m.) experienced more lapses of attention, elevated lane deviation at night, and higher sleepiness measured by subjective questionnaires.

A significant amount of research emphasizes the association between time of the day (such as night driving) and the development of fatigue(Cavuoto and Megahed 2017). Night drving is often accompanied by changes in shift scheduling, inadequate sleep, sleep apnea and disorder. For example, Pack et al. (1995) reported that the crashes in which the drivers fell asleep occurred primarily from mid-night to 7 a.m. and from 2 p.m. to 4 p.m.. Mitler et al. (1997) monitored 80 truck drivers in North America using 24-hour electrophysiologic measures. They found that drivers an average had 5.18 hours of sleep in bed per day and 4.78 hours of electrophysiologically validated sleep per day, which were significantly less than needed to stay alert on job. It was also suggested that late-night or early-morning work were detrimental to the drivers' sleep. Pahukula, Hernandez, and Unnikrishnan (2015) investigated the risk factors associated with crashes on Texas urban freeways between 2006

and 2010. They ran separate random-effects logistic regressions on five time periods: early morning (12 a.m. to 4 a.m.), morning (5 a.m. to 9 a.m.), mid-day (10 a.m. to 3 p.m.), afternoon (4 p.m. to 8 p.m.), and evening (9 p.m. to 11 p.m.). The results revealed major differences in traffic flow, light conditions, surface conditions, time of year, and the percentage of trucks on road among the five different models.

2.3.2 Driver characteristics

A study by Pack et al. (1995) revealed that the drivers were 25 years of age or younger in over a half of the 5,000 highway crashes in which they fell asleep whiling driving. Age-related safety (Duke, Guest, and Boggess 2010). Teenager driver (Ouimet et al. 2014). Another driver-related risk factor of driving safety critical events is the age of drivers. It is reported that younger drivers will experience more crashes than older drivers (Cantor et al. 2010).

On the other hand, older drivers can also experience an increased risk of crash. To meet the huge demand services and supply chain management, it is very common to extend the retirement age or reemploy retired workers in developing countries (Popkin et al. 2008). Aging drivers increase the chance of the safety critical events in three aspects: impaired eyesight, prolonged reaction time to exogenous stimuli, and vulnerability to fatigue (Di Milia et al. 2011). Aged drivers are associated with eyesight diseases or functionality impairment, such as cataracts, narrowed peripheral vision and decreasing visual acuity (Di Milia et al. 2011). In addition, working for truck companies often means irregular shifts and taking the night schedules, which disrupt the circadian time-keeping systems, especially for the aged workers (Moneta et al. 1996).

Aged drivers may find it much more difficult to adjust for the sleep-wake cycle to keep pace with the schedule required by the employer company. Therefore, this disruption of the circadian systems, in turn, increases the chances to feel sleepy or fatigue for workers. It is indicated by research that the "critical age" of shiftwork intolerance is about 45 to 50 years, at which sleep disorder, persisting fatigue and digestive problems become the most obvious

(Di Milia et al. 2011). Young drivers are much better in the sense of physical health and resistance to fatigue compared with aged drivers, however, they are more vulnerable regarding the experience of driving. A study conducted by Clarke suggested that young drivers (17 – 19 years old), especially males, have significantly more accidents than other drivers during the hours of darkness, on rural curves, and rear-end shunts compared with male drivers aged 20 -25 years (Clarke et al. 2006). The reasons for these young driver accidents were not fully explained, but could largely be attributed to inexperience.

One more risk factor that could explain driving safety critical events is drivers' gender. Gender has been suggested to be related with outcomes in medical treatment, education, sports and other fields, and there is no exception for truck drivers' safety. In the first place, women are more likely to suffer from fatigue compared with men. A study found that women in general have 1.4 times higher chance of complaining of fatigue than men (Fjell et al. 2008). However, females are found to have longer sleeping hours than their male counterparts of the same race (Lauderdale et al. 2006). In that study, it was found that the mean sleep hours for white females was 6.7 hours compared with 6.1 hours for white males, and 5.9 hours for black female compared with 5.1 hours for black males even after adjusting for socioeconomic status, lifestyle and sleep apnea (Lauderdale et al. 2006). Gender differences are huge in terms of working conditions. Females had significantly fewer working hours per week, with 47 hours versus 52 hours per week (Rotenberg et al. 2008). In general, women tend to work fewer hours within a week but are more prone to feel fatigue and have a higher risk of traffic incidences.

2.3.3 Traffic

Although it is generally believed that lower **speed limits** can reduce the chances of traffic crashes, studies on speed limits and traffic fatalities showed inconsistent results (Neeley and Richardson Jr 2009; Korkut, Ishak, and Wolshon 2010; Zhu and Srinivasan 2011; Davis et al. 2015).

Speed (Tseng et al. 2016).

For example, based on multinomial logit and negative binomial models, Dong et al. (2017) used vehicle and driver characteristics, traffic, environment, and road geometry to predict the frequency and severity of large truck-involved crashes.

Liu, Boyle, and Banerjee (2018) used artificial neural network classification and regression tree, and support vector machine to predict crash rate levels of motor carriers.

2.3.4 Weather

(Naik et al. 2016)

Weather has both direct and indirect effects on drivers' safety critical events. On one hand, the increase of ambient temperature places risks on drivers' occupational safety, and possibly leads to cognition loss, heat stroke, and impairment of wakefulness. Evidence showed that the risk of mistakes and safety critical events increase in hot weather (Kjellstrom et al. 2009; Basagaña et al. 2015). Leard and Roth found that for a day with temperature above 80F there is a 9.5% increase in fatality rates compared with a day at 50-60 F (Leard, Roth, and others 2015). A literature review found that 11 out of 13 studies indicated an increase in unintentional injuries associated with high temperatures (Kampe, Kovats, and Hajat 2016). On the other hand, real-time extreme weather conditions such as heavy rain, fog, storm, and snow can either impair the driver's visual capability or reduce the safety of driving on the road (Chang and Chen 2005; Al-Ghamdi 2007; Baker and Reynolds 1992). It is to noted that the cumulative time of driving in such extreme weather conditions could increase the chances of safety critical events. Studies that explore the association between precipitation and driving safety critical events consistently find a negative relationship. The positive linear relationship between precipitation and traffic accidents can be observed in both driver accidents and pedestrian accidents (Al-Ghamdi 2007; Graham and Glaister 2003). Abdel-Aty et al. used detector and sensor data to successfully predict more than 70% of accidents with low visibility conditions (Abdel-Aty et al. 2012). The common problem for

the literature exploring the relationship between ambient weather and safety driving critical events is the failure to include the cumulative effect of weather conditions. Instead, they all use an indicator variable to represent whether extreme weather happened during the trip or not, which could lead to potential bias in prediction models.

Ghasemzadeh et al. (2018)

2.3.5   Road Geometry

Dong et al. (2015)

Islam, Jones, and Dye (2014) did a comprehensive analysis of truck at-fault crashes.

Uddin and Huynh (2017) lighting conditions.

Meuleners et al. (2017) determinants of truck crashes in Australia.

2.4   Predictive models

2.4.1   Overview

The most commonly used statistical models for traffic safety are logistic regression and Poisson regression. Logistic regression is commonly used to predict crash likelihood (probability) using real-time data, for example traffic and weather at 5-minute intervals (Wang, Abdel-Aty, and Lee 2017). In contrast, Poisson regression is used to predict the crash frequency (the number of crashes) within a time period using aggregated data such as average data traffic and precipitation. I will briefly introduce the two models and then compare the two cultures of predictive models in statistical and machine learning perspective.

The parameterization of a binary logistic regression is shown in Model (2.1).

$$Y_i \sim \text{Bernoulli}(p_i)$$
$$\log(\frac{p_i}{1 - p_i}) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_k x_k$$

(2.1)

Where Y_i is a binary variable that indicates whether an event occurred or not in the i-th observation. p_i is the mean parameter of a Bernoulli distribution, which is constrained on $[0, 1]$. The logit transformation of p_i then has the range from $-\infty$ to $+\infty$, which equals a linear combination of the predictors $x_1, x_2, \cdots, x_k$ and associated parameters $\beta_0, \beta_1, \cdots, \beta_k$.

The most commonly used outcomes for binary logistic regressions are injury versus non-injury crashes or fatal versus non-fatal crashes (Savolainen et al. 2011). For example, Cong Chen et al. (2016) used a two-level hierarchical Bayesian logistic model to predict the likelihood of high-severity crashes compared to low-severity crashes in New Mexico, accounting for both crash-level and driver-level effects. They found that road curve, functional and disabled vehicle damage, single-vehicle crashes, female, older drivers, drug or alcohol involvement were associated with increased odds of severe crashes. Considering the rare-event natural of crashes, Theofilatos et al. (2016) used logistic regression with rare events bias correction and Firth method to study significant risk factors for crashes in Greece. They found a negative association between crash likelihood and speed in crash locations. The proportion of trucks on the road was included in their model but not found to be significant. Other traffic safety studies using logistic regressions include but were not limit to Moudon et al. (2011), Meuleners et al. (2017), Ahmed et al. (2018). There are two excellent systematic reviews on traffic crash likelihood predictions by Roshandel, Zheng, and Washington (2015) and Xu et al. (2015).

Other variants of a binary logistic regression are binary probit models (Lee and Abdel-Aty 2008; Yu and Abdel-Aty 2014), ordered logistic or probit models (Xie, Zhang, and Liang 2009; Zhu and Srinivasan 2011), multinomial logit models (Ye and Lord 2011). There are only minor difference between a probit model and a logistic model. A logistic model uses the inverse logit of the linear predictors to calculate the probability of an event, as shown in Equation (2.2); a probit model uses the cumulative normal density function of the linear predictors to calculate the probability, as shown in Equation (2.3). The error function $\mathrm{erf}(x)$

is an integral without an analytical solution: $\text{erf}(x) = \frac{2}{\pi} \int_0^x e^{-t^2} dt$.

$$p = \text{logit}^{-1}(\mathbf{X}'\beta) = \frac{\exp(\mathbf{X}'\beta)}{1 + \exp(\mathbf{X}'\beta)} \tag{2.2}$$

$$p = \Phi(\mathbf{X}'\beta) = \frac{1}{2}\left[1 + \text{erf}(\frac{\mathbf{X}'\beta}{\sqrt{2}})\right] \tag{2.3}$$

Ordered logistic or probit regressions aim to model an ordered multi-category outcome variable. The most common case is study the severity of crashes, such as no-injury crashes, minor-injury crashes, and fatal-injury crashes (Zhu and Srinivasan 2011). These ordered models account for the ranked nature of different severity levels but make the proportional odds assumption (Rifaat, Tay, and De Barros 2012). When the proportional odds assumption is violated, researchers often switch to multinomial logit or probit models, in which the outcome variable is deemed as nominal.

On the other hand, the parameterization of a Poisson regression is shown in model (2.4).

$$Y_i^{\star} \sim \text{Poisson}(\mu_i)$$
$$\log \mu_i = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_k x_k \tag{2.4}$$

Where $Y_i^{\star}$ is the number of events in the i-th observation, which must be a non-negative integer. μ_i is the mean and variance of the Poisson distribution, and it must be a non-negative numeric value. The logarithm of μ_i transforms μ_i into the range of $(-\infty, +\infty)$, which equals a linear combination of the predictors $x_1, x_2, \cdots, x_k$ and associated parameters $\beta_0, \beta_1, \cdots, \beta_k$. Note that the mean parameter equals the variance parameter in the Poisson distribution, which is often violated in real-life data. When the variance of the data is greater than expected, it is called overdispersion. Otherwise, it is called underdispersion. Overdispersion is much more common than underdispersion in statistical practice.

When researchers have crash data that are aggregated over a long time period such as one

year, it often makes sense to study the number of crashes instead of whether a crash occurred or not since they are often more than one crash. The most commonly used statistical model is therefore Poisson model, as it well handles count data that are right-skewed, long tailed, and only have non-negative integer values. For example, Cantor et al. (2010) used Poisson regressions to explore the association between the rate of crashes driver-level characteristics among 560,695 commercial truck drivers in the United States. They found that past safety performance, out-of-service rate, body mass index (BMI), age, and the number of unique companies were strong predictors of the rate of truck crashes.

Other variants of a Poisson model includes negative binomial models, quasi-Poisson models, and zero-inflated Poisson or negative binomial models (Lord 2006; Mohammadi, Samaranayake, and Bham 2014). Negative binomial or quasi-Poisson models are developed to account for the overdispersion and underdispersion in count data, for which a Poisson model fails to account. Zero-inflated Poisson or negative binomial models are developed to account for the feature of rare events in traffic crash data (Lord, Washington, and Ivan 2005, 2007; Washington, Karlaftis, and Mannering 2010; Dong et al. 2014). There is an excellent review paper on statistical models for crash frequency data by Lord and Mannering (2010).

There are two cultures in current statistical or data science field, explanation and prediction (Shmueli and others 2010; Breiman and others 2001). The pro-explanation culture has long been adopted by most desciplines, such as epidemiology, economics, and psychology. In these desciplines, researchers commonly use generalized linear models, such as logistic regression and Poisson regression, to explain the association between the outcome and predictor variables. In contrast, the pro-prediction culture has recently been adopted in data science disciplines, in which they use blackbox algorithms such as random forests, decision trees, and neural networks to achieve similarly high prediction accuracy in training and testing sets. Pro-explanation models tend to excel at explaining the association between predictors and the outcome variable and being less likely to overfit the data. However, compared with machine learning and deep learning algorithms, pro-explanation models are less likely to cap-

ture potential interaction between predictor variables since they are conceptual framework driven. Therefore, pro-explanation models generally have less prediction accuracy compared with black-box algorithms.

2.4.2 Bayesian models

In contrast to traditional frequentist models that view parameters as unknown but fixed values, Bayesian models view parameters as random variables that have probability distributions (Gelman et al. 2013). Researchers have subjective prior beliefs (a probability distribution) on these parameters $p(\theta)$ before they collect any data. After observing the data $\mathbf{X}$, the researchers could change their prior beliefs. Therefore, the posterior distribution $p(\theta|\mathbf{X})$ is an an unconditional distribution that is a compromise between the prior beliefs and the data. This compromise is given analytically by the Bayes Theorem (Equation (2.5)).

$$
\begin{aligned}
p(\theta|\mathbf{X}) &= \frac{p(\theta)p(\mathbf{X}|\theta)}{p(\mathbf{X})} \\
&= \frac{p(\theta)p(\mathbf{X}|\theta)}{\int p(\theta)p(\mathbf{X}|\theta)d\theta}
\end{aligned}
\tag{2.5}
$$

Where $p(\mathbf{X}|\theta)$ is the likelihood function, which reflects the data generating process that gives rise to the data observed. The denominator $\int p(\theta)p(\mathbf{X}|\theta)d\theta$ is a normalizing constant that makes the posterior distribution integrates to one. The prior and likelihood function are straightforward since they both have analytical forms. The trickiest part of Bayesian inference is the normalizing constant in the denominator (Gelman et al. 2013; Kruschke 2014).

The normalizing constant need to make the posterior distribution integrate to one since the posterior is supposed to be a probability density distribution. When there are more than two parameters in the model, the normalizing constant often becomes intractable since it involves integration in multiple dimensions. Modern Bayesian inference often uses numerical methods such as Markov chain Monte Carlo (MCMC) methods to directly sample from this

posterior distribution, or the integrated Laplace approximation to approximate this constant. However, this numerical methods often fail or take an inhibitively long time to solve the problem with the presence of high-dimensional data or very tall data in this era.

There are several strengths of Bayesian models over traditional Frequentist models. First, the probabilistic distribution of parameters, posterior credibles intervals, and posterior predictive distributions account for the uncertainty in parameters and the data generating process. They also have straightforward and intuitive interpretations. Second, Bayesian models incorporate prior information $p(\theta)$ into the statistical model, which can be useful when there is sufficient prior background information. This prior distribution (regularizing priors) is particularly useful for estimation in high-dimensional, sparse data settings, and complex statistical models such as hierarchical models (Betancourt and Girolami 2015; McElreath 2018). Lastly, Bayesian models are scalable to complex data generating process. This is because modern Bayesian estimation is powered by numerical methods and simulation, which in essence only requires researchers to specify the priors and likelihood function. The difficulty of written the likelihood function is minimal compared to traditional Frequentist approaches such as the maximum likelihood estimation, which scales with the complexity of models (Lambert 2018).

2.4.3 Hierarchical models

Most studies on traffic safety assume that the sampling unit is a spatial-temporal segment, which is a specific section of a road with relatively high rate of crashes during a period. However, it is not sufficient to only study the occasions where the crashes are more likely to occur; we must also study the non-crashes and compare them with crashes. On the other hand, these studies that focus on road segments ignore driver-level unobserved effects. It is reported that the chance of having crashes for truck drivers with crash history in the past year is nearly twice as high as those without crash history in the past year (Cantor et al. 2010). Most motor carrier insurance companies and employers also view historical safety

events as an important measure of the driver's performance. Therefore, it is more natural to use driver-focused models to account for unobserved variation and characteristics associated with vehicle drivers (Huang and Abdel-Aty 2010).

In the Bayesian perspective, a hierarchical model is a statistical model with the probability distribution of one parameter depends on another parameter (Kruschke and Vanpaemel 2015). Suppose we have a model with two parameters α, β and data D. The joint prior distribution of the two parameters is $p(\alpha, \beta)$. According to the Bayes Theorem, the posterior distribution is proportional to the product of the prior distribution and the likelihood function: $P(\alpha, \beta | D) \propto P(\alpha, \beta) P(D | \alpha, \beta)$. In a hierarchical model setting, the product can be factored as a chain of products among parameters, also known as conditional independence, such as $P(\alpha, \beta) P(D | \alpha, \beta) = P(D | \beta) P(\beta | \alpha) P(\alpha)$. In this parameterization, the parameter α is known as the hyperparameter because it gives rise to the parameter β (the parameter of a parameter) (Kruschke and Vanpaemel 2015).

Model (2.6) demonstrates a random-intercept hierarchical logitistic regression that predicts the likelihood of safety events. The outcome $Y_{i,d(i)}$ is a binary variable that indicates whether a safety event occurred or not, and it has a Bernoulli distribution with the mean parameter $p_{i,d(i)}$. The logit transformation of $p_{i,d(i)}$ can then be predicted by k variables $x_1, x_2, \cdots, x_k$. The random intercept $\beta_{0,d(i)}$ determines that this is hierarchical model since they vary across different drivers $d(i)$. This model assumes that these random intercepts are sampled from a population of drivers with the mean of μ_0 and standard deviation of σ_0, which are known as hyperparameters.

$$Y_{i,d(i)} \sim \text{Bernoulli}(p_{i,d(i)})$$
$$\log \frac{p_{i,d(i)}}{1 - p_{i,d(i)}} = \beta_{0,d(i)} + \beta_1 x_{1,i} + \beta_2 x_{2,i} + \cdots + \beta_k x_{k,i} \tag{2.6}$$
$$\beta_{0,d(i)} \sim N(\mu_0, \sigma_0^2), \quad k = 1, 2, \cdots, D$$

Compared with traditional fixed-effects models that either pool all groups of data or estimate separate models individually for each group, a hierarchical model has the advantage of

partial pooling across different groups (McElreath 2018). This partial pooling shrinks group-level parameter estimates towards the group mean and shares information across groups. Therefore, with reasonable assumptions on the data generating process, estimates from a hierarchical model are generally more robust to extreme observations and reasonably accurate for those groups with sparse data (Gelman and Hill 2006; Lambert 2018).

Hierarchical models also come with costs. They are particularly known for its complexity to estimate to coefficients in both Frequetist maximum likelihood and Bayesian estimation. The de facto way of current Bayesian estimation is Markov chain Monte Carlo (MCMC). However, in the hierarchical model setting, it is difficult for MCMC to efficiently sample from the posterior distributions of hyperparameters due to the correlation between different levels of parameters, as well as the large number of parameters created by the hierarchical structure.

2.4.4 Markov chain Monte Carlo (MCMC)

In modern statistics, Bayesian inference almost indispensably relies on Markov chain Monte Carlo (MCMC) sampling to overcome the intractable denominator in the Bayes Theroem (Equation (2.5)). A **Monte Carlo simulation** is a technique to understand a target distribution by generating a large amount of random values from that distribution (Kruschke 2014). A **Markov chain** has the property that the probability distribution of the observation i only depends on the previous observation $i - 1$, not on any one prior to observation $i - 1$, as demonstrated in Equation (2.7).

$$P\left(X_i = x_i | X_1 = x_1, X_2 = x_2, \ldots, X_{i-1} = x_{i-1}\right) = P\left(X_i = x_i | X_{i-1} = x_{i-1}\right) \qquad (2.7)$$

Integrating Markov chains and Monte Carlo simulations, the MCMC method can characterize an unknown unconditional distribution without knowing its all mathematical properties by sampling from the distribution (Van Ravenzwaaij, Cassey, and Brown 2018). It has

been widely applied in multiple fields such as statistics, physics, chemistry, and computer science (Craiu and Rosenthal 2014). The most notable application of MCMC is probably in Bayesian inference, in which it has been used to draw samples from the posterior distribution and calculate relevant statistics (such as mean, standard deviation, and intervals).

The first proposal of using MCMC dates to the paper by Metropolis et al. (1953), in which they tried to solve an intractable integral with a random walk MCMC. The Metropolis algorithm starts with a randomly defined initial value of the parameter θ. From a pre-defined symmetric proposal probability distribution $p(\theta|\mathbf{x})$, it then draw a proposal parameter value $\theta^{(\text{prop})}$, which only depends on the current parameter value $\theta^{(t)}$. This proposal value will be accepted with the probability of α defined in Equation (2.8).

$$\alpha = \min \left(1, \frac{p(\theta^{(\text{prop})}|\mathbf{x})}{p(\theta^{(t)}|\mathbf{x})} \right) \tag{2.8}$$

This proposal and acceptance with probability steps will be iterated for a pre-define number of times. When the Metropolis algorithm reaches a steady state, these proposal values are random values drawn from the posterior distribution of parameter θ, which can be used to describe and characterize the posterior distribution.

After decades of successful empirical trials in physics, Hastings (1970) proposed a more generalized form of the Metropolis algorithm, in which the proposal distribution can be arbitrary, but the acceptance probability $\alpha^\star$ is modified as shown in Equation (2.9). This Metropolis-Hasting (MH) algorithm is the most classic and widely-known MCMC algorithm used in multiple fields.

Let $p(\theta|\mathbf{X})$ be the posterior distribution we want to know, then the *Metropolis-Hasting algorithm* is:

1. Let $\theta^{(1)}$ denote an initial value for the continuous state Markov chain,

2. Set $t = 1$,

3. Let q be the proposal density which can depend on the current state $\theta^{(t)}$. Simulate one

observation $\theta^{(\text{prop})}$ from $q(\theta^{(\text{prop})}|\theta^{(t)})$,

4. Compute the following probability:

$$\alpha^{\star} = \min\left(1, \frac{p\left(\theta^{(\text{prop})}|\boldsymbol{x}\right)}{p\left(\theta^{(t)}|\boldsymbol{x}\right)} \frac{q\left(\theta^{(t)}|\theta^{(\text{prop})}\right)}{q\left(\theta^{(\text{prop})}|\theta^{(t)}\right)}\right) \tag{2.9}$$

5. Set $\theta^{(t+1)} = \theta^{(\text{prop})}$ with the probability of $\alpha^{\star}$; otherwise set $\theta^{(t+1)} = \theta^{(t)}$. Set $t \leftarrow t+1$ and return to 3 until the desired number of iterations is reached.

Although the M-H algorithm is simple and powerful for performing MCMC, its performance highly depends on the choice of the proposal distribution. When there are a few parameters in the model and the proposal distribution is not well-designed, the M-H algorithm will have a very low acceptance rate, which makes the M-H algorithm very inefficient. In view of this issue, Gibbs sampler was proposed with the idea that the proposed values are always accepted and each parameter is updated one at a time by generating samples from the conditional distributions (Geman and Geman 1987; Gelfand and Smith 1990; Lambert 2018). The development of the software *Bayesian inference Using Gibbs Sampler (BUGS)* (Lunn et al. 2000, 2009) was critical in increasing the popularity of applied Bayesian analyses considering its support for a wide variety of statistical distributions, automatic application of the Gibbs Sampler, and numerous textbooks, tutorials and discussion.

Suppose $\theta = [\theta_1, \theta_2, \cdots, \theta_k]$ is a k-dimensional parameter. Let $\mathbf{X}$ denote the data. The *Gibss sampling* algorithm is then:

1. Begin with an estimate $\theta^{(0)} = [\theta_1^{(0)}, \theta_2^{(0)}, \cdots, \theta_k^{(0)}]$ in the parameter space,

2. Set $t = 1$,

3. Simulate $\theta_1^{(t)}$ from $p(\theta_1|\theta_2^{(t-1)}, \theta_3^{(t-1)}, \cdots, \theta_k^{(t-1)}, \mathbf{X})$,

4. Simulate $\theta_2^{(t)}$ from $p(\theta_2|\theta_1^{(t)}, \theta_3^{(t-1)}, \cdots, \theta_k^{(t-1)}, \mathbf{X})$,

5. $\cdots$,

6. Simulate $\theta_k^{(t)}$ from $p(\theta_k|\theta_1^{(t)}, \theta_3^{(t)}, \cdots, \theta_{k-1}^{(t)}, \mathbf{X})$,

7. Set $t \leftarrow t+1$ and repeat steps $3-6$ for a pre-specified number of iterations and make

sure the Gibbs sampler reaches the steady state for a sufficient number of iterations.

The generality of the M-H algorithm and Gibbs sampler and the simplicity in software packages in `R` or `BUGS` help them gain popularity among applied researchers in the recent 30 years. However, as more and more data are available in applied field, the performance of the two most popular MCMC methods has been widely criticized (Betancourt 2019). The performance of the M-H algorithm crucially depends on the proposal distribution. An efficient proposal distribution in M-H algorithm should generate random draws with less auto-correlation, which enables more effective exploration of the parameter space (Quiroz 2015). On the other hand, the performance of the Gibss Sampler crucially depends on the parameter structure. If there is a significant correlation between parameter estimates, the Gibbs Sampler will become very inefficient as the geometry of the distribution is not aligned with the stepping directions of each sampler (Lambert 2018).

2.5   Scalable Bayesian models

Recent ten years witnessed an explosive growth of data size and dimensionality. This poses a major challenge to Bayesian methods using MCMC. Traditional MCMC algorithm need to evaluate the entire data at each step of iteration, which could be expensive for computation in the case of tall data (Bardenet, Doucet, and Holmes 2017). In applied analysis, researchers often need to set thousands of iterations to reach stable posterior distribution, which takes hours or days to implement a single model. Besides, when the researchers have high dimensional data where high-probability regions are concentrated on a extremely limited region of sample space, it would very hard for random-walk MCMC to generate samples from these small regions (Barp et al. 2018). Hierarchical models even complicate this issue by adding random parameters for each subgroup, which further grows the dimensionality of parameter space. Furthermore, when there is high correlation between different parameters that often occur in the case of many parameters, neither the M-H algorithm or Gibbs sam-

pler can efficiently generate samples from the posterior distribtion. All the aforementioned problems motivate researchers in different fields to develop different scalable algorithms to make Bayesian inference for big data.

2.5.1 Hamiltonian Monte Carlo (HMC)

The M-H algorithm and Gibbs sampler can be very inefficient in big data settings because of sparse high-density parameter space, high costs of evaluating the entire data at each step, or a high correlation between parameters. Originally proposed by Duane et al. (1987) with the name of Hybrid Monte Carlo, the Hamiltonian Monte Carlo (HMC) modifies the random-walk behavior in M-H algorithm into a deterministic one by adding auxiliary momentum parameters p_n, thus more efficiently explores the high-density regions in big data settings compared to the traditional M-H algorithm or the Gibbs sampler (Betancourt 2017; Wang, Broccardo, and Song 2019). Although HMC was originally proposed in 1987 (Duane et al. 1987), it is only widely adopted by applied researchers in the recent five years, thanks to the development of the No-U-Turn Sampler (NUTS) (Hoffman and Gelman 2014) and the statistical programming language `Stan` (Carpenter et al. 2017).

Let $\boldsymbol{q}$ denote the position vector and $\boldsymbol{p}$ denote the momentum vector in the conservative dynamics physics system. Note that $\boldsymbol{q}$ and $\boldsymbol{q}$ must have the same length. The combination $(\boldsymbol{q}, \boldsymbol{p})$ then defines a position-momentum phase space, which can be calculated using the conditional distribution (Neal and others 2011; Betancourt 2017):

$$\pi(\boldsymbol{p}, \boldsymbol{q}) = \pi(\boldsymbol{p}|\boldsymbol{q})\pi(\boldsymbol{q})$$

This joint distribution can also be defined in terms of the *Hamiltonian*:

$$\pi(\boldsymbol{p}, \boldsymbol{q}) = e^{-H(\boldsymbol{p}, \boldsymbol{q})}$$

After a little bit of transformation, we have:

$$H(\boldsymbol{p}, \boldsymbol{q}) = -\log \pi(\boldsymbol{p}, \boldsymbol{q})$$
$$= -\log \pi(\boldsymbol{p}|\boldsymbol{q}) - \log \pi(\boldsymbol{q}) \tag{2.10}$$
$$= K(\boldsymbol{p}, \boldsymbol{q}) + V(\boldsymbol{q})$$

In the perspective of physics, the *Hamiltonian* $H(\boldsymbol{p}, \boldsymbol{q})$ is the total energy of the system, which composes of two parts: *kinetic energy* $K(\boldsymbol{p}, \boldsymbol{q})$ and *potential energy* $V(\boldsymbol{q})$. Note that the potential energy $V(\boldsymbol{q}) = -\log \pi(\boldsymbol{q})$ is essentially the negative log of the posterior distribution of the parameter posterior density $\boldsymbol{q}$.

In a static system, the Hamiltonian is a constant. The evolution of this system is governed by the *Hamiltonian equations*:

$$\frac{d\boldsymbol{q}}{dt} = \frac{\partial H}{\boldsymbol{p}} = \frac{\partial K}{\boldsymbol{p}}$$
$$\frac{d\boldsymbol{p}}{dt} = -\frac{\partial H}{\boldsymbol{q}} = -\frac{\partial K}{\boldsymbol{q}} - \frac{\partial V}{\boldsymbol{q}} \tag{2.11}$$

It turns out that we can randomly generate high density proposals in the parameters space by taking advantage of the Hamiltonian system. Here is the general idea if the *HMC algorithm* (Lambert 2018):

1. Let $\theta^{(0)}$ denote a random initial value from a proposal distribution,

2. Set $t = 1$,

3. Generate a random initial momentum m from a proposal distribution (typically a multivariate normal distribution),

4. Use the leapfrog algorithm to solve the trajectory moving over the high-density posterior parameter space under the Hamiltonian mechanism for a period,

5. Calculate the new momentum m' and new position $\theta^{(\mathrm{prop})}$

6. Compute the following probability:

$$\alpha^H = \min\left(1, \frac{p\left(\theta^{(\text{prop})}|\boldsymbol{x}\right)p(\theta^{(\text{prop})})}{p\left(\theta^{(t)}|\boldsymbol{x}\right)p(\theta^{(t)})}\frac{q(m')}{q(m)}\right) \tag{2.12}$$

7. Set $\theta^{(t+1)} = \theta^{(\text{prop})}$ with the probability of α^H; otherwise set $\theta^{(t+1)} = \theta^{(t)}$. Set $t \leftarrow t+1$ and return to 3 until the desired number of iterations is reached.

The HMC is essentially an improved form of M-H algorithm by using the Hamiltonian to generate effective proposals instead of naive random-walk and revised form of the acceptance probability (Equation (2.12)).

Two parameters need to be tuned when implementing the HMC: step size ϵ and the optimal trajectory length T. The optimal trajectory length is the product of the number of steps L and step size ϵ (Neal and others 2011; Monnahan, Thorson, and Branch 2017). The step size ϵ decides how similarly the sympletic methods (typically the leapfrog algorithm) imitates the true unnormalized posterior density. If ϵ is too small, it will take a lot of steps for the leapfrog algorithm to explore the posterior space. If ϵ is too big, the leapfrog algorithm will loop around and return to a place near its original step. The trajectory length $T = \epsilon L$, which need to be tuned in similar style with ϵ: if L is too short, it will be hard to simulate distant proposal and the algorithm is inefficient; if L is too long, the trajectory will loop back and become computationally inefficient. Hand tuning these two parameters was the major obstacle to implement HMC for applied researchers.

The No-U-Turn Sampler (NUTS) proposed by Hoffman and Gelman (2014) solves the difficulty of hand tuning ϵ and T in static HMC. NUTS calculates the optimal step size ϵ and number of steps L through a tree building algorithm (Monnahan, Thorson, and Branch 2017). The tree depth k is defined as the number of doublings, resulting in 2^k leapfrog steps to build the trajectory. This k is then decided by repeating the doubling iterations until the trajectory 'makes a U-turn' (loops back) or diverges (the Hamiltonian expands to infinity). Therefore, the NUTS can automatically create trajectories that can efficiently explore the

high-density parameter space without having to hand tune ϵ and T.

2.5.2   Integrated Laplace Approximation (INLA)

(Rue, Martino, and Chopin 2009; Lindgren, Rue, and Lindström 2011; Martins et al. 2013; De Coninck et al. 2016; Rue et al. 2017; Verbosio et al. 2017; Kourounis, Fuchs, and Schenk 2018)

2.5.3   Subsampling MCMC

Stochastic Gradient HMC

    (Quiroz et al., n.d., 2016; Gunawan et al. 2018; Quiroz, Tran, et al. 2018; Quiroz, Kohn, et al. 2018; Dang et al. 2017; Quiroz 2015)

    T. Chen, Fox, and Guestrin (2014)

2.6   Conceptual framework

Cantor et al. (2010) suggested three factors that cause truck crashes: driver factors, vehicle factors (type and condition), and environmental factors.

    Roshandel, Zheng, and Washington (2015) proposed five factors that affect traffic safety: (a) behavioral characteristics of the driver, e.g., impairment, fatigue, distractions; (b) vehicle — the condition of the vehicle; (c) traffic — the traffic conditions; (d) geometry — geometric characteristics of the road, e.g. curve, hill, ramps, etc.; and (e) environmental — characteristics of the surrounding environment, such as weather conditions (rain, snow, night-time driving, etc.). Traffic conditions are the most studied of these and we focus on discussing them in this subsection.

2.7 Gaps in literature

- A focus on crashes instead of precursors of crashes

- A focus on road segments rather than drivers

- A focus on case-control comparison given the rareness of truck crashes rather than rates

2.8 Research aims

1. **Aim1**: Explore the association between truck crashes and critical events.

2. **Aim2**: Predict critical events using hierarchical statistical models.

3. **Aim3**: scalable hierarchical Bayesian models using subsampling MCMC.

Chapter 3

METHODS

3.1 Data source

3.1.1 Real-time ping

The J.B. Hunt Transport Services, a trucking and transportation company in the United States, provided real-time ping data on 496 truck drivers who conducted regional work (REG, JBI00) from April 1st, 2015 to March 29th, 2016 to the research team. A small device was installed in each of their truck, which will ping irregularly (typically every 5-30 minutes). Each **ping** will collect real-time data on the vehicle number, date and time, latitude, longitude, driver identity number, and speed at that second. In total, 13,187,289 pings were collected and provided to the research team.

For each of the truck drivers, if the ping data showed that the truck was not moving for more than 20 minutes, the ping data were separated into two different trips. These ping data were then aggregated into different trips. A **trip** is therefore defined as a continuous period of driving without stop. The average length of a trip in this study is 2.31 hours with the standard deviation of 1.8 hours.

After the ping data were aggregated into trips, these trips data were then further divded into different shifts according to an eight-hour rest time for each driver. A **shift** is defined

as a long period of driving with potentially less than 8 hours' stops. The Shift_ID column in Table 1 shows different shifts, separated by an eight-hour threshold. The average length of a shift in this study is 8.42 hours with the standard deviation of 2.45 hours.

3.1.2 Truck crashes and critical events

The company regularly collected real-time GPS location and time-stamped critical events data for all their trucks. Four types of critical events were recorded in the critical events data:

1. Hard brake
2. Headway
3. Rolling stability
4. Collision mitigation

Once some thresholds with regard to the driver's operation were met, the sensor will be triggered and critical events will be recorded in their database. There were 12,458 critical events created by the 496 truck drivers during the study period.

3.1.3 Driver demographics

These variables can be linked to the trips and crashes table via a common unique identifier.

3.1.4 Weather data from the `DarkSky API`

We retrieved weather variables from the Dark Sky API: *precipitation intensity*, *precipitation probability*, *wind speed*, and *visibility*. The DarkSky API allows the users to query historic minute-by-minute weather data anywhere on the globe (The Dark Sky Company, LLC 2019). According to the official document, the Dark Sky API is supported by a wide range of weather data sources, which are aggregated together to provide the most precise weather data possible

for a given location (The Dark Sky API 2019). Among several different weather data providers we tested, the Dark Sky API provides the most accurate and complete weather variables.

To reduce the cost of querying weather data for the 13 million historic ping data, we rounded the latitudes and longitudes coordinates to two decimal places, which were worth up to 1.1 kilometers We also rounded the time to the nearest hour and ignore those stopping pings. This scaled the original 13 million ping data down to 4.9 million unique latitude-longitude-date-time combinations. We used the R package `darksky` to obtain weather variables for these reduced 4.9 million unique combinations (Rudis 2018). The weather data for these combinations were then merged back to the origin data. A minimal example of R code to retrieve weather data from the DarkSky API can be found in Appendix 8.

3.1.5   Road geometry data from the `OpenStreetMap`

We queried two road geometry variables from the `OpenStreetMap` (OSM) project: *speed limit* and *the number of lanes.* The OSM data are collaboratively collected by over two million registered users via manual survey, GPS devices, aerial photography, and other open-access sources (Wikipedia contributors 2019). The OpenStreetMap Foundation supports a website to make the data freely available to the public under the Open Database License.

We queried the speed limit and the number of lanes by specifying a bounding box by defining a center point, as well as the width and height in meters in the `center_bbox()` function available from the `osmar` R package (Eugster and Schlesinger 2013). We used real-time longitudes and latitudes as the center point and defined a 100×100 meters box to retrieve the two variables. If the 100×100 meters box is too small to have any road geometry data, we expanded the box to 500×500 and then 1000×1000 to obtain geometry data. If the OSM API returned data from multiple geometry structures, we took the mean as the returned values. The R code to retrieve road geometry data can be found in Appendix 8.

3.2 Data Cleaning

Federal Motor Carrier Safety Administration (2017) demands that property-carrying drivers should not drive more than 14 consecutive hours after coming on duty after 10 hours of rest. This 14-hour restriction cannot be extended by off-duty time. (Goel 2014; Anderson et al. 2017)

3.3 Analytical Plan for Aim 1

This part will be based on truck driver's trips data and crashes data.

3.4 Analytical Plan for Aim 2

In this study, we use vehicle drivers as the sampling unit and adopt hierarchical statistical models that accounts for both driver-level variation and trip-level variation. The workflow is to sample a certain number of drivers from a population of drivers, observe their driving trips or shifts for a specific period, then compare the safety events with non-events, and make conclusions on risk factors associated with these safety events.

3.4.1 Logistic regression

$$Y_{i,d} \sim \text{Bernoulli}(p_i)$$
$$\log \frac{p_{i,d}}{1 - p_{i,d}} = \beta_{0,d} + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_k x_k \tag{3.1}$$
$$\beta_{0,d} \sim N(\mu_0, \sigma_0^2), \quad k = 1, 2, \cdots, D$$

3.4.2 Poisson regression

3.4.3 Non-homogeneous Poisson process

Mean function of a point process:

$$\Lambda(t) = E(N(t))$$

$\Lambda(t)$ is the expected number of failures through time t.

Rate of Occurence of Failures (ROCOF): When Λ is differentiable, the ROCOF is:

$$\mu(t) = \frac{d}{dt}\Lambda(t)$$

The ROCOF can be interpreted as the instantaneous rate of change in the expected number of failures.

Intensity function: The intensity function of a point process is

$$\lambda(t) = \lim_{\Delta t \to 0} \frac{P(N(t, t + \Delta t] \geq 1)}{\Delta t}$$

When there is no simultaneous events, ROCOF is the same as intensity function.

Nonhomogeneous Poisson Process (NHPP): The NHPP is a Poisson process whose intensity function is non-constant.

Power law process (PLP): When the intensity function of a NHPP is:

$$\lambda(t) = \frac{\beta}{\theta}\left(\frac{t}{\theta}\right)^{\beta-1}$$

Where $\beta > 0$ and $\theta > 0$, the process is called the power law process (PLP).

Therefore, the mean function $\Lambda(t)$ is the integral of the intensity function:

$$\Lambda(t) = \int_0^t \lambda(t)dt = \int_0^t \frac{\beta}{\theta}\left(\frac{t}{\theta}\right)^{\beta-1} = \left(\frac{t}{\theta}\right)^{\beta}$$

There are two forms of truncation in a NHPP:

1. **Failure truncation**: When testing stops after a predetermined number of failures, the data are said to be failure truncated.

2. **Time truncation**: Data are said to be time truncated when testing stops at a prede-

termined time t.

The first event

The cumulative density function (cdf) of time to the first event is $F(t_1)$: $F_1(t_1) = P(T_1 \leq t_1) = 1 - S(t_1)$.

The survival function for the first event $S_1(t_1)$ is:

$$S_1(t_1) = P(T_1 > t_1)$$

$$= P(N(0, t_1) = 0) \quad N \text{ is the number of events}$$

$$= e^{-\int_0^{t_1} \lambda_u du} (e^{-\int_0^{t_1} \lambda_u du})^0 / 0!$$

$$= e^{-\int_0^{t_1} \lambda_u du}$$

The probability density function (pdf) of time to the first event can be calculated by taking the first order derivative of the cdf $F_1(t_1)$:

$$f_1(t_1) = \frac{d}{dt_1} F_1(t_1)$$

$$= \frac{d}{dt_1} [1 - S_1(t_1)]$$

$$= -\frac{d}{dt_1} S_1(t_1)$$

$$= -\frac{d}{dt_1} e^{-\int_0^{t_1} \lambda(u) du}$$

$$= -(-\lambda_{t_1}) e^{-\int_0^{t_1} \lambda(u) du}$$

$$= \lambda(t_1) e^{-\int_0^{t_1} \lambda(u) du}$$

If this NHPP is a PLL, we plug in the intensity function $\lambda(t) = (\beta/\theta)(t/\theta)^{\beta-1}$, then we have:

$$f_1(t_1) = \frac{\beta}{\theta} (\frac{t_1}{\theta})^{\beta-1} e^{-(\frac{t_1}{\theta})^\beta}, \quad t_1 > 0$$

This pdf is identical with the pdf of Weibull distribution, so we have:

$$T_1 \sim \text{Weibull}(\beta, \theta)$$

The second event

The Survival function of the second event given the first event occurred at t_2 is:

$$S_2(t_2|t_1) = P(T_2 > t_2|T_1 = t)$$

$$= P(N(t_1, t_2) = 0|T_1 = t_1)$$

$$= e^{-\int_{t_1}^{t_2} \lambda_u du} [\int_{t_1}^{t_2} \lambda_u du]^0/0!$$

$$= e^{-\int_{t_1}^{t_2} \lambda_u du}$$

The we can derive the pdf of t_2 conditioned on t_1

$$f(t_2|t_1) = -\frac{d}{dt_2} S_2(t_2)$$

$$= -\frac{d}{dt_2} e^{-\int_{t_1}^{t_2} \lambda(u)du}$$

$$= \lambda(t_2) e^{-\int_{t_1}^{t_2} \lambda(u)du} \tag{3.2}$$

$$= \frac{\beta}{\theta}(\frac{t_2}{\theta})^{\beta-1} e^{-[(\frac{t_2}{\theta})^\beta - (\frac{t_1}{\theta})^\beta]}$$

$$= \frac{\frac{\beta}{\theta}(\frac{t_2}{\theta})^{\beta-1} e^{-(t_2/\theta)^\beta}}{e^{-(t_1/\theta)^\beta}}, \quad t_2 > t_1$$

In the *failure truncated case*, we know the total number of events n before the experiment starts. We can get the joint likelihood function for $t_1 < t_2 < \cdots < t_n$ in the failure truncated

case based on Equation 3.2.

$$
\begin{aligned}
f(t_1, t_2, \cdots, t_n) &= f(t_1)f(t_2|t_1)f(t_3|t_1, t_2)\cdots f(t_n|t_1, t_2, \ldots, t_{n-1}) \\
&= \lambda(t_1)e^{-\int_0^{t_1}\dot\lambda(u)du}\lambda(t_2)e^{-\int_{t_1}^{t_2}\dot\lambda(u)du}\cdots\lambda(t_n)e^{-\int_{t_{n-1}}^{t_n}\lambda(u)du} \\
&= \Big(\prod_{i=1}^{n}\lambda(t_i)\Big)e^{-\int_0^t\lambda(u)du} \\
&= \Big(\prod_{i=1}^{n}\frac{\beta}{\theta}(\frac{t_i}{\theta})^{\beta-1}\Big)e^{-(t_n/\theta)^\beta}, \quad t_1 < t_2 < \cdots < t_n
\end{aligned}
$$

(3.3)

The log-likelihood function in the failure truncated case is therefore:

$$
\log \ell = n\log\beta - n\beta\log\theta + (\beta - 1)\Big(\sum_{i=1}^{n}\log t_i\Big) - \Big(\frac{t_n}{\theta}\Big)^\beta
$$

In the *time truncated case*, we assume that the truncated time is τ. The derivation of $f(t_1, t_2, \cdots, t_n|n)$ is messy in math, we directly give the conclusion here:

$$
f(t_1, t_2, \cdots, t_n|n) = n!\prod_{i=1}^{n}\frac{\lambda(t_i)}{\Lambda(\tau)}
$$

Therefore, the joint likelihood function for $f(n, t_1, t_2, \cdots, t_n)$ is:

$$
\begin{aligned}
f(n, t_1, t_2, \cdots, t_n) &= f(n)f(t_1, t_2, \cdots, t_n|n) \\
&= \frac{e^{-\int_0^\tau\lambda(u)du}[\int_0^\tau\lambda(u)du]^n}{n!}n!\frac{\prod_{i=1}^{n}\lambda(t_i)}{[\Lambda(\tau)]^n} \\
&= \Big(\prod_{i=1}^{n}\lambda(t_i)\Big)e^{-\int_0^\tau\lambda(u)du} \\
&= \Big(\prod_{i=1}^{n}\frac{\beta}{\theta}(\frac{t_i}{\theta})^{\beta-1}\Big)e^{-(\tau/\theta)^\beta}, \\
n &= 0, 1, 2, \cdots, \quad 0 < t_1 < t_2 < \cdots < t_n
\end{aligned}
$$

(3.4)

The log likelihood function l is then:

$$
\begin{aligned}
l &= \log\left(\left(\prod_{i=1}^{n} \frac{\beta}{\theta}(\frac{t_i}{\theta})^{\beta-1}\right) e^{-(\tau/\theta)^\beta} \right) \\
&= \sum_{i=1}^{n} \log\left(\frac{\beta}{\theta}(\frac{t_i}{\theta})^{\beta-1} \right) - (\frac{\tau}{\theta})^\beta \\
&= n\log\beta - n\beta\log\theta + (\beta-1)\left(\sum_{i=1}^{n} \log t_i \right) - \left(\frac{\tau}{\theta}\right)^\beta
\end{aligned}
\tag{3.5}
$$

3.5   Analytical Plan for Aim 3

Chapter 4

THE PROBABLE CONTENT

This dissertation will be completed according to the three-paper model. The expected chapter headings are as follows:

Chapter 1: Introduction: The problem

A. Epidemiology of head and neck squamous cell carcinoma and its association with second primary malignancies

1. Cancer

2. Head and neck squamous cell carcinoma

3. Definition of second primary malignancies

4. second primary malignancies and Head and neck squamous cell carcinoma

B. Risk factors for second primary malignancies

1. Tobacco use

2. Excessive alcohol use

3. Human papillomavirus infection

C. Existing literature and gaps that exist in the area of second primary malignancies in patients with head and neck squamous cell carcinoma

Chapter 2: Literature review

A. Introduction

B. Methods

1. Source of data and eligibility/exclusion criteria

2. Definitions of primary and secondary outcomes and covariates

3. Statistical analysis

C. Results

1. Description of studies included in the study

2. Primary outcome results

3. Secondary outcome results

4. Publication bias and study quality assessment

D. Discussion/Conclusions

1. Implications

2. Strengths and limitations

3. Future research

Chapter 3: Aim 1 - Truck crashes and critical events

A. Introduction

B. Methods

1. Source of data and eligibility/exclusion criteria

2. Definitions of primary and secondary outcomes and covariates

3. Statistical analysis

C. Results

1. Demographics of study population

2. Primary outcome results

3. Secondary outcome results

D. Discussion/Conclusions

1. Implications

2. Strengths and limitations

3. Future research

Chapter 4: Aim 2 - Statistical models predicting truck critical events

A. Introduction

B. Methods

1. Source of data and eligibility/exclusion criteria

2. Definitions of primary and secondary outcomes and covariates

3. Statistical analysis

C. Results

1. Demographics of study population

2. Primary outcome results

3. Secondary outcome results

D. Discussion/Conclusions

1. Implications

2. Strengths and limitations

3. Future research

Chapter 5: Aim 3 - Subsampling Markov Chain Monte Carlo methods

Chapter 6: DISCUSSION

A. Conclusion

B. Strengths and limitation

C. Future research

Chapter 5

TRUCK CRASHES AND CRITICAL EVENTS

Chapter 6

THREE STATISTICAL MODELS PREDICTING TRUCK CRITICAL EVENTS

6.1 Hierarchical logistic model

6.1.1 Model set up

6.1.2 Bayesian estimation based on simulated data

6.2 Hierarchical Poisson model

6.3 Hierarchical power law process

*Mean function of a point process**:

$$\Lambda(t) = E(N(t))$$

$\Lambda(t)$ is the expected number of failures through time t.

Rate of Occurence of Failures (ROCOF): When Λ is differentiable, the ROCOF is:

$$\mu(t) = \frac{d}{dt}\Lambda(t)$$

The ROCOF can be interpreted as the instantaneous rate of change in the expected number of failures.

Intensity function: The intensity function of a point process is

$$\lambda(t) = \lim_{\Delta t \to 0} \frac{P(N(t, t + \Delta t] \geq 1)}{\Delta t}$$

When there is no simultaneous events, ROCOF is the same as intensity function.

Nonhomogeneous Poisson Process (NHPP): The NHPP is a Poisson process whose intensity function is non-constant.

Power law process (PLP): When the intensity function of a NHPP is:

$$\lambda(t) = \frac{\beta}{\theta}\left(\frac{t}{\theta}\right)^{\beta-1}$$

Where $\beta > 0$ and $\theta > 0$, the process is called the power law process (PLP).

Therefore, the mean function $\Lambda(t)$ is the integral of the intensity function:

$$\Lambda(t) = \int_0^t \lambda(t)dt = \int_0^t \frac{\beta}{\theta}\left(\frac{t}{\theta}\right)^{\beta-1} = \left(\frac{t}{\theta}\right)^{\beta}$$

6.3.1

6.3.2 Bayesian estimation based on simulated data

Chapter 7

HIERARCHICAL BAYESIAN MODELS USING SUBSAMPLING MARKOVE CHAIN MONTE CARLO METHODS

- Comparison with results given by frequentist models using REML (`lmer4`)
- Comparison with results given by Bayesian models using HMC (`rstan`)
- Comparison with results given Bayesian models using INLA (`INLA`)

Chapter 8

DISCUSSION

APPENDIX

Data query

Weather data

```
gps_sample =
  structure(list(
    from_lat = c(41.3473127, 41.8189037, 32.8258477, 40.6776808,
                 40.2366043, 41.3945561, 32.6320605, 40.5413856,
                 33.6287422, 40.0692742, 41.347986, 37.7781459,
                 43.0843081, 41.48026, 43.495149, 41.5228684,
                 41.5763081, 47.6728665, 41.0918361, 41.1537819),
    from_lon = c(-74.2850908, -73.0835104, -97.0306677, -75.1450753,
                 -76.9367494, -72.8589916, -96.8538145, -74.8547061,
                 -113.7671634, -76.762612, -74.284785, -77.4615586,
                 -76.0977384, -73.2107541, -73.7727896, -74.0739204,
                 -88.1529175, -117.3224667, -74.1554972, -74.1887031),
    beg_time = structure(
      c(1453101738, 1437508088, 1436195038, 1435243088, 1454270680,
        1432210106, 1438937772, 1446486480, 1450191622, 1449848630,
        1457597084, 1432870446, 1457968284, 1451298724, 1431503502,
        1443416864, 1438306368, 1445540454, 1452619392, 1436091072),
      class = c("POSIXct", "POSIXt"), tzone = "UTC")),
    .Names = c("from_lat", "from_lon", "beg_time"),
    row.names = c(NA, 20L),
    class = c("tbl_df", "tbl", "data.frame"))
gps_sample
```

```
library(darksky)
add_var = function(dat){
  dat[,c("time", "summary", "icon", "precipIntensity",
         "precipProbability", "temperature", "apparentTemperature",
         "dewPoint", "humidity", "pressure", "windSpeed", "windGust",
         "windBearing", "cloudCover", "visibility")] = NA
  return(dat)
}
```

```r
for(i in 1:nrow(gps_sample)){
  t = get_forecast_for(gps_sample$from_lat[i], gps_sample$from_lon[i],
                       gps_sample$beg_time[i])
  gps_sample$summary[i] = ifelse(is.null(t[[3]]$summary), NA,
                                 t[[3]]$summary)
  gps_sample$icon[i] = ifelse(is.null(t[[3]]$icon), NA, t[[3]]$icon)
  gps_sample$precipIntensity[i] = ifelse(is.null(t[[3]]$precipIntensity),
                                         NA, t[[3]]$precipIntensity)
  gps_sample$precipProbability[i] = ifelse(is.null(t[[3]]$precipProbability),
                                           NA, t[[3]]$precipProbability)
  gps_sample$temperature[i] = ifelse(is.null(t[[3]]$temperature), NA,
                                     t[[3]]$temperature)
  gps_sample$apparentTemperature[i] = ifelse(is.null(
    t[[3]]$apparentTemperature), NA, t[[3]]$apparentTemperature)
  gps_sample$dewPoint[i] = ifelse(is.null(t[[3]]$dewPoint), NA,
                                  t[[3]]$dewPoint)
  gps_sample$humidity[i] = ifelse(is.null(t[[3]]$humidity), NA,
                                  t[[3]]$humidity)
  gps_sample$pressure[i] = ifelse(is.null(t[[3]]$pressure), NA,
                                  t[[3]]$pressure)
  gps_sample$windSpeed[i] = ifelse(is.null(t[[3]]$windSpeed), NA,
                                   t[[3]]$windSpeed)
  gps_sample$windGust[i] = ifelse(is.null(t[[3]]$windGust), NA,
                                  t[[3]]$windGust)
  gps_sample$windBearing[i] = ifelse(is.null(t[[3]]$windBearing), NA,
                                     t[[3]]$windBearing)
  gps_sample$cloudCover[i] = ifelse(is.null(t[[3]]$cloudCover), NA,
                                    t[[3]]$cloudCover)
  gps_sample$visibility[i] = ifelse(is.null(t[[3]]$visibility), NA,
                                    t[[3]]$visibility)
}
```

Road geometry data

```r
pacman::p_load(osmar, stringr)
src <- osmsource_api(url = "https://api.openstreetmap.org/api/0.6/")
road_data = function(i = 5, width = 100, data = df3){
  bb <- center_bbox(data$lon_short[i], data$lat_short[i],
                    width, width)
  ua = get_osm(bb, source = src)
  ua
  road_inf <- data.frame(ua$ways$tags)
  colnames(road_inf) <- c("ID", "Key", "Value")
```

```r
  road_inf$Key <- as.character(road_inf$Key)
  road_inf$Value <- as.character(road_inf$Value)
  row_speed <-  which(road_inf$Key == "maxspeed", arr.ind=TRUE)
  row_lane <- which(road_inf$Key == "lanes", arr.ind=TRUE)

  max_speed <- as.numeric(str_extract(road_inf[row_speed,"Value"],
                                      "[[:digit:]]+"))
  num_lanes <- as.numeric(str_extract(road_inf[row_lane,"Value"],
                                      "[[:digit:]]+"))
  return(c(mean(max_speed), mean(num_lanes)))
}

loop_data = function(start_index = 1, loop_length = 100000){
  end_index = start_index + loop_length
  out_data = data.frame(matrix(0, ncol = 2, nrow = loop_length))
  df_index_diff = start_index-1

  for (i in start_index:end_index) {
    out_data[i-df_index_diff,] = road_data(i, data = df)
    print(paste0(end_index - i, " remained (",
                 round((end_index - i)*100/
                        (end_index-df_index_diff), 3), "%)"))
  }

  return(out_data)
}
```

```r
df = data.table::fread("data/20190605_ping_compressed_3digits.csv")
dfcontainer12 = loop_data(start_index = 1)
```

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