

# Bayesian hierarchical models for NHPP using **rstan**

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*2019-07-10*

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## 1 Model setting

Let  $T_{d,s,i}$  denote the time to the  $d$ -th driver's  $s$ -th shift's  $i$ -th critical event. The total number critical events of  $d$ -th driver's  $s$ -th shift is  $n_{d,s}$ . The ranges of these notations are:

- $i = 1, 2, \dots, n_{d,S_d}$ ,
- $s = 1, 2, \dots, S_d$ ,
- $d = 1, 2, \dots, D$ .

We assume the times of critical events within the  $d$ -th driver's  $s$ -th shift were generated from a non-homogeneous Poisson process (NHPP) with a power law process (PLP), with a fix rate parameter  $\beta$  and varying scale parameters  $\theta_{d,s}$  across drivers. The data generating process is then:

$$T_{d,s,1}, T_{d,s,2}, \dots, T_{d,s,n_{d,s}} \sim \text{PLP}(\beta, \theta_{d,s})$$

$$\beta \sim \text{Gamma}(1, 1)$$

$$\log \theta_{d,s} = \gamma_{0d} + \gamma_1 x_{d,s,1} + \gamma_2 x_{d,s,2} + \dots + \gamma_k x_{d,s,k}$$

$$\gamma_{01}, \gamma_{02}, \dots, \gamma_{0D} \sim \text{i.i.d. } N(\mu_0, \sigma_0^2)$$

$$\gamma_1, \gamma_2, \dots, \gamma_k \sim \text{i.i.d. } N(0, 10^2)$$

$$\mu_0 \sim N(0, 10^2)$$

$$\sigma_0 \sim \text{Gamma}(1, 1)$$

## 2 Simulating data

1. Random intercepts  $\gamma_{01}, \gamma_{02}, \dots, \gamma_{0D}$ . The standard deviation of  $\mu_0$  was intentionally set to small number 2 to make  $\theta_{d,s}$  fall into a reasonably small range. If I otherwise set it as 10,  $\theta_{d,s}$  may be more than  $10^5$  due to the exponentiation, which may not be realistic in real-life data.

$$\mu_0 = 0, \quad \sigma_0 = 0.5$$

$$\sigma_0 \sim \text{Gamma}(1, 1)$$

$$\gamma_{01}, \gamma_{02}, \dots, \gamma_{0D} \sim \text{i.i.d. } N(\mu_0, \sigma_0^2)$$

2. fixed parameters: 3 fixed parameters  $\gamma_1, \gamma_2, \gamma_3$ .

$$\gamma_1, \gamma_2, \gamma_3 \sim \text{i.i.d. } N(0, 10^2)$$

3. The number of observations in the  $d$ -th driver:  $N_d$ .

$$N_d \sim \text{Poisson}(100)$$

4. Data: 3 predictor variables  $x_{d,s,1}, x_{d,s,2}, x_{d,s,3}$ .

$$x_{d,s,1} \sim N(0, 1)$$

$$x_{d,s,2} \sim \text{Gamma}(1, 1)$$

$$x_{d,s,3} \sim \text{Poisson}(0.2)$$

5. Scale parameters of a NHPP (random effects):  $\theta_{d,s}$ .

$$\theta_{d,s} = \text{EXP}(\gamma_{0d} + \gamma_1 x_{d,s,1} + \gamma_2 x_{d,s,2} + \gamma_3 x_{d,s,3})$$

6. Shape parameter of a NHPP (fixed effect):  $\beta \sim \text{Gamma}(1, 1)$ . Set

$$\beta = 1.5$$

7. Simulate a NHPP based on  $\beta$  and  $\theta_{d,s}$ .

$$T_{d,s,1}, T_{d,s,2}, \dots, T_{d,s,n_{d,s}} \sim \text{PLP}(\beta, \theta_{d,s})$$

```

pacman::p_load(rstan, tidyverse, data.table)
source("functions/NHPP_functions.R")

set.seed(123)
D = 10 # the number of drivers
K = 3 # the number of predictor variables

# 1. Random-effect intercepts
# hyperparameters
mu0 = 0
sigma0 = 0.5
r_OD = rnorm(D, mean = mu0, sd = sigma0)

# 2. Fixed-effects parameters
R_K = rnorm(K, mean = 0, sd = 0.5)

# 3. The number of observations (shifts) in the  $d$ -th driver:  $N_{\{d\}}$ 
N_K = rpois(D, 10)
N = sum(N_K) # the total number of obs
id = rep(1:D, N_K)

# 4. Generate data:  $x_1, x_2, \dots, x_K$ 
sim1 = function(group_sizes = N_K){
  ntot = sum(group_sizes)

  int1 = rep(1, ntot)
  x1 = rnorm(ntot, 0, 1)
  x2 = rgamma(ntot, 1, 1)
  x3 = rpois(ntot, 0.2)

  return(data.frame(int1, x1, x2, x3))
}
X = sim1(N_K)

# 5. Scale parameters of a NHPP
# 5a. parameter matrix:  $P$ 
P = cbind(r0 = rep(r_OD, N_K),
          t(replicate(N, R_K)))
M_logtheta = P*X

# returned parameter for each observed shift
beta = 1.5
theta = exp(rowSums(M_logtheta))
round(theta, 3)

## [1] 0.284 1.721 1.440 0.403 1.672 1.265 0.795 2.269 1.485 1.498 2.208
## [12] 1.502 0.875 0.817 0.713 0.777 1.063 0.626 9.235 4.786 1.793 3.045
## [23] 2.627 3.822 2.729 2.755 2.249 2.171 6.624 2.214 6.020 0.975 3.326
## [34] 1.140 2.079 1.334 1.228 0.852 0.792 0.560 1.887 1.434 1.364 2.426
## [45] 4.139 0.976 0.270 5.343 1.880 2.676 4.558 3.342 1.340 3.739 1.500
## [56] 1.794 1.804 1.680 1.937 1.783 0.728 2.232 0.790 0.474 1.108 1.203
## [67] 0.910 0.646 0.507 1.667 0.597 2.714 1.961 0.772 0.441 0.662 1.144
## [78] 0.740 0.659 0.568 0.916 0.606

```

```

sim_hier_plp_tau = function(){
  t_list = list()
  len_list = list()
  tau_vector = rnorm(N, 10, 1.3)

  for (i in 1:N) {
    t_list[[i]] = sim_plp_tau(tau_vector[i], beta, theta[i])
    len_list[[i]] = length(t_list[[i]])
  }

  event_dat = data.frame(
    shift_id = rep(1:N, unlist(len_list)),
    event_time = Reduce(c, t_list)
  )

  start_end_dat = data.frame(
    shift_id = 1:N,
    start_time = rep(0, N),
    end_time = tau_vector #difference2
  )

  return(list(event_dat = event_dat,
             start_end_dat = start_end_dat,
             shift_length = unlist(len_list)))
}

df = sim_hier_plp_tau()

hier_dat = list(
  N = nrow(df$event_dat),
  K = nrow(df$start_end_dat),
  D = id, #driver index
  tau = df$start_end_dat$end_time,
  event_time = df$event_dat$event_time,
  s = df$shift_length, #the number of events in each shift
  x1 = X[,2], x2 = X[,3], x3 = X[,4]
)

```

### 3 Stan code

```
functions{
  real nhpp_log(vector t, real beta, real theta, real tau){
    vector[num_elements(t)] loglik_part;
    real loglikelihood;
    for (i in 1:num_elements(t)){
      loglik_part[i] = log(beta) - beta*log(theta) + (beta - 1)*log(t[i]);
    }
    loglikelihood = sum(loglik_part) - (tau/theta)^beta;
    return loglikelihood;
  }
}

data {
  int<lower=0> N; //total # of obs
  int<lower=0> K; //total # of shifts
  int<lower=0> D[K]; //driver index, this must be an array
  vector<lower=0>[K] tau; //truncated time
  vector<lower=0>[N] event_time; //failure time
  int s[K]; //group sizes
  vector[K] x1;
  vector[K] x2;
  vector[K] x3;
}

parameters{
  real<lower=0> beta;
  vector[K] r0; // random intercept
  vector[3] r; // fixed parameters
  real mu0; // hyperparameter
  real<lower=0> sigma0; // hyperparameter
}

transformed parameters{
  vector<lower=0>[K] theta;
  for (k0 in 1:K){
    theta[k0] = exp(r0[ D[k0] ] + x1[k0]*r[1] + x2[k0]*r[2] + x3[k0]*r[3]);
  }
}

model{
  int position;
  position = 1;
  for (k in 1:K){
    if(s[k] == 0) continue;
    segment(event_time, position, s[k]) ~ nhpp(beta, theta[k], tau[k]);
    position = position + s[k];
  }
  beta ~ gamma(1, 1);
  r0 ~ normal(mu0, sigma0);
  r ~ normal(0, 10);
  mu0 ~ normal(0, 10);
  sigma0 ~ gamma(1, 1);
  theta ~ gamma(1, 0.01);
}
```

## 4 Estimated results

```
f = stan("stan/nhpp_plp_tau_ML.stan",
        chains = 1, iter = 1000, data = hier_dat, refresh = 0)

## DIAGNOSTIC(S) FROM PARSER:
## Info (non-fatal):
## Left-hand side of sampling statement (~) may contain a non-linear transform of a parameter or local v
## If it does, you need to include a target += statement with the log absolute determinant of the Jacob
## Left-hand-side of sampling statement:
##      theta ~ gamma(...)

pacman::p_load(magrittr)
est = broom::tidy(f)

pull_est = function(var = "theta", est_obj = f){
  z = est_obj %>%
    broom::tidy() %>%
    filter(grepl(var, term)) %>%
    pull(estimate) %>%
    round(3)
  return(z)
}
```

Estimated values:

- Hyperparameters:  $\hat{\mu}_0$ : 0.062,  $\hat{\sigma}_0$ : 0.525
- Individual level parameters:  $\gamma_1, \gamma_2, \gamma_3$ : 0.624, 0.161, 0.167
- Rate parameter  $\beta$ : 1.497
- $\theta$ : 0.291, 1.744, 1.412, 0.415, 1.776, 1.299, 0.812, 2.219, 1.552, 1.558, 2.167, 1.518, 0.893, 0.825, 0.726, 0.766, 1.052, 0.605, 9.54, 4.92, 1.691, 2.908, 2.535, 3.893, 2.663, 2.79, 2.278, 2.205, 6.664, 2.213, 6.186, 0.961, 3.387, 1.14, 1.969, 1.338, 1.164, 0.849, 0.746, 0.552, 1.773, 1.435, 1.315, 2.383, 4.348, 0.982, 0.27, 5.345, 1.841, 2.499, 4.644, 3.184, 1.309, 3.646, 1.432, 1.721, 1.78, 1.558, 1.935, 1.662, 0.732, 2.121, 0.794, 0.475, 1.134, 1.192, 0.909, 0.653, 0.479, 1.617, 0.556, 2.664, 1.895, 0.711, 0.445, 0.656, 1.15, 0.75, 0.671, 0.561, 0.922, 0.595

## 5 Further improvement

In Stan code:

- Need a data matrix  $X$ ,
- Need matrix multiplication,

In data:

- Need a driver index  $d = 1, 2, \dots, K$  for each shift  $k$
- Need a data matrix  $X$