

Chapter 1

2025-09-28

Exercises

Exercise 1.1

Exercise 1.11

Consider the gambler's ruin process where at each wager, the gambler wins with probability p and loses with probability $q = 1 - p$. The gambler stops when reaching $\$n$ or losing all their money. If the gambler starts with $\$k$, with $0 < k < n$, find the probability of eventual ruin. See example 1.10.

Solution: Let p_k be the probability of winning the overall game with current wallet $\$k$, and q_k be corresponding probability of losing. At current wallet $\$k$, the chance of winning restarts at the next step up or down, so we condition on the gambler winning or losing this round (LOTP).

$$p_k = p \cdot p_{k+1} + q \cdot p_{k-1} \quad p \cdot p_k + q \cdot p_k = p \cdot p_{k+1} + q \cdot p_{k-1} \quad p(p_{k+1} - p_k) = q(p_k - p_{k-1}) \quad p_k - p_{k-1} = \frac{q}{p}(p_k - p_{k-1})$$

Telescoping and multiplying both sides by p/q :

$$p_1 - p_0 = \frac{p}{q}(p_2 - p_1) = \left(\frac{p}{q}\right)^2(p_3 - p_2) = \cdots = \left(\frac{p}{q}\right)^{n-1}(p_n - p_{n-1})$$

At the lower extreme, $p_1 - p_0 = p_1$ since $p_0 = 0$. $p_1 = \frac{p}{q}(p_2 - p_1)$, so $p_1 + \frac{p}{q}p_1 = p_2$ or $p_2 = \frac{1}{q}p_1$. Then,

Exercise 1.12

In n rolls of a fair die, let X be the number of times 1 is rolled, and Y the number of times 2 is rolled. Find the conditional distribution of X given $Y = y$.

Solution: Since all the 2s have been counted, $X | Y$ can take on values $\{1, 3, 4, 5, 6\}$ with equal probability. There are $n - y$ remaining rolls, and the probability of getting a 1 is $1/5$, so $X | Y = y \sim \text{Binom}(n - y, 1/5)$. Note that the support for $X | Y = y$ are the integers $0 \leq y \leq n$, as $X | Y = y$ is Binomial.

Exercise 1.13

Random variables X and Y have joint density function

$$f(x, y) = 3y, \quad \text{for } 0 < x < y < 1$$

(a) Find the conditional density of Y given $X = x$.

Solution: The marginal density:

$$f(x) = \int_Y f(x, y) dy = \int_x^1 3y dy = \frac{3y^2}{2} \Big|_x^1 = \frac{3 - 3x^2}{2}$$

The conditional density:

$$f(y | x) = \frac{f(x, y)}{f(x)} = \frac{3y}{(3 - 3x^2)/2} = \frac{2y}{1 - x^2} \quad \text{for } 0 < x < y < 1$$

(b) Find the conditional density of Y given $X = x$. Describe the conditional distribution.

Solution: The marginal distribution:

$$f(y) = \int_X f(x, y) dx = \int_0^y 3y dx = 3yx|_0^y = 3y^2$$

The conditional distribution:

$$f(x | y) = \frac{f(x, y)}{f(y)} = \frac{3y}{3y^2} = \frac{1}{y} \text{ for } x < y < 1$$

Note that this conditional distribution is a function of y , and does not depend on x except for the bounds. Once Y has been fixed, then X can take on any value with uniform probability across the region $0 < x < y$.

Exercise 1.14

Random variables X and Y have joint density function

$$f(x, y) = 4e^{-2x}, \quad \text{for } 0 < y < x < \infty$$

(a) Find the conditional density of X given $Y = y$.

Solution: The marginal density $f_Y(y)$ is:

$$f_Y(y) = \int_X f(x, y) dx = \int_y^\infty 4e^{-2x} dx = -2e^{-2x}|_y^\infty = 2e^{-2y}$$

The conditional density is:

$$f_{X|Y=y}(x | y) = \frac{f(x, y)}{f(y)} = \frac{4e^{-2x}}{2e^{-2y}} = 2e^{-2(x-y)}$$

(b) Find the conditional density of Y given $X = x$. Describe the conditional distribution.

Solution: The marginal density:

$$f_X(x) = \int_Y f(x, y) dy = \int_0^x 4e^{-2x} dy = 4ye^{-2x}|_0^x = 4xe^{-2x}$$

The conditional density:

$$f_{Y|X=x}(y | x) = \frac{f(x, y)}{f(x)} = \frac{4e^{-2x}}{4xe^{-2x}} = \frac{1}{x}$$

This is a uniform distribution. Given that X is a certain value, y can take any value between 0 and x , so this is a uniform distribution on the interval $(0, x)$ which is captured in the original limits of the joint density function.

Exercise 1.15

Let X and Y be uniformly distributed on the disk of radius 1 centered at the origin. Find the conditional distribution of Y given $X = x$.

Solution: The joint distribution is:

$$f_{X,Y}(x, y) = \frac{1}{\pi} \text{ for } x^2 + y^2 \leq 1$$

From the definition of conditional probability:

$$f_{Y|X=x}(y | x) = \frac{f_{X,Y}(x, y)}{f_X(x)}$$

The marginal distribution $f_X(x)$ can be found by integrating out the Y variable from the joint distribution

$$f_X(x) = \int_Y f_{X,Y}(x, y) dy = \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy = \frac{2\sqrt{1-x^2}}{\pi}$$

The conditional distribution is therefore:

$$f_{Y|X=x}(y | x) = \frac{2\sqrt{1-x^2}/\pi}{1/\pi} = 2\sqrt{1-x^2} \text{ for } -1 \leq x \leq 1$$

Exercise 1.16

A poker hand consists of five cards drawn from a standard 52-card deck. Find the expected number of aces in a poker hand given that the first card drawn is an ace.

Solution: Let X be the total number of aces drawn and A be the event that the first card is an ace. If A , then there are 3 aces to be drawn in the remaining 4 cards, so $X | A \sim \text{HGeom}(w = 3, b = 48, n = 4)$, with an additional ace at the beginning. Therefore,

$$E(X | A) = 1 + 4 \cdot \frac{3}{52} \approx 1.23$$

Exercise 1.17

Let X be a Poisson random variable with $\lambda = 3$. Find $E(X | X > 2)$.

Solution: The expression is:

$$E(X | X > 2) = \frac{\sum_{x=3}^{\infty} x \cdot P(X = x)}{P(X > 2)}$$

The numerator is:

$$\begin{aligned} \sum_{x=3}^{\infty} x \cdot P(X = x) &= \underbrace{\sum_{x=0}^{\infty} x \cdot P(X = x)}_{=E(X)} - 0 \cdot P(X = 0) - 1 \cdot P(X = 1) - 2 \cdot P(X = 2) \\ &= 3 - P(X = 1) - 2P(X = 2) \\ &= 3 - \frac{3^1 e^{-3}}{1!} - 2 \cdot \frac{3^2 e^{-3}}{2!} \\ &= 3 - e^{-3}(3 + 9) \\ &= 3 - 12e^{-3} \end{aligned}$$

The denominator is:

$$\begin{aligned} P(X > 2) &= 1 - P(X \leq 2) \\ &= 1 - P(X = 0) - P(X = 1) - P(X = 2) \\ &= 1 - \frac{3^0 e^{-3}}{0!} - \frac{3^1 e^{-3}}{1!} - \frac{3^2 e^{-3}}{2!} \\ &= 1 - e^{-3}(1 + 3 + 9/2) \\ &= 1 - e^{-3}(17/2) \end{aligned}$$

Combining both, we have:

$$E(X | X > 2) = \frac{3 - 12e^{-3}}{1 - 8.5e^{-3}}$$

Exercise 1.18

From the definition of conditional expectation given an event, show that

$$E(I_B | A) = P(B | A)$$

Solution:

$$E(I_B | A) = \frac{E(I_B I_A)}{P(A)} = \frac{E(I_{A \cap B})}{P(A)} = \frac{P(A \cap B)}{P(A)} = P(B | A)$$

Exercise 1.19 NOT DONE

Exercise 1.20

A fair coin is flipped repeatedly. (a) Find the expected number of flips needed to get three heads in a row.

Solution: These events partition the sample space: T, HT, HHT , and HHH . Let X be the flips to get 3 in a row.

$$\begin{aligned} E(X) &= E(X | T)P(T) + E(X | HT)P(HT) + E(X | HHT)P(HHT) + E(X | HHH)P(HHH) \\ &= E(X | T)\frac{1}{2} + E(X | HT)\frac{1}{4} + E(X | HHT)\frac{1}{8} + E(X | HHH)\frac{1}{8} \end{aligned}$$

If we get a tails, we restart. For simplicity, let $a = E(X)$. So $E(X | T) = 1 + E(X)$, $E(X | HT) = 2 + E(X)$, $E(X | HHT) = 3 + E(X)$, and $E(X | HHH) = 3$:

$$\begin{aligned} a &= \frac{1+a}{2} + \frac{2+a}{4} + \frac{3+a}{8} + \frac{3}{8} \\ a &= \frac{1}{2} + \frac{2}{4} + \frac{3}{8} + \frac{3}{8} + a\left(\frac{1}{2} + \frac{1}{4} + \frac{1}{8}\right) \\ \frac{a}{8} &= \frac{7}{4} \\ a &= 14 \end{aligned}$$

Exercise 1.21

Let T be a nonnegative, continuous random variable. Show

$$\int_0^\infty P(T > t)dt = \int_0^\infty \left(\int_t^\infty f_T(x)dx \right) dt$$

The inner integral runs from t to ∞ , and the outer integral from 0 to ∞ . Swapping the order of integration, the inner integral now runs from t to ∞ , while the outer integral runs from 0 to x (note that the bounds of integration coincide at the line $x = t$ if we imagine x axis on the horizontal axis and t on the vertical axis).

$$\int_0^\infty \left(\int_t^\infty f_T(x)dx \right) dt = \int_0^\infty \left(\int_0^x f_T(x)dt \right) dx = \int_0^\infty f_T(x) \left(\int_0^x dt \right) dx = \int_0^\infty f_T(x)x dx = \int_0^\infty x f_T(x)dx = E(T)$$

Exercise 1.22

Let $E(Y | X)$ when (X, Y) is uniformly distributed on the following regions. (a) The rectangle $[a, b] \times [c, d]$.

Solution: Since this is a rectangle and (X, Y) is uniformly distributed, there is an equal likelihood of Y appearing between $[c, d]$ regardless of where X is, so $E(Y) = E(Y | X) = \frac{c+d}{2}$.

(b) The triangle with vertices $(0,0), (1,0), (1,1)$.

Solution: The joint distribution is:

$$f_{X,Y}(x, y) = 2 \text{ for } 0 < y < x < 1$$

The marginal distribution of X is:

$$f_X(x) = \int_Y 2dy = \int_0^x 2dy = 2x$$

The conditional distribution of $Y \mid X$ is:

$$f_{Y|X}(y \mid x) = \frac{f_{X,Y}(x, y)}{f_X(x)} = \frac{2}{2x} = \frac{1}{x}$$

$$E(Y \mid X) = \int_Y y \cdot f_{Y|X}(y \mid x) dy = \int_0^x y \cdot \frac{1}{x} dy = \int_0^x \frac{y}{x} dy = \frac{y^2}{2x} \Big|_0^x = \frac{x^2}{2x} = \frac{x}{2}$$

(c) The disc of radius 1 centered at the origin.

Solution: The joint distribution is:

$$f_{X,Y}(x, y) = \frac{1}{\pi} \text{ for } x^2 + y^2 \leq 1$$

The marginal distribution of X :

$$f_X(x) = \int_Y f_{X,Y}(x, y) dy = \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy$$

$$= \frac{y}{\pi} \Big|_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}}$$

$$= \frac{2\sqrt{1-x^2}}{\pi}$$

Then the conditional distribution of $Y \mid X$ is:

$$f_{Y|X}(y \mid x) = \frac{f_{X,Y}(x, y)}{f_X(x)} = \frac{1/\pi}{2\sqrt{1-x^2}/\pi} = \frac{1}{2\sqrt{1-x^2}}$$

The conditional expectation $E(Y \mid X)$:

$$E(Y \mid X) = \int_Y y \cdot f_{Y|X}(y \mid x) dy = \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{y}{2\sqrt{1-x^2}} dy = \frac{y^2}{2\sqrt{1-x^2}} \Big|_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} = \frac{1-x^2 - (1-x^2)}{2\sqrt{1-x^2}} = 0$$

Exercise 1.23

Let X_1, X_2, \dots be an i.i.d sequence random variables with common mean μ . Let $S_n = X_1 + \dots + X_n$ for $n \geq 1$.

(a) Find $E(S_m \mid S_n)$, for $m \leq n$.

Solution: Since $S_m = X_1 + \dots + X_m$ and $S_n = S_m + X_{m+1} + \dots + X_n$:

$$E(S_m \mid S_n) = E(S_n - X_n - \dots - X_{m+1} \mid S_n)$$

$$= E(S_n \mid S_n) - E(X_n \mid S_n) - \dots - E(X_{m+1} \mid S_n)$$

$$= S_n - E(X_n \mid S_n) - \dots - E(X_{m+1} \mid S_n)$$

By symmetry, $E(X_i \mid S_n) = S_n/n$, so $E(S_m \mid S_n) = S_n - S_n(n-m)/n = S_n(m/n)$.

(b) Find $E(S_m | S_n)$, for $m > n$.

Solution:

$$E(S_m | S_n) = E(S_n + X_{n+1} + \cdots + X_m | S_n) = S_n + E(X_{n+1} | S_n) + \cdots + E(X_m | S_n)$$

Since each X_i is independent of S_n for $i > n$, then $E(S_m | S_n) = S_n + (m - n)\mu$.

Exercise 1.24

Prove the law of total expectation $E(Y) = E(E(Y | X))$ for the continuous case.

Solution:

$$E(E(Y | X)) = \int_x E(Y | X = x) f_X(x) dx = \int_x \left(\int_y y f_{Y|X}(y | x) dy \right) f_X(x) dx$$

The conditional distribution $f_{Y|X} = \frac{f_{X,Y}(x,y)}{f_X(x)}$.

$$\begin{aligned} & \int_x \left(\int_y y f_{Y|X}(y | x) dy \right) f_X(x) dx \\ &= \int_x \left(\int_y y \frac{f_{X,Y}(x,y)}{f_X(x)} dy \right) f_X(x) dx \\ &= \int_x \left(\int_y y f_X(x) \frac{f_{X,Y}(x,y)}{f_X(x)} dy \right) dx \\ &= \int_x \int_y y f_{X,Y}(x,y) dy dx \\ &= \int_y y \left(\int_x f_{X,Y}(x,y) dx \right) dy \\ &= \int_y y f_Y(y) dy \\ &= E(Y) \end{aligned}$$

Exercise 1.25

Let X and Y be independent exponential random variables with respective parameters 1 and 2. Find $P(X/Y < 3)$ by conditioning.

Solution:

$$\begin{aligned}
P(X/Y < 3) &= \int_0^\infty P(X/Y < 3 \mid Y = y) f_Y(y) dy \\
&= \int_0^\infty P(X/y < 3 \mid Y = y) f_Y(y) dy \\
&= \int_0^\infty P(X < 3y) P(Y = y) dy \\
&= \int_0^\infty (1 - e^{-3y}) 2e^{-2y} dy \\
&= \int_0^\infty 2e^{-2y} - 2e^{-5y} dy \\
&= -e^{-2y} \Big|_0^\infty + \frac{2}{5} e^{-5y} \Big|_0^\infty \\
&= 1 - \frac{2}{5} \\
&= \frac{3}{5}
\end{aligned}$$

Exercise 1.26

The density of X is $f(x) = xe^{-x}$, for $x > 0$. Given $X = x$, Y is uniformly distributed on $(0, x)$. Find $P(Y < 2)$ by conditioning on X .

Solution:

We have $f_{Y|X}(y \mid x) = 1/x$, and $F_{Y|X}(y \mid x) = \frac{y}{x}$ for $0 < y < x$, and 1 for $y \geq x$.

If $X < 2$, then Y will always be between 0 and 2, so $P(Y < 2 \mid X < 2) = 1$. If $X \geq 2$, then $P(Y < 2 \mid X \geq 2) = 2/x$.

$$\begin{aligned}
P(Y < 2) &= \int_0^\infty P(Y < 2 \mid X = x) f_X(x) dx \\
&= \int_0^\infty P(Y < 2 \mid X = x) x e^{-x} dx \\
&= \underbrace{\int_0^2 1 \cdot x e^{-x} dx}_{P(0 < X < 2)} + \int_2^\infty y/x \cdot x e^{-x} dx \\
&= + \int_2^\infty 2e^{-x} dx \\
&= (1 - e^{-2}) - 2e^{-x} \Big|_2^\infty dx \\
&= 1 - e^{-2} + 2e^{-2}
\end{aligned}$$

```

exercise_1_26 <- function() {
  x <- rexp(1000, 1)
  y <- runif(1000, rep(0, 1000), x)
  mean(y < 2)
}

```

Exercise 1.27

A restaurant receives N customers per day, where N is a random variable with mean 200 and standard deviation 40. The amount spent by each customer is normally distributed with mean \$15 and standard

deviation \$3. The amounts that customers spend are independent of each other and independent of N . Find the mean and standard deviation of the total amount spent at the restaurant per day.

Solution: Let X_i be the amount that the i customer spends, and $S = X_1 + \cdots + X_N$.

$$\begin{aligned} E(S) &= E(E(S | N)) = E(E(X_1 + \cdots + X_N | N)) = E(NX_i) = E(N)E(X_i) = 200 \cdot 15 = 3000 \\ \text{Var}(S) &= E(\text{Var}(S | N)) + \text{Var}(E(S | N)) \\ &= E(\text{Var}(X_1 + \cdots + X_N | N)) + \text{Var}(E(X_1 + \cdots + X_N | N)) \\ &= E(\text{Var}(X_1 | N) + \cdots + \text{Var}(X_N | N)) + \text{Var}(E(X_1 | N) + \cdots + E(X_N | N)) \\ &= E(N \cdot 3^2) + \text{Var}(N \cdot 15) \\ &= 200 \cdot 3^2 + 15^2 \cdot 1600 \end{aligned}$$

So the standard deviation is $\sqrt{361,800} = 30\sqrt{201} \approx 601.50$.

```
nsim <- 1000
Ns <- rbinom(nsim, size = 25/(1-0.125), prob = 0.125)
Ss <- sapply(Ns, function(n) sum(rnorm(n, mean = 15, sd = 3)))
mean(Ss) # should be 3000

## [1] 3016.168

sd(Ss) # should be around 601.5

## [1] 603.4623
```

Exercise 1.28

On any day, the number of accidents on the highway has a Poisson distributions with parameter Λ . The parameter Λ varies from day to day and is itself a random variable. Find the mean and variance of the number of accidents per day when Λ is uniformly distributed on $(0,3)$.

Solution: Let the total number of accidents be T , thus $T | \Lambda \sim \text{Pois}(\Lambda)$.

$$\begin{aligned} E(T) &= E(E(T | \Lambda)) = E(\Lambda) = \frac{3}{2} \\ \text{Var}(T) &= E(\text{Var}(T | \Lambda)) + \text{Var}(E(T | \Lambda)) \\ &= E(\Lambda) + \text{Var}(\Lambda) \\ &= \frac{3}{2} + \frac{3^2}{12} = \frac{27}{12} \\ &= \frac{9}{4} \end{aligned}$$

Exercise 1.29

If X and Y are independent, does $\text{Var}(Y | X) = \text{Var}(Y)$?

Solution: Yes, since $\text{Var}(Y | X) = E(Y^2 | X) - E(Y | X)^2 = E(Y^2) - E(Y)^2 = \text{Var}(Y)$.

Exercise 1.30

Assume that $Y = g(X)$ is a function of X . Find simple expressions for (a) $E(Y | X)$.

Solution: $E(g(X) | X) = g(X)$

(b) $\text{Var}(Y | X)$.

Solution: $\text{Var}(Y | X) = E(Y^2 | X) - E(Y | X)^2 = E(g(X)^2 | X) - E(g(X) | X)^2 = g(X)^2 - g(X)^2 = 0$.

Exercise 1.31 NOT DONE

Consider a sequence of i.i.d. Bernoulli trials with success parameter p . Let X be the number of trials needed until the first success occurs. Then, X has a geometric distribution with parameter p . Find the variance of X by conditioning on the first trial.

Solution: Let T_1 be the outcome of the first trial. If the first trial succeeds, then $X = 1$. Otherwise, if the trial fails, then we restart the game. The cases are:

$$X = \begin{cases} 0 & \text{with probability } p \\ 1 + X & \text{with probability } 1 - p \end{cases}$$

$$\text{Var}(X) = E(\text{Var}(X \mid T_1)) + \text{Var}(E(X \mid T_1)) = E()$$

Exercise 1.32

R: Simulate flipping three fair coins and counting the number of heads X .

(a) Use your simulation to estimate $P(X = 1)$ and $E(X)$.

Solution:

```
nsim <- 10000
heads_from_fair_coin <- rbinom(nsim, 3, 0.5)
mean(heads_from_fair_coin == 1) # P(X=1)
```

```
## [1] 0.3776
```

```
mean(heads_from_fair_coin) # E(X)
```

```
## [1] 1.4913
```

(b) Modify the above to allow for a biased coin where $P(\text{Heads}) = 3/4$.

Solution:

```
nsim <- 10000
p_biased_coin <- 0.75
heads_from_biased_coin <- rbinom(nsim, 3, p_biased_coin)
mean(heads_from_biased_coin == 1) # P(X=1)
```

```
## [1] 0.1394
```

```
mean(heads_from_biased_coin) # E(X)
```

```
## [1] 2.2598
```

Exercise 1.33

R: Cards are drawn from a standard deck, with replacement, until an ace appears. Simulate the mean and variance of the number of cards required.

Solution: