# Multivariate Bayesian Variable Selection and Prediction

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## Chapter 1

## Foreword

I am Caleb Jin. After I read this paper, Multivariate Bayesian variable selection and prediction (Brown and Vannucci, 1998), I write down the nodes of the key idea and R code to realize it.

### Chapter 2

# Multivariate Bayesian Variable Selection and Prediction

Consider a multivariate vairable regression model,

$$Y = XB + E,$$

where  $\mathbf{Y} = (\mathbf{y}_1, \mathbf{y}_2, \dots, \mathbf{y}_n)^{\top}$  is the  $n \times q$  response matrix,  $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_p)$  is the  $n \times p$  design matrix,  $\mathbf{B} = (\boldsymbol{\beta}_1, \dots, \boldsymbol{\beta}_p)^{\top}$  is a  $p \times q$  unknown coefficient matrix, and  $\mathbf{E} = (\mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_n)^{\top}$  is an  $n \times q$  error matrix with  $\mathbf{e}_i \overset{i.i.d}{\sim} \mathcal{N}_q(\mathbf{0}, \boldsymbol{\Sigma})$ . To identify the variables selected among p regressors, (Brown and Vannucci, 1998) introduced a latent binary vector  $\boldsymbol{\xi} = (\xi_1, \dots, \xi_p)^{\top}$  with

$$\xi_j = \begin{cases} 1 & \text{if } \beta_j \text{ is active} \\ 0 & \text{if } \beta_j \text{ is not active} \end{cases}.$$

The conjugate priors are considered as follows:

$$\pi(\mathbf{B}_{\boldsymbol{\xi}}|\boldsymbol{\Sigma},\boldsymbol{\xi}) \sim \mathcal{MN}(\mathbf{0}, \mathbf{H}_{\boldsymbol{\xi}}, \boldsymbol{\Sigma})$$

$$\pi(\boldsymbol{\Sigma}) \sim \mathcal{IW}(\mathbf{Q}, \delta)$$

$$\pi(\boldsymbol{\xi}) \stackrel{ind}{\sim} \prod_{j=1}^{p} \mathrm{Ber}(\omega_{j}),$$

where  $\mathbf{Q}$ ,  $\mathbf{H}_{\boldsymbol{\xi}}$  and  $\delta$  are hyperparameters. Then, by Bayesian theorem, the marginal posterior distribution can be obtained by

$$\begin{split} &m(\xi|\mathbf{Y}) \propto m(\mathbf{Y}|\xi)\pi(\xi) \\ &= \iint f(\mathbf{Y}|\mathbf{B}_{\xi}, \mathbf{\Sigma})\pi(\mathbf{B}_{\xi}|\mathbf{\Sigma}, \xi)\pi(\mathbf{\Sigma})d\mathbf{B}d\mathbf{\Sigma}\pi(\xi) \\ &= \iint (2\pi)^{-\frac{n_2}{2}} |\mathbf{\Sigma}|^{-\frac{n}{2}} \exp\left[-\frac{1}{2}tr\{(\mathbf{Y} - \mathbf{X}_{\xi}\mathbf{B}_{\xi})^{\top}(\mathbf{Y} - \mathbf{X}_{\xi}\mathbf{B}_{\xi})\}\mathbf{\Sigma}^{-1}\right] \times \\ &(2\pi)^{-\frac{|\xi|}{2}} |\mathbf{\Sigma}|^{-\frac{|\xi|}{2}} |\mathbf{H}_{\xi}|^{-\frac{n}{2}} \exp\left[\frac{1}{2}tr\{(\mathbf{Y} - \mathbf{X}_{\xi}\mathbf{B}_{\xi})^{\top}(\mathbf{Y} - \mathbf{X}_{\xi}\mathbf{B}_{\xi})\}\mathbf{\Sigma}^{-1}\right] \times \\ &|\mathbf{Q}|^{\frac{\delta}{2}} 2^{-\frac{1}{2}\delta\eta} \xi_{\eta}^{-1}(\frac{\delta}{2})|\mathbf{\Sigma}|^{-\frac{\delta+\eta+1}{2}} \exp\{-\frac{1}{2}tr(\mathbf{Q}\mathbf{\Sigma}^{-1})\}d\mathbf{B}_{\xi}d\mathbf{\Sigma}\pi(\xi) \\ &\propto (2\pi)^{-\frac{|\xi|}{2}} |\mathbf{H}_{\xi}|^{-\frac{q}{2}} \iint |\mathbf{\Sigma}|^{-\frac{1}{2}(n+|\xi|+\delta+q+1)} \exp\{-\frac{1}{2}tr(\mathbf{Q}\mathbf{\Sigma}^{-1})\} \\ &\exp\left[-\frac{1}{2}tr\left\{(\mathbf{B}_{\xi}^{\top}(\mathbf{X}_{\xi}^{\top}\mathbf{X}_{\xi} + \mathbf{H}_{\xi}^{-1})\mathbf{B}_{\xi} - 2\mathbf{B}_{\xi}^{\top}\mathbf{X}_{\xi}^{\top}\mathbf{Y}\right)\mathbf{\Sigma}^{-1}\right\}\right] \times \\ &\exp\left[-\frac{1}{2}tr\left\{(\mathbf{H}_{\xi}^{\top}(\mathbf{X}_{\xi}^{\top}\mathbf{X}_{\xi} + \mathbf{H}_{\xi}^{-1})\mathbf{B}_{\xi} - 2\mathbf{B}_{\xi}^{\top}\mathbf{X}_{\xi}^{\top}\mathbf{Y}\right)\mathbf{\Sigma}^{-1}\right\}\right] d\mathbf{B}_{\xi}d\mathbf{\Sigma}\pi(\xi) \\ &= (2\pi)^{-\frac{|\xi|}{2}} |\mathbf{H}_{\xi}|^{-\frac{q}{2}} \iint |\mathbf{\Sigma}|^{-\frac{1}{2}(n+|\xi|+\delta+q+1)} \exp\{-\frac{1}{2}tr[(\mathbf{Y}^{\top}\mathbf{Y} + \mathbf{Q})\mathbf{\Sigma}^{-1}]\} \times \\ &\exp\left[-\frac{1}{2}tr\left\{(\mathbf{B}_{\xi}^{\top}\mathbf{K}_{\xi}\mathbf{B}_{\xi} - 2\mathbf{B}_{\xi}^{\top}\mathbf{K}_{\xi}\mathbf{K}_{\xi}^{-1}\mathbf{X}_{\xi}^{\top}\mathbf{Y}\right)\mathbf{\Sigma}^{-1}\right\}\right] d\mathbf{B}_{\xi}d\mathbf{\Sigma}\pi(\xi) \\ &= (2\pi)^{-\frac{|\xi|}{2}} |\mathbf{H}_{\xi}|^{-\frac{q}{2}} \iint |\mathbf{\Sigma}|^{-\frac{1}{2}(n+|\xi|+\delta+q+1)} \times \\ &\exp\left[-\frac{1}{2}tr[(\mathbf{Y}^{\top}\mathbf{Y} - \hat{\mathbf{B}}_{\xi}^{\top}\mathbf{K}_{\xi}\hat{\mathbf{B}}_{\xi} + \mathbf{Q})\mathbf{\Sigma}^{-1}]\right] d\mathbf{B}_{\xi}d\mathbf{\Sigma}\pi(\xi) \\ &= (2\pi)^{-\frac{|\xi|}{2}} |\mathbf{H}_{\xi}|^{-\frac{q}{2}} (2\pi)^{\frac{|\xi|}{2}} |\mathbf{K}_{\xi}|^{-\frac{q}{2}} \times \\ &\int |\mathbf{\Sigma}|^{\frac{|\xi|}{2}}|\mathbf{\Sigma}|^{-\frac{1}{2}(n+|\xi|+\delta+q+1)} \exp\left\{-\frac{1}{2}tr[(\mathbf{Y}^{\top}\mathbf{Y} - \tilde{\mathbf{B}}_{\xi}^{\top}\mathbf{K}_{\xi}\tilde{\mathbf{B}}_{\xi} + \mathbf{Q})\mathbf{\Sigma}^{-1}]\right\} d\mathbf{\Sigma}\pi(\xi) \\ &= |\mathbf{H}_{\xi}|^{-\frac{q}{2}}|\mathbf{K}_{\xi}|^{-\frac{q}{2}} \int |\mathbf{\Sigma}|^{-\frac{1}{2}(n+\delta+q+1)} \times \\ &\exp\left\{-\frac{1}{2}tr[(\mathbf{Y}^{\top}\mathbf{Y} - \mathbf{Y}^{\top}\mathbf{X}_{\xi}\mathbf{K}_{\xi}^{-1}\mathbf{X}_{\xi}^{\top}\mathbf{Y} + \mathbf{Q})\mathbf{\Sigma}^{-1}]\right\} d\mathbf{\Sigma}\pi(\xi) \\ &= |\mathbf{H}_{\xi}\mathbf{K}_{\xi}|^{-\frac{q}{2}} \int |\mathbf{\Sigma}|^{-\frac{1}{2}(n+\delta+q+1)} \times \\ &\exp\left\{-\frac{1}{2}tr[(\mathbf{Y}^{\top}\mathbf{Y} - \mathbf{Y}^{\top}\mathbf{X}_{\xi}\mathbf{K}_{\xi}^{-1}\mathbf{X}_{\xi}^{\top}\mathbf{Y} + \mathbf{Q})\mathbf{\Sigma}^{-1}]\right\} d\mathbf{\Sigma}\pi(\xi) \\ &= |\mathbf{H}_{\xi}\mathbf{K}_{\xi}|^{-\frac{q}{2}} \int |\mathbf{\Sigma}|^{-\frac{1}{2}(n+\delta+q+1)} \times \\ &\exp\left\{-\frac{1}{2}tr[(\mathbf{Y}^{\top}\mathbf{Y} - \mathbf{X}_{\xi}\mathbf{K}_{\xi}^{-1}\mathbf{X}_{\xi$$

where  $\mathbf{K}_{\boldsymbol{\xi}} = \mathbf{X}_{\boldsymbol{\xi}}^{\top} \mathbf{X}_{\boldsymbol{\xi}} + \mathbf{H}_{\boldsymbol{\xi}}^{-1}$  and  $g(\boldsymbol{\xi}|\mathbf{Y})$  is the proportional form of  $\pi(\boldsymbol{\xi}|\mathbf{Y})$ . (Brown and Vannucci, 1998) set  $\mathbf{Q} = k\mathbf{I}_q$  and  $\mathbf{H}_{\boldsymbol{\xi}} = c(\mathbf{X}_{\boldsymbol{\xi}}^{\top} \mathbf{X}_{\boldsymbol{\xi}})^{-1}$  and  $\omega_j = \omega$ , then  $\pi(\boldsymbol{\xi}) = \prod_{j=1}^p \omega_j^{\xi_j} (1 - \omega_j)^{1-\xi_j} = \omega^{|\boldsymbol{\xi}|} (1 - \omega)^{p-|\boldsymbol{\xi}|}$ . Apply  $\mathbf{Q} = k\mathbf{I}_q$  and  $\mathbf{H}_{\boldsymbol{\xi}} = c(\mathbf{X}_{\boldsymbol{\xi}}^{\top} \mathbf{X}_{\boldsymbol{\xi}})^{-1}$  into (2.1), hence we have

$$\log g(\boldsymbol{\xi}|\mathbf{Y}) = -\frac{n+\delta}{2}\log|k\mathbf{I}_q + \mathbf{Y}^{\top}\mathbf{Y} - \frac{c}{c+1}\mathbf{Y}^{\top}\mathbf{X}_{\boldsymbol{\xi}}(\mathbf{X}_{\boldsymbol{\xi}}^{\top}\mathbf{X}_{\boldsymbol{\xi}}^{-1})\mathbf{X}_{\boldsymbol{\xi}}^{\top}\mathbf{Y}| - \frac{q|\boldsymbol{\xi}|}{2}\log(c+1) + |\boldsymbol{\xi}|\log\omega + (p-|\boldsymbol{\xi}|)\log(1-\omega).$$

Exhaustive computation of  $g(\boldsymbol{\xi}|\mathbf{Y})$  for  $2^p$  values of  $\boldsymbol{\xi}$  becomes prohibitive even in a super computer when p is larger than 40. In this circumstances, using MCMC sampling to explore marginal posterior is possible. (Brown and Vannucci, 1998) use Gibbs sampler to generate each  $\boldsymbol{\xi}$  value component-wise from full conditional distributions  $\pi(\xi_j|\boldsymbol{\xi}_{-j},\mathbf{Y})$ , where  $\boldsymbol{\xi}_{-j}=\boldsymbol{\xi}\setminus\xi_j$ . It is easy to show that

$$\pi(\xi_j = 1 | \boldsymbol{\xi}_{-j}, \mathbf{Y}) = \frac{\theta_j}{\theta_j + 1},$$

where

$$\theta_j = \frac{g(\xi_j = 1, \boldsymbol{\xi}_{-j} | \mathbf{Y})}{g(\xi_j = 0, \boldsymbol{\xi}_{-j} | \mathbf{Y})}.$$

The complete algorithm of (Brown and Vannucci, 1998) is described in Algorithm 3:

#### Algorithm 3 Brown: MCMC sampling for $\xi$

- 1. Initialize  $\boldsymbol{\xi} = (\xi_1, \xi_2, \dots, \xi_p)$ .
- 2. Repeat for  $t = 1, 2, \ldots$  until convergence
  - 1).  $a \leftarrow q(\xi_1^{(t)} = 1, \boldsymbol{\xi}_{-1}^{(t)} | \mathbf{Y}) \text{ and } b \leftarrow q(\xi_1^{(t)} = 0, \boldsymbol{\xi}_{-1}^{(t)} | \mathbf{Y}).$
  - 2).  $\theta_1 \leftarrow \frac{a}{b}$ ;  $p_1 \leftarrow \frac{\theta_1}{\theta_{1+1}}$ .
  - 3).  $\xi^* \sim \text{Ber}(p_1); \, \xi_1^{(t)} \leftarrow \xi^*.$
  - 4). Repeat for  $j = 2, \ldots, p$ :

i) 
$$\kappa = a\xi^* + b(1 - \xi^*); \tilde{\boldsymbol{\xi}} \leftarrow \boldsymbol{\xi}^{(t)}.$$

ii) If 
$$\xi_j^{(t)} = 1$$
,  
 $\tilde{\xi}_j \leftarrow 0$ ;  $a \leftarrow \kappa$ ;  $b \leftarrow g(\tilde{\boldsymbol{\xi}}|\mathbf{Y})$ .  
else

$$\tilde{\xi}_j \leftarrow 1; \ a \leftarrow g(\tilde{\boldsymbol{\xi}}|\mathbf{Y}); \ b \leftarrow \kappa.$$

iii) 
$$\theta_j \leftarrow \frac{a}{b}$$
;  $p_j \leftarrow \frac{\theta_j}{\theta_j + 1}$ .

iv) 
$$\xi^* \sim Ber(p_j); \, \xi_j^{(t)} \leftarrow \xi^*.$$

5). Get  $\xi^{(t)}$ 

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For the hyperparameter, we follow the same setting with (Brown and Vannucci, 1998), which are  $\delta=2+q, \omega=20/p, k=2$  and c=10. Note that since our definition of inverse-wishart distribution is different from that of (Brown and Vannucci, 1998), the values of q and k we set here are after adjustment to make them consistent with (Brown and Vannucci, 1998).

# **Bibliography**

Brown, P. J. and Vannucci, M. (1998). Multivariate bayesian variable selection and prediction. *Journal of the Royal Statistical Society: Series B (Methodological)*, pages 627–641.