

Approaches for Bayesian Variable Selection (SSVS)

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Chapter 1

Foreword

I am Caleb Jin. After I read this paper, **Approaches for Bayesian Variable Selection (SSVS)**(George and McCulloch, 1997) and (George and McCulloch, 1993), I write down the nodes of the key idea and R code to realize it.

Chapter 2

Approaches for Bayesian Variable Selection (SSVS)

Consider a high dimensional linear regression model as follows:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon} \quad (2.1)$$

where $\mathbf{y} = (y_1, y_2, \dots, y_n)^\top$ is the n -dimensional response vector, $\mathbf{X} = [1, \mathbf{M}] = [\mathbf{x}_1, \dots, \mathbf{x}_p]$ is the $n \times p$ design matrix, and $\boldsymbol{\epsilon} \sim \mathcal{N}_n(\mathbf{0}, \sigma^2 \mathbf{I}_n)$. Note that as $p > n$, \mathbf{X} is not full rank.

From (2.1), the likelihood function is given as

$$\mathbf{y}|\boldsymbol{\beta}, \sigma^2 \sim \mathcal{N}(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I}_n).$$

We define the prior as follows:

$$\beta_j|\sigma^2, \gamma_j \stackrel{ind}{\sim} (1 - \gamma_j)\mathcal{N}(0, \sigma^2\nu_0) + \gamma_j\mathcal{N}(0, \sigma^2\nu_1),$$

where ν_0 and ν_1 will be chosen to be small and large, respectively. Note that the likelihood is independent of $\boldsymbol{\gamma} = (\gamma_1, \dots, \gamma_p)$. Assume

$$\sigma^2 \sim \mathcal{IG}\left(\frac{a}{2}, \frac{b}{2}\right),$$

which is also independent of $\boldsymbol{\gamma}$. We consider

$$\gamma_j \stackrel{iid}{\sim} \text{Ber}(\omega).$$

To make our model robust to the choice of ω , we will assign the following prior on ω .

$$\omega \sim \mathcal{B}(c_1, c_2),$$

where we will use $c_1 = c_2 = 1$, which leads to the uniform distribution. Recall the density function of beta distribution is proportional $\pi(w) \propto w^{c_1-1}(1-w)^{c_2-1}$.

It is easy to show that the **full conditionals** are as follows:

- 1)

$$\beta_j | \boldsymbol{\beta}_{-j}, \sigma^2, \boldsymbol{\gamma}, \mathbf{y} \sim \mathcal{N}(\tilde{\beta}_j, \frac{\sigma^2}{\mu_j}).$$

where $\mu_j = x_j^\top x_j + \frac{1}{\nu_{\gamma_j}}$ and $\tilde{\beta}_j = \mu_j^{-1} x_j^\top (\mathbf{y} - \mathbf{X}_{-j} \boldsymbol{\beta}_{-j})$.

- 2)

$$\gamma_j | \gamma_{-j}, \boldsymbol{\beta}, \sigma^2, \mathbf{y} \sim \text{Ber} \left(\frac{\nu_1^{-\frac{1}{2}} \exp \left(-\frac{1}{2\sigma^2\nu_1} \beta_j^2 \right) \omega}{\nu_0^{-\frac{1}{2}} \exp \left(-\frac{1}{2\sigma^2\nu_0} \beta_j^2 \right) (1-\omega) + \nu_1^{-\frac{1}{2}} \exp \left(-\frac{1}{2\sigma^2\nu_1} \beta_j^2 \right) \omega} \right).$$

- 3)

$$\sigma^2 | \boldsymbol{\beta}, \boldsymbol{\gamma}, \mathbf{y} \sim \mathcal{IG}(a^*, b^*),$$

where $a^* = \frac{1}{2}(n + p + a)$ and $b^* = \frac{1}{2} \left(\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2 + \sum_{j=1}^p \frac{\beta_j^2}{\nu_{\gamma_j}} + b \right)$.

- 4)

$$w | \boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma^2, \mathbf{y} \sim \mathcal{B} \left(\sum_{j=1}^p \gamma_j + c_1, p - \sum_{j=1}^p \gamma_j + c_2 \right).$$

To speed up, we consider the following conditionals:

- 1')

$$\boldsymbol{\beta} | \sigma^2, \boldsymbol{\gamma}, \mathbf{y} \sim \mathcal{N}(\tilde{\boldsymbol{\beta}}, \sigma^2 (\mathbf{X}^\top \mathbf{X} + \mathbf{V}_{\boldsymbol{\gamma}}^{-1})^{-1}),$$

where $\mathbf{V}_{\boldsymbol{\gamma}} = \text{diag}(v_{\gamma_j})_{j=0}^p$ and $\tilde{\boldsymbol{\beta}} = (\mathbf{X}^\top \mathbf{X} + \mathbf{V}_{\boldsymbol{\gamma}}^{-1})^{-1} \mathbf{X}^\top \mathbf{y}$.

- 2')

$$\gamma | \boldsymbol{\beta}, \sigma^2, \mathbf{y} \sim \prod_{j=0}^p \text{Ber} \left(\frac{\nu_1^{-\frac{1}{2}} \exp \left(-\frac{1}{2\sigma^2\nu_1} \beta_j^2 \right) \omega}{\nu_0^{-\frac{1}{2}} \exp \left(-\frac{1}{2\sigma^2\nu_0} \beta_j^2 \right) (1-\omega) + \nu_1^{-\frac{1}{2}} \exp \left(-\frac{1}{2\sigma^2\nu_1} \beta_j^2 \right) \omega} \right).$$

- 3')

$$\sigma^2 | \boldsymbol{\beta}, \boldsymbol{\gamma}, \mathbf{y} \sim \mathcal{IG}(a^*, b^*).$$

where $a^* = \frac{1}{2}(n + p + a)$ and $b^* = \frac{1}{2} \left(\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2 + \sum_{j=1}^p \frac{\beta_j^2}{\nu_{\gamma_j}} + b \right)$.

- 4')

$$w | \boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma^2, \mathbf{y} \sim \mathcal{B} \left(\sum_{j=1}^p \gamma_j + c_1, p - \sum_{j=1}^p \gamma_j + c_2 \right).$$

2.1 Rcode

```

library(invgamma)
p <- 100
n = 100
power <- numeric()
e<-rnorm(n, mean = 0, sd = sqrt(2)) # error
X<-matrix(data = rnorm(n * p, 0, 1), nrow = n, ncol = p)
Beta <- c(1, 2, rep(0,(p - 3)), 3)# exclude intercept beta0
y <- X %*% Beta + e # true model
true.gamma<-as.numeric(Beta != 0)
#setup for initial values####
hat.beta <- as.numeric(solve(t(X) %*% (X) + diag(1, p)) %*% t(X) %*% y) #p-dim vector
hat.gamma <- rep(1, p)
hat.sig2 <- mean((y - X %*% hat.beta)^2)
#setup for priors #####
w <- 0.5
v0 <- 0.001
v1 <- 1000
v01 <- c(v0, v1)
a0 <- 1
b0 <- 1
#####
MC.size <- 2000 + 3000
hat.BETA <- matrix(0, MC.size, p) # to store beta for each iteration
hat.Gamma <- matrix(0, MC.size, p) # to store z for each iteration
hat.Sig2 <- rep(0, MC.size) # to store variance for each iteration
for (goh in 1:MC.size) {
  # Gibbs sampling
  # 1) for beta_j
  for (j in 1:p) {
    mu_j <- t(X[, j]) %*% X[, j] + 1/v01[(hat.gamma[j] + 1)]
    y.star <- y - X[, -j] %*% hat.beta[-j]
    tilde.beta.j <- as.numeric((1/mu_j) * t(X[, j]) %*% y.star)
    var.beta <- as.numeric(hat.sig2/mu_j)
    hat.beta[j] <- rnorm(1, tilde.beta.j, sqrt(var.beta)) #sampling from beta_j/others
  }
  # 2) gamma_j
  p.j <- dnorm(hat.beta, 0, sqrt(v1 * hat.sig2)) * w
  q.j <- dnorm(hat.beta, 0, sqrt(v0 * hat.sig2)) * (1 - w)
  prob.j <- p.j/(p.j + q.j)
  hat.gamma <- rbinom(p, 1, prob.j)
  hat.Gamma[goh, ] <- hat.gamma
  hat.BETA[goh, ] <- hat.beta
  # 3) sig2

```

```

a.star <- 1/2 * (n + p + a0)
v.z_j <- hat.gamma * v1 + (1 - hat.gamma) * v0
b.star <- 1/2 * (sum((y - X %*% hat.beta)^2) + sum(hat.beta^2/v.z_j) + b0)
hat.sig2 <- rinvgamma(1, shape = a.star, rate = b.star)
par(mfrow=c(1,1))
plot(hat.gamma, main = paste("rep:", goh))
points(true.gamma, col = 2, pch = "*")
}
colMeans(hat.Gamma)>0.5

```

2.2 Appendix

- For $\beta_j | \beta_{-j}, \sigma^2, \gamma, \mathbf{y}$ ($j = 1, 2, \dots, p$), we have

$$\begin{aligned}
\pi(\beta_j | \beta_{-j}, \sigma^2, \gamma, \mathbf{y}) &\propto f(\mathbf{y} | \beta, \sigma^2) \pi(\beta | \gamma, \sigma^2) \\
&= f(\mathbf{y} | \beta_{\gamma}, \sigma^2) \prod_{k=1}^p \pi(\beta_k | Z_k, \sigma^2) \\
&\propto f(\mathbf{y} | \beta, \sigma^2) \pi(\beta_j | \gamma_j, \sigma^2) \\
&\propto \exp\left(-\frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}\beta\|^2\right) \exp\left(-\frac{1}{2\sigma^2 \nu_{\gamma_j}} \beta_j^2\right) \\
&= \exp\left(-\frac{1}{2\sigma^2} \|\mathbf{y} - \mathbf{X}_{-j}\beta_{-j} - x_j\beta_j\|^2\right) \exp\left(-\frac{1}{2\sigma^2 \nu_{\gamma_j}} \beta_j^2\right) \\
&= \exp\left(-\frac{1}{2\sigma^2} \|\mathbf{y}^* - x_j\beta_j\|^2\right) \exp\left(-\frac{1}{2\sigma^2 \nu_{\gamma_j}} \beta_j^2\right) \\
&= \exp\left[-\frac{1}{2\sigma^2} \left(\mathbf{y}^{*\top} \mathbf{y}^* - 2\beta_j x_j^\top \mathbf{y}^* + \beta_j x_j^\top x_j \beta_j + \frac{1}{\nu_{\gamma_j}} \beta_j^2\right)\right] \\
&\propto \exp\left[-\frac{1}{2\sigma^2} \left(\beta_j^2 (x_j^\top x_j + \frac{1}{\nu_{\gamma_j}}) - 2\beta_j x_j^\top \mathbf{y}^*\right)\right] \\
&= \exp\left[-\frac{1}{2\sigma^2} \left(\beta_j^2 (x_j^\top x_j + \frac{1}{\nu_{\gamma_j}}) - 2\beta_j (x_j^\top x_j + \frac{1}{\nu_{\gamma_j}}) (x_j^\top x_j + \frac{1}{\nu_{\gamma_j}})^{-1} x_j^\top \mathbf{y}^*\right)\right] \\
&= \exp\left[-\frac{a}{2\sigma^2} (\beta_j^2 - 2a\beta_j \tilde{\beta}_j)\right] \\
&\propto \exp\left[-\frac{a}{2\sigma^2} (\beta_j - \tilde{\beta}_j)^2\right]
\end{aligned}$$

where $\mathbf{y}^* = \mathbf{y} - \mathbf{X}_{-j}\beta_{-j}$, $\mu_j = x_j^\top x_j + \frac{1}{\nu_{\gamma_j}}$, $\tilde{\beta}_j = \mu_j^{-1} x_j^\top \mathbf{y}^*$. Hence,

$$\beta_j | \beta_{-j}, \sigma^2, \gamma, \mathbf{y} \sim \mathcal{N}(\tilde{\beta}_j, \frac{\sigma^2}{\mu_j}).$$

- For $\pi(\gamma_j|\gamma_{-j}, \beta, \sigma^2, \mathbf{y})$ ($j = 1, 2, \dots, p$) we have

$$\begin{aligned}
\pi(\gamma_j|\gamma_{-j}, \beta, \sigma^2, \mathbf{y}) &\propto \pi(\gamma, \beta|\sigma^2) \\
&= \pi(\beta|\sigma^2, \gamma)\pi(\gamma) \\
&= \prod_{k=1}^p \pi(\beta_k|\sigma^2, z_k) \prod_{i=1}^p \pi(z_i) \\
&\propto \pi(\beta_j|\sigma^2, z_j)\pi(\gamma_j) \\
&\propto \nu_{\gamma_j}^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_{\gamma_j}}\beta_j^2\right) \omega^{\gamma_j}(1-\omega)^{1-\gamma_j}.
\end{aligned}$$

Note that

$$\begin{aligned}
\pi(\gamma_j = 0|\gamma_{-j}, \beta, \sigma^2, \mathbf{y}) &= C\nu_0^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_0}\beta_j^2\right) (1-\omega); \\
\pi(\gamma_j = 1|\gamma_{-j}, \beta, \sigma^2, \mathbf{y}) &= C\nu_1^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_1}\beta_j^2\right) \omega.
\end{aligned}$$

This implies that

$$\begin{aligned}
\pi(\gamma_j = 1|\gamma_{-j}, \beta, \sigma^2, \mathbf{y}) &= \frac{P(\gamma_j = 1, \Omega)}{\sum_{\gamma_j} P(\gamma_j, \Omega)} \\
&= \frac{C\nu_1^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_1}\beta_j^2\right) \omega}{C\nu_0^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_0}\beta_j^2\right) (1-\omega) + C\nu_1^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_1}\beta_j^2\right) \omega} \\
&= \frac{\nu_1^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_1}\beta_j^2\right) \omega}{\nu_0^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_0}\beta_j^2\right) (1-\omega) + \nu_1^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_1}\beta_j^2\right) \omega}.
\end{aligned}$$

Hence,

$$\gamma_j|\gamma_{-j}, \beta, \sigma^2, \mathbf{y} \sim Ber\left(\frac{\nu_1^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_1}\beta_j^2\right) \omega}{\nu_0^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_0}\beta_j^2\right) (1-\omega) + \nu_1^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_1}\beta_j^2\right) \omega}\right).$$

- For $\sigma^2|\beta, \gamma, \mathbf{y}$, we have

$$\begin{aligned}
\pi(\sigma^2|\boldsymbol{\beta}, \boldsymbol{\gamma}, \mathbf{y}) &\propto f(\mathbf{y}|\boldsymbol{\beta}, \sigma^2, \boldsymbol{\gamma})\pi(\boldsymbol{\beta}, \sigma^2|\boldsymbol{\gamma}) \\
&= f(\mathbf{y}|\boldsymbol{\beta}, \sigma^2, \boldsymbol{\gamma})\pi(\boldsymbol{\beta}|\sigma^2, \boldsymbol{\gamma})\pi(\sigma^2|\boldsymbol{\gamma}) \\
&= f(\mathbf{y}|\boldsymbol{\beta}, \sigma^2)\pi(\boldsymbol{\beta}|\sigma^2, \boldsymbol{\gamma})\pi(\sigma^2) \\
&\propto (\sigma^2)^{-\frac{n}{2}} \exp\left(-\frac{1}{2\sigma^2}\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2\right) \times \prod_{j=1}^p \left[(\sigma^2\nu_{\gamma_j})^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2\nu_{\gamma_j}}\beta_j^2\right)\right] \\
&\quad \times (\sigma^2)^{-\frac{a}{2}-1} \exp\left(-\frac{b}{2\sigma^2}\right) \\
&\propto (\sigma^2)^{-\frac{n}{2}} \exp\left(-\frac{1}{2\sigma^2}\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2\right) \times (\sigma^2)^{-\frac{p}{2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{j=1}^p \frac{\beta_j^2}{\nu_{\gamma_j}}\right) \\
&\quad \times (\sigma^2)^{-\frac{a}{2}-1} \exp\left(-\frac{b}{2\sigma^2}\right) \\
&= (\sigma^2)^{-\frac{1}{2}(n+p+a)-1} \exp\left(-\frac{\frac{1}{2}(\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2 + \sum_{j=1}^p \frac{\beta_j^2}{\nu_{\gamma_j}} + b)}{\sigma^2}\right) \\
&= (\sigma^2)^{-a^*-1} \exp\left(-\frac{b^*}{\sigma^2}\right),
\end{aligned}$$

where $a^* = \frac{1}{2}(n + p + a)$ and $b^* = \frac{1}{2}(\|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2 + \sum_{j=1}^p \frac{\beta_j^2}{\nu_{\gamma_j}} + b)$.

Therefore,

$$\sigma^2|\boldsymbol{\beta}, \boldsymbol{\gamma}, \mathbf{y} \sim \mathcal{IG}(a^*, b^*).$$

- For $w|\boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma^2, \mathbf{y}$, we have

$$\begin{aligned}
\pi(w|\boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma^2, \mathbf{y}) &\propto \pi(\boldsymbol{\gamma}|w)\pi(w) \\
&\propto \left[\prod_{j=1}^p w^{\gamma_j} (1-w)^{1-\gamma_j} \right] w^{c_1-1} (1-w)^{c_2-1} \\
&\propto w^{\sum_{j=1}^p \gamma_j + c_1 - 1} (1-w)^{p - \sum_{j=1}^p \gamma_j + c_2 - 1}.
\end{aligned}$$

We therefore have

$$w|\boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma^2, \mathbf{y} \sim \mathcal{B}\left(\sum_{j=1}^p \gamma_j + c_1, p - \sum_{j=1}^p \gamma_j + c_2\right).$$

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