

HOMWORK 3:

WRITTEN EXERCISE PART

1 Multinomial Naïve Bayes [25/2 pts]

Consider the Multinomial Naïve Bayes model. For each point (\mathbf{x}, y) , $y \in \{0, 1\}$, $\mathbf{x} = (x_1, x_2, \dots, x_M)$ where each x_j is an integer from $\{1, 2, \dots, K\}$ for $1 \leq j \leq M$. Here K and M are two fixed integer.

Suppose we have N data points $\{(\mathbf{x}^{(i)}, y^{(i)}) : 1 \leq i \leq N\}$, generated as follows.

for $i \in \{1, \dots, N\}$:
 $y^{(i)} \sim \text{Bernoulli}(\phi)$
 for $j \in \{1, \dots, M\}$:
 $x_j^{(i)} \sim \text{Multinomial}(\theta_{y^{(i)}}, 1)$

Here $\phi \in \mathbb{R}$ and $\theta_k \in \mathbb{R}^K$ ($k \in \{0, 1\}$) are parameters. Note that $\sum_l \theta_{k,l} = 1$ since they are the parameters of a multinomial distribution.

Derive the formula for estimating the parameters ϕ and θ_k , as we have done in the lecture for the Bernoulli Naïve Bayes model. Show the steps.

let, $x^i = \theta_{p,k}$
and, A_k^i is frequency of x^i be k

As y is a Bernoulli distribution, hence

$$p(y|\phi) = \prod_{i=1}^n \phi^{y^{(i)}} (1 - \phi)^{1-y^{(i)}}$$

now, ϕ can be computed by minimizing the above expression.

as the expression is a product, hence we can take log on both sides for our convenience.

computing the derivative of log to be 0, we get

$$\begin{aligned} \frac{\partial \log(p(y|\phi))}{\partial \phi} &= \frac{\partial}{\partial \phi} (-\sum_{i=1}^n (y^{(i)} \log(\phi) + (1 - y^{(i)}) \log(1 - \phi))) \\ \implies -\sum_{i=1}^n \frac{y^{(i)}}{\phi} - \frac{(1-y^{(i)})}{(1-\phi)} &= 0 \end{aligned}$$

$$\implies \phi = \frac{\sum_{i=1}^n I(y^{(i)})}{N} = 1$$

now,

$$p(y|x, \theta, \phi) \propto p(y|\phi) p(x|\theta, \phi)$$

as per above derivation,

$$p(y|\phi) = \prod_{i=1}^n p(y^i|\phi)$$

also,

$$p(x^i|\theta) = \prod_{k=1}^K \theta_{y^{(i)},k}^{A_k^i}$$

so, posterior distribution of following has to be estimated

$$p(y|x, \theta, \phi) = \prod_{i=1}^n p(y^i|\phi) \prod_{k=1}^K \theta_{y^{(i)},k}^{A_k^i}$$

provided, $\sum_{k=1}^K \theta_{p,k} = 1$

$$\implies -\log(\text{posterior}(L)) = [\sum_{i=1}^n y^{(i)} \log(\phi) + (1 - y^{(i)}) \log(1 - \phi)] \sum_{k=1}^K A_k^i \log(\theta_{y^{(i)},k})$$

$$\implies = [\sum_{i=1}^n y^{(i)} \log(\phi) + (1 - y^{(i)}) \log(1 - \phi) \sum_{k=1}^K A_k^i \log(\theta_{y^{(i)},k})] - \lambda_0(1 - \sum_{k=1}^K \theta_{0,k}) - \lambda_1(1 - \sum_{k=1}^K \theta_{1,k})$$

summarizing above, we have taken added the Adding Lagrange multiplier
now, for best estimation, we need to minimize this expression

$\frac{\partial L}{\partial \theta}$ can be subdivided into

$\frac{\partial L}{\partial \theta_{0,k}}$; subject to $y^{(i)} = 0$, and $\frac{\partial L}{\partial \theta_{1,k}}$ subject to $y^{(i)} = 1$

$$\begin{aligned} \text{now, } \frac{\partial L}{\partial \theta_{0,k}} &= \frac{\partial}{\partial \theta_{0,k}} [\sum_{i=1}^n y^{(i)} \log(\phi) + (1 - y^{(i)}) \log(1 - \phi) \sum_{k=1}^K A_k^i \log(\theta_{y^{(i)},k})] + \lambda_0(1 - \sum_{k=1}^K \theta_{0,k}) + \lambda_1(1 - \sum_{k=1}^K \theta_{1,k}) \\ \implies &= \frac{\partial}{\partial \theta_{0,k}} \sum_{i=1}^n 0 * \log(\phi) + (1 - 0) \log(1 - \phi) \sum_{k=1}^K A_k^i \log(\theta_{y^{(i)},k}) + \lambda_0(1 - \sum_{k=1}^K \theta_{0,k}) + \lambda_1(1 - \sum_{k=1}^K \theta_{1,k}) \\ \implies &= \sum_{i=1, y^{(i)}=0}^n \phi * \frac{A_k^i}{\theta_{0,k}} - \lambda_0 = 0 \end{aligned}$$

$$\implies \theta_{0,k} = \sum_{i=1, y^{(i)}=0}^n \phi * \frac{A_k^i}{\lambda_0}$$

using, $\sum_{k=1}^K \theta_{p,k} = 1$

$$\implies \sum_{i=1, y^{(i)}=0}^n \sum_{k=1}^K \phi * \frac{A_k^i}{\lambda_0} = 1$$

now, as A_k^i is frequency of x^i , therefore
 $\sum_{k=1}^K A_k^i = M$

$$\implies \frac{\lambda_0}{\phi} = M \sum_{i=1}^n I(y^i = 0)$$

as derived before, $\theta_{0,k} = \sum_{i=1, y^{(i)}=0}^n \phi * \frac{A_k^i}{\lambda_0}$

$$\implies = \frac{\sum_{i=1, y^{(i)}=0}^n A_k^i}{M \sum_{i=1}^n I(y^i = 0)}$$

$$\implies \theta_{0,k} = \frac{\sum_{i=1}^n \sum_{j=1}^M I(y^{(i)} = 0, x_j^i = k)}{M \sum_{i=1}^n I(y^{(i)} = 0)}$$

Now, for computation of $\theta_{1,k}$ we need following replacements

$0 \rightarrow 1$

$\phi \rightarrow (1 - \phi)$

so, $\frac{\partial L}{\partial \theta_{1,k}}$

$$\begin{aligned} &= \frac{\partial}{\partial \theta_{1,k}} [\sum_{i=1}^n y^{(i)} \log(\phi) + (1 - y^{(i)}) \log(1 - \phi) \sum_{k=1}^K A_k^i \log(\theta_{y^{(i)},k})] + \lambda_0(1 - \sum_{k=1}^K \theta_{0,k}) + \lambda_1(1 - \sum_{k=1}^K \theta_{1,k}) \\ \implies &= \frac{\partial}{\partial \theta_{1,k}} \sum_{i=1}^n 1 * \log(\phi) + (1 - 1) \log(1 - \phi) \sum_{k=1}^K A_k^i \log(\theta_{y^{(i)},k}) + \lambda_0(1 - \sum_{k=1}^K \theta_{0,k}) + \lambda_1(1 - \sum_{k=1}^K \theta_{1,k}) \\ \implies &= \sum_{i=1, y^{(i)}=1}^n (1 - \phi) * \frac{A_k^i}{\theta_{1,k}} - \lambda_1 = 0 \end{aligned}$$

$$\implies \theta_{1,k} = \sum_{i=1, y^{(i)}=1}^n (1 - \phi) * \frac{A_k^i}{\lambda_1}$$

following similar steps, we can compute,

$$\Rightarrow \theta_{1,k} = \frac{\sum_{i=1}^n \sum_{j=1}^M I(y^{(i)} = 1, x_j^i = k)}{M \sum_{i=1}^n I(y^{(i)} = 1)}$$

2 Logistic Regression [25/2 pts]

Suppose for each class $i \in \{1, \dots, K\}$, the class-conditional density $p(\mathbf{x}|y = i)$ is normal with mean $\mu_i \in \mathbb{R}^d$ and identity covariance:

$$p(\mathbf{x}|y = i) = N(\mathbf{x}|\mu_i, \mathbf{I}).$$

Prove that $p(y = i|\mathbf{x})$ is a softmax over a linear transformation of \mathbf{x} . Show the steps.

We are given that class-conditional density, is normal.

$$\Rightarrow \text{class-conditional density} = p(\mathbf{x}|y = i)$$

$$\Rightarrow = N(x|\mu_i, I)$$

$$\Rightarrow = \frac{1}{2\pi^{d/2}} \exp\left(-\frac{1}{2}\|x - \mu_i\|^2\right)$$

assume that, $m_i = \log(p(\mathbf{x}|y = i)p(y = i))$

$$\Rightarrow = \frac{-1}{2}\|x - \mu_i\|^2 + \log \frac{1}{2\pi^{d/2}} + \log(p(y = i))$$

and, we know,

$$\frac{-1}{2}\|x - \mu_i\|^2 = \frac{-1}{2}x^T x + \mu_i^T x + \frac{-1}{2}\mu_i^T \mu_i$$

$$\Rightarrow m_i = \frac{-1}{2}x^T x + \mu_i^T x + \frac{-1}{2}\mu_i^T \mu_i + \log \frac{1}{2\pi^{d/2}} + \log(p(y = i))$$

assume,

$$n_i = \frac{-1}{2}\mu_i^T \mu_i + \log \frac{1}{2\pi^{d/2}} + \log(p(y = i))$$

$$\Rightarrow m_i = \frac{-1}{2}x^T x + (w^i)^T x + n_i$$

now, using the bayesian rule, $p(y = i|\mathbf{x})$ is :

$$p(y = i|\mathbf{x}) = \frac{p(\mathbf{x}|y = i)p(y = i)}{\sum_j p(\mathbf{x}|y = j)p(y = j)}$$

$$\text{hence, } p(y = i|\mathbf{x}) = \frac{e^{m_i}}{\sum_j e^{m_j}}$$

$$\Rightarrow = \frac{e^{\left(\frac{-1}{2}x^T x + (w^i)^T x + n_i\right)}}{\sum_j e^{\left(\frac{-1}{2}x^T x + (w^j)^T x + n_j\right)}}$$

$$\Rightarrow = \frac{e^{((w^i)^T x + n_i)}}{\sum_j e^{((w^j)^T x + n_j)}}$$

As this is a softmax over linear transformation of \mathbf{x} , Hence Proved.