

Fitting a Linear Mixed-Effects Model using Expectation-Maximization

Camden Lopez

2022-06-15

Problem Statement

Assume that for $i = 1, \dots, n$,

- $Y_i = X_i\beta + Z_ib_i + \epsilon_i$
- X_i is a fixed $n_i \times p$ matrix
- Z_i is a fixed $n_i \times q$ matrix
- $b_i \sim N_q(0, \Sigma_b)$,
- $\epsilon_i \sim N_{n_i}(0, \sigma^2 I_{n_i})$
- b_i and ϵ_i are independent
- (b_i, ϵ_i) and (b_j, ϵ_j) are independent for all $i \neq j$

Given (X_i, Z_i, Y_i) for $i = 1, \dots, n$, we want to estimate β , Σ_b , and σ^2 using the expectation-maximization (EM) algorithm.

The likelihood function is

$$L(\beta, \Sigma_b, \sigma^2) = \prod_{i=1}^n f(y_i|b_i, \beta, \sigma^2)g(b_i|\Sigma_b)$$

where f is the density of y_i conditional on b_i , and g is the density of b_i .

$$f(y_i|b_i, \beta, \sigma^2) = (2\pi)^{-n_i/2}(\sigma^2)^{-n_i/2} \exp \left[-\frac{1}{2\sigma^2}(y_i - X_i\beta - Z_ib_i)^\top (y_i - X_i\beta - Z_ib_i) \right]$$
$$g(b_i|\Sigma_b) = (2\pi)^{-q/2}|\Sigma_b|^{-1/2} \exp \left[-\frac{1}{2}b_i^\top \Sigma_b^{-1} b_i \right]$$

The log-likelihood is

$$\begin{aligned}
l(\beta, \Sigma_b, \sigma^2) &= \sum_{i=1}^n \left[\log f(y_i | b_i, \beta, \sigma^2) + \log g(b_i | \Sigma_b) \right] \\
&= \sum_{i=1}^n \left[-\frac{n_i}{2} \log(2\pi) - \frac{n_i}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} (y_i - X_i\beta - Z_i b_i)^\top (y_i - X_i\beta - Z_i b_i) \right. \\
&\quad \left. - \frac{q}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_b| - \frac{1}{2} b_i^\top \Sigma_b^{-1} b_i \right]
\end{aligned}$$

E-Step Derivation

For the EM algorithm, first we set some initial values of the parameters, $\beta^{(0)}$, $\Sigma_b^{(0)}$, and $\sigma^{2(0)}$. Then, given current estimates $\beta^{(k)}$, $\Sigma_b^{(k)}$, and $\sigma^{2(k)}$, we calculate $Q(\beta, \Sigma_b, \sigma^2 | \beta^{(k)}, \Sigma_b^{(k)}, \sigma^{2(k)})$, the expectation of $l(\beta, \Sigma_b, \sigma^2)$ over the distribution of $b = (b_1, \dots, b_n)$ conditional on $y = (y_1, \dots, y_n)$.

The conditional distribution of $b_i | y_i$ has pdf $f(y_i | b_i)g(b_i)/h(y_i)$ where h is the marginal pdf of y_i . We find the distribution of $b_i | y_i$ by looking at the kernel of $f(y_i | b_i)g(b_i)/h(y_i)$:

$$\begin{aligned}
\frac{f(y_i | b_i)g(b_i)}{h(y_i)} &\propto \exp \left\{ -\frac{1}{2\sigma^2} (y_i - X_i\beta - Z_i b_i)^\top (y_i - X_i\beta - Z_i b_i) - \frac{1}{2} b_i^\top \Sigma_b^{-1} b_i \right\} \\
&\quad (\text{letting } r_i = y_i - X_i\beta) \\
&= \exp \left\{ -\frac{1}{2} \left[\frac{1}{\sigma^2} r_i^\top r_i - \frac{2}{\sigma^2} b_i^\top Z_i^\top r_i + \frac{1}{\sigma^2} b_i^\top Z_i^\top Z_i b_i + b_i^\top \Sigma_b^{-1} b_i \right] \right\} \\
&= \exp \left\{ -\frac{1}{2} \left[b_i^\top \left(\Sigma_b^{-1} + \frac{1}{\sigma^2} Z_i^\top Z_i \right) b_i - \frac{2}{\sigma^2} b_i^\top Z_i^\top r_i + \frac{1}{\sigma^2} r_i^\top r_i \right] \right\}
\end{aligned}$$

Now we complete the square:

$$x^\top A x + x^\top b + c = (x - u)^\top A (x - u) + v, \text{ where } u = -\frac{1}{2} A^{-1} b \text{ and } v = c - \frac{1}{4} b^\top A^{-1} b.$$

We have

$$\begin{aligned}
A &= \Sigma_b^{-1} + \frac{1}{\sigma^2} Z_i^\top Z_i \\
b &= -\frac{2}{\sigma^2} Z_i^\top r_i \\
\Rightarrow u &= -\frac{1}{2} \left(\Sigma_b^{-1} + \frac{1}{\sigma^2} Z_i^\top Z_i \right)^{-1} \left(-\frac{2}{\sigma^2} Z_i^\top r_i \right) \\
&= \frac{1}{\sigma^2} \left(\Sigma_b^{-1} + \frac{1}{\sigma^2} Z_i^\top Z_i \right)^{-1} Z_i^\top r_i
\end{aligned}$$

We don't need to calculate v because it won't involve b_i .

Therefore,

$$\frac{f(y_i|b_i)g(b_i)}{h(y_i)} \propto \exp \left\{ -\frac{1}{2}[b_i - u_i]^\top A_i [b_i - u_i] \right\}$$

which means that $b_i|y_i \sim N(u_i, A_i^{-1})$ where u_i and A_i are defined above.

Now, to obtain the Q function from the log-likelihood, we'll use the following expectations:

$$\begin{aligned} \mathbb{E}_{b_i|y_i} b_i^\top Z_i^\top r_i &= u_i^\top Z_i^\top r_i \\ \mathbb{E}_{b_i|y_i} b_i^\top Z_i^\top Z_i b_i &= \text{tr}(Z_i^\top Z_i A_i^{-1}) + u_i^\top Z_i^\top Z_i u_i \\ &= \text{tr}(Z_i A_i^{-1} Z_i^\top) + u_i^\top Z_i^\top Z_i u_i \\ \mathbb{E}_{b_i|y_i} b_i^\top \Sigma_b^{-1} b_i &= \text{tr}(\Sigma_b^{-1} A_i^{-1}) + u_i^\top \Sigma_b^{-1} u_i \end{aligned}$$

We have

$$\begin{aligned} &Q(\beta, \Sigma_b, \sigma^2 | \beta^{(k)}, \Sigma_b^{(k)}, \sigma^{2(k)}) \\ &= \sum_{i=1}^n \left[-\frac{n_i}{2} \log(2\pi) - \frac{n_i}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} r_i^\top r_i + \frac{1}{\sigma^2} u_i^\top Z_i^\top r_i \right. \\ &\quad \left. - \frac{1}{2\sigma^2} \text{tr}(Z_i A_i^{-1} Z_i^\top) - \frac{1}{2\sigma^2} u_i^\top Z_i^\top Z_i u_i \right. \\ &\quad \left. - \frac{q}{2} \log(2\pi) - \frac{1}{2} \log |\Sigma_b| \right. \\ &\quad \left. - \frac{1}{2} \text{tr}(\Sigma_b^{-1} A_i^{-1}) - \frac{1}{2} u_i^\top \Sigma_b^{-1} u_i \right] \\ &= C - \frac{N}{2} \log(\sigma^2) - \frac{n}{2} \log |\Sigma_b| \\ &\quad - \frac{1}{2\sigma^2} \sum_{i=1}^n r_i^\top r_i + \frac{1}{\sigma^2} \sum_{i=1}^n u_i^\top Z_i^\top r_i - \frac{1}{2\sigma^2} \sum_{i=1}^n u_i^\top Z_i^\top Z_i u_i \\ &\quad - \frac{1}{2\sigma^2} \sum_{i=1}^n \text{tr}(Z_i A_i^{-1} Z_i^\top) \\ &\quad - \frac{1}{2} \sum_{i=1}^n \text{tr}(\Sigma_b^{-1} A_i^{-1}) - \frac{1}{2} \sum_{i=1}^n u_i^\top \Sigma_b^{-1} u_i \end{aligned}$$

where C is a constant, and $N = \sum_{i=1}^n n_i$.

M-Step Derivation

For the maximization step, we find $\beta, \Sigma_b, \sigma^2$ that maximize Q .

Setting the gradient of Q wrt β equal to 0 and solving, we have

$$\begin{aligned}
\nabla_{\beta} Q &= 0 \\
\Rightarrow 0 &= \frac{1}{\sigma^2} \sum_{i=1}^n X_i^T (y_i - X_i \beta) - \frac{1}{\sigma^2} \sum_{i=1}^n X_i^T Z_i u_i \\
\Rightarrow 0 &= \sum_{i=1}^n X_i^T y_i - \sum_{i=1}^n (X_i^T X_i) \beta - \sum_{i=1}^n X_i^T Z_i u_i \\
\Rightarrow \beta &= \left(\sum_{i=1}^n X_i^T X_i \right)^{-1} \sum_{i=1}^n X_i^T (y_i - Z_i u_i) \\
&= \text{OLS estimate after subtracting predicted random effects from } y_i
\end{aligned}$$

Setting the gradient of Q wrt Σ_b equal to 0 and solving, we have

Reference for derivative-wrt-matrix identities:

<https://www.ics.uci.edu/~welling/teaching/KernelsICS273B/MatrixCookBook.pdf>

$$\begin{aligned}
\nabla_{\Sigma_b} Q &= 0 \\
\Rightarrow 0 &= -\frac{n}{2} \Sigma_b^{-1} + \frac{1}{2} \sum_{i=1}^n (\Sigma_b^{-1} A_i^{-1} \Sigma_b^{-1}) + \frac{1}{2} \sum_{i=1}^n (\Sigma_b^{-1} u_i u_i^T \Sigma_b^{-1}) \\
\Rightarrow n \Sigma_b \Sigma_b^{-1} \Sigma_b &= \sum_{i=1}^n \Sigma_b \Sigma_b^{-1} u_i u_i^T \Sigma_b^{-1} \Sigma_b + \sum_{i=1}^n \Sigma_b \Sigma_b^{-1} A_i^{-1} \Sigma_b^{-1} \Sigma_b \\
\Rightarrow \Sigma_b &= \frac{1}{n} \sum_{i=1}^n u_i u_i^T + \frac{1}{n} \sum_{i=1}^n A_i^{-1} \\
&= \text{Covariance of conditional means of } b_i + \text{Mean of conditional covariances of } b_i
\end{aligned}$$

Finally, setting the gradient of Q wrt σ^2 , we have

$$\begin{aligned}
\nabla_{\sigma^2} Q &= 0 \\
\Rightarrow 0 &= -\frac{N}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n r_i^\top r_i - \frac{1}{2(\sigma^2)^2} 2 \sum_{i=1}^n r_i^\top Z_i u_i + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n u_i^\top Z_i^\top Z_i u_i \\
&\quad + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n \text{tr}(Z_i A_i^{-1} Z_i^\top) \\
\Rightarrow N\sigma^2 &= \sum_{i=1}^n (r_i - Z_i u_i)^\top (r_i - Z_i u_i) + \sum_{i=1}^n \text{tr}(Z_i A_i^{-1} Z_i^\top) \\
\Rightarrow \sigma^2 &= \frac{1}{N} \sum_{i=1}^n (y_i - X_i \beta - Z_i u_i)^\top (y_i - X_i \beta - Z_i u_i) + \frac{1}{N} \sum_{i=1}^n \text{tr}(Z_i A_i^{-1} Z_i^\top) \\
&= \text{Variance from deviations of } y_i \text{ from conditional mean} \\
&\quad + \text{Variance from random effects}
\end{aligned}$$

where we substitute β with the estimate derived above, $(\sum_{i=1}^n X_i^\top X_i)^{-1} \sum_{i=1}^n X_i^\top (y_i - Z_i u_i)$.

Algorithm

Putting this all together, an EM algorithm for estimating β , Σ_b , and σ^2 is

1. Set $\beta^{(0)}$ (e.g. with OLS estimates), $\Sigma_b^{(0)}$ (e.g. with I_q), and $\sigma^{2(0)}$ (e.g. with OLS estimate).
2. At iteration k (E step), given current estimates $\beta^{(k)}$, $\Sigma_b^{(k)}$, and $\sigma^{2(k)}$,
 - Calculate $A_i^{-1} = \left(\Sigma_b^{-1(k)} + \frac{1}{\sigma^{2(k)}} Z_i^\top Z_i \right)^{-1}$, $i = 1, \dots, n$.
 - Calculate $u_i = \frac{1}{\sigma^{2(k)}} A_i^{-1} Z_i^\top (y_i - X_i \beta^{(k)})$, $i = 1, \dots, n$.
3. (Iteration k , M step) Calculate new estimates using u_1, \dots, u_n and $A_1^{-1}, \dots, A_n^{-1}$:

$$\begin{aligned}
\beta^{(k+1)} &= \left(\sum_{i=1}^n X_i^\top X_i \right)^{-1} \sum_{i=1}^n X_i^\top (y_i - Z_i u_i) \\
\Sigma_b^{(k+1)} &= \frac{1}{n} \sum_{i=1}^n u_i u_i^\top + \frac{1}{n} \sum_{i=1}^n A_i^{-1} \\
\sigma^{2(k)} &= \frac{1}{N} \sum_{i=1}^n (y_i - X_i \beta^{(k+1)} - Z_i u_i)^\top (y_i - X_i \beta^{(k+1)} - Z_i u_i) + \frac{1}{N} \sum_{i=1}^n \text{tr}(Z_i A_i^{-1} Z_i^\top)
\end{aligned}$$

4. Continue with iterations of steps 2–3 until convergence.