

CS 446: Machine Learning

Homework 9

Due on Tuesday, April 10, 2018, 11:59 a.m. Central Time

1. **[16 points]** Gaussian Mixture Models & EM

Consider a Gaussian mixture model with K components ($k \in \{1, \dots, K\}$), each having mean μ_k , variance σ_k^2 , and mixture weight π_k . All these are parameters to be learned, and we subsume them in the set θ . Further, we are given a dataset $X = \{x_i\}$, where $x_i \in \mathbb{R}$. We also use $Z = \{z_i\}$ to denote the latent variables, such that $z_i = k$ implies that x_i is generated from the k^{th} Gaussian.

- (a) What is the log-likelihood of the data $\log p(X; \theta)$ according to the Gaussian Mixture Model? (use μ_k , σ_k , π_k , K , x_i , and X). Don't use any abbreviations.

Your answer:

$$\begin{aligned} \log p(x; \theta) &= \log \prod_{i=1}^N p(x_i; \theta) = \sum_{i=1}^N \log p(x_i; \theta) \\ &= \sum_{i=1}^N \log \left(\sum_{k=1}^K \pi_k \mathcal{N}(x_i | \mu_k, \sigma_k) \right) \end{aligned}$$

where,

$$\mathcal{N}(x_i | \mu_k, \sigma_k) = \left(\frac{1}{2\pi\sigma_k^2} \right)^{1/2} \exp \left\{ -\frac{(x_i - \mu_k)^2}{2\sigma_k^2} \right\}$$

- (b) For learning θ using the EM algorithm, we need the conditional distribution of the latent variables Z given the current estimate of the parameters $\theta^{(t)}$ (we will use the superscript (t) for parameter estimates at step t). What is the posterior probability $p(z_i = k | x_i; \theta^{(t)})$? To simplify, wherever possible, use $\mathcal{N}(x_i | \mu_k, \sigma_k)$ to denote a Gaussian distribution over $x_i \in \mathbb{R}$ having mean μ_k and variance σ_k^2 .

Your answer: According bayes theorem:

$$p(A_i | B) = \frac{p(B | A_i) p(A_i)}{\sum_j p(B | A_j) p(A_j)}$$

Here our prior is $p(A_i) : p(z_i = k) = \pi_k$ and conditional is $p(B | A_i) : p(x_i | \theta^{(t)}) = \mathcal{N}(x_i | \mu_k^{(t)}, \sigma_k^{(t)})$.

Hence, our posterior will be

$$p(z_i = k | x_i; \theta^{(t)}) = \frac{\pi_k^{(t)} \mathcal{N}(x_i | \mu_k^{(t)}, \sigma_k^{(t)})}{\sum_{k=1}^K \pi_k^{(t)} \mathcal{N}(x_i | \mu_k^{(t)}, \sigma_k^{(t)})}$$

- (c) Find $\mathbb{E}_{z_i | x_i; \theta^{(t)}} [\log p(x_i, z_i; \theta)]$. Denote $p(z_i = k | x_i; \theta^{(t)})$ as z_{ik} , and use all previous notation simplifications.

Your answer:

$$\begin{aligned}
\mathbb{E}[\log p(x_i, z_i; \theta)] &= \mathbb{E} \left[\log \left(\prod_{k=1}^K (\pi_k p(x_i; \theta_k))^{\mathbb{1}(z_i=k)} \right) \right] \\
&= \sum_{k=1}^K \mathbb{E}[\mathbb{1}(z_i = k)] \log(\pi_k p(x_i; \theta_k)) \\
&= \sum_{k=1}^K p(z_i = k; x_i, \theta^{(t)}) \log[\pi_k p(x_i; \theta_k)] \\
&= \sum_{k=1}^K z_{ik} \log[\pi_k] + \sum_{k=1}^K z_{ik} \log[p(x_i; \theta_k)]
\end{aligned}$$

- (d) $\theta^{(t+1)}$ is obtained as the maximizer of $\sum_{i=1}^N \mathbb{E}_{z_i|x_i; \theta^{(t)}}[\log p(x_i, z_i; \theta)]$. Find $\mu_k^{(t+1)}$, $\pi_k^{(t+1)}$, and $\sigma_k^{(t+1)}$, by using your answer to the previous question.

Your answer: Proof given below,

$$\begin{aligned}
\pi_k^{(t+1)} &= \hat{\pi}_k = \frac{1}{N} \sum_{i=1}^N z_{ik} \\
\mu_k^{(t+1)} &= \hat{\mu}_k = \frac{\sum_{i=1}^N z_{ik} x_i}{\sum_{i=1}^N z_{ik}} \\
\sigma_k^{2(t+1)} &= \hat{\sigma}_k^2 = \frac{\sum_{i=1}^N z_{ik} (x_i - \mu_k)^2}{\sum_{i=1}^N z_{ik}}
\end{aligned}$$

Proof: Auxiliary function:

$$Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) \triangleq \mathbb{E} \left[\sum_{i=1}^N \log p(x_i, z_i; \boldsymbol{\theta}) \right] = \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log[\pi_k] + \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log[p(\mathbf{x}_i; \boldsymbol{\theta}_k)]$$

Such that:

$$\sum_{k=1}^K \pi_k = 1$$

Hence,

$$L(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) : \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log[\pi_k] + \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log[p(\mathbf{x}_i; \boldsymbol{\theta}_k)] + \lambda \left(\sum_{k=1}^K \pi_k - 1 \right)$$

Now taking derivative of $L(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$ w.r.t π_k and equating it to 0,

$$\begin{aligned} \frac{\partial}{\partial \pi_k} L(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) &= 0 \\ \frac{\partial}{\partial \pi_k} \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log[\pi_k] + \frac{\partial}{\partial \pi_k} \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log[p(\mathbf{x}_i; \boldsymbol{\theta}_k)] + \frac{\partial}{\partial \pi_k} \lambda \left(\sum_{k=1}^K \pi_k - 1 \right) &= 0 \\ \sum_{i=1}^N \frac{z_{ik}}{\pi_k} + 0 + \lambda &= 0 \\ \hat{\pi}_k &= \frac{1}{-\lambda} \sum_{i=1}^N z_{ik} \end{aligned}$$

we also know $\sum_{i=1}^N (\lambda \sum_{k=1}^K \pi_k - 1) = 0$ which gives on simplification $\lambda = -N$

Now taking derivative of $L(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$ w.r.t $\boldsymbol{\theta}_k$ and equating it to 0

(considering only middle part of the L i.e neglecting parts that does not have $\boldsymbol{\theta} : \mu, \sigma$)

$$\begin{aligned} \frac{\partial}{\partial \mu_k} L(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) &= 0 \\ \frac{\partial}{\partial \mu_k} \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log \left(\left(\frac{1}{2\pi\sigma_k^2} \right)^{1/2} \exp \left\{ -\frac{(x_i - \mu_k)^2}{2\sigma_k^2} \right\} \right) &= 0 \\ \sum_{i=1}^N z_{ik} \frac{(x_i - \mu_k)}{\sigma_k^2} &= 0 \\ \hat{\mu}_k &= \frac{\sum_{i=1}^N z_{ik} x_i}{\sum_{i=1}^N z_{ik}} \end{aligned}$$

Proof (continued):

Similarly,

$$\begin{aligned}\frac{\partial}{\partial \sigma_k} L(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) &= 0 \\ \frac{\partial}{\partial \sigma_k} \sum_{i=1}^N \sum_{k=1}^K z_{ik} \log \left(\left(\frac{1}{2\pi\sigma_k^2} \right)^{1/2} \exp\left\{-\frac{(x_i - \mu_k)^2}{2\sigma_k^2}\right\} \right) &= 0 \\ \sum_{i=1}^N z_{ik} \sigma_k^2 &= \sum_{i=1}^N z_{ik} (x_i - \mu_k)^2 \\ \hat{\sigma}_k^2 &= \frac{\sum_{i=1}^N z_{ik} (x_i - \mu_k)^2}{\sum_{i=1}^N z_{ik}}\end{aligned}$$

(e) How are kMeans and Gaussian Mixture Model related? (There are three conditions)

Your answer: K means can be obtained from E-M on GMM via :

- Uniform Mixture Weights i.e $1/K$
- Diagonal Covariances i.e $\sigma_k^2 = cI$
- $c \downarrow 0$