# Appendix: Non-parametric Econometrics

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#### 1. Motivation

- Standard regression models covered so far assumed that the relationship between y and X is linear
- Non-parametric approaches relax the linearity assumption but this advantage does come with costs
- We will cover in this lecture:
  - Densities estimation
  - Non-parametric regression
- Identification of causal effects using non-parametric techniques can be specially useful when treatment is multivalued and the researcher is not willing to impose functional forms

Even when the goal is not to establish a causal relationship, non-parametric approaches can potentially provide a better approximation to the CEF when it is no linear

#### 2. Estimating densities

Assume a researcher wants to estimate the labor income distribution Y. Let's assume that Y is a random variable such that:

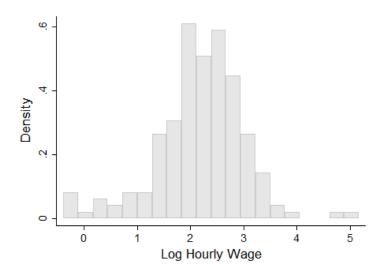
$$F(y) = Pr(Y \le y) \tag{1}$$

$$f(y) = \frac{dF(y)}{dy} \tag{2}$$

• We want to estimate f(y). In particular, we are interested in estimating  $f(y_0)$  for an arbitrary point  $y_0$ 

- What about using a simple histogram?
  - Divide the support of y into a bunch of evenly spaced bins
  - Count how many observations fall into each bin
  - Create a bar graph where the height of each bar is proportional to the number of observations that fall into the corresponding bin
- What is wrong about the histogram?
  - Unlike typical real density functions, it is discontinuous

### Histogram for labor income in US(1993)



- Another option is to follow a parametric approach
  - Fit a theoretical distribution (Example: Pareto Distribution)

$$f(y; \alpha) = \begin{cases} \frac{\alpha y_m^{\alpha}}{y^{\alpha+1}} & \text{if } y \ge y_m \\ 0 & \text{if } y < y_m \end{cases}$$

- f(.) depends on  $\alpha$ , so if an estimation of  $\widehat{\alpha}$  were available, an estimation of  $\widehat{f}(y^*, \widehat{\alpha})$  would be possible
- The problem is parametric: estimating  $\widehat{f}(y^*, \widehat{\alpha})$  is reduced to estimating  $\widehat{\alpha}$
- Problem: It requires to know f(.), but in many situations is exactly what we wanted to know in the first place (income distribution)
- An estimator that is smooth as you move across different values of y would be typically preferred

- How do we get something like this?
  - 1 Take some window, say of radius 1 unit (total 2 units wide), and move it along the *y*-axis
  - 2 As we move it along, count how many observations fall within the window
  - 3 For any given point  $y^*$ , define the estimate of  $f(y^*)$ ,  $\widehat{f}(y^*)$ , as the number of observations that fall within the moving window when it is centered at  $y^*$
  - 4 Although this partially solves the smoothing problem,  $\widehat{f}(y^*)$  will be somewhat discontinuous as observations drop into and out the window
  - 5 This can be solved by augmenting the moving window with a moving weighting function (known as *Kernel*) that places more weight on an observation as it moves closer to the center of the window

#### 2.1 Kernel density estimation

A kernel density estimator, introduced by Rosenblatt(1956), generalizes the histogram estimate by using an alternative weighting function:

$$\widehat{f}(y_0) = \frac{1}{n} \sum_{i=1}^{N} \frac{1}{h} K\left(\frac{Y_i - y_0}{h}\right)$$
 (3)

- Where h is a parameter known as bandwidth and K(.) satisfies the following conditions:
  - 1 K(.) is symmetric around 0 and is continuous
  - **2**  $K(z) \geq 0$

  - $\int sK(s)\mathrm{d}s=0$
  - $\int s^2 K(s) ds = \mu < \infty$

- To illustrate how this estimator works, consider the case triangle kernel:
  - Suppose h = 1 and K(z) = 1(|z| < 1)(1 |z|) (i.e., weight is a linear function of the distance)
  - At a given point  $y_0$ , we have:

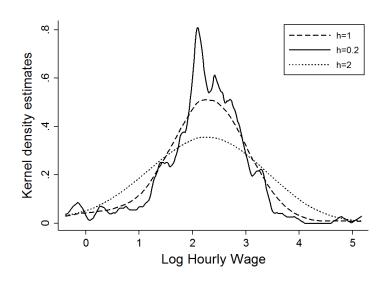
$$\widehat{f}(y_0) = \frac{1}{n} \sum_{i=1}^{N} 1(|Y_i - y_0| < 1)(1 - |Y_i - y_0|) \tag{4}$$

■ Where  $|Y_i - y_0|$  is the Euclidean distance metric

- $\widehat{f}(y_0)$  measures how close is on average each observation in the sample with respect to  $y_0$ 
  - If  $Y_i$  falls more than one unit of distance away from  $y_0$ , it receives a weight of zero and has not input into the sum
  - If Y<sub>i</sub> falls less than one unit of distance away from y<sub>0</sub>, it receives a positive weight which is increasing as Y<sub>i</sub> gets closer to y<sub>0</sub>
  - Computing this summation at all possible values of support of the density leads to  $\widehat{f}(y)$
- What about the role of h and K(.)?

- h plays a critical role in estimating the density:
  - What would happen if we choose a small *h*?
  - What would happen if we choose a large *h*?
- Trade-off between bias and variance!

#### Density estimates using a triangular Kernel with different h



- Less relevant is the role of K(.)
  - Uniform (or rectangular):

$$K(z) = \frac{1}{2}\mathbf{1}(|z| < 1)$$
 (5)

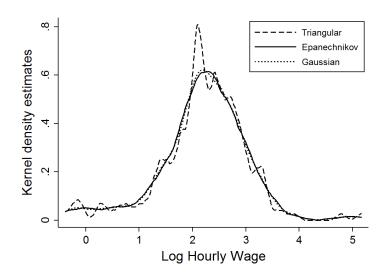
Epanechnikov (or quadratic)

$$K(z) = \frac{3}{4}(1 - z^2).\mathbf{1}(|z| < 1)$$
 (6)

Gaussian (or normal)

$$K(z) = (2\pi)^{\frac{-1}{2}} exp(-z^2/2)$$
 (7)

#### Density estimates using a different Kernel with h = 0.2



#### 3. Non-parametric regression

Assume a researcher wants to estimate the following model:

$$y_i = m(x_i) + u_i \tag{8}$$

• where  $m(x_i)$  is an unknown function and  $\mathbb{E}(u_i|x_i)=0$  such that:

$$\mathbb{E}(y_i|x_i) = m(x_i) \tag{9}$$

- A researcher wants to estimate  $m(x_i)$  using a sample  $(y_i, x_i)$
- Recall the linear solution:

$$m(x_i) = \alpha + \beta x \tag{10}$$



■ How to recover  $\alpha$  and  $\beta$  without knowing the true functional form?

#### 3.1 Kernel regression

- Suppose that for a distinct value of the regressor, say  $x_0$ , there are  $N_0$  observations on  $y_i$ . An estimator for  $m(x_0)$  is the sample average for these  $N_0$  values
- Problem: this estimator might be too noisy in finite samples if  $N_0$  is too small
- A way to solve this issue is looking not only to values of  $y_i$  for  $x_0$ , but also the observed values of  $y_i$  when x is close to  $x_0$

This leads to the local weighted average estimator:

$$\widehat{m}(x_0) = \sum_{i=1}^{N} w(x_i, x_0, h).y_i = \frac{\frac{1}{Nh} \sum_{i=1}^{N} 1\left(\frac{x_i - x_0}{h}\right) y_i}{\frac{1}{Nh} \sum_{i=1}^{N} 1\left(\frac{x_i - x_0}{h}\right)}$$
(11)

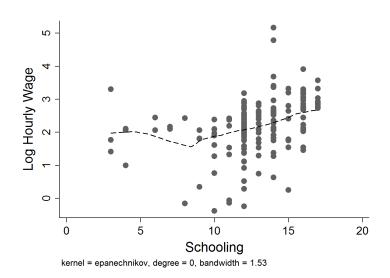
■ Problem: this estimator gives equal weights to all observations close to  $x_0$ 

This leads to the Kernel regression estimator:

$$\widehat{m}(x_0) = \frac{\frac{1}{Nh} \sum_{i=1}^{N} K\left(\frac{x_i - x_0}{h}\right) y_i}{\frac{1}{Nh} \sum_{i=1}^{N} K\left(\frac{x_i - x_0}{h}\right)}$$
(12)

- This estimator is also known as the Nadaraya-Watson
- h play a similar role as in the case of densities

#### Example: Kernel regression



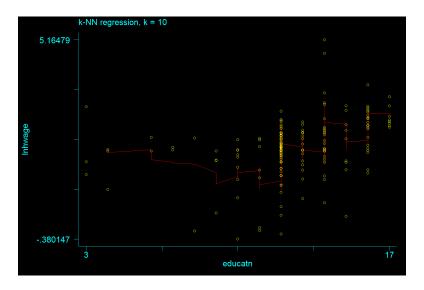
#### 3.2 Nearest neighbors regression

- Consider alternative weights. For instance, the researcher could fix the number of observations around a given point to be included in the analysis
- **k-nearest neighbor estimator** is the equally weighted average of y values for the k observations of  $x_i$  closest to  $x_0$

$$\widehat{m}_{k-NN}(x_0) = \frac{1}{k} \sum_{i=1}^{N} 1(x_i \in N_k(x_0)).y_i$$
 (13)

This estimator is a Kernel estimator with uniform weights

#### Example: Nearest neighbors regression (k=10)



#### 3.3 Local linear and local polynomial regression

- Kernel regression is a local constant estimator (it assumes  $m(x_0)$  equals a constant in the local neighborhood of  $x_0$ )
- We can instead let  $m(x_0)$  be linear in the local neighborhood of  $x_0$ :

$$m(x_0) = a_0 + b_0(x_i - x_0)$$
 (14)

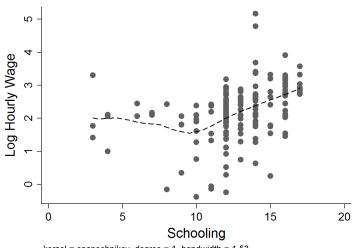
■ The **local linear regression estimator** minimizes:

$$\sum_{i=1}^{N} K\left(\frac{x_i - x_0}{h}\right) (y_i - a_0 - b_0(x_i - x_0))^2$$
 (15)

This can be easily generalized to the case of local polynomial estimator of degree p which minimizes:

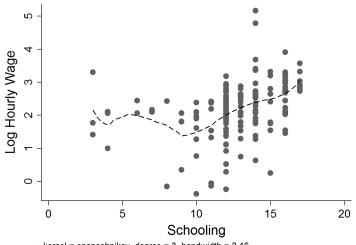
$$\sum_{i=1}^{N} K\left(\frac{x_i - x_0}{h}\right) (y_i - a_0 - a_1(x_i - x_0) - \dots - a_p \frac{(x_i - x_0)^p}{p!})^2$$
 (16)

#### Example: Local linear regression



kernel = epanechnikov, degree = 1, bandwidth = 1.53

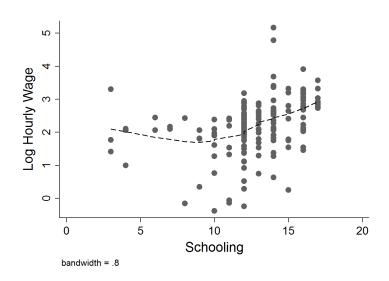
#### Example: Local polynomial regression (degree 3)



## 3.4 Locally weighted scatterplot smoothing (LOWESS) regression

- This is a popular variant of local polynomial estimation in (16):
  - Uses a variable bandwidth determined by the distance from  $x_0$  to the k-th nearest neighbor
  - Uses a tricubic kernel
  - Downweights observations with large residuals

#### Example: LOWESS regression



#### 4. The curse of dimensionality

- Non-parametric techniques typically work well with univariate distributions or with low-dimensional multivariate distributions
- It is hardly the case in most empirical research that scholars work with few covariates
- There is a curse of dimensionality:
  - Sparsity of the data grows exponentially with the number of covariates
  - The more dimensions, the less data to estimate conditional expectations
- There is an important trade-off to evaluate: a researcher can relax the linearity assumption, but it comes with the cost of the curse of dimensionality