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CSCI 2824: Discrete Structures

Lecture 28: Bayes Theorem and the Law of Total Probability

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Probability Theory

Conditional Probability: The probability that E occurs given that F occurred:

$$p(E|F) = \frac{p(E \cap F)}{p(F)}$$

- **Multiplication Rule:** $p(E \cap F) = p(E|F)p(F)$

Independence: Events A and B are independent if

$$p(A|B) = p(A) \quad \bullet$$

$$p(B|A) = p(B) \quad \bullet$$

$$p(A \cap B) = p(A)p(B) \quad \bullet$$

Bayes' Theorem

$$\text{key: } p(E \cap F) = p(F \cap E)$$

From the idea of conditional probability, $p(E|F) = \frac{p(E \cap F)}{p(F)} \cdot \left\{ p(F|E) = \frac{p(F \cap E)}{p(E)} \right.$

$$\Rightarrow p(E \cap F) = p(E|F)p(F) \quad \text{AND} \quad p(F \cap E) = p(F|E)p(E)$$

$$p(E \cap F) = p(F \cap E) \quad \checkmark$$

$$\Rightarrow p(E|F)p(F) = p(F|E)p(E) \quad \checkmark$$

$$\Rightarrow p(F|E) = \boxed{\frac{p(E|F)p(F)}{p(E)}}$$

Bayes' Theorem

This formula is known as **Bayes' Theorem**. $p(F | E) = \frac{p(E | F) p(F)}{p(E)}$

Thomas Bayes



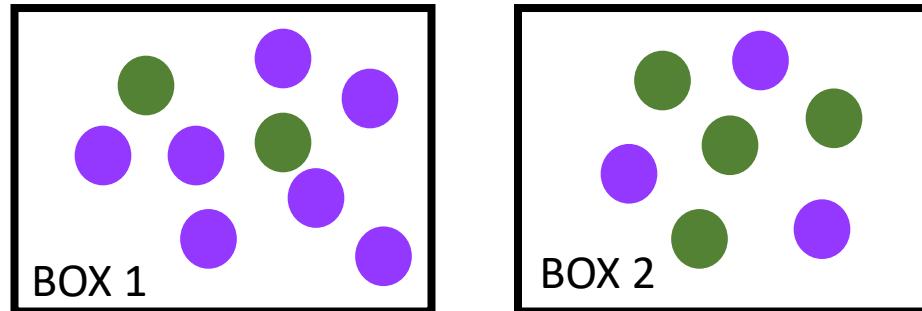
Portrait purportedly of Bayes used in a 1936 book,^[1] but it is doubtful whether the portrait is actually of him.^[2] No earlier portrait or claimed portrait survives.

Bayes' Theorem

Example: Suppose we have two boxes filled with green and purple balls.

Box 1: 2 green balls, 7 purple balls

Box 2: 4 green balls, 3 purple balls



Suppose Anna selects a ball by first choosing one of the two boxes at random. She then selects one of the balls in this box at random. If Anna has selected a purple ball, what is the probability that she selected a ball from the first box?

• $P(\text{ball from } B_1 \mid \text{ball is purple})$

P = event Anna picks a purple ball

B_1 = event Anna picks from Box 1

$\overline{B_1}$ = event Anna picks from Box 2

Bayes' Theorem:

$$p(B_1|P) = \frac{p(P|B_1)p(B_1)}{p(P)}$$

\overline{P} = event that green ball is chosen.

Bayes' Theorem

$$P(P|B_1) = \frac{P(P \cap B_1)}{P(B_1)}$$

We need to calculate $p(P)$.

Let's define $\overline{B_1}$ as the event that Anna selects from Box 2.

Let \bar{P} be the event that Anna has selected a green ball.

- Note that: $P = (\underline{P} \cap B_1) \cup (\underline{P} \cap \overline{B_1})$ and that $((\underline{P} \cap B_1) \cap (\underline{P} \cap \overline{B_1})) = \emptyset$

$$\begin{aligned}|P| &= |(\underline{P} \cap B_1) \cup (\underline{P} \cap \overline{B_1})| \\&= |\underline{P} \cap B_1| + |\underline{P} \cap \overline{B_1}| - |\cancel{(\underline{P} \cap B_1) \cap (\underline{P} \cap \overline{B_1})}| \end{aligned}$$

$$P(P) = P(P \cap B_1) + P(P \cap \overline{B_1})$$

$$P(P) = P(P|B_1) P(B_1) + P(P|\overline{B_1}) P(\overline{B_1})$$

definition of
conditional probability

Law of Total Probability

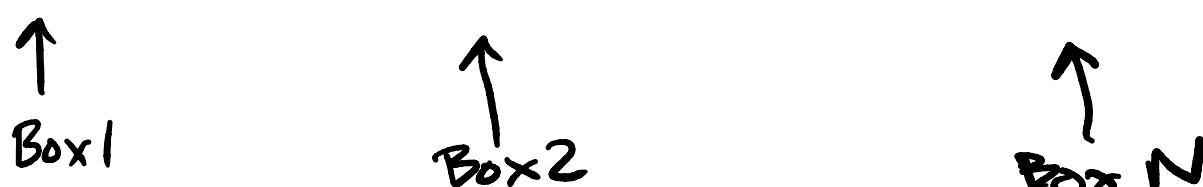
Law of Total Probability:

$$p(P) = p(P|B_1)p(B_1) + p(P|\overline{B_1})p(\overline{B_1})$$

To generalize:

If we can break the set for our event F up into $F = \bigcup_{i=1}^N F_i$, where $F_i \cap F_j = \emptyset$ for $i \neq j$, then the probability of some other event E , $P(E)$, is:

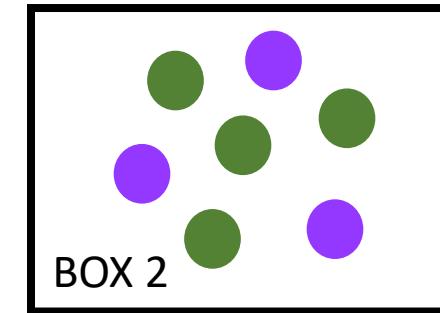
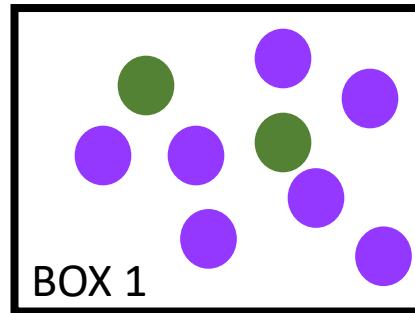
$$p(E) = p(E|F_1)p(F_1) + p(E|F_2)p(F_2) + \dots + p(E|F_N)p(F_N) = \sum_{i=1}^N p(E|F_i)p(F_i)$$



Bayes' Theorem

Back to our example:

$$p(B_1 | P) = \frac{p(P|B_1) p(B_1)}{p(P)}$$



Putting it all together:

$$\begin{aligned} p(B_1 | P) &= \frac{p(P|B_1)p(B_1)}{p(P|B_1)p(B_1) + p(P|\overline{B_1})p(\overline{B_1})} \\ &= \frac{\frac{7}{9} \cdot \frac{1}{2}}{\frac{7}{9} \cdot \frac{1}{2} + \frac{3}{7} \cdot \frac{1}{2}} = \frac{\frac{7}{18}}{\frac{7}{18} + \frac{3}{14}} \quad \boxed{\approx .6447} \end{aligned}$$

Bayes' Theorem

The crux of Bayesian reasoning is the following:

1. Without the observation that Anna picked a purple ball, you would have guessed that the probability of picking from Box 1 was 0.5
2. By assimilating this data, you were able to update this belief about the probability of this event. (to 0.645)

Bayes' Theorem

Example: Cancer Testing. Suppose we know that 1% of the people over the age of 40 have cancer. And assume that 90% of the people who have cancer will test positive for cancer, if tested. Finally suppose that 8% of people who do not have cancer will also test positive (false positives).

What is the probability that a person who tests positive for cancer actually has cancer?

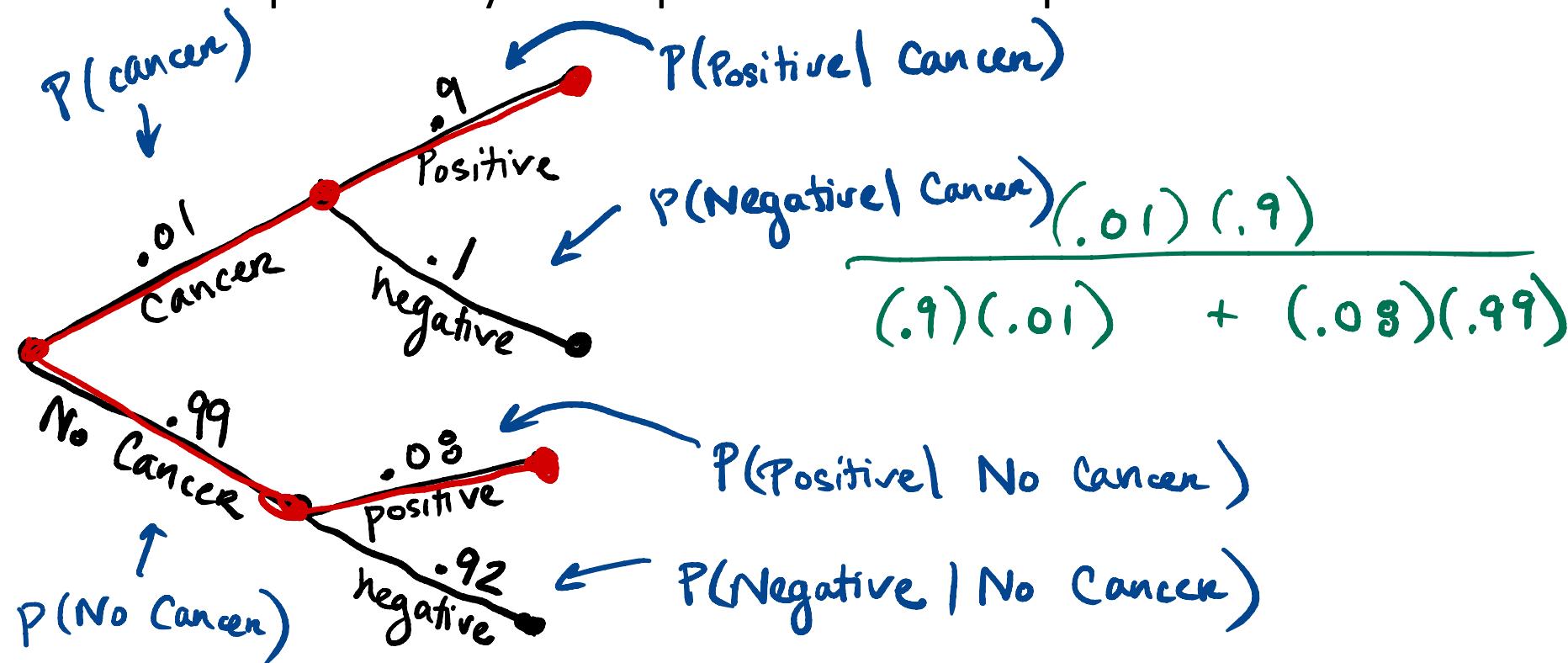
$$\begin{aligned} P(\text{Cancer} \mid \text{Positive}) &= \frac{P(\text{Positive} \mid \text{Cancer}) P(\text{cancer})}{P(\text{positive})} \\ &= \frac{P(\text{Positive} \mid \text{Cancer}) P(\text{cancer})}{P(\text{Positive} \mid \text{cancer}) P(\text{cancer}) + P(\text{Positive} \mid \text{No Cancer}) P(\text{No Cancer})} \\ &= \frac{(.9)(.01)}{(.9)(.01) + (.08)(.99)} \approx .10204082 \end{aligned}$$

expand using the Law of Total Probability.

Bayes' Theorem

Example: Cancer Testing. Suppose we know that 1% of the people over the age of 40 have cancer. And assume that 90% of the people who have cancer will test positive for cancer, if tested. Finally suppose that 8% of people who do not have cancer will also test positive (false positives).

What is the probability that a person who tests positive for cancer actually has cancer?



Bayes' Theorem

From Bayes' Theorem:

$$p(C|\text{positive}) = \frac{p(\text{positive}|C) p(C)}{p(\text{positive})}$$

From the Law of Total Probability:

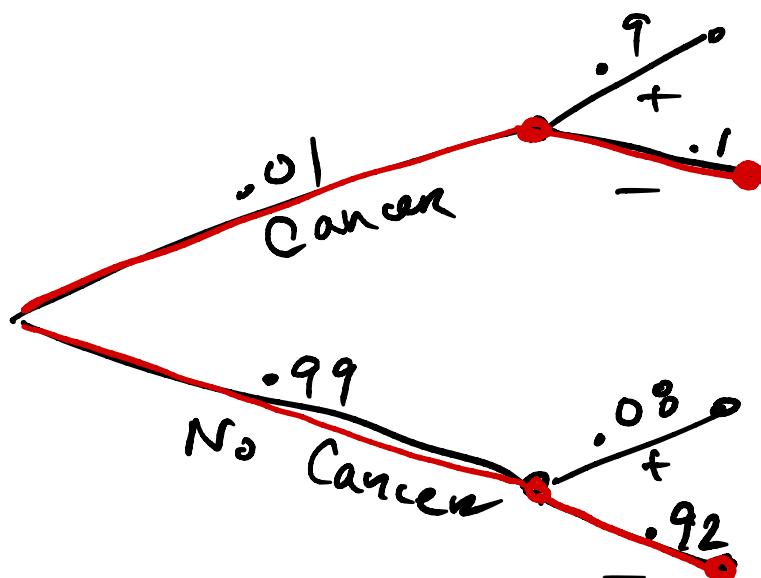
$$p(\text{positive}) = p(\text{positive} | C) p(C) + p(\text{positive} | \bar{C}) p(\bar{C})$$

Bayes' Theorem

Example: Cancer Testing - continued.

What is the probability that a person who tests negative for cancer is actually cancer free?

$$p(\text{no cancer} | \text{negative}) = \frac{p(\text{negative} | \text{no cancer})p(\text{no cancer})}{p(\text{negative} | \text{no cancer})p(\text{no cancer}) + p(\text{negative} | \text{cancer})p(\text{cancer})}$$



$$= \frac{(.92)(.99)}{(.92)(.99) + (.1)(.01)}$$

$$\approx .9989$$

Bayes' Theorem

Generalized Bayes' Theorem:

We can generalize Bayes' theorem to the situation where there are more than two boxes of balls (going back to the ball-drawing example).

Suppose that E is an event from sample space S and F_1, F_2, \dots, F_n are mutually disjoint events such that $S = \cup_{k=1}^n F_k$. Then

$$p(F_i | E) = \frac{p(E | F_i) p(F_i)}{\sum_{k=1}^n p(E | F_k) p(F_k)}$$

For example, the case where there are three bins of balls (B_1, B_2, B_3), and picking a red ball (R):

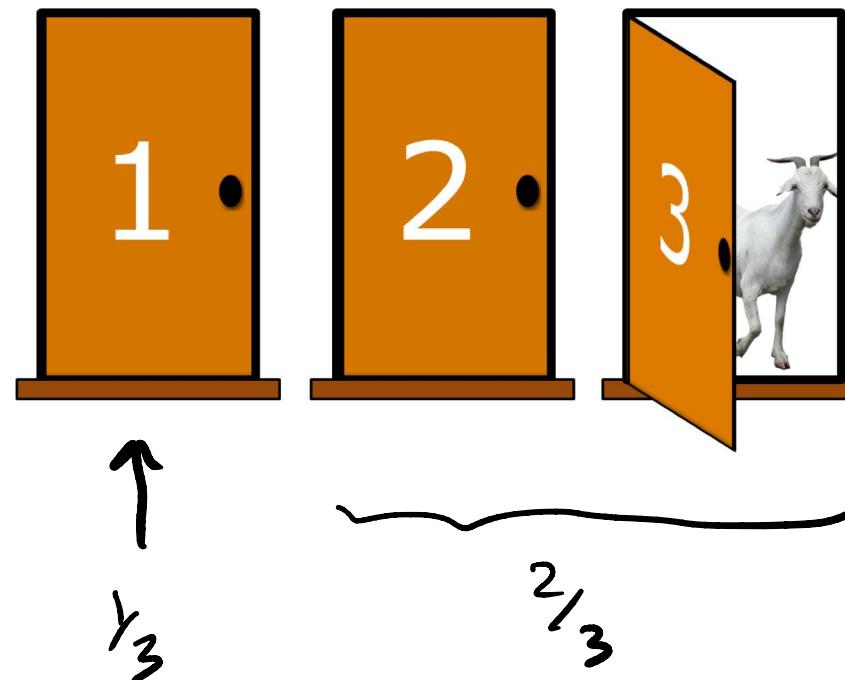
$$p(B_1 | R) = \frac{p(R | B_1) p(B_1)}{\underbrace{p(R | B_1) p(B_1)}_{\text{bin 1?}} + \underbrace{p(R | B_2) p(B_2)}_{\text{bin 2?}} + \underbrace{p(R | B_3) p(B_3)}_{\text{bin 3?}}}$$

Bayes' theorem

Example: Monty Hall Problem

You're on a game show, and you're given the choice of three doors. Behind one of the doors is a car; behind the others are goats. You pick a door - say, Door 1 - and the host, Monty, who knows what's behind all the doors, opens another door - say, Door 3 - which has a goat behind it. He then offers you the choice to switch to Door 2.

Do you switch?



Bayes' theorem

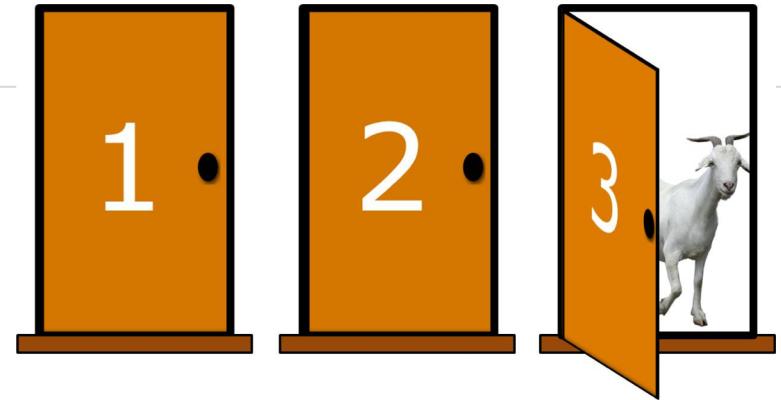
Example: Monty Hall Problem

Do you switch? Answer: Yes! Always!

Why?

It turns out that if you don't switch doors, then your probability of winning is $\frac{1}{3}$

But if you switch doors, then your probability of winning is $\frac{2}{3}$



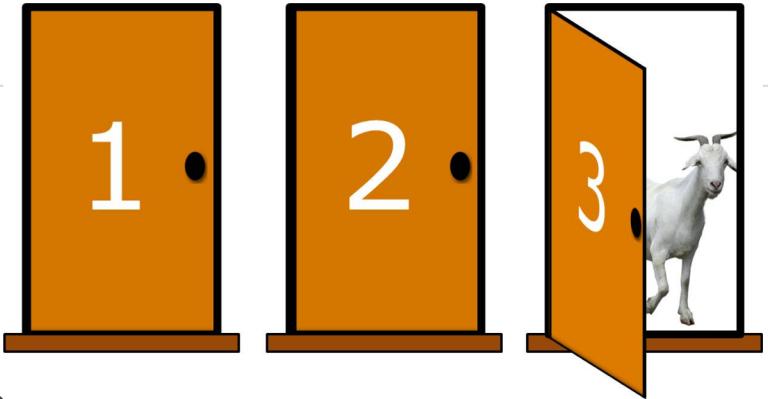
- There was originally $\frac{1}{3}$ probability that your Door 1 was the winner, and $\frac{2}{3}$ probability that Door 2 or Door 3 was the winner
- Monty then tells you that it isn't Door 3.
- This doesn't change the fact that there's $\frac{2}{3}$ total probability between Doors 2 & 3

Bayes' theorem

Example: Monty Hall Problem

Let's work this out using Bayes' Theorem:

Assume you pick Door 1 and Monty shows you there's a goat behind Door 3



Let D_i be the event that the car is behind Door i

Let M_i be the event that Monty reveals a goat behind Door i

⇒ We want to calculate $p(D_1 | M_3)$ and $p(D_2 | M_3)$ and decide which door to go with, 1 or 2

Bayes' Theorem gives:

$$p(D_i | M_3) = \frac{p(M_3 | D_i) p(D_i)}{\sum_{k=1}^3 p(M_3 | D_k) p(D_k)}$$

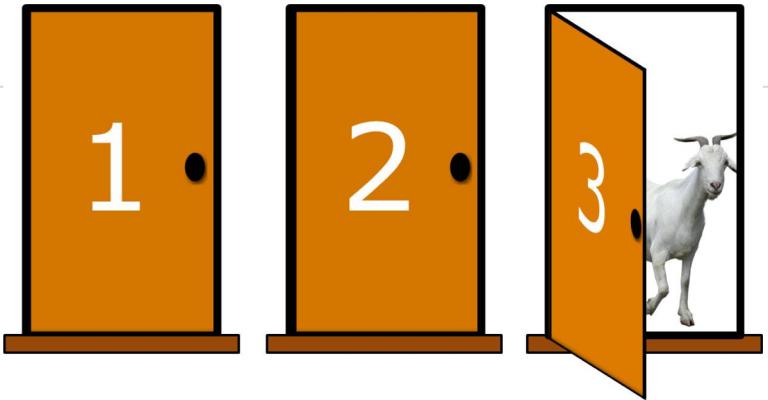
$$p(M_3 | D_1) p(D_1) + p(M_3 | D_2) p(D_2) + p(M_3 | D_3) p(D_3)$$

Bayes' theorem

Example: Monty Hall Problem

Bayes' Theorem:

$$p(D_i \mid M_3) = \frac{p(M_3 \mid D_i) p(D_i)}{\sum_{k=1}^3 p(M_3 \mid D_k) p(D_k)}$$



We can assume that initially, there's an equal probability the car is behind any given door:

$$p(D_1) = p(D_2) = p(D_3) = \frac{1}{3}$$

What should $p(M_3 \mid D_1)$, $p(M_3 \mid D_2)$ and $p(M_3 \mid D_3)$ be?

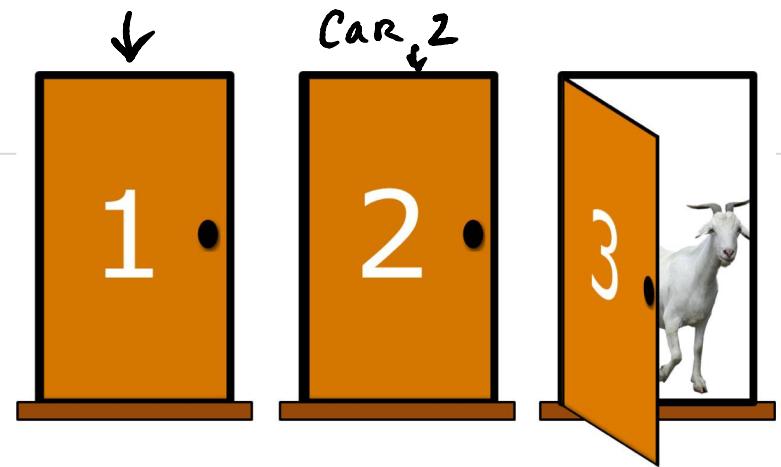
→ $p(M_3 \mid D_3) = 0$, because Monty would never show you which door has the car

Bayes' theorem

Example: Monty Hall Problem

Bayes' Theorem:

$$p(D_i \mid M_3) = \frac{p(M_3 \mid D_i) p(D_i)}{\sum_{k=1}^3 p(M_3 \mid D_k) p(D_k)}$$



We can assume that initially, there's an equal probability the car is behind any given door:

$$p(D_1) = p(D_2) = p(D_3) = \frac{1}{3}$$

What should $p(M_3 \mid D_1)$, $p(M_3 \mid D_2)$ and $p(M_3 \mid D_3)$ be?

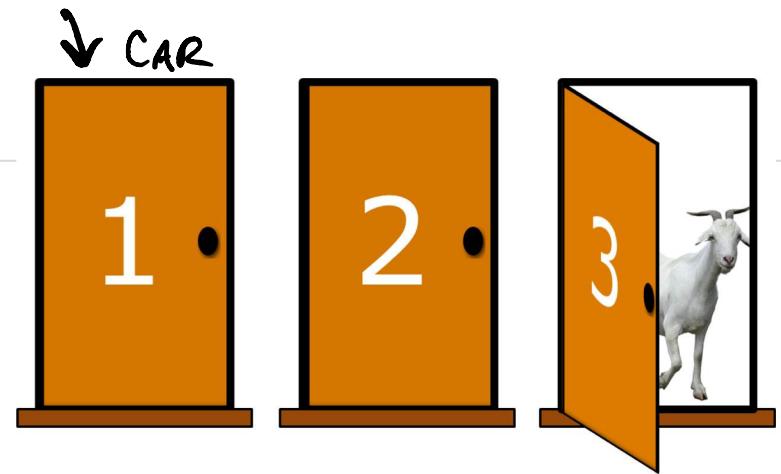
$\Rightarrow \underline{p(M_3 \mid D_2) = 1}$, because you guessed Door 1, so Monty can only show Doors 2 or 3
and the car is behind Door 2, so he can only reveal Door 3

Bayes' theorem

Example: Monty Hall Problem

Bayes' Theorem:

$$p(D_i \mid M_3) = \frac{p(M_3 \mid D_i) p(D_i)}{\sum_{k=1}^3 p(M_3 \mid D_k) p(D_k)}$$



We can assume that initially, there's an equal probability the car is behind any given door:

$$p(D_1) = p(D_2) = p(D_3) = \frac{1}{3}$$

What should $p(M_3 \mid D_1)$, $p(M_3 \mid D_2)$ and $p(M_3 \mid D_3)$ be?

→ $p(M_3 \mid D_1) = \frac{1}{2}$, because you guessed Door 1, so Monty can show either Doors 2 or 3 and will with equal probability

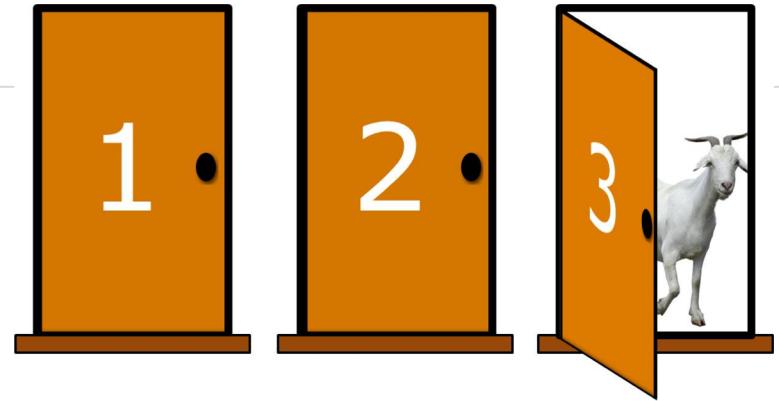
Bayes' theorem

Example: Monty Hall Problem

Bayes' Theorem:

$$p(D_i \mid M_3) = \frac{p(M_3 \mid D_i) p(D_i)}{\sum_{k=1}^3 p(M_3 \mid D_k) p(D_k)}$$

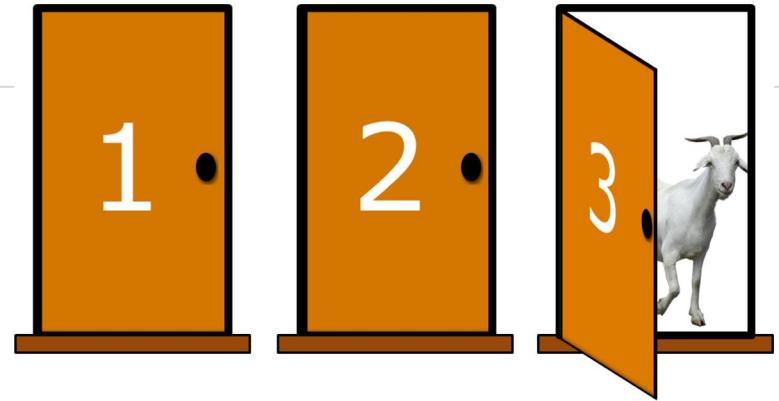
Putting all this together, we find:



Bayes' theorem

Example: Monty Hall Problem

Bayes' Theorem: $p(D_i \mid M_3) = \frac{p(M_3 \mid D_i) p(D_i)}{\sum_{k=1}^3 p(M_3 \mid D_k) p(D_k)}$



Putting all this together, we find:

$$p(D_1 \mid M_3) = \frac{p(M_3 \mid D_1) p(D_1)}{p(M_3 \mid D_1) p(D_1) + p(M_3 \mid D_2) p(D_2) + p(M_3 \mid D_3) p(D_3)}$$

$$= \frac{\frac{1}{2} \cdot \frac{1}{3}}{\frac{1}{2} \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3}} \quad \left. \right\}$$

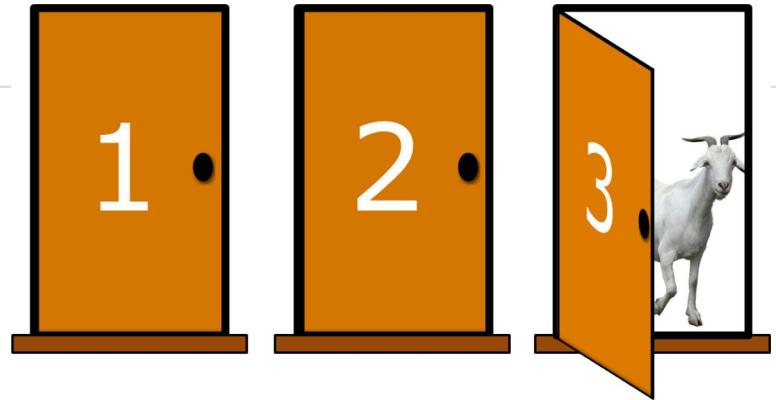
$$\Rightarrow p(D_1 \mid M_3) = \frac{1}{3}$$

Bayes' theorem

Example: Monty Hall Problem

Bayes' Theorem:

$$p(D_i \mid M_3) = \frac{p(M_3 \mid D_i) p(D_i)}{\sum_{k=1}^3 p(M_3 \mid D_k) p(D_k)}$$



Similarly...

$$p(D_2 \mid M_3) = \frac{p(M_3 \mid D_2) p(D_2)}{p(M_3 \mid D_1) p(D_1) + p(M_3 \mid D_2) p(D_2) + p(M_3 \mid D_3) p(D_3)}$$

$$= \frac{1 \cdot \frac{1}{3}}{\frac{1}{2} \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3}}$$

$$\Rightarrow p(D_2 \mid M_3) = \frac{2}{3}$$

Next: Applications of Bayes' Theorem!