

I225E Statistical Signal Processing

14. Signal Processing II

MAWALIM and UNOKI

candylin@jaist.ac.jp and unoki@jaist.ac.jp

School of Information Science

Wiener filter and Kalman filter

Kalman filter is an important generalization of Wiener filter.

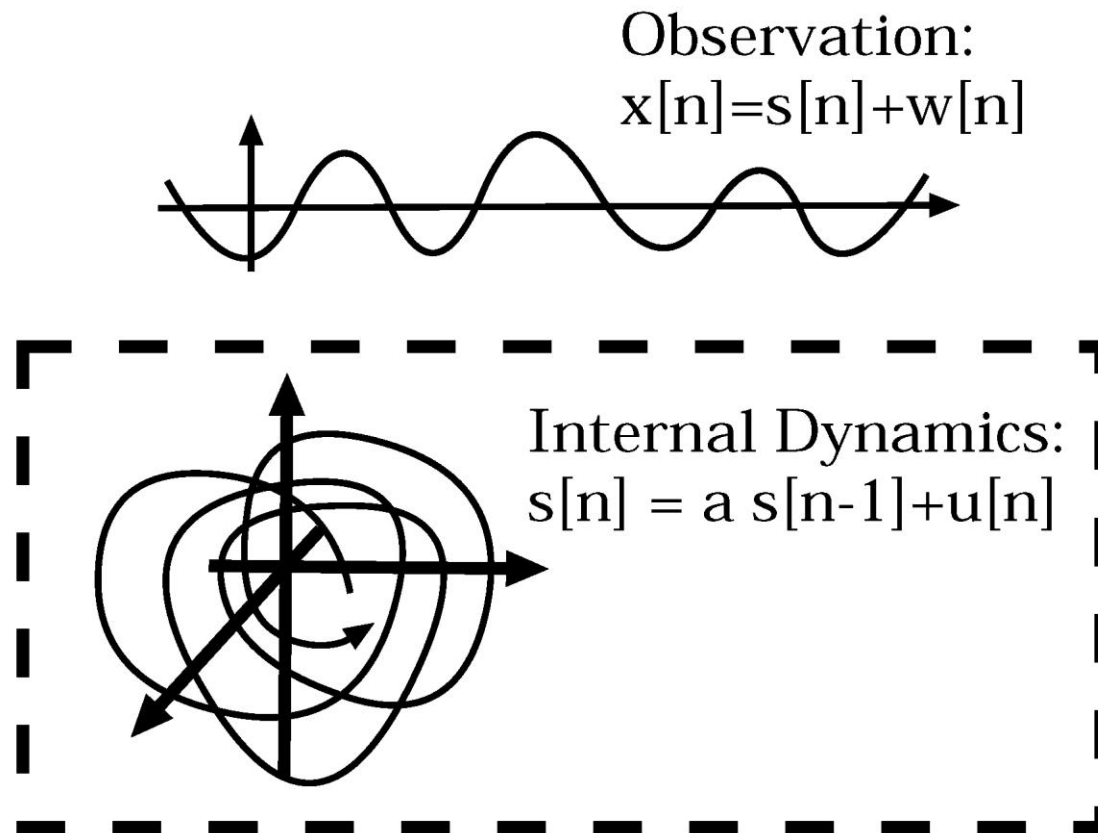
■ *Wiener Filter*

- WSS (Wide-sense-stationary) Process
- Data from infinite past
- Scalar signals
- Non-adaptive

■ *Kalman Filter*

- Gauss-Markov Process
- Data from a specific point in time
- Vector signals
- Adaptive (model may evolve over time)

Estimate internal system state from observed data



2. Scalar Kalman Filter

First-order Gauss-Markov model

$$x[n] = s[n] + w[n],$$

$$s[n] = as[n-1] + u[n].$$

From observed data $\mathbf{X}[n] = [x[0], x[1], \dots, x[n]]^T$, estimate $s[n]$ ($n \geq 0$)

- Constant a is known ($|a| < 1$).
- $u[n] \sim N(0, \sigma_u^2)$, $w[n] \sim N(0, \sigma_w^2)$, $s[-1] \sim N(0, \sigma_s^2)$.
- $s[-1]$, $u[n]$, $w[n]$ are all independent from each other.
- Denote estimate of $s[n]$ based on $\mathbf{X}[m] = [x[0], x[1], \dots, x[m]]^T$ by $\hat{s}[n|m]$.

Find estimator $\hat{s}[n|n]$ that minimizes mean square error

$$E[(s[n] - \hat{s}[n|n])^2].$$

Computational Procedure of Kalman Filter

■ Prediction

$$\hat{s}[n|n-1] = a\hat{s}[n-1|n-1]$$

■ Minimum prediction error

$$M[n|n-1] = a^2 M[n-1|n-1] + \sigma_u^2$$

■ Kalman gain

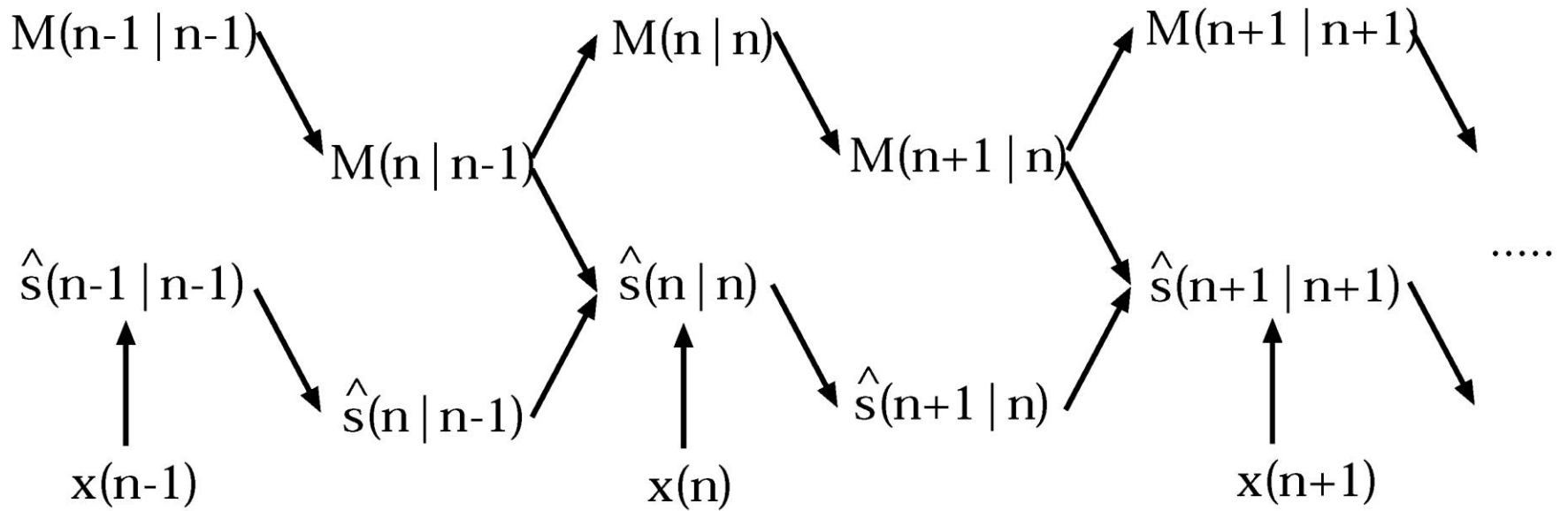
$$K[n] = \frac{M[n|n-1]}{\sigma_w^2 + M[n|n-1]}$$

■ Correction

$$\hat{s}[n|n] = \hat{s}[n|n-1] + K[n](x[n] - \hat{s}[n|n-1])$$

■ Minimum mean square error

$$M[n|n] = (1 - K[n])M[n|n-1]$$



3. Dynamical Model

Theorem: Gauss-Markov model

$$\mathbf{s}[n] = \mathbf{A}\mathbf{s}[n-1] + \mathbf{B}\mathbf{u}[n] \quad n \geq 0.$$

■ $\mathbf{A} \in \mathbb{R}^{p \times p}, \mathbf{B} \in \mathbb{R}^{p \times r}$

Eigenvalues of \mathbf{A} have amplitude smaller than 1.

■ $\mathbf{s}[n] \in \mathbb{R}^{p \times 1}, \mathbf{s}[-1] \sim N(\mu_s, \mathbf{C}_s)$

■ $\mathbf{u}[n] \in \mathbb{R}^{r \times 1}, \mathbf{u}[n] \sim N(0, \mathbf{Q})$

■ $\mathbf{s}[-1], \mathbf{u}[-1]$ are all independent from each other.

Then, mean and covariance of signal $\mathbf{s}[n]$ are

■ **Mean:** $E\{\mathbf{s}[n]\} = \mathbf{A}^{n+1}\boldsymbol{\mu}_s$

■ **Covariance:** $\mathbf{C}_s[m, n] =$

$$E\{(\mathbf{s}[m] - E\{\mathbf{s}[m]\})(\mathbf{s}[n] - E\{\mathbf{s}[n]\})^T\}$$

$(m \geq n) \quad \quad \quad = \mathbf{A}^{m+1}\mathbf{C}_s(\mathbf{A}^{n+1})^T + \sum_{k=m-n}^m \mathbf{A}^k \mathbf{B} \mathbf{Q} \mathbf{B}^T (\mathbf{A}^{n-m+k})^T$

$(m < n) \quad \quad \quad \mathbf{C}_s[m, n] = \mathbf{C}_s^T[n, m]$

$(m = n) \quad \quad \quad \mathbf{C}[n] = \mathbf{C}_s[n, n]$

$$= \mathbf{A}^{n+1}\mathbf{C}_s(\mathbf{A}^{n+1})^T + \sum_{k=0}^n \mathbf{A}^k \mathbf{B} \mathbf{Q} \mathbf{B}^T (\mathbf{A}^k)^T$$

Time evolution of mean and covariance is

$$E\{\mathbf{s}[n]\} = \mathbf{A}E\{\mathbf{s}[n-1]\},$$
$$\mathbf{C}[n] = \mathbf{A}\mathbf{C}[n-1]\mathbf{A}^T + \mathbf{B}\mathbf{Q}\mathbf{B}^T.$$

4. Derivation of scalar Kalman filter

First-order Gauss-Markov model

$$x[n] = s[n] + w[n]$$

$$s[n] = as[n-1] + u[n]$$

From observed data $X[n] = [x[0], x[1], \dots, x[n]]^T$, estimate $s[n] (n \geq 0)$.

- Constant a is known ($|a| < 1$).
- $u[n] \sim N(0, \sigma_u^2)$, $w[n] \sim N(0, \sigma_w^2)$, $s[-1] \sim N(0, \sigma_s^2)$.
- $s[-1], u[n], w[n]$ are all independent from each other.
- Denote estimate of $s[n]$ based on $X[m] = [x[0], x[1], \dots, x[m]]^T$ by $\hat{s}[n|m]$.
- Denote error by $\tilde{x}[n] = x[n] - \hat{x}[n|n-1]$.



With respect to minimum mean square error (MMSE)

$$E\{(s[n] - \hat{s}[n|n])^2\},$$

Corresponding minimum mean square estimation is

$$\hat{s}[n|n] = E\{s[n]|X[n]\}.$$

Basic properties of minimum mean square error estimator used for derivation:

- With respect to uncorrelated data x_1, x_2 , minimum mean square error estimator $\hat{\theta}$ is (in case of $E\{\theta\} = 0$),

$$\hat{\theta} = E\{\theta|x_1, x_2\} = E\{\theta|x_1\} + E\{\theta|x_2\}.$$


- For $\theta = \theta_1 + \theta_2$, minimum mean square error estimator $\hat{\theta}$ is,

$$\begin{aligned}\hat{\theta} &= E\{\theta|x\} \\ &= E\{\theta_1 + \theta_2|x\} \\ &= E\{\theta_1|x\} + E\{\theta_2|x\}.\end{aligned}$$

$$\begin{aligned}\tilde{s}[n|n] &= E\{s[n]|X[n]\} = E\{s[n]|X[n-1], x[n]\} \\ &= E\{s[n]|X[n-1], \tilde{x}[n] + \hat{x}[n]\}\end{aligned}$$

because $\hat{x}[n]$ is represented by linear
summation of $\{x[0], x[1], \dots, x[n-1]\}$

$$\begin{aligned}&= E\{s[n]|X[n-1], \tilde{x}[n]\} \\ &= E\{s[n]|X[n-1]\} + E\{s[n]|\tilde{x}[n]\}.\end{aligned}$$


$$\begin{aligned}\tilde{s}[n|n-1] &= E\{s[n]|\mathbf{X}[n-1]\} \\ &= E\{as[n-1] + u[n]|\mathbf{X}[n-1]\} \\ &= aE\{s[n-1]|\mathbf{X}[n-1]\} \\ &\quad \text{because } E\{u[n]|\mathbf{X}[n-1]\} = E\{u[n]\} = 0 \\ &= a\hat{s}[n-1|n-1]\end{aligned}$$

$$\begin{aligned}\hat{s}[n|n] &= \hat{s}[n|n-1] + E\{s[n]|\tilde{x}[n]\}, \\ \hat{s}[n|n-1] &= a\hat{s}[n-1|n-1]\end{aligned}$$



Since $E\{s[n]|\tilde{x}[n]\}$ is MMSE estimator of $s[n]$ based on $\tilde{x}[n]$,

$$\begin{aligned} E\{s[n]|\tilde{x}[n]\} &= K[n]\tilde{x}[n] \\ &= K[n](x[n] - \hat{x}[n|n-1]) \end{aligned}$$

$$K[n] = \frac{E\{s[n]\tilde{x}[n]\}}{E\{\tilde{x}^2[n]\}}$$

$$\begin{aligned} \hat{x}[n|n-1] &= E\{x[n]|\mathbf{X}[n-1]\} \\ &= E\{s[n] + w[n]|\mathbf{X}[n-1]\} \\ &= E\{s[n]|\mathbf{X}[n-1]\} + E\{w[n]|\mathbf{X}[n-1]\} \\ &= \hat{s}[n|n-1] + \hat{w}[n|n-1] \\ &= \hat{s}[n|n-1]. \quad (\text{from } \hat{w}[n|n-1] = 0) \end{aligned}$$



Summarizing above

$$\begin{aligned}\hat{s}[n|n] &= \hat{s}[n|n-1] + K[n](x[n] - \hat{x}[n|n-1]) \\ &= \hat{s}[n|n-1] + K[n](x[n] - \hat{s}[n|n-1]), \\ \hat{s}[n|n-1] &= a\hat{s}[n-1|n-1].\end{aligned}$$

Denominator and numerator of $K[n] = \frac{E\{s[n]\tilde{x}[n]\}}{E\{\tilde{x}^2[n]\}}$ are

$$E\{s[n]\tilde{x}[n]\} = E\{(s[n] - \hat{s}[n|n-1])\tilde{x}[n]\}$$

from orthogonality principle, $E\{\hat{s}[n|n-1]\tilde{x}[n]\} = 0$

using $\tilde{x}[n] = x[n] - \hat{x}[n|n-1] = x[n] - \hat{s}[n|n-1]$

$$\begin{aligned}
&= E\{(s[n] - \hat{s}[n|n-1])(x[n] - \hat{s}[n|n-1])\} \\
&= E\{(s[n] - \hat{s}[n|n-1])(s[n] + w[n] - \hat{x}[n|n-1])\} \\
&\quad \text{using } E\{(s[n] - \hat{s}[n|n-1])w[n]\} = 0 \\
&= E\{(s[n] - \hat{s}[n|n-1])(s[n] - \hat{s}[n|n-1])\} \\
&= E\{(s[n] - \hat{s}[n|n-1])^2\}.
\end{aligned}$$

$$\begin{aligned}
E\{\tilde{x}^2[n]\} &= E\{(x[n] - \hat{x}[n|n-1])^2\} \\
&= E\{(s[n] - \hat{s}[n|n-1] + w[n])^2\} \\
&= \sigma_w^2 + E\{(s[n] - \hat{s}[n|n-1])^2\}
\end{aligned}$$

Hence

$$K[n] = \frac{E\{(s[n] - \hat{s}[n|n-1])^2\}}{\sigma_w^2 + E\{(s[n] - \hat{s}[n|n-1])^2\}}$$

$$= \frac{M[n|n-1]}{\sigma_w^2 + M[n|n-1]},$$


where

$$M[n|n-1] = E\{(s[n] - \hat{s}[n|n-1])^2\}.$$

$$\begin{aligned} M[n|n-1] &= E\{(s[n] - \hat{s}[n|n-1])^2\} \\ &= E\{(as[n-1] + u[n] - \hat{s}[n|n-1])^2\} \\ &= E\{(a(s[n-1] - \hat{s}[n-1|n-1]) + u[n])^2\} \end{aligned}$$

because $E\{(s[n-1] - \hat{s}[n-1|n-1])u[n]\} = 0$

$$= a^2 M[n-1|n-1] + \sigma_u^2.$$



$$\begin{aligned}
 M[n|n] &= E\{(s[n] - \hat{s}[n|n])^2\} \\
 &= E\{(s[n] - \hat{s}[n|n-1] - K[n](x[n] - \hat{s}[n|n-1]))^2\} \\
 &= E\{(s[n] - \hat{s}[n|n-1])^2\} \\
 &\quad - 2K[n]E\{(s[n] - \hat{s}[n|n-1])(x[n] - \hat{s}[n|n-1])\} \\
 &\quad + K^2[n]E\{(x[n] - \hat{s}[n|n-1])^2\} \\
 &= M[n|n-1] - 2K[n]M[n|n-1] \\
 &\quad + K[n]M[n|n-1] \\
 &= (1 - K[n])M[n|n-1].
 \end{aligned}$$



In summary, scalar Kalman filter is obtained as:

$$\hat{s}[n|n-1] = a\hat{s}[n-1|n-1],$$

$$M[n|n-1] = a^2M[n-1|n-1] + \sigma_u^2,$$

$$K[n] = \frac{M[n|n-1]}{\sigma_w^2 + M[n|n-1]},$$

$$\hat{s}[n|n] = \hat{s}[n|n-1] + K[n](x[n] - \hat{s}[n|n-1]),$$

$$M[n|n] = (1 - K[n])M[n|n-1].$$

4. Extension to vector form

Gauss-Markov model

$$x[n] = \mathbf{h}^T \mathbf{s}[n-1] + w[n]$$

$$\mathbf{s}[n] = \mathbf{A}\mathbf{s}[n-1] + \mathbf{B}\mathbf{u}[n], \quad n \geq 0$$

Estimate $\mathbf{s}[n]$ from observed data $[x[0], x[1], \dots, x[n]]^T$

- $\mathbf{A} \in \mathbb{R}^{p \times p}$, $\mathbf{B} \in \mathbb{R}^{p \times r}$, $\mathbf{h}[n] \in \mathbb{R}^{p \times 1}$

- $\mathbf{s}[n] \in \mathbb{R}^{p \times 1}$, $\mathbf{s}[-1] \sim N(\mu_s, \mathbf{C}_s)$.

- $\mathbf{u}[n] \in \mathbb{R}^{r \times 1}$, $\mathbf{u}[n] \sim N(0, \mathbf{Q})$

- $w[n] \in \mathbb{R}^{1 \times 1}$, $w[n] \sim N(0, \sigma_w^2)$

Vector Kalman filter:

- **Prediction** $\hat{\mathbf{s}}[n|n-1] = \mathbf{A}\hat{\mathbf{s}}[n-1|n-1],$

- **Minimum prediction error**

$$\mathbf{M}[n|n-1] = \mathbf{A}\mathbf{M}[n-1|n-1]\mathbf{A}^T + \mathbf{B}\mathbf{Q}\mathbf{B}^T,$$

- **Kalman gain** $\mathbf{K}[n] = \frac{\mathbf{M}[n|n-1]\mathbf{h}[n]}{\sigma_w^2 + \mathbf{h}^T[n]\mathbf{M}[n|n-1]\mathbf{h}[n]},$

- **Correction**

$$\hat{\mathbf{s}}[n|n] = \hat{\mathbf{s}}[n|n-1] + \mathbf{K}[n](x[n] - \mathbf{h}^T[n]\hat{\mathbf{s}}[n|n-1]),$$

- **Minimum mean square error**

$$\mathbf{M}[n|n] = (\mathbf{I} - \mathbf{K}[n]\mathbf{h}^T[n])\mathbf{M}[n|n-1].$$

Applications

- Satellite control
- Autopilot
- Economy (especially macroeconomics), Time Series Econometrics
- Inertial navigation system
- Car navigation
- Weather forecast

Kalman filter

- Process and measurement equations:

$$\begin{cases} \mathbf{x}_{n+1} = \mathbf{A}\mathbf{x}_n + \mathbf{w}_n \\ \mathbf{z}_n = \mathbf{C}\mathbf{x}_n + \mathbf{v}_n \end{cases}$$

where \mathbf{w} and \mathbf{v} are Gaussian noises with mean $\mathbf{0}$ and covariance \mathbf{Q} and \mathbf{R} .

$$\mathbf{w}_n \sim N(\mathbf{0}, \mathbf{Q}), \quad \mathbf{v}_n \sim N(\mathbf{0}, \mathbf{R})$$

Problem: Given the posterior probability at step n

$$p(\mathbf{x}_n | \mathbf{Z}_{1:n}) = p(\mathbf{x}_n | \mathbf{z}_1, \dots, \mathbf{z}_n),$$

Compute the posterior probability at step $n+1$

$$p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1}) = p(\mathbf{x}_{n+1} | \mathbf{z}_1, \dots, \mathbf{z}_n, \mathbf{z}_{n+1}).$$

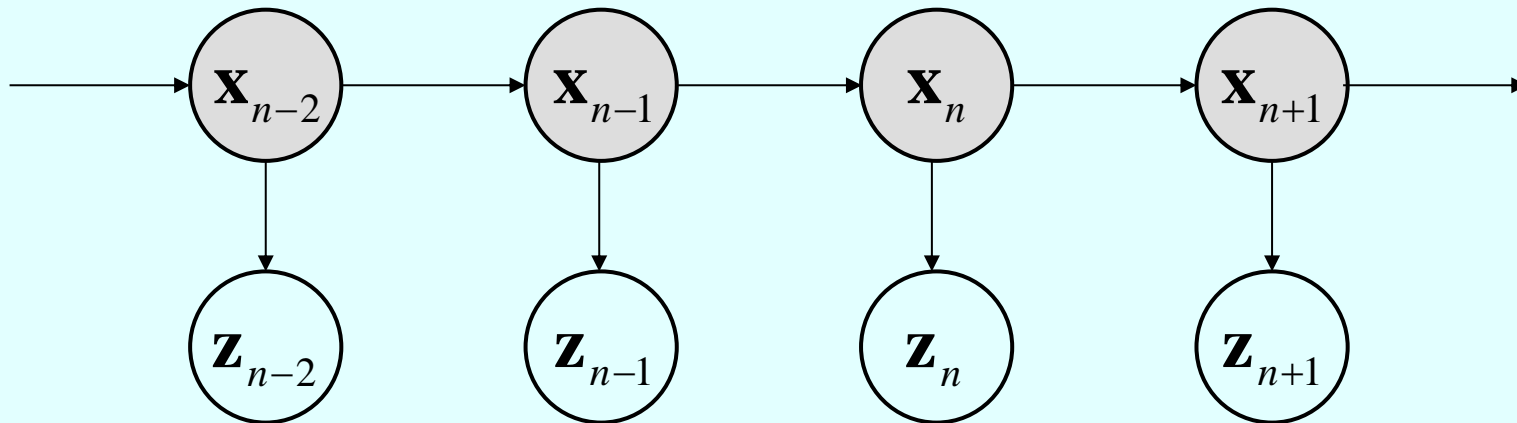
Kalman filter

- Process and measurement equations:

$$\begin{cases} \mathbf{x}_{n+1} = \mathbf{A}\mathbf{x}_n + \mathbf{w}_n \\ \mathbf{z}_n = \mathbf{C}\mathbf{x}_n + \mathbf{v}_n \end{cases}$$

where \mathbf{w} and \mathbf{v} are Gaussian noises with mean $\mathbf{0}$ and covariance \mathbf{Q} and \mathbf{R} .

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Kalman filter

- Problem: Given the posterior probability at step n

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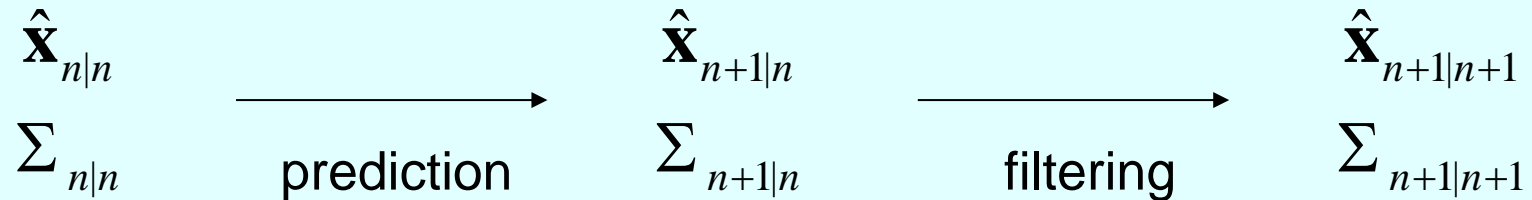
$$p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1}) = p(\mathbf{x}_{n+1} | \mathbf{z}_1, \dots, \mathbf{z}_n, \mathbf{z}_{n+1}).$$

$$p(\mathbf{x}_n | \mathbf{Z}_{1:n}) \xrightarrow{\text{prediction}} p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n}) \xrightarrow{\text{filtering}} p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1})$$

$$\begin{array}{ccccc} \hat{\mathbf{x}}_{n|n} & & \hat{\mathbf{x}}_{n+1|n} & & \hat{\mathbf{x}}_{n+1|n+1} \\ \Sigma_{n|n} & \xrightarrow{\text{prediction}} & \Sigma_{n+1|n} & \xrightarrow{\text{filtering}} & \Sigma_{n+1|n+1} \end{array}$$

Kalman filter

Kalman filter: result



prediction step:

$$\begin{aligned}\hat{\mathbf{x}}_{n+1|n} &= \mathbf{A}\hat{\mathbf{x}}_{n|n} \\ \Sigma_{n+1|n} &= \mathbf{A}\Sigma_{n|n}\mathbf{A}^T + \mathbf{Q}\end{aligned}$$

filtering step:

$$\begin{aligned}\hat{\mathbf{x}}_{n+1|n+1} &= \hat{\mathbf{x}}_{n+1|n} + \mathbf{K}_{n+1} \left(\mathbf{z}_{n+1} - \mathbf{C}\hat{\mathbf{x}}_{n+1|n} \right) \\ \Sigma_{n+1|n+1} &= (\mathbf{I} - \mathbf{K}_{n+1}\mathbf{C})\Sigma_{n+1|n}\end{aligned}$$

$$\mathbf{K}_{n+1} = \Sigma_{n+1|n}\mathbf{C}^T \left(\mathbf{C}\Sigma_{n+1|n}\mathbf{C}^T + \mathbf{R} \right)^{-1}$$

(Simplified) derivation based on mean and variance (1/3)

Prediction step:

$$\hat{\mathbf{x}}_{n+1|n} = \mathbb{E}[\mathbf{x}_{n+1} | \mathbf{Z}_{1:n}] = \mathbb{E}[\mathbf{A}\mathbf{x}_n + \mathbf{w}_n | \mathbf{Z}_{1:n}] = \mathbf{A}\hat{\mathbf{x}}_{n|n}$$

$$\begin{aligned}\Sigma_{n+1|n} &= \text{Cov}[\mathbf{x}_{n+1} | \mathbf{Z}_{1:n}] \\ &= \mathbb{E}\left[\left(\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n}\right)\left(\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n}\right)^T | \mathbf{Z}_{1:n}\right] \\ &= \mathbb{E}\left[\left\{\mathbf{A}\left(\mathbf{x}_n - \hat{\mathbf{x}}_{n|n}\right) + \mathbf{w}_n\right\}\left\{\mathbf{A}\left(\mathbf{x}_n - \hat{\mathbf{x}}_{n|n}\right) + \mathbf{w}_n\right\}^T | \mathbf{Z}_{1:n}\right] \\ &= \mathbf{A}\Sigma_{n|n}\mathbf{A}^T + \mathbf{Q}\end{aligned}$$

$$\boxed{\begin{aligned}\hat{\mathbf{x}}_{n+1|n} &= \mathbf{A}\hat{\mathbf{x}}_{n|n} \\ \Sigma_{n+1|n} &= \mathbf{A}\Sigma_{n|n}\mathbf{A}^T + \mathbf{Q}\end{aligned}}$$

(Simplified) derivation based on mean and variance (2/3)

■ Filtering step:

Assume that the posterior mean has a linear form:

$$\hat{\mathbf{x}}_{n+1|n+1} = \mathbb{E}[\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1}] = \hat{\mathbf{x}}_{n+1|n} + \mathbf{K}_{n+1} (\mathbf{z}_{n+1} - \mathbf{C}\hat{\mathbf{x}}_{n+1|n}).$$

Then, the Kalman gain \mathbf{K} is determined so as to minimize the trace of covariance matrix:

$$\text{Cov}[\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1}].$$

By noting

$$\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n+1} = (\mathbf{I} - \mathbf{K}_{n+1}\mathbf{C})(\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n}) + \mathbf{K}_{n+1}\mathbf{v}_{n+1},$$

we obtain

$$\text{Cov}[\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1}] = (\mathbf{I} - \mathbf{K}_{n+1}\mathbf{C})\Sigma_{n+1|n}(\mathbf{I} - \mathbf{K}_{n+1}\mathbf{C})^T + \mathbf{K}_{n+1}\mathbf{R}\mathbf{K}_{n+1}^T$$

(Simplified) derivation based on mean and variance (3/3)

Finding the Kalman gain:

$$\begin{aligned} 0 &= \frac{\partial}{\partial \mathbf{K}_{n+1}} \text{tr Cov}[\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1}] \\ &= \frac{\partial}{\partial \mathbf{K}_{n+1}} \text{tr} \left\{ (\mathbf{I} - \mathbf{K}_{n+1} \mathbf{C}) \Sigma_{n+1|n} (\mathbf{I} - \mathbf{K}_{n+1} \mathbf{C})^T + \mathbf{K}_{n+1} \mathbf{R} \mathbf{K}_{n+1}^T \right\} \\ &= -2(\mathbf{I} - \mathbf{K}_{n+1} \mathbf{C}) \Sigma_{n+1|n} \mathbf{C}^T + 2\mathbf{K}_{n+1} \mathbf{R} \end{aligned}$$

$$\therefore \mathbf{K}_{n+1} = \Sigma_{n+1|n} \mathbf{C}^T (\mathbf{C} \Sigma_{n+1|n} \mathbf{C}^T + \mathbf{R})^{-1}$$

$$\begin{aligned} \hat{\mathbf{x}}_{n+1|n+1} &= \hat{\mathbf{x}}_{n+1|n} + \mathbf{K}_{n+1} (\mathbf{z}_{n+1} - \mathbf{C} \hat{\mathbf{x}}_{n+1|n}) \\ \Sigma_{n+1|n+1} &= (\mathbf{I} - \mathbf{K}_{n+1} \mathbf{C}) \Sigma_{n+1|n} \end{aligned}$$

(More rigorous) derivation based on probability densities (1/3)

Prediction step:

$$\begin{aligned} p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n}) &= \int d\mathbf{x}_n p(\mathbf{x}_{n+1}, \mathbf{x}_n | \mathbf{Z}_{1:n+1}) \\ &= \int d\mathbf{x}_n p(\mathbf{x}_{n+1} | \mathbf{x}_n) p(\mathbf{x}_n | \mathbf{Z}_{1:n+1}) \\ &= \int d\mathbf{x}_n \frac{1}{(2\pi)^{n/2} |\mathbf{Q}|} \exp\left(-\frac{1}{2}(\mathbf{x}_{n+1} - \mathbf{A}\mathbf{x}_n)^T \mathbf{Q}^{-1} (\mathbf{x}_{n+1} - \mathbf{A}\mathbf{x}_n)\right) \\ &\quad \times \frac{1}{(2\pi)^{n/2} |\Sigma_{n|n}|} \exp\left(-\frac{1}{2}(\mathbf{x}_n - \hat{\mathbf{x}}_{n|n})^T \Sigma_{n|n}^{-1} (\mathbf{x}_n - \hat{\mathbf{x}}_{n|n})\right) \\ &= N(\mathbf{x}_{n+1}; \hat{\mathbf{x}}_{n+1|n}, \Sigma_{n+1|n}) \end{aligned}$$

$$\begin{aligned} \hat{\mathbf{x}}_{n+1|n} &= \mathbf{A}\hat{\mathbf{x}}_{n|n} \\ \Sigma_{n+1|n} &= \mathbf{A}\Sigma_{n|n}\mathbf{A}^T + \mathbf{Q} \end{aligned}$$

(More rigorous) derivation based on probability densities (2/3)

■ Filtering step:

$$\begin{aligned} p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n+1}) &= p(\mathbf{x}_{n+1} | \mathbf{z}_{n+1}, \mathbf{Z}_{1:n}) \\ &= \frac{p(\mathbf{z}_{n+1} | \mathbf{x}_{n+1}) p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n})}{p(\mathbf{z}_{n+1} | \mathbf{Z}_{1:n})} \\ &\propto p(\mathbf{z}_{n+1} | \mathbf{x}_{n+1}) p(\mathbf{x}_{n+1} | \mathbf{Z}_{1:n}) \\ &\propto \exp\left(-\frac{1}{2}(\mathbf{z}_{n+1} - \mathbf{C}\mathbf{x}_{n+1|n})^T \mathbf{R}^{-1}(\mathbf{z}_{n+1} - \mathbf{C}\mathbf{x}_{n+1|n})\right) \exp\left(-\frac{1}{2}(\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n})^T \Sigma_{n+1|n}^{-1}(\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n})\right) \\ &= \exp\left(-\frac{1}{2}(\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n+1})^T \Sigma_{n+1|n+1}^{-1}(\mathbf{x}_{n+1} - \hat{\mathbf{x}}_{n+1|n+1})\right) \end{aligned}$$

$$\begin{aligned} \hat{\mathbf{x}}_{n+1|n+1} &= \left(\Sigma_{n+1|n}^{-1} + \mathbf{C}^T \mathbf{R}^{-1} \mathbf{C}\right)^{-1} \left(\Sigma_{n+1|n}^{-1} \hat{\mathbf{x}}_{n+1|n} + \mathbf{C}^T \mathbf{R}^{-1} \mathbf{z}_{n+1}\right) \\ \Sigma_{n+1|n+1} &= \left(\Sigma_{n+1|n}^{-1} + \mathbf{C}^T \mathbf{R}^{-1} \mathbf{C}\right)^{-1} \end{aligned}$$

(More rigorous) derivation based on probability densities (3/3)

■ Filtering step:

$$\begin{aligned}\Sigma_{n+1|n+1} &= \left(\Sigma_{n+1|n}^{-1} + \mathbf{C}^T \mathbf{R}^{-1} \mathbf{C} \right)^{-1} \\ &= \Sigma_{n+1|n} - \Sigma_{n+1|n} \mathbf{C}^T \left(\mathbf{R} + \mathbf{C} \Sigma_{n+1|n} \mathbf{C}^T \right)^{-1} \mathbf{C} \Sigma_{n+1|n} \\ &= (\mathbf{I} - \mathbf{K}_{n+1} \mathbf{C}) \Sigma_{n+1|n} \\ \hat{\mathbf{x}}_{n+1|n+1} &= \left(\Sigma_{n+1|n}^{-1} + \mathbf{C}^T \mathbf{R}^{-1} \mathbf{C} \right)^{-1} \left(\Sigma_{n+1|n}^{-1} \hat{\mathbf{x}}_{n+1|n} + \mathbf{C}^T \mathbf{R}^{-1} \mathbf{z}_{n+1} \right) \\ &= \hat{\mathbf{x}}_{n+1|n} + \mathbf{K}_{n+1} \left(\mathbf{z}_{n+1} - \mathbf{C} \hat{\mathbf{x}}_{n+1|n} \right)\end{aligned}$$

$$\begin{aligned}\hat{\mathbf{x}}_{n+1|n+1} &= \hat{\mathbf{x}}_{n+1|n} + \mathbf{K}_{n+1} \left(\mathbf{z}_{n+1} - \mathbf{C} \hat{\mathbf{x}}_{n+1|n} \right) \\ \Sigma_{n+1|n+1} &= (\mathbf{I} - \mathbf{K}_{n+1} \mathbf{C}) \Sigma_{n+1|n}\end{aligned}$$