

I225E Statistical Signal Processing

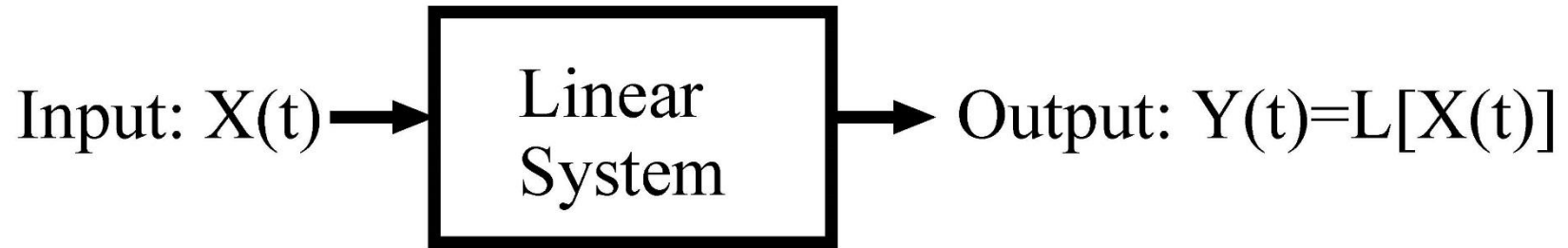
5. Stochastic Process and Systems II

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3. Linear System



$$Y(t) = L[X(t)]$$

■ Properties of linear system

(b1) For any $a_1, a_2, X_1(t), X_2(t)$,

$$L[a_1 X_1(t) + a_2 X_2(t)] = a_1 L[X_1(t)] + a_2 L[X_2(t)].$$

(b2) A system is called **time-invariant**, if its response to $X(t + c)$ equals $Y(t + c)$.

(b3) Output of linear time-invariant system is a convolution

$$Y(t) = X(t) * \mathbf{h}(t) = \int_{-\infty}^{\infty} X(t - \alpha) \mathbf{h}(\alpha) d\alpha$$

Here, $\mathbf{h}(t) = L[\delta(t)]$ represents impulse response.

(b4) Linear time-invariant system has following properties.

- i) If input $X(t)$ is a normal process, then output $Y(t)$ is also a normal process.
- ii) If input $X(t)$ is strict sense stationary, then output $Y(t)$ is also strict sense stationary.

■ Fundamental theorem

For any linear system, $E\{L[X(t)]\} = L[E\{X(t)\}]$

This means that output average $\eta_Y(t)$ equals to output of a system, to which input average $\eta_X(t)$ is input.

$$\eta_Y(t) = L[\eta_X(t)]$$

(Intuitive proof)

Supposing that i th trial is $Y(t, \omega_i) = L[X(t, \omega_i)]$

$$\begin{aligned} E\{Y(t)\} &\approx \frac{Y(t, \omega_1) + Y(t, \omega_2) + \cdots + Y(t, \omega_n)}{n} \\ &= \frac{L[X(t, \omega_1)] + L[X(t, \omega_2)] + \cdots + L[X(t, \omega_n)]}{n} \\ &= L \left[\frac{X(t, \omega_1) + X(t, \omega_2) + \cdots + X(t, \omega_n)}{n} \right] \\ &\approx L[E\{X(t)\}] \end{aligned}$$

(b5) Using the convolution of (b3), the theorem is described as

$$\begin{aligned} E\{Y(t)\} &= \int_{-\infty}^{\infty} E\{X(t - \alpha)\}h(\alpha)d\alpha \\ &= \int_{-\infty}^{\infty} \eta_X(t - \alpha)h(\alpha)d\alpha = \eta_X(t) * h(t) \end{aligned}$$

(b6) Setting $\tilde{X}(t) = X(t) - \eta_X(t)$ and $\tilde{Y}(t) = Y(t) - \eta_Y(t)$,

$$\tilde{Y}(t) = L[\tilde{X}(t)]$$

[Proof]

$$\begin{aligned} \tilde{Y}(t) &= Y(t) - \eta_Y(t) \\ &= L[X(t)] - L[\eta_X(t)] \\ &= L[X(t) - \eta_X(t)] \\ &= L[\tilde{X}(t)] \end{aligned}$$

(b7) With respect to input $X(t) = f(t) + v(t)$ ($f(t)$ is a deterministic signal; $E\{v(t)\} = 0$), the output average is

$$\begin{aligned}\eta_Y(t) &= E\{Y(t)\} = E\{L[X(t)]\} = L[E\{X(t)\}] \\ &= L[E\{f(t) + v(t)\}] \\ &= L[E\{f(t)\} + E\{v(t)\}] \\ &= L[f(t)] = f(t) * h(t)\end{aligned}$$

■ Autocorrelation of output

Consider input $X(t)$ and output $Y(t)$ for a linear system. We represent output autocorrelation $R_{YY}(t_1, t_2)$ in terms of input autocorrelation $R_{XX}(t_1, t_2)$.

■ Theorem

$$(A) R_{XY}(t_1, t_2) = L_{t_2}[R_{XX}(t_1, t_2)].$$

$$(B) R_{YY}(t_1, t_2) = L_{t_1}[R_{XY}(t_1, t_2)].$$

[Proof]

Multiplying both side of $Y(t) = L_t[X(t)]$ by $X(t_1)$,

$$X(t_1)Y(t) = X(t_1)L_t[X(t)] = L_t[X(t_1)X(t)].$$

Taking expectation,

$$\begin{aligned} E\{X(t_1)Y(t)\} &= E\{L_t[X(t_1)X(t)]\} \\ &= L_t[E\{X(t_1)X(t)\}]. \end{aligned}$$

Substitution of $t = t_2$ yields (A).

In a similar manner, multiplication of $Y(t) = L_t[X(t)]$ by $Y(t_2)$ and so on yields (B).

(c1) In the form of convolution,

$$\begin{aligned} \text{(a)} \quad R_{XY}(t_1, t_2) &= R_{XX}(t_1, t_2) * h(t_2) \\ &= \int_{-\infty}^{\infty} R_{XX}(t_1, t_2 - \alpha) h(\alpha) d\alpha \end{aligned}$$

$$\begin{aligned} \text{(b)} \quad R_{YY}(t_1, t_2) &= R_{XY}(t_1, t_2) * h(t_1) \\ &= \int_{-\infty}^{\infty} R_{XY}(t_1 - \alpha, t_2) h(\alpha) d\alpha \end{aligned}$$

(c) Combining (a) and (b),

$$\begin{aligned} R_{YY}(t_1, t_2) &= \\ &\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} R_{XX}(t_1 - \alpha, t_2 - \beta) h(\alpha) h(\beta) d\alpha d\beta \end{aligned}$$

(c2) For cross-covariance, the following holds.

$$C_{XY}(t_1, t_2) = C_{XX}(t_1, t_2) * h(t_2)$$

$$C_{YY}(t_1, t_2) = C_{XY}(t_1, t_2) * h(t_1)$$

[Proof]


Denoting $\tilde{X}(t) = X(t) - \eta_X(t)$, $\tilde{Y}(t) = Y(t) - \eta_Y(t)$,

$$C_{XX}(t_1, t_2) = E\{\tilde{X}(t_1)\tilde{X}(t_2)\} = R_{\tilde{X}\tilde{X}}(t_1, t_2)$$

$$C_{XY}(t_1, t_2) = R_{\tilde{X}\tilde{Y}}(t_1, t_2),$$

$$C_{YY}(t_1, t_2) = R_{\tilde{Y}\tilde{Y}}(t_1, t_2).$$

From (b6), $\tilde{Y}(t) = L[\tilde{X}(t)]$. Hence, (a) and (b) can be applied to $R_{\tilde{X}\tilde{X}}$, $R_{\tilde{X}\tilde{Y}}$, and $R_{\tilde{Y}\tilde{Y}}$.



(c3) Above results can be extended to the case the impulse response $\mathbf{h}(t)$ is complex function.

In this case,

$$R_{XY}(t_1, t_2) = R_{XX}(t_1, t_2) * \mathbf{h}^*(t_2)$$

$$R_{YY}(t_1, t_2) = R_{XY}(t_1, t_2) * \mathbf{h}(t_1)$$

Response to white noise

Theorem

Consider white noise, whose autocorrelation is given by $R_{XX}(t_1, t_2) = q(t_1)\delta(t_1 - t_2)$, as input to linear system. The output power is given by

$$E\{Y^2(t)\} = q(t) * h^2(t) = \int_{-\infty}^{\infty} q(t - \alpha)h^2(\alpha)d\alpha$$

[Proof]

Substitution of $R_{XX}(t_1, t_2) = q(t_1)\delta(t_1 - t_2)$ into convolution form of (A), (B) yields,

$$\begin{aligned} R_{XY}(t_1, t_2) &= q(t_1)\delta(t_1 - t_2) * h(t_2) = q(t_1)h(t_2 - t_1) \\ R_{YY}(t_1, t_2) &= \int_{-\infty}^{\infty} q(t_1 - \alpha)h[t_2 - (t_1 - \alpha)]h(\alpha)d\alpha. \end{aligned}$$

Substitution of $t_1 = t_2 = t$ gives the output power.

(d1) If input $X(t)$ is stationary white noise, $q(t) = q$.

Therefore,

$$E\{Y^2(t)\} = qE, \quad E = \int_{-\infty}^{\infty} h^2(\alpha) d\alpha$$

(d2) If correlation time of impulse response $h(t)$ is shorter than that of $q(t)$,

$$E\{Y^2(t)\} \approx q(t) \int_{-\infty}^{\infty} h^2(\alpha) d\alpha = Eq(t).$$

(d3) If $R_{XX}(\tau) = q\delta(\tau)$ and $X(t)$ is injected at $t = 0$.

$$q(t) = qU(t)$$

Therefore,

$$E\{Y^2(t)\} \approx q \int_{-\infty}^t h^2(\alpha) d\alpha$$

Digression on Dirac's delta function

- Properties of delta function $\delta(t)$:

$$\int_{-\infty}^{+\infty} dt \delta(t) = 1.$$

$$\int_{-\infty}^{+\infty} dt \delta(t) f(t) = f(0).$$

- Delta function as a limit of infinitesimally narrow Gaussian:

$$\delta(t) = \lim_{\sigma \rightarrow 0} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{t^2}{2\sigma^2}}$$

- Delta function has a uniform (white) spectrum:

$$\delta(t) = \frac{1}{2\pi} \int_{-\infty}^{+\infty} d\omega e^{i\omega t}$$

Example 1

Consider a linear system having impulse response

$$h(t) = e^{-ct}U(t)$$

Input a white noise, whose autocorrelation is given by $R_{WW}(\tau) = q\delta(\tau)$, at $t = 0$. Calculate autocorrelation $R_{YY}(t_1, t_2)$ ($0 < t_1 < t_2$) of the output $Y(t)$.

Answer Denoting input as $X(t)$, $X(t) = W(t)U(t)$.

(I) Case of $t_1 < 0$ or $t_2 < 0$: Since there is no input,

$$R_{XX}(t_1, t_2) = R_{XY}(t_1, t_2) = R_{YY}(t_1, t_2) = 0$$

(II) Case of $0 < t_1 < t_2$:

$$R_{XX}(t_1, t_2) = R_{WW}(t_1 - t_2) = q\delta(t_1 - t_2)$$

Therefore,

$$\begin{aligned} R_{XY}(t_1, t_2) &= \int_{-\infty}^{t_2} R_{XX}(t_1, t_2 - \alpha) h(\alpha) d\alpha \\ &= \int_{-\infty}^{t_2} q \delta(t_1 - t_2 + \alpha) e^{-c\alpha} U(\alpha) d\alpha \\ &= q e^{-c(t_2 - t_1)} \end{aligned}$$

$$\begin{aligned} R_{YY}(t_1, t_2) &= \int_{-\infty}^{t_1} R_{XY}(t_1 - \alpha, t_2) h(\alpha) d\alpha \\ &= \int_{-\infty}^{t_1} q e^{-c(t_2 - t_1 + \alpha)} e^{-c\alpha} U(\alpha) d\alpha \\ &= q e^{-c(t_2 - t_1)} \int_0^{t_1} e^{-2c\alpha} d\alpha \\ &= \frac{q}{2c} e^{-c(t_2 - t_1)} (1 - e^{-2ct_1}) \end{aligned}$$

$$E\{Y^2(t)\} = R_{YY}(t, t) = \frac{q}{2c} (1 - e^{-2ct}).$$

Example 2: Differentiator

Consider the output properties (mean and autocorrelation) of differentiator:

$$Y(t) = L[X(t)] = X'(t).$$

■ **Mean** is,

$$\eta_Y(t) = L[\eta_X(t)] = \eta'_X(t)$$

■ **Autocorrelation** is

$$R_{XY}(t_1, t_2) = L_{t_2}[R_{XX}(t_1, t_2)] = \frac{\partial R_{XX}(t_1, t_2)}{\partial t_2}$$

$$R_{YY}(t_1, t_2) = L_{t_1}[R_{XY}(t_1, t_2)] = \frac{\partial R_{XY}(t_1, t_2)}{\partial t_1}$$

Hence

$$R_{YY}(t_1, t_2) = L_{t_1}[R_{XY}(t_1, t_2)] = \frac{\partial^2 R_{XX}(t_1, t_2)}{\partial t_1 \partial t_2}$$

If $X(t)$ is wide-sense stationary, its **mean**, $\eta_X(t)$, is a constant. Therefore, $\eta_Y(t) = 0$. Since **autocorrelation** is a function of $\tau = t_1 - t_2$,

$$\frac{\partial R_{XY}(t_1, t_2)}{\partial t_2} = - \frac{dR_{XX}(\tau)}{d\tau}$$
$$\frac{\partial^2 R_{XX}(t_1, t_2)}{\partial t_1 \partial t_2} = - \frac{d^2 R_{XX}(\tau)}{d\tau^2}$$

Hence,

$$R_{XY}(\tau) = -R'_{XX}(\tau),$$
$$R_{YY}(\tau) = -R''_{XX}(\tau),$$

Example 3: Differential Equation

Consider output properties (mean and autocorrelation) of the following differential equation driven by noise $X(t)$.

$$a_n Y^{(n)}(t) + a_{n-1} Y^{(n-1)}(t) + \cdots + a_0 Y(t) = X(t)$$

The initial condition is given by

$$Y(0) = Y^{(1)}(0) = \cdots = Y^{(n)}(0) = 0$$

■ Mean:

Taking the expectation $E\{\cdot\}$ for both sides,

$$\begin{aligned} a_n E\{Y^{(n)}(t)\} + a_{n-1} E\{Y^{(n-1)}(t)\} + \cdots + a_0 E\{Y(t)\} \\ = E\{X(t)\} \end{aligned}$$

Since $\frac{d}{dx}$ and $E\{\cdot\}$ are commutative, $E\{Y^{(k)}(t)\} = \eta_Y^{(k)}(t)$

$$a_n \eta_Y^{(n)}(t) + a_{n-1} \eta_Y^{(n-1)}(t) + \cdots + a_0 \eta_Y(t) = \eta_X(t)$$

Because of $Y^{(k)}(0) = 0$, $\eta_Y^{(k)}(0) = 0$. The mean $\eta_Y^{(k)}(t)$ is obtained by integrating the differential equation for the following initial condition:

$$\eta_Y(0) = \eta_Y^{(1)}(0) = \cdots = \eta_Y^{(n)}(0) = 0$$

■ Correlation:

By substituting $t = t_2$ and multiplying by $X(t_1)$, the differential equation becomes

$$\begin{aligned} X(t_1) [a_n Y^{(n)}(t_2) + a_{n-1} Y^{(n-1)}(t_2) + \cdots + a_0 Y(t_2)] \\ = X(t_1) X(t_2) \end{aligned}$$

By taking the expectation and using the formula of

$$E\{X(t_1)Y^{(k)}(t_2)\} = \frac{\partial^k R_{XY}(t_1, t_2)}{\partial t_2^k},$$

$$a_n \frac{\partial^n R_{XY}(t_1, t_2)}{\partial t_2^n} + a_{n-1} \frac{\partial^{n-1} R_{XY}(t_1, t_2)}{\partial t_2^{n-1}} + \dots \\ + a_0 R_{XY}(t_1, t_2) = R_{XX}(t_1, t_2),$$

Since $X(t_1)Y^{(k)}(0) = 0$, initial condition is given by

$$R_{XY}(t_1, 0) = \frac{\partial R_{XY}(t_1, 0)}{\partial t_2} = \dots = \frac{\partial^n R_{XY}(t_1, 0)}{\partial t_2^n} = 0$$

Autocorrelations $\frac{\partial^n R_{XY}(t_1, t_2)}{\partial t_2^n}$ are obtained by solving the differential equation for the above initial condition.

In a similar manner, by substituting $t = t_1$ and multiplying by $Y(t_2)$,

$$\begin{aligned} & [a_n Y^{(n)}(t_1) + a_{n-1} Y^{(n-1)}(t_1) + \cdots + a_0 Y(t_1)] Y(t_2) \\ & = X(t_1) Y(t_2) \end{aligned}$$

By taking the expectation and by using the formula

$$\begin{aligned} E\{Y^{(k)}(t_1) Y(t_2)\} &= \frac{\partial^k R_{YY}(t_1, t_2)}{\partial t_1^k}, \\ a_n \frac{\partial^n R_{YY}(t_1, t_2)}{\partial t_1^n} &+ a_{n-1} \frac{\partial^{n-1} R_{YY}(t_1, t_2)}{\partial t_1^{n-1}} + \cdots \\ &+ a_0 R_{YY}(t_1, t_2) = R_{XY}(t_1, t_2). \end{aligned}$$

Since $Y(0)Y^{(k)}(t_2) = 0$, initial condition is given by

$$R_{YY}(0, t_2) = \frac{\partial R_{YY}(0, t_2)}{\partial t_1} = \dots = \frac{\partial^n R_{YY}(0, t_2)}{\partial t_1^n} = 0$$

Autocorrelations $\frac{\partial^n R_{YY}(t_1, t_2)}{\partial t_1^n}$ are obtained by solving the differential equation for the above initial condition.