I225E Statistical Signal Processing

4. Stochastic Process and Systems I

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Agenda

- Review
- Practice Exercise Discussion
- Introduction to Stochastic System
- Memoryless System

REVIEW OF PROBABILITY THEORY & BASICS OF STOCHASTIC PROCESS

Stochastic Process

Definition: A stochastic process is a family of random variables,

$$\{X(t): t \in T\},\$$

where t usually denotes time.

Discrete-time Process

Definition: if the set *T* is finite or countable.

In practice, this generally means $T = \{0, 1, 2, 3, ...\}$

Continuous-time Process

Definition: if *T* is not finite or countable.

In practice, this generally means $T = [0, \infty)$, or T = [0, K] for some K.

State Space

Definition: The state space, S, is the set of real values that X(t) can take.

Every X(t) takes a value in \mathbb{R} , but S will often be a smaller set: $S \subseteq \mathbb{R}$.

For example, if X(t) is the outcome of a coin tossed at time t, then the state space is $S = \{0, 1\}$.

Discrete and Continuous

The state space *S* is *discrete* if it is finite or countable. Otherwise, it is *continuous*.

Independent Random Variables

Consider two discrete random variables X and Y. We say that X and Y are independent if

$$P(X = x, Y = y) = P(X = x)P(Y = y)$$
, for all x, y.

Remember, the concept of independent events A and B.

$$P(A,B) = P(A \cap B) = P(A)P(B)$$

Expected value (=mean=average)

$$EX = E[X] = E(X) = \mu X$$

- Theorem:
 - E[aX + b] = aE[X] + b, for all $a, b \in \mathbb{R}$
 - $E[X_1 + X_2 + \dots + X_n] = E[X_1] + \dots + E[X_n]$, for any set of random variables X_1, X_2, \dots, X_n .
- Expected value of a function of a random variable:

$$E[g(X)] = \sum_{x_k \in R_X} g(x_k) P_X(x_k)$$

Variance

$$Var(X) = E[(X - \mu X)2].$$

- $Var(X) = E[X^2] [EX]^2$
- Standard Deviation

$$SD(X) = \sigma X = \sqrt{Var(X)}.$$

- Theorem:
 - For a random variable X and real numbers a and b, $Var(aX + b) = a^2Var(X)$
 - If $X_1, X_2, ..., X_n$ are independent random variables and $X = X_1 + X_2 + \cdots + X_n$, then

$$Var(X) = Var(X_1) + Var(X_2) + \dots + Var(X_n)$$

Special Distribution (Discrete)

- **■** Bernouli distribution
- Notation: $X \sim Bernouli(p)$.
- **Description:** A random variable X is a Bernoulli random variable with parameter p, where 0 .
- Probability function:

$$f_X(x) = \begin{cases} p & \text{for } x = 1\\ 1 - p & \text{for } x = 0\\ 0 & \text{otherwise} \end{cases}$$

- Mean: E(X) = p.
- Variance: Var(X) = pq, where q = 1 p.

Special Distribution (Discrete)

- Binomial distribution
- Notation: $X \sim Binomial(n, p)$.
- **Description:** number of successes in *n* independent trials, each with probability *p* of success.
- Probability function:

$$f_X(x) = P(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$$
, for $x = 0, 1, ..., n$.

- \blacksquare Mean: E(X) = np.
- Variance: Var(X) = np(1-p) = npq, where q = 1-p.
- Sum: If $X \sim Binomial(n, p), Y \sim Binomial(m, p)$, and X and Y are independent, then

$$X + Y \sim Binomial(n + m, p).$$

Special Distribution (Discrete)

- Poisson distribution
- Notation: $X \sim Poisson(\lambda)$.
- Probability function:

$$f_X(x) = P(X = x) = \frac{\lambda^x}{x!} e^{-\lambda}$$
, for $x = 0, 1, 2, ...$

- Mean: $E(X) = \lambda$.
- Variance: $Var(X) = \lambda$.
- Sum: If $X \sim Poisson(\lambda)$, $Y \sim Poisson(\mu)$, and X and Y are independent, then

$$X + Y \sim Poisson(\lambda + \mu)$$
.

- Uniform distribution
- **Notation:** $X \sim Uniform(a, b)$.
- Probability density function (PDF):

$$f_X(x) = \frac{1}{b-a}$$
, for $a < x < b$

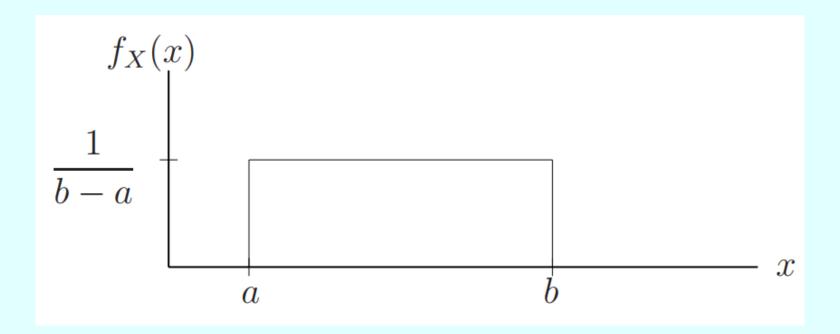
Cumulative distribution function:

$$F_X(x) = P(X \le x) = \frac{x-a}{b-a}, \text{ for } a < x < b$$

$$F_X(x) = 0, \text{ for } x \le a, \text{ and } F_X(x) = 1, x \ge b$$

- Mean: $E(x) = \frac{a+b}{2}$
- Variance: $Var(X) = \frac{(b-a)^2}{12}$

■ Uniform distribution (a, b)



- Exponential distribution
- **Notation:** $X \sim Exponential(\lambda)$.
- Probability density function (PDF):

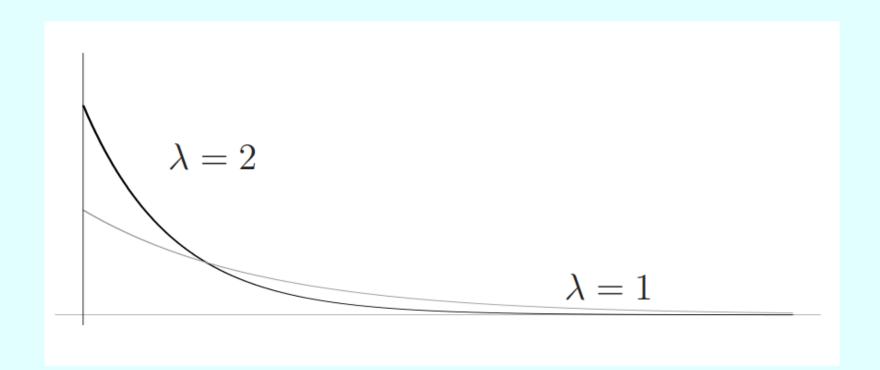
$$f_X(x) = \lambda e^{-\lambda x}$$
, for $0 < x < \infty$

Cumulative distribution function:

$$F_X(x) = P(X \le x) = 1 - e^{-\lambda x}$$
, for $0 < x < \infty$
 $F_X(x) = 0$, for $x \le 0$

- Mean: $E(x) = \frac{1}{\lambda}$
- Variance: $Var(X) = \frac{1}{\lambda^2}$

Exponential distribution (λ)



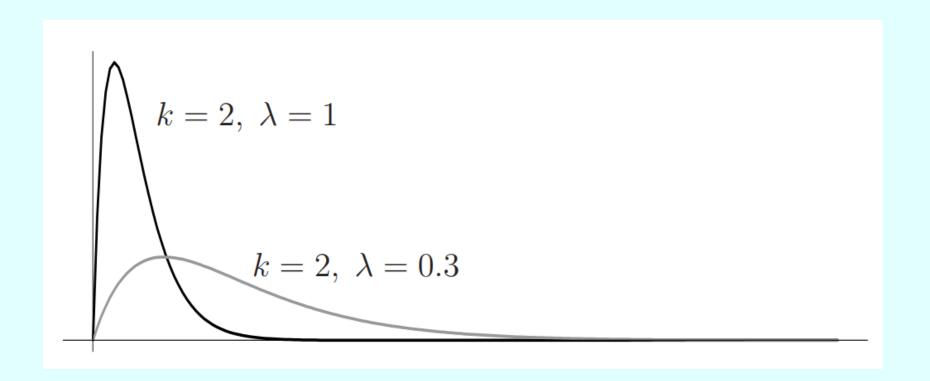
- Gamma distribution
- **Notation:** $X \sim Gamma(k, \lambda)$.
- Probability density function (PDF):

$$f_X(x) = \frac{\lambda^k}{\Gamma(k)} x^{k-1} e^{-\lambda x}$$
, for $0 < x < \infty$

where $\Gamma(k) = \int_0^\infty y^{k-1} e^{-y} dy$ (the Gamma function).

- **Cumulative distribution function:** no closed form.
- Mean: $E(x) = \frac{k}{\lambda}$
- Variance: $Var(X) = \frac{k}{\lambda^2}$

■ Gamma distribution (k, λ)

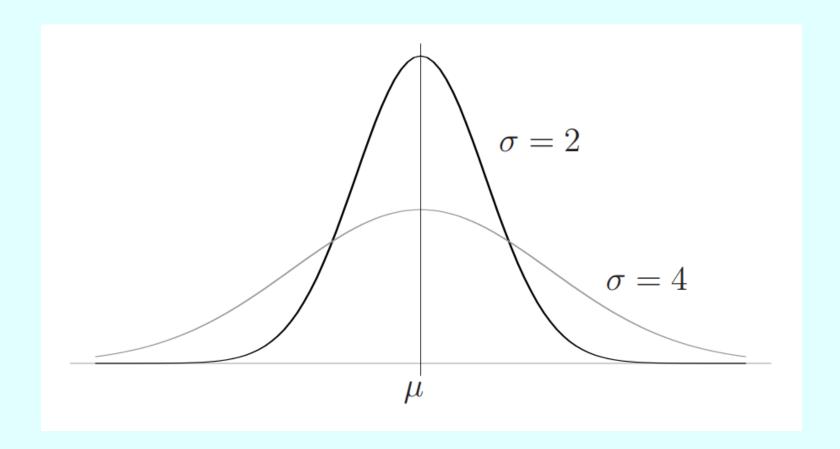


- Normal distribution
- Notation: $X \sim Normal(\mu, \sigma^2)$.
- Probability density function (PDF):

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}}, \text{ for } -\infty < x < \infty$$
where $\Gamma(k) = \int_0^\infty y^{k-1} e^{-y} dy$ (the Gamma function).

- **Cumulative distribution function:** no closed form.
- Mean: $E(x) = \mu$
- **Variance:** $Var(X) = \sigma^2$

Normal distribution (μ, σ^2)



Autocorrelation Function (ACF)

Autocorrelation of X(t)

$$\rho_{X}(t_{1}, t_{2}) = Corr(X_{t_{1}}, X_{t_{2}}) = R_{XX}(t_{1}, t_{2}) = E\{X(t_{1})X(t_{2})\} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_{1}x_{2}f_{X}(x_{1}, x_{2}; t_{1}, t_{2})dx_{1}dx_{2}$$

where sample autocorrelation is

$$\bar{R}_{XX}(t_1, t_2) = \frac{1}{n} \sum_{t=1}^{n} X(t + t_1) X(t + t_2)$$

Covariance Function

The autocovariance function $\gamma_X(t_1, t_2)$ of a stochastic process $\{X_t\}$ measures the covariance between the process at two different time points t_1 and t_2 :

$$\gamma_X(t_1, t_2) = Cov(X_{t_1}, X_{t_2}) = C_{XX}(t_1, t_2)$$

$$= E\left[\left(X_{t_1} - \mu_X(t_1) \right) \left(X_{t_2} - \mu_X(t_2) \right) \right]$$

$$= R_{XX}(t_1, t_2) - \eta_X(t_1) \eta_X(t_2)$$

- For a <u>weakly stationary</u> process, the autocovariance function depends only on the time difference (lag) $k = t_2 t_1$.
- In the case of $t_1=t_2=t$, $\mathcal{C}_{XX}(t_1,t_2)$ is equal to variance of

$$X(t) \to C_{XX}(t,t) = E\{X(t)X(t)\} - \eta_X^2(t) = Var(X(t))$$

Complex process

X(t) = Y(t) + jZ(t): complex variable X(t) is composed of real part Y(t) and imaginary part Z(t).

$$Corr(X_{t_1}, X_{t_2}) = R_{XX}(t_1, t_2) = E\{X(t_1)X^*(t_2)\}$$

$$Corr(X_{t_1}, X_{t_2}) = R_{XX}(t_2, t_1) = R_{XX}^*(t_1, t_2)$$

$$Corr(X_t, X_t) = R_{XX}(t, t) = E\{|X(t)|^2\} \ge 0$$

$$Cov(X_{t_1}, X_{t_2}) = C_{XX}(t_1, t_2) = R_{XX}(t_1, t_2) - \eta_X(t_1)\eta_X^*(t_2)$$

Correlation coefficient

$$r(t_1, t_2) = \frac{C_{XX}(t_1, t_2)}{\sqrt{C_{XX}(t_1, t_1)C_{XX}(t_2, t_2)}}$$

Previous Exercise

[Exercise]

A complex stochastic process Z(t) is given by:

$$Z(t) = X(t) + iY(t)$$

Where:

X(t) and Y(t) are real-valued random processes.

i is the imaginary unit.

$$E[X(t)] = 1$$
, $E[Y(t)] = 2$, for all t.

$$Var[X(t)] = 4$$
, $Var[Y(t)] = 9$, for all t.

$$Cov(X(t), Y(t)) = 0$$
, for all t .

The autocorrelation function of X(t) is $R_{XX}(t_1, t_2) = 4 \cdot \exp(-\frac{|t_1 - t_2|}{2})$.

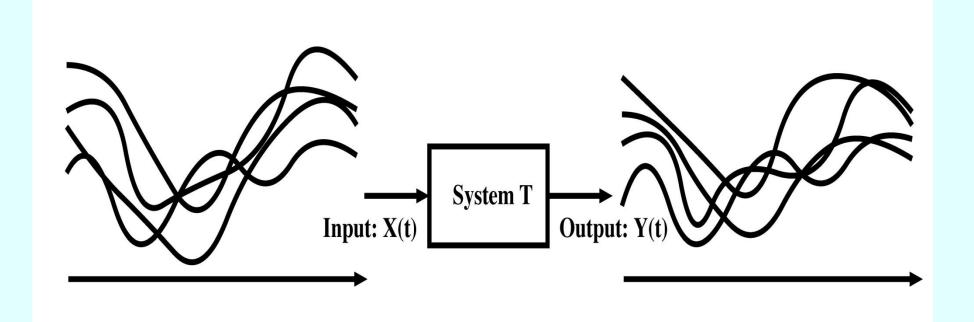
The autocorrelation function of Y(t) is $R_{YY}(t_1, t_2) = 9 \cdot \cos(\pi * (t1 - t2))$.

Previous Exercise

[Exercise]

- a) Calculate the mean function of the complex process Z(t).
- b) Calculate the autocovariance function of the complex process Z(t).
- c) Calculate the autocorrelation function of the complex process Z(t).
- d) Calculate the correlation coefficient between $Z(t_1)$ and $Z(t_2)$.

1. System with stochastic input



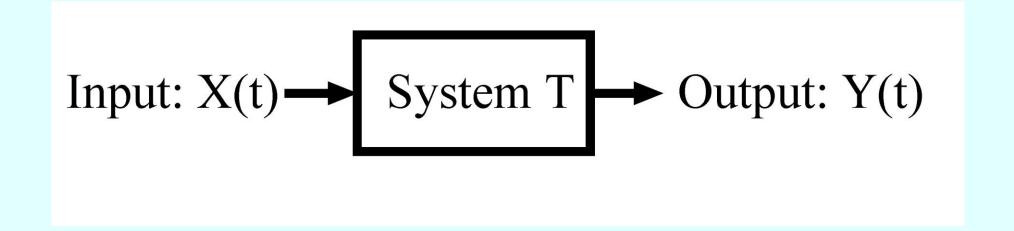
System

Given a stochastic process X(t) as input, Y(t) represents its output.

$$Y(t) = T[X(t)]$$

Purpose:

If statistical properties of input X(t) are known, study the statistical properties of output Y(t)



$$Y(t) = T[X(t)]$$

System Dynamics

Deterministic system:

System operates only on variable t, treating outcome ω as a parameter. Namely,

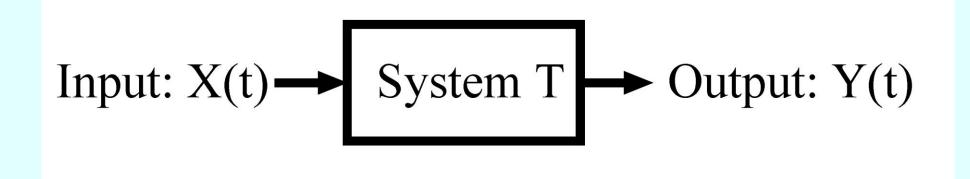
If
$$X(t, \omega_1) = X(t, \omega_2)$$
, then $Y(t, \omega_1) = Y(t, \omega_2)$,

Stochastic system:

System operates on both t and ω . Namely,

Even if
$$X(t, \omega_1) = X(t, \omega_2), Y(t, \omega_1) \neq Y(t, \omega_2)$$
.

Example: Physical element of the system or coefficient of the system equation is stochastic.



$$Y(t) = T[X(t)]$$

- This lecture deals with only deterministic systems.
- In deterministic systems, transformation T may depend on t. To emphasize this, sometimes denoted as

$$\boldsymbol{Y}(t) = T_t[\boldsymbol{X}(t)]$$

referred to as a time-dependent system.

Deterministic System

Memoryless System

System with Memory

$$\mathbf{Y}(t) = g[\mathbf{X}(t)]$$

Time-Varying System

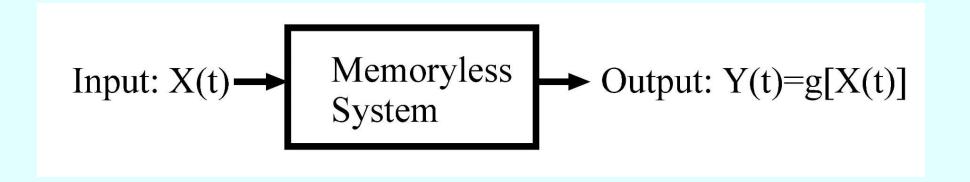
Time-Invariant System

Linear System Y(t) = L[X(t)]

Linear Time-Invariant system

$$\mathbf{Y}(t) = \mathbf{X}(t) * \mathbf{h}(t)$$

2. Memoryless System



Output $Y(t_1)$ at time $t = t_1$ depends only upon the simultaneous state of input $X(t_1)$, but not upon past or future state of X(t)

$$Y(t) = g[X(t)]$$

Mean of the Output

The mean (or expected value) of the output process Y(t), denoted by E[Y(t)] or $\mu_Y(t)$, can be found using the law of the unconscious statistician:

(a1)
$$\mu_Y(t) = E\{Y(t)\} = E[g(X(t))] =$$

$$\int_{-\infty}^{\infty} g(x) f_X(x, t) dx$$

where $f_X(x,t)$ is the first-order probability density function (PDF) of the input process X(t) at time t.

Correlation of the Output

- The autocorrelation function of the output process Y(t), denoted by $R_Y(t_1,t_2)=E[Y(t_1)Y(t_2)]$, measures the statistical dependence between the output at two different times t_1 and t_2 .
- For a memoryless system, $Y(t_1)$ depends only on $X(t_1)$, and $Y(t_2)$ depends only on $X(t_2)$. Therefore, the autocorrelation of the output is:

(a2)
$$E\{Y(t_1)Y(t_2)\} = E\{g(X(t_1))g(X(t_2))\}$$

= $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x_1)g(x_2)f_X(x_1, x_2; t_1, t_2)dx_1dx_2$

(a3) nth-order density of Y(t), $f_Y(y_1, y_2, \cdots, y_n; t_1, t_2, \cdots, t_n)$ is obtained from nth-order density of X(t), $f_X(x_1, x_2, \cdots, x_n; t_1, t_2, \cdots, t_n)$ through transformation $Y(t_1) = g[X(t_1)], Y(t_2) = g[X(t_2)], \cdots, Y(t_n) = g[X(t_n)].$ If the following system $y_1 = g[x_1], y_2 = g[x_2], \cdots, y_n = g[x_n]$ has a unique solution $\mathbf{x} = [x_1, x_2, \cdots, x_n], n$ th-order density of

Y(t) is obtained as $f_Y(y_1,y_2,\cdots,y_n;t_1,t_2,\cdots,t_n) = \frac{f_X(x_1,x_2,\cdots,x_n;t_1,t_2,\cdots,t_n)}{|I_X(x_1,x_2,\cdots,x_n)|}$

where \boldsymbol{J} is Jacobian $\boldsymbol{J} = [g'(x_1)g'(x_2)\cdots g'(x_n)].$

When more than two solutions exist, summation of the corresponding terms $\frac{f_X}{|J|}$ gives the *n*th-order density.

Digression on coordinate transformation

Let us consider *n*-dim variables $\mathbf{x} = (x_1, \dots, x_n)^T$ and $\mathbf{y} = (y_1, \dots, y_n)^T$ and the mapping $y_i = y_i(x_1, \dots, x_n) = y_i(\mathbf{x})$. In this case, their infinitesimal volume are related as

$$dy_1 \cdots dy_n = |J(x_1, \cdots, x_n)| dx_1 \cdots dx_n$$

where the matrix is called the Jacobian:

$$J = \begin{pmatrix} \frac{\partial y_1}{\partial x_1} & \dots & \frac{\partial y_1}{\partial x_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial y_n}{\partial x_1} & \dots & \frac{\partial y_n}{\partial x_n} \end{pmatrix}$$

Accordingly, their probability densities are related as

$$f_{\mathbf{y}}(y_1, \dots, y_n) = \frac{1}{|J(x_1, \dots, x_n)|} f_{\mathbf{x}}(x_1, \dots, x_n)$$

Appendix

Following has been used for the derivation of the density in (a3).

With respect to random variables $X = [X_1, X_2, \dots, X_n]$, n functions

$$Y_1 = g_1(X), Y_2 = g_2(X), \dots, Y_n = g_n(X),$$

are given. For n random numbers $Y = [Y_1, Y_2, \cdots, Y_n]$, we determine their joint density $f_Y(y_1, y_2, \cdots, y_n)$, where y_1, y_2, \cdots, y_n represent a specific set of numbers.

To find the density, we solve the system

$$g_1(X) = y_1, g_2(X) = y_2, \dots, g_n(X) = y_n.$$

If the system has no solution, then $f_Y(y_1, y_2, \dots, y_n) = 0$. If the system has a single solution $\mathbf{x} = [x_1, x_2, \dots, x_n]$, the density can be obtained by substituting the solution into following formula

$$f_Y(y_1, y_2, \dots, y_n) = \frac{f_X(x_1, x_2, \dots, x_n)}{|J(x_1, x_2, \dots, x_n)|},$$

$$J(x_1, x_2, \dots, x_n) = \begin{vmatrix} \frac{\partial g_1}{\partial x_1} & \dots & \frac{\partial g_1}{\partial x_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial g_n}{\partial x_1} & \dots & \frac{\partial g_n}{\partial x_n} \end{vmatrix}$$

If more than two solutions exist, the density is given by summation of all the corresponding terms

$$f_Y = \frac{f_X}{|J|}\Big|_{X=X_1} + \frac{f_X}{|J|}\Big|_{X=X_2} + \cdots$$

[Exercise]

Imagine you have a transformation that takes a point (u, v) in the uv-plane and maps it to a point (x, y) in the xy-plane. Let's say this transformation is defined by the following equations:

$$x(u,v) = u^2 v$$
$$y(u,v) = u + v^2$$

Calculate the Jacobian matrix of the output variables (x and y) with respect to the input variables (u and v).

[Answer]

$$J(u,v) = \begin{pmatrix} 2uv & u^2 \\ 1 & 2v \end{pmatrix}$$

(a4) If input X(t) is strict sense stationary, output Y(t) is also strict sense stationary.

[Proof] According to (a3), nth-order density of Y(t) is given as

$$f_{Y}(y_{1}, y_{2}, \cdots, y_{n}; t_{1}, t_{2}, \cdots, t_{n}) = \frac{f_{X}(x_{1}, x_{2}, \cdots, x_{n}; t_{1}, t_{2}, \cdots, t_{n})}{|\boldsymbol{J}(x_{1}, x_{2}, \cdots, x_{n})|}$$

Since X(t) is strict sense stationary, its density is invariant to a shift of the origin in time, denominator is independent of t. Therefore, f_Y is also invariant to time-shift. This proves that Y(t) is strict sense stationary.

- From the properties of strict sense stationary
- (i) First-order density of Y(t) is independent of t $\rightarrow f_Y(y;t) = f_Y(y)$
- (ii) Second-order density is a function of time lag $\tau = t_1 t_2$ $\rightarrow f_V(y_1, y_2; t_1, t_2) = f_V(y_1, y_2; \tau)$

Example of memoryless system

Square-law detector

Square-law detector is a memoryless system whose output equals

$$Y(t) = X^2(t).$$

Using the density $f_X(x,t)$ of input X(t), find the density $f_Y(y,t)$ of output Y(t).

First-order density

If y > 0, solutions of $y = x^2$ are $x = \pm \sqrt{y}$.

The corresponding Jacobian matrices are $J = \frac{dx^2}{dx} = 2x = \pm 2\sqrt{y}$.

Hence

$$f_Y(y;t) = \frac{f_X}{|J|} \Big|_{x=\sqrt{y}} + \frac{f_X}{|J|} \Big|_{x=-\sqrt{y}}$$
$$= \frac{1}{2\sqrt{y}} \Big[f_X(\sqrt{y};t) + f_X(-\sqrt{y};t) \Big]$$

Second-order density:

If $y_1 > 0$, $y_2 > 0$, solutions of $y_1 = x_1^2$, $y_2 = x_2^2$ are $(\pm \sqrt{y_1}, \pm \sqrt{y_2})$. Since the corresponding Jacobian matrices are $J = \pm 4\sqrt{y_1y_2}$,

$$f_Y(y_1, y_2; t_1, t_2) = \frac{1}{4\sqrt{y_1y_2}} \sum f_X(\pm \sqrt{y_1}, \pm \sqrt{y_2}; t_1, t_2)$$