analisys_modena_cancer

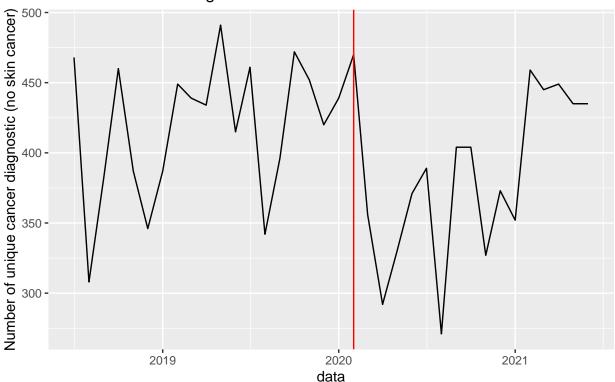
Rmarkdwons settings

Loading the dataset of month incidence

Exploratory data analysis:

Don't know how to automatically pick scale for object of type ts. Defaulting to continuous.

Diagnostic cancer by month, from July 2018 to June 2021, in Modena diagnostic structure

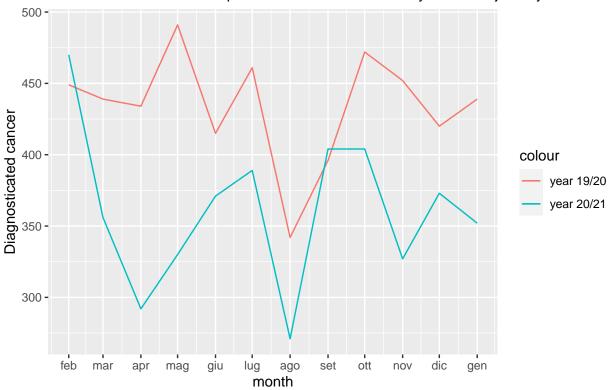


Dataframe of 12 months before and 12 anfter the covid pandemic

date_labels = "%b") +
theme(plot.title = element_text(size = 11))

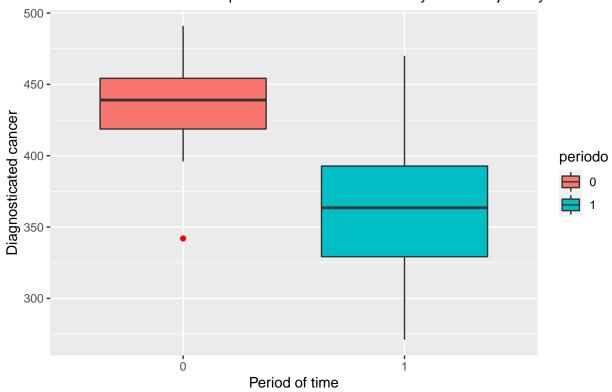
```
# 12 months dataser
# horizontal datafrmae
pre = dati all$conteggio[8:19]
post = dati all$conteggio[20:31]
# two columns, one before and one after covid
df_1920_12 = as.data.frame(t(rbind(pre,post)))
colnames(df_1920_12) = c("before", "after") # columns names
#### Vertical Dataframae
df_{1920v_{12}} = as.data.frame(matrix(data=c(df_{1920_{12}},1],df_{1920_{12}},2],rep(0,12),rep(1,12)),
                                    nrow = 24, ncol = 2, byrow = F)
colnames(df_1920v_12) =c("conteggio", "periodo")
df_1920v_12$periodo = as.factor(df_1920v_12$periodo)
Second dataframe with 5 months and the correspondent 5 months in 2020 and 2019 (pre-covid)
# last 5 month
# horizontal datafrmae
pre = dati all$conteggio[8:12]
post = dati_all$conteggio[20:24]
post2 = dati_all$conteggio[32:36]
df_1920_5 = as.data.frame(t(rbind(pre,post,post2)))
colnames(df_1920_5) = c("before", "after", "last5") # columns names
#### Vertical dataframe
df_1920v_5 = as.data.frame(matrix(data=c(pre,post,post2,rep(0,5),rep(1,5),rep(2,5)),
                                    nrow = 15, ncol = 3, byrow = F)
colnames(df_1920v_5) =c("conteggio","periodo") # columns names
df_1920v_5$periodo = as.factor(df_1920v_5$periodo)
Plot 2 - 12 months comparasion
time_mesi=seq(as.Date("2019-2-1"), as.Date("2020-1-1"), by = "months")
ggplot(df_1920_12, aes(time_mesi)) +
  ggtitle("Cancer diagnosticated, between febrauary 2019
          and march 2020 compared with count from february 2020 and january 2021") +
  geom_line(aes(y = before, colour = "year 19/20")) +
  geom_line(aes(y = after, colour = "year 20/21")) +
  xlab("month") +
  ylab("Diagnosticated cancer") +
  scale x date(breaks = '1 month',
```

Cancer diagnosticated, between febrauary 2019 and march 2020 compared with count from february 2020 and january 2021



Plot 3 - Boxplot comparasion

Cancer diagnosticated, between febrauary 2019 and march 2020 compared with count from february 2020 and january 2021



Mean difference observed

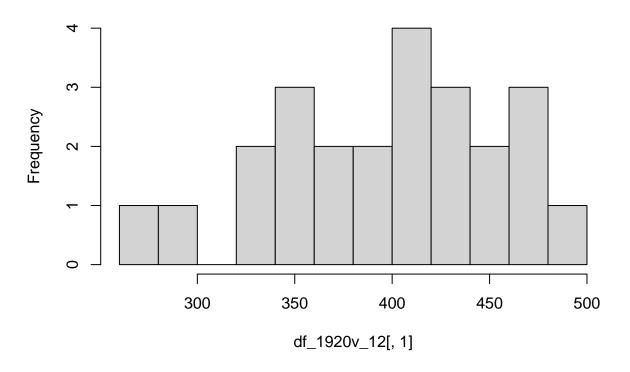
```
n1_12 = sum(df_1920_12$before);n1_12 # numerosità nei 12 mesi prima
## [1] 5210
n2_12 = sum(df_1920_12$after);n2_12 # numerosità nei 12 mesi after
## [1] 4339
1-(n2_12/n1_12) # rapporto che denota un calo del 16.7%
```

[1] 0.1671785

Data distribution analysis

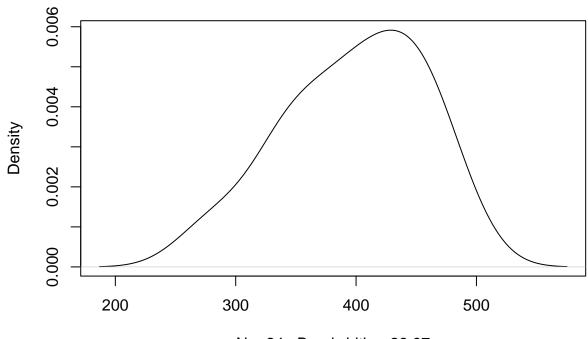
 $hist(df_1920v_12[,1], breaks = 8)$

Histogram of df_1920v_12[, 1]



```
# Kernel Density Plot
d <- density(df_1920v_12$conteggio) # returns the density data
plot(d) # plots the results</pre>
```

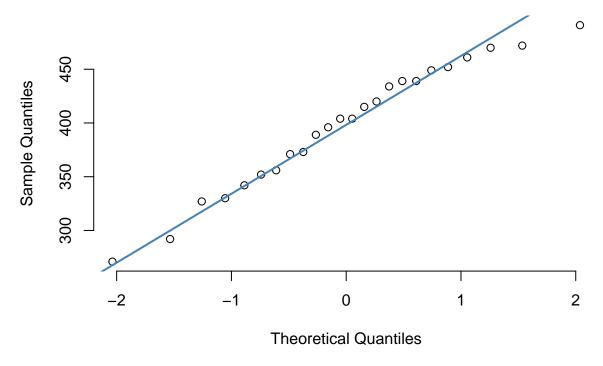
density.default(x = df_1920v_12\$conteggio)



N = 24 Bandwidth = 28.07

```
qqnorm(df_1920v_12$conteggio, pch = 1, frame = FALSE)
qqline(df_1920v_12$conteggio, col = "steelblue", lwd = 2)
```

Normal Q-Q Plot



```
# qqplot(df_1920v$conteggio)
shapiro.test(df_1920v_12$conteggio)
##
##
    Shapiro-Wilk normality test
##
## data: df_1920v_12$conteggio
## W = 0.96812, p-value = 0.6208
shapiro.test(df_1920_12$before)
##
##
    Shapiro-Wilk normality test
##
## data: df_1920_12$before
## W = 0.93423, p-value = 0.4271
shapiro.test(df_1920_12$after)
##
##
    Shapiro-Wilk normality test
##
  data: df_1920_12$after
## W = 0.97775, p-value = 0.9731
```

Normality can be assumed but the nature of the data and the dimension of the sample suggest the try a distribution for positive counting data like Poisson or Negative Binomial.

In our case Poisson can't be used because the variance is much bigger compared to the mean of the distribution.

(violation of one of the assumption for the Poisson distribution)

Checking mean and variance difference

```
mean(df_1920_12$before)

## [1] 434.1667

mean(df_1920_12$after)

## [1] 361.5833

var(df_1920_12)

## before after

## before 1498.697 714.803

## after 714.803 2880.629
```

Test for equal variance in both the sample

```
bartlett.test(df_1920v_12$conteggio~df_1920v_12$periodo)
```

```
##
## Bartlett test of homogeneity of variances
##
## data: df_1920v_12$conteggio by df_1920v_12$periodo
## Bartlett's K-squared = 1.1036, df = 1, p-value = 0.2935
```

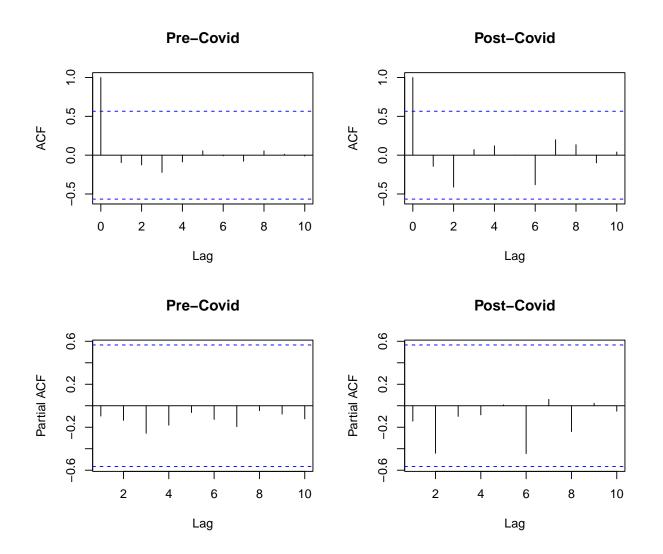
Confirmed the equal variance.

Now watch the autocorrelation function and partial autocorrelation function for check the independence assumption of our observation.

```
par(mfrow=c(2,2)) # 2 disegni in un riquadro

plt_ts_1 = acf(ts(df_1920_12[,1]), plot=F)
plot(plt_ts_1, main="Pre-Covid")
plt_ts_2 = acf(ts(df_1920_12[,2]), plot=F)
plot(plt_ts_2, main="Post-Covid")

plt_ts_1 = pacf(ts(df_1920_12[,1]), plot=F)
plot(plt_ts_1, main="Pre-Covid")
plt_ts_2 = pacf(ts(df_1920_12[,2]), plot=F)
plot(plt_ts_2, main="Post-Covid")
```



No big signal of correlation between observation in different times, we know about august seasonality.

Dummy variables for model the strong seasonality of august

```
# dummy
dm=make.dummy(length(df_1920v_12[,1]),start=2, freq=12)
# dm[1:6,]
# give names to the dummies
nomi=c("dm4","dm5","dm6","dm7","dm8","dm9","dm10","dm11","dm12","dm1","dm2","dm3")
dimnames(dm)=list(NULL,nomi)
colnames(dm)<-nomi</pre>
```

Linear model with normal assumption, t-test and Wilconox-Mann test

```
# 12 month
# Linear model
summary(lm(df_1920v_12$conteggio~df_1920v_12$periodo+dm[,8]))
##
## Call:
## lm(formula = df_1920v_12$conteggio ~ df_1920v_12$periodo + dm[,
## 8])
```

```
##
## Residuals:
     Min
              1Q Median
                            3Q
## -77.89 -19.04
                 0.00 18.67 100.11
## Coefficients:
                        Estimate Std. Error t value Pr(>|t|)
                                      11.15 39.691 < 2e-16 ***
## (Intercept)
                          442.47
## df_1920v_12$periodo1
                          -72.58
                                      15.42 -4.707 0.00012 ***
## dm[, 8]
                          -99.68
                                      27.89 -3.574 0.00179 **
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 37.77 on 21 degrees of freedom
## Multiple R-squared: 0.6245, Adjusted R-squared: 0.5888
## F-statistic: 17.47 on 2 and 21 DF, p-value: 3.412e-05
# T test:
t.test(df_1920v_12$conteggio~df_1920v_12$periodo, var.equal =T)
##
##
   Two Sample t-test
##
## data: df_1920v_12$conteggio by df_1920v_12$periodo
## t = 3.7995, df = 22, p-value = 0.0009823
## alternative hypothesis: true difference in means between group 0 and group 1 is not equal to 0
## 95 percent confidence interval:
   32.96509 112.20157
## sample estimates:
## mean in group 0 mean in group 1
         434.1667
                          361.5833
##
# Wilcoxon-Mann-Whitney test:
wilcox.test(df_1920v_12$conteggio~df_1920v_12$periodo)
## Warning in wilcox.test.default(x = c(449, 439, 434, 491, 415, 461, 342, : cannot
## compute exact p-value with ties
##
##
  Wilcoxon rank sum test with continuity correction
##
## data: df_1920v_12$conteggio by df_1920v_12$periodo
## W = 125, p-value = 0.002426
## alternative hypothesis: true location shift is not equal to 0
Test with negative binomial assumption for overdispersed counting data
# 12 month
bn_covid = glm.nb(df_1920v_12$conteggio~df_1920v_12$periodo+dm[,8])
summary(bn_covid)
##
## Call:
## glm.nb(formula = df_1920v_12$conteggio ~ df_1920v_12$periodo +
       dm[, 8], init.theta = 175.0351894, link = log)
```

Deviance Residuals:

```
##
                   1Q
                         Median
                                                Max
                                       3Q
## -2.42351 -0.53273 -0.02723
                                  0.47449
                                            2.77921
##
## Coefficients:
##
                        Estimate Std. Error z value Pr(>|z|)
                                    0.02641 230.765 < 2e-16 ***
## (Intercept)
                         6.09452
## df_1920v_12$periodo1 -0.18359
                                    0.03710 -4.949 7.47e-07 ***
                                    0.06979 -4.057 4.98e-05 ***
## dm[, 8]
                        -0.28314
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for Negative Binomial(175.0352) family taken to be 1)
##
      Null deviance: 64.633 on 23 degrees of freedom
##
## Residual deviance: 24.213 on 21 degrees of freedom
## AIC: 248.08
##
## Number of Fisher Scoring iterations: 1
##
##
##
                 Theta: 175.0
##
             Std. Err.: 73.6
##
## 2 x log-likelihood: -240.077
# percentage difference after covid can be calculated with 1-exp(-0.18359)
1-\exp(-0.18359)
## [1] 0.167723
Check for correct assumption of Negative binomial instead of Poisson
poi_covid <- glm(df_1920v_12$conteggio ~ df_1920v_12$periodo, family = "poisson")
pchisq(2 * (logLik(bn_covid) - logLik(poi_covid)), df = 1, lower.tail = FALSE)
```

'log Lik.' 6.578825e-18 (df=4)

confirm the correct assumption of negative binomial instead of Poisson.

Both models confirm the mean difference between the two sample. Same conclusion.