

Lecture 11

September 27, 2018

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- ▶ We first construct a reduced model which incorporates the hypothesis in the full model M . Call this reduced model m . We then look at the quantity:

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- ▶ It makes sense to reject the null hypothesis if T is large. To answer the question: *how large is large?*
- ▶ We rely on the assumption of normality of the errors i.e., $e \sim N(0, \sigma^2 I)$ to assert that $T \sim F_{p-q, n-p-1}$ under H_0

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- ▶ Suppose we do not want to assume normality of errors. Is there any way to obtain a p -value? This is possible in some cases via permutation tests. We provide two examples below.

Testing for all explanatory variables

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- ▶ Under the null hypothesis, we assume that if the response variable y has no relation to the explanatory variables.
- ▶ Therefore, it is plausible to assume that under the null hypothesis, the values of the response variable y_1, \dots, y_n are randomly distributed between the n subjects without relation to the predictors.

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3. Repeat the above pair of steps a large number of times.
4. This results in a large number of values of the test statistic (one for each permutation of the response values). Let us call them T_1, \dots, T_N . The p -value is calculated as the proportion of T_1, \dots, T_N that exceed the original test statistic value T (T is calculated with the actual unpermuted response values y_1, \dots, y_n).

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- ▶ The idea is to do this by calculating the statistic after permuting the response values.

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- ▶ Because once the response values are permuted, all association between the response and explanatory variables breaks down so that the values of

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- ▶ The p -value is then calculated as the proportion of these values larger than the observed value.

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- ▶ How to do this without normality?
- ▶ We can follow the permutation test by permuting the values of x_1 .
- ▶ For each permutation, we calculate the t -statistic and the p -value is the proportion of these t -values which are larger than the observed t -value in absolute value.

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- ▶ The problem with permutation tests is that they cannot be used to test arbitrary hypotheses about β (for example, there is no natural permutation test for the hypothesis $H_0 : \beta_1 = \beta_2 + \beta_5, \beta_2 - 4\beta_3 = 2$).

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- ▶ Moreover, there is no way of obtaining confidence intervals for components of β by permutation methods.
- ▶ The bootstrap is a general technique that can be used to obtain confidence intervals and to carry out hypothesis tests without reliance on the normal regression theory.

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 3. The sample median $\hat{\mu}_3$.

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- ▶ This gives us that $\hat{\mu}_1 - \mu$ is approximately distributed according to the normal distribution with mean 0 and variance σ^2/n .

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- ▶ The standard error of $\hat{\mu}_1$ can then be approximated by $\hat{\sigma}n^{-1/2}$ where

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- ▶ Even if F is completely known, it is not easy to write down a formula for the asymptotic distribution of $\hat{\mu}_j - \mu$.
- ▶ But if F is known, one can use simulation on the computer to exactly determine the distribution of $\hat{\mu}_j - \mu$.

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4. The values $\hat{\mu}_2^{(j)} - 5$ for $j = 1, \dots, N$ form a (very large) sample from the distribution of $\hat{\mu}_2 - \mu$. These determine the distribution of $\hat{\mu}_2 - \mu$.

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 1. One can assume that \hat{F} has a parametric form such as a normal distribution and then estimate the parameters from the observed data X_1, \dots, X_n . This is called *parametric bootstrap*.

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- ▶ Different versions of the bootstrap correspond to different choices of the estimate \hat{F} . There are two broad choices for \hat{F} :
 1. One can assume that \hat{F} has a parametric form such as a normal distribution and then estimate the parameters from the observed data X_1, \dots, X_n . This is called *parametric bootstrap*.
 2. One can try to estimate F nonparametrically. The most common way of doing this is to take \hat{F} to be the *empirical distribution* of the data X_1, \dots, X_n . This is called *nonparametric bootstrap*. Usually people mean this when referring to the bootstrap.

The Bootstrap

- ▶ The empirical distribution of the data X_1, \dots, X_n is a discrete probability measure that gives the mass $1/n$ to each data point X_1, \dots, X_n . Its distribution function is given by

$$\hat{F}(t) := \frac{1}{n} (\text{number of points } X_1, \dots, X_n \text{ that are } \leq t).$$

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- ▶ Remember the notion of empirical distribution. The empirical distribution of the data X_1, \dots, X_n is a discrete probability measure that gives the mass $1/n$ to each data point X_1, \dots, X_n .

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 - 3.1 Generate n observations $X_1^{(j)}, \dots, X_n^{(j)}$ from \hat{F} .
 - 3.2 Compute the statistic T from the generated observations $X_1^{(j)}, \dots, X_n^{(j)}$. Call the computed value $T^{(j)}$.

The Bootstrap Algorithm for i.i.d data

- ▶ Its distribution function is given by

$$\hat{F}(t) := \frac{1}{n} (\text{number of points } X_1, \dots, X_n \text{ that are } \leq t).$$

- ▶ The following is the bootstrap algorithm for approximating the distribution of $T - \mu$:
 1. Estimate the distribution of the data F by the empirical distribution \hat{F} .
 2. Calculate the value of the parameter μ for the empirical distribution \hat{F} . Let us call this value $\mu(\hat{F})$.
 3. Fix a large number N (say $N = 5000$) and repeat the following steps for $j = 1, \dots, N$.
 - 3.1 Generate n observations $X_1^{(j)}, \dots, X_n^{(j)}$ from \hat{F} .
 - 3.2 Compute the statistic T from the generated observations $X_1^{(j)}, \dots, X_n^{(j)}$. Call the computed value $T^{(j)}$.
 4. The empirical distribution of the values $T^{(j)} - \mu(\hat{F})$ for $j = 1, \dots, N$ is an estimate of the distribution of $T - \mu$.