

$$a) P(Z=1; \theta) = w \quad P(Z=2; \theta) = 1-w$$

joint likelihood

$$P(X=x, Z=1; \theta) = P(Z=1; \theta) \cdot P(X=x|Z=1; \theta) = w \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_1)^2}{2}}$$

$$P(X=x, Z=2; \theta) = P(Z=2; \theta) P(X=x|Z=2; \theta) = (1-w) \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_2)^2}{2}}$$

marginal likelihood

$$P(X=x; \theta) = \sum_{k=1}^2 P(X=x, Z=k; \theta) = \sum_{k=1}^2 P(Z=k; \theta) P(X=x|Z=k; \theta)$$

$$= w \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_1)^2}{2}} + (1-w) \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_2)^2}{2}}$$

log-likelihood

$$\ell(X=x; \theta) = \log \sum_{k=1}^2 P(X=x, Z=k; \theta)$$

n i.i.d.

$$\ell(X_1=x_1, \dots, X_n=x_n; \theta) = \log \left(\prod_{i=1}^n \sum_{k=1}^2 P(X_i=x_i, Z=k; \theta) \right) = \sum_{i=1}^n \log \sum_{k=1}^2 P(X_i=x_i, Z=k; \theta)$$

$$(b) \ell(X_i; \theta) = \log P(X_i = x_i; \theta) = \log \sum_{k=1}^2 P(X=x_i, Z=k; \theta) \quad \text{let } q_i \text{ be some distribution}$$

$$= \log \sum_{k=1}^2 \frac{P(X=x_i, Z=k; \theta)}{q_i(k)} \cdot q_i(k) \quad \sum_{k=1}^2 q_i(k) = 1, q_i(k) \geq 0$$

$$= \log \mathbb{E} \frac{P(X=x_i, Z=k; \theta)}{q_i(k)}$$

Since \log is concave, by Jensen's inequality

$$\mathbb{E} \phi(x) \leq \phi(\mathbb{E}(x))$$

$$\geq \mathbb{E} \log \frac{P(X=x_i, Z=k; \theta)}{q_i(k)} = F_i(\theta; q_i)$$

$$= \sum_{k=1}^2 q_i(k) \cdot \log \frac{P(X=x_i, Z=k; \theta)}{q_i(k)}$$

$$= \sum_{k=1}^2 q_i(k) \log P(X_i, k; \theta) + \sum_{k=1}^2 q_i(k) \log \left(\frac{1}{q_i(k)} \right)$$

$$\ell(\{X_i\}_{i=1}^n; \theta) \geq F(\theta; q) = \sum_{i=1}^n F_i(\theta; q_i)$$

c). for above Jensen's inequality to hold with equality, we want the expectation to be taken over a constant, such that equality holds

for a single observation X_i , let $\frac{P(X=X_i, Z=z_i; \theta)}{q_i(Z=z_i)} = C$

$$q_i(Z=z_i) \propto P(X=X_i, Z=z_i; \theta)$$

Since $q_i(Z=z_i)$ is a distribution, its summation is 1

$$\begin{aligned} q_i(Z=z_i) &= \frac{P(X=X_i, Z=z_i; \theta)}{\sum_{i=1}^n P(X=X_i, Z=z_i; \theta)} \\ &= \frac{P(X=X_i, Z=z_i; \theta)}{P(X=X_i; \theta)} \end{aligned}$$

$$= P(Z=z_i | X=X_i; \theta)$$

Then for n i.i.d. sample, given θ^t

$$q^{t+1}(Z_1=z_1, \dots, Z_n=z_n) = \prod_{i=1}^n P(Z=z_i | X=X_i; \theta^t)$$

(d) for given θ^t , we can calculate $P(X=x_i; \theta^t)$
 $P(Z=k; X=x_i; \theta^t)$

$$P(Z=1|X=x_i; \theta^t) = \frac{P(Z=1, X=x_i; \theta^t)}{P(X=x_i; \theta^t)} = \frac{P(Z=1, X=x_i; \theta^t)}{\sum_{k=1}^2 P(X=x_i, Z=k; \theta^t)}$$

$$= \frac{w^t \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_1^t)^2}{2}}}{w^t \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_1^t)^2}{2}} + (1-w^t) \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_2^t)^2}{2}}}$$

$$P(Z=2|X=x_i; \theta^t) = \frac{(1-w^t) \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_2^t)^2}{2}}}{w^t \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_1^t)^2}{2}} + (1-w^t) \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-u_2^t)^2}{2}}}$$

c)

$$\sum_{i=1}^n \ell(X_i; \theta, q_i^{t+1}) = L(\theta; q^{t+1}) = \sum_{i=1}^n \sum_{k=1}^2 q_{i,k} \log p(X_i, k; \theta)$$

$$= \sum_{i=1}^n \left[q_i^{t+1} \cdot \log(w \cdot \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{(X-u_1)^2}{2}}) + (1-q_i^{t+1}) \left(\log(1-w) \frac{1}{\sqrt{2\pi}} e^{-\frac{(X-u_2)^2}{2}} \right) \right]$$

$$= \sum_{i=1}^n \left[q_i^{t+1} \left(\log w - \log(\sqrt{2\pi}) - \frac{(X-u_1)^2}{2} \right) + (1-q_i^{t+1}) \left(\log(1-w) - \log(\sqrt{2\pi}) - \frac{(X-u_2)^2}{2} \right) \right]$$

$$= \sum_{i=1}^n \left[q_i^{t+1} \left(\log w - \frac{(X-u_1)^2}{2} \right) - q_i^{t+1} \log(\sqrt{2\pi}) + q_i^{t+1} \log(\sqrt{2\pi}) - \log(\sqrt{2\pi}) \right. \\ \left. + (1-q_i^{t+1}) \left(\log(1-w) - \frac{(X-u_2)^2}{2} \right) \right]$$

$$= \sum_{i=1}^n \left[q_i^{t+1} \left(\log w - \frac{(X-u_1)^2}{2} \right) + (1-q_i^{t+1}) \left(\log(1-w) - \frac{(X-u_2)^2}{2} \right) - \log(\sqrt{2\pi}) \right]$$

$$= C + \sum_{i=1}^n \left[q_i^{t+1} \left(\log w - \frac{(X-u_1)^2}{2} \right) + (1-q_i^{t+1}) \left(\log(1-w) - \frac{(X-u_2)^2}{2} \right) \right]$$

(f)

$$\nabla_{u_1} \mathcal{L}(\theta; q^{t+1}) = \nabla_{u_1} \sum_{i=1}^n q_i^{t+1} \left(\log w - \frac{(X_i - u_1)^2}{2} \right)$$

$$= \nabla_{u_1} \sum_{i=1}^n -q_i^{t+1} \frac{(X_i - u_1)^2}{2}$$

$$= \sum_{i=1}^n q_i^{t+1} \cdot (X_i - u_1) = 0$$

$$\sum_{i=1}^n q_i^{t+1} X_i = \sum_{i=1}^n q_i^{t+1} u_1$$

$$u_1^{t+1} = \frac{\sum_{i=1}^n q_i^{t+1} X_i}{\sum_{i=1}^n q_i^{t+1}}$$

$$\nabla_{u_2} \mathcal{L}(\theta; q^{t+1}) = \nabla_{u_2} \sum_{i=1}^n (1 - q_i^{t+1}) \left(\log(1 - w) - \frac{(X_i - u_2)^2}{2} \right)$$

$$= \nabla_{u_2} \sum_{i=1}^n -(1 - q_i^{t+1}) \frac{(X_i - u_2)^2}{2}$$

$$= \sum_{i=1}^n (1 - q_i^{t+1}) (X_i - u_2) = 0$$

$$\sum_{i=1}^n (1 - q_i^{t+1}) X_i = \sum_{i=1}^n (1 - q_i^{t+1}) u_2$$

$$u_2^{t+1} = \frac{\sum_{i=1}^n (1 - q_i^{t+1}) X_i}{\sum_{i=1}^n (1 - q_i^{t+1})}$$

$$\nabla_w \mathcal{L}(\theta; q^{t+1}) = \sum_{i=1}^n \left[q_i^{t+1} \cdot \frac{1}{w} + (1 - q_i^{t+1}) \cdot \frac{-1}{(1-w)} \right] = 0$$

$$\frac{1}{w} \sum_{i=1}^n q_i^{t+1} = \frac{1}{1-w} \sum_{i=1}^n (1 - q_i^{t+1})$$

$$\frac{1-w}{w} \sum_{i=1}^n q_i^{t+1} = n - \sum_{i=1}^n q_i^{t+1}$$

$$\frac{1}{w} - 1 = \frac{n}{\sum_{i=1}^n q_i^{t+1}} - 1$$

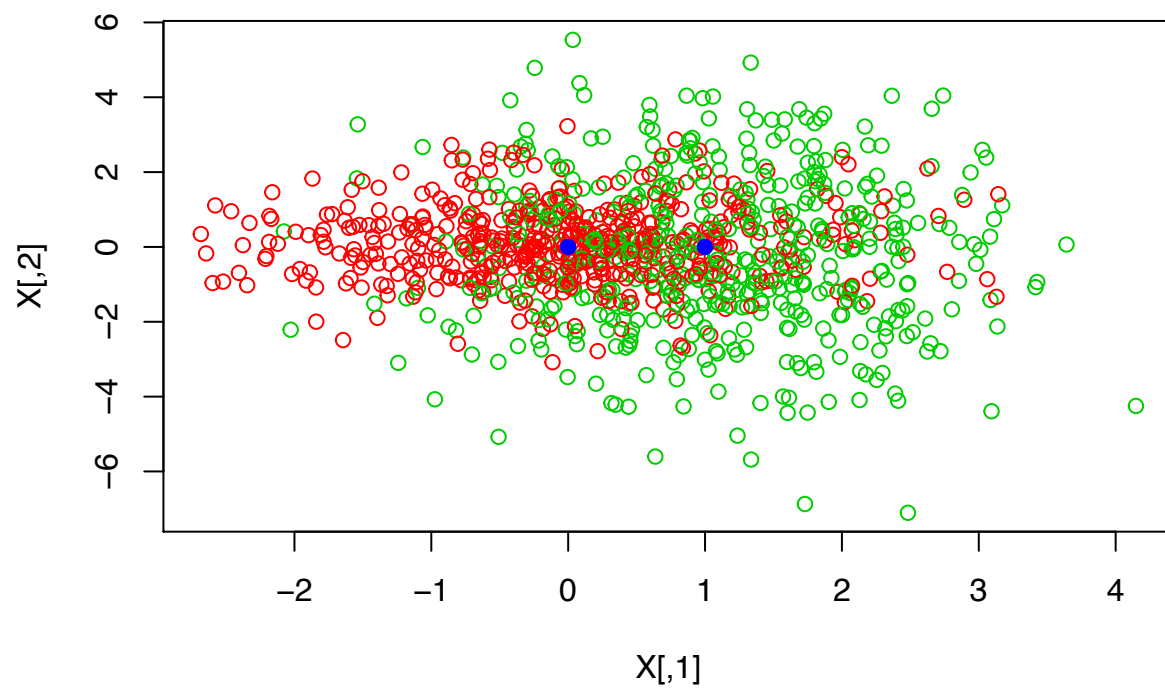
$$w^{t+1} = \frac{\sum_{i=1}^n q_i^{t+1}}{n}$$

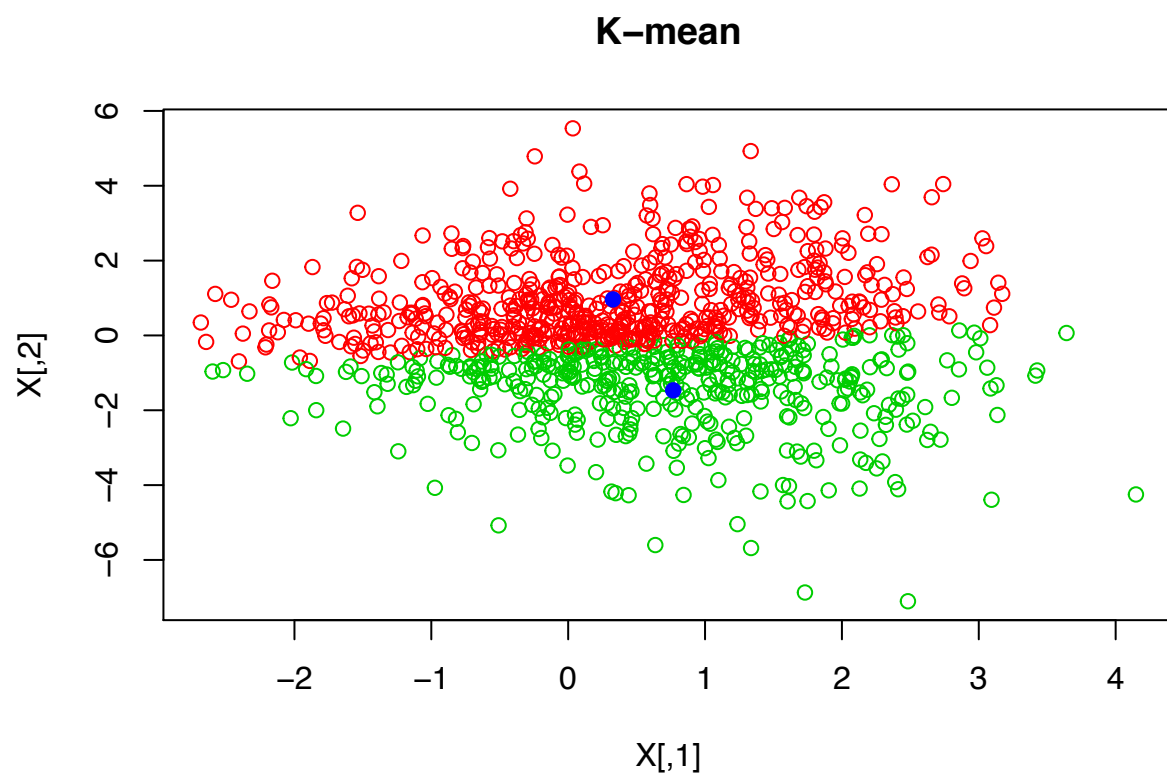
problem(g)

caojilin

3/4/2019

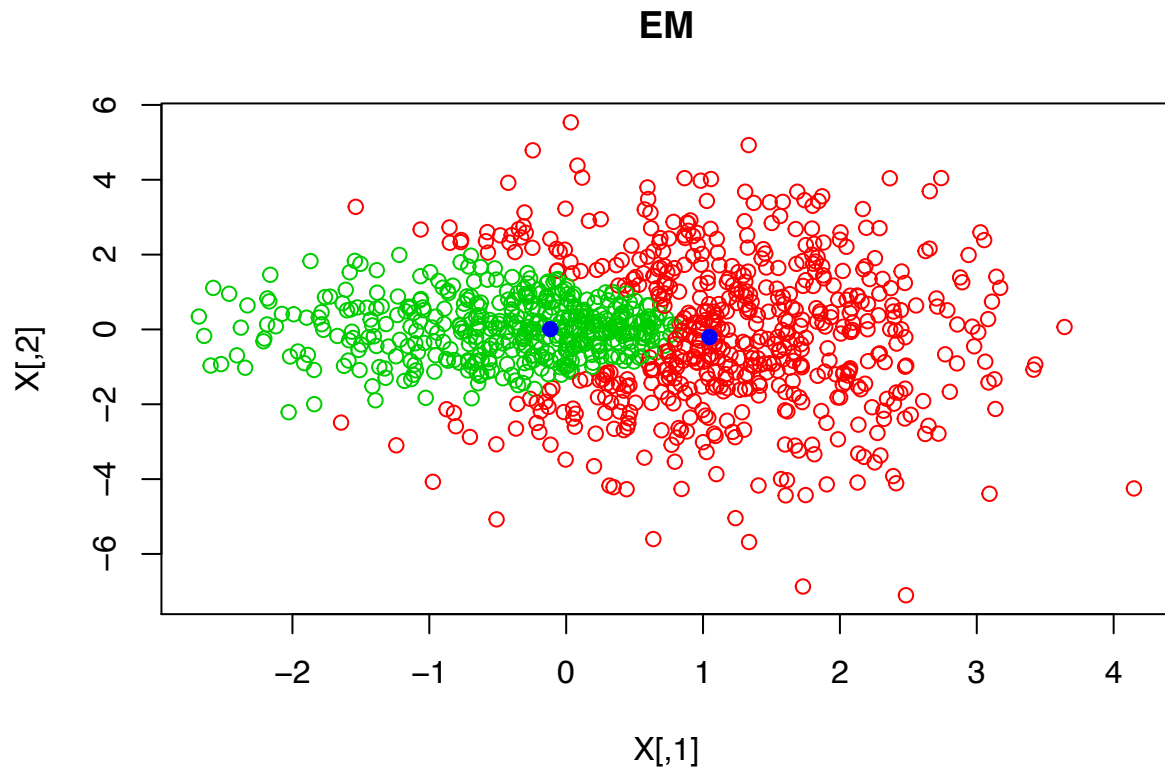
Original





```
## [1] "EM centers"
```

```
##           [,1]      [,2]  
## [1,]  1.0467685 -0.203373117  
## [2,] -0.1157559  0.008598925
```

K-means only calculates Euclidean distance, which is like a non-parametric approach, while EM assumes there exist underlying distributions and uses likelihood to calculate, which is more like a parametric approach. We know that original data is almost non-separable. K-mean gives us a “hard” classification while EM gives us a “soft” classification. Thus the estimated labels are different.