

# Appendix - Forecasting Arctic Sea Ice

TSA FINAL PROJECT - Time Series Analysis (MATH1318)

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*04/06/2019*

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# 1 Data pre-processing

- Data Source: <https://sites.google.com/site/arcticseaicegraphs/>

## 1.1 Load libraries

## 1.2 Load data set

```
sea_ice_orig <- read.csv("sea_ice_arctic.csv")

colnames(sea_ice_orig) <- c('Year', 'Jan', 'Feb', 'Mar', 'Apr', 'May', 'Jun', 'Jul', 'Aug', 'Sep', 'Oct', 'Nov', 'Dec')

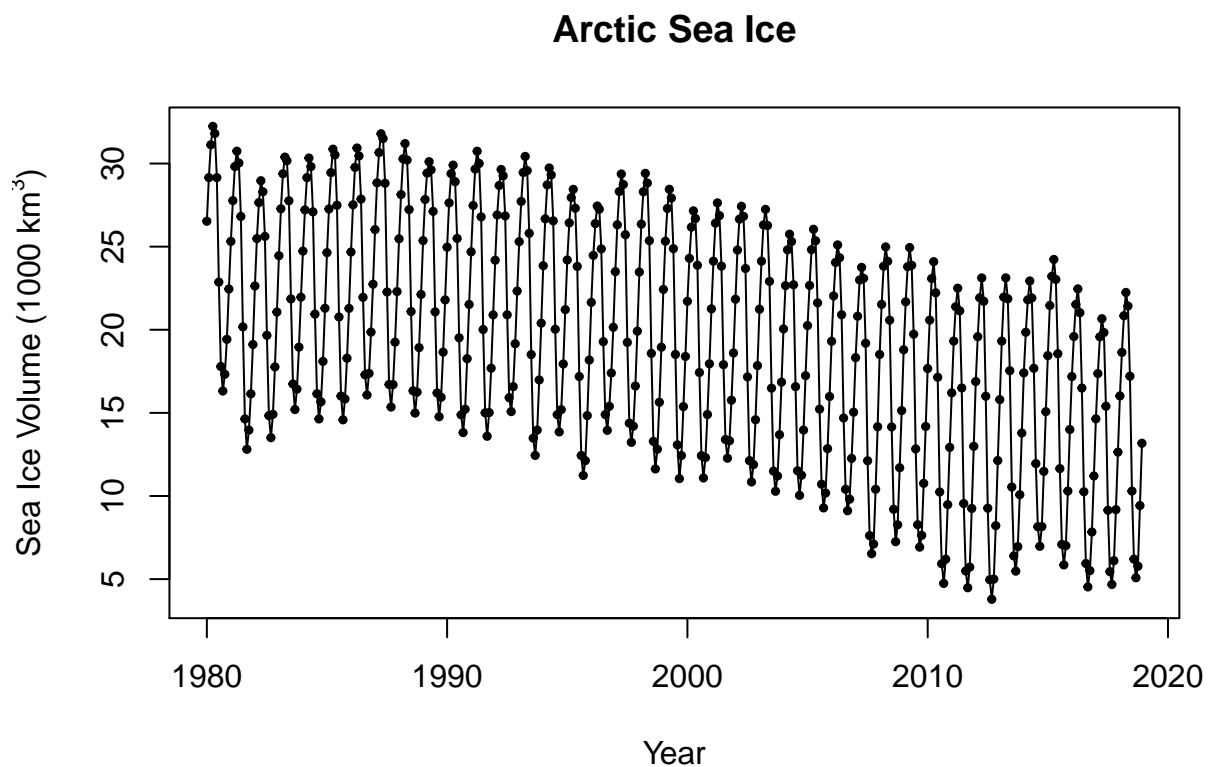
sea_ice_orig <- sea_ice_orig %>% gather('Jan', 'Feb', 'Mar', 'Apr', 'May', 'Jun', 'Jul', 'Aug', 'Sep', 'Oct', 'Nov', 'Dec')
```

## 1.3 Convert to time series

```
sea_ice_orig_ts <- ts(sea_ice_orig$volume, start = 1980, frequency = 12)
```

## 1.4 Plot data

```
plot(sea_ice_orig_ts, type = "o", pch = 19, cex = 0.5, xlab = "Year", ylab = expression(paste("Sea Ice Volume (1000 km3)")))
```



## 1.5 Load truncated data set

```
sea_ice_long <- read.csv("sea_ice_arctic2.csv")

colnames(sea_ice_long) <- c('Year', 'Jan', 'Feb', 'Mar', 'Apr', 'May', 'Jun', 'Jul', 'Aug', 'Sep', 'Oct')

sea_ice_long <- sea_ice_long %>% gather('Jan', 'Feb', 'Mar', 'Apr', 'May', 'Jun', 'Jul', 'Aug', 'Sep',
```

## 1.6 Convert truncated data set to time series

```
sea_ice_long_ts <- ts(sea_ice_long$volume, start = 1986, frequency = 12)
lambda = 0.85
sea_ice_long_BC <- (sea_ice_long_ts^lambda-1)/lambda

sea_ice <- read.csv("sea_ice_arctic3.csv")

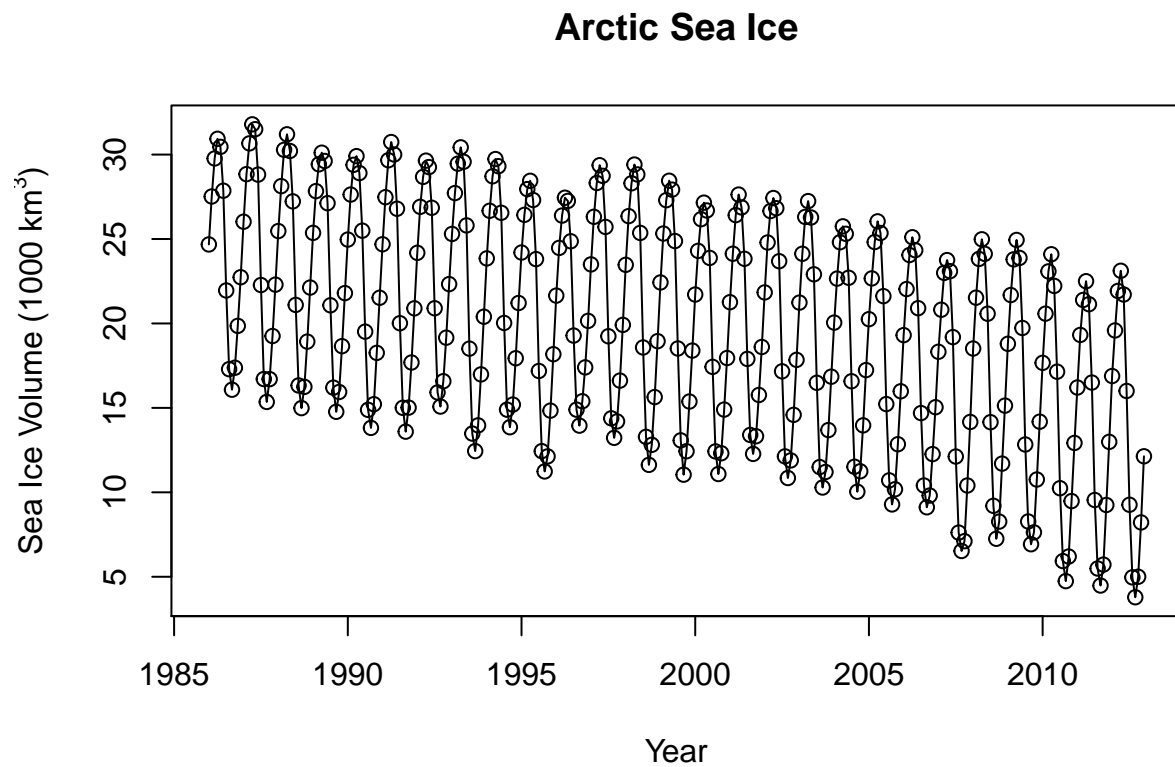
colnames(sea_ice) <- c('Year', 'Jan', 'Feb', 'Mar', 'Apr', 'May', 'Jun', 'Jul', 'Aug', 'Sep', 'Oct', 'Nov')

sea_ice <- sea_ice %>% gather('Jan', 'Feb', 'Mar', 'Apr', 'May', 'Jun', 'Jul', 'Aug', 'Sep', 'Oct', 'Nov')

sea_ice_ts <- ts(sea_ice$volume, start = 1986, frequency = 12)
```

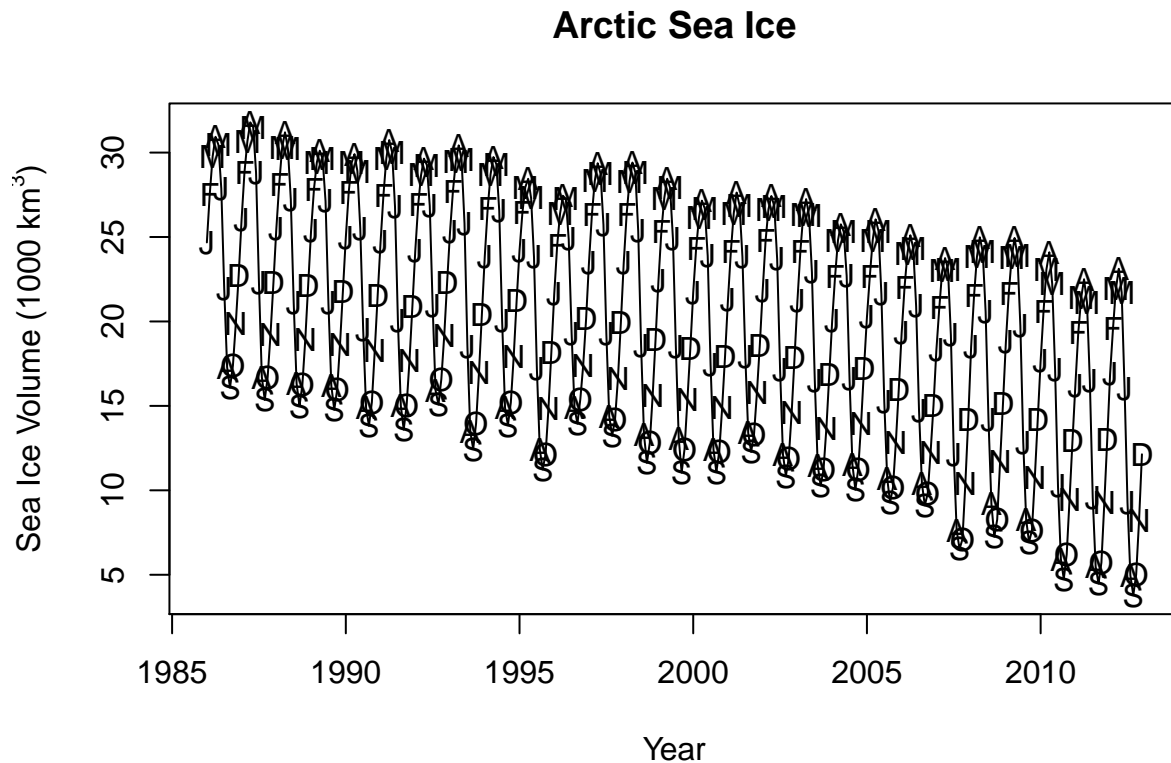
## 1.7 Plot truncated time series data

```
plot(sea_ice_ts, type = "o", xlab = "Year", ylab = expression(paste("Sea Ice Volume (1000 km"3", ")")
```



#### 1.8 Annotated plot time series of truncated data to examine seasonality

```
plot(sea_ice_ts, ylab=expression(paste("Sea Ice Volume (1000 km"3", ")")), xlab = 'Year', type = "l",
points(y=sea_ice_ts,x=time(sea_ice_ts),pch=as.vector(season(sea_ice_ts)))
```



- From the plot of the output we know:
  1. Seasonality is present.
  2. There is a downward trend (non-stationary).
  3. There is no obvious change in variance.
  4. It is not possible to determine behaviour due to the presence of seasonality.

## 2 Harmonic-Trend model

- Set up indicator variables that indicate the month to which each of the data points pertains before estimating parameters.

```
month.=season(sea_ice_ts) # period added to improve table display and this line sets up indicators
model2=lm(sea_ice_ts~month.-1) # -1 removes the intercept term
summary(model2)
```

```
##
## Call:
## lm(formula = sea_ice_ts ~ month. - 1)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -7.840 -2.387  0.302  2.705  5.265
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
```

```
## month.January    21.9192    0.6247    35.09    <2e-16 ***
## month.February   24.6609    0.6247    39.48    <2e-16 ***
## month.March      26.6870    0.6247    42.72    <2e-16 ***
## month.April      27.6801    0.6247    44.31    <2e-16 ***
## month.May        26.9158    0.6247    43.09    <2e-16 ***
## month.June       23.5491    0.6247    37.70    <2e-16 ***
## month.July       17.1039    0.6247    27.38    <2e-16 ***
## month.August     12.3081    0.6247    19.70    <2e-16 ***
## month.September  11.0253    0.6247    17.65    <2e-16 ***
## month.October    12.1887    0.6247    19.51    <2e-16 ***
## month.November   14.9781    0.6247    23.98    <2e-16 ***
## month.December   18.2201    0.6247    29.17    <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.246 on 312 degrees of freedom
## Multiple R-squared:  0.9768, Adjusted R-squared:  0.9759
## F-statistic: 1093 on 12 and 312 DF,  p-value: < 2.2e-16
```

All of the parameters corresponding to months are statistically significant at 5% level.

```
month.=season(sea_ice_ts) # period added to improve table display and this line sets up indicators
model3=lm(sea_ice_ts~month.) # -1 removes the intercept term
summary(model3)
```

```
##
## Call:
## lm(formula = sea_ice_ts ~ month.)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -7.840 -2.387  0.302  2.705  5.265
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    21.9192    0.6247   35.088 < 2e-16 ***
## month.February  2.7417    0.8834    3.103  0.00209 **
## month.March     4.7678    0.8834    5.397 1.34e-07 ***
## month.April     5.7609    0.8834    6.521 2.81e-10 ***
## month.May       4.9966    0.8834    5.656 3.51e-08 ***
## month.June      1.6300    0.8834    1.845  0.06598 .
## month.July     -4.8153    0.8834   -5.451 1.02e-07 ***
## month.August    -9.6110    0.8834  -10.879 < 2e-16 ***
## month.September -10.8939    0.8834  -12.331 < 2e-16 ***
## month.October   -9.7304    0.8834  -11.014 < 2e-16 ***
## month.November  -6.9411    0.8834   -7.857 6.42e-14 ***
## month.December  -3.6990    0.8834   -4.187 3.68e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.246 on 312 degrees of freedom
## Multiple R-squared:  0.7779, Adjusted R-squared:  0.7701
## F-statistic: 99.33 on 11 and 312 DF,  p-value: < 2.2e-16
```

## 2.1 Fit of cosine curve to average monthly Arctic sea ice series.

```
har.=harmonic(sea_ice_ts,1) # calculate cos(2*pi*t) and sin(2*pi*t)
t1 <- time(sea_ice_ts)
model4=lm(sea_ice_ts~har. + t1)
summary(model4)
```

```
##
## Call:
## lm(formula = sea_ice_ts ~ har. + t1)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.9501 -0.9066  0.0295  0.9480  3.0424
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    794.03512    20.30584    39.10  <2e-16 ***
## har.cos(2*pi*t)    2.09440     0.11190    18.72  <2e-16 ***
## har.sin(2*pi*t)    7.91857     0.11194    70.74  <2e-16 ***
## t1              -0.38724     0.01016   -38.13  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.424 on 320 degrees of freedom
## Multiple R-squared:  0.9561, Adjusted R-squared:  0.9557
## F-statistic: 2326 on 3 and 320 DF, p-value: < 2.2e-16
```

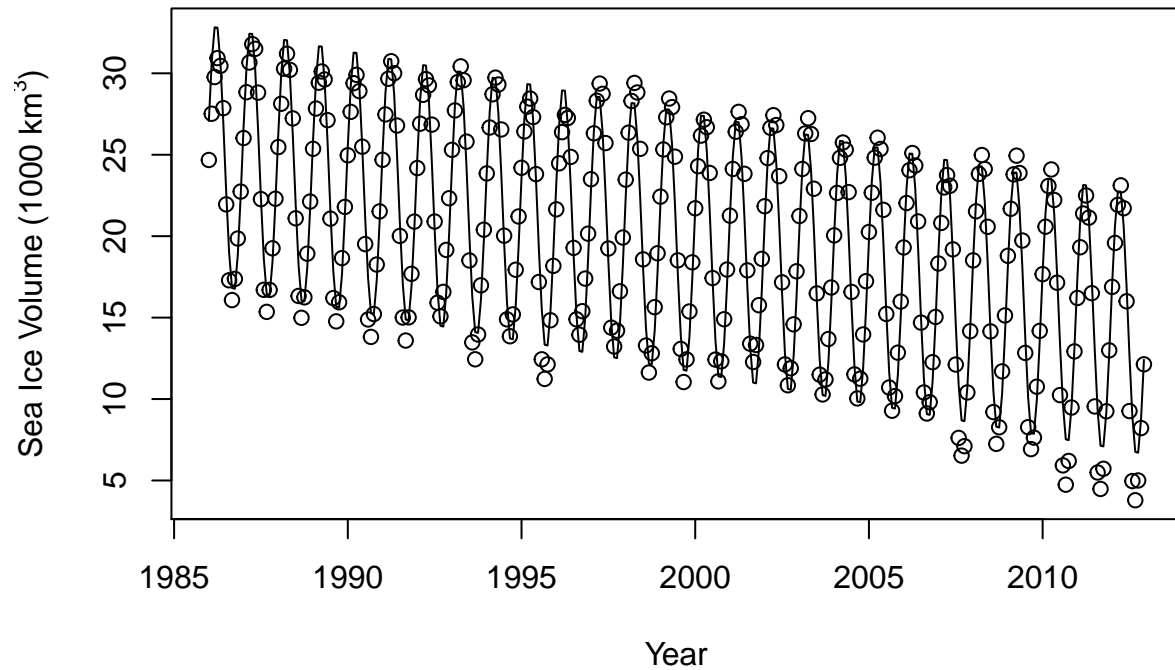
Low p-values indicate a high degree of significance for the determined parameters. The R-squared value is high (0.96) which is very good but care should be taken to ensure this is not due to overfitting.

## 2.2 Plot of fitted Harmonic-Trend curve along with observed average monthly Arctic sea ice series.

```
plot(ts(fitted(model4),freq=12,start=c(1986,1)),xlab = 'Year', ylab=expression(paste("Sea Ice Volume (1012 m3)")),ylim=range(c(fitted(model4),sea_ice_ts)),main="Fitted model to Sea Ice Volume") # ylim ensures that the points(sea_ice_ts)
```



## Fitted model to Sea Ice Volume

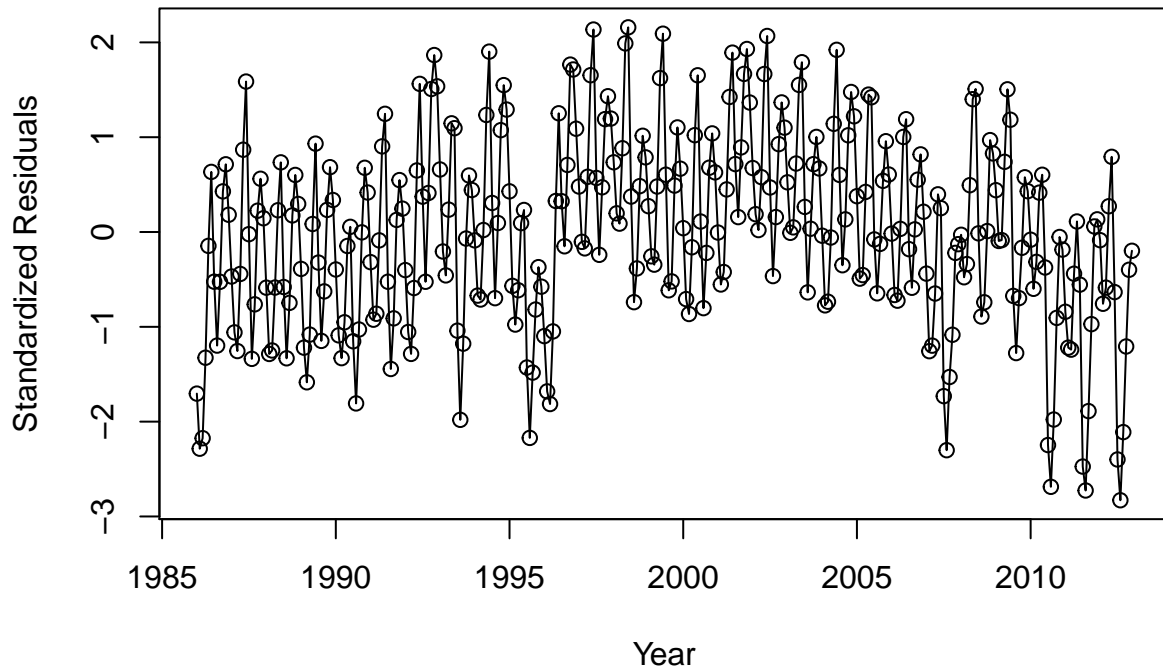


\* The Harmonic-Trend model appears to fit reasonably well to the average monthly Arctic sea ice data up to 2007 but there is some deviation beyond this point.

### 2.3 Time series plot of standardized residuals.

```
plot(y=rstudent(model4),x=as.vector(time(sea_ice_ts)), xlab='Year',ylab='Standardized Residuals',type='o')
```

## Time series plot of residuals

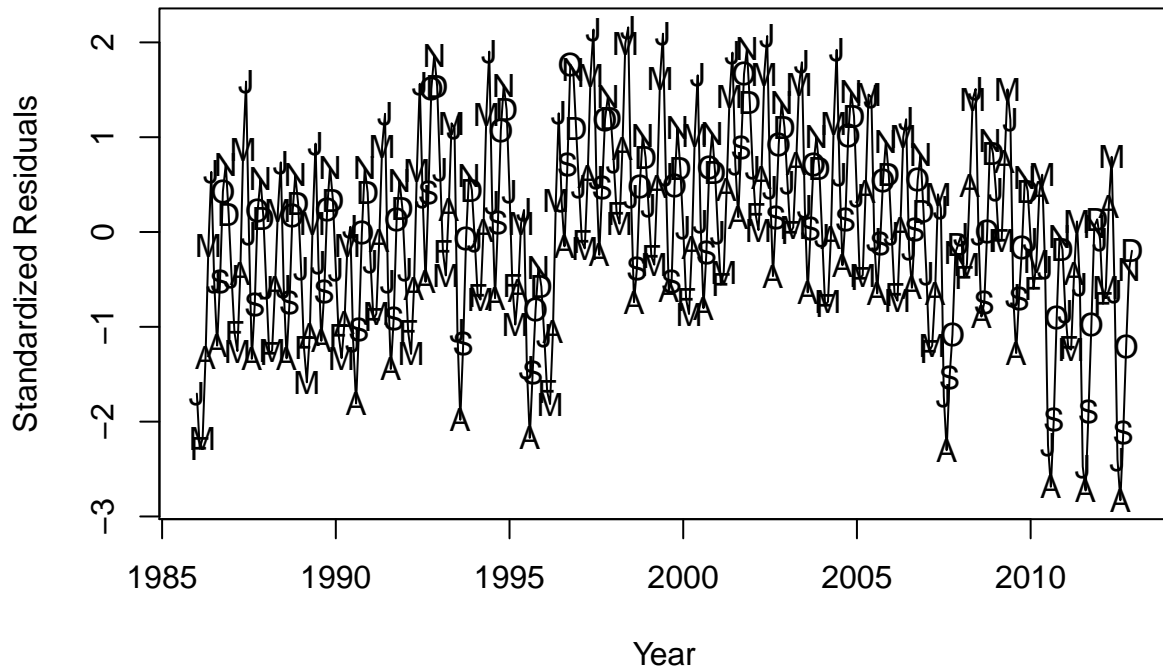


\* There was a departure from randomness in the plot of the standardized residuals. Therefore, labeled months to determine if there is a trend present.

### 2.4 Labeled months in plot of standardized residuals.

```
plot(y=rstudent(model4),x=as.vector(time(sea_ice_ts)),xlab='Year', ylab='Standardized Residuals',type='l')
points(y=rstudent(model4),x=as.vector(time(sea_ice_ts)), pch=as.vector(season(sea_ice_ts)))
```

## Time series plot of residuals

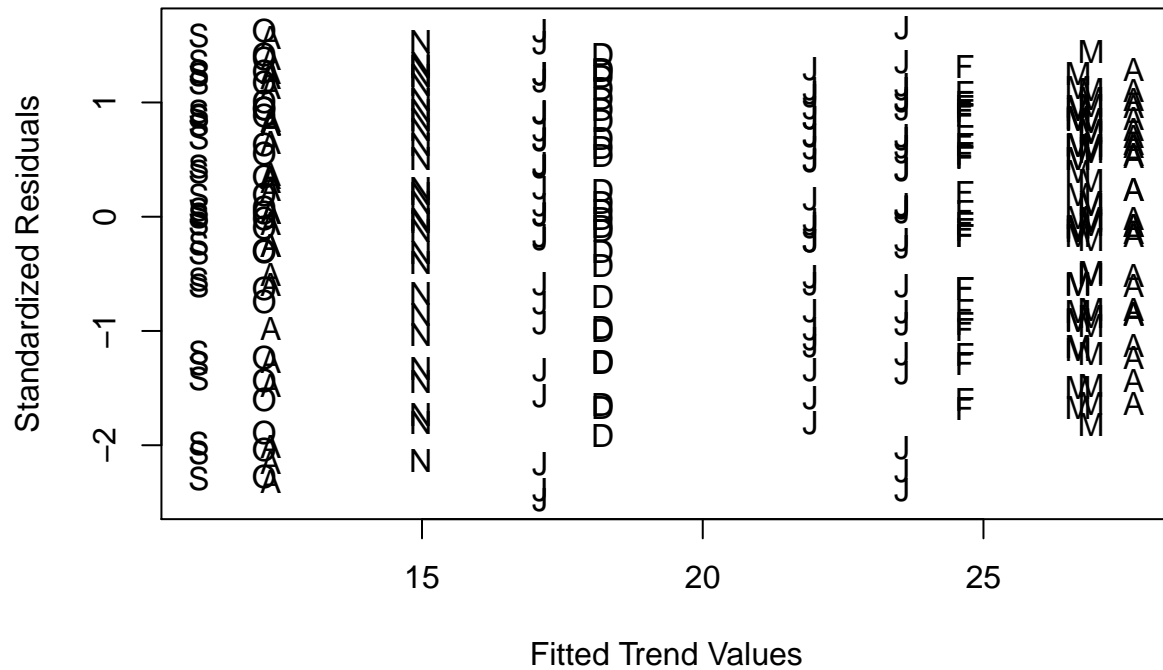


\* Some high and low points but appears more random than the original series.

## 2.5 Plot of standardized residuals with labels.

```
plot(y=rstudent(model3),x=as.vector(fitted(model3)), xlab='Fitted Trend Values', ylab='Standardized Residuals')
points(y=rstudent(model3),x=as.vector(fitted(model3)),pch=as.vector(season(sea_ice_ts)))
```

## Time series plot of standardised residuals



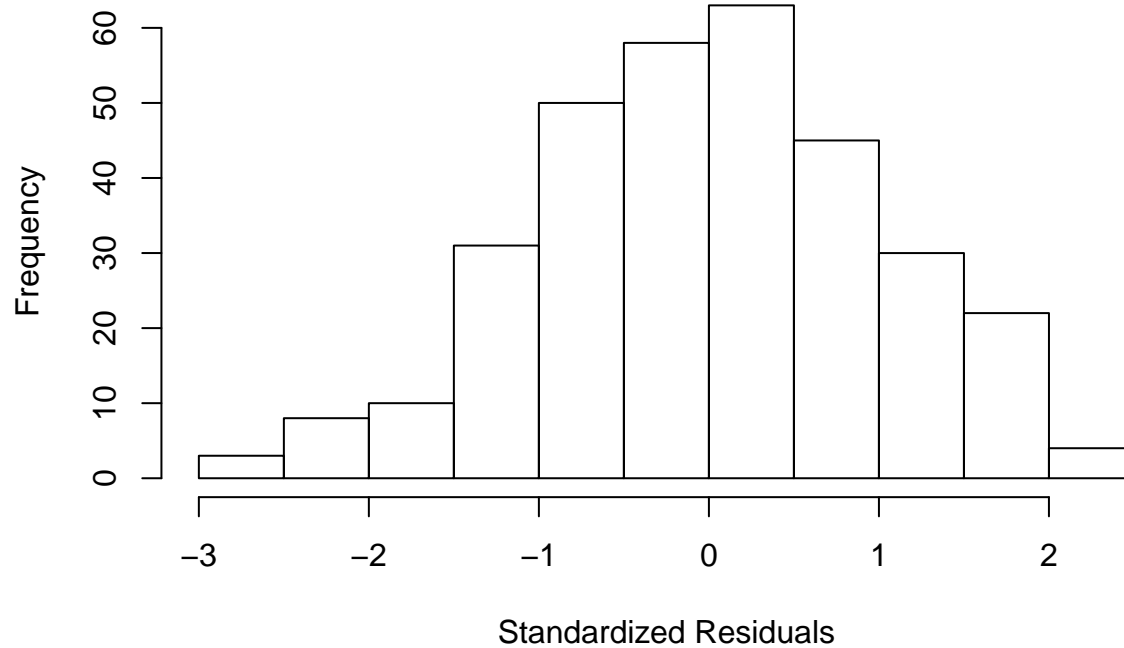
\* There was a similar spread of labels across the plot. The plot does not indicate any dramatic patterns that would cause us to doubt the seasonal means model.

## 2.6 Normality of standardized residuals

### 2.6.1 Histogram of standardized residuals for Harmoic-Trend model.

```
hist(rstudent(model4), xlab='Standardized Residuals', main = 'Histogram of Standardized Residuals')
```

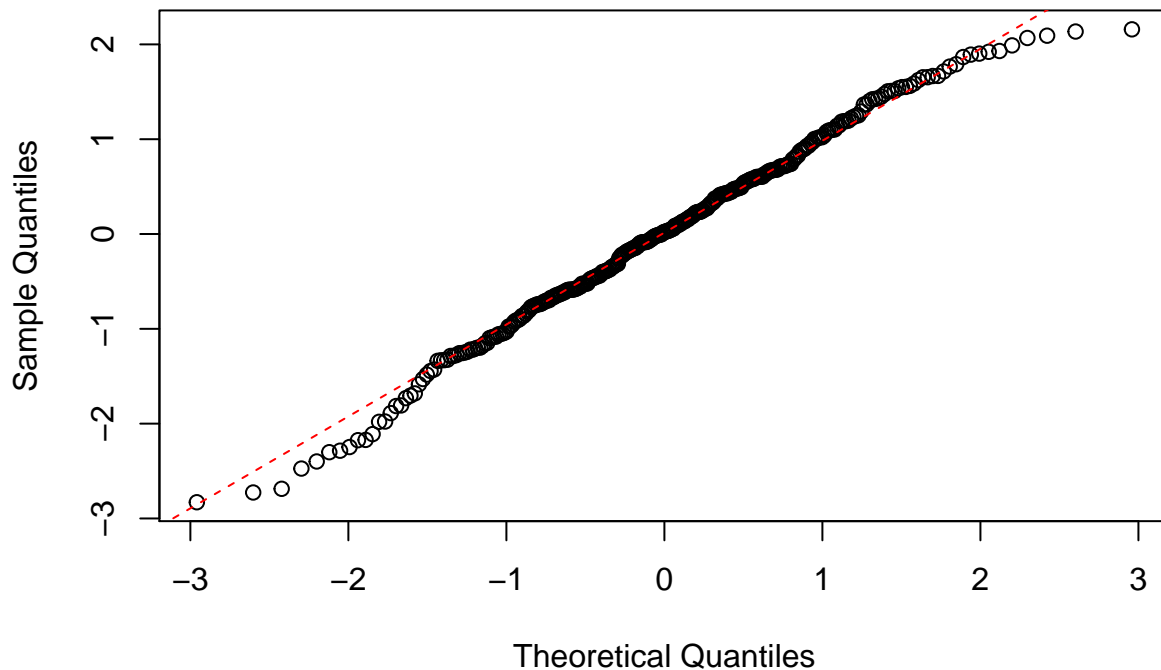
## Histogram of Standardized Residuals



### 2.6.2 Q-Q plot of standardized residuals for Harmonic-Trend model.

```
y = rstudent(model4)
qqnorm(y)
qqline(y, col = 2, lwd = 1, lty = 2)
```

## Normal Q-Q Plot

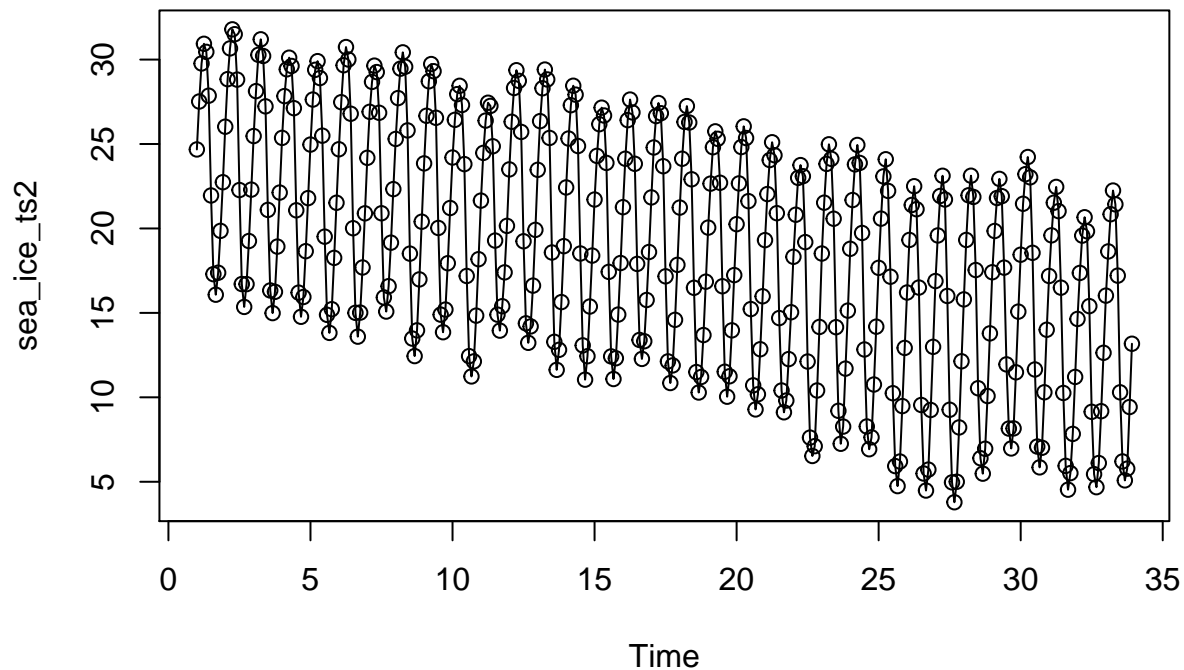


```
shapiro.test(model4$residuals)
```

```
##  
## Shapiro-Wilk normality test  
##  
## data:  model4$residuals  
## W = 0.99224, p-value = 0.08909
```

- The histogram, Q-Q plot of standardized residuals for the Harmonic-Trend model shows a normal distribution. The Shapiro-Wilk test gave a p-value of 0.089 indicating that we can not reject the null hypothesis that the stochastic component of this model is normally distributed.

```
sea_ice2 = read.csv("sea_ice_arctic2.csv")  
colnames(sea_ice2) <- c('Year', 'Jan', 'Feb', 'Mar', 'Apr', 'May', 'June', 'July', 'Aug', 'Sep', 'Oct', 'Nov', 'Dec')  
sea_ice2 <- sea_ice2 %>% gather('Jan', 'Feb', 'Mar', 'Apr', 'May', 'June', 'July', 'Aug', 'Sep', 'Oct', 'Nov', 'Dec')  
sea_ice_ts2 <- ts(sea_ice2$volume, start = 1, frequency = 12)  
plot(sea_ice_ts2, type = "o")
```



## 2.7 Three-year forecast of Arctic sea ice volumes from 2013 to 2015 using the Harmonic-Trend model.

```

har.=harmonic(sea_ice_ts2,1) # calculate cos(2*pi*t) and sin(2*pi*t)
t3 <- time(sea_ice_ts2)
t1 = har.[,1] # To make it easier assign harmonic variables to separate variables
t2 = har.[,2]
model4=lm(sea_ice_ts2~t3+t1+t2) # Fit the model with separate variables
# We need to create continuous time for 12 months starting from the first month of 2013
t = c(34.000, 34.083, 34.167 ,34.250, 34.333, 34.417 ,34.500, 34.583, 34.667, 34.750, 34.833, 34.917, 35.000)
t1 = cos(2*pi*t)
t2 = sin(2*pi*t)
t3 <- t
t3

## [1] 34.000 34.083 34.167 34.250 34.333 34.417 34.500 34.583 34.667 34.750
## [11] 34.833 34.917 35.000 35.083 35.167 35.250 35.333 35.417 35.500 35.583
## [21] 35.667 35.750 35.833 35.917 36.000 36.083 36.167 36.250 36.333 36.417
## [31] 36.500 36.583 36.667 36.750 36.833 36.917 37.000

new = data.frame(t3, t1 , t2) # Step 1
# Notice here that I'm using the same variable names "t1" and "t2" as in the
# fitted model above, where the name of the variables showing sine and cosine
# components are also "t1" and "t2". To run the predict() function properly,
# the names of variables in fitted model and "new" data frame

```

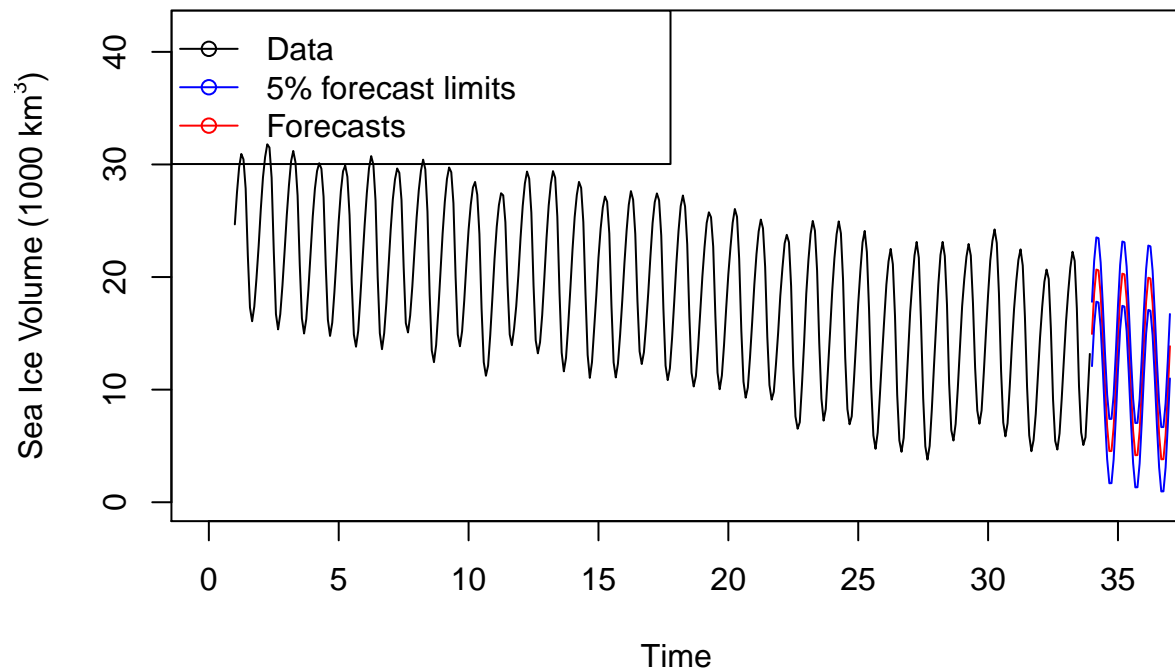
```
# must be the same!!!
forecasts = predict(model4, new, interval = "prediction")
print(forecasts)
```

```
##          fit          lwr          upr
## 1  14.946060 12.0910334 17.801086
## 2  18.577035 15.7217116 21.432358
## 3  20.664983 17.8094369 23.520529
## 4  20.598113 17.7424564 23.453770
## 5  18.422072 15.5664148 21.277729
## 6  14.656961 11.8013863 17.512536
## 7  10.383300  7.5278354 13.238764
## 8   6.692016  3.8366324  9.547400
## 9   4.543033  1.6876495  7.398417
## 10  4.549595  1.6941007  7.405089
## 11  6.665328  3.8096123  9.521043
## 12 10.369404  7.5133780 13.225429
## 13 14.582757 11.7263900 17.439123
## 14 18.213732 15.3570505 21.070414
## 15 20.301680 17.4447628 23.158597
## 16 20.234810 17.3777759 23.091845
## 17 18.058769 15.2017343 20.915803
## 18 14.293658 11.4367106 17.150606
## 19 10.019997  7.1631661 12.876828
## 20  6.328713  3.4719678  9.185458
## 21  4.179730  1.3229849  7.036476
## 22  4.186292  1.3294297  7.043154
## 23  6.302025  3.4449284  9.159121
## 24 10.006101  7.1486763 12.863525
## 25 14.219454 11.3616690 17.077238
## 26 17.850429 14.9923119 20.708546
## 27 19.938377 17.0800112 22.796743
## 28 19.871507 17.0130178 22.729997
## 29 17.695466 14.8369763 20.553956
## 30 13.930355 11.0719573 16.788753
## 31  9.656694  6.7984193 12.514969
## 32  5.965410  3.1072257  8.823595
## 33  3.816427  0.9582427  6.674612
## 34  3.822989  0.9646811  6.681297
## 35  5.938722  3.0801669  8.797277
## 36  9.642798  6.7838971 12.501698
## 37 13.856151 10.9968705 16.715431
```

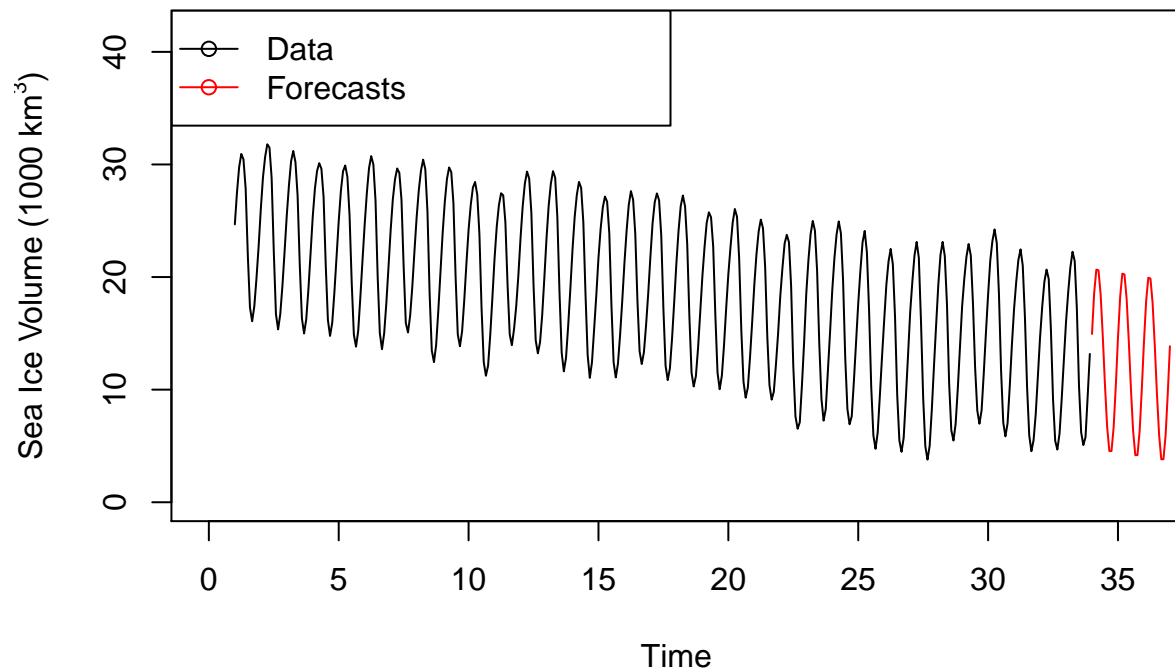
## 2.8 Plot of 3-year forecast of Arctic sea ice.

```
plot(sea_ice_ts2, xlim = c(0,36), ylim = c(0, 42), ylab = expression(paste("Sea Ice Volume (1000 km"3
# Here we convert the forecasts and prediction limits to monthly time series!
lines(ts(as.vector(forecasts[,1]), start = c(34,1), frequency = 12), col="red", type="l")
lines(ts(as.vector(forecasts[,2]), start = c(34,1), frequency = 12), col="blue", type="l")
lines(ts(as.vector(forecasts[,3]), start = c(34,1), frequency = 12), col="blue", type="l")
legend("topleft", lty=1, pch=1, col=c("black","blue","red"), text.width = 15,
      c("Data", "5% forecast limits", "Forecasts"))
```





```
plot(sea_ice_ts2, xlim = c(0,36), ylim = c(0, 42), ylab = expression(paste("Sea Ice Volume (1000 km"^(3)
# Here we convert the forecasts and prediction limits to monthly time series!
lines(ts(as.vector(forecasts[,1]), start = c(34,1), frequency = 12), col="red", type="l")
legend("topleft", lty=1, pch=1, col=c("black","red"), text.width = 15,
      c("Data","Forecasts"))
```



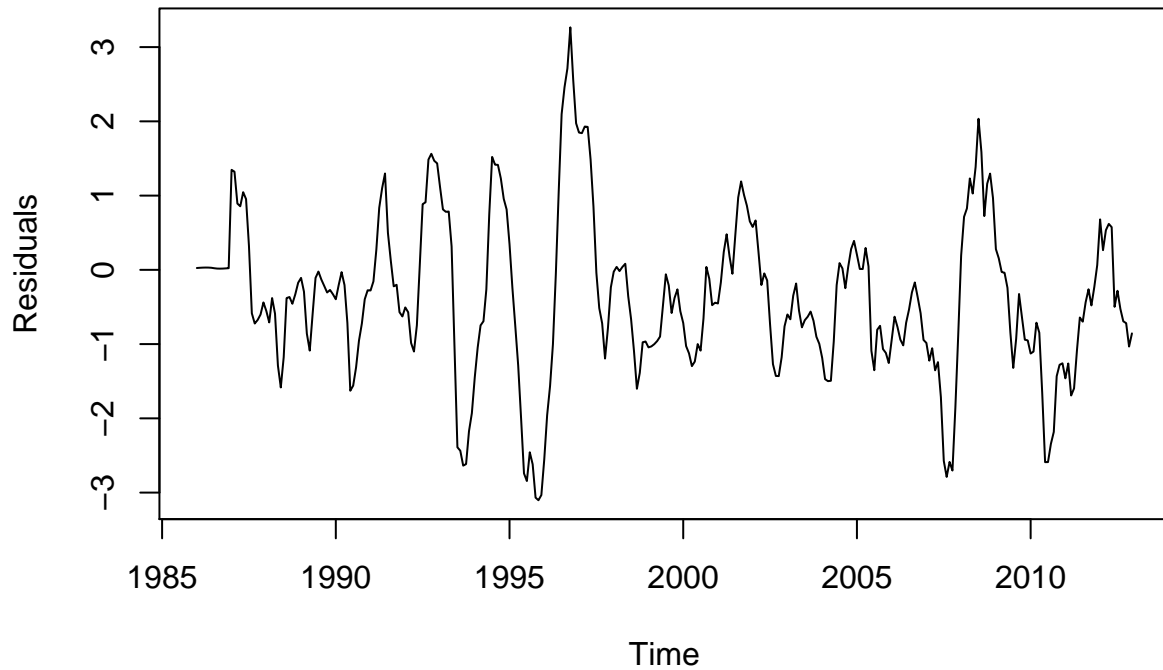
### 3 Nonstationary seasonal ARIMA (SARIMA) model (Residuals Approach).

#### 3.1 Fit of SARIMA models.

##### 3.1.1 Initial fit of SARIMA(0,0,0) $\times$ (0,1,0)<sub>12</sub> model

```
m100_sea_ice = arima(sea_ice_ts,order=c(0,0,0),seasonal=list(order=c(0,1,0), period=12))
res_m100 = residuals(m100_sea_ice);
plot(res_m100,xlab='Time', ylab='Residuals',main="Time series plot of the residuals")
```

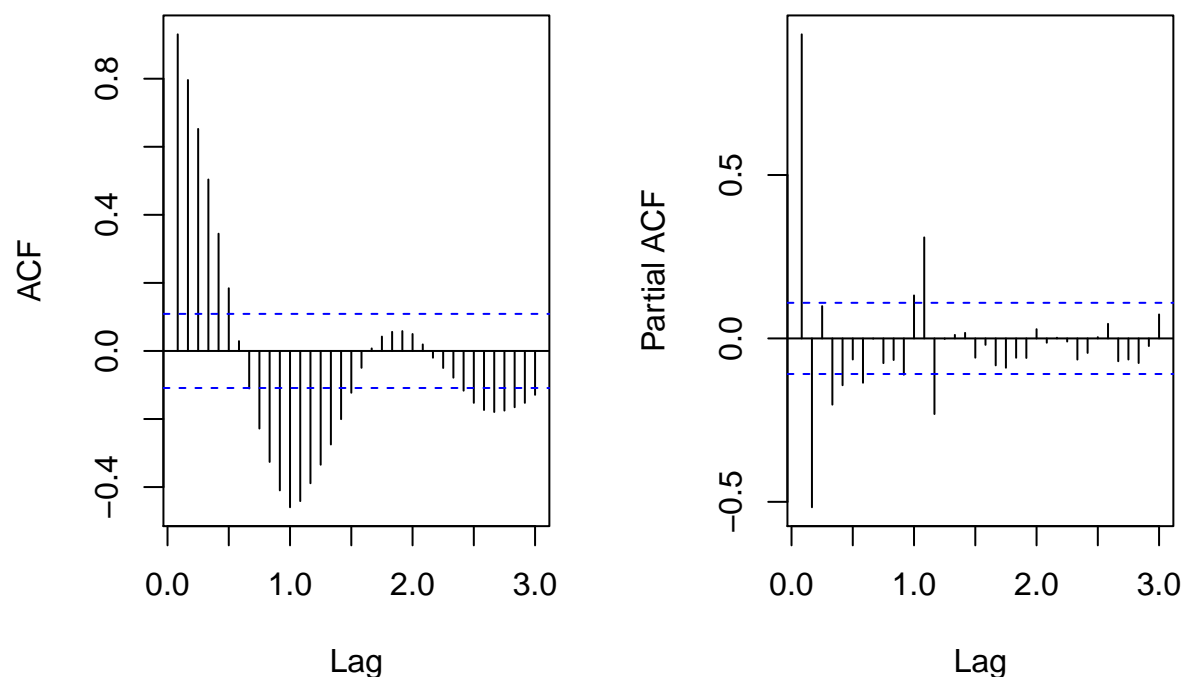
### Time series plot of the residuals



#### 3.1.2 ACF and PACF plots of residuals for SARIMA(0,0,0)x(0,1,0)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m100, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m100, lag.max = 36, main = "The sample PACF of the residuals")
```

**The sample ACF of the residual:**      **The sample PACF of the residual**

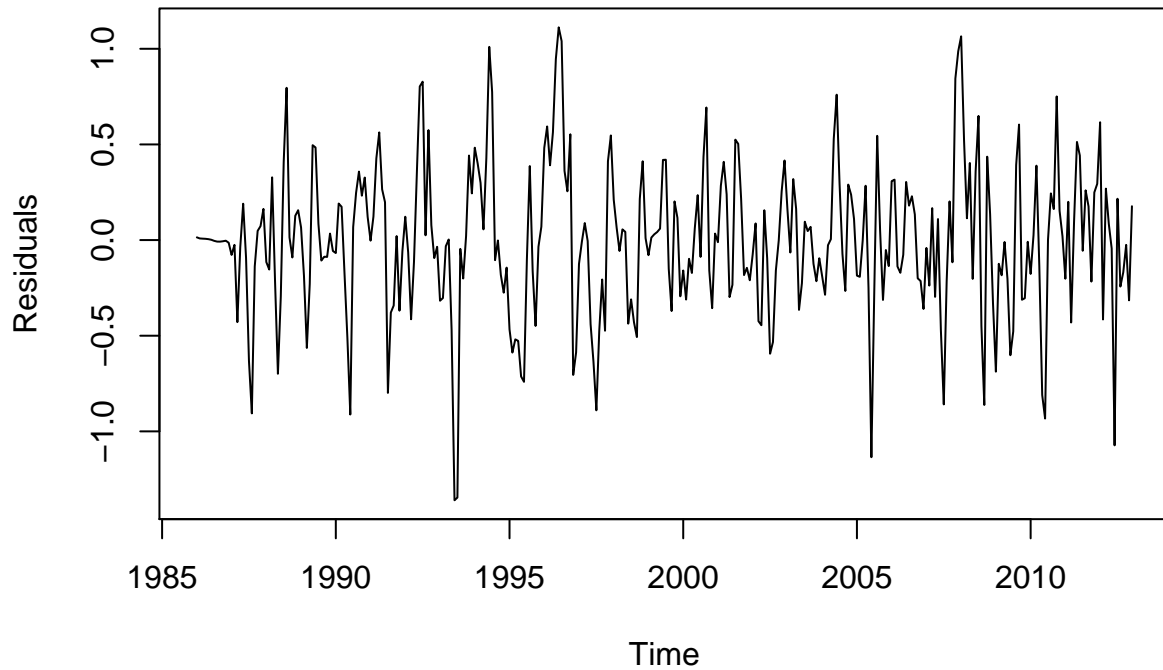


- There is now no pattern implying the existence of a seasonal trend.
- However, the slowly decaying pattern prior to the first period at 1s implies the existence of an ordinary trend in the ACF plot of the residuals.
- We will attempt to remove this ordinary trend by fitting a  $\text{SARIMA}(0,1,0) \times (0,1,0)_{12}$  model.

### 3.1.3 Fit of $\text{SARIMA}(0,1,0) \times (0,1,0)_{12}$ model

```
m200_sea_ice = arima(sea_ice_ts,order=c(0,1,0),seasonal=list(order=c(0,1,0), period=12))
res_m200 = residuals(m200_sea_ice);
plot(res_m200,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

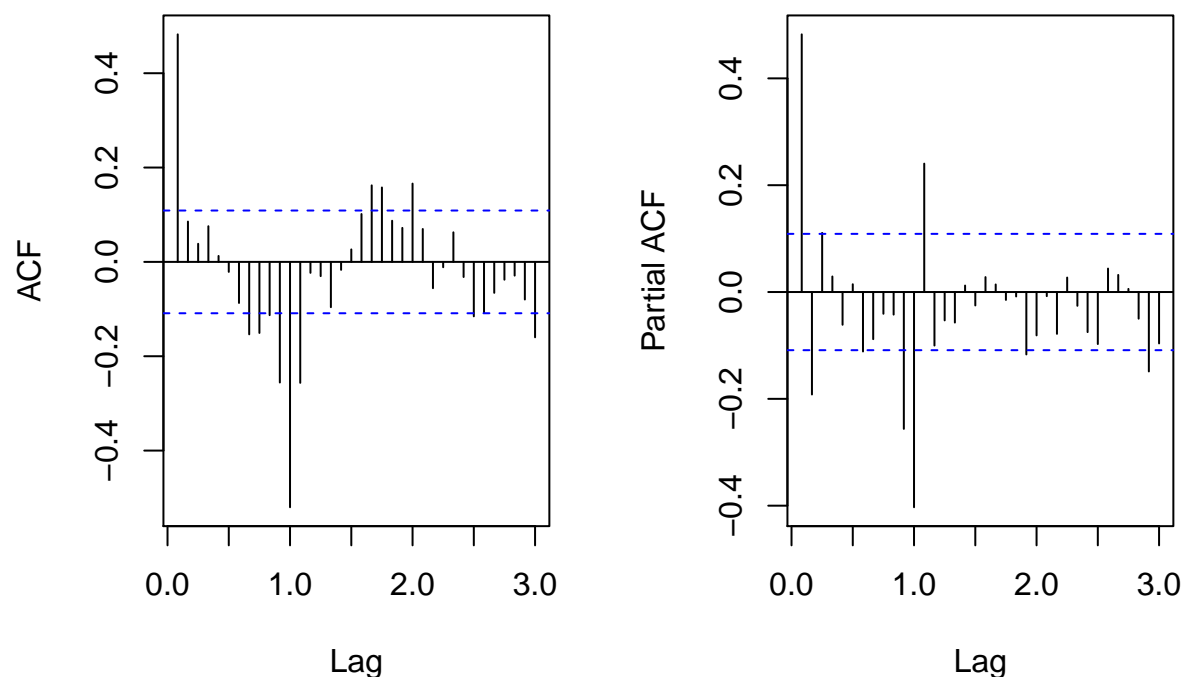
### Time series plot of the residuals



#### 3.1.4 ACF and PACF plots for SARIMA(0,1,0)x(0,1,0)<sub>12</sub> model.

```
par(mfrow=c(1,2))
acf(res_m200, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m200, lag.max = 36, main = "The sample PACF of the residuals")
```

## The sample ACF of the residual:      The sample PACF of the residual

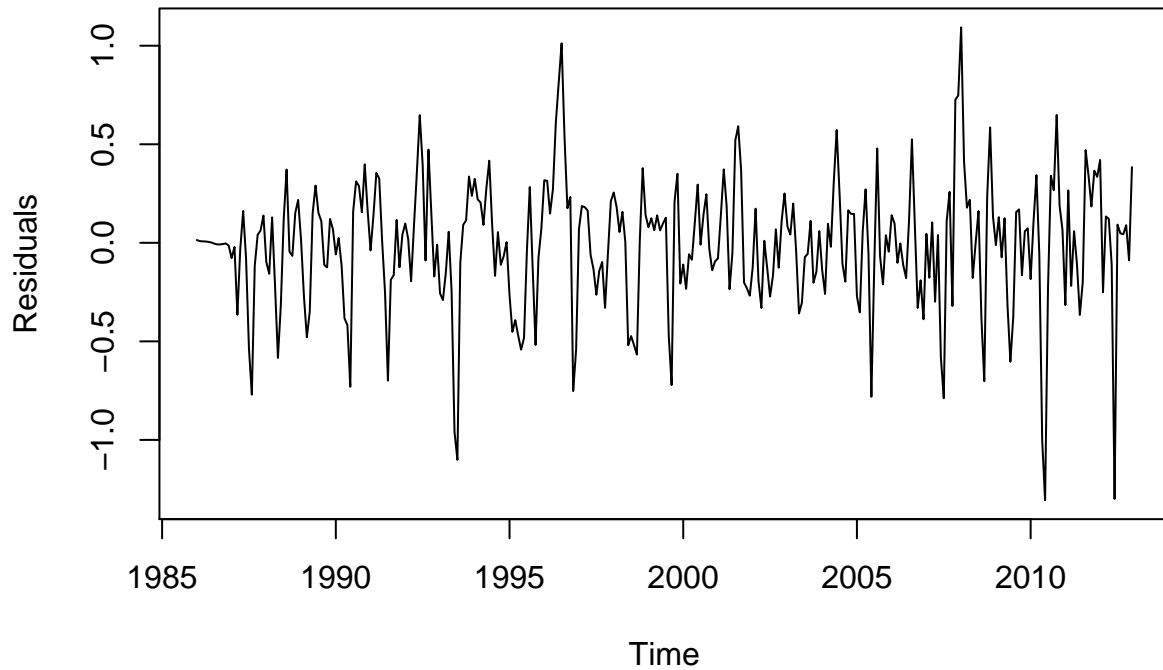


- No evidence of an ordinary trend remaining in the residuals.
- There is a decreasing pattern in lags 1, 2, 3, ... in the SARMA component of the PACF plot. The correlation at lag 1 in the ACF plot is significant. This implies the existence of an SMA(1) component.
- ACF - 1 significant lag at 1s and 1 in first part prior to 1s ( $q=1$ ,  $Q=1$ )
- PACF - 1 significant lag and 1 not so significant lag in first part and 1 significant after 1s ( $p=1,2$ ,  $P=1$ )
- Now tried to fit a  $SARIMA(0,1,0) \times (0,1,1)_{12}$  model to try to remove the remaining seasonal component in the residuals.

### 3.1.5 Fit of $SARIMA(0,1,0) \times (0,1,1)_{12}$ model.

```
m300_sea_ice = arima(sea_ice_ts,order=c(0,1,0),seasonal=list(order=c(0,1,1), period=12))
res_m300 = residuals(m300_sea_ice);
plot(res_m300,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

### Time series plot of the residuals

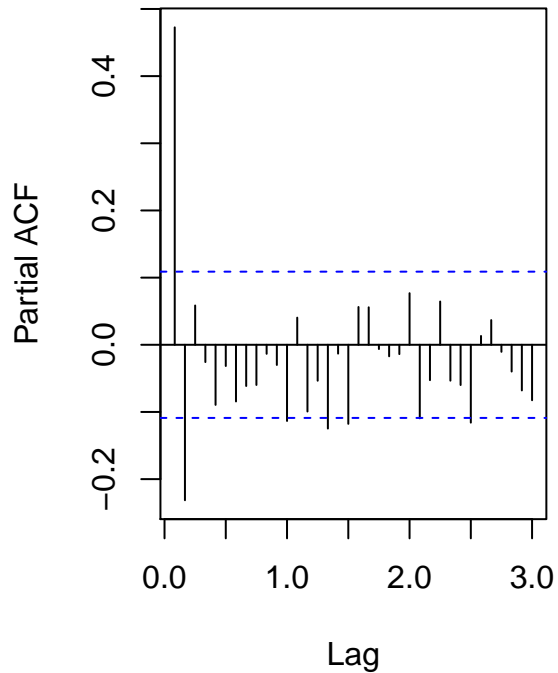
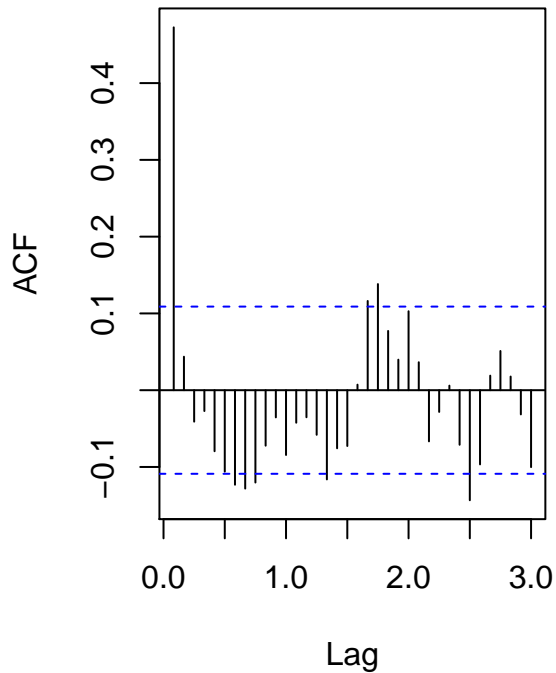


#### 3.1.6 ACF AND PACF plots for SARIMA(0,1,0)x(0,1,1)<sub>12</sub> model.

```
par(mfrow=c(1,2))
acf(res_m300, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m300, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:

The sample PACF of the residual



- The autocorrelations, especially the first seasonal lag (lag 1) in the ACF plot, become insignificant after adding the seasonal component.
- The ACF and PACF plots can be used to determine the orders of the ARMA component since there are no highly significant correlations at lags  $s$ ,  $2s$ ,  $3s$ , ...
- The ACF plot displays one significant and 3 less significant autocorrelation ( $q=1,2,3$ ) while the PACF plot has two significant autocorrelations ( $p=1,2$ ). This suggests a  $\text{ARMA}(1,1)$ ,  $\text{ARMA}(1,2)$ ,  $\text{ARMA}(2,1)$ ,  $\text{ARMA}(1,3)$  and  $\text{ARMA}(2,3)$  models.
- We will now fit  $\text{SARIMA}(1,1,1)\times(0,1,1)_{12}$ ,  $\text{SARIMA}(1,1,2)\times(0,1,1)_{12}$  models,  $\text{SARIMA}(2,1,1)\times(0,1,1)_{12}$ ,  $\text{SARIMA}(1,1,3)\times(0,1,1)_{12}$  and  $\text{SARIMA}(2,1,3)\times(0,1,1)_{12}$  models.

### 3.1.6.1 EACF analysis of the residuals.

```
eacf(res_m300)
```

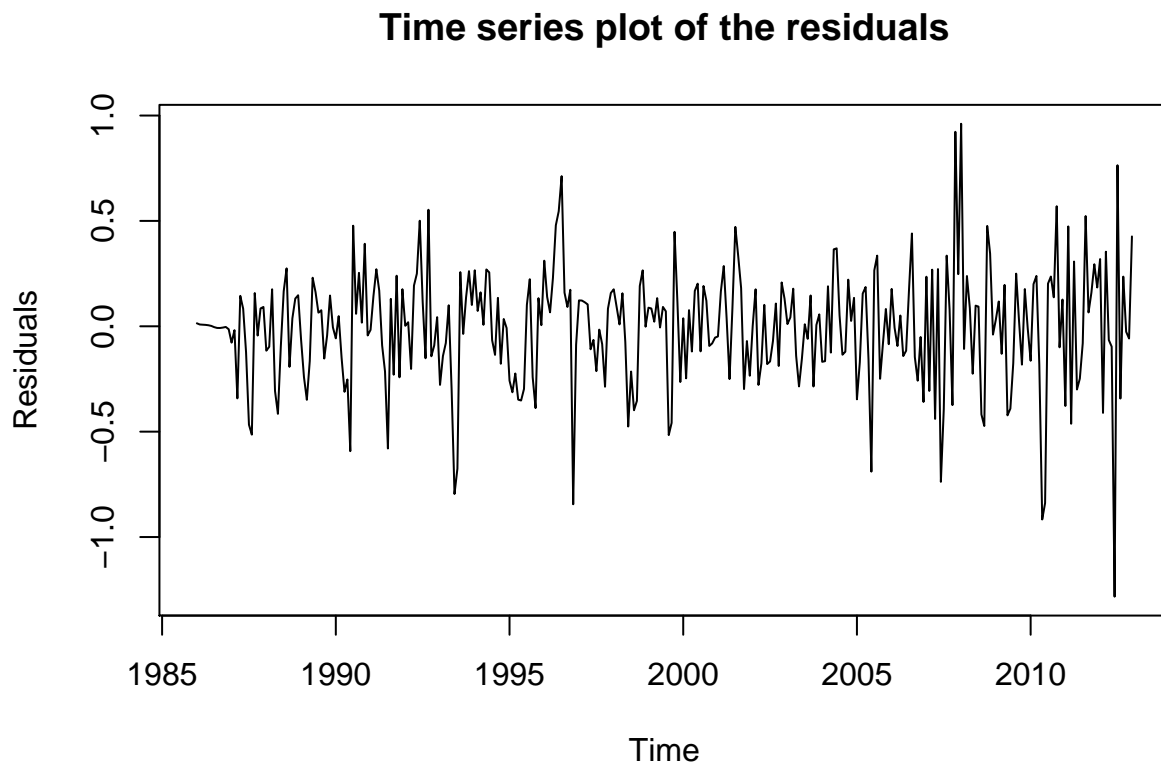
```
## AR/MA
##   0 1 2 3 4 5 6 7 8 9 10 11 12 13
## 0 x o o o o o x x x o o o o o
## 1 x x o o o o o o o o o o o o
## 2 x o x o o o o o o o o o o o
## 3 x o x x o o o o o o o o o o
## 4 x x x x o o o o o o o o o o
## 5 x x o x o o o o o o o o o o
## 6 x x o x o o o o o o o o o o
## 7 x x x x o o o o o o o o o o
```



- The tentative models are specified as
- SARIMA(1,1,1)x(0,1,1)<sub>12</sub>
- SARIMA(1,1,2)x(0,1,1)<sub>12</sub>
- SARIMA(1,1,3)x(0,1,1)<sub>12</sub>
- SARIMA(2,1,2)x(0,1,1)<sub>12</sub>
- From the EACF, we will include
- SARIMA(0,1,1)x(0,1,1)<sub>12</sub>
- SARIMA(0,1,2)x(0,1,1)<sub>12</sub>

### 3.1.7 Fit of SARIMA(0,1,1)x(0,1,1)<sub>12</sub> model

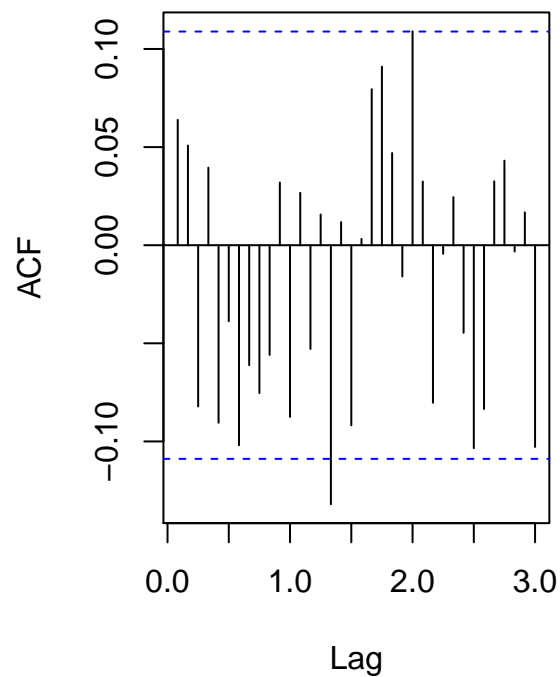
```
m1_sea_ice = arima(sea_ice_ts,order=c(0,1,1),seasonal=list(order=c(0,1,1), period=12))
res_m1 = residuals(m1_sea_ice);
plot(res_m1,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```



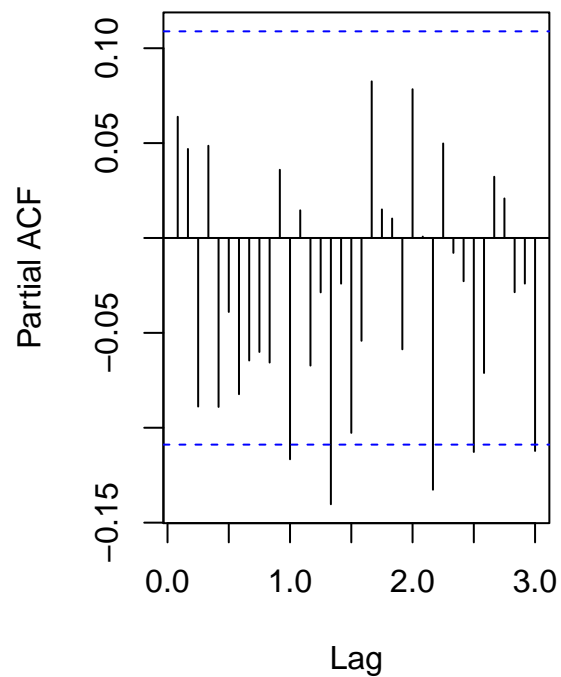
### 3.1.8 ACF and PACF plots for SARIMA(0,1,1)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m1, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m1, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:



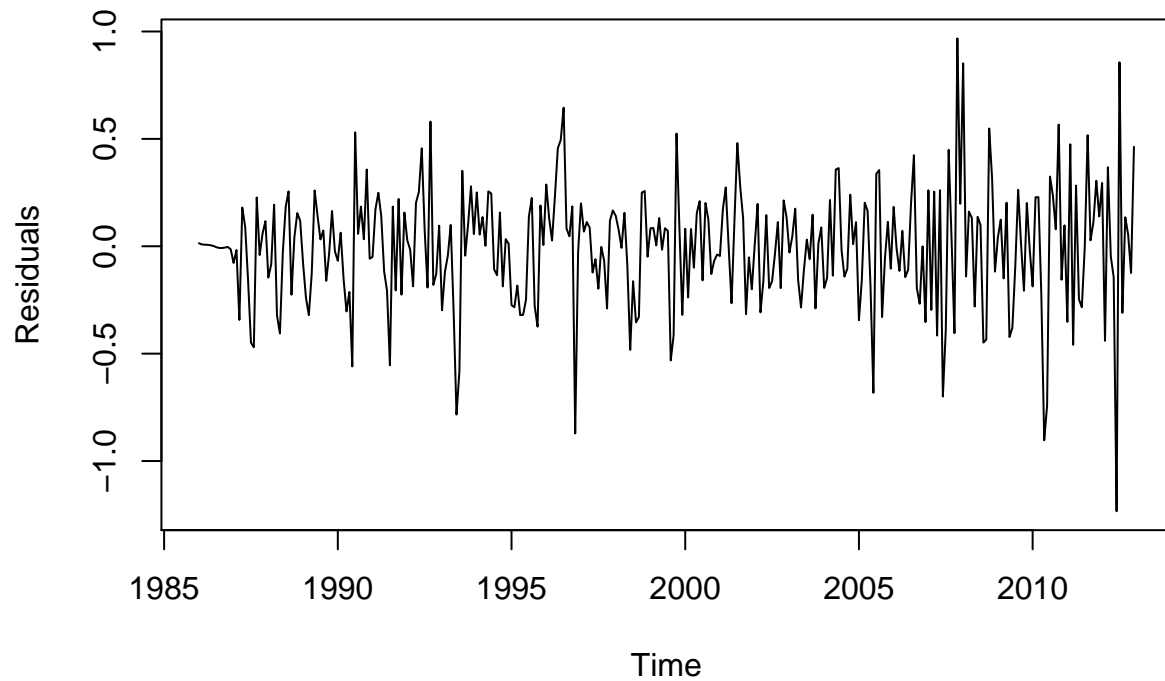
The sample PACF of the residual



### 3.1.9 Fit of SARIMA(0,1,2) $\times$ (0,1,1)<sub>12</sub> model

```
m2_sea_ice = arima(sea_ice_ts,order=c(0,1,2),seasonal=list(order=c(0,1,1), period=12))
res_m2 = residuals(m2_sea_ice);
plot(res_m2,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

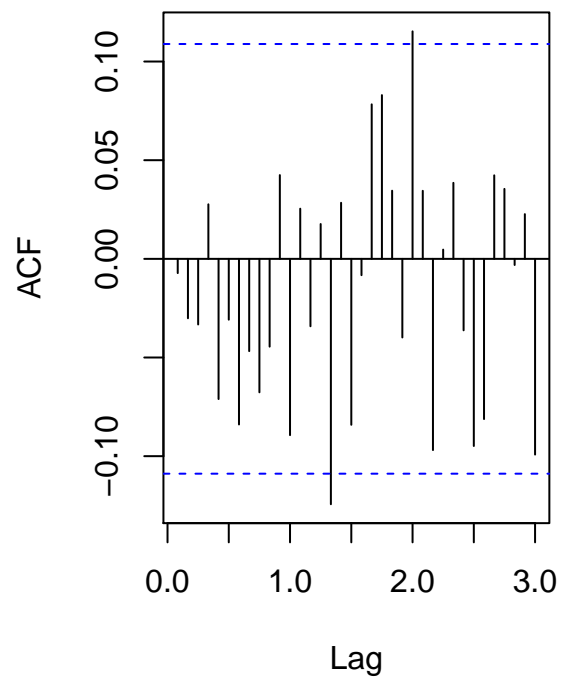
### Time series plot of the residuals



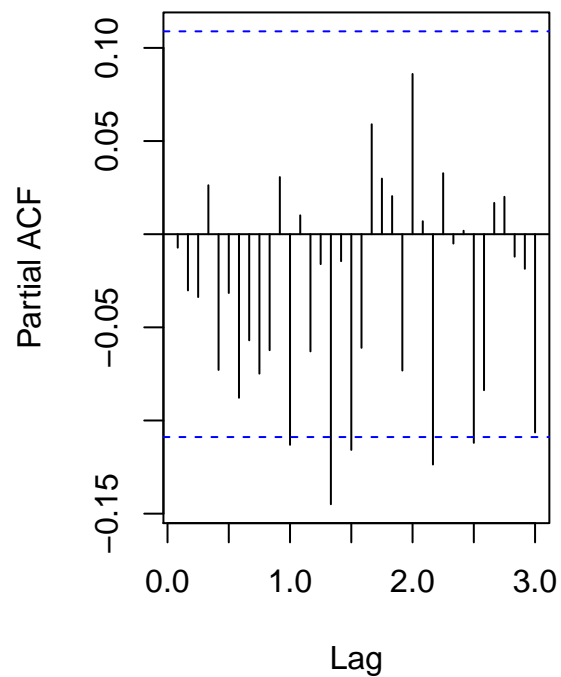
#### 3.1.10 ACF and PACF plots for SARIMA(0,1,2)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m2, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m2, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:



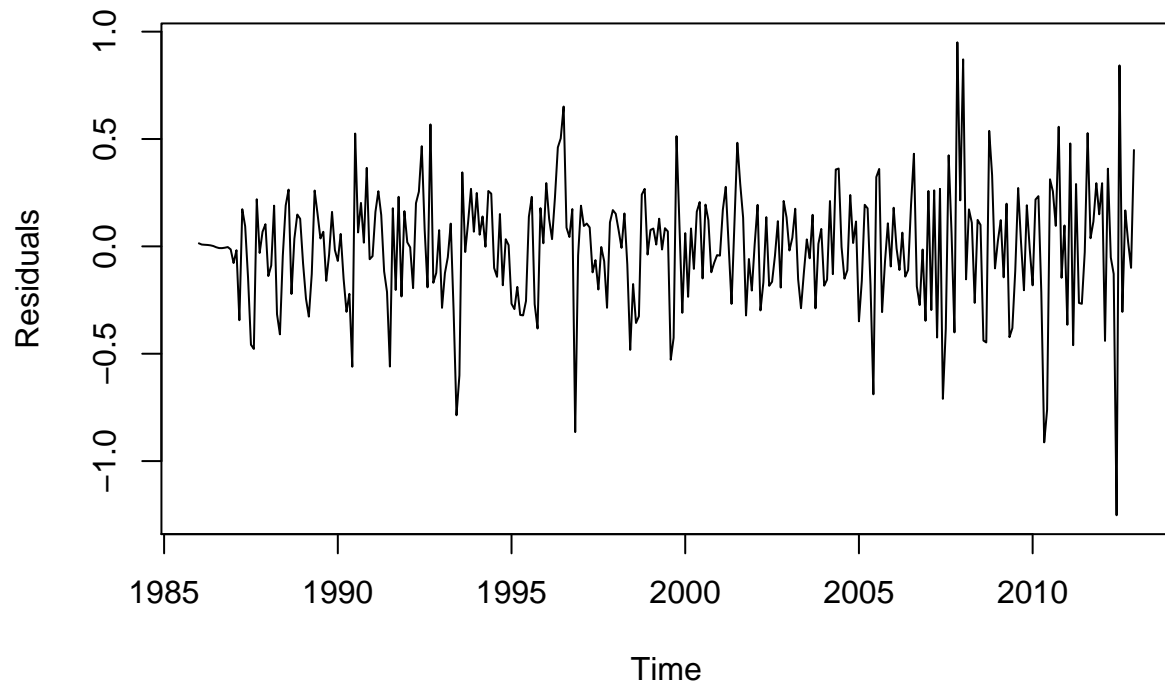
The sample PACF of the residual



### 3.1.11 Fit of SARIMA(1,1,1)x(0,1,1)<sub>12</sub> model

```
m3_sea_ice = arima(sea_ice_ts,order=c(1,1,1),seasonal=list(order=c(0,1,1), period=12))
res_m3 = residuals(m3_sea_ice);
plot(res_m3,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

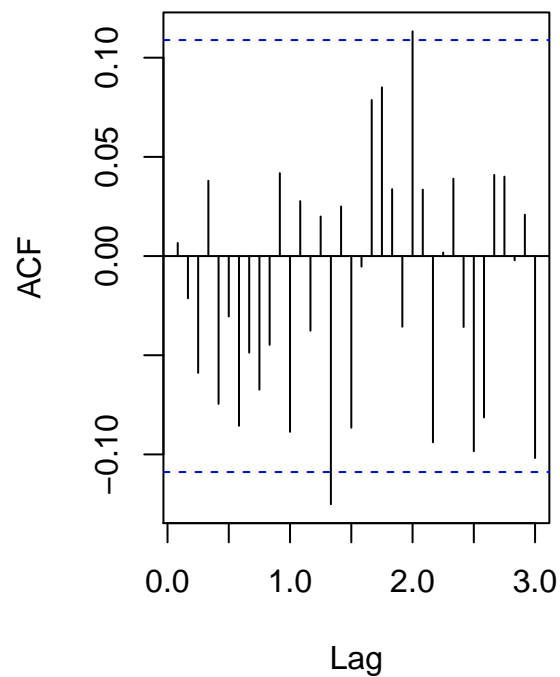
### Time series plot of the residuals



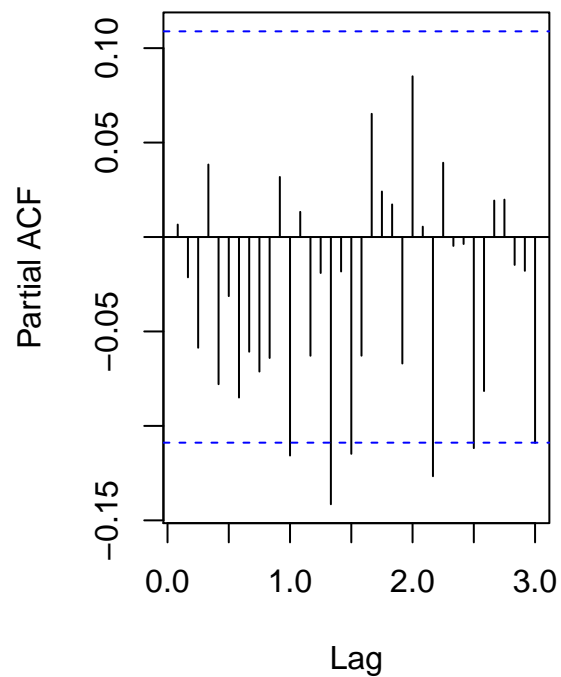
#### 3.1.12 ACF and PACF plots for SARIMA(1,1,1)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m3, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m3, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:



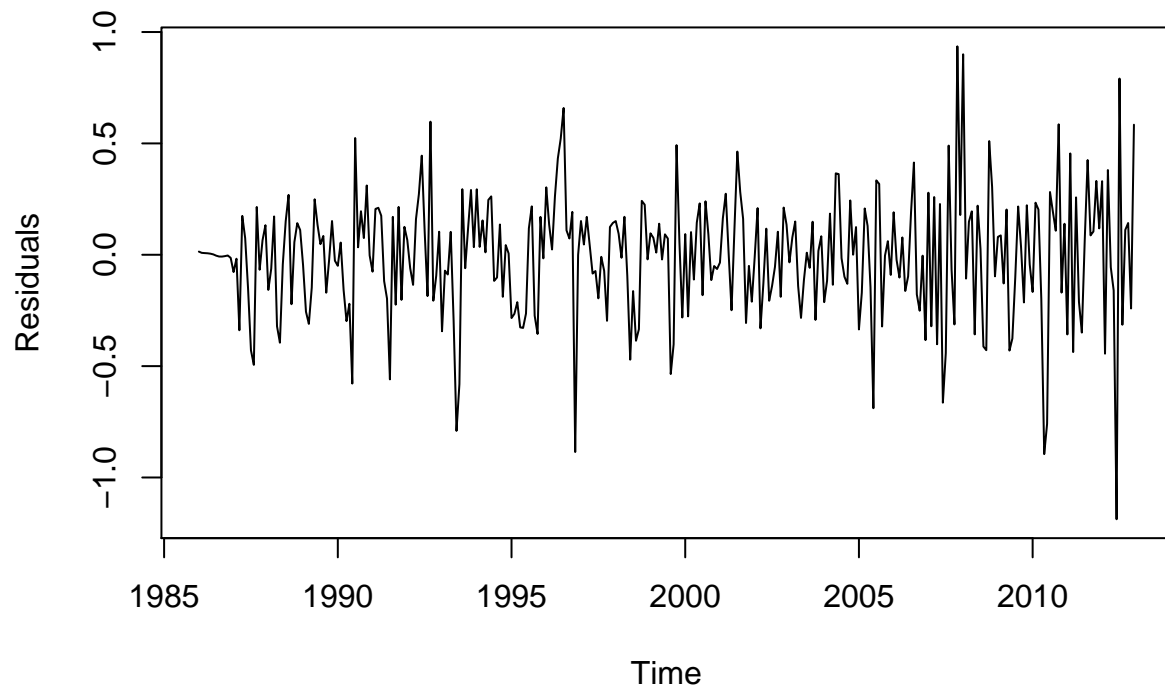
The sample PACF of the residual



### 3.1.13 Fit of SARIMA(1,1,2)x(0,1,1)<sub>12</sub> model

```
m4_sea_ice = arima(sea_ice_ts,order=c(1,1,2),seasonal=list(order=c(0,1,1), period=12))
res_m4 = residuals(m4_sea_ice);
plot(res_m4,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

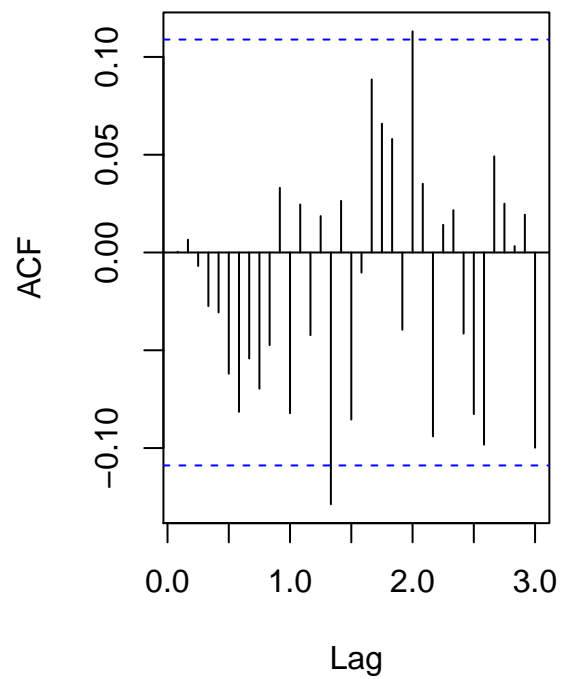
### Time series plot of the residuals



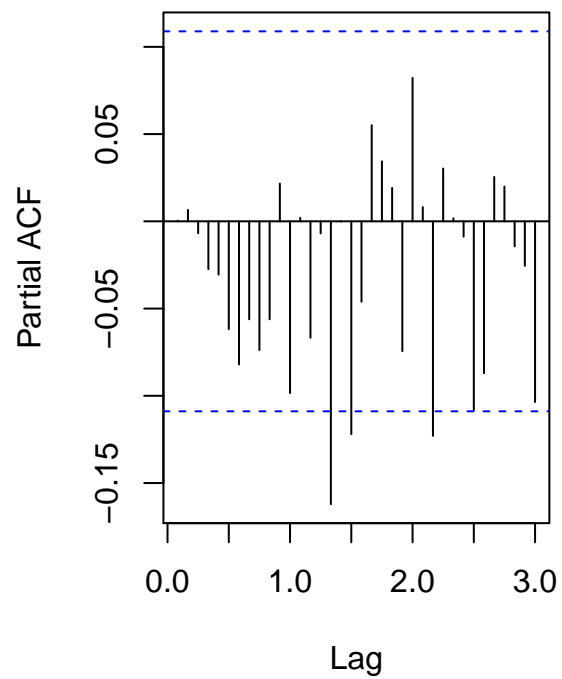
#### 3.1.14 ACF and PACF plots for SARIMA(1,1,2)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m4, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m4, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:



The sample PACF of the residual

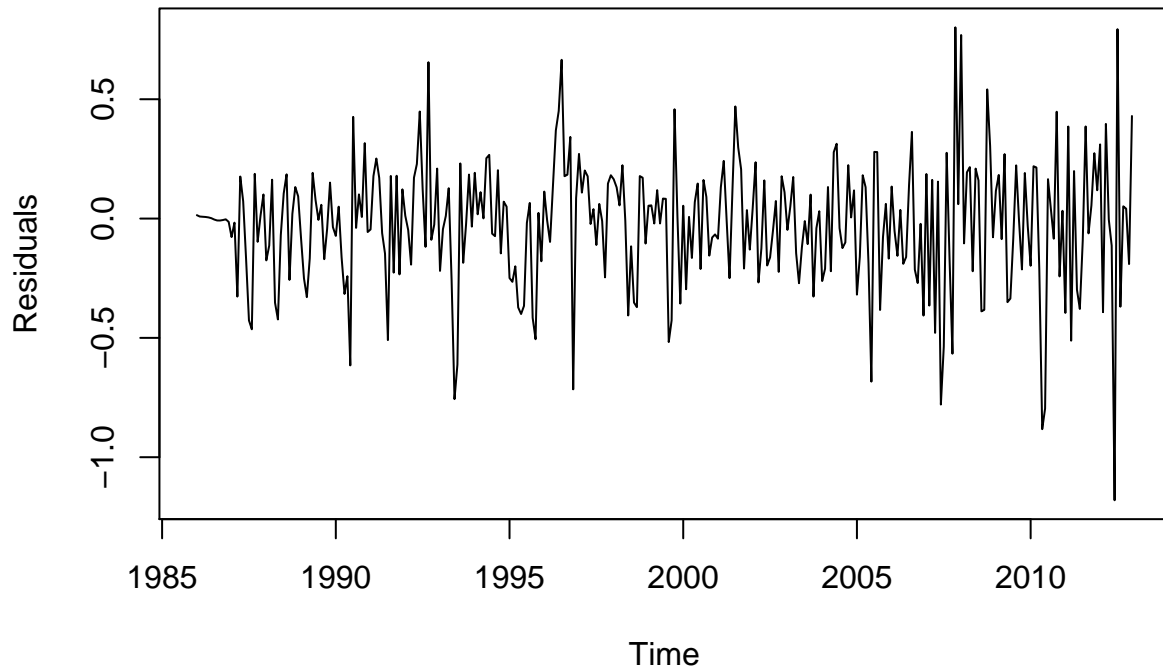


### 3.1.15 Fit of SARIMA(1,1,3)x(0,1,1)<sub>12</sub> model

```
m5_sea_ice = arima(sea_ice_ts,order=c(1,1,3),seasonal=list(order=c(0,1,1), period=12))
res_m5 = residuals(m5_sea_ice);
plot(res_m5,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```



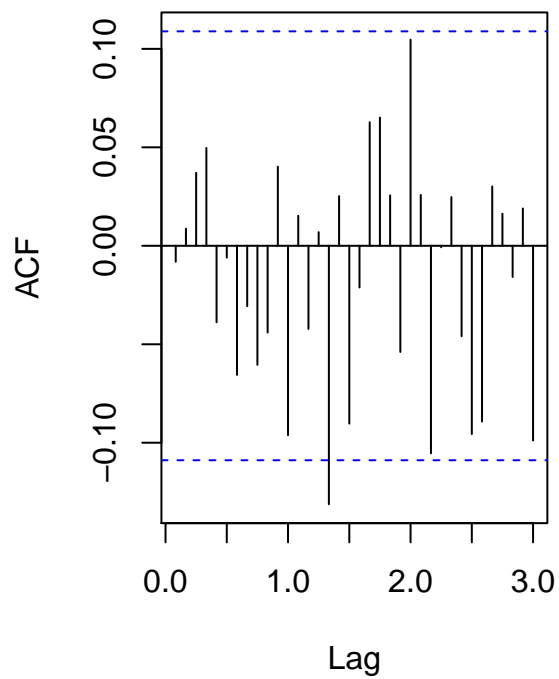
### Time series plot of the residuals



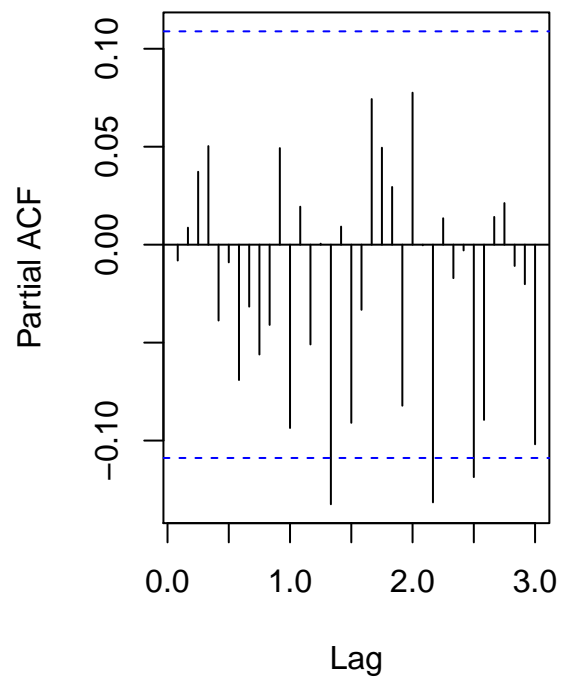
#### 3.1.16 ACF and PACF plots for SARIMA(1,1,3)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m5, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m5, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:



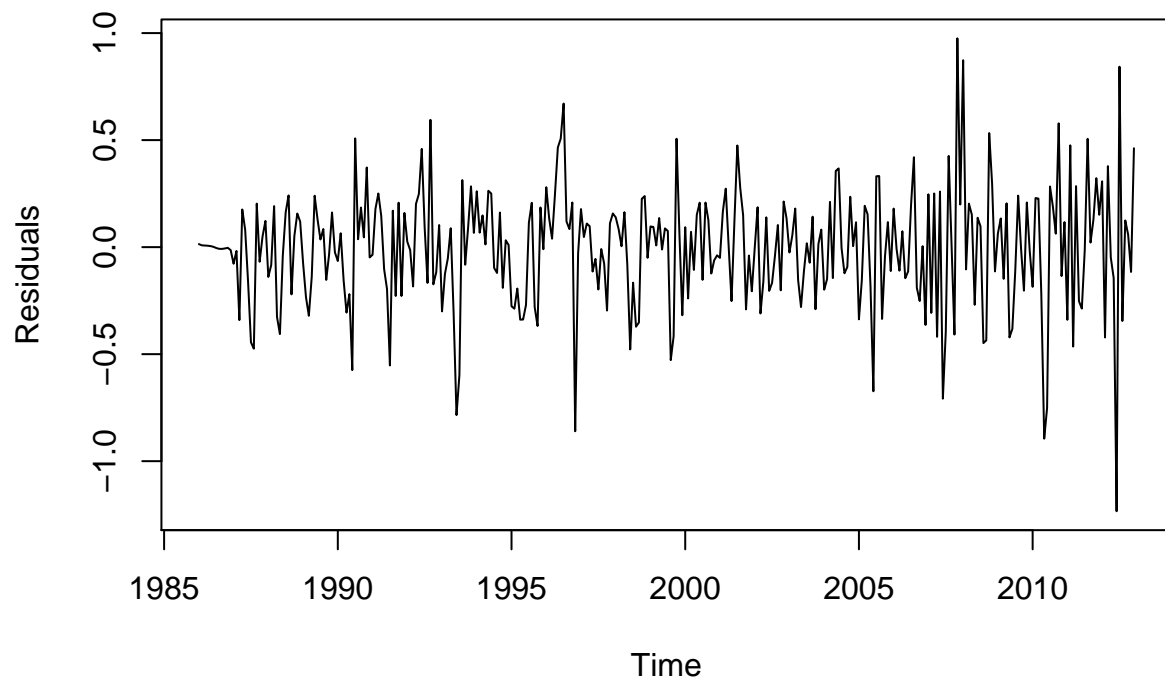
The sample PACF of the residual



### 3.1.17 Fit of SARIMA(2,1,1)x(0,1,1)<sub>12</sub> model

```
m6_sea_ice = arima(sea_ice_ts,order=c(2,1,1),seasonal=list(order=c(0,1,1), period=12))
res_m6 = residuals(m6_sea_ice);
plot(res_m6,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

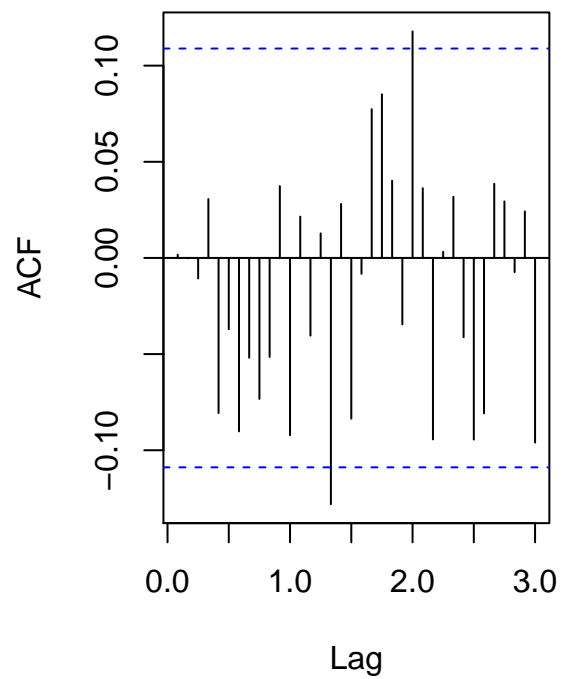
### Time series plot of the residuals



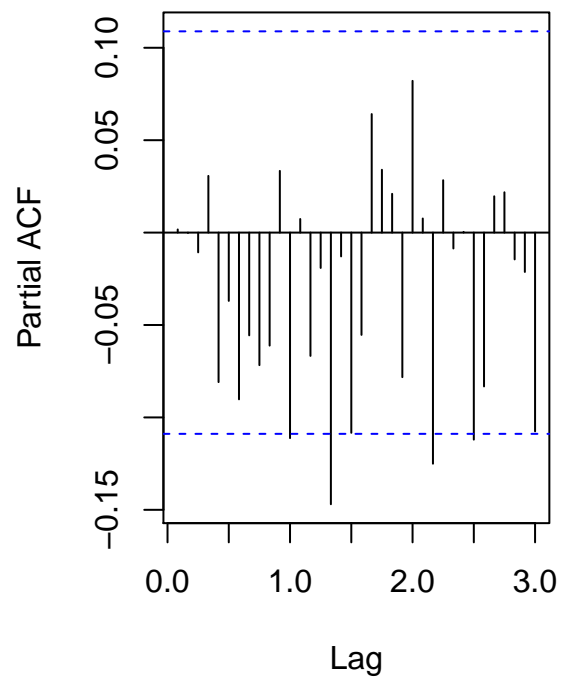
#### 3.1.18 ACF and PACF plots for SARIMA(2,1,1)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m6, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m6, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:



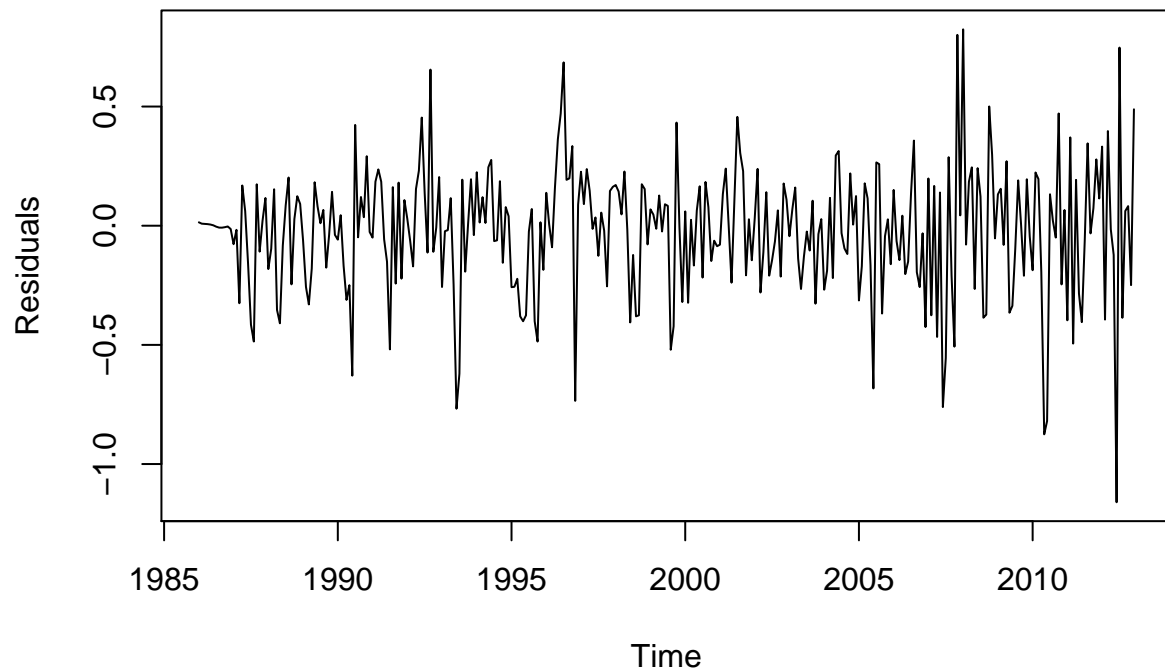
The sample PACF of the residual



### 3.1.19 Fit of SARIMA(2,1,3)x(0,1,1)<sub>12</sub> model

```
m7_sea_ice = arima(sea_ice_ts,order=c(2,1,3),seasonal=list(order=c(0,1,1), period=12))
res_m7 = residuals(m7_sea_ice);
plot(res_m7,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

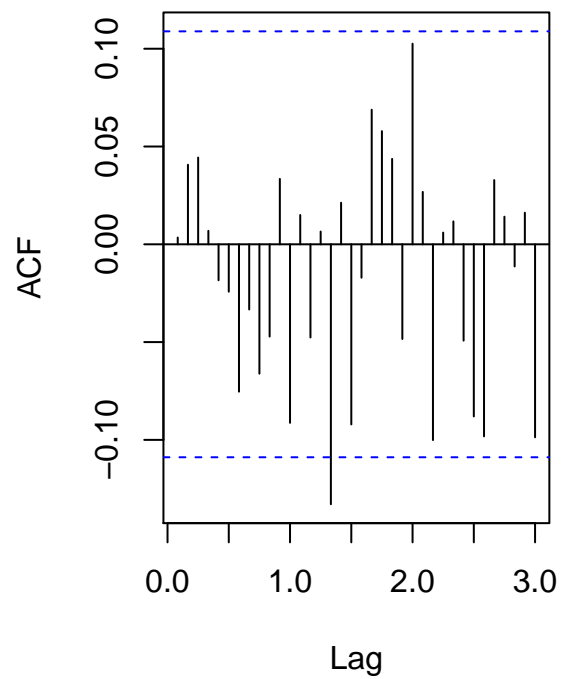
### Time series plot of the residuals



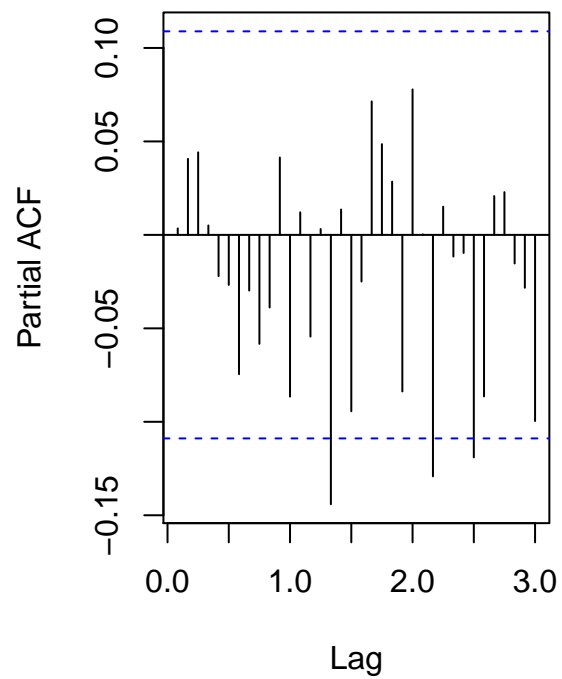
#### 3.1.20 ACF and PACF plots for SARIMA(2,1,3)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m7, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m7, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:



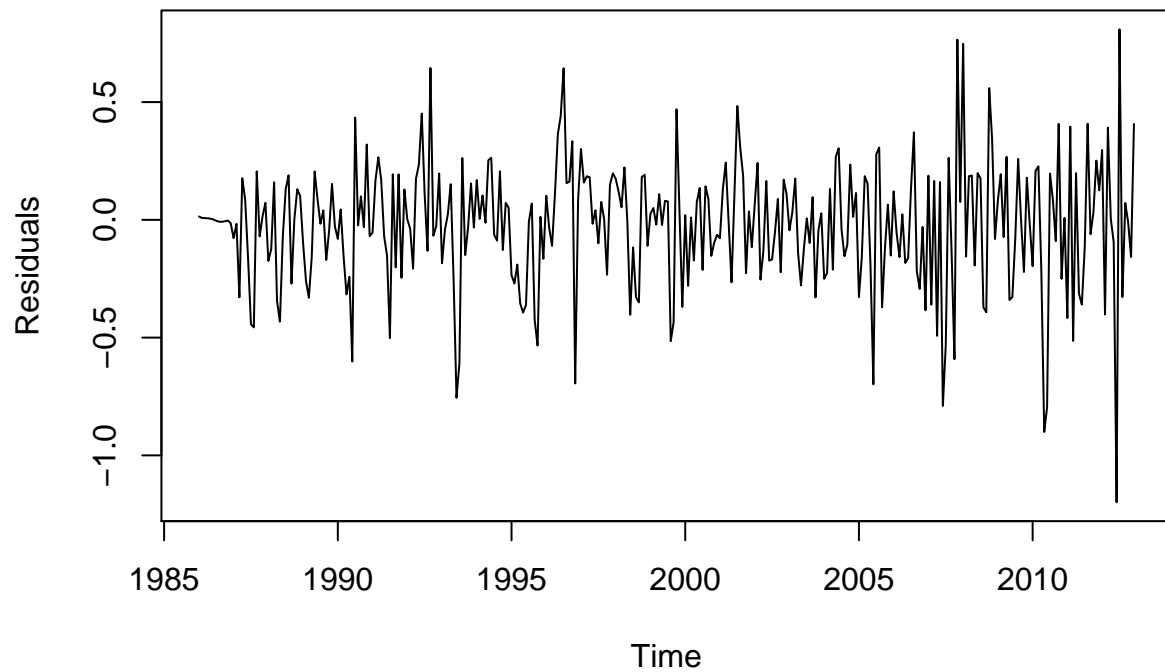
The sample PACF of the residual



### 3.1.21 Fit of SARIMA(2,1,2)x(0,1,1)<sub>12</sub> model

```
m8_sea_ice = arima(sea_ice_ts,order=c(2,1,2),seasonal=list(order=c(0,1,1), period=12))
res_m8 = residuals(m8_sea_ice);
plot(res_m8,xlab='Time',ylab='Residuals',main="Time series plot of the residuals")
```

### Time series plot of the residuals

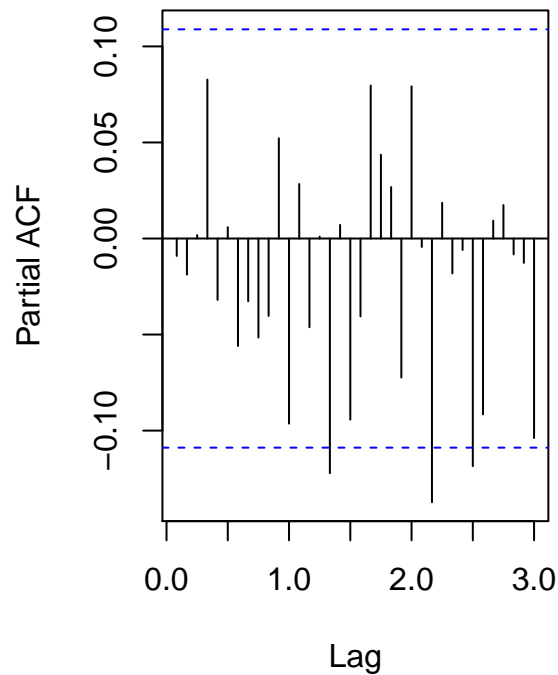
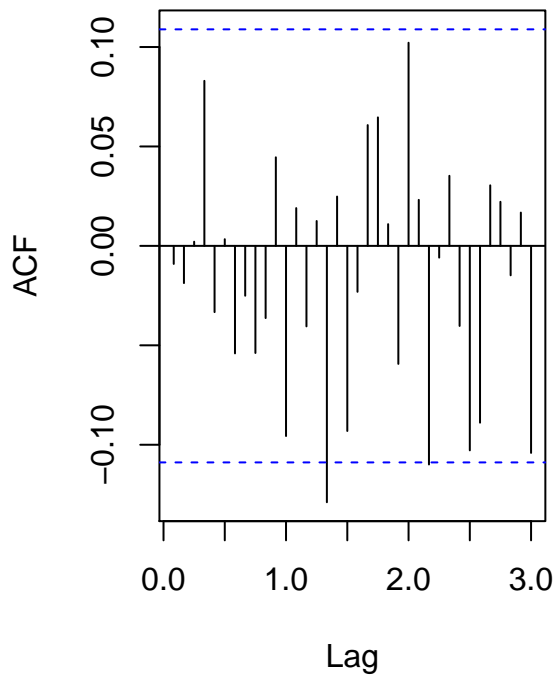


#### 3.1.22 ACF and PACF plots for SARIMA(2,1,2)x(0,1,1)<sub>12</sub> model

```
par(mfrow=c(1,2))
acf(res_m8, lag.max = 36, main = "The sample ACF of the residuals")
pacf(res_m8, lag.max = 36, main = "The sample PACF of the residuals")
```

The sample ACF of the residual:

The sample PACF of the residual



- The residuals for the  $\text{SARIMA}(1,1,2) \times (0,1,1)_{12}$  model are closer to white noise. However, there are a number of significant autocorrelations in the ACF and PACF plots of the residuals which will be considered in more detail latter.
- Therefore we can conclude that the orders are:  $p=1$ ,  $d=1$ ,  $q=2$ ,  $P=0$ ,  $D=1$ ,  $Q=1$  and  $s=12$  for the  $\text{SARIMA}(p,d,q) \times (P,D,Q)_s$  model.

### 3.2 Model Fitting - ML estimates and Conditional Least Squares for SARIMA models.

#### 3.2.1 ML estimates for $\text{SARIMA}(0,1,1) \times (0,1,1)_{12}$ model

```
m1_sea_ice_ts = arima(sea_ice_ts, order=c(0,1,1), seasonal=list(order=c(0,1,1), period=12))
coeftest(m1_sea_ice_ts)
```

```
##
## z test of coefficients:
##
##      Estimate Std. Error z value Pr(>|z|)
## ma1    0.533392   0.041952  12.714 < 2.2e-16 ***
## sma1 -0.640394   0.044802 -14.294 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```



### 3.2.2 ML estimates for SARIMA(0,1,2)x(0,1,1)<sub>12</sub> model

```
m2_sea_ice_ts = arima(sea_ice_ts,order=c(0,1,2), seasonal=list(order=c(0,1,1), period=12))
coeftest(m2_sea_ice_ts)

##
## z test of coefficients:
##
##      Estimate Std. Error  z value Pr(>|z|)
## ma1    0.613225   0.059311  10.3392 < 2e-16 ***
## ma2    0.124154   0.056944   2.1803 0.02924 *
## sma1 -0.636502   0.045190 -14.0851 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### 3.2.3 ML estimates for SARIMA(1,1,1)x(0,1,1)<sub>12</sub> model

```
m3_sea_ice_ts = arima(sea_ice_ts,order=c(1,1,1), seasonal=list(order=c(0,1,1), period=12))
coeftest(m3_sea_ice_ts)

##
## z test of coefficients:
##
##      Estimate Std. Error  z value  Pr(>|z|)
## ar1    0.178782   0.090896   1.9669   0.0492 *
## ma1    0.417993   0.080069   5.2204 1.786e-07 ***
## sma1 -0.636858   0.045063 -14.1327 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### 3.2.4 ML estimates for SARIMA(1,1,2)x(0,1,1)<sub>12</sub> model

```
m4_sea_ice_ts = arima(sea_ice_ts,order=c(1,1,2), seasonal=list(order=c(0,1,1), period=12))
coeftest(m4_sea_ice_ts)

##
## z test of coefficients:
##
##      Estimate Std. Error  z value  Pr(>|z|)
## ar1   -0.647143   0.197411  -3.2782 0.001045 **
## ma1    1.260159   0.195445   6.4476 1.136e-10 ***
## ma2    0.477516   0.110513   4.3209 1.554e-05 ***
## sma1 -0.639266   0.046287 -13.8109 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### 3.2.5 ML estimates for SARIMA(1,1,3)x(0,1,1)<sub>12</sub> model

```
m5_sea_ice_ts = arima(sea_ice_ts,order=c(1,1,3), seasonal=list(order=c(0,1,1), period=12))
coeftest(m5_sea_ice_ts)
```

```
##
## z test of coefficients:
##
##      Estimate Std. Error  z value  Pr(>|z|)
## ar1    0.842707   0.040071  21.0305 < 2.2e-16 ***
## ma1   -0.297435   0.067382  -4.4142 1.014e-05 ***
## ma2   -0.487598   0.049078  -9.9352 < 2.2e-16 ***
## ma3   -0.190600   0.059194  -3.2199 0.001282 **
## sma1  -0.636035   0.045983 -13.8319 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### 3.2.6 ML estimates for SARIMA(2,1,1)<sub>x</sub>(0,1,1)<sub>12</sub> model

```
m6_sea_ice_ts = arima(sea_ice_ts,order=c(2,1,1), seasonal=list(order=c(0,1,1), period=12))
coeftest(m6_sea_ice_ts)
```

```
##
## z test of coefficients:
##
##      Estimate Std. Error  z value Pr(>|z|)
## ar1    0.356707   0.187181   1.9057 0.05669 .
## ar2   -0.129080   0.108771  -1.1867 0.23534
## ma1    0.246321   0.184888   1.3323 0.18277
## sma1  -0.638723   0.045102 -14.1616 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### 3.2.7 ML estimates for SARIMA(2,1,3)<sub>x</sub>(0,1,1)<sub>12</sub> model

```
m7_sea_ice_ts = arima(sea_ice_ts,order=c(2,1,3), seasonal=list(order=c(0,1,1), period=12))
coeftest(m7_sea_ice_ts)
```

```
##
## z test of coefficients:
##
##      Estimate Std. Error  z value Pr(>|z|)
## ar1    0.484978   2.111859   0.2296 0.8184
## ar2    0.323860   1.924049   0.1683 0.8663
## ma1    0.055727   2.062558   0.0270 0.9784
## ma2   -0.648039   0.972311  -0.6665 0.5051
## ma3   -0.378331   1.064944  -0.3553 0.7224
## sma1  -0.637609   0.047261 -13.4912 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

### 3.2.8 ML estimates for SARIMA(2,1,2)<sub>x</sub>(0,1,1)<sub>12</sub> model

```
m8_sea_ice_ts = arima(sea_ice_ts,order=c(2,1,2), seasonal=list(order=c(0,1,1), period=12))
coeftest(m8_sea_ice_ts)
```

```
##
## z test of coefficients:
##
##      Estimate Std. Error  z value  Pr(>|z|)
## ar1    1.177434   0.091309  12.8951 < 2.2e-16 ***
## ar2   -0.298959   0.089459  -3.3419 0.0008322 ***
## ma1   -0.638129   0.089610  -7.1212 1.07e-12 ***
## ma2   -0.342867   0.088469  -3.8756 0.0001064 ***
## sma1  -0.635259   0.045760 -13.8824 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

- The Maximum likelihood (ML) estimates were determined to be highly significant for SARIMA(0,1,1)x(0,1,1)<sub>12</sub>, SARIMA(0,1,2)x(0,1,1)<sub>12</sub>, SARIMA(1,1,1)x(0,1,1)<sub>12</sub>, SARIMA(1,1,2)x(0,1,1)<sub>12</sub> models, SARIMA(1,1,3)x(0,1,1)<sub>12</sub> models and SARIMA(2,1,2)x(0,1,1)<sub>12</sub> models.
- The best model was SARIMA(1,1,2)x(0,1,1)<sub>12</sub>. The ACF and PACF plots were mostly white noise but contained some significant residuals.

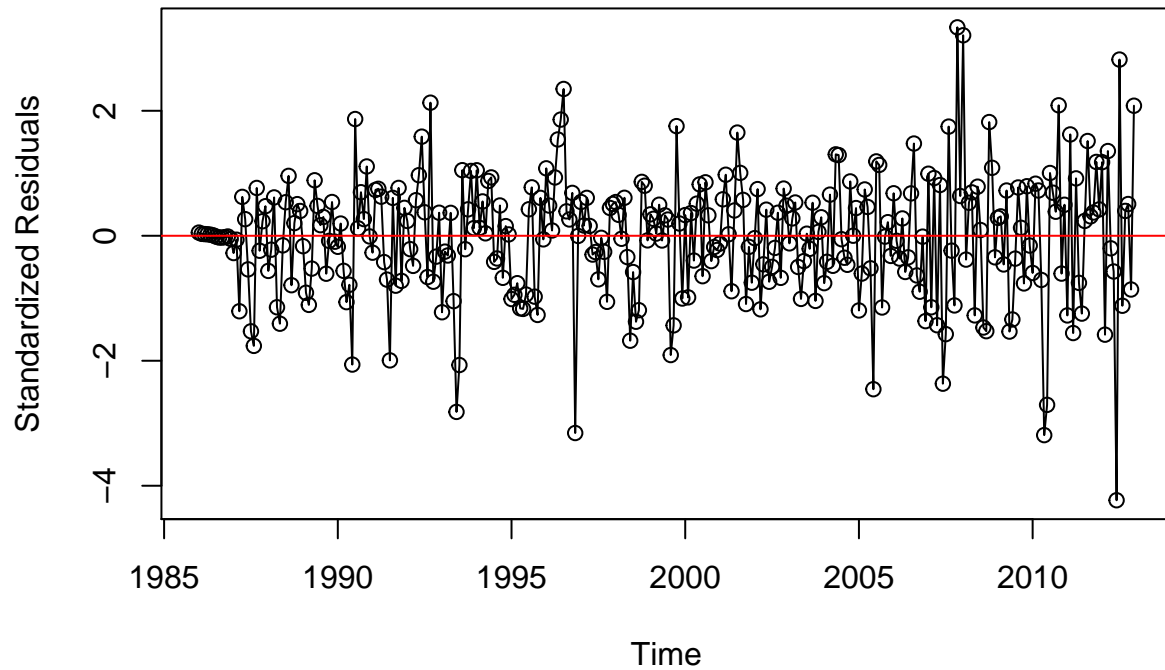
### 3.3 Diagnostic Check of models

#### 3.3.1 SARIMA(1,1,2)x(0,1,1)<sub>12</sub> model

#### 3.3.2 Time series plot for standardized residuals.

```
plot(window(rstandard(m4_sea_ice),start=c(1986, 1)),
      ylab='Standardized Residuals',type='o',
      main="Residuals from the SARIMA(1,1,2)x(0,1,1)_12 Model")
abline(h=0, col = 'red')
```

### Residuals from the SARIMA(1,1,2)x(0,1,1)\_12 Model

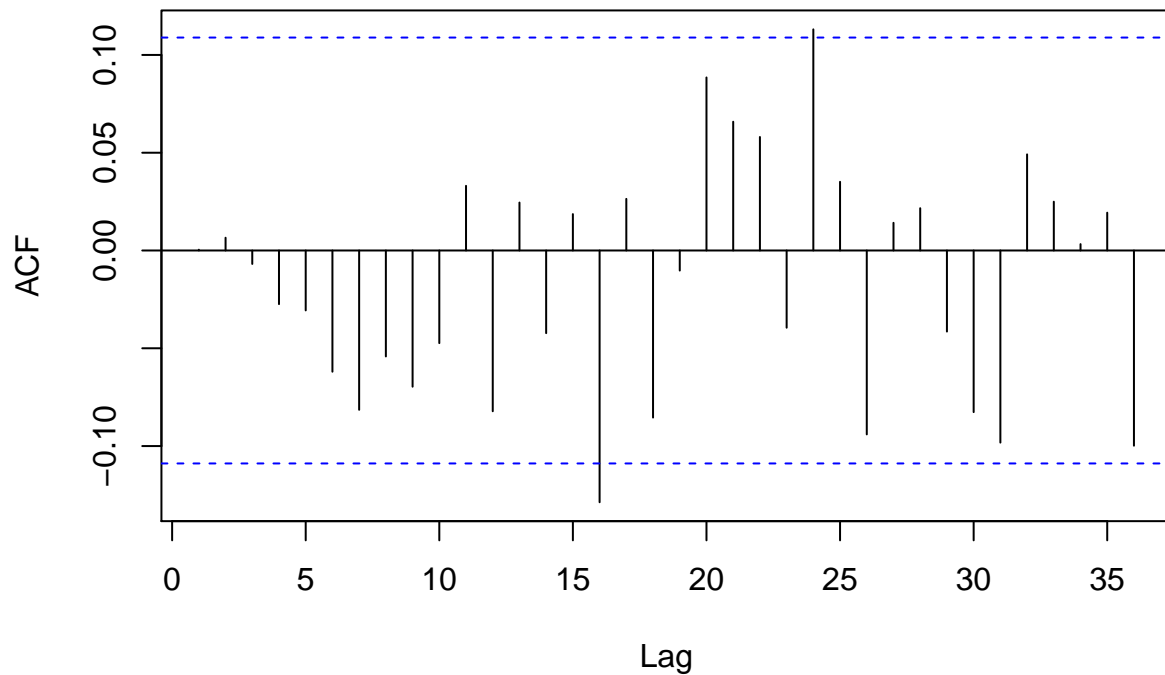


- Plot suggests no major abnormalities with this model(except around 1995) although there are several outliers that may need to be investigated in more detail.

#### 3.3.3 ACF plot of standardized residuals.

```
acf(as.vector(window(rstandard(m4_sea_ice),start=c(1986,1))),  
lag.max=36,  
main="ACF of Residuals from the SARIMA(1,1,2)x(0,1,1)_12 Model")
```

### ACF of Residuals from the SARIMA(1,1,2)x(0,1,1)\_12 Model



- Besides the slightly significant autocorrelations at lag 16 there is no sign of violation of the independence of residuals.

#### 3.3.3.1 Box-Ljung test.

```
Box.test(window(rstandard(m4_sea_ice), start = c(1986,1)), lag = 16, type = "Ljung-Box", fitdf = 0)
```

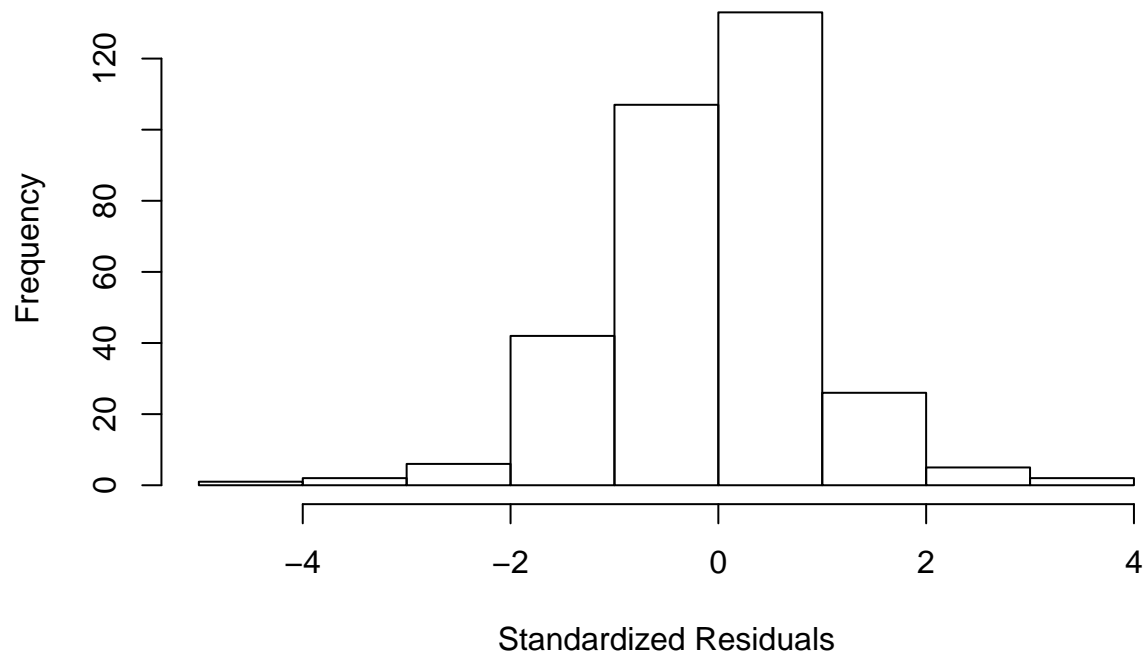
```
##
## Box-Ljung test
##
## data: window(rstandard(m4_sea_ice), start = c(1986, 1))
## X-squared = 16.708, df = 16, p-value = 0.4047
```

- Overall, the Ljung-Box test indicates that there is no problem in terms of independence of errors.

#### 3.3.3.2 Histogram of standardized residuals.

```
hist(window(rstandard(m4_sea_ice),start=c(1986,1)), xlab = 'Standardized Residuals', ylab = 'Frequency')
```

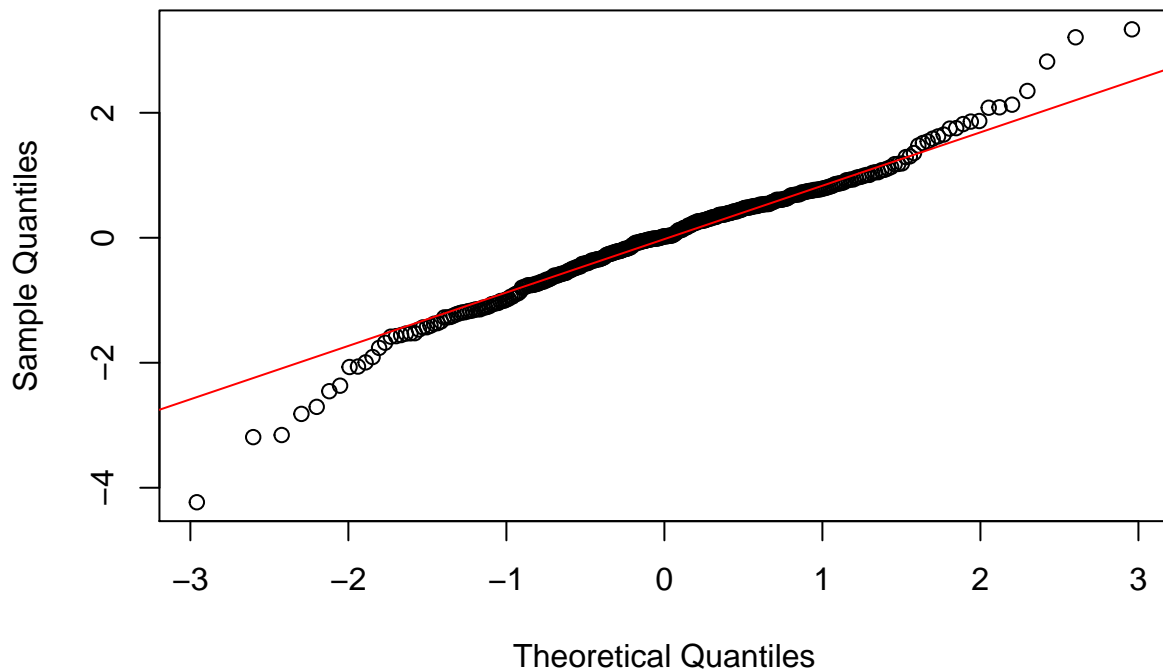
### Residuals of SARIMA(1,1,2)x(0,1,1)\_12 Model



#### 3.3.3.3 Q-Q plot of standardized residuals.

```
qqnorm(window(rstandard(m4_sea_ice),start=c(1986,1)),main="Q-Q plot for Residuals: SARIMA(1,1,2)x(0,1,1,1)",col=2)
qqline(window(rstandard(m4_sea_ice),start=c(1986,1)), col=2)
```

### Q-Q plot for Residuals: SARIMA(1,1,2)x(0,1,1)<sub>12</sub> Model



#### 3.3.3.4 Shapiro-Wilk test for normality of standardized residuals.

```
shapiro.test(window(rstandard(m4_sea_ice), start=c(1986,1), end=c(2012,12)))
```

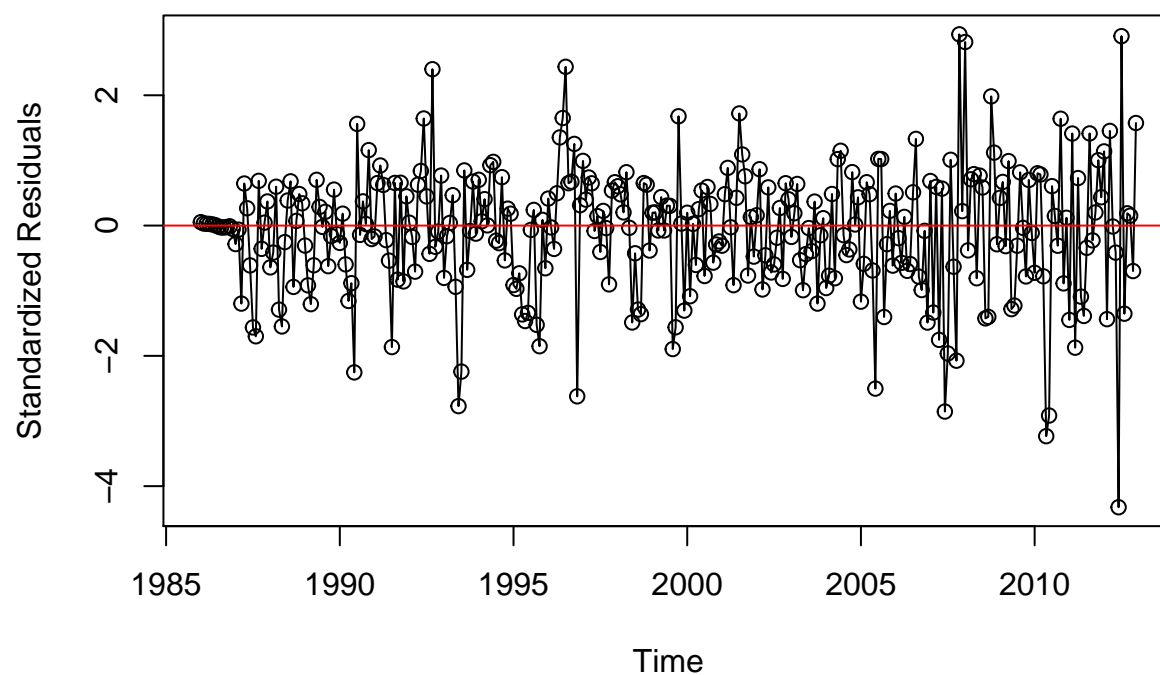
```
##
## Shapiro-Wilk normality test
##
## data:  window(rstandard(m4_sea_ice), start = c(1986, 1), end = c(2012, 12))
## W = 0.97724, p-value = 5.201e-05
```

#### 3.3.4 SARIMA(1,1,3)x(0,1,1)<sub>12</sub> model

##### 3.3.4.1 Time series plot for standardized residuals.

```
plot(window(rstandard(m5_sea_ice), start=c(1986, 1)),
      ylab='Standardized Residuals', type='o',
      main="Residuals from the SARIMA(1,1,3)x(0,1,1)_12 Model")
abline(h=0, col = 'red')
```

### Residuals from the SARIMA(1,1,3)x(0,1,1)\_12 Model



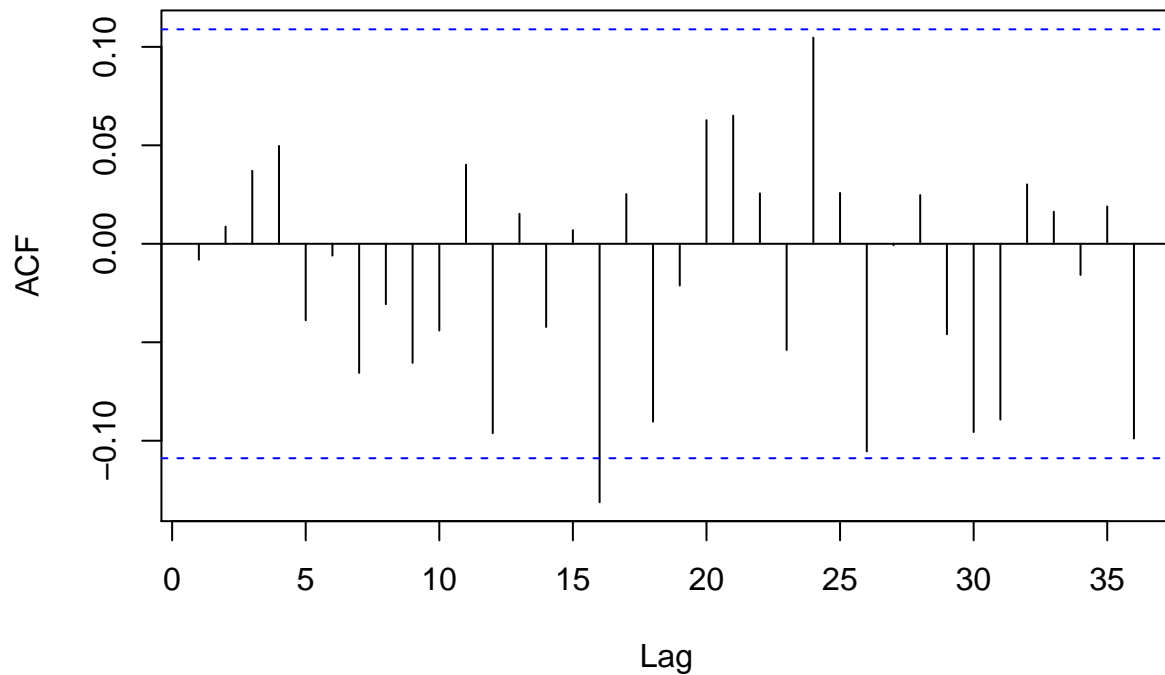
- Plot suggests no major abnormalities with this model(except around 1995) although there are several outliers that may need to be investigated in more detail.

#### 3.3.4.2 ACF plot of standardized residuals.

```
acf(as.vector(window(rstandard(m5_sea_ice),start=c(1986,1))),  
lag.max=36,  
main="ACF of Residuals from the SARIMA(1,1,3)x(0,1,1)_12 Model")
```



### ACF of Residuals from the SARIMA(1,1,3)x(0,1,1)\_12 Model



- Besides the slightly significant autocorrelations at lag 16 there is no sign of violation of the independence of residuals.

#### 3.3.4.3 Box-Ljung test.

```
Box.test(window(rstandard(m5_sea_ice), start = c(1986,1)), lag = 16, type = "Ljung-Box", fitdf = 0)
```

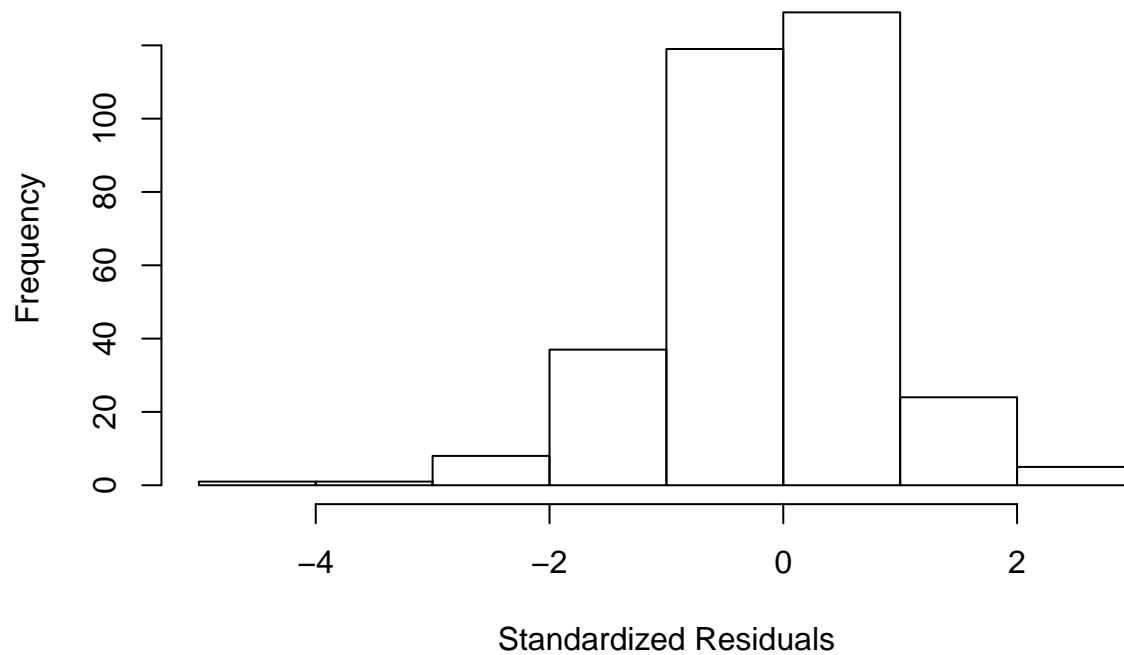
```
##  
## Box-Ljung test  
##  
## data: window(rstandard(m5_sea_ice), start = c(1986, 1))  
## X-squared = 15.733, df = 16, p-value = 0.4717
```

- Overall, the Ljung-Box test indicates that there is no problem in terms of independence of errors.

#### 3.3.4.4 Histogram of standardized residuals.

```
hist(window(rstandard(m5_sea_ice), start=c(1986,1)), xlab = 'Standardized Residuals', ylab = 'Frequency')
```

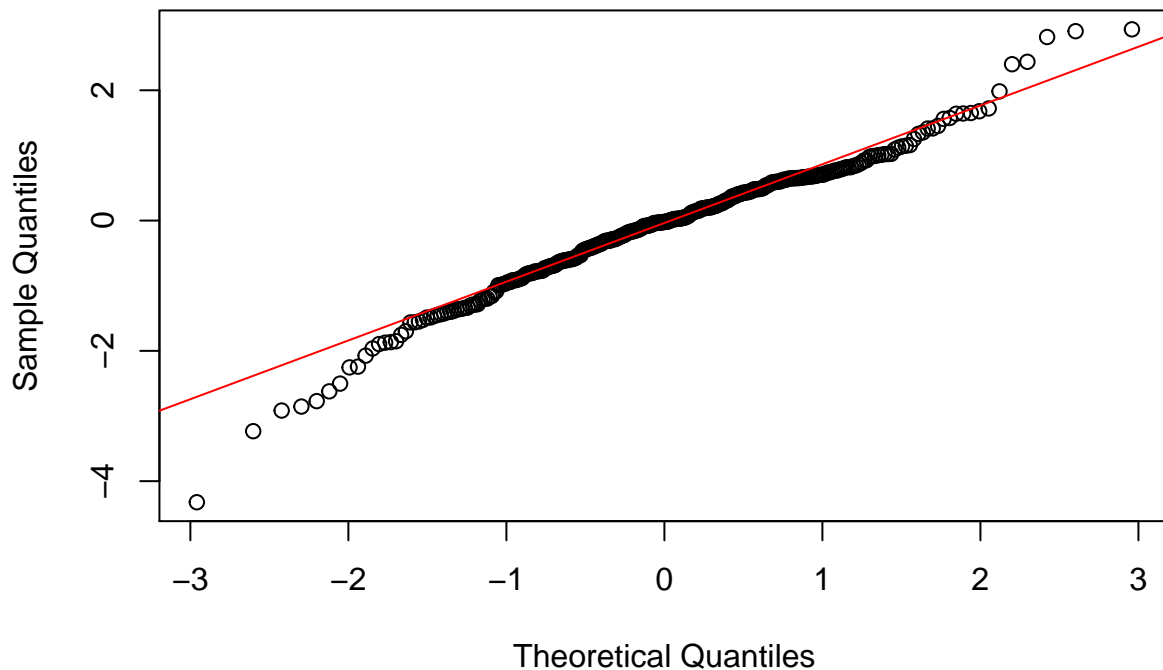
### Residuals of SARIMA(1,1,3)x(0,1,1)\_12 Model



#### 3.3.4.5 Q-Q plot of standardized residuals.

```
qqnorm(window(rstandard(m5_sea_ice),start=c(1986,1)),main="Q-Q plot for Residuals: SARIMA(1,1,3)x(0,1,1,1)",col=2)
qqline(window(rstandard(m5_sea_ice),start=c(1986,1)), col=2)
```

### Q-Q plot for Residuals: SARIMA(1,1,3)x(0,1,1)<sub>12</sub> Model



#### 3.3.4.6 Shapiro-Wilk test for normality of standardized residuals.

```
shapiro.test(window(rstandard(m5_sea_ice), start=c(1986,1), end=c(2012,12)))
```

```
##  
## Shapiro-Wilk normality test  
##  
## data: window(rstandard(m5_sea_ice), start = c(1986, 1), end = c(2012, 12))  
## W = 0.97567, p-value = 2.698e-05
```

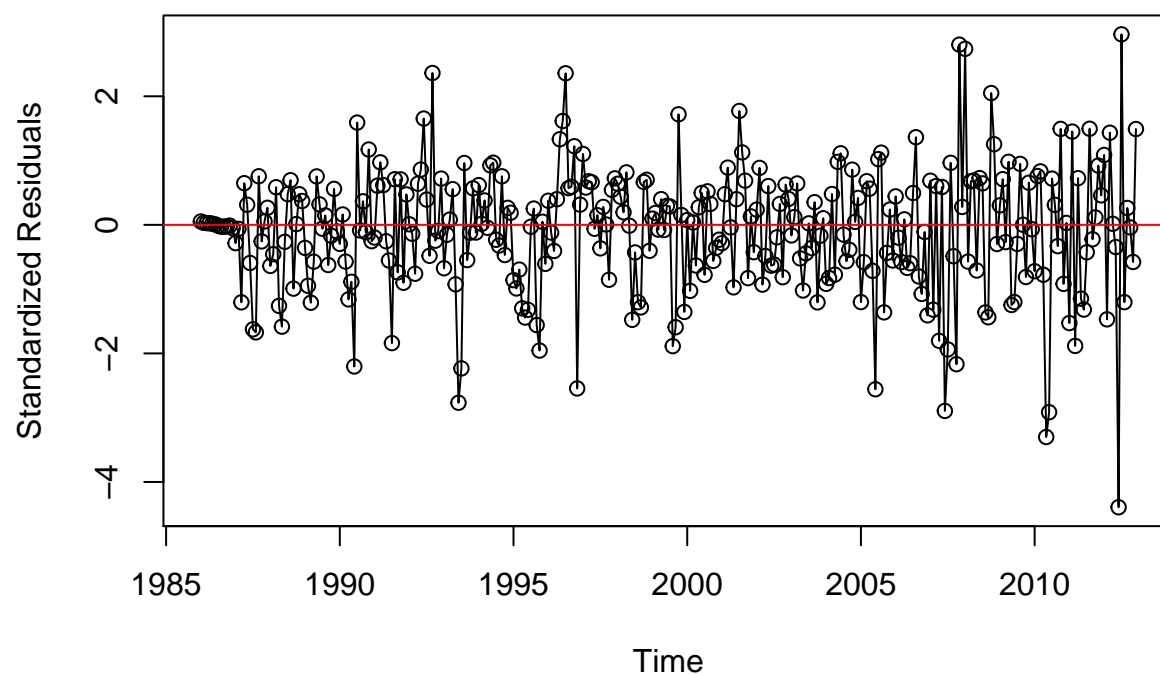
- Although we have mostly white noise residuals, the large-valued residuals make it impossible to conclude the normality of residuals by either the Q-Q plot or Shapiro test at 5% level of significance.

### 3.3.5 SARIMA(2,1,2)x(0,1,1)<sub>12</sub> model

#### 3.3.5.1 Time series plot for standardized residuals.

```
plot(window(rstandard(m8_sea_ice), start=c(1986, 1)),  
     ylab='Standardized Residuals', type='o',  
     main="Residuals from the SARIMA(2,1,2)x(0,1,1)12 Model")  
abline(h=0, col = 'red')
```

### Residuals from the SARIMA(2,1,2)x(0,1,1)\_12 Model

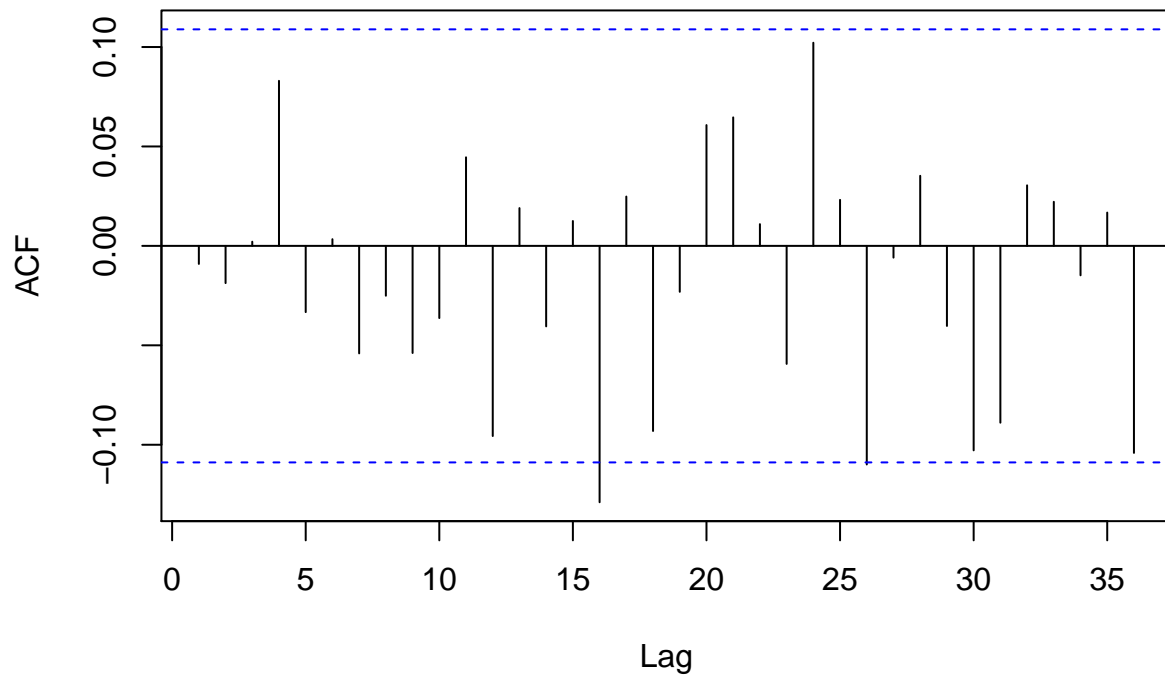


- Plot suggests no major abnormalities with this model(except around 1995) although there are several outliers that may need to be investigated in more detail.

#### 3.3.5.2 ACF plot of standardized residuals.

```
acf(as.vector(window(rstandard(m8_sea_ice),start=c(1986,1))),  
lag.max=36,  
main="ACF of Residuals from the SARIMA(2,1,2)x(0,1,1)_12 Model")
```

### ACF of Residuals from the SARIMA(2,1,2)x(0,1,1)\_12 Model



- Besides the slightly significant autocorrelations at lag 16 there is no sign of violation of the independence of residuals.

#### 3.3.5.3 Box-Ljung test.

```
Box.test(window(rstandard(m8_sea_ice), start = c(1986,1)), lag = 16, type = "Ljung-Box", fitdf = 0)
```

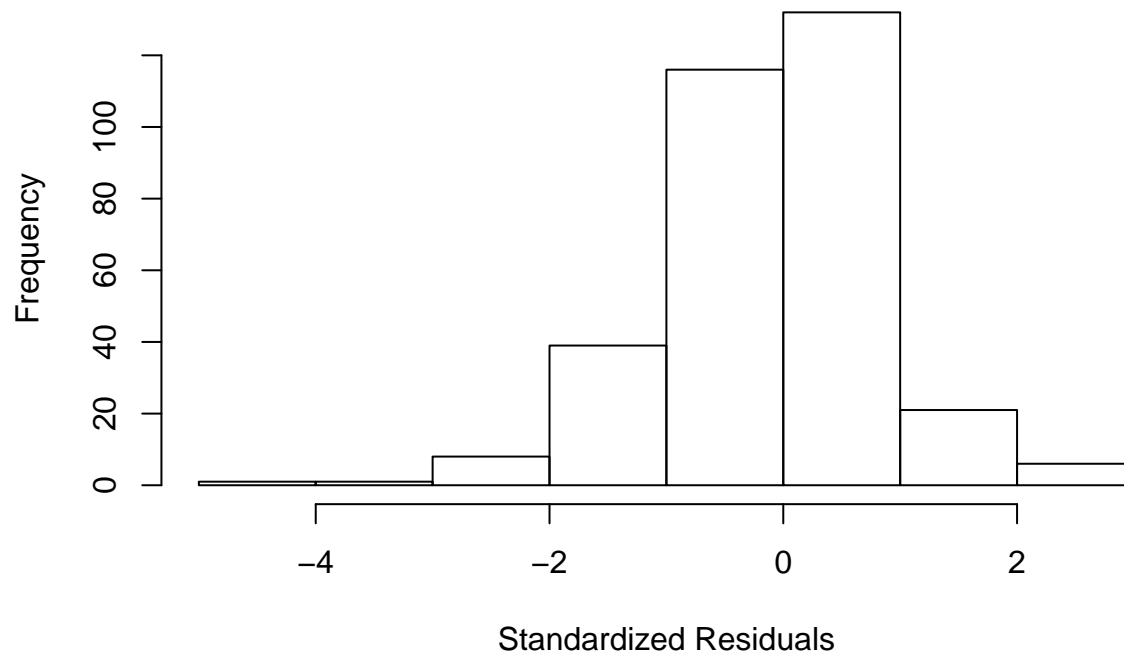
```
##
## Box-Ljung test
##
## data: window(rstandard(m8_sea_ice), start = c(1986, 1))
## X-squared = 15.587, df = 16, p-value = 0.4821
```

- Overall, the Ljung-Box test indicates that there is no problem in terms of independence of errors.

#### 3.3.5.4 Histogram of standardized residuals.

```
hist(window(rstandard(m8_sea_ice),start=c(1986,1)), xlab = 'Standardized Residuals', ylab = 'Frequency')
```

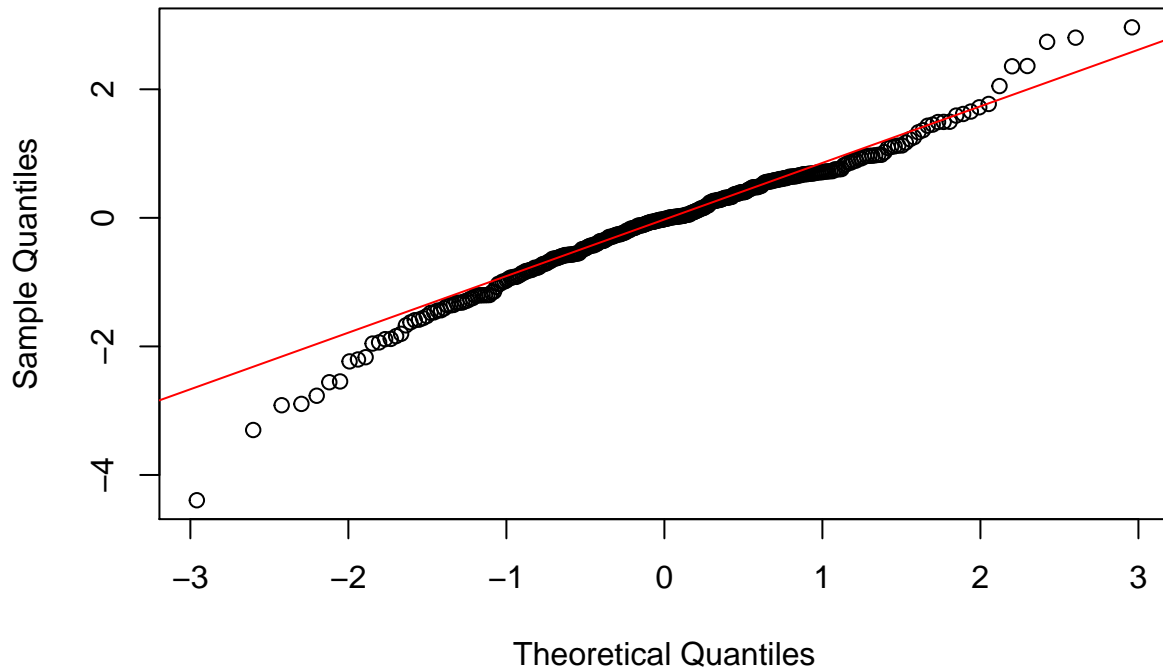
### Residuals of SARIMA(2,1,2)x(0,1,1)\_12 Model



#### 3.3.5.5 Q-Q plot of standardized residuals.

```
qqnorm(window(rstandard(m8_sea_ice),start=c(1986,1)),main="Q-Q plot for Residuals: SARIMA(2,1,2)x(0,1,1)",col=2)
qqline(window(rstandard(m8_sea_ice),start=c(1986,1)), col=2)
```

### Q-Q plot for Residuals: SARIMA(2,1,2)x(0,1,1)<sub>12</sub> Model



#### 3.3.5.6 Shapiro-Wilk test for normality of standardized residuals.

```
shapiro.test(window(rstandard(m8_sea_ice), start=c(1986,1), end=c(2012,12)))
```

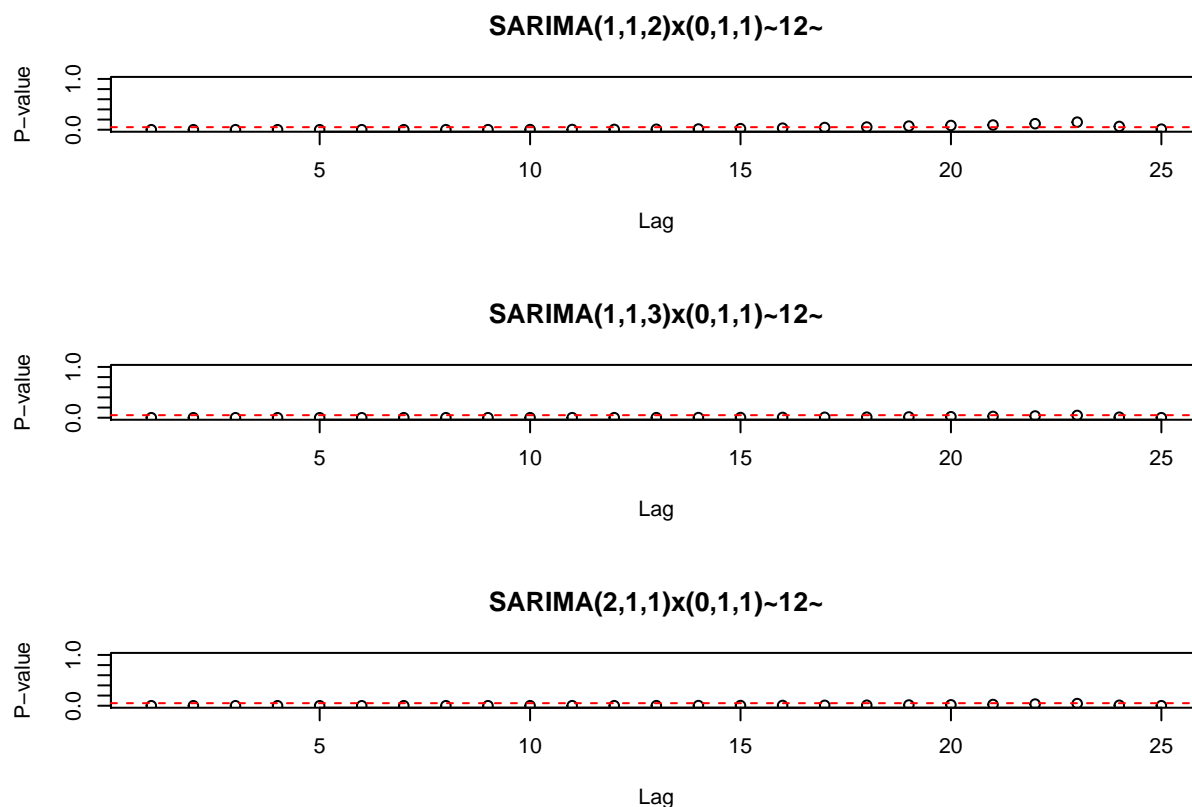
```
##
##  Shapiro-Wilk normality test
##
## data:  window(rstandard(m8_sea_ice), start = c(1986, 1), end = c(2012,      12))
## W = 0.97615, p-value = 3.299e-05
```

- Although we have mostly white noise residuals, the large-valued residuals make it impossible to conclude the normality of residuals by either the Q-Q plot or Shapiro test at 5% level of significance.
- The identified SARIMA models contained a number of slightly significant autocorrelations for their residuals and the residuals did not appear to normally distributed. Therefore we decided to check for the presence of an ARCH component.

#### 3.3.6 Check for ARCH component in residuals of SARIMA model.

##### 3.3.6.1 McLeod-Li test and Q-Q plot for the SARIMA(1,1,2)x(0,1,1)<sub>12</sub> model.

```
par(mfrow=c(3,1))
McLeod.Li.test(y=res_m4, main = 'SARIMA(1,1,2)x(0,1,1)~12~')
McLeod.Li.test(y=res_m5, main = 'SARIMA(1,1,3)x(0,1,1)~12~')
McLeod.Li.test(y=res_m8, main = 'SARIMA(2,1,1)x(0,1,1)~12~')
```



- McLeod-Li tests are mostly highly significant and the normality assumption is highly violated.

### 3.4 Consider overfitting (compare with SARIMA models $\text{SARIMA}(1,1,3)\times(0,1,1)_{12}$ and $\text{SARIMA}(2,1,2)\times(0,1,1)_{12}$ )

#### 3.4.1 ML estimates for $\text{SARIMA}(1,1,4)\times(0,1,1)_{12}$ model

```
m9_sea_ice = arima(sea_ice_ts, order=c(1,1,4), seasonal=list(order=c(0,1,1), period=12))
coeftest(m9_sea_ice)
```

```
##
## z test of coefficients:
##
##      Estimate Std. Error  z value  Pr(>|z|)
## ar1  -0.621139   0.247964  -2.5050  0.012247 *
## ma1   1.235405   0.253429   4.8748 1.089e-06 ***
## ma2   0.467886   0.178312   2.6240 0.008691 **
## ma3   0.016506   0.094981   0.1738 0.862040
## ma4   0.020480   0.071640   0.2859 0.774978
## sma1 -0.638814   0.046483 -13.7430 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```



### 3.4.2 ML estimates for SARIMA(3,1,3)x(0,1,1)<sub>12</sub> model

```
m10_sea_ice = arima(sea_ice_ts, order=c(3,1,3), seasonal=list(order=c(0,1,1), period=12))
coeftest(m10_sea_ice)
```

```
##
## z test of coefficients:
##
##      Estimate Std. Error  z value  Pr(>|z|)
## ar1   0.359364   0.201405   1.7843   0.074378 .
## ar2   0.555213   0.199476   2.7834   0.005380 **
## ar3  -0.133740   0.110446  -1.2109   0.225932
## ma1   0.193508   0.200443   0.9654   0.334344
## ma2  -0.776139   0.108064  -7.1822 6.858e-13 ***
## ma3  -0.383398   0.131857  -2.9077   0.003641 **
## sma1 -0.636961   0.046565 -13.6790 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

- It was not possible to get significant coefficients in the over-fitted model indicating that SARIMA(1,1,2)x(0,1,1)<sub>12</sub> is the most suitable model.

### 3.5 Comparison of AIC values for different SARIMA models.

- m4\_sea\_ice\_ts SARIMA(1,1,2)x(0,1,1)<sub>12</sub>, m5\_sea\_ice\_ts SARIMA(1,1,3)x(0,1,1)<sub>12</sub>, m8\_sea\_ice\_ts SARIMA(2,1,2)x(0,1,1)<sub>12</sub>

```
AIC(m4_sea_ice_ts, m5_sea_ice_ts, m8_sea_ice_ts)
```

```
##           df          AIC
## m4_sea_ice_ts  5 108.83175
## m5_sea_ice_ts  6  96.29905
## m8_sea_ice_ts  6  96.18930
```

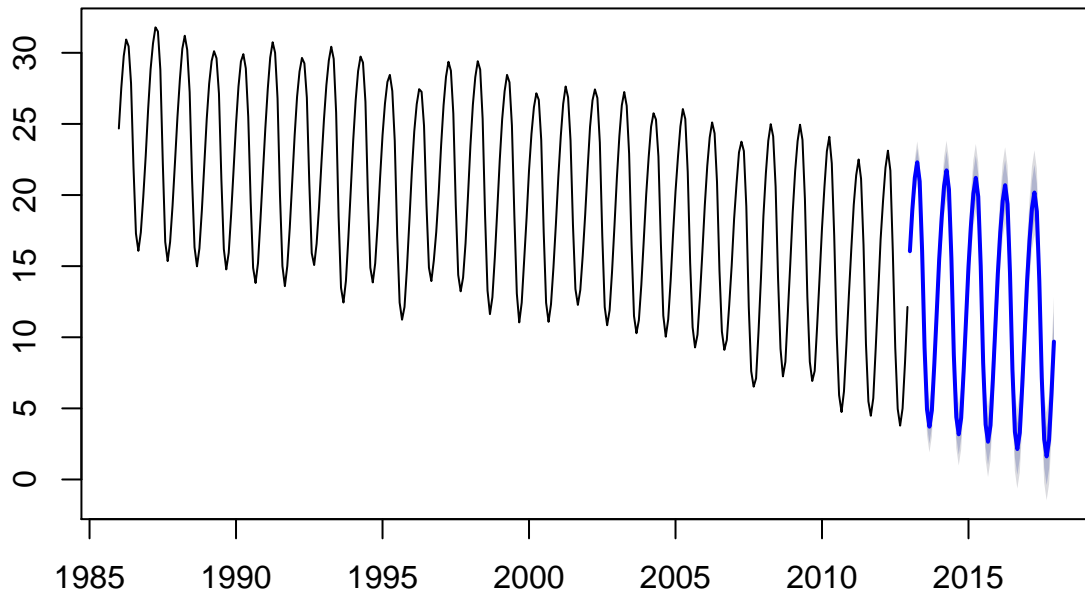
- Values obtained from the AIC analysis indicate that the best models are SARIMA(1,1,3)x(0,1,1)<sub>12</sub> and SARIMA(2,1,2)x(0,1,1)<sub>12</sub>.

### 3.6 Prediction of Seasonal Arctic Sea-Ice

#### 3.6.1 Five year forecast for SARIMA(1,1,3)x(0,1,1)<sub>12</sub>.

```
sea_ice_for_113 = Arima(sea_ice_ts, order=c(1,1,3), seasonal=list(order=c(0,1,1), period=12))
future_113 = forecast(sea_ice_for_113, h = 60)
plot(future_113, main = 'Five Year Forecast - SARIMA(1,1,3)x(0,1,1)_12')
```

### Five Year Forecast – SARIMA(1,1,3)x(0,1,1)<sub>12</sub>



\* The 5-year forecast is shown as a blue line and the forecast limits are grey.

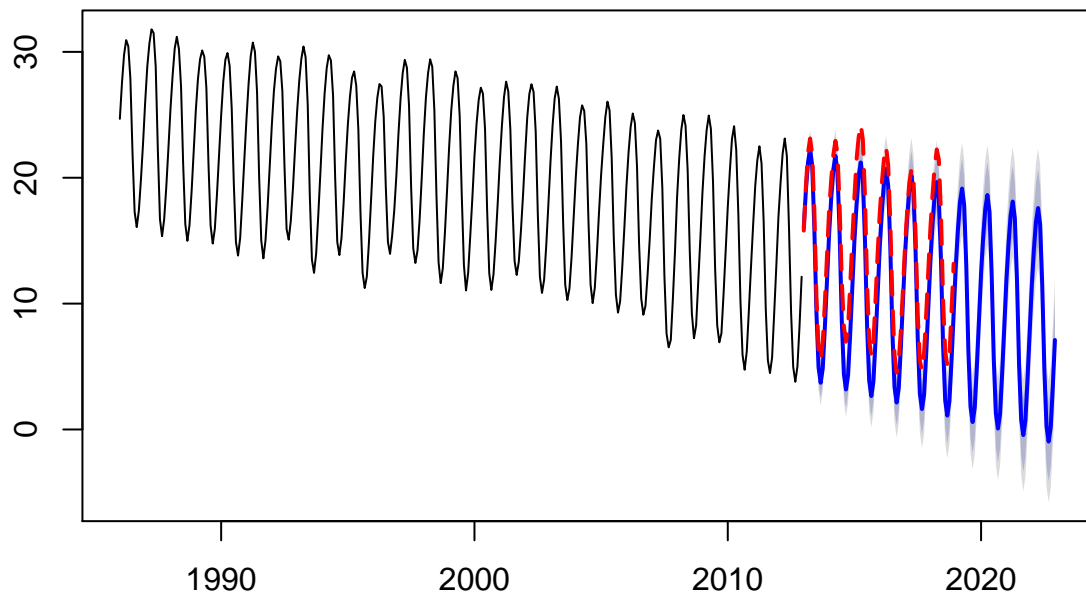
#### 3.6.2 Long-term (10 year) forecast for SARIMA(1,1,3)x(0,1,1)<sub>12</sub>.

```
sea_ice_for_long_113 = Arima(sea_ice_ts,order=c(1,1,3),seasonal=list(order=c(0,1,1), period=12))

future_long_113 = forecast(sea_ice_for_long_113, h = 120)

par(mfrow=c(1,1))
plot(future_long_113, main = 'Ten Year Forecast - SARIMA(1,1,3)x(0,1,1)_12')
lines(window(sea_ice_long_ts,start=c(2013, 1)), col = 2, lty = 5, lwd = 2, type = 'l')
```

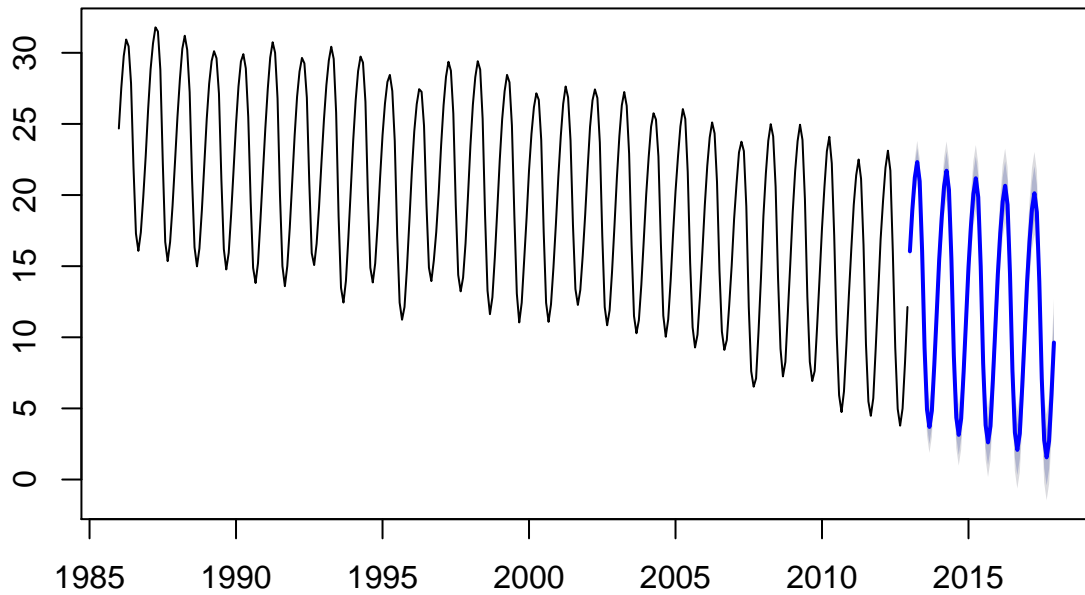
### Ten Year Forecast – SARIMA(1,1,3)x(0,1,1)<sub>12</sub>



#### 3.6.3 Five year forecast for SARIMA(2,1,2)x(0,1,1)<sub>12</sub>.

```
sea_ice_for_212 = Arima(sea_ice_ts,order=c(2,1,2),seasonal=list(order=c(0,1,1), period=12))  
  
future_212 = forecast(sea_ice_for_212, h = 60)  
plot(future_212, main = 'Five Year Forecast - SARIMA(2,1,2)x(0,1,1)_12')
```

### Five Year Forecast – SARIMA(2,1,2)x(0,1,1)<sub>12</sub>



\* The 5-year forecast is shown as a blue line and the forecast limits are grey.

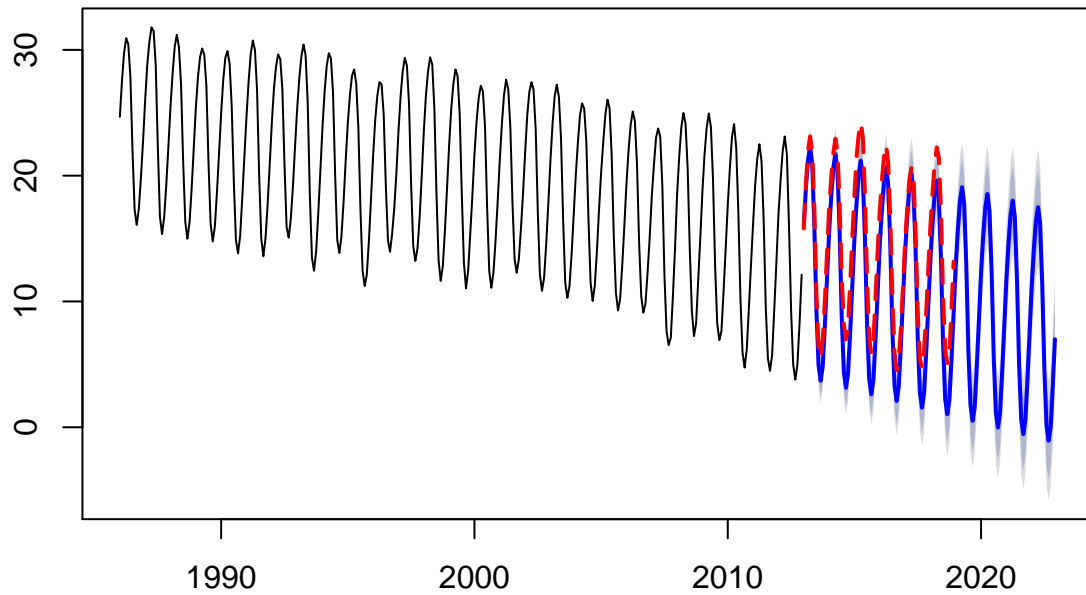
#### 3.6.4 Long-term (10 year) forecast for SARIMA(2,1,2)x(0,1,1)<sub>12</sub>.

```
sea_ice_for_long_212 = Arima(sea_ice_ts,order=c(2,1,2),seasonal=list(order=c(0,1,1), period=12))

future_long_212 = forecast(sea_ice_for_long_212, h = 120)

par(mfrow=c(1,1))
plot(future_long_212, main = 'Ten Year Forecast - SARIMA(2,1,2)x(0,1,1)_12')
lines(window(sea_ice_long_ts,start=c(2013, 1)), col = 2, lty = 5, lwd = 2, type = 'l')
```

### Ten Year Forecast – SARIMA(2,1,2)x(0,1,1)\_12



- The truncated data (2013-2018) has been overlayed (dashed red line) onto the 10-year forecasts for Arctic sea-ice volumes. The data fit well within the forecast limits for both models and highlights the potential for forecasts to gain a idea of future Arctic sea-ice volumes.