

# GB13604 - Maths for Computer Science

## Lecture 9 – Probability, Part II

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This course is based on Mathematics for Computer Science, Spring 2015, by Albert Meyer and Adam Chlipala, Massachusetts Institute of Technology OpenCourseWare.



# Introduction

- Independence and Causality
- Random Variables
- Expectation

# Independence and Causality

# Independent Events

- If it is raining, the probability that people will take the bus from tsukuba station is higher than if it is sunny.

$$\Pr(\text{Take the bus} \mid \text{raining}) \neq \Pr(\text{Take the bus} \mid \text{sunny})$$

- If your father and mother have type A blood, the probability that you have type A blood is higher.

$$\Pr(A_{me} \mid A_P \& A_M) \geq \Pr(A_{me})$$

- If you flip two separate coins, the probability of the second coin being heads is not influenced by the result of the first coin.

$$\Pr(H_2 \mid H_1) = \Pr(H_2 \mid \bar{H}_1) = \Pr(H_2)$$

# Independent Events

Two events,  $A$  and  $B$  are **independent** if the probability of  $A$  **does not depend** on the occurrence of  $B$ .

## Mathematical Definitions:

- **Definition 1:** Events  $A$  and  $B$  are **independent** iff:

$$\Pr(A) = \Pr(A|B) \quad (1)$$

- **Definition 2:** Events  $A$  and  $B$  are **independent** iff:

$$\Pr(A) \cdot \Pr(B) = \Pr(A \cap B) \quad (2)$$

# Proof of Equivalence of Definitions 1 and 2

Definition 1:

$$\Pr(A) = \Pr(A|B) \text{ iff}$$

Conditional Prob Definition:

$$\Pr(A) = \frac{\Pr(A \cap B)}{\Pr(B)} \text{ iff}$$

Algebra:

$$\Pr(A) \cdot \Pr(B) = \Pr(A \cap B).$$

# Independence: Characteristics

## Simmetry:

$$\Pr(A) \cdot \Pr(B) = \Pr(A \cap B).$$

If  $A$  is independent from  $B$ , **then**  $B$  is independent from  $A$ .

## Zero probability event:

If  $\Pr(B) = 0$  then:

- **Definition 1:** Does not work
- **Definition 2:**  $\Pr(A) \cdot 0 = \Pr(A \cap \emptyset) = 0$ .

So,  $B$  is **independent from all events!**

# Independence: Complement

A independent of B means probability of A does not change if B happens or not. **Therefore** A should also be independent of **Not B**.

**Lemma:** A independent of B **iff** A independent of NOT B.

**Proof:** Proof can be derived using algebra from:

$$\Pr(A - B) = \Pr(A) - \Pr(A \cap B)$$



# Mutual Independence: Idea

**Experiment:** Two coins are thrown

- **Event 1:** Coin 1 is heads:  $\Pr(H_1) = (0, 0), (0, 1), (1, 0), (1, 1) = 0.5$
- **Event 2:** Coin 2 is heads:  $\Pr(H_2) = (0, 0), (0, 1), (1, 0), (1, 1) = 0.5$
- **Event 3:** # of heads is odd:  
 $\Pr(O) = (0, 0), (0, 1), (1, 0), (1, 1) = 0.5$

# Mutual Independence: Idea

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 $\Pr(O) = (0, 0), (0, 1), (1, 0), (1, 1) = 0.5$

Events  $H_1, H_2, H_1, O, H_2, O$  are, **2-by-2**, independent:

- $\Pr(O|H_1) = (1, 0), (1, 1) = 0.5$
- $\Pr(O|H_2) = (0, 1), (1, 1) = 0.5$
- $\Pr(H_1|H_2) = (0, 1), (1, 1) = 0.5$

These 3 events are **2-way independent**.

# Mutually Independent Events – Definition

Events  $A_1, A_2, A_3, \dots, A_n$  are **Mutually Independent** when:

The probability that  $A_i$  occurs is not changed by the occurrence of the other events.

# Mutually Independent Events – Example

**Experiment:** You throw the coin  $n$  times

**Event:**  $H_i$  the  $i$ -th throw is heads

What happens in the 5th flip is independent from the 1st flip, or the 7th flip, etc.

$$\Pr(H_5) = \Pr(H_5 | H_1 \cap H_3 \cap H_7 \cap \dots)$$

# Mutually Independent Events – Math Def

Events  $A_1, A_2, \dots, A_n$  are **mutually independent** when:

**Definition 1:**

$$\Pr(A_i) = \Pr(A_i | A_1 \cap A_2 \cap \dots \cap A_j \cap \dots \cap A_n); j \neq i \quad (3)$$

**Definition 2:**

$$\Pr(A_1 \cap A_2 \cap A_3 \cap \dots \cap A_n) = \Pr(A_1) \cdot \Pr(A_2) \cdot \Pr(A_3) \cdot \dots \cdot \Pr(A_n) \quad (4)$$

# K-wise independence

**Experiment:** Two coins are thrown

- **Event 1:** Coin 1 is heads:  $\Pr(H_1) = (0, 0), (0, 1), (1, 0), (1, 1) = 0.5$
  - **Event 2:** Coin 2 is heads:  $\Pr(H_2) = (0, 0), (0, 1), (1, 0), (1, 1) = 0.5$
  - **Event 3:** # of heads is odd:  
 $\Pr(O) = (0, 0), (0, 1), (1, 0), (1, 1) = 0.5$
- 
- Each pair  $(H_1, H_2), (H_1, O), (H_2, O)$  is mutually independent.
  - However,  $(H_1, H_2, O)$  are not mutually independent:  
 $\Pr(O|H_1 \cap H_2) = 0$

# K-wise independence

**Experiment:**  $n$  coins are thrown

- **Event  $i$ :** Coin  $i$  is heads:  $\Pr(H_i) = 0.5$
- **Event  $i + 1$ :** # of heads is odd:  $\Pr(O) = 0.5$

Any subset of  $i$  events is mutually independent.  
However, the set of all  $i + 1$  events is **not**.

# K-wise independence

A set of  $m$  events is  $k$ -wise independent if:

$$A_1, A_2, \dots, A_m$$

Any subset of  $k$  events is mutually independent.



# Random Variables

# Playing tRPGs

In the game **3:16**, you **win a space battle** if you roll a d10 dice, and the result is below the value **FA**, which you decide before the game.

- If you lose one battle **you lose your armor**
- If you lose another battle **you are wounded**
- If you lose a third battle **you die**



# Events and Numbers

Until now, we approached probability from a **Set** and **Events** point of view:

- **Event 1:** Result of the dice:  $\{1, 2, 3, 4, 5, 6, 7, 8, 9, 10\}$
- **Event 2:** Value of **FA**:  $\{1, 2, 3, 4, 5, 6, 7, 8, 9, 10\}$
- **Event 3:** You win a combat with  $FA = 6$ :  
 $\{1, 6\}, \{2, 6\}, \{3, 6\}, \dots$
- **Event 4:** 3 victories in 5 battles with  $FA = 6$ :  
 $\{E_3, E_3, E_3, \bar{E}_3, \bar{E}_3\}, \{E_3, E_3, \bar{E}_3, E_3, \bar{E}_3\}, \dots$

# Events and Numbers

However, many times we want to see probability in terms of **Numbers**:

- What is the probability that I can **survive 7 battles** if my FA is 6?
- What is the **minimum FA** I need to have more than .8 chance to survive 5 battles?
- **How many battles in a row** I expect to win if my FA is 6?

Intuitively, a **Random Variable** is a number that comes from a **Random Process**

- # of hours to the next **System Crash**
- # **faulty pixels** in the monitor
- # heads in a **coin flip**

# Random Variable: Formal Definition

A random variable is a **total function** from a **sample space** to a **numeric domain**.

$$R : S \rightarrow \mathbb{R} \text{ or } R : S \rightarrow \mathbb{N}, \text{ etc...} \quad (5)$$

So a random *variable* is really a **function**...  
(... maybe that is a bad name)

# Random Variable: Formal Example

**Experiment:** We throw three *fair* coins independently.

- Variable  $C ::=$  Number of heads (**C**ount)
- Variable  $M ::=$  1 if all coins **M**atch, 0 if they don't **M**atch
- $M = 1$  when  $\{H, H, H\}$  or  $\{T, T, T\}$ ,  
 $\Pr(M = 1) = 2/8$
- $C = 1$  when  $\{H, T, T\}, \{T, H, T\}, \{T, T, H\}$ ,  
 $\Pr(C = 1) = 3/8$
- $\Pr(C \cdot M > 0) = \Pr(C > 0 \cap M > 0) = 1/8$

# “Indicator” random variables

An **indicator random variable** takes a value of 0 or 1 for a subset of the sample space.

- Variable  $M ::= 1$  if all coins **M**atch, 0 if they don't **M**atch
- Variable  $O ::= 1$  if number of heads is **O**dd.

Other **integer** random variables **partition** the sample space:

- $C = 0 \rightarrow \{T, T, T\}$
- $C = 1 \rightarrow \{H, T, T\}, \{T, H, T\}, \{T, T, H\}$
- $C = 2 \rightarrow \{H, T, H\}, \{H, H, T\}, \{T, H, H\}$
- $C = 3 \rightarrow \{H, H, H\}$

# Independence and Random Variables

So, a **Random Variable** is a **function** that assigns a value to a set of outcomes.

Two or more **Random Variables**  $R_1, R_2, \dots, R_n$  are mutually independent **iff**  $(R_1 = a_1), (R_2 = a_2), \dots, (R_n = a_n)$  are all **mutually independent events**.

**Alternatively:**

$$\Pr(R_1 = a_1 \cap R_2 = a_2 \cap R_3 = a_3 \cap \dots) = \Pr(R_1 = a_1) \cdot \Pr(R_2 = a_2) \dots$$

For all values of  $a_1$



# Independence and Random Variables

Are R.V.s  $C$  and  $M$  independent?

- Variable  $M ::= 1$  if all coins **M**atch, 0 if they don't **M**atch
- Variable  $C ::=$  number of heads (**C**ount).

**Answer:** NO!

- $\Pr(M = 1) \cdot \Pr(C = 1) > 0$
- $\Pr(M = 1 \cap C = 1) = 0$

# Independence and Random Variables

Are R.V.s  $O$  and  $M$  independent?

- Variable  $M ::= 1$  if all coins **M**atch, 0 if they don't **M**atch
- Variable  $O ::= 1$  if # heads is **O**dd, 0 if it is even.

**Answer:** YES!

- $\Pr(M = 1) \cdot \Pr(O = 1) = \frac{2}{8} \cdot \frac{1}{2} = \frac{1}{8}$
- $\Pr(M = 1 \cap O = 1) = \frac{1}{8}$
- Don't forget to test for the other combinations:  
 $(M = 0, O = 1), (M = 1, O = 0), \dots$
- To prove **NOT** independence, you just need one counterexample.

# Uniform Random Variables

A **Uniform** Random Variable is one where the **all values are equally likely**.

- $D_n ::=$  result of a **fair** dice with  $n$  sides.

$$\Pr(D_6 = 1) = \Pr(D_6 = 2) = \dots = \Pr(D_6 = 6) = \frac{1}{6}$$

- $S_4 ::=$  lottery number with 4 digits.

$$\Pr(S_4 = 0000) = \Pr(S_4 = 1234) = \dots = \Pr(S_4 = 9998) = 0.0001$$

# Binomial Random Variable

**Experiment:** We throw a biased coin  $n$  times, each time with  $p$  chance to turn heads.

**Event:**  $B_{n,p} ::= \#$  of heads in  $n$  mutually independent flips.

For  $n = 5, p = 2/3$ ,  $\Pr(HHTTH)$ :

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For  $n = 5, p = 2/3$ ,  $\Pr(HHTTH)$ :

- $= \Pr(H) \cdot \Pr(H) \cdot \Pr(T) \cdot \Pr(T) \cdot \Pr(H)$

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For  $n = 5, p = 2/3$ ,  $\Pr(HHTTH)$ :

- $= \Pr(H) \cdot \Pr(H) \cdot \Pr(T) \cdot \Pr(T) \cdot \Pr(H)$
- $= \frac{2}{3} \cdot \frac{2}{3} \cdot \frac{1}{3} \cdot \frac{1}{3} \cdot \frac{2}{3}$

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- $= \Pr(H) \cdot \Pr(H) \cdot \Pr(T) \cdot \Pr(T) \cdot \Pr(H)$
- $= \frac{2}{3} \cdot \frac{2}{3} \cdot \frac{1}{3} \cdot \frac{1}{3} \cdot \frac{2}{3}$
- $= \frac{2^3}{3} \cdot \frac{1^2}{3}$

# Binomial Random Variable

**Experiment:** We throw a biased coin  $n$  times, each time with  $p$  chance to turn heads.

**Event:**  $B_{n,p} ::= \#$  of heads in  $n$  mutually independent flips.

For  $n = 5, p = 2/3$ :

- $\Pr(\text{ONE seq with } i \text{ H and } (n-i) \text{ T}):$   
 $p^i(1-p)^{n-i}$
- $\Pr(\text{ANY seq with } i \text{ H and } (n-i) \text{ T}):$   
 $\binom{n}{i} p^i(1-p)^{n-i}$



# Probability Density Function

The **Probability Density Function (PDF)** of a random variable  $R$  defines the probability that  $R$  will have a given value  $i$

$$\text{PDF}_R(i) ::= \Pr(R = i)$$

**Example 1:** for the binomial random variable:

$$\text{PDF}_{B_{n,p}}(i) = \binom{n}{i} p^i (1-p)^{n-i}$$

**Example 2:** for the uniform random variable:

$$\text{PDF}_U(v) = \text{constant}$$

(for  $v$  in range of  $U$ )

# Cumulative Distribution

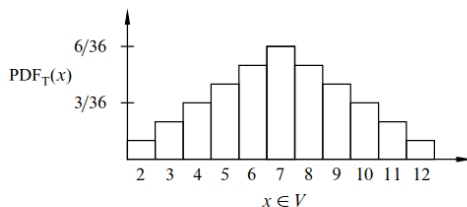
The **Cumulative Distribution Function** is the function that measures the probability that random variable  $R$  will have a value **smaller or equal** than  $i$

$$\text{CDF}_R(i) ::= \Pr(R \leq i)$$

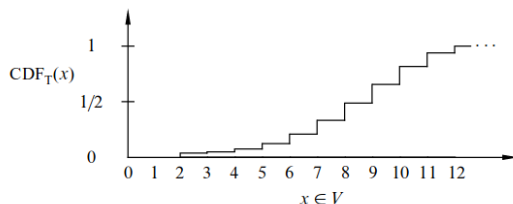
## Important points

- PDF and CDF capture similar information.
- PDF and CDF **do not involve** the actual sample space.
- Many different **experiment** have **very similar PDFs**

# PDF and CDF example: 2d6



**Figure 18.1** The probability density function for the sum of two 6-sided dice.



**Figure 18.2** The cumulative distribution function for the sum of two 6-sided dice.

# Expectation

## Game: Carnival Dice

Consider the following game: Choose a number of 1 to 6 and roll 3 dice. Each dice that rolls the same number you win \$1. If no dice roll that number, you lose \$1.

**Example:** You choose number 5:

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**Example:** You choose number 5:

- 2, 3, 4: Lose \$1

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- 2, 3, 4: Lose \$1
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## Game: Carnival Dice

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**Example:** You choose number 5:

- 2, 3, 4: Lose \$1
- 5, 4, 6: Win \$1
- 5, 5, 1: Win \$2



# Game: Carnival Dice

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**Example:** You choose number 5:

- 2, 3, 4: Lose \$1
- 5, 4, 6: Win \$1
- 5, 5, 1: Win \$2
- 5, 5, 5: Win \$3

# Game: Carnival Dice

Consider the following game: Choose a number of 1 to 6 and roll 3 dice. Each dice that rolls the same number you win \$1. If no dice roll that number, you lose \$1.

**Example:** You choose number 5:

- 2, 3, 4: Lose \$1
- 5, 4, 6: Win \$1
- 5, 5, 1: Win \$2
- 5, 5, 5: Win \$3
- Is this a fair game?

# Game: Carnival Dice Analysis

- $\Pr(\text{zero 5s}) = \frac{5^3}{6^3} = \frac{125}{216} = \text{Lose \$1}$
- $\Pr(\text{one 5}) = \binom{3}{1} \frac{5^2}{6^2} \frac{1}{6} = \frac{75}{216} = \text{Win \$1}$
- $\Pr(\text{two 5}) = \binom{3}{2} \frac{5}{6} \frac{1}{6}^2 = \frac{15}{216} = \text{Win \$2}$
- $\Pr(\text{three 5}) = \frac{1}{6}^3 = \frac{1}{216} = \text{Win \$3}$

# Carnival Game: Law of Averages

Every **216 Games**, you expect **125 games** with 0 matches, **75 games** with 1 match, **15 games** with 2 matches, and only **one game** with 3 matches!

$$\frac{216 \times (-1) + 75 \times (1) + 15 \times (2) + 1 \times (3)}{216} = \frac{-17}{216} = -0.08$$

So on average, you expect to **lose \$0.08** per game. Not a fair game!

# Carnival Game: Impossible expectations

Note that, for the carnival game, you **expect to lose 0.08 per game** but **the result -0.08 does not exist!**

**Another example:** Rolling one 6-sided die:

$\Pr(1) = 1/6$ ,  $\Pr(2) = 1/6$ ,  $\Pr(3) = 1/6$ , ...

$$\text{Expected value} = \frac{1 \cdot 1 + 1 \cdot 2 + 1 \cdot 3 + 1 \cdot 4 + 1 \cdot 5 + 1 \cdot 6}{6} = \frac{21}{6} = 3.5$$

You expect an **average result** of 3.5, but the 3.5 result does not exist!

# Expectation: Definitions

The **Expected Value** of a variable  $R$  is the **Average Value** of  $R$ , weighted by the probabilities.

$$E[R] ::= \sum_{v \in \text{range}(R)} v \cdot \Pr[R = v] \quad (6)$$

So,  $E[\$ \text{ at carnival game}] = -0.08$

**Note 1:** We can use sums in this definition because the **probability space is countable**.

**Note 2:** For **continuous probability spaces** we need to use integrals.

# Total Expectation

Law of Total Expectation is useful for reasoning by cases:

- **Conditional Expectation:**  $E[R|A] = \sum v \cdot \Pr[R = v|A]$
- **Total Expectation:**  $E[R] = E[R|A] \cdot \Pr[A] + E[R|\bar{A}] \cdot \Pr[\bar{A}]$

In other words: The total expectation is The expectation when  $A$  happens plus The expectation when  $A$  does not happen.

# General Total Expectation

More generally:

$$E[R] = E[R|A_1] \cdot \Pr[A_1] + E[R|A_2] \cdot \Pr[A_2] + \dots + E[R|A_n] \cdot \Pr[A_n] \quad (7)$$

when  $\Pr[A_1] + \Pr[A_2] + \dots + \Pr[A_n] = 1$



# Total Expectation: How many heads?

What is **expected number of Heads** when we flip a coin  $n$  times with **bias  $p$** ?

A coin with **bias  $p$** :

- Will turn **Heads** with probability  $p$
- Will turn **Tails** with probability  $q = 1 - p$

$E[n]$  expectation of heads after  $n$  flips - **study by cases**:

- **Case 1:** First coin is Heads:  $E[n|H_1] = 1 + E[n - 1]$
- **Case 2:** First coin is Tails:  $E[n|T_1] = E[n - 1]$

# Total Expectation: How many heads?

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By total expectation:

- $E[n] = E[n|H_1] \cdot p + E[n|T_1] \cdot q$

# Total Expectation: How many heads?

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By total expectation:

- $E[n] = E[n|H_1] \cdot p + E[n|T_1] \cdot q$
- $E[n] = (1 + E[n - 1]) \cdot p + (E[n - 1]) \cdot q$

# Total Expectation: How many heads?

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By total expectation:

- $E[n] = E[n|H_1] \cdot p + E[n|T_1] \cdot q$
- $E[n] = (1 + E[n - 1]) \cdot p + (E[n - 1]) \cdot q$
- $E[n] = p + E[n - 1](p + q) = p + E[n - 1]$

# Total Expectation: How many heads?

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By total expectation:

- $E[n] = E[n|H_1] \cdot p + E[n|T_1] \cdot q$
- $E[n] = (1 + E[n - 1]) \cdot p + (E[n - 1]) \cdot q$
- $E[n] = p + E[n - 1](p + q) = p + E[n - 1]$
- $E[n] = p + E[n - 1] = p + p + E[n - 2] = 3p + E[n - 3] = \dots = np + E[0] = np$

# Mean Time To Failure

Suppose that a computer has a **chance to fail** of  $p$  every hour.

```
A problem has been detected and Windows has been shut down to prevent damage
to your computer.
```

```
If this is the first time you've seen this Stop error screen,
restart your computer. If this screen appears again, follow these steps:
```

```
Check to be sure you have adequate disk space. If a driver is
identified in the Stop message, disable the driver or check
with the manufacturer for driver updates. Try changing video
adapters.
```

```
Check with your hardware vendor for any BIOS updates. Disable
BIOS memory options such as caching or shadowing. If you need
to use Safe Mode to remove or disable components, restart your
computer, press F8 to select Advanced Startup options, and then
select Safe Mode.
```

```
Technical information:
```

```
*** STOP: 0x0000007E (0xFFFFFFFFC0000047, 0xFFFFF80002EB5B48)
```

How many hours does it take for the computer to fail on average?

# Mean Time To Failure

- Chance to fail in the first hour =  $p$
- Chance to fail in the second hour =  $qp$
- Chance to fail in the third hour =  $q^2p$
- ...
- Chance to fail in the  $n$ -th hour =  $q^{n-1}p$

## Geometric Distribution / Geometric Sum



# Mean Time To Failure

How many hours does it take for the computer to fail on average?

- $E[F] = \sum_{n>0} n \cdot \Pr[F = n]$
- $E[F] = \sum_{n>0} nq^{n-1}p = p \sum_{n>0} nq^{n-1}$
- $E[F] = p \cdot \frac{1}{(1-q)^2}$
- $E[F] = p \cdot \frac{1}{p^2} = \frac{1}{p}$

So, in average, the server will fail after  $\frac{1}{p}$  hours.

# Class Summary

- **Random Variable** is a total function that describes a numeric value for a set of probability events.
- **Probability Density Function** describes the probability of a Random Variable to have a certain value  $a$
- **Expectation** is the weighted average of a PDF over its domain.

# Next Class

- No Exercise Sheet for Next class!
- Sample Exam;
- Lecture Evaluation;

Please come!