

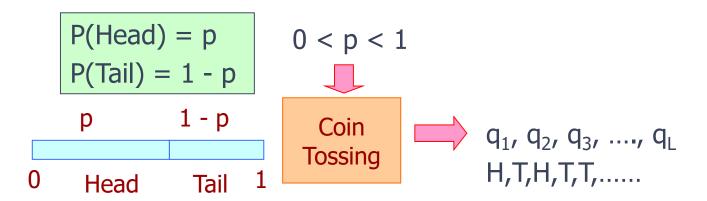
Observations from Random Process

- Observable output generated by realworld random process
- o Discrete vs. Continuous
 - Character, gender, class, ...
 - Speech, temperature,...

HIDDEN MARKOV MODEL

- A type of *stochastic model*
- First published in mathematic journals
 - Baum and his colleagues, late 1960s
- Applied to speech recognition
 - Baker at CMU/Jelinek at IBM, 1970s
- A general approach
 - Can be applied to many other areas

A SIMPLE STOCHASTIC PROCESS



- The outcomes $q_1, q_2,, q_T$ can be obtained through random tests
 - q_i has Two states: Head (H) and Tail (T)
- Each tossing is *independent* of the results of previous tossings
 - Random variables $q_1, q_2, ..., q_L$ are i.i.d.
- States (H,T): possible outcomes of random variables q_t

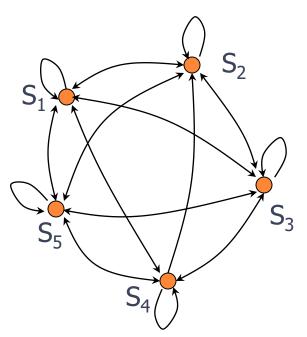
DICE TOSSING

- Fair dice: (1/6, 1/6, 1/6, 1/6, 1/6, 1/6) for all faces
- Every tossing is independent of the others
 - Current tossing does not depend on previous results
- Number of states: 6
 - $q_t = S_i$, i = 1, 2, ..., 6
 - q_t random variables
 - S_i outcome symbols (number of points)

DICE TOSSING (CONT'D)

- o 6 dices numbered from 1 to 6
 - With different distributions (probably unfair)
 - o Dice 1: (1/3, 1/10, ...)
 - o Dice 2: (1/4., 1/6, ...)
 - **O** . . .
 - Dice 6: (1/5, 1/8, ...)
 - The choice of dice depends on the outcome of previous tossing
 - E.g. if $q_t = S_4$ (4 points) in t_{th} tossing, then choose Dice 4 for $(t+1)_{th}$ tossing
 - q_t depends on q_{t-1}
 - $\circ q_1 \rightarrow q_2 \rightarrow q_3 \rightarrow \dots \rightarrow q_T$

DISCRETE MARKOV PROCESS



A system with 5 states $S_1 - S_5$.

- System is in one of N distinct states $(S_1, S_2, ..., S_N)$ at any time, and **state sequence** $q_1, q_2, ..., q_T$ is observed
- First-order Markov Chain

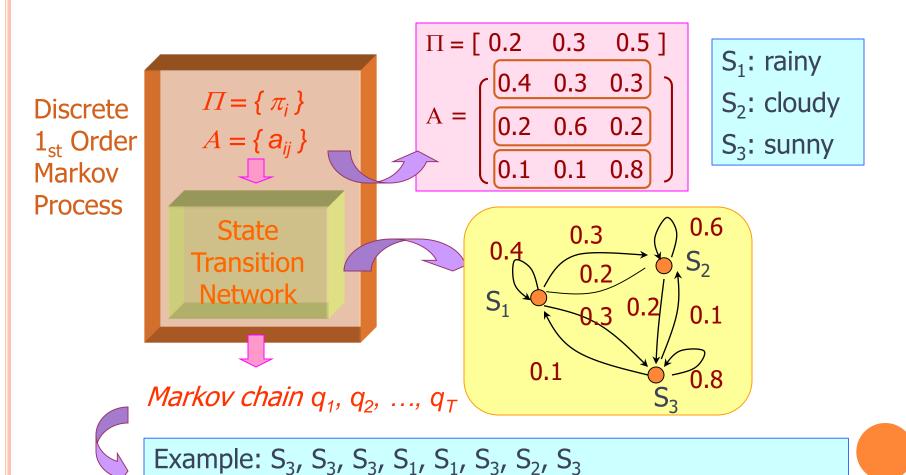
•
$$P(q_t=S_j|q_{t-1}=S_i, q_{t-2}=S_k...) = P(q_t=S_j|q_{t-1}=S_i)$$

•
$$a_{ij} \equiv P(q_t = S_j | q_{t-1} = S_i)$$
 $1 \le i, j \le N$
 $a_{ij} > 0, \ \Sigma_{j=1}^N a_{ij} = 1$

•
$$\pi_i \equiv P(q_1 = S_i)$$
 $1 \leq i \leq N$

- State transition depends on *only the previous state*
- State transits stochastically
- Imagine: N dices with N faces each

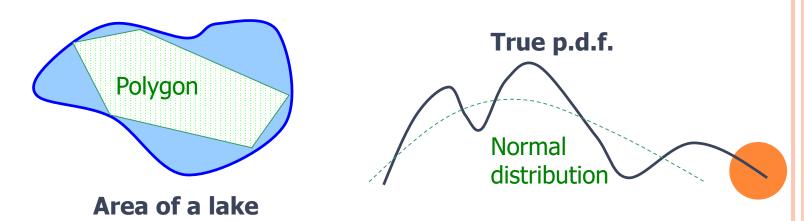
MARKOV PROCESS SIMULATOR



sunny, sunny, rainy, rainy, sunny, cloudy, sunny,...

Model vs. Reality

- \circ Model \neq Reality
 - Does God use Markov Process to generate weather patterns?
 - Real-world is usually *unknown* and thus *approximated* by some model
 - Amount of parameters may influence the accuracy



DISCUSSIONS

- Simulation
 - It is possible to simulate a (true) Markov process, and generate random observations accordingly
- Modeling the real world
 - In modeling the true random process in real world, we have no idea whether the true (unknown) random process is Markov process or not.
 - We simply model it!
 - The observations can be used to estimate the model parameters such that the model can best account for the statistics of the observations

Estimation of Model Parameters

 $\hat{a}_{ij} = \frac{c_{ij}}{\sum_k c_{ik}}$

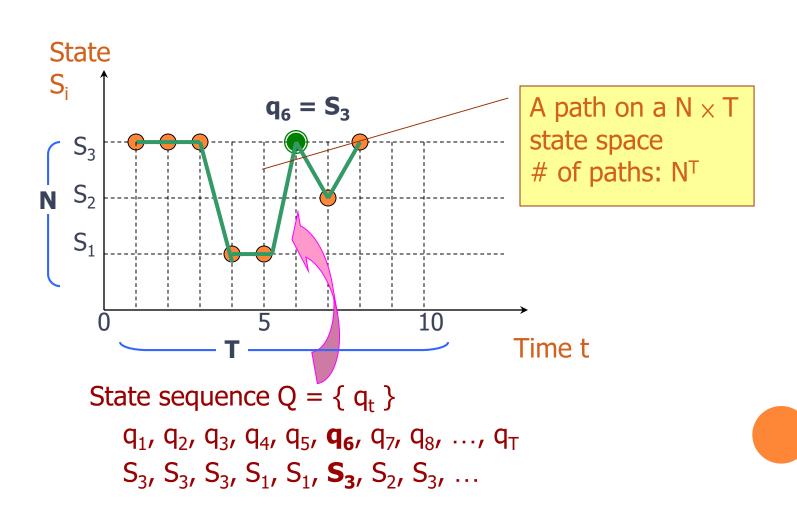
 c_{ik} is the occurrence count of transitions from state i to state k in all observation sequences

 $\dot{\pi}_j = \frac{c_j}{\sum_k c_k}$

 c_k is the occurrence count that the first state in a state sequence is state k for all observation sequences

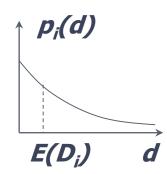
- Example
 - $S_1, S_2, S_2, S_3, S_1, S_1, S_3, ...$
 - c_{12} ++, c_{22} ++, c_{23} ++, c_{31} ++, c_{11} ++, c_{13} ++, ...

STATE DIAGRAM



PROBABILITY OF STATE SEQUENCE

- The probability of Markov model (Π, A) generating the state sequence $Q = S_3, S_3, S_3, S_1, S_1, S_2, S_3$
 - $P(Q|\Pi,A) = P(S_3) \cdot P(S_3|S_3) \cdot P(S_3|S_3) \cdot P(S_1|S_3) \cdot P(S_1|S_3) \cdot P(S_1|S_3) \cdot P(S_1|S_3) \cdot P(S_1|S_3) \cdot P(S_2|S_3) \cdot P(S_3|S_2) = \pi_3 \cdot a_{33} \cdot a_{33} \cdot a_{31} \cdot a_{11} \cdot a_{13} \cdot a_{32} \cdot a_{23}$



- System stays in S_i for time D_i (random variable)

 - Expectation $E(D_i) = \Sigma_d(d \cdot p_i(d)) = 1/(1-a_{ii})$ $a_{ii} \uparrow$, $E(D_i) \uparrow$
- Transition matrixes: modeling the duration

FROM MARKOV MODEL TO HMM

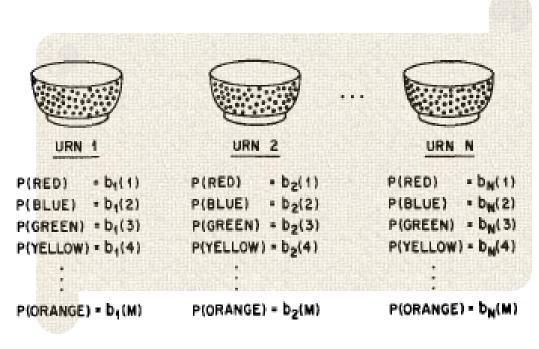
• Markov Model

• State transits stochastically and state sequence can be observed

o Hidden Markov Model

- State transits stochastically, but state sequence *cannot be observed*
- State sequence is unseen { q_t }
 rainy, cloudy, sunny
- Observation sequence { Ot }
 very hot (VH), hot(H), warm(W), cool(CL), cold(CD)

URN AND BALL MODEL



O . {GREEN, GREEN, BLUE, RED, YELLOW, RED,, BLUE }

- State transition
 - Roll the dice to select a urn
 - Multiple dices
- Observation
 - Pick up a ball from the selected urn
- The urns(j) has an associated distribution of balls, b_i(O_t)
- Double stochastic processes

ELEMENTS OF HMM

- N: number of states
 - $S_1, S_2, ..., S_N$ (outcome symbols, not random)
- Initial state distribution $\Pi = \{\pi_i\}$
 - $\pi_i = P(q_1 = S_i)$
- State transition distribution $A = \{a_{ij}\}$
 - $a_{ij} = P(q_{t+1} = S_i | q_t = S_i)$
- State observation distribution $B = \{b_j(O_t)\}$
 - $b_i(O_t) = P(O_t|q_t = S_i)$
 - The output O_t depends on the current state q_t
 - Could be discrete or continuous functions

$$\circ \lambda = (\Pi, A, B)$$

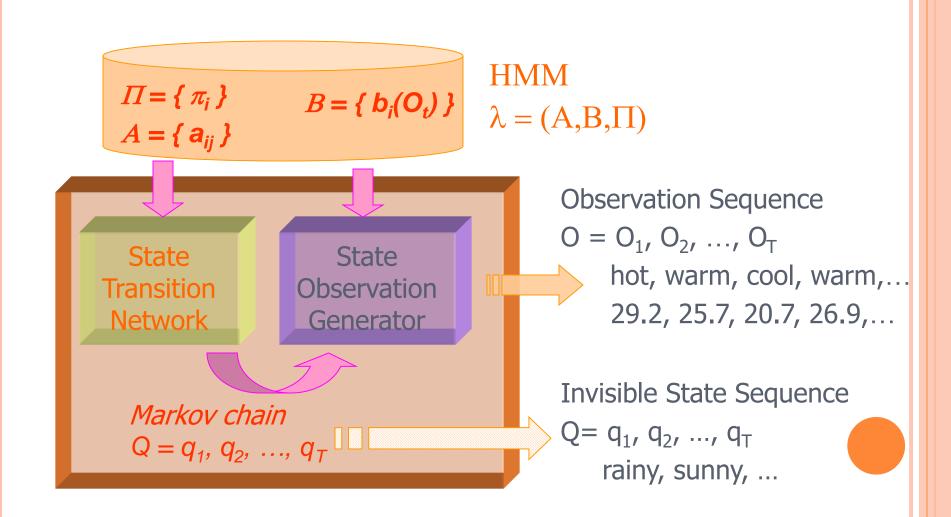
OBSERVATION DISTRIBUTION

- Discrete: Probability weighting function
 - $b_i(O_t = v_k)$ k= 1, 2, ..., K
 - $v_{k,s}$ may be obtained through encoding/quantization
 - Ex. $O_t = v_k$: very hot, hot, warm, cool, cold, very code
- o Continuous: Probability density function
 - $b_j(O_t=x)$ could be a continuous function in any region which satisfied $\int b_i(O_t=x)dx = 1$
 - Ex. temperature could be 34, 31.5, 28, 29.3,
 - Example: weighted Gaussian mixtures
 - $b_i(O_t=x) = \Sigma_{m=1} c_{im} \cdot \mathcal{N}(x; m_{im}, \sigma_{im}^2)$ $-\infty < x < \infty$

SEMI-CONTINUOUS OBSERVATION DISTR.

- Discrete : $b_j(O_t = v_k) = b_{jk}$
 - Quantization error, insufficient training (e.g. $b_{kj} = 0$)
- Continuous: $\mathcal{N}(m_{im}, \sigma_{im}^2)$ for state j & mixture m
 - Computation load is dependent of the numbers of models, states, or mixtures
 - 500 HMMs, 4 states, 10 mix → 20000 computations!
- Semi-Continuous : $b_j(O_t=x) = \Sigma_{k=1}b_{jk} \cdot \mathcal{N}(x;m_k,\sigma_k^2)$
 - Common Gaussian mixtures are shared among all states, but with all states have different mixture weights
 - Trade-off between discrete and continuous distribution
 - Can reduce the effect of quantization errors and insufficient training due to parameter sharing
 - m_k and σ_k are obtained from observations clustered to codeword ν_k (codebook size is fixed)

HMM SIMULATOR



CONCEPTS OF HIDDEN STATES

- Economic conditions
 - Hidden States: economic conditions
 - Observation: economic indexes
- Moods
 - Hidden States: mood states
 - Observation: color of dressing, face expression, interaction style
- Periods of life
 - Hidden States: stages of growth
 - Observation: weight, height, and other features
- Seasons
 - 4 seasons

TO EVALUATE THE LEARNED

- Given HMM λ and unknown observation O
 - Calculation of $P(O | \lambda)$
 - $i^* = argmax_i P(O | \lambda_i)$: to find the HMM that most probably generates O
- Use for Recognition
 - With observed sequence O and the HMMs for some λ_i 's, calculate $P(O \,|\, \lambda_i)$
 - Decide which model is most probable to generate the unknown observation O

TO LEARN FROM OBSERVATION

- We have the observation O, and *assume* it is generated by a HMM.

 - HMM λ^* can *best represent* the observation O
 - The stochastic characteristics of observation O are caught and kept in model parameters of λ^*
 - O $\rightarrow \lambda^* = (\Pi, A, B)$
- Example
 - Temperature sequence O \rightarrow HMM λ^*

TO GUESS THE UNSEEN STATES Q

- Assume observation O is generated by an HMM λ .
 - Is it possible to reconstruct the *state sequence* Q^* that is optimal in some sense?
- Example
 - Temperature (O) & HMM (λ) \rightarrow weather (Q*) • C, H, H, VH,... \rightarrow rainy, sunny, sunny, sunny, ...

FORMULATE PROBLEMS

- Evaluating: $P(O | \lambda)$
- o Learning: O → λ^*
- o Labeling: $O,λ → Q^*$
- \circ O = O₁, O₂, ..., O_T (random variables)
- \circ Q = q₁, q₂, ..., q_T (random variables)
- $\circ \lambda$: model parameters
- Important Algorithms
 - Forward-backward procedure
 - Expectation-Maximization (EM) algorithm
 - Viterbi algorithm

JOINT PROBABILITY P(O,Q)

•
$$P(Q|\lambda) = P(q_1) \cdot \Pi_{t=2 \sim T} P(q_t|q_t, \lambda)$$

= $\pi_{q_1} \cdot a_{q_1 q_2} \cdot a_{q_2 q_3} \cdot \dots \cdot a_{q_{T-1} q_T} (Q = q_1, q_2, \dots, q_T)$
• $P(O|Q,\lambda) = \Pi_{t=1 \sim T} P(O_t|q_t, \lambda)$
= $b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdot \dots \cdot b_{q_T}(O_T)$
• $P(O,Q|\lambda) = P(Q|\lambda) \cdot P(O|\lambda,Q)$
= $\pi_{q_1} \cdot b_{q_1}(O_1) \cdot a_{q_1 q_2} \cdot b_{q_2}(O_2) \cdot \dots \cdot a_{q_{T-1} q_T} \cdot b_{q_T}(O_T)$

ISSUE OF COMPUTATION COMPLEXITY

- $P(O|\lambda) = \Sigma_{Q} P(O,Q|\lambda)$ $= \Sigma_{Q} P(Q|\lambda) \cdot P(O|\lambda,Q)$ $= \Sigma_{Q} [\pi_{q_{1}} \cdot b_{q_{1}}(O_{1}) \cdot a_{q_{1}q_{2}} \cdot b_{q_{2}}(O_{2}) \cdot \dots a_{q_{T-1}q_{T}} \cdot b_{q_{T}}(O_{T})]$
- o $P(O|\lambda)$ is the summation of $P(O,Q|\lambda)$ over <u>ALL paths</u> on the state space
 - Note: path Q is *hidden* in HMM (cannot be seen)
 - $P(X) = \Sigma_Y P(X, Y)$
- On N × T state space, there are totally N^T paths
 - Infeasible if every path is calculated independently
 - (2T-1)*N^T multiplications and N^T-1 additions

FORWARD PROCEDURE

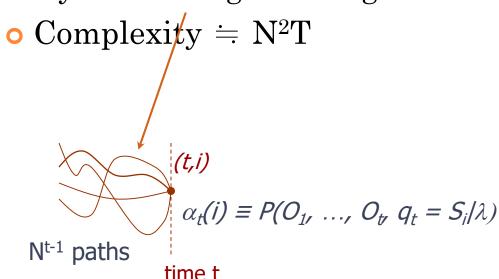
$$\begin{array}{l} \bullet \ \alpha_{t}(j) \equiv P(O_{1}, \ ..., \ O_{t}, \ q_{t} = S_{j} | \lambda) \\ \bullet \ \alpha_{1}(j) = P(O_{1}, \ q_{1} = S_{j} | \lambda) \\ = P(q_{1} = S_{j} | \lambda) \cdot P(O_{1} | \lambda, \ q_{1} = S_{j}) \\ = \pi_{j} \cdot b_{j}(O_{1}) \\ \bullet \ \alpha_{t+1}(j) \equiv P(O_{1}, \ ..., \ O_{t+1}, \ q_{t+1} = S_{j} | \lambda) \\ = P(O_{1}, \ ..., \ O_{t}, \ q_{t+1} = S_{j} | \lambda) \cdot P(O_{t+1} | \lambda, \ \bigcirc_{\gamma}, \ ..., \ \bigcirc_{p} q_{t+1} = S_{j}) \\ = [\Sigma_{i} P(O_{1}, \ ..., \ O_{t}, \ q_{t} = S_{i}, \ q_{t+1} = S_{j} | \lambda)] \cdot b_{j}(O_{t+1}) \\ = \{\Sigma_{i} [P(O_{1}, \ ..., \ O_{t}, \ q_{t} = S_{j} | \lambda) \cdot P(q_{t+1} = S_{j} | \lambda, \ \bigcirc_{\gamma}, \ ..., \ \bigcirc_{p} q_{t} = S_{i})]\} \cdot b_{j}(O_{t+1}) \\ = [\Sigma_{i} (\alpha_{t}(i) \cdot \alpha_{i})] \cdot b_{j}(O_{t+1}) \end{array}$$

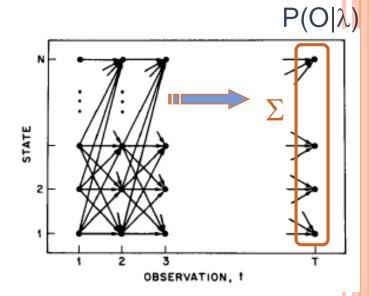
$\circ P(O|\lambda) = \Sigma_i \alpha_T(i)$

Stochastic Dependency

FORWARD PROCEDURE (CONT'D)

- Calculate $\alpha_t(i)$ for every *grid location* (t,i) on state space in left-to-right direction
- $\alpha_t(i)$ is the <u>temporary summation</u> of $P(O, Q|\lambda)$ over <u>all partial paths terminating at (t, i)</u>
- Dynamic Programming

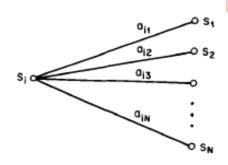




BACKWARD PROCEDURE

$$\begin{split} \bullet \ \ \beta_{t}(i) & \equiv P(O_{t+1}, \ ..., \ O_{T} \ | q_{t} = S_{i}, \lambda) \\ & = \Sigma_{j} \ P(O_{t+1}, \ ..., \ O_{T}, \ q_{t+1} = S_{j} \ | \ q_{t} = S_{i}, \lambda) \\ & = \Sigma_{j} \ \{ \ P(q_{t+1} = S_{j} \ | \ q_{t} = S_{i}, \lambda) \cdot P(O_{t+1}, \ ..., \ O_{T} \ | \ q_{t} = S_{j}, q_{t+1} = S_{j}, \lambda) \} \\ & = \Sigma_{j} \ \{ \ a_{ij} \cdot [P(O_{t+1} \ | \ q_{t+1} = S_{j}, \lambda) \cdot P(O_{t+2}, \ ..., \ O_{T} \ | \ q_{t+1} = S_{j}, \ O_{t+1}, \lambda)] \} \\ & = \Sigma_{j} \ [\ a_{ij} \cdot b_{j}(O_{t+1}) \cdot \beta_{t+1} \ (j) \] \\ & \bullet \ \beta_{T-1}(i) = P(O_{T} \ | \ q_{T-1} = S_{i}, \lambda) = \Sigma_{j} \ P(O_{T}, \ q_{T} = S_{j}, \ | \ q_{T-1} = S_{i}, \lambda) \\ & = \Sigma_{j} \ [P(q_{T} = S_{j}, \ | \ q_{T-1} = S_{i}, \lambda) \cdot P(O_{T} \ | \ q_{T} = S_{j}, \lambda)] \\ & = \Sigma_{i} \ [a_{ij} \cdot b_{i}(O_{T})] \end{split}$$

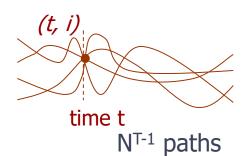
- It is there appropriate to assign $\beta_T(i) \equiv 1$ such that the inductive formula holds for t = T-1
- $\beta_t(i)$ is calculated for every grid point (t,i) in right-to-left direction



FIND OPTIMAL STATES Q_T*

- $P(O, q_t = S_i | \lambda) = P(O_1, ..., O_t, O_{t+1}, ..., O_T, q_t = S_i | \lambda)$ = $P(O_1, ..., O_t, q_t = S_i | \lambda) \cdot P(O_{t+1}, ..., O_T | q_t = S_i, O_{j,..., O_t} \lambda)$ = $\alpha_t(i) \cdot \beta_t(i)$
 - Summation of P(O,Q|λ) over <u>all paths passing (t, i)</u>

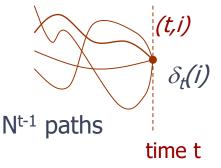
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q_t^* \equiv \operatorname{argmax}_i \{ P(q_t = S_i | O, \lambda) \}
= \operatorname{argmax}_i \{ P(O, q_t = S_i | \lambda) / P(O | \lambda) \}
= \operatorname{argmax}_i \{ \gamma_t(i) \}
\bullet \ \gamma_t(i) \equiv P(q_t = S_i | O, \lambda) = \alpha_t(i) \cdot \beta_t(i) / P(O | \lambda)
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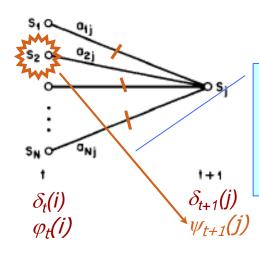
- o $q_{t's}^*$ are the states optimized for each time individually
- o $P(q_1^*, q_2^*, ..., q_T^*, O|\lambda)$ is NOT the path with highest probability $P(Q, O|\lambda)$ among all paths Q's

FIND OPTIMAL PATH - VITERBI ALGORITHM

- o $P^* = \max_{Q} P(O, Q|\lambda)$, $Q^* = \operatorname{argmax}_{Q} P(O, Q|\lambda)$ Q^* is the path with the highest probability P^*
- o $\delta_t(i) \equiv \max_{q_1,q_2,...,q_t=S_i} P(O_1, ..., O_t, q_t=S_i \mid \lambda)$
- o $\delta_t(i)$ is the <u>temporary maximum</u> of $P(O, Q|\lambda)$ for <u>all partial paths terminating at (t, i)</u>



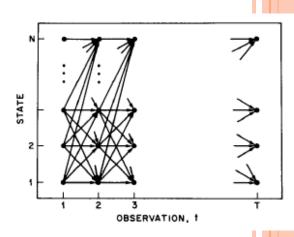
- $\delta_t(j) = \pi_j \cdot b_j(O_1), \ \varphi_t(j) = 0$
- $\delta_{t+1}(j) = \max_{i} \left[\delta_{t}(i) \cdot a_{ij} \right] \cdot b_{j}(O_{t+1})$ $\psi_{t+1}(j) = \operatorname{argmax}_{i} \left[\delta_{t}(i) \cdot a_{ij} \right]$
- $P^* = \max_i [\delta_T(i)]$ $q_T' = \operatorname{argmax}_i [\delta_T(i)]$ $q_t' = \varphi_{t+1}(q_{t+1}') \text{ backtracking}$ $Q^* = q_1', q_2', ..., q_T'$



At any time, only the winner can survive and be memorized

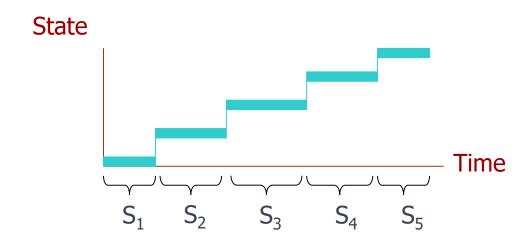
VITERBI ALGORITHM (CONT'D)

- Take *log*, the multiplications becomes additions
 - $\delta'_{t+1}(j) = \max_{i} [\delta'_{t}(i) + \log(a_{ij})] + \log[b_{j}(O_{t+1})]$ $\psi_{t+1}(j) = \operatorname{argmax}_{i} [\delta'_{t}(i) + \log[a_{ij})]$
 - Avoid underflow for long observation sequence
- Concept analogy
 - To pick up the maximum points on the street
 - Can proceed in directions \rightarrow and \triangleright
 - Get log(a_{ii}) points by transit from i to j
 - Get log[b_i(O_t)] points on the cross (t, j)



VITERBI ALGORITHM (CONT'D)

- For left-to-right HMM, the optimal path Q* obtained from Viterbi algorithm can segment the observation sequence into different states
- o Can be used for automatic segmentation

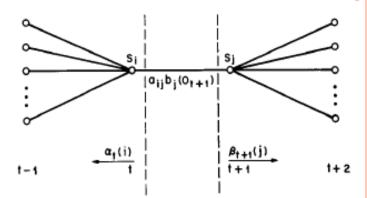


EM ALGORITHM

- Expectation-Maximization (8.30)
- Find λ^* by iterative procedure
 - $P(O | \lambda') > P(O | \lambda)$ and can find local maximum
 - Converge when the change of likelihood is small
- $\gamma_t(i) \equiv P(q_t = S_i | O, \lambda)$ the contribution of O_t for state i
- $\varepsilon_t(i, j) \equiv P(q_t = S_i, q_{t+1} = S_i | O, \lambda)$

$$\gamma_t(i) = \frac{\alpha_t(i) \ \beta_t(i)}{P(O|\lambda)}$$

$$\xi_t(i,j) = \frac{\alpha_t(i) \ a_{ij}b_j(O_{t+1}) \ \beta_{t+1}(j)}{P(O|\lambda)}$$



EM ALGORITHM (CONT'D)

o Baum-Welch Reestimation Procedure

$$\overline{\pi}_i = \text{expected frequency (number of times) in state } S_i \text{ at time } (t=1) = \gamma_1(i)$$

$$\overline{a}_{ij} = \frac{\text{expected number of transitions from state } S_i \text{ to state } S_j}{\text{expected number of transitions from state } S_i}$$

$$= \frac{\sum\limits_{t=1}^{T-1} \xi_t(i,j)}{\sum\limits_{t=1}^{T-1} \gamma_t(i)}$$

$$\sum\limits_{t=1}^{T-1} \gamma_t(i)$$

$$\overline{b}_{j}(k) = \frac{\text{expected number of times in state } j \text{ and observing symbol } v_{k}}{\text{expected number of times in state } j}$$

$$= \frac{\sum\limits_{t=1}^{T} \gamma_{t}(j)}{\sum\limits_{t=1}^{T} \gamma_{t}(j)}. \quad \text{For Discrete state observation distribution}$$

$$= \frac{\sum\limits_{t=1}^{T} \gamma_{t}(j)}{\sum\limits_{t=1}^{T} \gamma_{t}(j)}. \quad b_{j}(O_{\underline{t}} = \underline{v}_{k})$$

EM ALGORITHM (CONT'D)

For Continuous state observation distribution

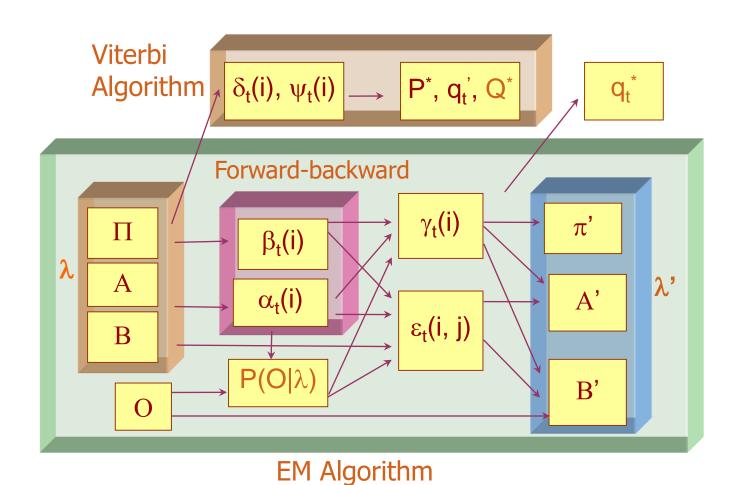
$$b_{j}(O_{t}=x) = \Sigma_{k} c_{jk} \cdot N(x; \underline{\mu}_{jk}, U_{jk})$$

$$\bar{c}_{jk} = \frac{\sum\limits_{t=1}^{T} \gamma_t(j,k) \cdot \mathbf{O}_t}{\sum\limits_{t=1}^{T} \sum\limits_{k=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{T} \gamma_t(j,k) \cdot \mathbf{O}_t}{\sum\limits_{t=1}^{T} \sum\limits_{k=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{T} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})'}{\sum\limits_{t=1}^{T} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{T} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})'}{\sum\limits_{j=1}^{T} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j) \beta_t(j)}{\sum\limits_{j=1}^{N} \alpha_t(j) \beta_t(j)} = \frac{\sum\limits_{m=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{m=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j) \beta_t(j)}{\sum\limits_{j=1}^{N} \alpha_t(j) \beta_t(j)} = \frac{\sum\limits_{m=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{m=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j) \beta_t(j)}{\sum\limits_{j=1}^{N} \alpha_t(j) \beta_t(j)} = \frac{\sum\limits_{m=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{m=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{j=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{j=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{j=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{j=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{j=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k) \cdot (\mathbf{O}_t - \mathbf{\mu}_{jk})(\mathbf{O}_t - \mathbf{\mu}_{jk})}{\sum\limits_{t=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k)}{\sum\limits_{j=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k)}{\sum\limits_{t=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k)}{\sum\limits_{t=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k)}{\sum\limits_{t=1}^{N} \gamma_t(j,k)} = \frac{\sum\limits_{t=1}^{N} \gamma_t(j,k)}{\sum\limits_{t=1}$$

belongs to state j

mixture k if in state j

DEPENDENCY DIAGRAM



HMM FOR PATTERN RECOGNITION

- Assume there are M patterns to be recognized, and we have observations for these patterns. Then, we can train the model λ_i for i-th pattern using the observations of this pattern
- $oi^* = argmax_i P(O | \lambda_i) i=1, 2, ..., M$

TOPOLOGY OF HMM

$$A = \begin{bmatrix} a_{11} & a_{12} & a_{13} & a_{14} \\ a_{21} & a_{22} & a_{23} & a_{24} \\ a_{31} & a_{32} & a_{33} & a_{34} \\ a_{41} & a_{42} & a_{43} & a_{44} \end{bmatrix}.$$

$$A = \begin{bmatrix} a_{11} & a_{12} & a_{13} & 0 \\ 0 & a_{22} & a_{23} & a_{24} \\ 0 & 0 & a_{33} & a_{34} \\ 0 & 0 & 0 & a_{44} \end{bmatrix}.$$

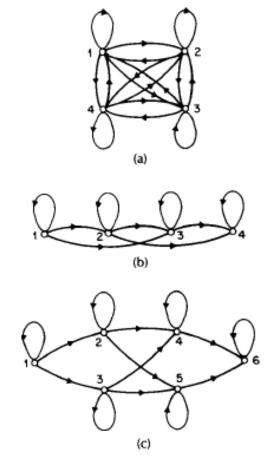


Fig. 7. Illustration of 3 distinct types of HMMs. (a) A 4-state ergodic model. (b) A 4-state left-right model. (c) A 6-state parallel path left-right model.

SCALING ISSUE

• If the length of observation sequence is long, there may arise the problem of underflow since all probabilities are less than 1.

- $P(O \mid \lambda) = \Sigma_Q \pi_{q_1} \cdot b_{q_1}(O_1) \cdot a_{q_1 q_2} \cdot b_{q_2}(O_2) \cdot \dots \cdot a_{q_{T-1} q_T} \cdot b_{q_T}(O_T)$
- $S_t = 1/\Sigma_i \ \alpha_t(i)$ is used for scaling $\alpha_t(i)$ and $\beta_t(i)$ at time t in Forward-backward procedure
 - It can be proved scaling will NOT influence the reestimation formula
- In Viterbi algorithm, probability is taken "log" and then accumulated

MULTIPLE OBSERVATION SEQUENCES

$$\overline{a_{ij}} = \frac{\sum_{k=1}^{K} \frac{1}{P_k} \sum_{t=1}^{T_k-1} \alpha_t^k(i) \ a_{ij} b_j(O_{t+1}^{(k)}) \ \beta_{t+1}^k(j)}{\sum_{k=1}^{K} \frac{1}{P_k} \sum_{t=1}^{T_k-1} \alpha_t^k(i) \ \beta_t^k(i)}$$

$$\overline{b_j}(\ell) = \frac{\sum_{k=1}^{K} \frac{1}{P_k} \sum_{t=1}^{T_k-1} \alpha_t^k(i) \ \beta_t^k(i)}{\sum_{k=1}^{K} \frac{1}{P_k} \sum_{t=1}^{T_k-1} \alpha_t^k(i) \ \beta_t^k(i)}$$

- The Reestimation formula can be modified for multiple observation sequences
- $P_k = P(O^{(k)} | \lambda)$

INITIAL MODEL

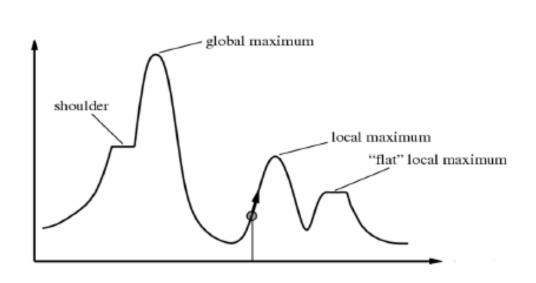
- Good initial estimate is important for avoiding local maximum
- \circ Random or uniform distributed parameters are adequate for Π and A, but not enough for B.
- To obtain better initial estimate
 - Manual segmentation
 - Segmental K-means (Viterbi segmentation)

LIMITATION OF HMM

- First-Order Assumption
 - Markov Chain
 - Leading to exponential duration model
- Conditional Independence Assumption
 - The Observation depends ONLY on the CURRENT state
 - Maximum entropy Markov model (MEMM)
 - Conditional random field (CRF)

LIMITATION OF HMM (CONT'D)

 Local optimum guaranteed by EM algorithm

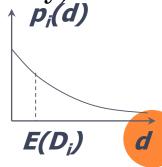


GENERALIZATION OF MM/HMM

- Observation O can be vector instead of scalar
- Concept of time can be generalized as other physical dimension
 - Example: x-axis in Euclidean space
 - State sequence $Q = \{ q_t \}$
 - Observation sequence $O = \{O_t\}$
 - Recognition of patterns which vary with some dimension (e.g. image, gene pattern, ...)

GENERALIZATION OF MM/HMM

- State can be flexibly defined
 - Number of state might be very large
- Better duration model
 - Use Gaussian or Gamma function as duration distribution instead of exponential function (a_{ij} is exponential duration distribution effectively)
 - Slight improvement achieved



APPLICATIONS OF HMM

- Statistical Model for *Trajectory*
- Speech recognition/synthesis
 - Spectrum trajectory
- Image recognition
 - row trajectory
- Natural language processing
 - Word segmentation / POS tagging
- Gesture recognition
 - 2D/3D trajectory
 - Action recognition (multi-point trajectory)
- Melody recognition/synthesis
 - Note trajecotry