Lab 4

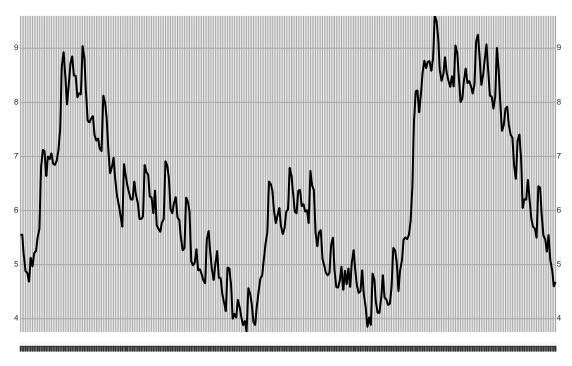
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```
library(ggplot2)
library(forecast)
## Warning: package 'forecast' was built under R version 3.4.3
## Attaching package: 'forecast'
## The following object is masked from 'package:ggplot2':
##
##
       autolayer
library(reshape2)
library(xts)
## Warning: package 'xts' was built under R version 3.4.3
## Loading required package: zoo
## Warning: package 'zoo' was built under R version 3.4.1
## Attaching package: 'zoo'
## The following objects are masked from 'package:base':
##
       as.Date, as.Date.numeric
##
library(tseries)
## Warning: package 'tseries' was built under R version 3.4.3
setwd('C:/Users/Samir/Documents/MIDS/StatsF17/lab 4/')
lab4data <- read.csv('Lab4-series2.csv')</pre>
```

EDA

xts_x 1990–01–01 / 2015–11–01

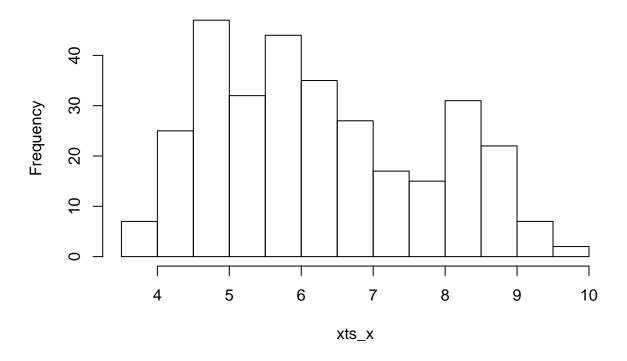


Jan 1990 Jan 1992 Jan 1994 Jan 1996 Jan 1998 Jan 2000 Jan 2002 Jan 2004 Jan 2006 Jan 2008 Jan 2010 Jan 2012 Jan 2014

At first glance, appears to be very non-stationary with strong seasonal trends. Seasonal trends appear to be yearly.

hist(xts_x)

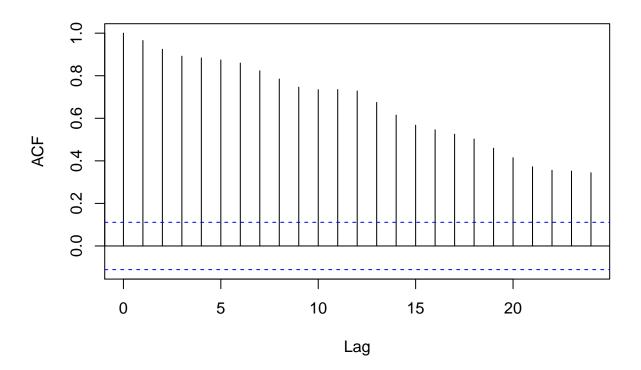
Histogram of xts_x



Histogram doesn't seem to suggest any log transformations are necessary.

acf(xts_x)

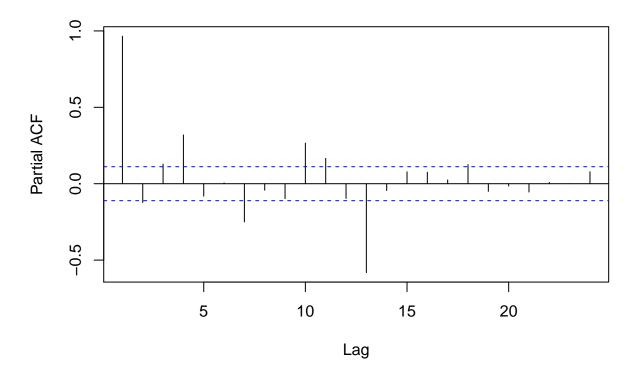
Series xts_x



ACF has a gradual decline suggesting an AR model with p of at least 1 would be usful.

pacf(xts_x)

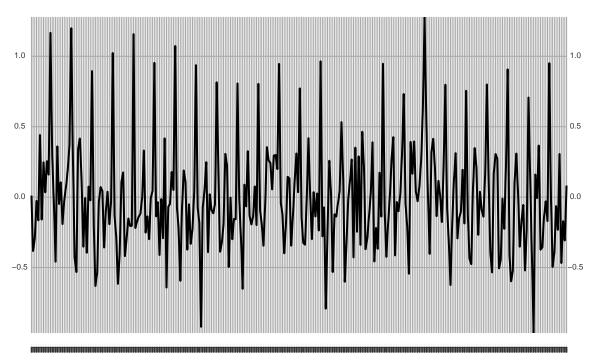
Series xts_x



The strongest PACF occurs at lag 13. (Does this confirm the yearly seasonal trend? I would have expected the significance at lag = 12) There are quite a few other significant PACFs at different lags which suggest non-seasonal, so we will have to test many different parameters.

```
xts_x.diff = diff(xts_x)
xts_x.diff <- xts_x.diff[!is.na(xts_x.diff)]
plot(xts_x.diff)</pre>
```

xts_x.diff 1990-02-01 / 2015-11-01



Feb 1990 Feb 1992 Jan 1994 Jan 1996 Jan 1998 Jan 2000 Jan 2002 Jan 2004 Jan 2006 Jan 2008 Jan 2010 Jan 2012 Jan 2014

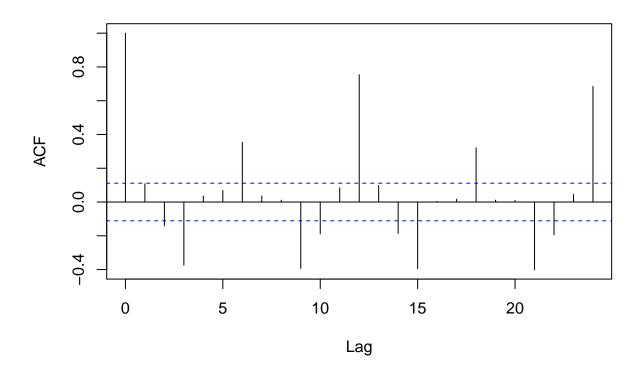
```
adf.test(xts_x.diff)
```

```
## Warning in adf.test(xts_x.diff): p-value smaller than printed p-value
##
## Augmented Dickey-Fuller Test
##
## data: xts_x.diff
## Dickey-Fuller = -5.6045, Lag order = 6, p-value = 0.01
## alternative hypothesis: stationary
```

Plotting the first difference makes it stationary, which is seen from the plot and confirmed by the significance ADF test. But there are clear spikes from the seasonal trend we have to account for.

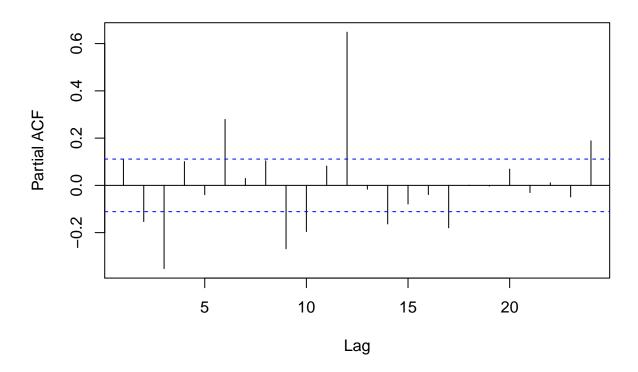
```
acf(xts_x.diff)
```

Series xts_x.diff



pacf(xts_x.diff)

Series xts x.diff



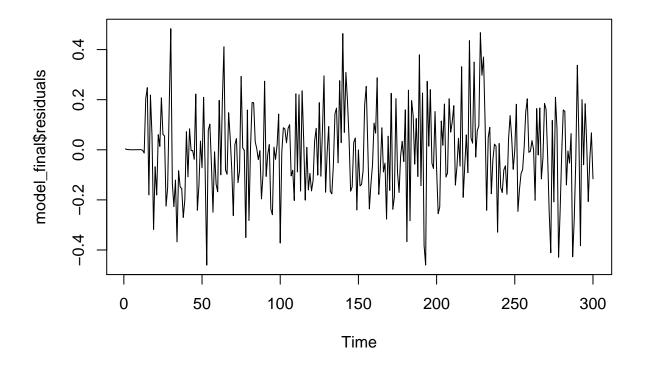
Both the ACF and PACF for the differences series show a huge spike at lag 12 confirming the yearly seasonal trend.

Loop to find optimal ARIMA

```
x_train <- lab4data[1:300,]$x</pre>
x_test <- lab4data[301:311,]$x</pre>
results = data.frame(p=NA, d=NA, q=NA, P=NA, D=NA, Q=NA, AIC=NA, RMSE=NA, MAPE=NA)
upperLimit = 2
start = Sys.time()
for(p in 0:upperLimit){
  for(d in 1:1){
    for(q in 0:upperLimit){
      for(P in 0:upperLimit){
        for(D in 1:1){
          for(Q in 0:upperLimit){
tryCatch({Arima.out <- Arima(x_train,</pre>
                   order = c(p,d,q),
                    seasonal = list(order=c(P,D,Q), period=12))
AIC <- Arima.out$aic
s <- as.data.frame(summary(Arima.out))</pre>
RMSE <- s$RMSE
MAPE <- s$MAPE
```

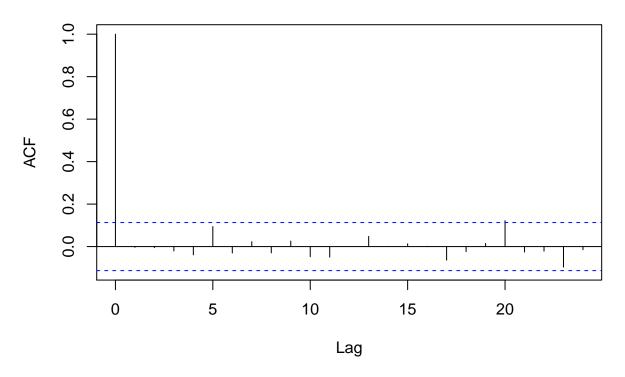
```
}, warning = function(w){
  AIC <- NA
  RMSE <- NA
  MAPE <- NA
  }, error = function(e){
  AIC <- NA
  RMSE <- NA
 MAPE <- NA
})
result <- data.frame(p=p, d=d, q=q, P=P, D=D, Q=Q, AIC=AIC, RMSE=RMSE, MAPE=MAPE)
results <- rbind(results, result)</pre>
        }
     }
    }
  }
}
results <- results[2:nrow(results),]
end = Sys.time()
end-start
results[results$AIC==min(results$AIC, na.rm=T),]
##
     pdqPDQ
                        AIC
                                 RMSE
                                         MAPE
## 73 2 1 1 2 1 2 -129.3362 0.1753747 2.28454
results[results$RMSE==min(results$RMSE, na.rm=T),]
     pdqPDQ
                        AIC
                                 RMSE
                                          MAPE
## 64 2 1 0 2 1 2 -121.4514 0.1743903 2.291954
results[results$MAPE==min(results$MAPE, na.rm=T),]
      pdqPDQ
                       AIC
                                RMSE
                                         MAPE
## 55 1 1 2 2 1 2 -129.156 0.1755049 2.283458
Using minimum AIC/RMSE/MAPE all gives different answers...
model final <- Arima.out <- Arima(lab4data[1:300,]$x,
                  order = c(2,1,1),
                   seasonal = list(order=c(2,1,2), period=12))
summary(model_final)
## Series: lab4data[1:300, ]$x
## ARIMA(2,1,1)(2,1,2)[12]
##
## Coefficients:
##
            ar1
                                                               sma2
                    ar2
                                             sar2
                                                      sma1
                             ma1
                                     sar1
##
         0.6849 0.1652 -0.6501
                                 -0.8035
                                           0.1778
                                                   -0.0008
                                                            -0.8962
## s.e. 0.0902 0.0633
                        0.0844
                                  0.0755 0.0739
                                                    0.0839
                                                             0.0751
## sigma^2 estimated as 0.03295: log likelihood=72.67
## AIC=-129.34 AICc=-128.82 BIC=-100.06
##
## Training set error measures:
##
                          ME
                                  RMSE
                                             MAE
                                                         MPE
                                                                MAPE
                                                                         MASE
```

```
## Training set -0.008473527 0.1753747 0.1367246 -0.09983992 2.28454 0.485703
## ACF1
## Training set -0.003564555
plot(model_final$residuals)
```



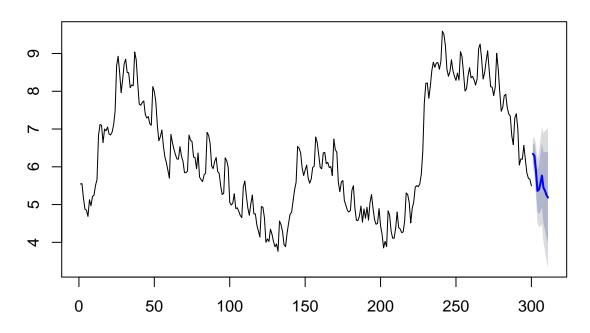
acf(model_final\$residuals)

Series model_final\$residuals



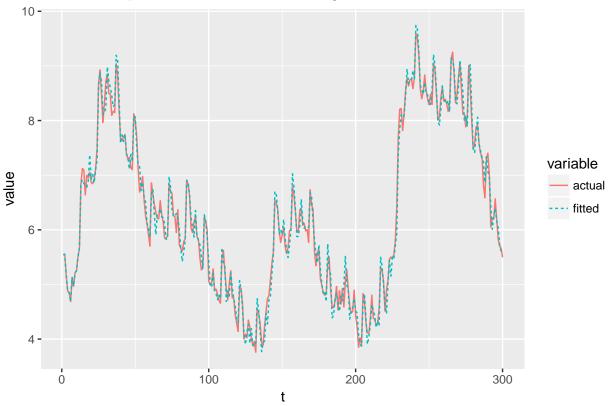
```
adf.test(model_final$residuals)
## Warning in adf.test(model_final$residuals): p-value smaller than printed p-
## value
##
## Augmented Dickey-Fuller Test
##
## data: model_final$residuals
## Dickey-Fuller = -6.1841, Lag order = 6, p-value = 0.01
## alternative hypothesis: stationary
The residuals of the final model, based on the plot, ACF, and ADF test, are stationary.
model_forecast <- forecast(model_final, h = 11)
plot(model_forecast)</pre>
```

Forecasts from ARIMA(2,1,1)(2,1,2)[12]



```
predicted <- as.numeric(model forecast$mean)</pre>
actual <- lab4data[301:311,]$x</pre>
mean(abs((actual-predicted)/actual) * 100)
## [1] 5.504884
model_final_df <- data.frame(actual = x_train, fitted = as.numeric(model_final$fitted))</pre>
df_melt <- cbind(melt(model_final_df), rbind(cbind(1:300), cbind(1:300)))</pre>
## No id variables; using all as measure variables
head(df_melt)
     variable value rbind(cbind(1:300), cbind(1:300))
##
## 1
       actual 5.544
       actual 5.555
                                                       2
## 3
       actual 5.172
                                                       3
## 4
       actual 4.878
## 5
       actual 4.851
       actual 4.686
colnames(df_melt) <- c("variable", "value", "t")</pre>
df_melt$variable <- factor(df_melt$variable, levels=c("actual", "fitted"))</pre>
ggp <- ggplot(df_melt, aes(x=t, y=value, group=variable, color=variable))</pre>
ggp+geom_line(aes(linetype=variable))+
  ggtitle("Actual and predicted values for training set")
```

Actual and predicted values for training set



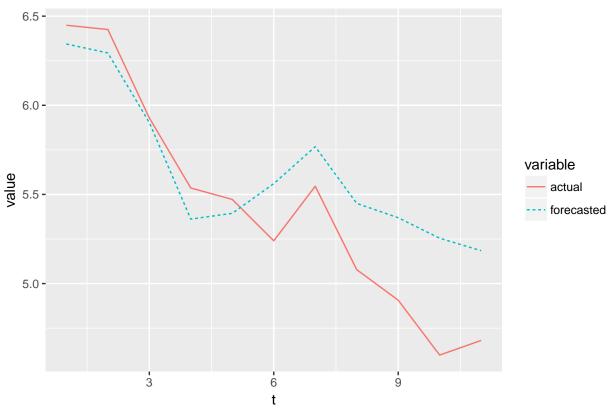
 $\mbox{\tt \#\#}$ No id variables; using all as measure variables

```
head(df_melt)
```

```
colnames(df_melt) <- c("variable", "value", "t")
df_melt$variable <- factor(df_melt$variable, levels=c("actual", "forecasted"))

ggp <- ggplot(df_melt, aes(x=t, y=value, group=variable, color=variable))
ggp+geom_line(aes(linetype=variable))+
ggtitle("Actual and forecasted values for test set")</pre>
```

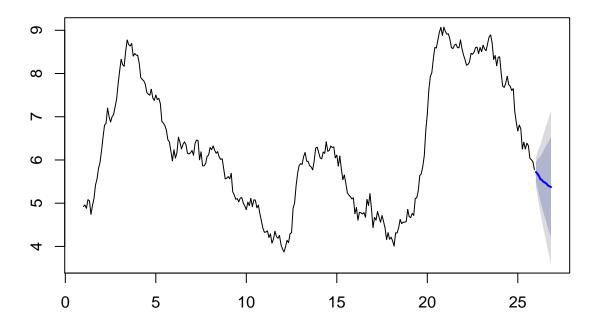
Actual and forecasted values for test set



```
ts2 <- ts(lab4data[0:300,]$x, frequency=12)
dsts2 <- seasadj(stl(ts2, s.window="periodic"))</pre>
auto.arima.out <- auto.arima(dsts2, seasonal=T)</pre>
summary(auto.arima.out)
## Series: dsts2
## ARIMA(1,1,2)(1,0,0)[12]
##
## Coefficients:
                             ma2
##
            ar1
                                     sar1
                     ma1
         0.8714 -0.8421 0.1377 0.0777
##
## s.e. 0.0604 0.0813 0.0591 0.0602
## sigma^2 estimated as 0.03117: log likelihood=96.08
## AIC=-182.16 AICc=-181.95 BIC=-163.66
##
## Training set error measures:
##
                                                         MPE
                                                                           MASE
                          ME
                                  RMSE
                                              MAE
                                                                 MAPE
## Training set 0.0003683998 0.1750703 0.1388201 0.01981815 2.326253 0.171446
##
## Training set -0.0006323352
model_forecast <- forecast(auto.arima.out, h = 11)</pre>
```

plot(model_forecast)

Forecasts from ARIMA(1,1,2)(1,0,0)[12]



The autoarima function with seasonal decomposition has a much lower AIC, suggesting better in-sample fit, but the forecast looks bad. The coefficient of the seasonal component in the arima output is very small and not significant which probably explains the bad fit... not worth it?