An Introduction to Bayesian Statistics

David Carlson

May 16, 2022

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- 4) Carry out the investigation or experiment
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- 6) Repeat 3—5 as necessary

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 - Experimental design, survey sampling

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 - Statistical inference
- Bayesian statistics is particularly well-suited for step 2 and 5
 - Step 2: Prior information
 - Step 5: Prior information → posterior information

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Economic Applications of Bayesian Statistics

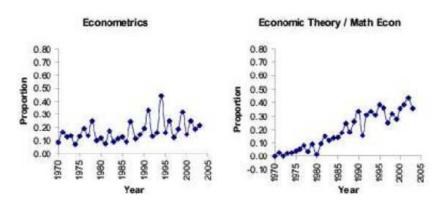


Figure 5: Econometrica Containing "Bayes" or "Bayesian"

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Economics

- Economics
- Marketing

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- Marketing
- Education

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- Economics
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- The list goes on...

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- No notion of infinitely repeating an event of interest
- Using as much prior information as possible as well as personal judgment (placing a bet)
- Once an outcome is revealed, prior information is updated

Table of Frequentist vs. Bayesian Interpretations

	Frequentist statistics	Bayesian statistics
Definition of the p value	The probability of observing the same or more extreme data assuming that the null hypothesis is true in the population	The probability of the (null) hypothesis
Large samples needed?	Usually, when normal theory-based methods are used	Not necessarily
Inclusion of prior knowledge possible?	No	Yes
Nature of the parameters in the model	Unknown but fixed	Unknown and therefore random
Population parameter	One true value	A distribution of values reflecting uncertainty
Uncertainty is defined by	The sampling distribution based on the idea of infinite repeated sampling	Probability distribution for the population parameter
Estimated intervals	Confidence interval: Over an infinity of samples taken from the population, 95% of these contain the true population value	Credibility interval: A 95% probability that the population value is within the limits of the interval

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- Frequentist: 95 of 100 replications of same experiment capture the fixed but unknown parameter
- Bayesian: Probability that a parameter lies in the credible interval

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- Prior reflects knowledge about parameters before observing current data
- Science can be accumulative!



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Bayes' Theorem

$$P(A|B) = \frac{P(B|A) \times P(A)}{P(B)}$$



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- Enzyme-Linked Immuno Sobert Assay (ELISA)
 - Diagnosis tool for HIV
 - Low cost
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- Let $A = \{\text{the patient is positive}\}$. Is it proper to use Pr(A)?
- How should we update the uncertainty after a test?

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• Data: Test result (positive/negative)

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- Update the uncertainty using Bayes' theorem

$$P(A|+) = \frac{P(+|A)P(A)}{P(+|A)P(A) + P(+|\neg A)P(\neg A)}$$

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- False negative rate: P(A|-) = 0.00052

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$$f(\theta|y) = \frac{f(\theta, y)}{f(y)}$$
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Bayesian

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 - ★ If we switch the role of two hypotheses, frequentists cannot compute p-value
 - ► Bayesian: No trouble

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Disadvantages of Bayesian



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 - Can easily incorporate prior information
 - Inferences are conditional on the actual data
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 - In principle, all problems can be solved

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- Empirical Bayes



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- Scale invariant prior for scale family

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Beta-binomial model



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Some Common Conjugate Priors

Likelihood	Prior	Posterior
$X \theta \sim \mathcal{N}(\theta, \sigma^2)$	$\theta \sim \mathcal{N}(\mu, \tau^2)$	$\theta X \sim \mathcal{N}(\frac{\tau^2}{\sigma^2 + \tau^2} X + \frac{\sigma^2}{\sigma^2 + \tau^2} \mu, \frac{\sigma^2 \tau^2}{\sigma^2 + \tau^2})$
$X \theta \sim \mathcal{B}(n,\theta)$	$\theta \sim \mathcal{B}e(\alpha, \beta)$	$\theta X \sim \mathcal{B}e(\alpha + x, n - x + \beta)$
$X_1, \ldots, X_n \theta \sim \mathcal{P}(\theta)$	$\theta \sim \mathcal{G}a(\alpha, \beta)$	$\theta X_1,\ldots,X_n \sim \mathcal{G}a(\sum_i X_i + \alpha, n + \beta).$
$X_1,\ldots,X_n \theta \sim \mathcal{NB}(m,\theta)$	$\theta \sim \mathcal{B}e(\alpha, \beta)$	$\theta X_1,\ldots,X_n \sim \mathcal{B}e(\alpha+mn,\beta+\sum_{i=1}^n x_i)$
$X \sim \mathcal{G}(n/2, 2\theta)$	$\theta \sim \mathcal{IG}(\alpha, \beta)$	$\theta X \sim \mathcal{IG}(n/2 + \alpha, (x/2 + \beta^{-1})^{-1})$
$X_1,\ldots,X_n \theta\sim\mathcal{U}(0,\theta)$	$\theta \sim \mathcal{P}a(\theta_0, \alpha)$	$\theta X_1, \dots, X_n \sim \mathcal{P}a(\max\{\theta_0, x_1, \dots, x_n\}\alpha + n)$
$X \theta \sim \mathcal{N}(\mu, \theta)$	$\theta \sim \mathcal{IG}(\alpha, \beta)$	$\theta X \sim \mathcal{IG}(\alpha + 1/2, \beta + (\mu - X)^2/2)$
$X \theta \sim \mathcal{G}a(\nu,\theta)$	$\theta \sim \mathcal{G}a(\alpha, \beta)$	$\theta X \sim \mathcal{G}a(\alpha + \nu, \beta + x)$

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Likelihood Principle

• Likelihood principle: In the inference about θ , after y is observed, all relevant experimental information is contained in the likelihood function for the observed y. Furthermore, two likelihood functions contain the same information about θ if they are proportional to each other

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- Consider testing the fairness of a coin:

$$H_0: \theta = \frac{1}{2} \text{ vs. } H_1: \theta > \frac{1}{2}$$

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Likelihood Principle

- Likelihood principle: In the inference about θ , after y is observed, all relevant experimental information is contained in the likelihood function for the observed y. Furthermore, two likelihood functions contain the same information about θ if they are proportional to each other
- Consider testing the fairness of a coin:

$$H_0: \theta = \frac{1}{2} \text{ vs. } H_1: \theta > \frac{1}{2}$$

 Data: An experiment is conducted and 9 heads and 3 tails are observed

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- ullet Bayesian method has no difficulty o the same conclusion under both scenarios

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 - This is the same as if we observed both batches together

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Simulate Normal Random Variables: Box–Muller Transformation

 We require two random variables, U and V, uniformly distributed on [0,1]. Set

$$R = \sqrt{-2\log V},$$
$$\theta = 2\pi U,$$

and

$$Z_1 = R \cos \theta$$
,
 $Z_2 = R \sin \theta$.

Then they are independent standard normal variables. To obtain two standard normal variables with correlation ρ , take

$$X = Z_1$$

$$Y = \rho Z_1 + \sqrt{1 - \rho^2} Z_2$$

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• Response: *y*

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iid



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- $E(\mathbf{y}|\mathbf{X}) = f(\mathbf{X})$
- $\mathbf{y}|\mathbf{X} \sim \mathcal{N}_n(\mathbf{X}\boldsymbol{\beta}, \sigma^2 \boldsymbol{I})$



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Frequentist Inference

Ordinary least squares

$$\hat{oldsymbol{eta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$$
 $\hat{\sigma}^2 = s^2 = rac{1}{n-k}(\mathbf{y} - \mathbf{X}\hat{oldsymbol{eta}})'(\mathbf{y} - \mathbf{X}\hat{oldsymbol{eta}})$

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Bayesian Inference

Noninformative prior:

$$f(\boldsymbol{\beta}, \sigma^2 | \mathbf{X}) \propto \sigma^{-2}$$

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Posterior:

$$eta, \sigma^2 | \mathbf{y} \sim \mathcal{N}(\hat{eta}, V_{eta} \sigma^2) \ rac{(n-k)s^2}{\sigma^2} | \mathbf{y} \sim \chi^2_{n-k} \ \sigma^2 | \mathbf{y} \sim Inv - \chi^2 (n-k, s^2)$$

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Bayesian Inference

Noninformative prior:

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Posterior:

$$\begin{split} \boldsymbol{\beta}, \sigma^2 | \mathbf{y} &\sim \mathcal{N}(\hat{\boldsymbol{\beta}}, V_{\boldsymbol{\beta}} \sigma^2) \\ \frac{(n-k)s^2}{\sigma^2} | \mathbf{y} &\sim \chi_{n-k}^2 \\ \sigma^2 | \mathbf{y} &\sim \textit{Inv} - \chi^2 (n-k, s^2) \end{split}$$

• Marginal posterior of $\boldsymbol{\beta}|\mathbf{y}$ is the multivariate t-distribution with n-k degrees of freedom

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Hierarchical Model

 Powerful technique for describing complex models. Idea is to break the model down into smaller easier understood pieces, which when put together describes the joint distribution of all data and parameters

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 - Non-hierarchical models with few parameters generally don't fit the data well

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 - Hierarchical models can often fit data with a small number of parameters but can also do well in prediction

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Observed datum is X

$$\mathbf{X}|\theta \sim \mathcal{N}(\theta,1)$$



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 - Second-stage prior: $\tau^2 | \alpha \sim gamma(\alpha, 1)$
 - ▶ Third-stage prior: $\alpha \sim \textit{Exp}(1)$

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Hierarchical Linear Model

$$Y|X, \beta, \Sigma \sim \mathcal{N}(X\beta, \Sigma)$$
$$\beta|X_{\beta}, \alpha, \Sigma_{\beta} \sim \mathcal{N}(X_{\beta}\alpha, \Sigma_{\beta})$$
$$\alpha|\alpha_{0}, \Sigma_{\alpha} \sim \mathcal{N}(\alpha_{0}, \Sigma_{\alpha})$$

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J groups



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- J groups
- Data in group j: $Y_{1j}, \ldots, Y_{n_j j}$

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- Data in group j: $Y_{1j}, \ldots, Y_{n_j j}$
- $Y_{ij}|\beta_j, \sigma^2 \sim \mathcal{N}(\beta_j, \sigma^2)$ independent, $j=1,\ldots,J, \ i=1,\ldots,n_j$

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- $\alpha \sim \mathcal{N}(\alpha_0, \sigma_\alpha^2)$



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 - ► Simulate independent samples from the posterior distributions

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- ► Gibbs sampling
- Metropolis-Hastings
- Hamiltonian dynamics

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• Used for multiparameter models



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• Used for multiparameter models

• Parameter: $\theta = (\theta_1, \dots, \theta_k)$

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 - **...**

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 - **...**
 - ▶ Draw θ_k from $p(\theta_k|\theta_1,\theta_2,\ldots,\theta_{k-1},y)$

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Gibbs Sampler (cont.)

• Full conditional distribution $p(\theta_j|\theta_{-j},y)$, where $\theta_{-j}=(\theta_1,\ldots,\theta_{j-1},\theta_{j+1},\ldots,\theta_k)$

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Gibbs Sampler (cont.)

- Full conditional distribution $p(\theta_j | \theta_{-j}, y)$, where $\theta_{-j} = (\theta_1, \dots, \theta_{j-1}, \theta_{j+1}, \dots, \theta_k)$
- In iteration t, draw $\theta_j^t = p(\theta_j | \theta_{-j}^t, y)$, where $(\theta_1^t, \dots, \theta_{j-1}^t, \theta_{j+1}^{t-1}, \dots, \theta_k^{t-1})$



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Gibbs Sampler (cont.)

- Full conditional distribution $p(\theta_j | \theta_{-j}, y)$, where $\theta_{-j} = (\theta_1, \dots, \theta_{j-1}, \theta_{j+1}, \dots, \theta_k)$
- In iteration t, draw $\theta_j^t = p(\theta_j | \theta_{-j}^t, y)$, where $(\theta_1^t, \dots, \theta_{j-1}^t, \theta_{j+1}^{t-1}, \dots, \theta_k^{t-1})$
- ullet Each $heta_j$ is updated conditional on the latest values of heta

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General Property of Gibbs Sampler

Output: A dependent sequence

$$\theta^{(1)} = \{\theta_1^{(1)}, \dots, \theta_p^{(1)}\}$$

$$\theta^{(2)} = \{\theta_1^{(2)}, \dots, \theta_p^{(2)}\}$$

$$\vdots$$

$$\theta^{(S)} = \{\theta_1^{(S)}, \dots, \theta_p^{(S)}\}$$

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 $m{ heta}(\mathcal{S})$ depends on $m{ heta}^{(0)},\ldots,m{ heta}^{(\mathcal{S}-1)}$ only through $m{ heta}^{(\mathcal{S}-1)}$

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$$\theta^{(S)} = \{\theta_1^{(S)}, \dots, \theta_p^{(S)}\}\$$

- $\theta^{(S)}$ depends on $\theta^{(0)}, \dots, \theta^{(S-1)}$ only through $\theta^{(S-1)}$
- $\theta^{(S)}$ is conditionally independent of $\theta^{(0)}, \dots, \theta^{(S-2)}$ given $\theta^{(S-1)}$

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• Output: A dependent sequence

$$\theta^{(1)} = \{\theta_1^{(1)}, \dots, \theta_p^{(1)}\}\$$

$$\theta^{(2)} = \{\theta_1^{(2)}, \dots, \theta_p^{(2)}\}\$$

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- For the models in this class, the sampling distribution of $\theta^{(S)}$ approaches the target distribution as $S \to \infty$, regardless of starting value

$$Pr(heta^{(S)} \in A
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 as $S
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General Property of Gibbs Sampler (cont.)

More importantly, for most functions g of interest,

$$\frac{1}{S} \sum_{s=1}^{S} g(\theta^{(s)}) \to E[g(\theta)] = \int g(\theta) p(\theta) d\theta \text{ as } S \to \infty$$

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- Hence, we call this method Markov chain Monte Carlo (MCMC)

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Bayesian data analysis using Monte Carlo methods

- Bayesian data analysis using Monte Carlo methods
 - Data analysis: The statistical part

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- When the posterior distribution is complicated, we can "look at" the posterior by studying Monte Carlo samples from the posterior

Monte Carlo and MCMC sampling algorithms

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- Approximation: The use of Monte Carlo procedures to approximate integrals

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• Length of a chain



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- Multiple chains (in parallel)
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- We will deal with these later

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