

ISLR Chapter 3 Exercises

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```
library(ISLR)
library(ggplot2)
library(boot)
```

Conceptual

Question 1

$$Var(\alpha X + (1 - \alpha)Y) = \alpha^2 Var(X) + (1 - \alpha)^2 Var(Y) + \alpha(1 - \alpha)Cov(X, Y)$$

Differentiate with respect to α and set equal to 0:

$$2\alpha Var(X) - 2(1 - \alpha)Var(Y) + 2(-\alpha + 1 - \alpha)Cov(X, Y) = 2\alpha Var(X) - 2(1 - \alpha)Var(Y) + 2(1 - 2\alpha)Cov(X, Y) = 2\alpha Var(X) - 2Var(Y) + 2\alpha Var(Y) + 2Cov(X, Y) - 4\alpha Cov(X, Y) = 0$$

$$\alpha(Var(X) + Var(Y) - 2Cov(X, Y)) - Var(Y) + Cov(X, Y) = 0$$

$$\alpha = \frac{Var(Y) - Cov(X, Y)}{Var(X) + Var(Y) - 2Cov(X, Y)}$$

Question 2

a $1 - \frac{1}{n}$

b $1 - \frac{1}{n}$

c The probability that the bootstrap sample is not in the sample is the multiplication of the probabilities that a given bootstrap observation is not the j th observation. This probability is

$$(1 - \frac{1}{n})^n$$

```
print(1 - (1 - 1/5) ^ 5)
```

d-f

```
## [1] 0.67232
```

```
print(1 - (1 - 1/100) ^ 100)
```

```
## [1] 0.6339677
```

```
print(1 - (1 - 1/10000) ^ 10000)
```

```
## [1] 0.632139
```

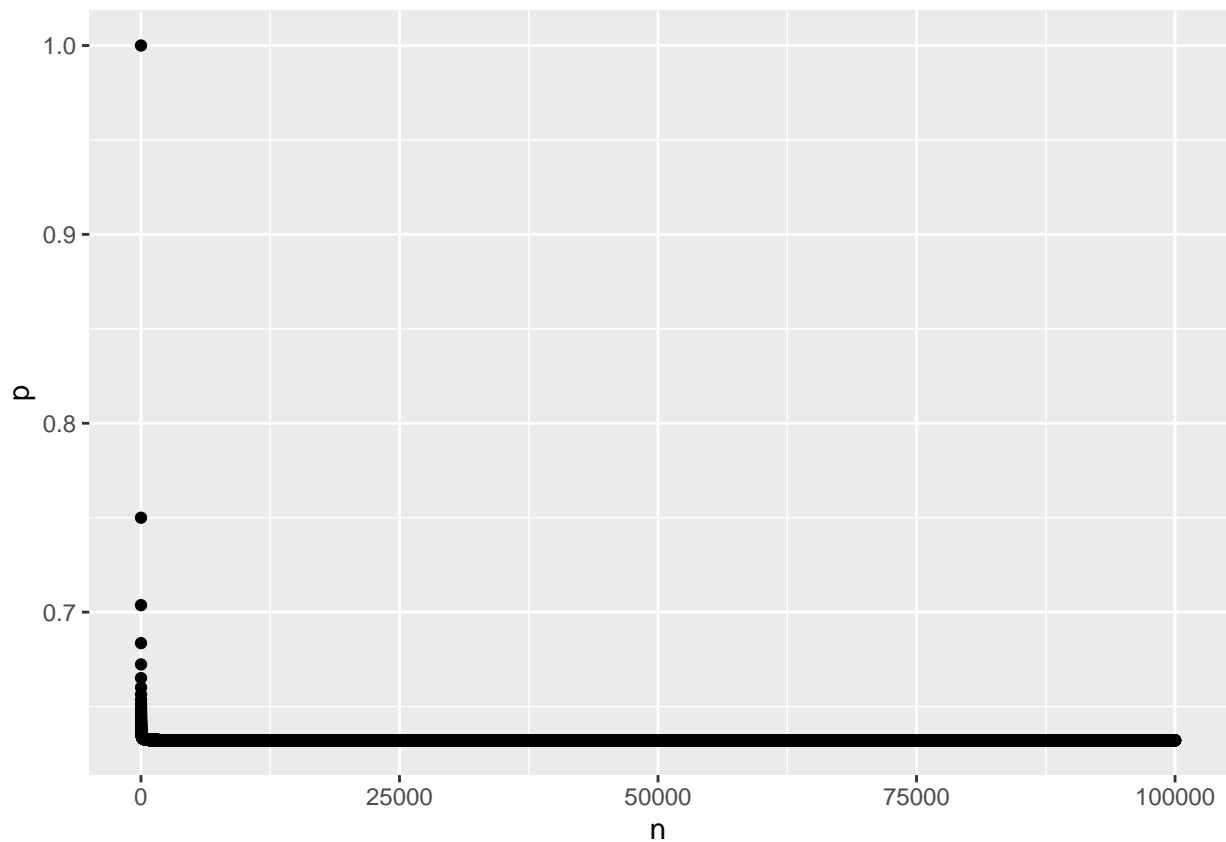
As the number of observations increases, the probability converges to $\lim_{n \rightarrow \infty} 1 - (1 - \frac{1}{n})^n = 1 - \frac{1}{e}$

g

```
df <- data.frame(n = seq(100000))
```

```
df$p <- 1 - (1 - 1/df$n) ^ df$n
```

```
ggplot2::ggplot(data = df) +  
  ggplot2::geom_point(ggplot2::aes(x = n, y = p))
```



We see the line approaching the horizontal asymptote that corresponds to the limit above.

h

```
n_samples <- 100000
```

```
results <- rep(0, n_samples)
```

```
samples <- sample(
```

```
seq(100000), function(idx) {
  any(sample(100, replace = TRUE) == 4)
}
)

print(mean(samples))
```

```
## [1] 0.63423
```

Question 3

a K-fold cross validation proceeds as follows. First, the dataset is randomly split into K distinct subsets whose union is the entire dataset. Each subset, or **fold** should have the same number of observations (in the case where the dataset size is not divisible by K, some folds will have one more observation than other folds.) Then for each fold, we take the union of the other folds and train a model using those folds. That model is then used to predict on the fold in question to generate an error estimate. Finally, the error estimates from each fold are averaged for a final error estimate.

b

i K-fold cross validation has lower bias compared to the validation set approach, as it uses more of the data for each model and hence will tend to underestimate the error less. It also has lower variance than the validation set approach, as the validation set approach uses only half the data and hence will be quite variable across different datasets.

ii K-fold cross validation has lower variance than LOOCV because the models are less correlated with one another, as each model in LOOCV is trained on the same dataset as each other model with the exception of one observation. When averaging many correlated quantities, the variance of the mean is higher than if they were independent. In K-fold cross-validation, less observations are shared between models, so this variance effect is not as large. K-fold cross-validation will have higher bias than LOOCV though, since each model is trained on a smaller portion of the data than in LOOCV.

iii If we have already performed cross-validation and K is large, we could use each of the models we trained to predict on the test point and record the standard deviation of these predictions. Alternatively, we could create a larger number of bootstrap samples from the training set and build a model on each of those, and then use the to predict on the test point. Again we could record the standard deviation of those predictions.

Applied

Question 5

a

```
df_default <- Default
model_full <- glm(default ~ income + balance, data = df_default, family = "binomial")

print(summary(model_full))
```

```
##
## Call:
## glm(formula = default ~ income + balance, family = "binomial",
##      data = df_default)
##
## Coefficients:
```

```
##               Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.154e+01  4.348e-01 -26.545  < 2e-16 ***
## income      2.081e-05  4.985e-06   4.174 2.99e-05 ***
## balance     5.647e-03  2.274e-04  24.836  < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##    Null deviance: 2920.6  on 9999  degrees of freedom
## Residual deviance: 1579.0  on 9997  degrees of freedom
## AIC: 1585
##
## Number of Fisher Scoring iterations: 8
```

b - d

```
create_train_test_split <- function(df, seed) {
  set.seed(seed)
  nrows <- nrow(df)
  train_idx <- sample(nrows, nrows %/% 2)

  df_train <- df[train_idx, ]
  df_test <- df[-train_idx, ]

  list(df_train = df_train, df_test = df_test)
}

build_model_and_validate <- function(df_train, df_test, vars) {

  form <- as.formula(paste("default ~", paste(vars, collapse = " + ")))
  model <- glm(form, data = df_train, family = "binomial")

  preds <- predict(model, df_test, type = "response")
  preds <- ifelse(preds >= 0.5, "Yes", "No")

  mean(preds == df_test$default)
}

run_analysis <- function(df, seed, vars) {
  df_list <- create_train_test_split(df, seed)
  do.call(build_model_and_validate, c(df_list, list(vars = vars)))
}

seeds <- seq(3)
var_list <- list(c("income", "balance"), c("income", "balance", "student")) %>%
  setNames(., lapply(., paste, collapse = "_"))

mean accuracies <- vector("list", length(var_list)) %>%
  setNames(., names(var_list))

for (nm in names(var_list)) {
  mean accuracies[[nm]] <- sapply(seeds, run_analysis, df = df_default, vars = var_list[[nm]])
}
```

```

}

print(mean_accuracies)

## $income_balance
## [1] 0.9746 0.9762 0.9736
##
## $income_balance_student
## [1] 0.9740 0.9754 0.9728

```

```
print(sapply(mean_accuracies, mean))
```

```
##          income_balance income_balance_student
##          0.9748000          0.9740667
```

The test set error depends on the splitting of the data, which is determined by the seed. There is no evidence that including student as a covariate improves the model.

Question 6

a - d

```
print(summary(glm(default ~ income + balance, data = df_default, family = "binomial"))$coefficients)
```

```
##              Estimate Std. Error   z value    Pr(>|z|)
## (Intercept) -1.154047e+01 4.347564e-01 -26.544680 2.958355e-155
## income      2.080898e-05 4.985167e-06  4.174178 2.990638e-05
## balance     5.647103e-03 2.273731e-04 24.836280 3.638120e-136
```

```
run_logistic_regression <- function(df, index) {
  df <- df[index, ]
  coef(glm(default ~ income + balance, data = df, family = "binomial"))
}
```

```
set.seed(1)
print(boot::boot(df_default, run_logistic_regression, R = 50))
```

```
##
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
## Call:
## boot::boot(data = df_default, statistic = run_logistic_regression,
##           R = 50)
##
##
## Bootstrap Statistics :
##      original      bias    std. error
## t1* -1.154047e+01 -5.661486e-02 4.847786e-01
## t2*  2.080898e-05 -7.436578e-08 4.456965e-06
## t3*  5.647103e-03  1.854126e-05 2.639029e-04
```

The standard errors from the bootstrap are similar to the theoretical standard errors.

Question 7

a - e

```
df_weekly <- ISLR::Weekly

# Using the boot library
accuracy_cost <- function(y_true, y_pred) {
  y_pred <- ifelse(y_pred >= 0.5, 1, 0)
  mean(y_pred == y_true)
}

model <- glm(Direction ~ Lag1 + Lag2, data = df_weekly, family = "binomial")
loocv_estimate <- boot::cv.glm(data = df_weekly, glmfit = model, cost = accuracy_cost)

print(paste("LOOCV estimate from boot library:", loocv_estimate$delta[[1]]))

## [1] "LOOCV estimate from boot library: 0.550045913682276"

# Manually
predictions <- sapply(
  seq(nrow(df_weekly)),
  function(idx, df) {
    df_train <- df[-idx, ]
    df_test <- df[idx, ]

    model <- glm(Direction ~ Lag1 + Lag2, data = df_train, family = "binomial")

    predict(model, df_test, type = "response")
  },
  df = df_weekly
)

binary_y_true <- ifelse(df_weekly$Direction == "Up", 1, 0)
loocv_estimate <- accuracy_cost(binary_y_true, predictions)

print(paste("Manual LOOCV estimate:", loocv_estimate))

## [1] "Manual LOOCV estimate: 0.550045913682277"
```

Question 8

a

```
set.seed(1)

df <- data.frame(x = rnorm(100))
df$y <- df$x - 2 * df$x ^ 2 + rnorm(100)

print(paste("n:", nrow(df)))

## [1] "n: 100"

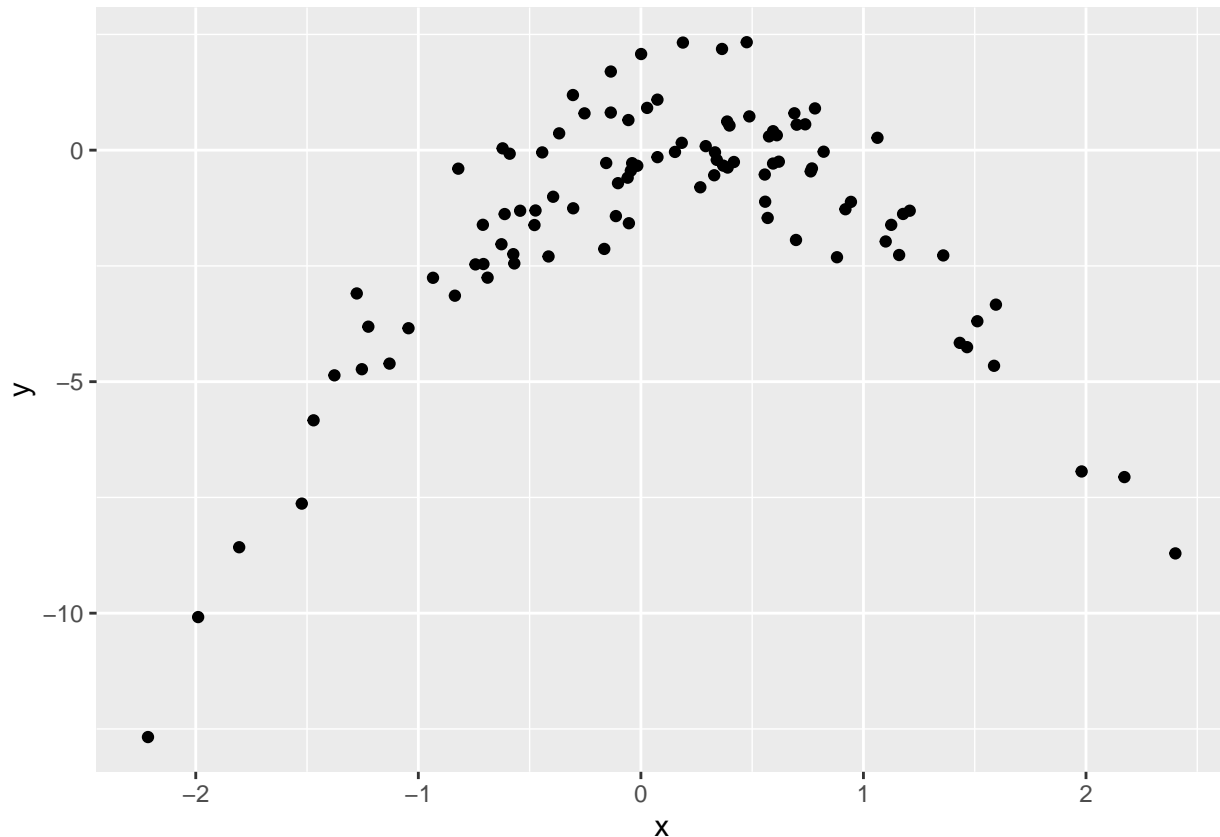
print(paste("p:", ncol(df)))
```

```
## [1] "p: 2"
```

$$y = x - 2x^2 + \epsilon$$

b

```
ggplot2::ggplot(data = df) +  
  ggplot2::geom_point(ggplot2::aes(x = x, y = y))
```



We see a quadratic relationship.

c Note that I am using orthogonal polynomials here.

```
seeds <- seq(3)  
poly_degrees <- seq(4)  
  
results <- vector("list", length = length(seeds)) %>%  
  setNames(seeds)  
for (seed in seeds) {  
  set.seed(seed)  
  results[[seed]] <- sapply(  
    poly_degrees,  
    function(degree, df, seed) {  
      model <- glm(y ~ poly(x, degree), data = df)  
      # Only need to print this once, since the results are the same across seeds.  
      if (identical(seed, 1)) {  
        print(summary(model))  
      }  
    }  
  )  
}
```

```

    }
    boot::cv.glm(df, model)$delta[[1]]
  },
  df = df,
  seed = seed
)
}

```

The results are the same regardless of seed because there is no randomness in LOOCV. The best MSE results from using a polynomial of degree 2. Degrees greater than 2 are not statistically significant, agreeing with the LOOCV results.

Question 9

a

```

df_boston <- Boston
mean_medv <- mean(df_boston$medv)
print(paste("Sample mean of medv:", mean_medv))

```

```
## [1] "Sample mean of medv: 22.5328063241107"
```

b From theory we know that the sample mean is normally distributed with mean equal to the true mean of the population, and variance equal to the variance of the population divided by the number of observations that comprise the mean.

```
print(paste("Estimated sd of sample mean of medv:", sd(df_boston$medv) / sqrt(nrow(df_boston))))
```

```
## [1] "Estimated sd of sample mean of medv: 0.408861147497535"
```

c

```

calculate_mean <- function(medv, index) {
  mean(medv[index])
}

set.seed(1)
sd_boot_estimate <- boot::boot(df_boston$medv, calculate_mean, 1000)
sd_boot_estimate <- sd(sd_boot_estimate$t)
print(paste("Standard deviation of mean using bootstrap:", sd_boot_estimate))

```

```
## [1] "Standard deviation of mean using bootstrap: 0.410662153886544"
```

We get a similar estimate of the standard error.

```

quantile_t <- qt(.975, nrow(df_boston) - 1)
print(
  paste(
    "Confidence interval for mean of medv using bootstrap:",
    paste(mean_medv + c(-quantile_t, quantile_t) * sd_boot_estimate, collapse = " - ")
  )
)

```


d

```
## [1] "Confidence interval for mean of medv using bootstrap: 21.7259896266039 - 23.3396230216174"
t_test <- t.test(df_boston$medv)
print(paste("Confidence interval for mean of medv using t.test:", paste(t_test$conf.int, collapse = " -
```

```
## [1] "Confidence interval for mean of medv using t.test: 21.7295280145786 - 23.3360846336428"
```

```
median_medv <- median(df_boston$medv)
print(paste("Sample median of medv:", median_medv))
```

e

```
## [1] "Sample median of medv: 21.2"
```

```
calculate_median <- function(medv, index) {
  median(medv[index])
}

set.seed(1)
sd_boot_estimate <- boot::boot(df_boston$medv, calculate_median, R = 1000)
sd_boot_estimate <- sd(sd_boot_estimate$t)

print(paste("SD of boot estimate for median:", sd_boot_estimate))
```

f

```
## [1] "SD of boot estimate for median: 0.377807538352579"
```

```
percentile_10_medv <- quantile(df_boston$medv, 0.1)
print(paste("Sample 10th percentile of medv:", percentile_10_medv))
```

g

```
## [1] "Sample 10th percentile of medv: 12.75"
```

```
calculate_10th_percentile <- function(medv, index) {
  quantile(medv[index], 0.1)
}

set.seed(1)
sd_boot_estimate <- boot::boot(df_boston$medv, calculate_10th_percentile, R = 1000)
sd_boot_estimate <- sd(sd_boot_estimate$t)

print(paste("SD of boot estimate for 10th percentile:", sd_boot_estimate))
```

g

```
## [1] "SD of boot estimate for 10th percentile: 0.476752643498789"
```