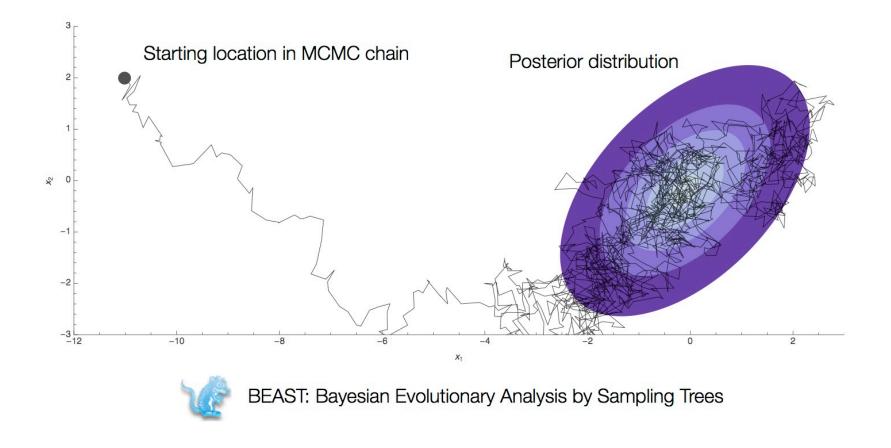
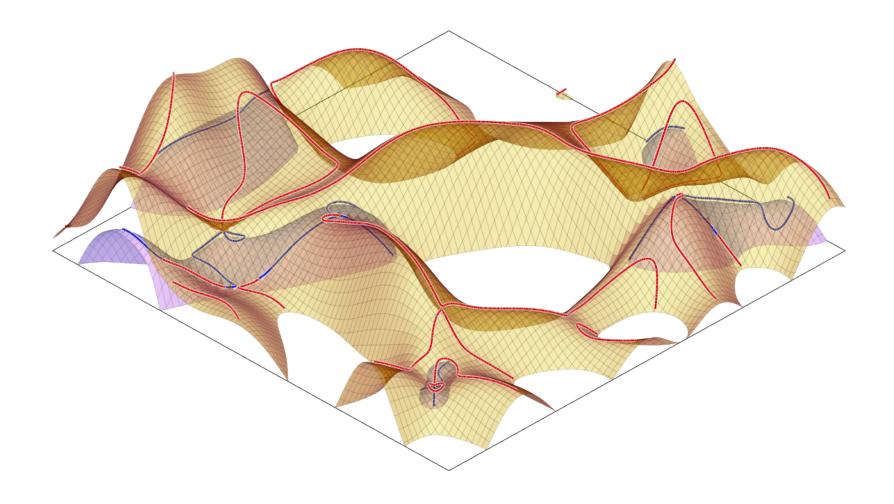
Markov Chain Monte Carlo

Methods, Algorithms, and Applications



Black magic (noun): code based on techniques that appear to work but which lack a theoretical explanation



$$F(T) = \left(\frac{m}{2\pi i \hbar \varepsilon}\right)^{N/2} \int_{-\infty}^{\infty} \left(\prod_{l=1}^{N-1} d\chi_l\right) \exp\left\{\frac{im}{2\hbar \varepsilon} \chi_j M_{jk} \chi_k\right\}$$

Applications of MCMC

- High-dimensional integrals/ probability distributions
- Computing large hierarchical models
- Rare event sampling
- Marginalization/ Parameter estimation
- Model comparison



Top 10 Most Influential Algorithm of the 20th Century (CiSE)

Markov Chain

$$p(\theta_{i+1} \mid \{\theta_i\}) = p(\theta_{i+1} \mid \theta_i)$$

"memoryless"

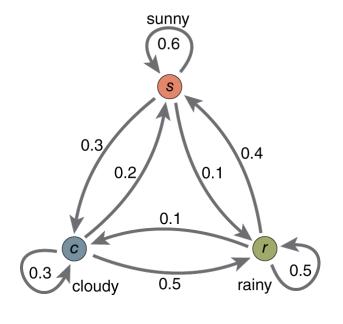
$$p(\theta_{i+1} \mid \theta_i) = p(\theta_i \mid \theta_{i+1})$$

reversible

$$p(\theta_{i+1}) = \int T(\theta_{i+1} \mid \theta_i) p(\theta_i) d\theta_i$$

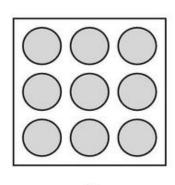
$$T(\theta_{i+1} \mid \theta_i) p(\theta_i) = T(\theta_i \mid \theta_{i+1}) p(\theta_{i+1})$$

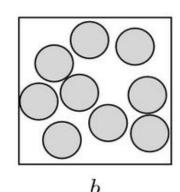
$$p_{acc}(\theta_i, \theta_{i+1}) = \frac{K(\theta_i \mid \theta_{i+1})p(\theta_{i+1})}{K(\theta_{i+1} \mid \theta_i)p(\theta_i)}$$



Metropolis-Hastings Algorithm

- Choose an appropriate initial model, or "state"
- Compute steps
- keep or reject for each "item"
- Calculate quantity of interest for new state
- Iterate as needed





$$a = \frac{f(x_{i+1})}{f(x_i)}$$

$$f(x) = \frac{1}{N} \sum_{i} f(x_i)$$

Metropolis-Hastings Algorithm

ADVANTAGES

- No curse of dimensionality
 - P(rejection) α D
- High-dimensionality
 - Only option

DISADVANTAGES

- Auto-correlation
 - Jumping width
 - Thinning
- Burn-in period
 - Local minima

Gibbs Sampling

- Multi-dimensional sampling → multiple lowdimensional samples
- Conditional distribution
- Advantages over M-H
 - Bayesian networks
 - Reduce autocorrelation
 - Collapsed Gibbs Sampling
 - Blocked Gibbs Sampling
 - Simulated annealing
 - Ordered overrelaxation

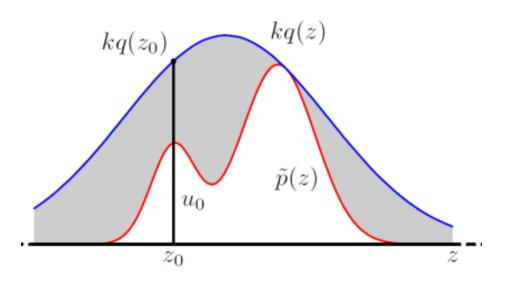
$$X^{(i)} \rightarrow X^{(i+1)}$$

$$p(x_1, x_2, ..., x_n)$$

$$p(x_j^{(i+1)} \mid x_1^{(i+1)}, ..., x_{j-1}^{(i+1)}, x_{j+1}^{(i)}, ..., x_n^{(i)})$$

Adaptive Rejection Sampling

- Log-concave density functions
- Piece-wise linear density functions
- Avoid evaluating p(z0)



MCMC with Python

- Emcee
- PyMC
 - Bayesian statistical models and fitting algorithms
- PyStan
 - Gradient-based MCMC algorithms for Bayesian inference
 - No-U-Turn Sampling (NUTS), Hamiltonian MC

	Complexity	Execution time (100,000 samples; includes burn-in)	Ease of Installation	Learning Curve/Ease of Use	Number of Features
emcee	Very lightweight	~6 sec	Pure python; installs easily with pip	Straightforward & Pythonic	not much built-in beyond basic MCMC sampling
pymc2	Lots of functionality & options	~17 sec	Requires fortran compiler; binaries available via conda	Pythonic, but lots of pymc- specific boilerplate	Lots of built- in functionality in Python
pystan	Large package; requires coding in Stan language	~20 sec compilation + ~6 sec computation	Requires C compiler + Stan installation; no binaries available	Not pure Python; must learn Stan model specification language	Lots of built- in functionality in Stan- specific language

Jake VanderPlas - Frequentism and Bayesianism: How to be Bayesian in Python

Emcee (The MCMC Hammer)

- Affine Invariant MCMC
 Ensemble Sampler
- Bayesian parameter estimation



```
import numpy as np
import emcee

def lnprob(x, ivar):
    return -0.5 * np.sum(ivar * x ** 2)

ndim, nwalkers = 10, 100
ivar = 1. / np.random.rand(ndim)
p0 = [np.random.rand(ndim) for i in range(nwalkers)]

sampler = emcee.EnsembleSampler(nwalkers, ndim, lnprob, args=[ivar])
sampler.run_mcmc(p0, 1000)
```

Affine Invariance

$$P(x) = \lambda^{-n} P(\lambda x)$$

$$x \to y = ax + b$$

$$P(x) \to P(ax + b) \propto P(x)$$

$$P(x) \propto \exp(\frac{-(x_1 - x_2)^2}{2\varepsilon} - \frac{(x_1 + x_2)^2}{2})$$

$$P(x) \propto \exp(\frac{-(y_1^2 + y_2^2)}{2})$$

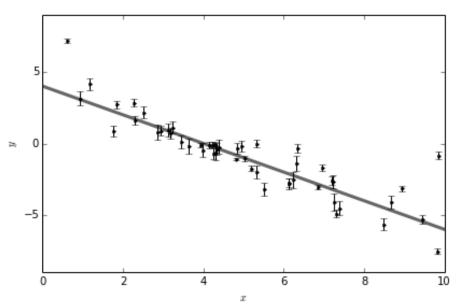
$$y_1 = \frac{x_1 - x_2}{\sqrt{\varepsilon}}, y_2 = x_1 + x_2$$

Example - Fitting a Line

```
import emcee
import corner
import numpy as np
import scipy.optimize as op
import matplotlib.pyplot as pl
from matplotlib.ticker import MaxNLocator

# Set "true" parameters
m_true = -1.0
b_true = 4.0
f_true = 0.50
```

Example – Fitting a Line



```
# Generate sample data from our "model"
np.random.seed(123)
N = 50
x = np.sort(10*np.random.rand(N))
yerr = 0.1+0.5*np.random.rand(N)
y = m_true*x+b_true
y += np.abs(f_true*y) * np.random.randn(N)
y += yerr * np.random.randn(N)
```

Example – Fitting a Line

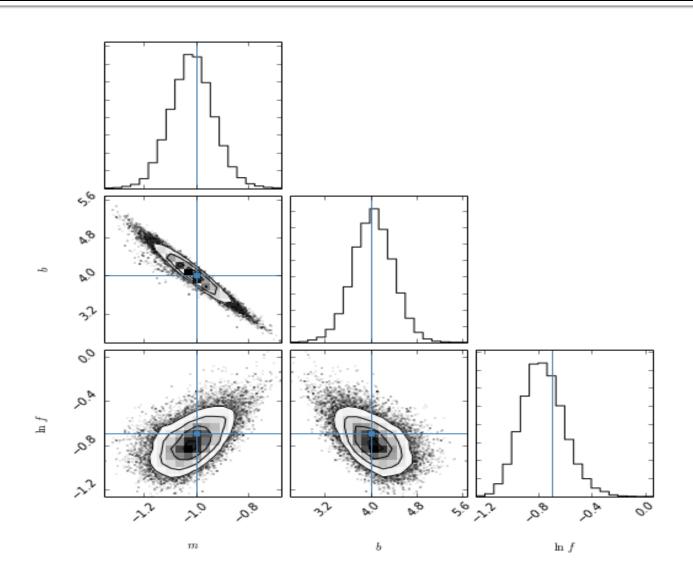
```
# Define the probability function as likelihood * prior.
def Inprior(theta):
    m, b, lnf = theta
    if -5.0 < m < 0.5 and 0.0 < b < 10.0 and -10.0 < lnf < 1.0:
        return 0.0
    return -np.inf
def lnlike(theta, x, y, yerr):
    m, b, lnf = theta
    model = m * x + b
    inv sigma2 = 1.0/(yerr**2 + model**2*np.exp(2*lnf))
    return -0.5*(np.sum((y-model)**2*inv sigma2 - np.log(inv sigma2)))
def lnprob(theta, x, y, yerr):
    lp = lnprior(theta)
    if not np.isfinite(lp):
        return -np.inf
    return lp + lnlike(theta, x, y, yerr)
```

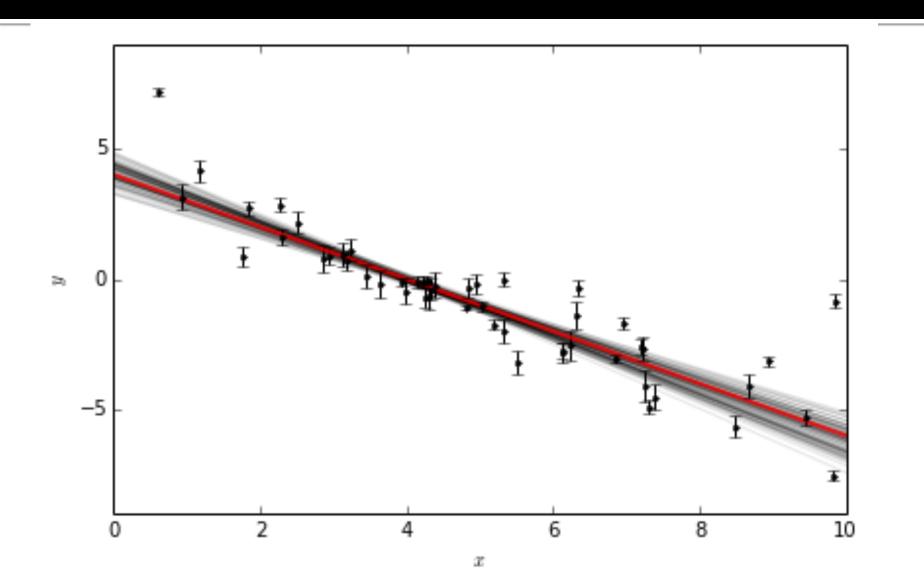
LSR and MLE Results

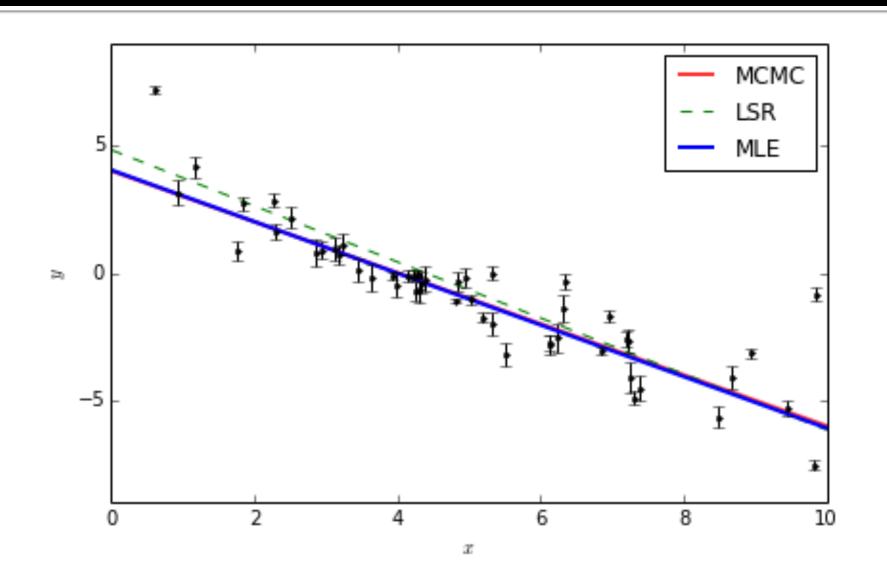
		Maximum Likelihood Estimation	True Values
Slope	-1.0907 ± 0.0162	-1.0115	-1.0
Intercept	4.8155 ± 0.0909	4.0283	4.0

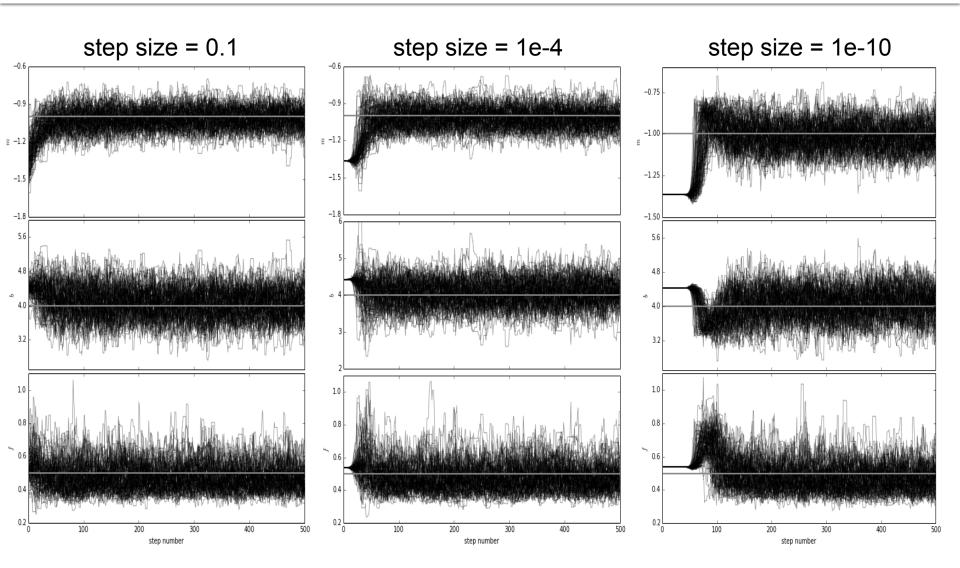
```
# Set up ensemble sampler and run MCMC
n_dim, n_walkers = 3, 100
p0 = result["x"] + np.random.normal(size=3)
pos = [p0 + 1e-4*np.random.randn(n_dim) for i in range(n_walkers)]
sampler = emcee.EnsembleSampler(n_walkers, n_dim, lnprob, args=(x, y, yerr))
sampler.run_mcmc(pos, 500, rstate0=np.random.get_state())
```

Slope	Intercept	
$-1.0182^{+0.0819}_{-0.0814}$	$4.0539^{+0.0828}_{-0.0647}$	









References

- 1. Ivezic Section 5.8
- 2. https://en.wikipedia.org/wiki/ Markov chain Monte Carlo
- 3. Goodman and Weare, "Ensemble Samplers with Affine Invariance", http://msp.org/camcos/2010/5-1/camcos-v5-n1-p04-s.pdf
- 4. Emcee, http://dan.iel.fm/emcee/current/