# **PCML Cheat Sheet**

#### 1 Math Prerequisites

- Bayes rule

$$p(A,B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}))$$

- Production of independent variables:

$$\operatorname{Var}(XY) = \mathbb{E}(X^2) \,\mathbb{E}(Y^2) - [\mathbb{E}(X)]^2 [\mathbb{E}(Y)]^2$$

Covariance matrix of a data vector x

$$\mathbf{\Sigma} = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \mathbb{E}(\mathbf{x})) (\mathbf{x}_n - \mathbb{E}(\mathbf{x}))^T$$

### 1.1 Convexity

- A function f(x) is convex, if for any  $x_1, x_2 \in \mathbf{X}$ and for any  $0 \le \lambda \le 1$ , we have :

$$f(\lambda x_1 + (1-\lambda)x_2) \leq \lambda f(x_1) + (1-\lambda)f(x_2)$$

The Hessian of a convex function is psd and for a strictly-convex function it's pd.

$$\mathbf{H}_{i,j} = d^2 f / dx_i dx_j$$

## 1.2 Linear Algebra

Condition number of a function measures how much the output value can change for a small change in the input. A matrix with a high condition number is said to be ill-conditioned. If **A** is normal  $(A^T A = AA^T)$  then

$$k(\mathbf{A}) = \left| \frac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

- A positive definite matrix is symmetric with all positive eigenvalues
- The real symmetric  $N \times N$  matrix **V** is said to be positive semidefinite if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} \ge 0$$

- for any real  $N \times 1$  vector a.
- positive definite if a<sup>T</sup>Va > 0

# 2 Cost functions

- Cost functions are used to learn parameters that explain the data well.
- It is essential to make sure that a global minimum exist → lower bounded

# Mean square error (MSE)

$$MSE(\boldsymbol{w}) = \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is convex thus it has only one global minumum value.
- MSE is not good when outliers are present.

#### Mean Absolute Error (MAE):

$$MAE = \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \leq \delta \\ \delta|z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

- Huber loss is convex, differentiable, and also

#### Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Non-convex, non-diff., but robust to outliers

 $Hinge = [1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$ Logistic loss

 $Logistic = log(1 - exp(y_n f(\mathbf{x}_n)))$ 

#### 3 Regression

Data consists of N pairs (yn, xn)

- 1.  $y_n$  the n'th output
- 2.  $\mathbf{x}_n$  is a vector of D inputs Prediction: predict the ouput for a new input
- Interpretation: understand the effect of inputs on output.
- Outliers are data that are far away from most of the other examples.

#### 3.1 Linear Regression

- Model that assume linear relationship between inputs and the ouput.

 $y_n \approx f(\mathbf{x}_n) := w_0 + w_1 x_{n,1} + \dots = \omega_0 + \mathbf{x}_n^T \mathbf{w}$ with  $\boldsymbol{w}$  the parameters of the model.

Variance grows only linearly with dimensionality

#### 4 Optimization

#### 4.1 Grid search

- Compute the cost over a grid of M points to find the minimum
- Exponential Complexity
- Hard to find a good range of values

#### 4.2 Gradient Descent

- Gradient descent uses only first-order information and takes steps in the direction of the gradient
- Given a cost function  $\mathcal{L}(w)$  we wish to find wthat minimizes the cost:

$$\min_{\boldsymbol{x}} \mathcal{L}(\boldsymbol{w})$$

#### 4.3 Batch Gradient Descent

- Take steps in the opposite direction of the gradient

$$\boldsymbol{w}^{(t+1)} \leftarrow \boldsymbol{w}^{(t)} - \gamma \nabla \mathcal{L}(\boldsymbol{w}^{(t)})$$

with  $\gamma > 0$  the learning rate.

With  $\gamma$  too big, method might diverge. With  $\gamma$ too small, convergence is slow.

#### 4.4 Gradients for MSE

- We define the error vector e:

$$e := y - Xw$$

and MSE as follows:

$$\mathcal{L}(\boldsymbol{w}) = \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{w})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

- then the gradient is given by

$$\nabla \mathcal{L}(\boldsymbol{w}) = -\frac{1}{N} \mathbf{X}^T \mathbf{e}$$

- 1. necessary: gradient equal zero:  $\frac{d\mathcal{L}(\boldsymbol{w}^*)}{dz} = 0$ 2. sufficient: Hessian matrix is positive definite:
- $\mathbf{H}(\boldsymbol{w}^*) = \frac{d^2 \mathcal{L}(\boldsymbol{w}^*)}{}$  $H(w^{+}) = \frac{1}{dwdw}T^{-}$ Very sensitive to illconditioning. Therefore,
- always normalize your feature otherwise step-size selection is difficult since different directions might move at different speed.
- Complexity: O(NDI) with I the number of iterations

#### 4.5 Stochastic Gradient Descent

In ML, most cost functions are formulated as a sum over the training examples:

$$\mathcal{L}(\boldsymbol{w}) = \frac{1}{N} \sum_{n=1}^{N} \mathcal{L}_{n}(\boldsymbol{w})$$

⇒ SGD algo is given by update rule:

$$m^{(t+1)} \leftarrow m^{(t)} - \gamma \nabla f - (m^{(t)})$$

Idea: Cheap but unbiased estimate of grad

$$\mathbb{E}[\nabla \mathcal{L}_n(w)] = \nabla(w)$$

### 4.6 Mini-batch SGD

Update direction  $(B \subset [N])$ :

$$g^{(t)} := \frac{1}{|B|} \sum_{n \in B} \nabla \mathcal{L}_n(w^{(t)})$$

Update rule :  $\boldsymbol{w}^{(t+1)} \leftarrow \boldsymbol{w}^{(t)} - \gamma \boldsymbol{\sigma}^{(t)}$ 

# 4.7 Subgradients (Non-Smooth OPT)

A vector  $\mathbf{g} \in \mathbb{R}^D$  s.t.

$$\mathcal{L}(u) \ge \mathcal{L}(w) + g^T(u - w) \quad \forall u$$

is the subgradient to  $\mathcal{L}$  at w. If  $\mathcal{L}$  is differentiable at w, we have  $g = \nabla \mathcal{L}(w)$ 

## 5 Least Squares

- In some cases, we can compute the minimum of the cost function analytically.
- use the first optimality conditions:

$$\nabla L(\mathbf{w}^*) = 0 \Rightarrow \mathbf{X}^T \mathbf{e} = \mathbf{X}^T (\mathbf{y} - \mathbf{X}\mathbf{w}) = 0$$

When  $\mathbf{X}^T\mathbf{X}$  is invertible, we have the closed-form expression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

– thus we can predict values for a new  $\mathbf{x}_m$ 

$$y_m := \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} \mathbf{w}^* = \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

- The **Gram matrix**  $\mathbf{X}^T \mathbf{X}$  is positive definite and is also invertible iff **X** has full column rank  $Complexity: O(ND^2 + D^3) \equiv O(ND^2)$ 

**X** can be rank deficient when D > N or when the comlumns  $\bar{\mathbf{x}}_d$  are nearly collinear. In this case, the matrix is ill-conditioned, leading to numerical issues.

# 6 Maximum Likelihood

Let define our mistakes ε<sub>n</sub> ~ N(0, σ<sup>2</sup>).

$$y_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

- Another way of expressing this:

$$\begin{split} p(\mathbf{y}|\mathbf{X}, \mathbf{w}) &= \prod_{n=1}^{N} p(y_n|\mathbf{x}_n, \mathbf{w}) \\ &= \prod_{n=1}^{N} \mathcal{N}(y_n|\mathbf{x}_n^T \mathbf{w}, \sigma^2) \end{split}$$

which defines the likelihood of observating y given X and w

Define cost with log-likelihood

$$\begin{split} \mathcal{L}_{lik}(\mathbf{w}) &= \log p(\mathbf{y}|\mathbf{X}, \mathbf{w}) \\ &= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} \left(y_n - \mathbf{x}_n^T \mathbf{w}\right)^2 + cnst \end{split}$$

Maximum likelihood estimator (MLE) gives another way to design cost functions

$$\underset{\boldsymbol{a} \in \mathcal{A}}{\operatorname{argmin}} \mathcal{L}_{MSE}(\boldsymbol{w}) = \underset{\boldsymbol{a} \in \mathcal{A}}{\operatorname{argmax}} \mathcal{L}_{lik}(\boldsymbol{w})$$

- MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.
- With Laplace distribution

$$p(y_n|\mathbf{x}_n, \mathbf{w}) = \frac{1}{2b} e^{-\frac{1}{b}|y_n - \mathbf{x}_n^T \mathbf{w}|}$$

$$\sum_{n} \log p(y_n | \mathbf{x}_n, \boldsymbol{w}) = \sum_{n} |y_n - \mathbf{x}_n^T \boldsymbol{w}| + cnst$$

# 7 Ridge Regression

- Linear models usually underfit. One way is to use nonlinear basis functions instead

$$y_n = w_0 + \sum_{j=1}^{M} w_j \phi_j(\mathbf{x}_n) = \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{w}$$

- This model is linear in w but nonlinear in x. Note that the dimensionality is now M, not D.
- Polynomial basis  $\phi(x_n) = [1, x_n, x_n^2, ..., x_n^M]$

The least square solution becomes 
$$oldsymbol{w}_{lse}^* = (\mathbf{\tilde{\Phi}}^T\mathbf{\tilde{\Phi}})^{-1}\mathbf{\tilde{\Phi}}^T\mathbf{y}$$

Complex models overfit easily. Thus we can choose simpler models by adding a regularization term which penalizes complex

$$\min_{\boldsymbol{w}} \left( \mathcal{L}(\boldsymbol{w}) + \frac{\lambda}{2N} \sum_{j=1}^{M} w_j^2 \right)$$

$$\boldsymbol{w}^* = \underset{\boldsymbol{w}}{\operatorname{argmin}} \left( \frac{1}{2} (\mathbf{y} - \mathbf{X} \boldsymbol{w})^T (\mathbf{y} - \mathbf{X} \boldsymbol{w}) + \frac{\lambda}{2} \boldsymbol{w}^T \boldsymbol{w} \right)$$

By differentiating and setting to zero we get

$$\begin{aligned} \mathbf{w}_{ridge} &= (\tilde{\mathbf{\Phi}}^T \tilde{\mathbf{\Phi}} + \mathbf{\Lambda})^{-1} \tilde{\mathbf{\Phi}}^T \mathbf{y} \\ \mathbf{\Lambda} &= \begin{bmatrix} 0 & \underline{0} \\ 0 & \lambda \overline{I}_m \end{bmatrix} \end{aligned}$$

Ridge regression improves the condition number of the Gram matrix since the eigenvalues of  $(\tilde{\Phi}^T \tilde{\Phi} + \lambda I_m)$  are at least  $\lambda$ 

#### Maximum-a-posteriori (MAP) estimator:

- Maximizes the product of the likelihood and the prior.

$$w_{MAP} = \underset{\mathbf{z}_{\mathbf{z}}}{\operatorname{argmax}} (p(\mathbf{y}|\mathbf{X}, \mathbf{\Lambda})p(\mathbf{w}|\mathbf{\Sigma}))$$

- Assume 
$$w_0 = 0$$

$$w_{ridge} = \underset{\boldsymbol{w}}{\operatorname{argmax}} \left( \log \left[ \prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T \boldsymbol{w}, \boldsymbol{\Lambda}) \times \mathcal{N}(\boldsymbol{w} | 0, \mathbf{I}) \right] \right) \text{ Hessian is the derivative of the gradient}$$

- Lasso regularizer forces some  $w_i$  to be strictly 0 and therefore forces sparsity in the model.

$$\begin{split} \min_{\pmb{w}} \frac{1}{2N} \sum_{n=1}^{N} \left(y_n - \tilde{\pmb{\phi}}(\mathbf{x}_n)^T \pmb{w}\right)^2, \\ \text{such that } \sum_{n=1}^{M} \left|w_i\right| \leq \tau \end{split}$$

#### 8 Cross-Validation

We should choose  $\lambda$  to minimize the mistakes that will be made in the future.

- We split the data into train and validation sets and we pretend that the validation set is the future data. We fit our model on the training set and compute a prediction-error on the validation set. This gives us an estimate of the
- generalization error. K-fold cross validation randomly partition the data into K groups. We train on K-1groups and test on the remaining group. We repeat this until we have tested on all K sets. We then average the results.
- Cross-validation returns an unbiased estimate of the generalization error and its variance.

# 9 Bias-Variance decomposition

The expected test error can be expressed as the sum of two terms

- Squared bias: The average shift of the predictions
- Variance: measure how data points vary around their average.

expected loss = 
$$(bias)^2 + variance + noise$$

- Both model bias and estimation bias are important
- Bidge regression increases estimation bias while reducing variance
- Increasing model complexity increases test error
- Small λ → low bias but large variance Large λ → large bias but low variance

$$err = \sigma^2 + E[f_{lse} - E[f_{lse}]]^2 + [f_{true} - E[f_{lse}]]^2$$

# 10 Logistic Regression

- Classification relates input variables x to
- discrete output variable y Binary classifier: we use y = 0 for  $C_1$  and
- y = 1 for  $C_2$ . Can use least-squares to predict  $\hat{y}_*$

$$\hat{y} = \left\{ egin{array}{ll} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* \geq 0.5 \end{array} \right.$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{-1}$$

$$p(y_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$

$$p(y_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$$

- The probabilistic model:

- The probabilistic model: 
$$p(\mathbf{y}|\mathbf{X},w) = \prod_{n=1}^{N} \sigma(\mathbf{x}_n^T w)^{y_n} \left(1 - \sigma(\mathbf{x}_n^T w)\right)^{1-y_n}$$
 - The log-likelihood:

$$\mathcal{L}_{MLE}(\boldsymbol{w}) = \sum_{n=1}^{N} \left( y_n \boldsymbol{x}_n^T \boldsymbol{w} - \log(1 + \exp(\boldsymbol{x}_n^T \boldsymbol{w})) \right) \qquad \qquad \frac{d\mathcal{L}}{d\boldsymbol{w}} = \boldsymbol{\mathbf{X}}^T [\boldsymbol{\mathbf{g}}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\boldsymbol{\mathbf{y}})]$$

- We can use the fact that

$$\frac{d}{dx}\log(1 + \exp(x)) = \sigma(x)$$

$$\mathbf{g} = \frac{d\mathcal{L}}{d\mathbf{w}} = \sum_{n=1}^{N} \left( \mathbf{x}_n y_n - \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w}) \right)$$
$$= -\mathbf{X}^T [\sigma(\mathbf{X} \mathbf{w}) - \mathbf{y}]$$

The negative of the log-likelihood  $-\mathcal{L}_{mle}(\boldsymbol{w})$  is

- Hessian of the log-likelihood

- We know that

$$\frac{d\sigma(t)}{d\sigma(t)} = \sigma(t)(1 - \sigma(t))$$

$$\begin{aligned} \mathbf{H}(\boldsymbol{w}) &= -\frac{d\mathbf{g}(\boldsymbol{w})}{d\boldsymbol{w}^T} = \sum_{n=1}^N \frac{d}{d\boldsymbol{w}^T} \sigma(\boldsymbol{x}_n^T \boldsymbol{w}) \boldsymbol{x}_n \\ &= \sum_{n=1}^N \boldsymbol{x}_n \sigma(\boldsymbol{x}_n^T \boldsymbol{w}) (1 - \sigma(\boldsymbol{x}_n^T \boldsymbol{w})) \boldsymbol{x}_n^T \end{aligned}$$

where S is a  $N \times N$  diagonal matrix with diagonals

$$S_{nn} = \sigma(\boldsymbol{x}_n^T \boldsymbol{w}) (1 - \sigma(\boldsymbol{x}_n^T \boldsymbol{w}))$$

- The negative of the log-likelihood is not strictly convex.
- Newton's Method
- Uses second-order information and takes steps in the direction that minimizes a quadratic approximation

$$\mathcal{L}(\boldsymbol{w}) = \mathcal{L}(\boldsymbol{w}^{(k)}) + \mathbf{g}_k^T (\boldsymbol{w} - \boldsymbol{w}^{(k)})$$
$$+ (\boldsymbol{w} - \boldsymbol{w}^{(k)})^T \mathbf{H}_L (\boldsymbol{w} - \boldsymbol{w}^{(k)})$$

and it's minimum is at  $\mathbf{w}^{k+1} = \mathbf{w}^{(k)} - \alpha_k \mathbf{H}_{L}^{-1} \mathbf{g}_k$ 

where  $\mathbf{g}_{L}$  is the gradient and  $\alpha_{L}$  the - Complexity:  $O((ND^2 + D^3)I)$ 

Penalized Logistic Regression 
$$\min_{\boldsymbol{w}} \left( -\sum_{n=1}^{N} \log p(y_n | \mathbf{x}_n^T \boldsymbol{w}) + \lambda \sum_{d=1}^{D} w_d^2 \right)$$

# 11 Generalized Linear Model

- Exponential family distribution 
$$p(\mathbf{y}|\boldsymbol{\eta}) = \frac{h(y)}{Z} \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

- Bernoulli distribution
$$p(y|\mu) = \mu^y (1-\mu)^{1-y}$$

 $= \exp(y \log(\frac{\mu}{1-\mu} + \log(1-\mu)))$ - there is a relationship between  $\eta$  and  $\mu$  throught the **link function** 

the link function 
$$\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$$

 Note that μ is the mean parameter of η Relationship between the mean  $\mu$  and  $\eta$  is defined using a link function a

$$\eta = \mathbf{g}(\mu) \Leftrightarrow \mu = \mathbf{g}^{-1}(\eta)$$
  
First and second derivatives of  $A(\eta)$  are related

to the mean and the variance  $\frac{dA(\eta)}{d\eta} = \mathbb{E}[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = \operatorname{Var}[\phi(\eta)]$ 

-  $A(\eta)$  is convex - The generalized maximum likelihood cost to

 $\min_{\boldsymbol{w}} \mathcal{L}(\boldsymbol{w}) = -\sum_{n}^{N} \log(p(y_n | \boldsymbol{x}_n^T \boldsymbol{w}))$ where  $p(y_n | \boldsymbol{x}_n^T \boldsymbol{w})$  is an exponential family

$$\frac{d\mathcal{L}}{d\mathbf{w}} = \mathbf{X}^T [\mathbf{g}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\mathbf{y})]$$

# 12 k-Nearest Neighbor (k-NN)

- The k-NN prediction for x is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$$

where  $nbh_k(\mathbf{x})$  is the neighborhood of  $\mathbf{x}$  defined by the k closest points  $\mathbf{x}_n$  in the training data

- Curse of dimensionality: Generalizing correctly becomes exponentially harder as the dimensionality grows.
- Gathering more inputs variables may be a bad

### 13 Support Vector Machine

- Combination of the kernel trick plus a modified loss function (Hinge loss)
- Solution to the dual problem is sparse and
- non-zero entries will be our support vectors. Kernelised feature vector where  $\mu_k$  are

$$\phi(\mathbf{x}) = [k(\mathbf{x}, \boldsymbol{\mu}_1), ..., k(\mathbf{x}, \boldsymbol{\mu}_K)]$$

- In practice we'll take a subset of data points to be prototype  $\rightarrow$  sparse vector machine.
- Assume  $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$\mathcal{L}(\boldsymbol{w}) = \min_{\boldsymbol{w}} \sum_{n=1}^{N} [1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{w}]_+ + \frac{\lambda}{2} \sum_{j=1}^{M} w_j^2$$

- Minimum doesn't change with a rescaling of w
- choose the hyperplane so that the distance from it to the nearest data point on each side is maximized
- Duality:
- Hard to minimize g(w) so we define

$$\mathcal{L}(\boldsymbol{w}) = \max_{\boldsymbol{\alpha}} G(\boldsymbol{w}, \boldsymbol{\alpha})$$

- we use the property that 
$$C[v_n]_+ = \max(0,Cv_n) = \max_{\alpha_n \in [0,C]} \alpha_n v_n$$

$$\min_{\boldsymbol{w}} \max_{\boldsymbol{\alpha} \in [0,C]^N} \sum_{n=1}^N \alpha_n (1-y_n \boldsymbol{\phi}_n^T \boldsymbol{w}) + \frac{1}{2} \sum_{j=1}^M w_j^2 \quad \text{such that } z_{nk} \in \{0,1\} \text{ and } \sum_{k=1}^K z_{nk} = 1 \text{ where } x_{nk} \in \{0,1\} \text{ and } x_{nk} \in \{$$

- This is differentiable, convex in w and concave in  $\alpha$
- Minimax theorem:

$$\min \max_{\alpha} G(w, \alpha) = \max_{\alpha} \min_{\alpha} G(w, \alpha)$$

because G is convex in  $\boldsymbol{w}$  and concave in  $\boldsymbol{\alpha}$ .

- Derivative w.r.t. w:

$$oldsymbol{
abla}_{oldsymbol{w}}G(oldsymbol{w},oldsymbol{a}) = -\sum_{n=1}^{N}lpha_{n}y_{n}oldsymbol{x}_{n} + \lambda oldsymbol{w}$$

$$w(\alpha) = \frac{1}{\lambda} \sum_{n=1}^{N} \alpha_n y_n x_n = \frac{1}{\lambda} X Y \alpha$$

Y := diag(y)

- Plugging  $w^*$  back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,1]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2\lambda} \boldsymbol{\alpha}^T \mathbf{Y} \boldsymbol{X}^T \boldsymbol{X} \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with  $\mathbf{K} = \boldsymbol{X}^T\boldsymbol{X}$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision

#### 14 Kernel Ridge Regression

- The following is true for ridge regression

$$w^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y}$$

$$= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \boldsymbol{\alpha}^*$$
(1)

- Complexity of computing w: (1)  $O(D^2N + D^3), (2) O(DN^2 + N^3)$
- Thus we have
- $w^* = \mathbf{X}\alpha^*$ , with  $w^* \in \mathbb{R}^D$  and  $\alpha^* \in \mathbb{R}^N$
- The representer theorem allows us to write an equivalent optimization problem in terms of  $\alpha$ .

$$\boldsymbol{\alpha} = \operatorname*{argmax}_{\boldsymbol{\alpha}} \left( -\frac{1}{2} \boldsymbol{\alpha}^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right) \quad \Big|$$

- $K = XX^T$  is called the **kernel matrix** or Gram matrix.
- If K is positive definite, then it's called a Mercer Kernel.
- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$
- If the kernel is Mercer, then there exists a function  $\phi(\mathbf{x})$  s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

#### Kernel trick:

- We can work directly with K and never have to worry about X
- Replace  $\langle \mathbf{x}, \mathbf{x}' \rangle$  with  $k(\mathbf{x}, \mathbf{x}')$ .
- Kernel function can be interpreted as a measure of similarity
- The evaluation of a kernel is usually faster with k than with d
- Kernelized rigde regression might be
- computationally more efficient in some cases Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Properties of a kernel to ensure the existance of
- K should be symmetric:  $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$ - K should be positive semidefinite.

$$\mathbf{y} = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N z_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

- 1. For all n, compute  $\mathbf{z}_n$  given  $\boldsymbol{\mu}$

$$z_{nk} = \left\{ \begin{array}{ll} 1 & \text{ if } k = \mathop{\rm argmin}_j ||\mathbf{x}_n - \boldsymbol{\mu}||_2^2 \\ 0 & \text{ otherwise} \end{array} \right.$$

$$\mu_k = \frac{\sum_{n=1}^{N} z_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} z_{nk}}$$

- A good initialization procedure is to choose the prototypes to be equal to a random subset of Kdata points
- Probabilistic model

$$p(\mathbf{z}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[ \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{z_{nk}}$$

- K-means as a Matrix Factorization

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{M}\mathbf{z}^T||_{\text{Frob}}^2$$

Computation can be heavy, each example can belong to only on cluster and clusters have to

# 16 Gaussian Mixture Models

Clusters can be elliptical using a full covariance matrix instead of isotropic covariance.

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{z}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[ \mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{z_{nk}}$$

Soft-clustering: Points can belong to several

$$p(z_n = k) = \pi_k$$
 where  $\pi_k > 0, \forall k, \sum_{k=1}^K \pi_k = 1$ 

- Joint distribution of Gaussian mixture model

$$\begin{split} p(\mathbf{X}, \mathbf{z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) &= \prod_{n=1}^{N} p(\mathbf{x}_n | \mathbf{r}_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{z}_n | \boldsymbol{\pi}) \\ &= \prod_{n=1}^{N} \prod_{k=1}^{K} [(\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k))^{z_n k}] \prod_{k=1}^{K} [\boldsymbol{\pi}]^{z_n k} \end{split}$$

 Unknown parameters are given by  $\theta = \{\mu, \Sigma, \pi\}$ 

 We get the marginal likelihood by marginalizing  $z_n$  out from the likelihood

$$\begin{split} p(\mathbf{x}_n|\boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, z_n = k|\boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(z_n = k|\boldsymbol{\theta}) p(\mathbf{x}_n|z_n = k, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{split}$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to
- To get maximum likelihood estimate of  $\theta$ , we

$$\max_{\pmb{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \pmb{\Sigma}_k)$$

#### 17 Expectation Maximization Algorithm

- [ALGORITHM] Start with  $\boldsymbol{\theta}^{(1)}$  and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous  $\boldsymbol{\theta}^{(t)}$  with equality when,

$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$

$$\theta^{(t+1)} = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$$

$$\boldsymbol{\mu}_k^{(t+1)} = \frac{\sum_{n=1}^N \gamma^{(i)}(r_{nk})\mathbf{x}_n}{\sum_{n=1}^N q_{kn}^{(t)}}$$

$$\boldsymbol{\Sigma}_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} q_{kn}^{(t)} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)})}{\sum_{n=1}^{N} q_{kn}^{(t)}} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)}) \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W} \mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ are orthogonal.}$$

$$\pi_k^{(t+1)} = \frac{1}{N} \sum_{n=1}^{N} q_{kn}^{(t)}$$

If covariance is diagonal → K-means

#### 18 Matrix factorization

- We have D movies and N users
- **X** is a matrix  $D \times N$  with  $x_{dn}$  the rating of n'th user for d'th movie.
- We project data vectors  $\mathbf{x}_n$  to a smaller dimension  $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
  - **Z** a  $N \times K$  matrix that gives features for the
  - W a D × K matrix that gives features for the movies

 $x_{dn} pprox \mathbf{w}_d^T \mathbf{z}_n$  We can add a regularizer and minimize the

$$\begin{split} \mathcal{L}(\mathbf{W}, \mathbf{Z}) &= \frac{1}{2} \sum_{(d, n) \in \Omega} [x_{dn} - (\mathbf{W} \mathbf{Z}^T)_{dn}]^2 \\ &+ \frac{\lambda_w}{2} \sum_{d=1}^{D} \mathbf{w}_d^T \mathbf{w}_d + \frac{\lambda_z}{2} \sum_{n=1}^{N} \mathbf{z}_n^T \mathbf{z}_n \end{split}$$

first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_K)^{-1} \mathbf{W}^T \mathbf{X}$$
$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_K)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- Complexity:  $O(DNK^2 + NK^3) \rightarrow O(DNK^2)$ 

- Probabilistic model

$$\prod_{(d,n)\in\Omega} \mathcal{N}(x_{dn}|\mathbf{w}_d^T\mathbf{z}_n,I) \times \prod_{n=1}^N \mathcal{N}(\mathbf{z}_n|0,\frac{1}{\lambda_z}I)$$

Since many ratings are missing we cannot normalize the data. A solution is to add offset

$$\frac{1}{2} \sum_{(d,n) \in \Omega} (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu)^2$$

## 19 Singular Value Decomposition

- Matrix factorization method  $\mathbf{X} = \mathbf{U}\mathbf{S}\mathbf{V}^T$ 
  - **U** is a unitary  $D \times D$  matrix
  - $\mathbf{V}$  is a unitary  $N \times N$  matrix
- S is a non-negative diagonal matrix of size  $D \times N$  which are called singular values appearing in a descending order.
- Columns of II and V are the left and right singular vectors respectively.
- Assuming D < N we have</li>

$$\mathbf{X} = \sum_{d=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information.

Dimensionality Reduction

Take the matrix  $S^{(K)}$  with the K first diagonal elements non zero. Then, rank-K approx:

$$\mathbf{x} \approx \mathbf{x}_K = \mathbf{U}\mathbf{S}^{(K)}\mathbf{V}^T$$

# 19.1 Principal Componement Analysis

- PCA is a dimensionality reduction method and

e data is zero mean
$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$

$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Thus the columns of matrix U are called the principal components and they decorrelate the covariance matrix.
- Using SVD, we can compute the matrices in the following way

$$\mathbf{W} = \mathbf{U}\mathbf{S}_{\mathbf{D}}^{1/2}, \mathbf{Z}^T = \mathbf{S}^{1/2}\mathbf{V}^T$$

# 20 Neural Net

Basic structure: One input layer of size D, L hidden layers of size K, and one output layer.

(feedforward network).  

$$r(l) = \phi \left( \sum_{l} w^{(l)} r^{(l-1)} + b^{(l)} \right)$$

- $x_j^{(l)} = \phi\left(\sum_i w_{i,j}^{(l)} x_i^{(l-1)} + b_j^{(l)}\right).$  NN can represent the Riemmann sum with only two layers ⇒ It's powerful!
- Cost function:

$$L = \frac{1}{N} \sum_{n=1}^{N} \left( y_n - f^{(L+1)} \circ \dots \circ f^{(1)}(\mathbf{x}_n^{(0)}) \right)^2$$
We can use SGD to minimize the cost function.

- Compact form:  $\mathbf{W}_{i,j}^{(l)} = w_{i,j}^{(l)}$ 

- Compact form: 
$$\mathbf{W}_{i,j}^{r,j} = \mathbf{w}_{i,j}^{r,j}$$
  
 $\mathbf{x}^{(l)} = f^{(l)}(\mathbf{x}^{(l-1)}) = \phi\left(\left(\mathbf{W}^{(l)}\right)^{T}\mathbf{x}^{(l-1)} + \mathbf{b}^{(l)}\right)$ 

#### 20.1 Backpropagation Algorithm

- Forward pass: Compute

$$\mathbf{z}^{(l)} = \left(\mathbf{W}^{(l)}\right)^T \mathbf{z}^{(l-1)} + \mathbf{b}^{(l)} \text{ with } \mathbf{z}^{(0)} = \mathbf{z}_n$$
and  $\mathbf{z}^{(l)} = \phi(\mathbf{z}^{(l)})$ .

 $\delta^{(L+1)} = -2(y_n - \mathbf{x}^{(L+1)})\phi'(z^{(L+1)})$  (if squared loss). Then compute

$$\begin{split} \boldsymbol{\delta}_{j}^{(l)} &= \frac{\partial \mathcal{L}_{n}}{\partial z_{j}^{(l)}} = \sum_{k} \frac{\partial \mathcal{L}_{n}}{\partial z_{k}^{(l+1)}} \frac{\partial z_{k}^{(l+1)}}{\partial z_{j}^{(l)}} \\ &= \sum_{l} \boldsymbol{\delta}_{k}^{(l+1)} \boldsymbol{W}_{j,k}^{(l+1)} \boldsymbol{\phi}'(\boldsymbol{z}_{j}^{(l)}) \end{split}$$

$$\begin{aligned} & n = 1 & \lambda z \\ & \times \prod_{d=1}^{D} \mathcal{N}(\mathbf{w}_{d}|0, \frac{1}{\lambda_{w}}I) \\ & \text{missing we cannot olution is to add offset} \end{aligned} \quad \begin{aligned} & = \sum_{k} \delta_{k}^{(\ell+1)} \mathbf{W}_{j,k}^{(\ell+1)} \phi'(z_{j}^{(\ell)}) \\ & - \text{Final Computation:} \\ & \frac{\partial \mathcal{L}_{n}}{\partial w_{i,j}^{(\ell)}} = \sum_{k} \frac{\partial \mathcal{L}_{n}}{\partial z_{k}^{(\ell)}} \frac{\partial z_{k}^{(\ell)}}{\partial w_{i,j}^{(\ell)}} = \frac{\partial \mathcal{L}_{n}}{\partial z_{j}^{(\ell)}} \frac{\partial z_{j}^{(\ell)}}{\partial w_{i,j}^{(\ell)}} \\ & = \delta_{j}^{(\ell)} \mathbf{w}_{i,j}^{(\ell-1)} \end{aligned}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}} \\ &= \delta_j^{(l)} \cdot 1 = \delta_j^{(l)} \end{split}$$

#### 20.2 Activation Functions

Sigmoid  $\phi(x) = \frac{1}{1+e^{-x}}$  Positive, bounded.

$$\phi'(x) \simeq 0$$
 for large  $|x| \Rightarrow$  Learning slo

 $\phi'(x) \simeq 0 \text{ for large } |x| \Rightarrow \text{Learning slow}.$   $\textbf{Tanh} \quad \tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = \phi(2x) - 1/2.$ Balanced, bounded. Learning slow too.

ReLU  $(x)_{+} = \max 0, x$  Positive, unbounded Derivate = 1 if x > 0, 0 if x < 0

Leaky ReLU  $f(x) = \max \alpha x, x$  Remove 0

derivative.  $\begin{aligned} \mathbf{Maxout} \quad f(x) &= max\mathbf{x}^T\mathbf{w}_1 + b_1, \dots, \mathbf{x}^T\mathbf{w}_k + b_k \\ & \text{(Generalization of ReLU)} \end{aligned}$ 

#### 20.3 Convulctional NN WHAT CAN I SAY???

20.4 Reg, Data Augmentation and Dropout DO WE NEED SOMETHING IN THIS??

# 21 Bayes Net

- Graph example: p(x, y, z) = p(y|x)p(z|x)p(x):  $(y \leftarrow x \rightarrow z)$
- D-Separation X and Y are D-separated by Z if every path from  $x \in X$  to  $y \in Y$  is blocked by
- Blocked Path if the path contains a variable
  - is in Z and is head-to-tail or tail-to-tail - the node is head-to-head and neither the
- node nor the descendant are in Z. Markov Blanket (which blocks node A from
  - the rest of the net) contains:
  - parents of A
  - children of A - parents of children of A