PCML Cheat Sheet

1 Math Prerequisites

- Bayes rule

$$p(A,B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

- Bayes rule: $P(A|B) \propto P(B|A)P(A)$
- Gaussian distribution

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})) \text{ definite: } \mathbf{H}(\boldsymbol{w}^*) = \frac{d^2 \mathcal{L}(\boldsymbol{w}^*)}{d\boldsymbol{w} d\boldsymbol{w}^T} - \text{Very sensitive to illconditioning}$$

- Production of independent variables:
- $Var(XY) = \mathbb{E}(X^2) \mathbb{E}(Y^2) [\mathbb{E}(X)]^2 [\mathbb{E}(Y)]^2$ Covariance matrix of a data vector x

$$\Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \mathbb{E}(\mathbf{x})) (\mathbf{x}_n - \mathbb{E}(\mathbf{x}))^T$$

1.1 Convexity

- The Hessian of a convex function is psd and for a strictly-convex function it's pd.

$$\mathbf{H}_{i,j} = d^2 f / dx_i dx_j$$

1.2 Linear Algebra

- Condition number If A is normal $(A^T A = AA^T)$ then

$$k(\mathbf{A}) = \left| \frac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

2 Cost functions

Mean square error (MSE):

$$MSE(\boldsymbol{w}) = \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

Mean Absolute Error (MAE):

$$MAE = \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \leq \delta \\ \delta|z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

$$Hinge = [1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$$

Logistic loss

$$Logistic = log(1 - exp(y_n f(\mathbf{x}_n)))$$

3 Optimization

3.1 Grid search

 Compute the cost over a grid of M points to find the minimum. Exponential Complexity. Hard to find a good range of values

3.2 Gradient Descent

- GD uses only first-order information and takes steps in the opposite direction of the gradient
- Given cost function $\mathcal{L}(w)$ we want to find w $w = \arg\min_{\boldsymbol{x}} \mathcal{L}(w)$

3.3 Batch Gradient Descent

 Take steps in the opposite direction of the gradient

$$\boldsymbol{w}^{(t+1)} \leftarrow \boldsymbol{w}^{(t)} - \gamma \nabla \mathcal{L}(\boldsymbol{w}^{(t)})$$

with $\gamma > 0$ the learning rate.

With γ too big, method might diverge. With γ too small, convergence is slow.

3.4 Gradients for MSE

We define the error vector e:

$$e := y - Xw$$

and MSE as follows:

$$\mathcal{L}(\boldsymbol{w}) = \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\mathbf{x}}_n^T \boldsymbol{w})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

- then the gradient is given by

$$\nabla \mathcal{L}(\boldsymbol{w}) = -\frac{1}{N} \mathbf{X}^T \mathbf{e}$$

- Optimality conditions:
- 1. necessary: gradient equal zero: $\frac{d\mathcal{L}(\boldsymbol{w}^*)}{d\mathcal{L}(\boldsymbol{w}^*)} = 0$
- 2. sufficient: Hessian matrix is positive
- (μ)) definite: $\mathbf{H}(\mathbf{w}^+) = \frac{1}{d\mathbf{w}d\mathbf{w}^T}$ Very sensitive to illconditioning \Rightarrow always normalize features.
- Complexity: O(NDI) with I the number of

3.5 Stochastic Gradient Descent

In ML, most cost functions are formulated as a sum over the training examples:

$$\mathcal{L}(\boldsymbol{w}) = \frac{1}{N} \sum_{n=1}^{N} \mathcal{L}_n(\boldsymbol{w})$$

⇒ SGD algo is given by update rule:

$$\boldsymbol{w}^{(t+1)} \leftarrow \boldsymbol{w}^{(t)} - \gamma \nabla \mathcal{L}_n(\boldsymbol{w}^{(t)})$$

Idea: Cheap but unbiased estimate of grad.

$$\mathbb{E}[\boldsymbol{\nabla}\mathcal{L}_n(\boldsymbol{w})] = \boldsymbol{\nabla}\mathcal{L}_n(\boldsymbol{w})$$

3.6 Mini-batch SGD

Update direction $(B \subset [N])$:

$$g^{(t)} := \frac{1}{|B|} \sum_{n \in B} \nabla \mathcal{L}_n(w^{(t)})$$

Update rule : $\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \mathbf{g}^{(t)}$

3.7 Subgradients (Non-Smooth optim) A vector $\boldsymbol{a} \in \mathbb{R}^D$ s.t.

$$\mathcal{L}(oldsymbol{u}) \geq \mathcal{L}(oldsymbol{w}) + oldsymbol{g}^T(oldsymbol{u} - oldsymbol{w}) \quad orall oldsymbol{u}$$

is the subgradient to \mathcal{L} at w. If \mathcal{L} is differentiable at w, we have $g = \nabla \mathcal{L}(w)$

4 Least Squares

- Use the first optimality conditions:

$$\nabla L(\boldsymbol{w}^*) = 0 \Rightarrow \boldsymbol{X}^T \boldsymbol{e} = \boldsymbol{X}^T (\boldsymbol{v} - \boldsymbol{X} \boldsymbol{w}) = 0$$

 When X^TX is invertible, we have the closed-form expression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{v}$$

- thus we can predict values for a new \mathbf{x}_m $y_m := \mathbf{x}_m^T \mathbf{w}^* = \mathbf{x}_m^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$

- The Gram matrix X^TX is pd and is also
- invertible iff X has full column rank.
- Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$
- **X** can be rank deficient when D > N or when the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. \Rightarrow matrix is ill-conditioned.

5 Maximum Likelihood

Let define our mistakes ε_n ~ N(0, σ²).

$$\to y_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

Another way of expressing this:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} p(y_n|\mathbf{x}_n, \mathbf{w})$$
$$= \prod_{n=1}^{N} \mathcal{N}(y_n|\mathbf{x}_n^T \mathbf{w}, \sigma^2)$$

which defines the likelihood of observating y given \mathbf{X} and \boldsymbol{w}

Define cost with log-likelihood $\mathcal{L}_{lik}(\boldsymbol{w}) = \log p(\mathbf{y}|\mathbf{X}, \boldsymbol{w})$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \mathbf{x}_n^T \mathbf{w})^2 + cnst$$

Maximum likelihood estimator (MLE) gives another way to design cost functions $\operatorname{argmin} \mathcal{L}_{MSE}(\boldsymbol{w}) = \operatorname{argmax} \mathcal{L}_{lik}(\boldsymbol{w})$

MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.

With Laplace distribution

$$p(y_n|\mathbf{x}_n, \mathbf{w}) = \frac{1}{2b} e^{-\frac{1}{b}|y_n - \mathbf{x}_n^T \mathbf{w}|}$$
$$\sum \log p(y_n|\mathbf{x}_n, \mathbf{w}) = \sum |y_n - \mathbf{x}_n^T \mathbf{w}| + cnst$$

6 Ridge Regression

Basis functions:

$$y_n = w_0 + \sum_{j=1}^{M} w_j \phi_j(\mathbf{x}_n) = \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{w}$$

- This model is linear in w but nonlinear in xDimension is now M, not N.
- Polynomial basis

$$\phi(x_n) = [1, x_n, x_n^2, ..., x_n^M]$$
 – The least square solution becomes

$$\min_{\boldsymbol{w}} \left(\mathcal{L}(\boldsymbol{w}) + \frac{\lambda}{2N} \sum_{j=1}^{M} w_j^2 \right)$$

$$\boldsymbol{w}^* = \operatorname*{argmin}_{\boldsymbol{w}} \left(\frac{1}{2} (\mathbf{y} - \mathbf{X} \boldsymbol{w})^T (\mathbf{y} - \mathbf{X} \boldsymbol{w}) + \frac{\lambda}{2} \boldsymbol{w}^T | \boldsymbol{w} \right)$$

- Note that w₀ is not penalized.
- By differentiating and setting to zero we get

$$egin{aligned} oldsymbol{w}_{ridge} &= (\mathbf{ ilde{\Phi}}^T\mathbf{ ilde{\Phi}} + \mathbf{\Lambda})^{-1}\mathbf{ ilde{\Phi}}^T\mathbf{y} \ oldsymbol{\Lambda} &= \left[egin{array}{c} 0 & 0 \ 0 & \lambda I_M \end{array}
ight] \end{aligned}$$

- Ridge regression improves the condition number of the Gram matrix since the eigenvalues of $(\tilde{\Phi}^T \tilde{\Phi} + \lambda I_m)$ are at least λ : SVD $\tilde{\Phi}^T \tilde{\Phi} = USU^T$ with S = diag(>0), then $\tilde{\Phi}^T \tilde{\Phi} + \Lambda = U(S + \Lambda)U^T$ has eigenvalues $> \lambda$.
- Maximum-a-posteriori (MAP)
 - Maximizes the product of the likelihood and the prior.
 - $\boldsymbol{w}_{MAP} = \operatorname{argmax} \left(p(\mathbf{y}|\mathbf{X}, \boldsymbol{\Lambda}) p(\boldsymbol{w}|\boldsymbol{\Sigma}) \right)$
 - Assume $w_0 = 0$

$$\boldsymbol{w}_{ridge} = \underset{\boldsymbol{w}}{\operatorname{argmax}} \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_{n} | \mathbf{x}_{n}^{T} \boldsymbol{w}, \boldsymbol{\Lambda}) \times \mathcal{N}(\boldsymbol{w} | \boldsymbol{0}, \boldsymbol{\mathbf{I}}) \right] \right) \begin{cases} S_{nn} = \sigma(\boldsymbol{x}_{n}^{T} \boldsymbol{w})(1 - \sigma(\boldsymbol{x}_{n}^{T} \boldsymbol{w})) \\ \text{The negative of the log-likelihood is not} \end{cases}$$

 Lasso regularizer forces some w_i to be strictly 0 and therefore forces sparsity in the

$$\min_{\boldsymbol{w}} \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{w})^2,$$
such that $\sum_{i=1}^{M} |w_i| \le \tau$

7 Bias-Variance decomposition - The expected test error can be expressed as the sum of two terms

- $expected loss = (bias)^2 + variance + noise$
- Model bias and estimation bias are important RR increases estimation bias and reduces var Model more complex increases test error

- Small λ → low bias but large variance Large λ → large bias but low variance $err = \sigma^2 + \mathbb{E}[f_{lse} - \mathbb{E}[f_{lse}]]^2 + [f_{true} - \mathbb{E}[f_{lse}]]^2$
- 8 Logistic Regression
- Classification relates input variables x to discrete output variable y
- Binary classifier: we use y = 0 for C_1 and y = 1 for \mathbf{C}_2 .

Can use least-squares to predict \(\hat{y}_*\)

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5\\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(y_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$
$$p(y_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$$

$$p(\mathbf{y}|\mathbf{X}, \boldsymbol{w}) = \prod_{n=1}^{N} \sigma(\boldsymbol{x}_{n}^{T} \boldsymbol{w})^{y_{n}} (1 - \sigma(\boldsymbol{x}_{n}^{T} \boldsymbol{w}))^{1 - y_{n}}$$

The log-likelihood:
$$\mathcal{L}_{MLE}(\boldsymbol{w}) = \sum_{n=1}^{N} \left(y_n \boldsymbol{x}_n^T \boldsymbol{w} - \log(1 + \exp(\boldsymbol{x}_n^T \boldsymbol{w})) \right)$$

- We can use the fact that
$$\frac{d}{dx}\log(1+\exp(x)) = \sigma(x)$$
- Gradient of the log-likelihood

$$\mathbf{g} = \frac{d\mathcal{L}}{d\mathbf{w}} = \sum_{n=1}^{N} \left(\mathbf{x}_n y_n - \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w}) \right)$$

 $= -\mathbf{X}^T [\sigma(\mathbf{X}\boldsymbol{w}) - \mathbf{y}]$ - The negative of the log-likelihood $-\mathcal{L}_{mle}(\boldsymbol{w})$

is convex Hessian of the log-likelihood

We know that

- we know that
$$\frac{d\sigma(t)}{dt} = \sigma(t)(1-\sigma(t))$$
 - Hessian is the derivative of the gradient

$$\begin{aligned} \mathbf{H}(\boldsymbol{w}) &= -\frac{d\mathbf{g}(\boldsymbol{w})}{d\boldsymbol{w}^T} = \sum_{n=1}^N \frac{d}{d\boldsymbol{w}^T} \sigma(\boldsymbol{x}_n^T \boldsymbol{w}) \boldsymbol{x}_n \\ &= \sum_{n=1}^N \boldsymbol{x}_n \sigma(\boldsymbol{x}_n^T \boldsymbol{w}) (1 - \sigma(\boldsymbol{x}_n^T \boldsymbol{w})) \boldsymbol{x}_n^T \end{aligned}$$

where **S** is a $N \times N$ diagonal matrix with

Newton's Method - Uses second-order information and takes

steps in the direction that minimizes a quadratic approximation $\mathcal{L}(\boldsymbol{w}) = \mathcal{L}(\boldsymbol{w}^{(k)}) + \nabla \mathcal{L}_{b}^{T}(\boldsymbol{w} - \boldsymbol{w}^{(k)})$

$$\mathcal{L}(\boldsymbol{w}) = \mathcal{L}(\boldsymbol{w}^{(k)}) + \nabla \mathcal{L}_k^* (\boldsymbol{w} - \boldsymbol{w}^{(k)})$$

 $+ \frac{1}{2} (\boldsymbol{w} - \boldsymbol{w}^{(k)})^T \mathbf{H}_k (\boldsymbol{w} - \boldsymbol{w}^{(k)})$

 $\nabla \mathcal{L}(w) = \nabla \mathcal{L}(w^{(k)}) + H_k(w - w^{(k)})$ equating to 0 gives the minimum:

 $\boldsymbol{w}^{k+1} = \boldsymbol{w}^{(k)} - \gamma_k \boldsymbol{H}_k^{-1} \nabla \mathcal{L}_k$ - γ_k added because it's all approximative.

 $\min_{\boldsymbol{w}} \left(-\sum_{n=1}^{\infty} \log p(y_n | \mathbf{x}_n^T \boldsymbol{w}) + \lambda \sum_{d=1}^{\infty} w_d^2 \right)$

- Complexity: $O((ND^2 + D^3)I)$ Penalized Logistic Regression

9 Generalized Linear Model

Exponential family distribution
$$p(\mathbf{y}|\mathbf{\eta}) = h(y) \exp(\mathbf{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\mathbf{\eta}))$$

$$p(\mathbf{y}|\boldsymbol{\eta}) = h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$
$$A(\boldsymbol{\eta}) = \ln\left[\int h(y)e^{\boldsymbol{\eta}^T \boldsymbol{\phi}(y)} dy\right] < \infty$$

is convex by Holder.

- Bernoulli distribution $p(y|\mu) = \mu^y (1-\mu)^{1-y}$

$$= \exp(y \log(\frac{\mu}{1-\mu} + \log(1-\mu)))$$

there is a relationship between η and μ throught the **link function**

$$\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$$

- Note that μ is the mean parameter of y- Relationship between the mean $\boldsymbol{\mu}$ and $\boldsymbol{\eta}$ is
- defined using a link function q

$$\eta = \mathbf{g}(\boldsymbol{\mu}) \Leftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta})$$

First and second derivatives of A(n) are related to the mean and the variance

$$\frac{dA(\eta)}{d\eta} = \mathbb{E}[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = \text{Var}[\phi(\eta)]$$

The generalized maximum likelihood cost to

$$\min_{\boldsymbol{w}} \mathcal{L}(\boldsymbol{w}) = -\sum_{n=1}^{N} \log(p(y_n | \boldsymbol{x}_n^T \boldsymbol{w}))$$

where $p(y_n | \boldsymbol{x}_n^T \boldsymbol{w})$ is an exponential family distribution

We obtain the solution

we obtain which solution
$$\frac{d\mathcal{L}}{d\boldsymbol{w}} = \mathbf{X}^T[\mathbf{g}^{-1}(\boldsymbol{X}\boldsymbol{w}) - \boldsymbol{\phi}(\mathbf{y})]$$
 since $\frac{dA(\eta)}{d\eta} = g^{-1}(\eta)$.

10 k-Nearest Neighbor (k-NN)

– The k-NN prediction for \mathbf{x} is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$$

where $nbh_k(\mathbf{x})$ is the neightborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n . For classification take majority vote instead of

Curse of dimensionality: Let $\mathcal{X} = [0, 1]^d$ $\mathcal{Y} = \{0, 1\}, S \sim \mathcal{D}^N, \eta(x) = p(y = 1 \mid x).$ If \mathcal{D} is known, best classifier is Bayes

$$f_*(x) = \mathbf{1}[\eta(x) > \frac{1}{2}]$$
. If
$$|\eta(x) - \eta(x')| \le c||x - x'||$$
 then, for f_S the 1-nearest neighbor classifier,

then, for
$$f_S$$
 the 1-nearest neighbor classifier, $\mathbb{E}[\mathcal{L}(f_S)] \leq 2\mathcal{L}(f_*) + c\mathbb{E}_{S,x \sim \mathcal{D}}[||x - nbh_1(x)||]$

11 Support Vector Machine

- Assume $y_n \in \{-1, 1\}$

Combination of the kernel trick plus a modified loss function (Hinge loss)

Solution to the dual problem is sparse and non-zero entries will be our support

vectors Kernelised feature vector where μ_k are centroids

voids
$$\phi(\mathbf{x}) = [k(\mathbf{x}, \boldsymbol{\mu}_1), ..., k(\mathbf{x}, \boldsymbol{\mu}_K)]$$

 In practice we'll take a subset of data points to be prototype \rightarrow sparse vector machine.

- SVM optimizes the following cost
$$\mathcal{L}(\boldsymbol{w}) = \min_{\boldsymbol{w}} \sum_{n=1}^{N} [1 - y_n \tilde{\phi}_n^T \boldsymbol{w}]_+ + \frac{\lambda}{2} \sum_{i=1}^{M} w_j^2$$

Loss is convex but not diff. Can be optim with subgradient descent.

choose the hyperplane so that the distance from it to the nearest data point on each side is maximized

- Duality:

- Hard to minimize g(w) so we define

$$\mathcal{L}(\boldsymbol{w}) = \max_{\boldsymbol{\alpha}} G(\boldsymbol{w}, \boldsymbol{\alpha})$$

- we use the property that $C[v_n]_+ = \max(0, Cv_n) = \max_{\alpha_n \in [0, C]} \alpha_n v_n$

- We can rewrite the problem as

$$\min_{\boldsymbol{w}} \max_{\boldsymbol{\alpha} \in [0,C]} \sum_{n=1}^{N} \alpha_n (1 - y_n \boldsymbol{\phi}_n^T \boldsymbol{w}) + \frac{1}{2} \sum_{j=1}^{M} w_j^{\text{23}} \quad \text{K-means}$$
- This is differentiable, convex in \boldsymbol{w} and

concave in α

Minimax theorem:

 $\min_{\boldsymbol{\alpha}} \max_{\boldsymbol{\alpha}} G(\boldsymbol{w}, \boldsymbol{\alpha}) = \max_{\boldsymbol{\alpha}} \min_{\boldsymbol{\alpha}} G(\boldsymbol{w}, \boldsymbol{\alpha})$

because G is convex in \boldsymbol{w} and concave in

- Derivative w.r.t. w:

$$\nabla_{\pmb{w}} G(\pmb{w},\pmb{\alpha}) = -\sum_{n=1}^N \alpha_n y_n \pmb{x}_n + \lambda \pmb{w}$$
 – Equating this to 0, we get:

$$w(\alpha) = \frac{1}{\lambda} \sum_{n=1}^{N} \alpha_n y_n x_n = \frac{1}{\lambda} X Y \alpha$$

Y := diag(y)

- Plugging w^* back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,1]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2\lambda} \boldsymbol{\alpha}^T \mathbf{Y}^T \boldsymbol{X}^T \boldsymbol{X} \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{X}^T \mathbf{X}$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision
- $\alpha_n = 0$ if x_n i far from the boundary in the right side. $\alpha_n \in (0,1)$ if x_n is exactly on the margin and $\alpha_n = 1$ if x_n is in the margin or on the wrong side.

12 Kernel Ridge Regression

- The following is true for ridge regression

$$w^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y}$$

$$= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \boldsymbol{\alpha}^*$$
(1)

since $(PQ + I_N)^{-1}P = P(QP + I_M)^{-1}$.

Complexity of computing \boldsymbol{w} : (1)

 $O(D^2N + D^3)$, (2) $O(DN^2 + N^3)$ Thus we have

 $\boldsymbol{w}^* = \mathbf{X} \boldsymbol{\alpha}^*, \quad \text{with } \boldsymbol{w}^* \in \mathbb{R}^D \text{ and } \boldsymbol{\alpha}^* \in \mathbb{R}^N$

The representer theorem allows us to write an equivalent optimization problem in terms

$$\alpha = \underset{\boldsymbol{\alpha}}{\operatorname{argmax}} \left(-\frac{1}{2} \boldsymbol{\alpha}^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

 $K = XX^T$ is called the **kernel matrix** or Gram matrix

If K is positive definite, then it's called a

Mercer Kernel.

- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$ - If the kernel is Mercer, then there exists a function $\phi(\mathbf{x})$ s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

Kernel trick:

- compute dot-product in R^m while remaining in \mathbb{R}^n
- Replace (x, x') with k(x, x')

- Common Kernel

- Polynomial Kernel: $(\gamma \langle \mathbf{x}_i, \mathbf{x}_i \rangle + r)^d$
- Radial Basis function kernel (RBF)
- $k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} \mathbf{x}')^T(\mathbf{x} \mathbf{x}'))$
- Sigmoid Kernel: $tanh(\langle \mathbf{x}_i, \mathbf{x}_i \rangle + r)$

- Properties of kernels to ensure the existance of a corresponding ϕ :
- symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$ - positive semi-definite.

$$\mathbf{y} = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances outside the cluster.

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N z_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

such that $z_{nk} \in \{0,1\}$ and $\sum_{k=1}^{K} z_{nk} = 1$ K-means algorithm (Coordinate Descent): Initialize μ_k , then iterate

1. For all n, compute \mathbf{z}_n given $\boldsymbol{\mu}$ $z_{nk} = \left\{ \begin{array}{ll} 1 & \text{if } k = \mathop{\mathrm{argmin}}_j ||\mathbf{x}_n - \boldsymbol{\mu}||_2^2 \\ 0 & \text{otherwise} \end{array} \right.$

2. For all k, compute μ_k given ${\bf z}$

$$u_k = \frac{\sum_{n=1}^{N} z_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} z_{nk}}$$

- $\mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}}$ A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{z}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{z_{nk}}$$

- K-means as a Matrix Factorization $\min \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{M}\mathbf{Z}^T||_{\text{Erob}}^2$

Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

14 Gaussian Mixture Models

Clusters can be elliptical using a full covariance matrix instead of isotropic

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{z}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{z_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining z_n to be a random

$$p(z_n = k) = \pi_k$$
 where $\pi_k > 0, \forall k, \sum_{k=1}^K \pi_k = 1$

- Joint distribution of Gaussian mixture model

$$\begin{split} p(\mathbf{X}, \mathbf{z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) &= \prod_{n=1}^{N} p(\mathbf{x}_n | \mathbf{r}_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{z}_n | \boldsymbol{\pi}) \\ &= \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\left(\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right)^{z_{nk}} \right] \prod_{k=1}^{K} \left[\boldsymbol{\pi} \right]^{z_{nk}} \end{split}$$

- z_n are called *latent* unobserved variables
- Unknown parameters are given by
- $\theta = \{\mu, \Sigma, \pi\}$
- We get the marginal likelihood by marginalizing z_n out from the likelihood

marginalizing
$$z_n$$
 out from the menhood
$$p(\mathbf{x}_n|\boldsymbol{\theta}) = \sum_{k=1}^K p(\mathbf{x}_n, z_n = k|\boldsymbol{\theta})$$
$$= \sum_{k=1}^K p(z_n = k|\boldsymbol{\theta})p(\mathbf{x}_n|z_n = k, \boldsymbol{\theta})$$
$$= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to $O(D^2K)$
- To get maximum likelihood estimate of θ , we

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

15 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\boldsymbol{\theta}^{(t)}$ with equality when,

$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$

this is the posterior of $z_n = k$ given x_n, θ . 2. Maximization step: Update θ

$$oldsymbol{ heta}^{(t+1)} = \operatorname{argmax} \mathcal{L}(oldsymbol{ heta}, oldsymbol{ heta}^{(t)})$$

$$\boldsymbol{\mu}_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk}) \mathbf{x}_{n}}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\boldsymbol{\Sigma}_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} q_{kn}^{(t)}(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)})(\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)})^{T}}{\sum_{n=1}^{N} q_{kn}^{(t)}} \quad \begin{array}{c} \text{Principal Componement Analysis} \\ - \text{PCA is a dimensionality reduction method and a method to decorrelate the data} \\ \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \text{Otherwise} \quad \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \mathbf{X} \approx \mathbf{X} = \mathbf{W}\mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \mathbf{X} \approx \mathbf{X} = \mathbf{X} \mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \mathbf{X} \approx \mathbf{X} \approx \mathbf{X} = \mathbf{X} \mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \mathbf{X} \approx \mathbf{X} \approx \mathbf{X} = \mathbf{X} \mathbf{Z}^{T} \text{ such that columns of } \mathbf{W} \text{ and on the order} \\ - \mathbf{X} \approx \mathbf$$

If covariance is diagonal → K-means

16 Matrix factorization

- We have D movies and N users
- \mathbf{X} is a matrix $D \times N$ with x_{dn} the rating of n'th user for d'th movie.
- We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
 - \mathbf{Z} a $N \times K$ matrix that gives features for the users
 - \mathbf{W} a $D \times K$ matrix that gives features for the movies

$$x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$$

- Non-convex and non-identifiable $(W = \beta W, Z = \frac{1}{\beta} Z).$
- We can add a regularizer and minimize the

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{(d,n) \in \Omega} [x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}]^2$$

$$+rac{\lambda_w}{2}\sum_{d=1}^D\mathbf{w}_d^T\mathbf{w}_d+rac{\lambda_z}{2}\sum_{n=1}^N\mathbf{z}_n^T\mathbf{z}_n$$

We can use coordinate descent algorithm, by first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_K)^{-1} \mathbf{W}^T \mathbf{X}$$

$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T\mathbf{Z} + \lambda_w \mathbf{I}_K)^{-1}\mathbf{Z}^T\mathbf{X}^T$$

- Complexity: $O(DNK^2 + NK^3) \rightarrow O(DNK^2)$
 - $\prod_{(d,n)\in\Omega} \mathcal{N}(\boldsymbol{x}_{dn}|\mathbf{w}_d^T\mathbf{z}_n,I) \times \prod_{n=1}^N \mathcal{N}(\mathbf{z}_n|0,\frac{1}{\lambda_z}I) \bigg|$

$$\times \prod_{d=1}^{\infty} \mathcal{N}(\mathbf{w}_d | 0, \frac{1}{\lambda_w} I)$$
missing we cannot

Since many ratings are missing we cannot normalize the data. A solution is to add $\frac{1}{2} \sum \left(x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu \right)^2$

17 Singular Value Decomposition

Matrix factorization method

$$X = USV^T$$

- \mathbf{U} is a unitary $D \times D$ matrix
- \mathbf{V} is a unitary $N \times N$ matrix
- S is a non-negative diagonal matrix of size $D \times N$ which are called singular values appearing in a descending order.
- Columns of U and V are the left and right singular vectors respectively.
- Assuming D < N we have

$$\mathbf{X} = \sum_{d=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information. Dimensionality Reduction

Take the matrix $S^{(K)}$ with the K first diagonal elements non zero. Then, rank-K

$$\mathbf{X} \approx \mathbf{X}_K = \mathbf{U}\mathbf{S}^{(K)}\mathbf{V}^T$$

PCA is a dimensionality reduction method

and a method to decorrelate the data $\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$ such that columns of \mathbf{W} are orthogonal.

If the data is zero mean

$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^{T} \Rightarrow \mathbf{X} \mathbf{X}^{T} = \mathbf{U} \mathbf{S}^{2} \mathbf{U}^{T}$$
$$\Rightarrow \mathbf{U}^{T} \mathbf{X} \mathbf{X}^{T} \mathbf{U} = \mathbf{U}^{T} \mathbf{U} \mathbf{S}^{2} \mathbf{U}^{T} \mathbf{U} = \mathbf{S}^{2}$$

- Thus the columns of matrix U are called the principal components and they decorrelate the covariance matrix.
- Using SVD, we can compute the matrices in the following way

$$\mathbf{W} = \mathbf{U}\mathbf{S}_D^{1/2}, \mathbf{Z}^T = \mathbf{S}^{1/2}\mathbf{V}^T$$

18 Neural Net

- Basic structure: One input layer of size D, L hidden layers of size K, and one output layer. (feedforward network).

$$x_{j}^{(l)} = \phi \left(\sum_{i} w_{i,j}^{(l)} x_{i}^{(l-1)} + b_{j}^{(l)} \right)$$

- Lemma: For
$$f: \mathbb{R}^D \to \mathbb{R}$$
 with $\int |w| |\tilde{f}(w)| dw < C$ with

$$\tilde{f}(w) = \int f(x)e^{-iw^Tx} dx$$
 is the Fourier

$$\begin{split} \tilde{f}(w) &= \int f(x) e^{-iw^T x} dx \text{ is the Fourier} \\ \text{transform. Then } \forall n > 0, \\ \exists f_n(x) &= \sum_{i=1}^n c_i \sigma(x^t w_i + b_i) + c_0 \text{ so that} \end{split}$$

$$\int_{|x| < r} (f - f_n)^2 \le \frac{(2Cr)^2}{n}, \text{ with } \sigma$$
 sigmoid-like

Cost function:

$$\mathcal{L} = \frac{1}{N} \sum_{n=1}^{N} \left(y_n - f^{(L+1)} \circ \dots \circ f^{(1)}(\boldsymbol{x}_n^{(0)}) \right)^2$$
We can use SGD to minimize the cost

function. Compact form:
$$\mathbf{W}_{i,j}^{(l)} = w_{i,j}^{(l)}$$

 $\mathbf{x}^{(l)} = f^{(l)}(\mathbf{x}^{(l-1)}) = \phi \left(\left(\mathbf{W}^{(l)} \right)^T \mathbf{x}^{(l-1)} + \mathbf{b}^{(l)} \right)$

18.1 Backpropagation Algorithm

- Forward pass: Compute

- Forward pass: Compute
$$\mathbf{z}^{(l)} = \left(\mathbf{W}^{(l)}\right)^T \mathbf{x}^{(l-1)} + \mathbf{b}^{(l)} \text{ with }$$

$$\begin{split} \boldsymbol{\delta}_{j}^{(l)} &= \frac{\partial \mathcal{L}_{n}}{\partial z_{j}^{(l)}} = \sum_{k} \frac{\partial \mathcal{L}_{n}}{\partial z_{k}^{(l+1)}} \frac{\partial z_{k}^{(l+1)}}{\partial z_{j}^{(l)}} \\ &= \sum_{l} \boldsymbol{\delta}_{k}^{(l+1)} \boldsymbol{W}_{j,k}^{(l+1)} \boldsymbol{\phi}'(z_{j}^{(l)}) \end{split}$$

Final Computation:

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}} \\ &= \delta_j^{(l)} \mathbf{x}_i^{(l-1)} \end{split}$$

$$\frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}}$$
$$= \delta_i^{(l)} \cdot 1 = \delta_i^{(l)}$$

18.2 Activation Functions

Sigmoid $\phi(x) = \frac{1}{1+e^{-x}}$ Positive, bounded

$$\phi'(x) \simeq 0$$
 for large $|x| \Rightarrow$ Learning slow.

Tanh $\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = 2\phi(2x) - 1.$

Balanced, bounded. Learning slow too. **ReLU** $(x)_{+} = \max\{0, x\}$ Positive, unbounded.

Derivate = 1 if x > 0, 0 if x < 0Leaky ReLU $f(x) = \max\{\alpha x, x\}$ Remove 0 derivative.

Maxout

 $f(x) = max\{\boldsymbol{x}^T \boldsymbol{w}_1 + b_1, ..., \boldsymbol{x}^T \boldsymbol{w}_k + b_k\}$ (Generalization of ReLU)

18.3 Convolutional NN

Sparse connections and weights sharing: reduce complexity. (e.g. pixels in pictures only depend on neighbours)

18.4 Reg, Data Augmentation and Dropout

- Regularization term: $\frac{1}{2}\sum_{l=1}^{L+1}\mu^{(l)}||W^{(l)}||_F^2$ Data Augm.: e.g. shift or rotation of pics
- Dropout: avoid overfit. Drop nodes randomly. (Then average multiple drop-NN)

19 Bayes Net

- Graph example: p(x, y, z) = p(y|x)p(z|x)p(x) $: (y \leftarrow x \rightarrow z)$
- Lemma: X and Y are independent given Z iff X and Y are D-separated by Z.
- **D-Separation** X and Y are D-separated by Z if every path from $x \in X$ to $y \in Y$ is blocked by Z.
- Blocked Path contains a variable that
- is in Z and is **head-to-tail** $(x \to z \to y)$ or tail-to-tail $(x \leftarrow z \rightarrow y)$.
- the node is **head-to-head** $(x \rightarrow a \leftarrow y)$ and neither the node nor the descendant
- are in Z. Markov Blanket (which blocks node A from the rest of the net) contains:
- parents of A
- children of A
- parents of children of A
- Lemma: Any x is independent of y ∉ MB(X) given MB(x).

 $f(x_1,...,x_6) =$

 $f_1(x_1, x_2, x_3) f_2(x_1, x_4, x_6) f_3(x_4) f_4(x_4, x_5),$ $\left[\sum_{x_2,x_3} f_1(x_1,x_2,x_3)\right] \left[\sum_{x_4} f_3(x_4) \left(\sum_{x_6} f_2(x_1,x_2,x_3)\right)\right] \left[\sum_{x_6} f_3(x_6) \left(\sum_{x_6} f_2(x_1,x_2,x_3)\right)\right] \left[\sum_{x_6} f_3(x_6) \left(\sum_{x_6} f_2(x_6) \left(\sum_{x_6} f_2($

If $x = [x_1, x_2]^T$ both gaussian. Then $x_1 | x_2 = a$ is gaussian with mean $\mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(a-\mu_2)$ and variance $\Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$ (note that

$$oldsymbol{z}^{(l)} = \left(oldsymbol{W}^{(l)}\right)^T oldsymbol{x}^{(l-1)} + oldsymbol{b}^{(l)} ext{ with }$$

 $\mathbf{x}^{(0)} = \mathbf{x}_n$ and $\mathbf{x}^{(l)} = \phi(\mathbf{z}^{(l)})$.

- Backward pass: Set $\delta^{(L+1)} = -2(y_n - \mathbf{x}^{(L+1)})\phi'(z^{(L+1)}) \text{ (if squared loss). Then compute}$