## ML Cheat Sheet

#### 1 Math Prerequisites

#### 1.1 Derivatives

$$\begin{array}{l} -\partial(\mathbf{XY}) = (\partial\mathbf{X})\mathbf{Y} + \mathbf{X}(\partial\mathbf{Y}) \\ -\partial\mathbf{Y} & \partial\mathbf{X} & = \frac{\partial\mathbf{u}(\mathbf{x})}{\partial\mathbf{x}} \frac{\partial\mathbf{g}(\mathbf{u})}{\partial\mathbf{u}} \frac{\partial\mathbf{f}(\mathbf{g})}{\partial\mathbf{g}} \end{array}$$

$$-\frac{\partial \mathbf{x}^T \mathbf{a}}{\partial \mathbf{x}} = \frac{\partial \mathbf{a}^T \mathbf{x}}{\partial \mathbf{x}} = \mathbf{a}$$
$$-\frac{\partial \mathbf{a}^T \mathbf{X} \mathbf{b}}{\partial \mathbf{x}} = \mathbf{a} \mathbf{b}^T$$

$$-\frac{\partial \mathbf{a}^T \mathbf{X} \mathbf{b}}{\partial \mathbf{X}} = \mathbf{a} \mathbf{b}^T$$
$$-\frac{\partial \mathbf{a}^T \mathbf{X}^T \mathbf{b}}{\partial \mathbf{X}} = \mathbf{b} \mathbf{a}^T$$

$$-\frac{\partial \mathbf{a} \mathbf{X}}{\partial \mathbf{X}} = \mathbf{b} \mathbf{a}^{T}$$

$$-\frac{\partial \mathbf{X}}{\partial \mathbf{Y}} = \mathbf{J}^{ij}, J^{ij} \text{ is the single entry matrix}$$

- 
$$\frac{\partial \mathbf{A}}{\partial X_{ij}} = \mathbf{J}^{ij}$$
,  $J^{ij}$  is the single entry matri  
-  $\frac{\partial \mathbf{b}^T \mathbf{X}^T \mathbf{X} \mathbf{c}}{\partial \mathbf{v}} = \mathbf{X} \left( \mathbf{b} \mathbf{c}^T + \mathbf{c} \mathbf{b}^T \right)$ 

$$-\frac{\partial \mathbf{B} \cdot \mathbf{A} \mathbf{c}}{\partial \mathbf{X}} = \mathbf{X} \left( \mathbf{b} \mathbf{c}^{T} + \mathbf{c} \mathbf{b}^{T} \right)$$
$$-\frac{\partial \mathbf{x}^{T} \mathbf{B} \mathbf{x}}{\partial \mathbf{x}} = \left( \mathbf{B} + \mathbf{B}^{T} \right) \mathbf{x}$$

$$-\frac{\partial}{\partial \mathbf{x}}(\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s}) = 2\mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s})$$

$$-\frac{\partial}{\partial \mathbf{X}} \|\mathbf{X}\|_{\mathrm{F}}^2 = \frac{\partial}{\partial \mathbf{X}} \operatorname{Tr} \left( \mathbf{X} \mathbf{X}^H \right) = 2\mathbf{X}$$

## 1.2 Linear Algebra

- positive definite (pd) if 
$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

$$- (\mathbf{x} - \mathbf{b})^T (\mathbf{x} - \mathbf{b}) = \|\mathbf{x} - \mathbf{b}\|_2^2$$

- 
$$\|\mathbf{X}\|_F = \|\mathbf{X}^T\|_F$$
  
1.3 Distributions

# Valid distribution p(x)>0, $\forall x$ and $\sum p(x)=1$ Model is identifiable iff $\theta_1=\theta_2\to P_{\theta_1}=P_{\theta_2}$

- Gaussian (Not convex):

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(x|\mu, \Sigma^2) = \frac{\exp(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu))}{\sqrt{(2\pi)^D \det(\Sigma)}}$$

- **Poisson**: P(k events in interval) =  $e^{-\lambda} \frac{\lambda^k}{k!}$ 

- Bernoulli:  $p(y|\mu) = \mu^{y}(1-\mu)^{1-y}$ 

#### 1.4 Convexity

A function f(x) is convex if

- for any 
$$\mathbf{x}_1, \mathbf{x}_2 \in \mathbf{X}$$
 and  $0 \le \lambda \le 1$ , we have :  $f(\lambda \mathbf{x}_1 + (1 - \lambda)\mathbf{x}_2) \le \lambda f(\mathbf{x}_1) + (1 - \lambda)f(\mathbf{x}_2)$ 

it is a sum of convex functions

- composition of convex and linear functions - f(x) = g(h(x)), g,h are convex, g increasing

- the Hessian H is positive semi-definite

#### 1.5 Others

Production of independent variables:

$$\operatorname{Var}(xy) = \mathbb{E}(x^2) \mathbb{E}(y^2) - [\mathbb{E}(x)]^2 [\mathbb{E}(y)]^2$$

Covariance matrix of a data vector x

$$\mathbf{\Sigma} = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \mathbb{E}(\mathbf{x})) (\mathbf{x}_n - \mathbb{E}(\mathbf{x}))^T$$

- Multi-class x

$$p(\mathbf{y}|\mathbf{X}, \beta) = \prod_{n=1}^{N} p(\mathbf{y}_{n}|\mathbf{x}_{n}, \beta)$$
$$= \prod_{n=1}^{K} \prod_{n=1}^{N} [p(\mathbf{y}_{n} = k|\mathbf{x}_{n}, \beta)]^{\tilde{y}_{n}k}$$

#### 2 Cost functions

#### Mean square error (MSE):

$$MSE(\boldsymbol{w}) = \frac{1}{N} \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is strictly convex thus it has only one global minumum value.

MSE is very prone to outliers.

#### Mean Absolute Error (MAE):

$$MAE = \frac{1}{N} \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

- MAE is more robust to outliers than MSE

## Huber less

$$Huber = \left\{ \begin{array}{cc} \frac{1}{2}z^2 & , |z| \leq \delta \\ \delta |z| - \frac{1}{2}\delta^2 & , |z| > \delta \end{array} \right.$$

- Huber loss is convex, differentiable, and also robust to outliers but hard to set  $\delta$ .

#### Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Non-convex, non-diff., but robust to outliers. Hinge loss:

 $[1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$ **Logistic loss**:  $\log(1 - \exp(y_n f(\mathbf{x}_n)))$ 

#### 3 Linear Regression

- Model that assume linear relationship

$$\mathbf{y}_n \approx f(\mathbf{x}_n) := \mathbf{w}_0 + \mathbf{w}_1 \mathbf{x}_{n1} + \dots = \mathbf{w}_0 + \mathbf{x}_n^T \mathbf{w}$$

 $\approx \tilde{\mathbf{x}}_{n}^{T} \mathbf{w}$  ,where  $\tilde{x}$  contains offset comp.

D > N problem: task is underdetermined.

#### 4 Optimization

- Local minimum:

 $L(w^*) \le L(w) \ \forall w : \|w - w^*\| < \epsilon$ 

- Global minimum:  $L(w^*) \leq L(w) \ \forall w$ 

#### 4.1 Grid search

- Compute the cost over a grid of V points. Exponential Complexity  $\mathcal{O}(|V|^D)$ , D is the dimension. Hard to find a good range of values. No guarantee to converge.

### 4.2 GD - Gradient Descent (Batch)

GD uses only first-order information and takes steps in the opposite direction of the gradient

 Given cost function L(w) we want to find w  $\mathbf{w} = \arg\min \mathcal{L}(\mathbf{w})$ 

Take steps in the opposite direction of the

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})$$

With  $\gamma$  too big, method might diverge. With  $\gamma$  too small, convergence is slow.

 Very sensitive to ill-conditioning ⇒ always  $normalize features \Rightarrow allow different$ directions to converge at same speed.

## 4.3 SGD - Stochastic Gradient Descent

SGD update rule (only n-th training example):

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Idea: Cheap but unbiased estimate of grad.

$$\mathbb{E}[\nabla \mathcal{L}_n(\mathbf{w})] = \nabla(\mathbf{w})$$

 $\begin{array}{ll} \text{Robbins-Monroe condition:} \\ - \ \gamma^{(t)} : \sum_{t=1}^{\infty} \gamma^{(t)} = \infty; \sum_{t=1}^{\infty} (\gamma^{(t)})^2 < \infty \\ - \ \text{e.g.} \ \gamma^{(t)} = 1/(t+1)^r, r \in (0.5,1) \end{array}$ 

#### 4.4 Mini-batch SGD

Update direction  $(B \subset [N])$ :

$$\boldsymbol{g}^{(t)} := \frac{1}{|B|} \sum_{n \in B} \boldsymbol{\nabla} \mathcal{L}_n(\mathbf{w}^{(t)})$$

Update rule:  $\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \mathbf{g}^{(t)}$ 

#### 4.5 Gradients for MSE

We define the error vector e:

$$e := y - Xw$$

and MSE as follows:

$$\mathcal{L}(\mathbf{w}) = \frac{1}{2N} \sum_{n=1}^{N} (\mathbf{y}_n - \tilde{\mathbf{x}}_n^T \mathbf{w})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

then the gradient is given by

$$\nabla \mathcal{L}(\mathbf{w}) = -\frac{1}{N} \mathbf{X}^T \mathbf{e}$$

1. necessary: gradient equal zero:  $\frac{d\mathcal{L}(\mathbf{w}^*)}{d\mathcal{L}(\mathbf{w}^*)} = 0$ 

2. sufficient: Hessian matrix is positive

definite:  $\mathbf{H}(\mathbf{w}^*) = \frac{d^2 \mathcal{L}(\mathbf{w}^*)}{d\mathbf{w} d\mathbf{w}^T} = \frac{1}{N} X^T X$ 

4.6 Subgradients (Non-Smooth OPT)

A vector  $\mathbf{g} \in \mathbb{R}^D$  s.t.

$$\mathcal{L}(\mathbf{u}) \ge \mathcal{L}(\mathbf{w}) + \mathbf{g}^T(\mathbf{u} - \mathbf{w}) \quad \forall \mathbf{u} \in \mathbb{R}^D$$

is the subgradient to 
$$\mathcal{L}$$
 at  $\mathbf{w}$ . If  $\mathcal{L}$  is

#### differentiable at $\mathbf{w}$ , we have $\mathbf{g} = \nabla \mathcal{L}(\mathbf{w})$ 4.7 Constrained Optimization

Find solution min  $\mathcal{L}(\mathbf{w})$  s.t.  $\mathbf{w} \in \mathcal{C}$ 

- Add proj. onto 
$$C$$
 after each step:  

$$P_{C}(\mathbf{w}') = \arg\min|\mathbf{v} - \mathbf{w}'|, \mathbf{v} \in C$$

$$\mathbf{w}^{(t+1)} = P_{C}[\mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})]$$

- Use penalty functions
- $\min \mathcal{L}(\mathbf{w}) + I_{\mathcal{C}}, I_{\mathcal{C}} = 0 \text{ if } \mathbf{w} \in \mathcal{C}, \text{ ow } + \infty$
- $-\min \mathcal{L}(\mathbf{w}) + \lambda |\mathbf{A}\mathbf{w} \mathbf{b}|$
- Stopping criteria when  $\mathcal{L}(\mathbf{w})$  close to 0

#### 4.8 Complexities for MSE/MAE per iteration

- $GD = \mathcal{O}(ND)$
- MB-GD= $\mathcal{O}(BD)$ -  $SGD = \mathcal{O}(D)$

#### 5 Least Squares

- Use the first optimality conditions:

$$\nabla L(\mathbf{w}^*) = 0 \Rightarrow \mathbf{X}^T \mathbf{e} = \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}) = 0$$

When  $\mathbf{X}^T \mathbf{X}$  is invertible, we have the closed-form expression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

 thus we can predict values for a new x<sub>m</sub>  $\mathbf{y}_m := \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} \mathbf{w}^* = \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ 

 The Gram matrix X<sup>T</sup>X is pd and is also invertible iff X has full column rank.

Complexity:  $O(ND^2 + D^3) \equiv O(ND^2)$ 

**X** can be rank deficient when D > N or when the comlumns  $\bar{\mathbf{x}}_d$  are nearly collinear.  $\Rightarrow$ matrix is ill-conditioned.

Can still solve using a linear system solver using normal equations:

$$\mathbf{X}^{\top}\mathbf{X}\mathbf{w} = \mathbf{X}^{\top}\mathbf{v}$$

#### 6 Maximum Likelihood (MLE)

Let define the noise ε<sub>n</sub> ~ N(0, σ<sup>2</sup>).

$$\rightarrow \mathbf{y}_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

Another way of expressing this:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} p(\mathbf{y}_n | \mathbf{x}_n, \mathbf{w})$$
$$= \prod_{n=1}^{N} \mathcal{N}(\mathbf{y}_n | \mathbf{x}_n^T \mathbf{w}, \sigma^2)$$

which defines the likelihood of observating y given  $\mathbf{X}$  and  $\mathbf{w}$ 

Define cost with log-likelihood

$$\mathcal{L}_{MLE}(\mathbf{w}) = \log p(\mathbf{y}|\mathbf{X}, \mathbf{w})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (\mathbf{y}_n - \mathbf{x}_n^T \mathbf{w})^2 + cnst$$

 Maximum likelihood estimator (MLE) gives another way to design cost functions

 $\operatorname{argmin} \mathcal{L}_{MSE}(\mathbf{w}) = \operatorname{argmax} \mathcal{L}_{MLE}(\mathbf{w})$ 

 MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.

 $\mathbf{w}_{\mathrm{MLE}} \to \mathbf{w}_{\mathrm{true}}$  for large amount of data

## 7 Ridge Regression and LASSO

- Add regularization term

$$\min_{\mathbf{w}} \mathcal{L}(\mathbf{w}) + \Omega(\mathbf{w})$$

This corresponds to MAP estimate with prior on weights -  $L_2$ -Reg. (Ridge):  $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_2^2$ 

 $-\rightarrow$  small values of  $\mathbf{w}_i$ , not sparse  $-\rightarrow \mathbf{w}^{\star} = (\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \text{ with } \lambda' = 2N\lambda$ 

$$- \rightarrow (\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1}$$
 exists (lifted eigenvalues)

-  $L_1$ -Reg. (Lasso):  $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_1$ 

→ sparsity of weight vector

- Maximum-a-posteriori (MAP) - (i) Posterior prob. ∝ Likelihood × Prior prob

$$\begin{split} p(\mathbf{y}|\mathbf{X}\mathbf{w}) &= \prod^{N} \mathcal{N}(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}, \sigma_{n}^{2}) \\ p(\mathbf{w}) &= \mathcal{N}(\mathbf{w}|0, \sigma_{0}^{2}\mathbf{I}_{D}) \\ \text{then} &\rightarrow \mathbf{w}^{\star} = \operatorname{argmax} p(\mathbf{y}|\mathbf{X}\mathbf{w}) \cdot p(\mathbf{w}) \end{split}$$

$$\mathbf{w}^{\star} = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{\mathbf{w}}^{N} \frac{1}{2\sigma^{2}} (\mathbf{y}_{n} - \mathbf{x}^{T} \mathbf{w})^{2} + \frac{1}{2\sigma^{2}} \|\mathbf{w}\|^{2}$$

#### 8 Bias-Variance decomposition

 The expected test error can be expressed as the sum of two terms

- Squared bias: The average shift of the predictions

- Variance: measure how data points vary around their average.

expected loss =  $(bias)^2 + variance + noise$ Model bias and estimation bias are important

RR increases estimation bias and reduces var Model more complex increases test error

 Small λ → low bias but large variance Large λ → large bias but low variance Simple → large bias but low variance

 Complex → low bias but large variance  $err = \sigma^2 + \mathbb{E}[f_{lse} - \mathbb{E}[f_{lse}]]^2 + [f_{true} - \mathbb{E}[f_{lse}]]$ 

#### 9 Logistic Regression

- Classification relates input variables x to discrete output variable y

Binary classifier: we use y = 0 for  $C_1$  and y = 1 for  $C_2$ .

- Can use least-squares to predict 
$$\hat{y}_*$$
  

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(\mathbf{y}_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$
$$p(\mathbf{y}_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$$

The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} \sigma(\mathbf{x}_{n}^{T}\mathbf{w})^{\mathbf{y}_{n}} (1 - \sigma(\mathbf{x}_{n}^{T}\mathbf{w}))^{1 - \mathbf{y}_{n}}$$

- The log-likelihood (w.r.t. MLE):

$$\mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \mathbf{y}_n \ln \sigma(\mathbf{x}_n^T \mathbf{w}) + (1 - \mathbf{y}_n) \ln (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

$$= -\sum_{n=1}^{N} \mathbf{y}_n \ln \sigma(\mathbf{x}_n^T \mathbf{w}) + (1 - \mathbf{y}_n) \ln (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$
The generalized maximum likelihood cost to expect the content of t

$$= \sum_{n=1}^{N} \ln[1 + \exp(\mathbf{x}_n^T \mathbf{w})] - \mathbf{y}_n \mathbf{x}_n^T \mathbf{w}$$

$$n=1$$

We can use the fact that
$$\frac{d}{dx}\log(1+\exp(x)) = \sigma(x)$$

- Gradient of the log-likelihood

$$\mathbf{g} = \boldsymbol{\nabla} \mathcal{L}(\mathbf{w}) = \sum_{n=1}^{N} \mathbf{x}_n (\sigma(\mathbf{x}_n^T \mathbf{w}) - \mathbf{y}_n)$$

- The negative of the log-likelihood  $-\mathcal{L}_{mle}(w)$ is convex

- Hessian of the log-likelihood - We know that

$$\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t))$$

- Hessian is the derivative of the gradient

$$\mathbf{H}(\mathbf{w}) = -\frac{d\mathbf{g}(\mathbf{w})}{d\mathbf{w}^T} = \sum_{n=1}^{N} \frac{d}{d\mathbf{w}^T} \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w})$$
$$= \sum_{n=1}^{N} \mathbf{x}_n \mathbf{x}_n^T \sigma(\mathbf{x}_n^T \mathbf{w}) (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

where S is a  $N \times N$  diagonal matrix with diagonals

$$S_{nn} = \sigma(\mathbf{x}_n^T \mathbf{w})(1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$
- The negative of the log-likelihood is not

strictly convex. Newton's Method

#### - Uses second-order information and takes steps in the direction that minimizes a

quadratic approximation (Taylor)  

$$\mathcal{L}(\mathbf{w}) = \mathcal{L}(\mathbf{w}^{(k)}) + \nabla \mathcal{L}_{k}^{T} (\mathbf{w} - \mathbf{w}^{(k)})$$

 $+(\mathbf{w}-\mathbf{w}^{(k)})^T\mathbf{H}_k(\mathbf{w}-\mathbf{w}^{(k)})$ and it's minimum is at  $\mathbf{w}^{k+1} = \mathbf{w}^{(k)} - \gamma_k \mathbf{H}_k^{-1} \nabla \mathcal{L}_k$ 

- Complexity:  $O((ND^2 + D^3)I)$ 

$$\underset{\mathbf{w}}{\operatorname{argmin}} - \sum_{n=1}^{N} \ln p(\mathbf{y}_{n} | \mathbf{x}_{n}^{T} \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|^{2}$$

#### 10 Exponential family distribution & Generalized Linear Model

Exponential family distribution

$$\begin{split} p(\mathbf{y}|\boldsymbol{\eta}) &= h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta})) \\ \mathbf{Bernoulli} \ \mathrm{distribution} \ \mathrm{example} \\ &\rightarrow \exp(\log(\frac{\mu}{1-\mu})y + \log(1-\mu))) \end{split}$$

(i) there is a relationship between  $\eta$  and  $\mu$ through the link function

th the link function 
$$\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$$

(ii) Note that  $\mu$  is the mean parameter of y(iii) Relationship between the mean  $\mu$  and  $\eta$ 

is defined using a link function 
$$g$$
  
 $\eta = \mathbf{g}(\mu) \Leftrightarrow \mu = \mathbf{g}^{-1}(\eta)$ 

Gaussian distribution example

exp
$$((\frac{\mu}{\sigma^2}, \frac{-1}{2\sigma^2})(y, y^2)^T - \frac{\mu^2}{2\sigma^2} - \frac{1}{2}\ln(2\pi\sigma^2))$$

link function 
$$\eta = (\eta_1 = \mu/\sigma^2, \eta_2 = -1/(2\sigma^2))^T$$

 $\mu = -\eta_1/(2\eta_2)$ ;  $\sigma^2 = -1/(2\eta_2)$ First and second derivatives of A(n) are

related to the mean and the variance 
$$\frac{dA(\eta)}{d\eta} = \mathbb{E}[\phi(\eta)], \ \, \frac{d^2A(\eta)}{d\eta^2} = \mathrm{Var}[\phi(\eta)]$$

minize is 
$$\min_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \log(p(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}))$$

where  $p(\mathbf{y}_n|\mathbf{x}_n^T\mathbf{w})$  is an exponential family

distribution We obtain the solution

$$\frac{d\mathcal{L}}{d\mathbf{w}} = \mathbf{X}^T [\mathbf{g}^{-1}(\mathbf{X}\mathbf{w}) - \phi(\mathbf{y})]$$

## 11 k-Nearest Neighbor (k-NN)

The k-NN prediction for x is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} \mathbf{y}_n$$

where  $nbh_k(\mathbf{x})$  is the neightborhood of  $\mathbf{x}$ defined by the k closest points  $\mathbf{x}_n$ .

Curse of dimensionality: Generalizing correctly becomes exponentially harder as the dimensionality grows.

Gathering more inputs variables may be bad

#### 12 Support Vector Machine

- Combination of the kernel trick plus a modified loss function (Hinge loss)
- Solution to the dual problem is sparse and non-zero entries will be our support vectors
- Kernelised feature vector where  $\mu_k$  are centroids

$$\phi(\mathbf{x}) = [k(\mathbf{x}, \boldsymbol{\mu}_1), ..., k(\mathbf{x}, \boldsymbol{\mu}_K)]$$

- In practice we'll take a subset of data points to be prototype -> sparse vector machine.
- Assume  $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$\mathcal{L}(\mathbf{w}) = \min_{\mathbf{w}} \sum_{n=1}^{N} [1 - \mathbf{y}_n \tilde{\boldsymbol{\phi}}_n^T \mathbf{w}]_+ + \frac{\lambda}{2} \|\mathbf{w}\|^2$$

- Minimum doesn't change with a rescaling of
- choose the hyperplane so that the distance from it to the nearest data point on each side is maximized

#### Duality:

- Hard to minimize  $g(\mathbf{w})$  so we define  $\mathcal{L}(\mathbf{w}) = \max_{\boldsymbol{\alpha}} G(\mathbf{w}, \boldsymbol{\alpha})$
- we use the property that

$$[\mathbf{v}_n]_+ = \max(0, \mathbf{v}_n) = \max_{\alpha_n \in [0, 1]} \alpha_n \mathbf{v}_n$$
- We can rewrite the problem as

We can rewrite the problem as 
$$\min_{\mathbf{w}} \max_{\alpha} \sum_{n=1}^{N} \alpha_n (1 - \mathbf{y}_n \boldsymbol{\phi}_n^T \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|_2^2$$

- This is differentiable, convex in  $\boldsymbol{w}$  and concave in a
- Minimax theorem:

 $\min_{\mathbf{w}} \max_{\mathbf{\alpha}} G(\mathbf{w}, \mathbf{\alpha}) = \max_{\mathbf{\alpha}} \min_{\mathbf{w}} G(\mathbf{w}, \mathbf{\alpha})$ because G is convex in  $\mathbf{w}$  and concave in

- Derivative w.r.t. w:

- Derivative w.r.t. **w**: 
$$\nabla_{\mathbf{w}} G(\mathbf{w}, \boldsymbol{\alpha}) = -\sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n + \lambda \mathbf{w}$$
- Equating this to 0, we get:

$$\mathbf{w}(\boldsymbol{\alpha}) = \frac{1}{\lambda} \sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n = \frac{1}{\lambda} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

$$\mathbf{Y} := \operatorname{diag}(\cdot)$$

Plugging w\* back in the dual problem

Plugging 
$$\mathbf{w}^*$$
 back in the dual problem 
$$\max_{\boldsymbol{\alpha} \in [0,1]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2\lambda} \boldsymbol{\alpha}^T \mathbf{Y} \mathbf{X} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with  $\mathbf{K} = \mathbf{X}^T \mathbf{X}$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision
- α is the slope of lines that are lower bound to Hinge loss
- Non support vector: Example that lies on the correct side, outside margin  $\alpha_n = 0$
- Essen. support vector: Example that lies on the margin  $\alpha_n \in (0,1)$
- Bound support vector: Example that lies strictly inside the margin or wrong side
- Use Coordinates Descent to find  $\alpha$ . Update one coordinate (argmin) at the time and others constant

#### 13 Kernel Ridge Regression

- The following is true for ridge regression  $\mathbf{w}^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{v} . (1)$ 

$$= \mathbf{X}^{T} (\mathbf{X} \mathbf{X}^{T} + \lambda \mathbf{I}_{D})^{-1} \mathbf{y} = \mathbf{X}^{T} \boldsymbol{\alpha}^{*}, (2)$$

$$= \mathbf{X}^{T} (\mathbf{X} \mathbf{X}^{T} + \lambda \mathbf{I}_{N})^{-1} \mathbf{y} = \mathbf{X}^{T} \boldsymbol{\alpha}^{*}, (2)$$

Complexity of computing w: (1)

$$O(D^2N + D^3), (2) O(DN^2 + N^3)$$

Thus we have

 $\mathbf{w}^* = \mathbf{X}^T \boldsymbol{\alpha}^*$ , with  $\mathbf{w}^* \in \mathbb{R}^D$  and  $\boldsymbol{\alpha}^* \in \mathbb{R}^N$ 

 The representer theorem allows us to write an equivalent optimization problem in terms

$$\alpha = \operatorname*{argmax}_{\boldsymbol{\alpha}} \left( -\frac{1}{2} \boldsymbol{\alpha}^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

- $\mathbf{K} = \mathbf{X}\mathbf{X}^T$  is called the **kernel matrix** o Gram matrix.
- If K is positive definite, then it's called a Mercer Kernel.
- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$  If the kernel is Mercer, then there exists a function  $\phi(\mathbf{x})$  s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

#### Kernel trick:

- compute dot-product in  $\mathbb{R}^m$  while remaining in  $\mathbb{R}^n$
- Replace (x, x') with k(x, x').

#### Common Kernel

- Polynomial Kernel:  $(\gamma \langle \mathbf{x}_i, \mathbf{x}_j \rangle + r)^d$ Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Sigmoid Kernel:  $tanh(\langle \mathbf{x}_i, \mathbf{x}_i \rangle + r)$
- Properties of kernels to ensure the existance of a corresponding  $\phi$ :
- symmetric:  $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$ - positive semi-definite.

$$\mathbf{y} = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

#### 14 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances outside the cluster.

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N z_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

- such that  $z_{nk} \in \{0,1\}$  and  $\sum_{k=1}^{K} z_{nk} = 1$  K-means algorithm (Coordinate Descent): Initialize  $\mu_k$ , then iterate
  - 1. For all n, compute  $\mathbf{z}_n$  given  $\boldsymbol{\mu}$

$$z_{nk} = \begin{cases} 1 & \text{if } k = \operatorname{argmin}_{j} ||\mathbf{x}_{n} - \boldsymbol{\mu}||_{2}^{2} \\ 0 & \text{otherwise} \end{cases}$$

2. For all k, compute  $\mu_k$  given  ${\bf z}$ 

$$u_k = \frac{\sum_{n=1}^{N} z_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} z_{nk}}$$

- $\mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}}$  A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{z}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[ \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{z_{nk}}$$

$$\begin{aligned} -\log p(\mathbf{x}_n|\boldsymbol{\mu},z) &= \sum_{}^{N} \sum_{}^{K} \frac{1}{2} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 \mathbf{z}_{nk} + c' \\ &- \text{K-means as a Matrix Factorization} \\ &\min_{\mathbf{z},\boldsymbol{\mu}} \mathcal{L}(\mathbf{z},\boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{M} \mathbf{Z}^T||_{\text{Frob}}^2 \end{aligned}$$

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{M}\mathbf{Z}^T||_{\text{Fro}}^2$$

Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

#### 15 Gaussian Mixture Models

- Clusters can be elliptical using a full covariance matrix instead of isotropic

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{z}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[ \mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{z_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining  $z_n$  to be a random

$$p(z_n = k) = \pi_k$$
 where  $\pi_k > 0, \forall k, \sum_{k=1}^K \pi_k = 1$ 

Joint distribution of Gaussian mixture model

$$p(\mathbf{X}, \mathbf{z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^{N} p(\mathbf{x}_n | \mathbf{r}_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{z}_n | \boldsymbol{\pi})$$

$$= \prod_{n=1}^{N} \prod_{k=1}^{K} \left[ \left( \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right)^{z_{nk}} \right] \prod_{k=1}^{K} \left[ \pi_{k} \right]^{z_{nk}}$$
-  $z_{n}$  are called  $latent$  unobserved variables

- Unknown parameters are given by
- $\theta = \{\mu, \Sigma, \pi\}$
- We get the marginal likelihood by marginalizing  $z_n$  out from the likelihood

$$\begin{split} p(\mathbf{x}_n|\boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, z_n = k|\boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(z_n = k|\boldsymbol{\theta}) p(\mathbf{x}_n|z_n = k, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{split}$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to  $O(D^2K)$
- To get maximum likelihood estimate of  $\theta$ , we

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

### 16 Expectation Maximization Algorithm

- [ALGORITHM] Start with  $\theta^{(1)}$  and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous  $\boldsymbol{\theta}^{(t)}$  with equality when,

$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$
 2. Maximization step: Update  $\boldsymbol{\theta}$ 

$$\boldsymbol{\theta}^{(t+1)} = \operatorname*{argmax}_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$$

$$\boldsymbol{\mu}_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk}) \mathbf{x}_{n}}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\Sigma_k^{(t+1)} = \frac{\sum_{n=1}^{N} q_{kn}^{(t)} (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)})^T}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\pi_k^{(t+1)} = \frac{1}{N} \sum_{n=1}^{N} q_{kn}^{(t)}$$

If covariance is diagonal → K-means.

#### 17 Matrix factorization

- We have D movies and N users
- **X** is a matrix  $D \times N$  with  $x_{dn}$  the rating of n'th user for d'th movie.
- We project data vectors  $\mathbf{x}_n$  to a smaller dimension  $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
- $\mathbf{Z}$  a  $N \times K$  matrix that gives features for the users
- $\mathbf{W}$  a  $D \times K$  matrix that gives features for the movies

 $x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$ We can add a regularizer and minimize the

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{(d,n) \in \Omega} [x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}]^2$$

$$+ \frac{\lambda_w}{2} \|\mathbf{W}\|_{\operatorname{Frob}}^2 + \frac{\lambda_z}{2} \|\mathbf{Z}\|_{\operatorname{Frob}}^2$$

**SGD**: For one fixed element (d, n) we derive entry (d', k) of **W** (if d = d' oth. 0):

$$\frac{\partial}{\partial w_{d',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] z_{nk}$$
And of **Z** (if  $n = n'$  oth. 0):

$$\frac{\partial}{\partial z_{n',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] w_{nk}$$

$$\mathbf{W}^{t+1} = \mathbf{W}^t - \gamma \nabla_w f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$

$$\mathbf{Z}^{t+1} = \mathbf{W}^t - \gamma \nabla_z f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$
 — We can use coordinate descent algorithm, by

first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_K)^{-1} \mathbf{W}^T \mathbf{X}$$

$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_K)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- Complexity:  $O(DNK^2 + NK^3) \rightarrow O(DNK^2)$ 

#### 18 Singular Value Decomposition

Matrix factorization method

$$\mathbf{X} = \mathbf{U}\mathbf{S}\mathbf{V}^T$$

- $\mathbf{U}$  is a unitary  $D \times D$  matrix
- V is a unitary N × N matrix
- S is a non-negative diagonal matrix of size  $D \times N$  which are called singular values appearing in a descending order.
- Columns of U and V are the left and right singular vectors respectively.
- Assuming D < N we have</li>

$$\mathbf{X} = \sum_{d=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information.

Truncated SVD: Take the matrix  $\mathbf{S}^{(K)}$  with the K first diagonal elements non zero. Then, rank-Kapprox:

$$\mathbf{X} \approx \mathbf{X}_K = \mathbf{U}\mathbf{S}^{(K)}\mathbf{V}^T$$

## 19 Principal Component Analysis

- PCA is a dimensionality reduction method and a method  $\underline{\mathbf{t}}$ o decorrelate the data  $\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$  such that columns of  $\mathbf{W}$  are
- orthogonal. If the data is zero mean

$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$

$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Thus the columns of matrix U are called the principal components and they decorrelate the covariance matrix
- Using SVD, we can compute the matrices in the following way

wing way
$$\mathbf{W} = \mathbf{U}\mathbf{S}_D^{1/2}, \mathbf{Z}^T = \mathbf{S}^{1/2}\mathbf{V}^T$$

- Not invariant under scalings of the feature = arbitrariness, → normalize X

#### 20 Neural Net

 Basic structure: One input layer of size D, L hidden layers of size K, and one output layer. (feedforward network).

$$\begin{aligned} x_j^{(l)} &= \phi\left(\sum_i w_{i,j}^{(l)} x_i^{(l-1)} + b_j^{(l)}\right). \\ &= \text{NN can represent the Rienmann sum with} \\ &\text{only two layers} \Rightarrow \text{It's powerful!} \end{aligned}$$

 $\frac{1}{N} \sum_{n=1}^{N} \left( y_n - f^{(L+1)} \circ \dots \circ f^{(1)}(\boldsymbol{x}_n^{(0)}) \right)^2$  We can use SGD to minimize the cost

#### 20.1 Backpropagation Algorithm

Forward pass: Compute

Cost function:

$$\mathbf{z}^{(l)} = \left(\mathbf{W}^{(l)}\right)^T \mathbf{x}^{(l-1)} + \mathbf{b}^{(l)} \text{ with }$$
  
 $\mathbf{x}^{(0)} = \mathbf{x}_n \text{ and } \mathbf{x}^{(l)} = \phi(\mathbf{z}^{(l)}).$ 

Backward pass: Set

 $\delta^{(L+1)} = -2(y_n - \boldsymbol{x}^{(L+1)})\phi'(z^{(L+1)})$  (if squared loss). Then compute

$$\begin{split} \boldsymbol{\delta}_{j}^{(l)} &= \frac{\partial \mathcal{L}_{n}}{\partial \boldsymbol{z}_{j}^{(l)}} = \sum_{k} \frac{\partial \mathcal{L}_{n}}{\partial \boldsymbol{z}_{k}^{(l+1)}} \frac{\partial \boldsymbol{z}_{k}^{(l+1)}}{\partial \boldsymbol{z}_{j}^{(l)}} \\ &= \sum_{l} \boldsymbol{\delta}_{k}^{(l+1)} \boldsymbol{W}_{j,k}^{(l+1)} \boldsymbol{\phi}'(\boldsymbol{z}_{j}^{(l)}) \end{split}$$

$$\frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\delta^{(l)}}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}} \\ &= \delta_j^{(l)} \cdot 1 = \delta_j^{(l)} \end{split}$$

#### 20.2 Activation Functions

Sigmoid  $\phi(x) = \frac{1}{1+e^{-x}}$  Positive, bounded

 $\phi'(x) \simeq 0$  for large  $|x| \Rightarrow$  Learning slow. Tanh  $\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = \phi(2x) - 1/2$ .

Balanced, bounded. Learning slow too. **ReLU**  $(x)_{+} = \max 0, x$  Positive, unbounded

Derivate = 1 if x > 0, 0 if x < 0Leaky ReLU  $f(x) = \max \alpha x, x$  Remove 0 derivative.

## $f(x) = \max \mathbf{x}^T \mathbf{w}_1 + b_1, ..., \mathbf{x}^T \mathbf{w}_k + b_k$ (Generalization of ReLU)

20.3 Convolutional NN Sparse connections and weights sharing: reduce complexity. (e.g. pixels in pictures only depend

## 20.4 Reg, Data Augmentation and

- Dropout – Regularization term:  $\frac{1}{2} \sum_{l=1}^{L+1} \mu^{(l)} ||W^{(l)}||_F^2$
- Weight decay is  $\Theta[t](1-\eta\mu)$  in:
- $\Theta[t+1] = \Theta[t] + \eta(\nabla \mathcal{L} + \mu \Theta[t])$ Data Augm.: e.g. shift or rotation of pics
- Dropout: avoid overfit. Drop nodes randomly. (Then average multiple drop-NN)

on neighbours)

- 21 Bayes Net - Graph example: p(x, y, z) = p(y|x)p(z|x)p(x)
- $: (y \leftarrow x \rightarrow z)$ D-Separation X and Y are D-separated by Z if every path from  $x \in X$  to  $y \in Y$  is
- blocked by Z. (→ independent) Blocked Path contains a variable that
- is in Z and is head-to-tail or tail-to-tail - the node is head-to-head and neither the
- node nor any of its descendants are in Z. Markov Blanket (which blocks node A from
  - parents of A
  - children of A
  - parents of children of A

the rest of the net) contains: