ML Cheat Sheet

1 Math Prerequisites

1.1 Derivatives

$$\begin{array}{l} - \ \partial(\mathbf{X}\mathbf{Y}) = (\partial\mathbf{X})\mathbf{Y} + \mathbf{X}(\partial\mathbf{Y}) \\ - \ \frac{\partial\mathbf{f}(\mathbf{g}(\mathbf{u}(\mathbf{x})))}{\partial\mathbf{x}} = \frac{\partial\mathbf{u}(\mathbf{x})}{\partial\mathbf{x}} \frac{\partial\mathbf{g}(\mathbf{u})}{\partial\mathbf{u}} \frac{\partial\mathbf{f}(\mathbf{g})}{\partial\mathbf{g}} \\ - \ \frac{\partial\mathbf{x}^T\mathbf{a}}{\partial\mathbf{x}} = \frac{\partial\mathbf{a}^T\mathbf{x}}{\partial\mathbf{x}} = \mathbf{a} \\ \frac{\partial\mathbf{x}^T\mathbf{y}}{\partial\mathbf{x}} = \frac{\partial\mathbf{x}^T\mathbf{y}}{\partial\mathbf{x}} \end{array}$$

$$- \frac{\partial \mathbf{a}^T \mathbf{X} \mathbf{b}}{\partial \mathbf{X}} = \mathbf{a} \mathbf{b}^T$$

$$- \frac{\partial \mathbf{a}^T \mathbf{X}^T \mathbf{b}}{\partial \mathbf{Y}} = \mathbf{b} \mathbf{a}^T$$

$$-\frac{\partial \mathbf{X}}{\partial X_{ij}} = \mathbf{J}^{ij}, J^{ij}$$
 is the single entry matrix

$$- \frac{\partial \mathbf{b}^{T} \mathbf{X}^{T} \mathbf{X} \mathbf{c}}{\partial \mathbf{X}} = \mathbf{X} \left(\mathbf{b} \mathbf{c}^{T} + \mathbf{c} \mathbf{b}^{T} \right)$$

$$- \frac{\partial \mathbf{x}^T \mathbf{B} \mathbf{x}}{\partial \mathbf{x}} = (\mathbf{B} + \mathbf{B}^T) \mathbf{x}$$

$$-\frac{\partial}{\partial \mathbf{x}}(\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s}) = 2\mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s})$$

$$- \frac{\partial}{\partial \mathbf{X}} \|\mathbf{X}\|_{\mathrm{F}}^2 = \frac{\partial}{\partial \mathbf{X}} \operatorname{Tr} \left(\mathbf{X} \mathbf{X}^H \right) = 2\mathbf{X}$$

1.2 Linear Algebra

- positive definite (pd) if
$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

$$- (\mathbf{x} - \mathbf{b})^T (\mathbf{x} - \mathbf{b}) = \|\mathbf{x} - \mathbf{b}\|_2^2$$

-
$$\|\mathbf{X}\|_F = \|\mathbf{X}^T\|_F$$

1.3 Distributions

Valid distribution p(x) > 0, $\forall x$ and $\sum p(x) = 1$ Model is identifiable iff $\theta_1 = \theta_2 \rightarrow P_{\theta_1} = P_{\theta_2}$

Gaussian (Not convex):

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$
$$\mathcal{N}(x|\mu, \Sigma^2) = \frac{\exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right)}{\sqrt{(2\pi)^D \det(\Sigma)}}$$

- **Poisson**: P(k events in interval) = $e^{-\lambda} \frac{\lambda^k}{k!}$

- Bernoulli: $p(y|\mu) = \mu^{y}(1-\mu)^{1-y}$

1.4 Convexity

A function f(x) is convex if

- for any $\mathbf{x}_1, \mathbf{x}_2 \in \mathbf{X}$ and $0 < \lambda < 1$, we have : $f(\lambda \mathbf{x}_1 + (1 - \lambda)\mathbf{x}_2) \le \lambda f(\mathbf{x}_1) + (1 - \lambda)f(\mathbf{x}_2)$ it is a sum of convex functions

- the Hessian H is positive semi-definite

1.5 Others

- Production of independent variables:

 $Var(xy) = \mathbb{E}(x^2) \mathbb{E}(y^2) - [\mathbb{E}(x)]^2 [\mathbb{E}(y)]^2$

Covariance matrix of a data vector x

$$\mathbf{\Sigma} = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \mathbb{E}(\mathbf{x})) (\mathbf{x}_n - \mathbb{E}(\mathbf{x}))^T$$

- Multi-class x

$$p(\mathbf{y}|\mathbf{X}, \beta) = \prod_{i=1}^{K} p(\mathbf{y}_{i}|\mathbf{x}_{i}, \beta)$$
$$= \prod_{i=1}^{K} \prod_{j=1}^{K} [p(\mathbf{y}_{i} = k|\mathbf{x}_{i}, \beta)]^{\tilde{y}_{i}k}$$

2 Cost functions

Cost functions are used to learn parameters that explain the data well.

 It is essential to make sure that a global minimum exist \rightarrow lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{w}) = \frac{1}{N} \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

MSE is convex thus it has only one global minumum value.

MSE is not good when outliers are present.

Mean Absolute Error (MAE):

$$MAE = \frac{1}{N} \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

Huber loss

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \le \delta \\ \delta|z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

- Huber loss is convex, differentiable, and also robust to outliers but hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Non-convex, non-diff., but robust to outliers. Hinge loss:

 $[1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$ Logistic loss: $\log(1 - \exp(y_n f(\mathbf{x}_n)))$

3 Regression

- Model that assume linear relationship

$$\mathbf{y}_n \approx f(\mathbf{x}_n) := \mathbf{w}_0 + \mathbf{w}_1 \mathbf{x}_{n1} + \dots = \mathbf{w}_0 + \mathbf{x}_n^T \mathbf{w}$$

 $\approx \tilde{\mathbf{x}}_n^T \mathbf{w}$, where \tilde{x} contains offset comp.

- Prediction: predict the ouput for a new
- Interpretation: understand the effect of inputs on output.

4 Optimization

4.1 Grid search

 Compute the cost over a grid of M points to find the minimum. Exponential Complexity. Hard to find a good range of values

4.2 GD - Gradient Descent (Batch)

 GD uses only first-order information and takes steps in the opposite direction of the

 Given cost function L(w) we want to find w $\mathbf{w} = \arg\min \mathcal{L}(\mathbf{w})$

- Take steps in the opposite direction of the

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})$$

 With γ too big, method might diverge. With γ too small, convergence is slow.

4.3 SGD - Stochastic Gradient Descent

In ML, most cost functions are formulated as a sum over the training examples:

$$\mathcal{L}(\mathbf{w}) = \frac{1}{N} \sum_{n=1}^{N} \mathcal{L}_n(\mathbf{w})$$

⇒ SGD update rule (only n-th training exam.):

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Idea: Cheap but unbiased estimate of grad.

$$\mathbb{E}[\nabla \mathcal{L}_n(\mathbf{w})] = \nabla(\mathbf{w})$$

4.4 Mini-batch SGD

Update direction $(B \subset [N])$:

$$\boldsymbol{g}^{(t)} := \frac{1}{|B|} \sum_{n \in B} \boldsymbol{\nabla} \mathcal{L}_n(\mathbf{w}^{(t)})$$

Update rule : $\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \mathbf{g}^{(t)}$

4.5 Gradients for MSE

We define the error vector e:

$$e := y - Xy$$

$$\mathcal{L}(\mathbf{w}) = \frac{1}{2N} \sum_{n=1}^{N} (\mathbf{y}_n - \tilde{\mathbf{x}}_n^T \mathbf{w})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

then the gradient is given by

$$\nabla \mathcal{L}(\mathbf{w}) = -\frac{1}{N} \mathbf{X}^T \mathbf{e}$$

Optimality conditions:

1. necessary: gradient equal zero: $\frac{d\mathcal{L}(\mathbf{w}^*)}{d\mathcal{L}(\mathbf{w}^*)} = 0$

2. sufficient: Hessian matrix is positive

definite: $\mathbf{H}(\mathbf{w}^*) = \frac{d^2 \mathcal{L}(\mathbf{w}^*)}{dt}$ definite: $\mathbf{h}(\mathbf{w}') = \frac{1}{d\mathbf{w}d\mathbf{w}^T}$ Very sensitive to illconditioning \Rightarrow always

normalize features.

Complexity: O(NDI) with I the number of

4.6 Subgradients (Non-Smooth OPT)

A vector $\mathbf{g} \in \mathbb{R}^D$ s.t.

$$\mathcal{L}(\mathbf{u}) \geq \mathcal{L}(\mathbf{w}) + \mathbf{g}^T(\mathbf{u} - \mathbf{w}) \quad \forall \mathbf{u} \in \mathbb{R}^D$$

is the subgradient to \mathcal{L} at \mathbf{w} . If \mathcal{L} is differentiable at \mathbf{w} , we have $\mathbf{g} = \nabla \mathcal{L}(\mathbf{w})$

4.7 Constrained Optimization

Find solution min $\mathcal{L}(\mathbf{w})$ s.t. $\mathbf{w} \in \mathcal{C}$

 Add proj. onto C after each step: $P_{\mathcal{C}}(\mathbf{w}') = \arg\min |\mathbf{v} - \mathbf{w}'|, \mathbf{v} \in \mathcal{C}$ $\mathbf{w}^{(t+1)} = P_{\mathcal{L}}[\mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})]$

- Use penalty functions

 $-\min \mathcal{L}(\mathbf{w}) + I_{\mathcal{C}}, I_{\mathcal{C}} = 0 \text{ if } \mathbf{w} \in \mathcal{C}, \text{ ow } + \infty$ $-\min \mathcal{L}(\mathbf{w}) + \lambda |\mathbf{A}\mathbf{w} - \mathbf{b}|$

Stopping criteria when $\mathcal{L}(\mathbf{w})$ close to 0

5 Least Squares

- Use the first optimality conditions:

$$\nabla L(\mathbf{w}^*) = 0 \Rightarrow \mathbf{X}^T \mathbf{e} = \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}) = 0$$

 When X^TX is invertible, we have the closed-form expression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

- thus we can predict values for a new \mathbf{x}_m $\mathbf{y}_m := \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} \mathbf{w}^* = \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$

The Gram matrix $\mathbf{X}^T \mathbf{X}$ is pd and is also

invertible iff X has full column rank. Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$ **X** can be rank deficient when D > N or when

the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. \Rightarrow matrix is ill-conditioned.

6 Maximum Likelihood (MLE)

Let define our mistakes ϵ_n ∼ N(0, σ²).

$$\rightarrow \mathbf{y}_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

Another way of expressing this:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} p(\mathbf{y}_n | \mathbf{x}_n, \mathbf{w})$$

which defines the likelihood of observating y

given \mathbf{X} and \mathbf{w} Define cost with log-likelihood

$$\mathcal{L}_{MLE}(\mathbf{w}) = \log p(\mathbf{y}|\mathbf{X}, \mathbf{w})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (\mathbf{y}_n - \mathbf{x}_n^T \mathbf{w})^2 + cnst$$

Maximum likelihood estimator (MLE) gives another way to design cost functions

 $\operatorname{argmin} \mathcal{L}_{MSE}(\mathbf{w}) = \operatorname{argmax} \mathcal{L}_{MLE}(\mathbf{w})$ MLE can also be interpreted as finding the

model under which the observed data is most likely to have been generated from.

$\mathbf{w}_{\mathrm{MLE}} ightarrow \mathbf{w}_{\mathrm{true}}$ for large amount of data

7 Ridge Regression (RR)

- Linear models usually overfit. We can penalize them with a regularization term $\min \mathcal{L}(\mathbf{w}) + \Omega(\mathbf{w})$

- L_2 -Reg. (Ridge): $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_2^2$ → small values of w_i, not sparse

 $-\rightarrow \mathbf{w}^{\star} = (\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \text{ with } \lambda' = 2N\lambda$ - \rightarrow No ill cond.. $(\mathbf{X}^T\mathbf{X} + \lambda'\mathbf{I})^{-1}$ exists

- L_1 -Reg. (Lasso): $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_1$ $-\rightarrow$ large values of \mathbf{w}_i , sparse - Maximum-a-posteriori (MAP) (i) Posterior prob. ∝ Likelihood × Prior prob

$$p(\mathbf{y}|\mathbf{X}\mathbf{w}) = \prod_{n=1}^{N} \mathcal{N}(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}, \sigma_{n}^{2})$$
$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|0, \sigma_{0}^{2}\mathbf{I}_{D})$$
then $\rightarrow \mathbf{w}^{*} = \operatorname{argmax} p(\mathbf{y}|\mathbf{X}\mathbf{w}) \cdot p(\mathbf{w})$

$$\mathbf{w}^{\star} = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{n=1}^{N} \frac{1}{2\sigma_{n}^{2}} (\mathbf{y}_{n} - \mathbf{x}^{T} \mathbf{w})^{2} + \frac{1}{2\sigma_{0}^{2}} ||\mathbf{w}||^{2}$$

8 Bias-Variance decomposition

 The expected test error can be expressed as the sum of two terms

- Squared bias: The average shift of the predictions

Variance: measure how data points vary around their average.

 $expected loss = (bias)^2 + variance + noise$ Model bias and estimation bias are important

RR increases estimation bias and reduces var

 Model more complex increases test error Small λ → low bias but large variance

 Large λ → large bias but low variance Simple → large bias but low variance

 Complex → low bias but large variance $err = \sigma^2 + \mathbb{E}[f_{lse} - \mathbb{E}[f_{lse}]]^2 + [f_{true} - \mathbb{E}[f_{lse}]]^2$

9 Logistic Regression

 Classification relates input variables x to discrete output variable y

Binary classifier: we use y = 0 for C_1 and y = 1 for C_2 .

- Can use least-squares to predict \hat{y}_*

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5\\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$
$$p(\mathbf{y}_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$

$$p(\mathbf{y}_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$$

- The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} \sigma(\mathbf{x}_{n}^{T}\mathbf{w})^{\mathbf{y}_{n}} (1 - \sigma(\mathbf{x}_{n}^{T}\mathbf{w}))^{1-\mathbf{y}_{n}}$$
- The log-likelihood (w.r.t. MLE):

The log-likelihood (w.r.t. MLE):
$$\mu = -\eta_1/(2\eta_2); \ \sigma = -1/(2\eta_2)$$

$$- \text{First and second derivatives of } A(\eta) \text{ a}$$

$$\mathcal{L}(\mathbf{w}) = -\sum_{n} \mathbf{y}_n \ln \sigma(\mathbf{x}_n^T \mathbf{w}) + (1 - \mathbf{y}_n) \ln(1 - \sigma(\mathbf{x}_n^T \mathbf{w})) \text{ related to the mean and the variance}$$

$$\begin{split} &= \sum_{n=1}^{N} \ln[1 + \exp(\mathbf{x}_n^T \mathbf{w})] - \mathbf{y}_n \mathbf{x}_n^T \mathbf{w} \\ &- \text{ We can use the fact that} \\ &\qquad \frac{d}{dx} \log(1 + \exp(x)) = \sigma(x) \\ &- \text{ Gradient of the log-likelihood} \end{split}$$

$$\frac{d}{d}\log(1+\exp(x)) = \sigma(x)$$

radient of the log-likelihood
$$\mathbf{g} = \nabla \mathcal{L}(\mathbf{w}) = \sum_{n=1}^{N} \mathbf{x}_n (\sigma(\mathbf{x}_n^T \mathbf{w}) - \mathbf{y}_n)$$

 $= \mathbf{X}^T [\sigma(\mathbf{X}w) - \mathbf{y}]$ The negative of the log-likelihood $-\mathcal{L}_{mle}(\boldsymbol{w})$

is convex Hessian of the log-likelihood

– We know that
$$\frac{d\sigma(t)}{dt} = \sigma(t)(1-\sigma(t))$$

- Hessian is the derivative of the gradient

$$\mathbf{H}(\mathbf{w}) = -\frac{d\mathbf{g}(\mathbf{w})}{d\mathbf{w}^T} = \sum_{n=1}^{N} \frac{d}{d\mathbf{w}^T} \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w})$$
$$= \sum_{n=1}^{N} \mathbf{x}_n \mathbf{x}_n^T \sigma(\mathbf{x}_n^T \mathbf{w}) (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

where **S** is a $N \times N$ diagonal matrix with

 $S_{nn} = \sigma(\mathbf{x}_n^T \mathbf{w})(1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$

The negative of the log-likelihood is not strictly convex.

- Newton's Method

- Uses second-order information and takes steps in the direction that minimizes a quadratic approximation (Taylor)

$$\mathcal{L}(\mathbf{w}) = \mathcal{L}(\mathbf{w}^{(k)}) + \nabla \mathcal{L}_k^T (\mathbf{w} - \mathbf{w}^{(k)})$$

$$\begin{aligned} &+(\mathbf{w}-\mathbf{w}^{(k)})^T\mathbf{H}_k(\mathbf{w}-\mathbf{w}^{(k)}) \\ \text{and it's minimum is at} & \\ &\mathbf{w}^{k+1}=\mathbf{w}^{(k)}-\gamma_k\mathbf{H}_k^{-1}\nabla\mathcal{L}_k \end{aligned}$$

- Complexity: $O((ND^2 + D^3)I)$

- Regularized Logistic Regression

$$\underset{\mathbf{w}}{\operatorname{argmin}} - \sum_{n=1}^{N} \ln p(\mathbf{y}_{n} | \mathbf{x}_{n}^{T} \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|^{2}$$

10 Exponential family distribution & Generalized Linear Model

Exponential family distribution

$$p(\mathbf{y}|\boldsymbol{\eta}) = h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

Bernoulli distribution example

$$\rightarrow \exp(\log(\frac{\mu}{1-\mu})y + \log(1-\mu)))$$
(i) there is a relationship between η and μ

through the link function

 $\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$ (ii) Note that μ is the mean parameter of y

(iii) Relationship between the mean
$$\mu$$
 and η is defined using a link function g

 $\eta = \mathbf{g}(\boldsymbol{\mu}) \Leftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta})$ - Gaussian distribution example

Gaussian distribution example
$$\exp((\frac{\mu}{\sigma^2}, \frac{-1}{2\sigma^2})(y, y^2)^T - \frac{\mu^2}{2\sigma^2} - \frac{1}{2}\ln(2\pi\sigma^2))$$

nk function
$$\eta = (\eta_1 = \mu/\sigma^2, \eta_2 = -1/(2\sigma^2))^T$$

 $\mu = -\eta_1/(2\eta_2)$; $\sigma^2 = -1/(2\eta_2)$ First and second derivatives of A(η) are

elated to the mean and the variance
$$\frac{dA(\eta)}{d\eta} = \mathbb{E}[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = \operatorname{Var}[\phi(\eta)]$$

– $A(\eta)$ is convex – The generalized maximum likelihood cost to

mimize is
$$\min_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \log(p(\mathbf{y}_n | \mathbf{x}_n^T \mathbf{w}))$$

where $p(\mathbf{y}_n|\mathbf{x}_n^T\mathbf{w})$ is an exponential family

- We obtain the solution
$$\frac{d\mathcal{L}}{d\mathbf{w}} = \mathbf{X}^T [\mathbf{g}^{-1} (\mathbf{X} \mathbf{w}) - \phi(\mathbf{y})]$$

11 k-Nearest Neighbor (k-NN)

The k-NN prediction for x is

of prediction for
$$\mathbf{x}$$
 is
$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} \mathbf{y}_n$$

where $nbh_k(\mathbf{x})$ is the neightborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n .

Curse of dimensionality: Generalizing correctly becomes exponentially harder as the dimensionality grows.

Gathering more inputs variables may be bad

12 Support Vector Machine

- Combination of the kernel trick plus a modified loss function (Hinge loss)
- Solution to the dual problem is sparse and non-zero entries will be our support
- Kernelised feature vector where μ_k are centroids

$$\phi(\mathbf{x}) = [k(\mathbf{x}, \boldsymbol{\mu}_1), ..., k(\mathbf{x}, \boldsymbol{\mu}_K)]$$

- In practice we'll take a subset of data points to be prototype \rightarrow sparse vector machine. Assume $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$\mathcal{L}(\mathbf{w}) = \min_{\mathbf{w}} \sum_{n=1}^{N} [1 - \mathbf{y}_n \tilde{\boldsymbol{\phi}}_n^T \mathbf{w}]_+ + \frac{\lambda}{2} \|\mathbf{w}\|^2$$

- choose the hyperplane so that the distance from it to the nearest data point on each side is maximized

Duality:

- Hard to minimize $g(\mathbf{w})$ so we define $\mathcal{L}(\mathbf{w}) = \max G(\mathbf{w}, \boldsymbol{\alpha})$
- we use the property that $[\mathbf{v}_n]_+ = \max(0, \mathbf{v}_n) = \max_{\alpha_n \in [0, 1]} \alpha_n \mathbf{v}_n$
- We can rewrite the problem as

$$\min_{\mathbf{w}} \max_{\alpha} \sum_{n=1}^{N} \alpha_{n} (1 - \mathbf{y}_{n} \boldsymbol{\phi}_{n}^{T} \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|_{2}^{2}$$

- This is differentiable, convex in \boldsymbol{w} and concave in α
- Minimax theorem:

 $\min_{\mathbf{w}} \max_{\mathbf{\alpha}} G(\mathbf{w}, \mathbf{\alpha}) = \max_{\mathbf{\alpha}} \min_{\mathbf{w}} G(\mathbf{w}, \mathbf{\alpha})$ because G is convex in \mathbf{w} and concave in

- Derivative w.r.t. w:

$$\begin{aligned} \boldsymbol{\nabla}_{\mathbf{w}} G(\mathbf{w}, \boldsymbol{\alpha}) &= -\sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n + \lambda \mathbf{w} \\ &- \text{ Equating this to 0, we get:} \end{aligned}$$

$$\mathbf{w}(\boldsymbol{\alpha}) = \frac{1}{\lambda} \sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n = \frac{1}{\lambda} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

Plugging w* back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,1]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2\lambda} \boldsymbol{\alpha}^T \mathbf{Y} \mathbf{X} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{X}^T \mathbf{X}$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision boundary
- α is the slope of lines that are lower bound to Hinge loss
- Non support vector: Example that lies on the correct side, outside margin $\alpha_n = 0$
- Essen. support vector: Example that lies on the margin $\alpha_n \in (0,1)$
- Bound support vector: Example that lies strictly inside the margin or wrong side
- Use Coordinates Descent to find α . Update one coordinate (argmin) at the time and others constant.

13 Kernel Ridge Regression

- The following is true for ridge regression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y} , (1)$$
$$= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \boldsymbol{\alpha}^* , (2)$$

- Complexity of computing w: (1) $O(D^2N + D^3)$, (2) $O(DN^2 + N^3)$ - Thus we have

 $\mathbf{w}^* = \mathbf{X}^T \boldsymbol{\alpha}^*, \text{ with } \mathbf{w}^* \in \mathbb{R}^D \text{ and } \boldsymbol{\alpha}^* \in \mathbb{R}^N$

- The representer theorem allows us to write an equivalent optimization problem in terms

of
$$\alpha$$
.

$$\alpha = \underset{\alpha}{\operatorname{argmax}} \left(-\frac{1}{2} \alpha^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \alpha + \alpha^T \mathbf{y} \right)$$

- $\mathbf{K} = \mathbf{X}\mathbf{X}^T$ is called the **kernel matrix** or Gram matrix.
- If K is positive definite, then it's called a Mercer Kernel.
- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$ If the kernel is Mercer, then there exists a function $\phi(\mathbf{x})$ s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

- Kernel trick:
- compute dot-product in R^m while remaining in \mathbb{R}^n Replace (x, x') with k(x, x').
- Common Kernel
- $\begin{array}{ll} & \text{Polynomial Kernel: } (\gamma \langle \mathbf{x}_i, \mathbf{x}_j \rangle + r)^d \\ & \text{Radial Basis function kernel (RBF)} \end{array}$

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Sigmoid Kernel: $tanh(\langle \mathbf{x}_i, \mathbf{x}_i \rangle + r)$
- Properties of kernels to ensure the existance of a corresponding ϕ :
- symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$
- positive semi-definite.

$$\mathbf{y} = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

14 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances outside the cluster.

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N z_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

- such that $z_{nk} \in \{0,1\}$ and $\sum_{k=1}^K z_{nk} = 1$ K-means algorithm (Coordinate Descent): Initialize μ_k , then iterate
 - 1. For all n, compute \mathbf{z}_n given $\boldsymbol{\mu}$

$$z_{nk} = \left\{ \begin{array}{ll} 1 & \text{ if } k = \mathop{\mathrm{argmin}}_j ||\mathbf{x}_n - \boldsymbol{\mu}||_2^2 \\ 0 & \text{ otherwise} \end{array} \right.$$

2. For all k, compute μ_k given \mathbf{z}

$$\mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}}$$

- A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{z}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \mathbf{I}) \right]^{z_{nk}}$$

$$-\log p(\mathbf{x}_{n}|\mu, z) = \sum_{k=1}^{N} \sum_{n=1}^{K} \frac{1}{2} \|\mathbf{x}_{n} - \mu_{k}\|^{2} \mathbf{z}_{nk} + c'$$

K-means as a Matrix Factorization

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{MZ}^T||_{\text{Frob}}^2$$

Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

15 Gaussian Mixture Models

- Clusters can be elliptical using a full covariance matrix instead of isotropic

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{z}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{z_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining z_n to be a random

$$p(z_n = k) = \pi_k$$
 where $\pi_k > 0, \forall k, \sum_{k=1}^K \pi_k = 1$

- Joint distribution of Gaussian mixture model

$$p(\mathbf{X}, \mathbf{z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^{N} p(\mathbf{x}_{n} | \mathbf{r}_{n}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{z}_{n} | \boldsymbol{\pi})$$

$$=\prod_{k=1}^{N}\prod_{k=1}^{K}[\left(\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k},\boldsymbol{\Sigma}_{k})\right)^{z_{nk}}]\prod_{k=1}^{K}[\boldsymbol{\pi}_{k}]^{z_{nk}}$$

- z_n are called latent unobserved variables Unknown parameters are given by
- $\theta = \{\mu, \Sigma, \pi\}$
- We get the marginal likelihood by marginalizing z_n out from the likelihood

$$\begin{aligned} p(\mathbf{x}_n|\boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, z_n = k|\boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(z_n = k|\boldsymbol{\theta}) p(\mathbf{x}_n|z_n = k, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{aligned}$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced
- To get maximum likelihood estimate of θ , we

$$\max_{\pmb{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \pmb{\mu}_k, \pmb{\Sigma}_k)$$

16 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\boldsymbol{\theta}^{(1)}$ and iterate
 - 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\boldsymbol{\theta}^{(t)}$ with equality when,

$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$
 2. Maximization step: Update $\boldsymbol{\theta}$

$$\boldsymbol{\theta}^{(t+1)} = \operatorname*{argmax}_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$$

$$\mu_k^{(t+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk}) \mathbf{x}_n}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\Sigma_k^{(t+1)} = \frac{\sum_{n=1}^N q_{kn}^{(t)} (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)})^T}{\sum_{n=1}^N q_{kn}^{(t)}}$$

$$\pi_k^{(t+1)} = \frac{1}{N} \sum_{n=1}^{N} q_{kn}^{(t)}$$

If covariance is diagonal → K-means.

17 Matrix factorization

- We have D movies and N users
- **X** is a matrix $D \times N$ with x_{dn} the rating of n'th user for d'th movie.
- We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
- \mathbf{Z} a $N \times K$ matrix that gives features for the users
- \mathbf{W} a $D \times K$ matrix that gives features for the movies

 $x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$ - We can add a regularizer and minimize the

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{(\mathbf{J}, \mathbf{r}) \in \Omega} [x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}]^2$$

$$+ \frac{\lambda_w}{2} \|\mathbf{W}\|_{\operatorname{Frob}}^2 + \frac{\lambda_z}{2} \|\mathbf{Z}\|_{\operatorname{Frob}}^2$$

SGD: For one fixed element (d, n) we derive entry (d', k) of **W** (if d = d' oth. 0):

$$\frac{\partial}{\partial w_{d',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] z_{nk}$$
And of **Z** (if $n = n'$ oth. 0):

$$\frac{\partial}{\partial z_{n',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] w_{nk}$$

$$\mathbf{W}^{t+1} = \mathbf{W}^t - \gamma \nabla_w f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$
$$\mathbf{Z}^{t+1} = \mathbf{W}^t - \gamma \nabla_z f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$

- We can use coordinate descent algorithm, by first minimizing w.r.t. Z given W and then minimizing W given Z. This is called

Alternating least-squares (ALS):
$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_x \mathbf{I}_K)^{-1} \mathbf{W}^T \mathbf{X}$$

$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_K)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- Complexity: $O(DNK^2 + NK^3) \rightarrow O(DNK^2)$

18 Singular Value Decomposition

Matrix factorization method

$$X = USV^T$$

- \mathbf{U} is a unitary $D \times D$ matrix
- V is a unitary N × N matrix
- S is a non-negative diagonal matrix of size $D \times N$ which are called **singular values** appearing in a descending order.
- Columns of U and V are the left and right singular vectors respectively.
- Assuming D < N we have

$$\mathbf{X} = \sum_{d=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information.

Truncated SVD: Take the matrix $\mathbf{S}^{(K)}$ with the K first diagonal elements non zero. Then, rank-Kapprox:

$$\mathbf{X} \approx \mathbf{X}_K = \mathbf{U}\mathbf{S}^{(K)}\mathbf{V}^T$$

19 Principal Component Analysis

- PCA is a dimensionality reduction method and a method to decorrelate the data $\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$ such that columns of \mathbf{W} are orthogonal.
- If the data is zero mean

$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$
$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Thus the columns of matrix U are called the principal components and they decorrelate the covariance matrix
- Using SVD, we can compute the matrices in the following way

$$\mathbf{W} = \mathbf{U}\mathbf{S}_D^{1/2}, \mathbf{Z}^T = \mathbf{S}^{1/2}\mathbf{V}^T$$

- Not invariant under scalings of the feature = arbitrariness, → normalize X

20 Neural Net

 Basic structure: One input layer of size D, L hidden layers of size K, and one output layer. (feedforward network).

$$x_{j}^{(l)} = \phi\left(\sum_{i} w_{i,j}^{(l)} x_{i}^{(l-1)} + b_{j}^{(l)}\right).$$
- NN can represent the Rienmann sum with

only two layers ⇒ It's powerful!

Cost function:

$$\mathcal{L} = \frac{1}{N} \sum_{n=1}^{N} \left(y_n - f^{(L+1)} \circ \dots \circ f^{(1)}(\boldsymbol{x}_n^{(0)}) \right)^2$$

We can use SGD to minimize the cost

20.1 Backpropagation Algorithm

Forward pass: Compute

$$\boldsymbol{z}^{(l)} = \left(\boldsymbol{W}^{(l)}\right)^T \boldsymbol{x}^{(l-1)} + \boldsymbol{b}^{(l)}$$
 with

 $\mathbf{x}^{(0)} = \mathbf{x}_n \text{ and } \mathbf{x}^{(l)} = \phi(\mathbf{z}^{(l)}).$ Backward pass: Set

- Backwara pass: Set
$$\delta^{(L+1)} = -2(y_n - \boldsymbol{x}^{(L+1)})\phi'(z^{(L+1)}) \text{ (if squared loss). Then compute}$$

$$\begin{split} \delta_j^{(l)} &= \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l+1)}} \frac{\partial z_k^{(l+1)}}{\partial z_j^{(l)}} \\ &= \sum_k \delta_k^{(l+1)} \mathbf{W}_{j,k}^{(l+1)} \phi'(z_j^{(l)}) \end{split}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}} \\ &= \delta_i^{(l)} \mathbf{x}_i^{(l-1)} \end{split}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}} \\ &= \delta_j^{(l)} \cdot 1 = \delta_j^{(l)} \end{split}$$

20.2 Activation Functions

Sigmoid $\phi(x) = \frac{1}{1+e^{-x}}$ Positive, bounded

 $\phi'(x) \simeq 0$ for large $|x| \Rightarrow$ Learning slow. Tanh $\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = \phi(2x) - 1/2$.

Balanced, bounded. Learning slow too. **ReLU** $(x)_{+} = \max 0, x$ Positive, unbounded

Derivate = 1 if x > 0, 0 if x < 0Leaky ReLU $f(x) = \max \alpha x, x$ Remove 0

derivative.

Maxout
$$f(x) = max \boldsymbol{x}^T \boldsymbol{w}_1 + b_1, ..., \boldsymbol{x}^T \boldsymbol{w}_k + b_k$$
(Generalization of ReLU)

20.3 Convolutional NN

Sparse connections and weights sharing: reduce complexity. (e.g. pixels in pictures only depend on neighbours)

20.4 Reg, Data Augmentation and Dropout

- Regularization term: $\frac{1}{2} \sum_{l=1}^{L+1} \mu^{(l)} ||W^{(l)}||_F^2$
- Weight decay is Θ[t](1 ημ) in:
- $\Theta[t+1] = \Theta[t] + \eta(\nabla \mathcal{L} + \mu \Theta[t])$ Data Augm.: e.g. shift or rotation of pics - Dropout: avoid overfit. Drop nodes
- randomly. (Then average multiple drop-NN)
- 21 Bayes Net - Graph example: p(x, y, z) = p(y|x)p(z|x)p(x) $: (y \leftarrow x \rightarrow z)$
- D-Separation X and Y are D-separated by Z if every path from $x \in X$ to $y \in Y$ is blocked by Z. (→ independent)
- Blocked Path contains a variable that
- is in Z and is head-to-tail or tail-to-tail
- the node is head-to-head and neither the node nor any of its descendants are in Z. Markov Blanket (which blocks node A from
- the rest of the net) contains:
- parents of A
- children of A
- parents of children of A