

ML Cheat Sheet

1 Math Prerequisites

1.1 Derivatives

- $\partial(\mathbf{XY}) = (\partial\mathbf{X})\mathbf{Y} + \mathbf{X}(\partial\mathbf{Y})$
- $\frac{\partial f(\mathbf{g}(\mathbf{u}(\mathbf{x})))}{\partial \mathbf{x}} = \frac{\partial \mathbf{u}(\mathbf{x})}{\partial \mathbf{x}} \frac{\partial \mathbf{g}(\mathbf{u})}{\partial \mathbf{u}} \frac{\partial f(\mathbf{g})}{\partial \mathbf{g}}$
- $\frac{\partial \mathbf{x}^T \mathbf{a}}{\partial \mathbf{x}} = \frac{\partial \mathbf{a}^T \mathbf{x}}{\partial \mathbf{x}} = \mathbf{a}$
- $\frac{\partial \mathbf{a}^T \mathbf{X} \mathbf{b}}{\partial \mathbf{X}} = \mathbf{a} \mathbf{b}^T$
- $\frac{\partial \mathbf{a}^T \mathbf{X}^T \mathbf{b}}{\partial \mathbf{X}} = \mathbf{b} \mathbf{a}^T$
- $\frac{\partial \mathbf{X}}{\partial X_{ij}} = \mathbf{J}^{ij}, \mathbf{J}^{ij}$ is the single entry matrix
- $\frac{\partial \mathbf{b}^T \mathbf{X}^T \mathbf{X} \mathbf{c}}{\partial \mathbf{X}} = \mathbf{X}(\mathbf{b} \mathbf{c}^T + \mathbf{c} \mathbf{b}^T)$
- $\frac{\partial \mathbf{x}^T \mathbf{B} \mathbf{x}}{\partial \mathbf{x}} = (\mathbf{B} + \mathbf{B}^T) \mathbf{x}$
- $\frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A} \mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A} \mathbf{s}) = 2 \mathbf{W} (\mathbf{x} - \mathbf{A} \mathbf{s})$
- $\frac{\partial}{\partial \mathbf{X}} \|\mathbf{X}\|_F^2 = \frac{\partial}{\partial \mathbf{X}} \text{Tr}(\mathbf{X} \mathbf{X}^H) = 2 \mathbf{X}$

1.2 Linear Algebra

- positive definite** (pd) if $\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$
- $(\mathbf{x} - \mathbf{b})^T (\mathbf{x} - \mathbf{b}) = \|\mathbf{x} - \mathbf{b}\|_2^2$
- $\|\mathbf{X}\|_F = \|\mathbf{X}^T\|_F$
- $\det \mathbf{X} = x_{11}x_{22} - x_{12}x_{21}$, if \mathbf{X} is 2×2

1.3 Distributions

Valid distribution $p(x) > 0, \forall x$ and $\sum p(x) = 1$
Model is identifiable iff $\theta_1 = \theta_2 \rightarrow P_{\theta_1} = P_{\theta_2}$

- Gaussian** (Not convex):

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

$$\mathcal{N}(x|\mu, \Sigma^2) = \frac{\exp\left(-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu)\right)}{\sqrt{(2\pi)^D \det(\Sigma)}}$$

- Poisson**: P(k events in interval) = $e^{-\lambda} \frac{\lambda^k}{k!}$
- Bernoulli**: $p(y|\mu) = \mu^y (1-\mu)^{1-y}$

1.4 Convexity

A function $f(x)$ is convex if

- for any $\mathbf{x}_1, \mathbf{x}_2 \in \mathbf{X}$ and $0 \leq \lambda \leq 1$, we have : $f(\lambda \mathbf{x}_1 + (1-\lambda)\mathbf{x}_2) \leq \lambda f(\mathbf{x}_1) + (1-\lambda)f(\mathbf{x}_2)$
- it is a sum of convex functions
- composition of convex and linear functions
- $f(x) = g(h(x))$, g,h are convex, g increasing
- the Hessian \mathbf{H} is positive semi-definite

2 Cost functions

Mean square error (MSE):

$$MSE(\mathbf{w}) = \frac{1}{N} \sum_{n=1}^N (y_n - f(\mathbf{x}_n))^2$$

- MSE is **strictly convex** thus it has only one global minimum value.
- MSE is very prone to outliers.

Mean Absolute Error (MAE):

$$MAE = \frac{1}{N} \sum_{n=1}^N |y_n - f(\mathbf{x}_n)|$$

- MAE is more robust to outliers than MSE.

Huber loss

$$Huber = \begin{cases} \frac{1}{2} z^2 & , |z| \leq \delta \\ \delta |z| - \frac{1}{2} \delta^2 & , |z| > \delta \end{cases}$$

- Huber loss is convex, differentiable, and also robust to outliers but hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 & , |z| < \delta \\ 0 & , |z| \geq \delta \end{cases}$$

Non-convex, non-diff., but robust to outliers.

Hinge loss:

$$[1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$$

Logistic loss: $\log(1 - \exp(y_n f(\mathbf{x}_n)))$

3 Optimization

- Local minimum:
- $L(\mathbf{w}^*) \leq L(\mathbf{w}) \forall \mathbf{w} : \|\mathbf{w} - \mathbf{w}^*\| < \epsilon$
- Global minimum: $L(\mathbf{w}^*) \leq L(\mathbf{w}) \forall \mathbf{w}$

3.1 Grid search

- Compute the cost over a grid of V points. Exponential complexity $\mathcal{O}(|V|^D)$. Hard to find a good value range. No guarantee to converge.

3.2 GD - Gradient Descent (Batch)

- GD uses only first-order information
- Given cost function $\mathcal{L}(\mathbf{w})$ we want to find $\mathbf{w} = \arg \min_{\mathbf{w}} \mathcal{L}(\mathbf{w})$

- Take steps in the opposite direction of gradient

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})$$

- With γ too big, method might diverge. With γ too small, convergence is slow.
- Very sensitive to ill-conditioning \Rightarrow always normalize features \Rightarrow allow different directions to converge at same speed.

3.3 SGD - Stochastic Gradient Descent

SGD update rule (only n-th training example):

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Idea: Cheap but unbiased estimate of grad.

$$\mathbb{E}[\nabla \mathcal{L}_n(\mathbf{w})] = \nabla \mathcal{L}(\mathbf{w})$$

Robbins-Monroe condition:

- $\gamma^{(t)} : \sum_{t=1}^{\infty} \gamma^{(t)} = \infty; \sum_{t=1}^{\infty} (\gamma^{(t)})^2 < \infty$
- e.g. $\gamma^{(t)} = 1/(t+1)^r, r \in (0.5, 1)$

3.4 Mini-batch SGD

Update direction ($B \subseteq [N]$):

$$\mathbf{g}^{(t)} := \frac{1}{|B|} \sum_{n \in B} \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Update rule : $\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \mathbf{g}^{(t)}$

3.5 Gradients for MSE

- Define error $\mathbf{e} := \mathbf{y} - \mathbf{X} \mathbf{w}$
- and MSE as follows:

$$\mathcal{L}(\mathbf{w}) = \frac{1}{2N} \sum_{n=1}^N (\mathbf{y}_n - \tilde{\mathbf{x}}_n^T \mathbf{w})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

- Optimality conditions:

- necessary*: $\frac{d\mathcal{L}(\mathbf{w}^*)}{d\mathbf{w}} = -\frac{1}{N} \mathbf{X}^T \mathbf{e} = 0$
- sufficient*: Hessian matrix is positive definite: $\mathbf{H}(\mathbf{w}^*) = \frac{d^2 \mathcal{L}(\mathbf{w}^*)}{d\mathbf{w} d\mathbf{w}^T} = \frac{1}{N} \mathbf{X}^T \mathbf{X}$

3.6 Subgradients (Non-Smooth OPT)

A vector $\mathbf{g} \in \mathbb{R}^D$ s.t.

$$\mathcal{L}(\mathbf{u}) \geq \mathcal{L}(\mathbf{w}) + \mathbf{g}^T (\mathbf{u} - \mathbf{w}) \quad \forall \mathbf{u} \in \mathbb{R}^D$$

is the subgradient to \mathcal{L} at \mathbf{w} . If \mathcal{L} is differentiable at \mathbf{w} , we have $\mathbf{g} = \nabla \mathcal{L}(\mathbf{w})$

3.7 Constrained Optimization

Find solution $\min \mathcal{L}(\mathbf{w})$ s.t. $\mathbf{w} \in \mathcal{C}$

- Add proj. onto \mathcal{C} after each step:
 $P_{\mathcal{C}}(\mathbf{w}') = \arg \min_{\mathbf{v}} \|\mathbf{v} - \mathbf{w}'\|, \mathbf{v} \in \mathcal{C}$
- $\mathbf{w}^{(t+1)} = P_{\mathcal{C}}[\mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})]$
- Use penalty functions

- $\min \mathcal{L}(\mathbf{w}) + I_{\mathcal{C}}, I_{\mathcal{C}} = 0$ if $\mathbf{w} \in \mathcal{C}$, ∞ otherwise
- $\min \mathcal{L}(\mathbf{w}) + \lambda \|\mathbf{A} \mathbf{w} - \mathbf{b}\|$
- Stopping criteria when $\mathcal{L}(\mathbf{w})$ close to 0

3.8 Iteration complexities for MSE/MAE

- GD= $\mathcal{O}(ND)$
- MB-GD= $\mathcal{O}(BD)$
- SGD= $\mathcal{O}(D)$

4 Least Squares

- Use the first optimality conditions:
 $\nabla \mathcal{L}(\mathbf{w}^*) = 0 \Rightarrow \mathbf{X}^T \mathbf{e} = \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}) = 0$
- When $\mathbf{X}^T \mathbf{X}$ is invertible, we have the closed-form expression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

- thus we can predict values for a new \mathbf{x}_m

$$\mathbf{y}_m := \mathbf{x}_m^T \mathbf{w}^* = \mathbf{x}_m^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

- The **Gram matrix** $\mathbf{X}^T \mathbf{X}$ is pd and is also invertible iff \mathbf{X} has full column rank.
- Complexity*: $\mathcal{O}(ND^2 + D^3) \equiv \mathcal{O}(ND^2)$
- \mathbf{X} can be rank deficient when $D > N$ or when the columns $\tilde{\mathbf{x}}_d$ are nearly collinear. \Rightarrow matrix is ill-conditioned.
- Can still solve using a linear system solver using normal equations:

$$\mathbf{X}^T \mathbf{X} \mathbf{w} = \mathbf{X}^T \mathbf{y}$$

5 Maximum Likelihood (MLE)

- Let define the noise $\epsilon_n \sim \mathcal{N}(0, \sigma^2)$.
 $\rightarrow \mathbf{y}_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$
- Another way of expressing this:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^N p(\mathbf{y}_n|\mathbf{x}_n, \mathbf{w}) = \prod_{n=1}^N \mathcal{N}(\mathbf{y}_n|\mathbf{x}_n^T \mathbf{w}, \sigma^2)$$

which defines the likelihood of observing \mathbf{y} given \mathbf{X} and \mathbf{w}

- Define cost with log-likelihood

$$\begin{aligned} \mathcal{L}_{MLE}(\mathbf{w}) &= \log p(\mathbf{y}|\mathbf{X}, \mathbf{w}) \\ &= -\frac{1}{2\sigma^2} \sum_{n=1}^N (\mathbf{y}_n - \mathbf{x}_n^T \mathbf{w})^2 + c n s t \end{aligned}$$

- Maximum likelihood estimator (MLE) gives another way to design cost functions
 $\arg \min_{\mathbf{w}} \mathcal{L}_{MSE}(\mathbf{w}) = \arg \max_{\mathbf{w}} \mathcal{L}_{MLE}(\mathbf{w})$
- MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.
- $\mathbf{w}_{MLE} \rightarrow \mathbf{w}_{true}$ for large amount of data

6 Ridge Regression and LASSO

- Add **regularization term**
 $\min_{\mathbf{w}} \mathcal{L}(\mathbf{w}) + \Omega(\mathbf{w})$
- L_2 -Reg. (Ridge): $\Omega(\mathbf{w}) = \lambda \|\mathbf{w}\|_2^2$
 \rightarrow small values of \mathbf{w}_i , not sparse
- $\rightarrow \mathbf{w}^* = (\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$ with $\lambda' = 2N\lambda$
- $(\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1}$ exists (lifted eigenvalues)
- L_1 -Reg. (Lasso): $\Omega(\mathbf{w}) = \lambda \|\mathbf{w}\|_1$
 \rightarrow sparsity of weight vector
- \rightarrow implicit model selection
- Maximum-a-posteriori (MAP)**
- (i) Posterior prob. \propto Likelihood \times Prior prob

$$p(\mathbf{y}|\mathbf{X} \mathbf{w}) = \prod_{n=1}^N \mathcal{N}(\mathbf{y}_n|\mathbf{x}_n^T \mathbf{w}, \sigma_n^2)$$

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \sigma_0^2 \mathbf{I}_D)$$

$$\text{then } \rightarrow \mathbf{w}^* = \arg \max_{\mathbf{w}} p(\mathbf{y}|\mathbf{X} \mathbf{w}) \cdot p(\mathbf{w})$$

$$\mathbf{w}^* = \arg \min_{\mathbf{w}} \sum_{n=1}^N \frac{1}{2\sigma_n^2} (\mathbf{y}_n - \mathbf{x}_n^T \mathbf{w})^2 + \frac{1}{2\sigma_0^2} \|\mathbf{w}\|^2$$

7 Model Selection

- Generalisation error: $L_D(f) = \mathbb{E}[l(y, f(x))]$, but D normally unknown.
- Instead approximate by $L_{S_{test}}(f_{train}) = \frac{1}{|S_{test}|} \sum_{S_{test}} l(y_n, f_{S_{train}}(x_n))$
- In expectation this equates the true error.
- Worst case, if comparing K models:

- $\mathbb{P}[\max_k |L_D(f_k) - L_{test}(f_k)| \geq \sqrt{\frac{(bia)^2 \ln(2K/\delta)}{2|S_{test}|}}] \leq \delta$
- Error decreases as $\mathcal{O}(1/\sqrt{|S_{test}|})$
- Error only goes up by $\sqrt{\ln(K)}$ for testing K models.
- use cross-validation for an efficient, unbiased estimate of generalisation error and variance.

8 Bias-Variance decomposition

- Simple** (e.g. large λ) \rightarrow large bias, but low variance
- Complex** (e.g. small λ) \rightarrow low bias, but large variance
- The expected squared loss between true model and learned model is a sum of three non-negative terms:
 $\mathbb{E}_S[(f(x) + \epsilon - f_S(x))^2] = \text{Var}[\epsilon] + \text{bias} + \text{variance}$
- Bias** = $(f(x) - \mathbb{E}_{S'}[f_{S'}(x)])^2$: Difference between actual value and expected prediction.
- Variance** = $\mathbb{E}_S[(\mathbb{E}_{S'}[f_{S'}(x)] - f_S(x))^2]$: variance of predictions between training sets.
- All terms are lower bounds for the error.
- Cannot do better than $\text{Var}[\epsilon]$.

9 Logistic Regression

- Binary classifier**: use $y \in \{0, 1\}$.
- Can use least-squares to predict \hat{y}_s

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_s < 0.5 \\ \mathbf{C}_2 & \hat{y}_s \geq 0.5 \end{cases}$$

- Logistic function**

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(\mathbf{y}_n = 1|\mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$

$$p(\mathbf{y}_n = 0|\mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$$

- The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^N \sigma(\mathbf{x}_n^T \mathbf{w})^{y_n} (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))^{1-y_n}$$

- The negative log-likelihood (w.r.t. MLE):

$$\mathcal{L}(\mathbf{w}) = - \sum_{n=1}^N \mathbf{y}_n \ln \sigma(\mathbf{x}_n^T \mathbf{w}) + (1 - \mathbf{y}_n) \ln(1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

$$= \sum_{n=1}^N \ln[1 + \exp(\mathbf{x}_n^T \mathbf{w})] - \mathbf{y}_n \mathbf{x}_n^T \mathbf{w}$$

- We can use the fact that

$$\frac{d}{dz} \ln(1 + \exp(z)) = \sigma(z)$$

- Gradient of the log-likelihood

$$\begin{aligned} \mathbf{g} &= \nabla \mathcal{L}(\mathbf{w}) = \sum_{n=1}^N \mathbf{x}_n (\sigma(\mathbf{x}_n^T \mathbf{w}) - \mathbf{y}_n) \\ &= \mathbf{X}^T [\sigma(\mathbf{X} \mathbf{w}) - \mathbf{y}] \end{aligned}$$

- The neg. log-likelihood $-\mathcal{L}_{mle}(\mathbf{w})$ is convex
- Hessian** of the neg. log-likelihood

$$\begin{aligned} &\text{We know that} \\ &\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t)) \end{aligned}$$

- Hessian is the derivative of the gradient

$$\mathbf{H}(\mathbf{w}) = \frac{d\mathbf{g}(\mathbf{w})}{d\mathbf{w}^T} = \sum_{n=1}^N \frac{d}{d\mathbf{w}^T} \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w})$$

$$\begin{aligned} &= \sum_{n=1}^N \mathbf{x}_n \mathbf{x}_n^T \sigma(\mathbf{x}_n^T \mathbf{w})(1 - \sigma(\mathbf{x}_n^T \mathbf{w})) \\ &= \tilde{\mathbf{X}}^T \mathbf{S} \tilde{\mathbf{X}} \end{aligned}$$

where \mathbf{S} is a $N \times N$ diagonal with

$$S_{nn} = \sigma(\mathbf{x}_n^T \mathbf{w})(1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

- The neg. log-likelihood is not strictly convex. ????

Newton's Method

- Uses second-order information and takes steps in the direction that minimizes a quadratic approximation (Taylor)

$$\mathcal{L}(\mathbf{w}) = \mathcal{L}(\mathbf{w}^{(k)}) + \nabla \mathcal{L}_k^T (\mathbf{w} - \mathbf{w}^{(k)}) + (\mathbf{w} - \mathbf{w}^{(k)})^T \mathbf{H}_k (\mathbf{w} - \mathbf{w}^{(k)})$$

and it's minimum is at
 $\mathbf{w}^{k+1} = \mathbf{w}^{(k)} - \gamma_k \mathbf{H}_k^{-1} \nabla \mathcal{L}_k$

- Complexity: $\mathcal{O}((ND^2 + D^3)I)$
- Regularized Logistic Regression**

- If data is linearly separable, there is no best weight vector \Rightarrow optimisation does not stop.
- \rightarrow use penalty term.

$$\arg \min_{\mathbf{w}} - \sum_{n=1}^N \ln p(\mathbf{y}_n|\mathbf{x}_n^T \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|^2$$

10 Exponential family distribution & Generalized Linear Model

- Exponential family distribution

$$p(\mathbf{y}|\boldsymbol{\eta}) = h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

- For proper normalisation ($\int p = 1$):

$$A(\boldsymbol{\eta}) = \ln \left[\int_{\mathbf{y}} h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y})) d\mathbf{y} \right]$$

- Bernoulli** distribution example

$$\rightarrow \exp(\log(\frac{\mu}{1-\mu})y + \log(1-\mu))$$

- (i) **link function** \mathbf{g} relates $\boldsymbol{\eta}$ and μ

$$\boldsymbol{\eta} = \mathbf{g}(\boldsymbol{\mu}) \Leftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta})$$

$$\bold$$

- dimension.
- b) In high dimensions, points are far from each other \Rightarrow choice of NN becomes essentially random.
Need radius
$$r = \sqrt{d \left(1 - \frac{1}{N}\right)}$$
to have at least one data point in r^d rectangle with $p \geq \frac{1}{2}$.
- NN performance:**
 - Bayes classifier: $f_*(x) = 1\{\mathbb{P}[y = 1|x] > \frac{1}{2}\}$
 - $\mathbb{E}_S[L(f_S)] \leq 2L(f_*) + 4c\sqrt{dN} \frac{1}{1+d}$

12 Support Vector Machine

- Assume $y_n \in \{-1, 1\}$ and optimise
$$\mathcal{L}(\mathbf{w}) = \min_{\mathbf{w}} \sum_{n=1}^N [1 - y_n x_n^T \mathbf{w}]_+ + \frac{\lambda}{2} \|\mathbf{w}\|^2$$
- Can be optimised using subgradient descent.
- Case: Linear separability:** We get a separating hyperplane, no point in the margin and $w, s.t$ margin is maximised ($2/\|\mathbf{w}\|$).
- This is called hard-margin compared to soft-margin formulation.
- Duality:**
 - Hard to minimize $g(\mathbf{w})$ so we define
$$\mathcal{L}(\mathbf{w}) = \max G(\mathbf{w}, \alpha)$$
 - we use the property that
$$[\mathbf{v}_n]_+ = \max(0, \mathbf{v}_n) = \max_{\alpha_n \in [0, 1]} \alpha_n \mathbf{v}_n$$
- We can rewrite the problem as
$$\min_{\mathbf{w}} \max_{\alpha} \sum_{n=1}^N \alpha_n (1 - y_n \phi_n^T \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|^2$$
- This is differentiable, convex in \mathbf{w} and concave in α
- Minimax theorem:**
$$\min_{\alpha} \max_{\mathbf{w}} G(\mathbf{w}, \alpha) = \max_{\mathbf{w}} \min_{\alpha} G(\mathbf{w}, \alpha)$$
because G is convex in \mathbf{w} and concave in α .
- Derivative w.r.t. \mathbf{w} :
$$\nabla_{\mathbf{w}} G(\mathbf{w}, \alpha) = - \sum_{n=1}^N \alpha_n y_n \mathbf{x}_n + \lambda \mathbf{w}$$

- Equating this to 0, we get:
$$\mathbf{w}(\alpha) = \frac{1}{\lambda} \sum_{n=1}^N \alpha_n y_n \mathbf{x}_n = \frac{1}{\lambda} \mathbf{X}^T \mathbf{Y} \alpha$$
- $\mathbf{Y} := \text{diag}(\mathbf{y})$
- Plugging \mathbf{w}^* back in the dual problem
$$\max_{\alpha \in [0, 1]^N} \alpha^T \mathbf{1} - \frac{1}{2\lambda} \alpha^T \mathbf{Y} \mathbf{X} \mathbf{X}^T \mathbf{Y} \alpha$$
- Data only enters as $\mathbf{K} = \mathbf{X}^T \mathbf{X}$.
- Non support vector:** Example that lies on the correct side, outside margin $\alpha_n = 0$
- Essen. support vector:** Example that lies on the margin $\alpha_n \in (0, 1)$
- Bound support vector:** Example that lies strictly inside the margin or wrong side $\alpha_n = 1$
- Use Coordinates ascent to find α . Update one coordinate (argmin) at the time and others constant.

13 Kernel Ridge Regression

- The following is true for ridge regression
$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y}, (1)$$
$$= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \alpha^*, (2)$$
- Complexity of computing \mathbf{w} : (1) $O(D^2 N + D^3)$, (2) $O(DN^2 + N^3)$
- Thus we have
$$\mathbf{w}^* = \mathbf{X}^T \alpha^*, \quad \text{with } \mathbf{w}^* \in \mathbb{R}^D \text{ and } \alpha^* \in \mathbb{R}^N$$

- Following representer theorem write:
$$\alpha = \arg \max_{\alpha} \left(-\frac{1}{2} \alpha^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \alpha + \alpha^T \mathbf{y} \right)$$
- $\mathbf{K} = \mathbf{X} \mathbf{X}^T$ is called the **kernel matrix** or **Gram matrix**.
- If \mathbf{K} is positive definite and symmetric, then it's called a **Mercer Kernel**.
- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$
- If the kernel is Mercer, then there exists a function $\phi(\mathbf{x})$ s.t.

- $$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$
- Kernel trick:**
 - compute dot-product in \mathbb{R}^m while remaining in \mathbb{R}^n
 - Replace $\langle \mathbf{x}, \mathbf{x}' \rangle$ with $k(\mathbf{x}, \mathbf{x}')$.
- Common Kernel**
 - $x \in \mathbb{R}, k(\mathbf{x}, \mathbf{x}') = (xx')^2 \Rightarrow \phi(x) = x^2$
 - Radial Basis function kernel (RBF)
$$k(\mathbf{x}, \mathbf{x}') = \exp\left(-\frac{1}{2} (\mathbf{x} - \mathbf{x}')^T (\mathbf{x} - \mathbf{x}')\right)$$
- Thus we get
$$\mathbf{y} = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$
- Creating new kernels:**
 - $\kappa(x, x') = a\kappa_1(x, x') + b\kappa_2(x, x')$
 - $\kappa(x, x') = \kappa_1(x, x')\kappa_2(x, x')$
 - $\kappa(x, x') = \kappa_1(f(x), f(x'))$

14 K-means

- $$\min_{\mathbf{z}, \mu} \mathcal{L}(\mathbf{z}, \mu) = \sum_{k=1}^K \sum_{n=1}^N z_{nk} \|\mathbf{x}_n - \mu_k\|_2^2$$
- such that $z_{nk} \in \{0, 1\}$ and $\sum_{k=1}^K z_{nk} = 1$
- K-means algorithm (Coordinate Descent): Initialize μ_k , then iterate
 - For all n , compute \mathbf{z}_n given μ
$$z_{nk} = \begin{cases} 1 & \text{if } k = \arg \min_j \|\mathbf{x}_n - \mu_j\|_2^2 \\ 0 & \text{otherwise} \end{cases}$$
 - For all k , compute μ_k given \mathbf{z}
$$\mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}}$$
- A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model
$$p(\mathbf{z}, \mu) = \prod_{n=1}^N \prod_{k=1}^K [\mathcal{N}(\mathbf{x}_n | \mu_k, \mathbf{I})]^{z_{nk}}$$
- $\log p(\mathbf{x}_n | \mu, \mathbf{z}) = \sum_{k=1}^N \frac{1}{2} \|\mathbf{x}_n - \mu_k\|^2 z_{nk} + c'$
- K-means as a Matrix Factorization
$$\min_{\mathbf{z}, \mu} \mathcal{L}(\mathbf{z}, \mu) = \|\mathbf{X} - \mathbf{M} \mathbf{Z}^T\|_{\text{Frob}}^2$$
- Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

15 Gaussian Mixture Models

- Clusters can be elliptical using a full instead of isotropic covariance matrix.
$$p(\mathbf{X} | \mu, \Sigma, \mathbf{z}) = \prod_{n=1}^N \prod_{k=1}^K [\mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)]^{z_{nk}}$$
- Soft-clustering:** Points can belong to several cluster by defining z_n to be a random variable.
$$p(z_n = k) = \pi_k \text{ where } \pi_k > 0, \forall k, \sum_{k=1}^K \pi_k = 1$$

- Joint distribution of Gaussian mixture model
$$p(\mathbf{X}, \mathbf{z} | \mu, \Sigma, \pi) = \prod_{n=1}^N p(\mathbf{x}_n | \mathbf{r}_n, \mu, \Sigma) p(\mathbf{z}_n | \pi)$$
$$= \prod_{n=1}^N \prod_{k=1}^K [(\mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k))]^{z_{nk}} \prod_{k=1}^K [\pi_k]^{z_{nk}}$$
- z_n are called *latent* unobserved variables
- Unknown parameters are $\theta = \{\mu, \Sigma, \pi\}$
- We get the **marginal likelihood** by marginalizing \mathbf{z}_n out from the likelihood
$$p(\mathbf{x}_n | \theta) = \sum_{k=1}^K p(\mathbf{x}_n, z_n = k | \theta)$$
$$= \sum_{k=1}^K p(z_n = k | \theta) p(\mathbf{x}_n | z_n = k, \theta)$$
$$= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)$$
- Without a latent variable model, number of parameters grow at rate $O(N)$
- After marginalization, the growth is reduced to $O(D^2 K)$
- To get maximum likelihood estimate of θ , we maximize
$$\max_{\theta} \sum_{n=1}^N \log \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)$$

16 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate
 - Expectation step:* Compute a lower bound to the cost such that it is tight at the previous $\theta^{(t)}$ with equality when,
$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}$$
 - Maximization step:* Update θ
$$\theta^{(t+1)} = \arg \max_{\theta} \mathcal{L}(\theta, \theta^{(t)})$$

$$\mu_k^{(t+1)} = \frac{\sum_{n=1}^N \gamma^{(i)}(r_{nk}) \mathbf{x}_n}{\sum_{n=1}^N q_{kn}^{(t)}}$$

$$\Sigma_k^{(t+1)} = \frac{\sum_{n=1}^N q_{kn}^{(t)} (\mathbf{x}_n - \mu_k^{(t+1)}) (\mathbf{x}_n - \mu_k^{(t+1)})^T}{\sum_{n=1}^N q_{kn}^{(t)}}$$
- $\pi_k^{(t+1)} = \frac{1}{N} \sum_{n=1}^N q_{kn}^{(t)}$
- If covariance is diagonal \rightarrow K-means.
- $q_{kn}^{(t)} = p(z_n = k | \mathbf{x}_n, \theta^{(t)})$ posterior of z_n

17 Matrix factorization

- Find $\mathbf{X} \approx \mathbf{W} \mathbf{Z}^T$
 - \mathbf{X} is $D \times N$ (e.g movies \times user)
 - \mathbf{Z} is $N \times K, \mathbf{W}$ is $D \times K$ matrix
- $$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{(d,n) \in \Omega} [x_{dn} - (\mathbf{W} \mathbf{Z}^T)_{dn}]^2 + \frac{\lambda_w}{2} \|\mathbf{W}\|_{\text{Frob}}^2 + \frac{\lambda_z}{2} \|\mathbf{Z}\|_{\text{Frob}}^2$$
- SGD:** For one fixed element (d, n) we derive entry (d', k) of \mathbf{W} (if $d = d'$ oth. 0):
$$\frac{\partial}{\partial w_{d',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W} \mathbf{Z}^T)_{dn}] z_{nk}$$
And of \mathbf{Z} (if $n = n'$ oth. 0):
$$\frac{\partial}{\partial z_{n',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W} \mathbf{Z}^T)_{dn}] w_{nk}$$

- updates:
$$\mathbf{W}^{t+1} = \mathbf{W}^t - \gamma \nabla_w f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$
$$\mathbf{Z}^{t+1} = \mathbf{W}^t - \gamma \nabla_z f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$
- We can use coordinate descent algorithm, by first minimizing w.r.t. \mathbf{Z} given \mathbf{W} and then minimizing \mathbf{W} given \mathbf{Z} . This is called **Alternating least-squares (ALS)**:
$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_K)^{-1} \mathbf{W}^T \mathbf{X}$$
$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_D)^{-1} \mathbf{Z}^T \mathbf{X}^T$$
- $\mathbb{O}(DNK + DK^2)$ and $\mathbb{O}(DNK + NK^2)$

18 Text Representation

- word2vec:** map every word to a vector $w_i \in \mathbb{R}^K$, K large, that captures its semantics.
- Topic model:** Documents consist of collections of topics
 - topic = probability distribution over words
 - use clustering to pick out representative topics
- Word representations by matrix factorisation**
 - typically use log counts from co-occurrence matrix
 - $\min_{w,z} L(w, z) = \frac{1}{2} \sum_{(d,n) \in \Omega} f_{dn} [x_{dn} - (WZ^T)_{dn}]^2$
 - f_{dn} : importance of entry
 - $f_{dn} = 1$ is okay, but better
 $f_{dn} = \min[1, (n_{dn}/N_{max}^{\alpha}), \alpha \in [0, 1], n_{dn}$ are counts.
 - this weighting is called GloVe (word2vec variant) and creates spatial analogies
 - training with SGD or ALS
 - Skip-Gram (original word2vec) uses binary classification to distinguish real from fake word pairs. Implicitly based on matrix factorisation.
 - FastText: supervised sentence classification.
 - Sentence as x_n bag-of-words representation, f is a linear classifier loss, $y_n \in \{0, 1\}$
 - $\min_{W,Z} L(W, Z) = \sum_{x_n} f(y_n, WZ^T x_n), Wis 1 \times K, Z|V| \times K$

19 Singular Value Decomposition

- Matrix factorization method $\mathbf{X} = \mathbf{U} \mathbf{S} \mathbf{V}^T$
 - \mathbf{U} orthonormal $D \times D$, \mathbf{V} orthonormal $N \times N$
 - \mathbf{S} contains (non-negative) singular values in diagonal in descending order: $D \times N$
 - Columns of \mathbf{U} and \mathbf{V} are the left and right **singular vectors** (eigenvectors of $\mathbf{X} \mathbf{X}^T$ and $\mathbf{X}^T \mathbf{X}$).
- Truncated SVD:**
Take the matrix $\mathbf{S}^{(K)}$ with the K first diagonal elements non zero.
$$\mathbf{X} \approx \mathbf{X}_K = \mathbf{U} \mathbf{S}^{(K)} \mathbf{V}^T$$

20 Principal Component Analysis

- dimensionality reduction and decorrelation
$$\|\mathbf{X} - \hat{\mathbf{X}}\|_F^2 \geq \|\mathbf{X} - \mathbf{U}_k \mathbf{U}_k^T \mathbf{X}\|_F^2 = \sum_{i>K} s_i^2$$
- If the data has zero mean
$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$
$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$
- Columns of \mathbf{U} are called **principal components** and decorrelate \mathbf{X} 's columns.
- Not invariant under scalings \rightarrow normalize \mathbf{X}
- Can compute \mathbf{U} and \mathbf{S} efficiently via $EVD(\mathbf{X} \mathbf{X}^T)$ or $EVD(\mathbf{X}^T \mathbf{X})$
- 21 Neural Net**
 - $x_j^{(l)} = \phi\left(\sum_i w_{i,j}^{(l)} x_i^{(l-1)} + b_j^{(l)}\right)$.

- NN with one hidden layer and sigmoid-like activation function can approximate any sufficiently smooth function on a bounded domain in average ($\leq \frac{(2Cr)^2}{n}$) and point-wise
- Cost function:
$$\mathcal{L} = \frac{1}{N} \sum_{n=1}^N \left(y_n - f^{(L+1)} \circ \dots \circ f^{(1)}(\mathbf{x}_n^{(0)}) \right)^2$$
We can use SGD to minimize the cost.
- 21.1 Backpropagation Algorithm**
 - Forward pass:* Compute
$$\mathbf{z}^{(l)} = \left(\mathbf{W}^{(l)} \right)^T \mathbf{x}^{(l-1)} + \mathbf{b}^{(l)}$$
with
 $\mathbf{x}^{(0)} = \mathbf{x}_n$ and $\mathbf{x}^{(l)} = \phi(\mathbf{z}^{(l)})$.
 - Backward pass:* Set
$$\delta^{(L+1)} = -2(y_n - \mathbf{x}^{(L+1)}) \phi'(z^{(L+1)})$$
 (if squared loss). Then compute

$$\delta_j^{(l)} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l+1)}} \frac{\partial z_k^{(l+1)}}{\partial z_j^{(l)}} = \sum_k \delta_k^{(l+1)} \mathbf{W}_{j,k}^{(l+1)} \phi'(z_j^{(l)})$$

$$\frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}} = \delta_j^{(l)} \mathbf{x}_i^{(l-1)}$$

$$\frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}} = \delta_j^{(l)} \cdot 1 = \delta_j^{(l)}$$

21.2 Activation Functions

- Sigmoid** $\phi(x) = \frac{1}{1+e^{-x}}$ Positive, bounded.
 $\phi'(x) \simeq 0$ for large $|x| \Rightarrow$ Learning slow.
- Tanh** $\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = \phi(2x) - 1/2$.
Balanced, bounded. Learning slow too.
- ReLU** $(x)_+ = \max(0, x)$ Positive, unbounded.
Derivate = 1 if $x > 0$, 0 if $x < 0$
- Leaky ReLU** $f(x) = \max \alpha x, x$ Remove 0 derivate.
- Maxout**
$$f(x) = \max_{\alpha} \mathbf{x}^T \mathbf{w}_1 + b_1, \dots, \mathbf{x}^T \mathbf{w}_k + b_k$$
(Generalization of ReLU)

21.3 Convolutional NN

Sparse connections and *weights sharing*: reduce parameters.

21.4 Reg, Data Augmentation and Dropout

- Regularization term: $\frac{1}{2} \sum_{l=1}^{L+1} \mu^{(l)} \|\mathbf{W}^{(l)}\|_F^2$
- Weight decay is $\Theta[t](1 - \eta\mu)$ in:
$$\Theta[t+1] = \Theta[t] + \eta(\nabla \mathcal{L} + \mu \Theta[t])$$
- Data Augm.: e.g. shift or rotation of pics
- Dropout: avoid overfit. Drop nodes randomly. (Then average multiple drop-NN or divide by dropout rate.)

22 Bayes Net

- Graph example: $p(x, y, z) = p(x)p(y|x)p(z|x)$
: ($y \leftarrow x \rightarrow z$)
- D-Separation** \mathbf{X} and \mathbf{Y} are D-separated by \mathbf{Z} if every path from $x \in \mathbf{X}$ to $y \in \mathbf{Y}$ is blocked by \mathbf{Z} . (\rightarrow independent)
- Blocked Path** contains a variable that
 - is in \mathbf{Z} and is **head-to-tail** or **tail-to-tail**
 - the node is **head-to-head** and neither the node nor any of its descendants are in \mathbf{Z} .
- Markov Blanket** (which blocks node A from the rest of the net) contains:
 - parents of A
 - children of A
 - parents of children of A