PCML Cheat Sheet

1 Math Prerequisites

- Bayes rule

$$p(A,B) = \underbrace{p(A|B)}_{\text{Lik.}} \underbrace{p(B)}_{\text{Prior}} = \underbrace{p(B|A)}_{\text{Post}} \underbrace{p(A)}_{\text{Marg. Lik.}}$$

- Bayes rule: $P(A|B) \propto P(B|A)P(A)$
- Gaussian distribution

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$
$$(\mathbf{x}|\mu, \mathbf{\Sigma}) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{1}{2\sigma^2}(\mathbf{x} - \mu)^T \mathbf{\Sigma}^{-1}(\mathbf{x} - \mu$$

Production of independent variables:

$$\operatorname{Var}(XY) = \mathbb{E}(X^{2}) \mathbb{E}(Y^{2}) - [\mathbb{E}(X)]^{2} [\mathbb{E}(Y)]^{2}$$

Covariance matrix of a data vector x

$$\Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \mathbb{E}(\mathbf{x})) (\mathbf{x}_n - \mathbb{E}(\mathbf{x}))^T$$

 A function f(x) is convex, if for any $x_1, x_2 \in \mathbf{X}$ and for any $0 < \lambda < 1$, we have :

$$f(\lambda x_1 + (1-\lambda)x_2) \le \lambda f(x_1) + (1-\lambda)f(x_2)$$

- The Hessian of a convex function is psd and for a strictly-convex function it's pd.

$$\mathbf{H}_{i,j} = d^2 f / dx_i dx_j$$

1.2 Linear Algebra

Condition number If A is normal $(A^T A = AA^T)$ then

$$k(\mathbf{A}) = \left| \frac{\lambda_{max}(\mathbf{A})}{\lambda_{min}(\mathbf{A})} \right|$$

- A positive definite matrix is symmetric with all positive eigenvalues
- The real symmetric $N \times N$ matrix V is said to be positive semidefinite if

$$\mathbf{a}^T \mathbf{V} \mathbf{a} \geq 0$$

for any real $N \times 1$ vector a.

positive definite if $\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$

2 Cost functions

- Cost functions are used to learn parameters that explain the data well.
- It is essential to make sure that a global minimum exist \rightarrow lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{w}) = \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is convex thus it has only one global minumum value
- MSE is not good when outliers are present.

Mean Absolute Error (MAE):

$$MAE = \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \le \delta \\ \delta|z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

 Huber loss is convex, differentiable, and also robust to outliers but hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Non-convex, non-diff., but robust to outliers. Hinge loss

$Hinge = [1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$ Logistic loss

 $Logistic = log(1 - exp(y_n f(\mathbf{x}_n)))$

3 Regression

Data consists of N pairs (yn, xn)

- 1. y_n the n'th output
- 2. \mathbf{x}_n is a vector of D inputs
- Prediction: predict the ouput for a new input vector.
- Interpretation: understand the effect of inputs on output.
- Outliers are data that are far away from most of the other examples.

3.1 Linear Regression

- Model that assume linear relationship between inputs and the ouput.

 $\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})^W \mathbf{x}^{-1}(\mathbf{x} - \boldsymbol{\mu})^W \mathbf$

 Variance grows only linearly with dimensionality

4 Optimization

4.1 Grid search

– Compute the cost over a grid of M points to find the minimum. Exponential Complexity. Hard to find a good range of values

4.2 Gradient Descent

- GD uses only first-order information and takes steps in the opposite direction of the gradient
- Given cost function $\mathcal{L}(w)$ we want to find w $w = \arg\min_{\boldsymbol{x}} \mathcal{L}(w)$

4.3 Batch Gradient Descent

- Take steps in the opposite direction of the

$$\boldsymbol{w}^{(t+1)} \leftarrow \boldsymbol{w}^{(t)} - \gamma \nabla \mathcal{L}(\boldsymbol{w}^{(t)})$$

with $\gamma > 0$ the learning rate.

With γ too big, method might diverge. With γ too small, convergence is slow.

4.4 Gradients for MSE

We define the error vector e:

$$e := y - Xw$$

- and MSE as follows:

$$\mathcal{L}(\boldsymbol{w}) = \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\boldsymbol{x}}_n^T \boldsymbol{w})^2 = \frac{1}{2N} \boldsymbol{e}^T \boldsymbol{e}$$

- then the gradient is given by

$$\nabla \mathcal{L}(\boldsymbol{w}) = -\frac{1}{N} \mathbf{X}^T \mathbf{e}$$

Optimality conditions:

- 1. necessary: gradient equal zero: $\frac{d\mathcal{L}(\boldsymbol{w}^*)}{d\mathcal{L}(\boldsymbol{w}^*)} = 0$
- 2. sufficient: Hessian matrix is positive
- definite: $\mathbf{H}(\boldsymbol{w}^*) = \frac{d^2 \mathcal{L}(\boldsymbol{w}^*)}{dt}$
- Very sensitive to illconditioning \Rightarrow always normalize features.
- Complexity: O(NDI) with I the number of iterations

4.5 Stochastic Gradient Descent

In ML, most cost functions are formulated as a sum over the training examples:

$$\mathcal{L}(\boldsymbol{w}) = \frac{1}{N} \sum_{n=1}^{N} \mathcal{L}_n(\boldsymbol{w})$$

⇒ SGD algo is given by update rule:

$$\boldsymbol{w}^{(t+1)} \leftarrow \boldsymbol{w}^{(t)} - \gamma \boldsymbol{\nabla} \mathcal{L}_n(\boldsymbol{w}^{(t)})$$

Idea: Cheap but unbiased estimate of grad.

$$\mathbb{E}[\boldsymbol{\nabla}\mathcal{L}_n(\boldsymbol{w})] = \boldsymbol{\nabla}(\boldsymbol{w})$$

4.6 Mini-batch SGD

Update direction $(B \subseteq [N])$:

$$\boldsymbol{g}^{(t)} \coloneqq \frac{1}{|B|} \sum_{n \in B} \boldsymbol{\nabla} \mathcal{L}_n(\boldsymbol{w}^{(t)})$$

Update rule: $\boldsymbol{w}^{(t+1)} \leftarrow \boldsymbol{w}^{(t)} - \gamma \boldsymbol{q}^{(t)}$

4.7 Subgradients (Non-Smooth OPT)

A vector $\mathbf{g} \in \mathbb{R}^D$ s.t.

$$\mathcal{L}(u) \ge \mathcal{L}(w) + g^T(u - w) \quad \forall u$$

is the subgradient to \mathcal{L} at w. If \mathcal{L} is differentiable at w, we have $g = \nabla \mathcal{L}(w)$

5 Least Squares

- Use the first optimality conditions:

$$\nabla L(\boldsymbol{w}^*) = 0 \Rightarrow \boldsymbol{X}^T \boldsymbol{e} = \boldsymbol{X}^T (\boldsymbol{y} - \boldsymbol{X} \boldsymbol{w}) = 0$$

– When $\mathbf{X}^T \mathbf{X}$ is invertible, we have the closed-form expression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{v}$$

- thus we can predict values for a new \mathbf{x}_m $y_m := \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} \mathbf{w}^* = \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$

- The Gram matrix $\mathbf{X}^T \mathbf{X}$ is pd and is also invertible iff X has full column rank.
- Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$
- X can be rank deficient when D > N or when the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. \Rightarrow matrix is ill-conditioned.

6 Maximum Likelihood

– Let define our mistakes $\epsilon_n \sim \mathcal{N}(0, \sigma^2)$.

$$\to y_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

- Another way of expressing this:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} p(y_n|\mathbf{x}_n, \mathbf{w})$$
$$= \prod_{n=1}^{N} \mathcal{N}(y_n|\mathbf{x}_n^T \mathbf{w}, \sigma^2)$$

which defines the likelihood of observating y given X and w

Define cost with log-likelihood

$$\mathcal{L}_{lik}(\boldsymbol{w}) = \log p(\mathbf{y}|\mathbf{X}, \boldsymbol{w})$$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \mathbf{x}_n^T \mathbf{w})^2 + cnst$$

Maximum likelihood estimator (MLE) gives another way to design cost functions

$$\operatorname{argmin} \mathcal{L}_{MSE}(m{w}) = \operatorname{argmax} \mathcal{L}_{lik}(m{w})$$

- MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.
- With Laplace distribution

$$p(y_n|\mathbf{x}_n, \mathbf{w}) = \frac{1}{2b} e^{-\frac{1}{b}|y_n - \mathbf{x}_n^T \mathbf{w}|}$$

$$\sum_{n} \log p(y_n | \mathbf{x}_n, \boldsymbol{w}) = \sum_{n} |y_n - \mathbf{x}_n^T \boldsymbol{w}| + cnst$$

7 Ridge Regression

Linear models usually overfit. One way is to use nonlinear basis functions instead.

$$y_n = w_0 + \sum_{j=1}^{M} w_j \phi_j(\mathbf{x}_n) = \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{w}$$

- This model is linear in w but nonlinear in x. Dimension is now M, not N.
- Polynomial basis

$$\phi(x_n) = [1, x_n, x_n^2, ..., x_n^M]$$
- The least square solution becomes

 $\mathbf{w}_{lse}^* = (\mathbf{\tilde{\Phi}}^T \mathbf{\tilde{\Phi}})^{-1} \mathbf{\tilde{\Phi}}^T \mathbf{y}$

- Complex models overfit easily. Thus we can penalize them with a regularization term

$$\min_{oldsymbol{w}} \left(\mathcal{L}(oldsymbol{w}) + rac{\lambda}{2N} \sum_{j=1}^{M} w_j^2
ight)$$

$$\boldsymbol{w}^* = \underset{\boldsymbol{w}}{\operatorname{argmin}} \left(\frac{1}{2} (\mathbf{y} - \mathbf{X} \boldsymbol{w})^T (\mathbf{y} - \mathbf{X} \boldsymbol{w}) + \frac{\lambda}{2} \boldsymbol{w}^T | \boldsymbol{w} \right)$$

Note that w₀ is not penalized.

By differentiating and setting to zero we get

$$\begin{aligned} \boldsymbol{w}_{ridge} &= (\boldsymbol{\tilde{\Phi}}^T \boldsymbol{\tilde{\Phi}} + \boldsymbol{\Lambda})^{-1} \boldsymbol{\tilde{\Phi}}^T \mathbf{y} \\ \boldsymbol{\Lambda} &= \begin{bmatrix} 0 & 0 \\ 0 & \lambda \boldsymbol{I}_m \end{bmatrix} \end{aligned}$$

- Ridge regression improves the condition number of the Gram matrix since the eigenvalues of $(\tilde{\Phi}^T \tilde{\Phi} + \lambda I_m)$ are at least λ
- Maximum-a-posteriori (MAP)
 - Maximizes the product of the likelihood and the prior.

$$w_{MAP} = \underset{w}{\operatorname{argmax}} (p(\mathbf{y}|\mathbf{X}, \mathbf{\Lambda})p(w|\mathbf{\Sigma}))$$

- Assume $w_0 = 0$

$$egin{align*} oldsymbol{w}_{ridge} = rgmax \left(\log \left[\prod_{n=1}^{N} \mathcal{N}(y_n | \mathbf{x}_n^T oldsymbol{w}, oldsymbol{\Lambda}) imes \mathcal{N}(oldsymbol{w} | 0, \mathbf{I}) \right] & \text{where S is diagonals} \\ S_{nr} & \text{constant} \end{aligned}$$

 Lasso regularizer forces some w_i to be strictly 0 and therefore forces sparsity in the

$$\min_{\boldsymbol{w}} \frac{1}{2N} \sum_{n=1}^{N} (y_n - \tilde{\boldsymbol{\phi}}(\mathbf{x}_n)^T \boldsymbol{w})^2,$$

such that $\sum_{i=1}^{M} |w_i| \leq \tau$

 The expected test error can be expressed as the sum of two terms

- Squared bias: The average shift of the
- predictions - Variance: measure how data points vary around their average.

expected loss =
$$(bias)^2$$
 + variance + noise

- Model bias and estimation bias are important
- RR increases estimation bias and reduces var
- Model more complex increases test error Small λ → low bias but large variance
- Large λ → large bias but low variance $err = \sigma^2 + \mathbb{E}[f_{lse} - \mathbb{E}[f_{lse}]]^2 + [f_{true} - \mathbb{E}[f_{lse}]]^2$

9 Logistic Regression

- Classification relates input variables x to
- discrete output variable yBinary classifier: we use y = 0 for C_1 and
- y = 1 for \mathbf{C}_2 . Can use least-squares to predict û*

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(y_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$
$$p(y_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$$

- The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \boldsymbol{w}) = \prod_{n=1}^{N} \sigma(\boldsymbol{x}_{n}^{T} \boldsymbol{w})^{y_{n}} (1 - \sigma(\boldsymbol{x}_{n}^{T} \boldsymbol{w}))^{1 - y_{n}}$$

$$\mathcal{L}_{MLE}(\boldsymbol{w}) = \sum_{n=1}^{N} \left(y_n \boldsymbol{x}_n^T \boldsymbol{w} - \log(1 + \exp(\boldsymbol{x}_n^T \boldsymbol{w})) \right)$$

We can use the fact that

$$\frac{d}{dx}\log(1+\exp(x)) = \sigma(x)$$

Gradient of the log-likelihood

$$\mathbf{g} = \frac{d\mathcal{L}}{d\mathbf{w}} = \sum_{n=1}^{N} \left(\mathbf{x}_n \mathbf{y}_n - \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w}) \right)$$
$$= -\mathbf{X}^T [\sigma(\mathbf{X} \mathbf{w}) - \mathbf{y}]$$

The negative of the log-likelihood $-\mathcal{L}_{mle}(\boldsymbol{w})$ is convex

Hessian of the log-likelihood

- We know that

- We know that $\frac{d\sigma(t)}{dt} = \sigma(t)(1-\sigma(t))$ - Hessian is the derivative of the gradient

$$\begin{aligned} \mathbf{H}(\boldsymbol{w}) &= -\frac{d\mathbf{g}(\boldsymbol{w})}{d\boldsymbol{w}^T} = \sum_{n=1}^N \frac{d}{d\boldsymbol{w}^T} \sigma(\boldsymbol{x}_n^T \boldsymbol{w}) \boldsymbol{x}_n \\ &= \sum_{n=1}^N \boldsymbol{x}_n \sigma(\boldsymbol{x}_n^T \boldsymbol{w}) (1 - \sigma(\boldsymbol{x}_n^T \boldsymbol{w})) \boldsymbol{x}_n^T \end{aligned}$$

where **S** is a $N \times N$ diagonal matrix with

 $S_{nn} = \sigma(\boldsymbol{x}_n^T \boldsymbol{w})(1 - \sigma(\boldsymbol{x}_n^T \boldsymbol{w}))$ The negative of the log-likelihood is not

strictly convex. Newton's Method

- Uses second-order information and takes steps in the direction that minimizes a quadratic approximation

$$\mathcal{L}(oldsymbol{w}) = \mathcal{L}(oldsymbol{w}^{(k)}) +
abla \mathcal{L}_k^T (oldsymbol{w} - oldsymbol{w}^{(k)})$$

$$\begin{aligned} &+(\boldsymbol{w}-\boldsymbol{w}^{(k)})^T\mathbf{H}_k(\boldsymbol{w}-\boldsymbol{w}^{(k)})\\ \text{and it's minimum is at} & \\ &\boldsymbol{w}^{k+1}=\boldsymbol{w}^{(k)}-\gamma_k\mathbf{H}_k^{-1}\nabla\mathcal{L}_k \end{aligned}$$

- Complexity: $O((ND^2 + D^3)I)$

Penalized Logistic Regression
$$\min_{\boldsymbol{w}} \left(-\sum_{n=1}^{N} \log p(y_n | \mathbf{x}_n^T \boldsymbol{w}) + \lambda \sum_{d=1}^{D} w_d^2 \right)$$

10 Generalized Linear Model

Exponential family distribution
$$p(\mathbf{y}|\boldsymbol{\eta}) = \frac{h(y)}{Z} \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

Bernoulli distribution

$$p(y|\mu) = \mu^{y} (1 - \mu)^{1 - y}$$

$$= \exp(y \log(\frac{\mu}{1 - \mu} + \log(1 - \mu)))$$

there is a relationship between η and μ

throught the link function
$$\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$$

Note that μ is the mean parameter of y

- Relationship between the mean
$$\mu$$
 and η is defined using a link function g

$$\eta = g(\mu) \Leftrightarrow \mu = g^{-1}(\eta)$$
First and second derivatives of $A(\eta)$ are related to the mean and the variance

 $\frac{dA(\eta)}{d\eta} = \mathbb{E}[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = \operatorname{Var}[\phi(\eta)]$

- $A(\eta)$ is convex - The generalized maximum likelihood cost to

$$\min_{\boldsymbol{w}} \mathcal{L}(\boldsymbol{w}) = -\sum_{n}^{N} \log(p(y_n | \boldsymbol{x}_n^T \boldsymbol{w}))$$

where $p(y_n|\boldsymbol{x}_n^T\boldsymbol{w})$ is an exponential family distribution

We obtain the solution
$$\frac{d\mathcal{L}}{d\mathbf{r}} = \mathbf{X}^T [\mathbf{g}^{-1}(\boldsymbol{\eta}) - \boldsymbol{\phi}(\mathbf{y})]$$

11 k-Nearest Neighbor (k-NN)

- The k-NN prediction for x is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} y_n$$

where $nbh_k(\mathbf{x})$ is the neightborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n .

- Curse of dimensionality: Generalizing correctly becomes exponentially harder as the dimensionality grows.
- Gathering more inputs variables may be bad

12 Support Vector Machine

- Combination of the kernel trick plus a modified loss function (Hinge loss)
- Solution to the dual problem is sparse and non-zero entries will be our support
- Kernelised feature vector where μ_k are centroids

$$\phi(\mathbf{x}) = [k(\mathbf{x}, \boldsymbol{\mu}_1), ..., k(\mathbf{x}, \boldsymbol{\mu}_K)]$$

- In practice we'll take a subset of data points to be prototype -> sparse vector machine. Assume $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$\mathcal{L}(\boldsymbol{w}) = \min_{\boldsymbol{w}} \sum_{n=1}^{N} [1 - y_n \tilde{\boldsymbol{\phi}}_n^T \boldsymbol{w}]_+ + \frac{\lambda}{2} \sum_{j=1}^{M} w_j^2$$

- Minimum doesn't change with a rescaling of
- choose the hyperplane so that the distance from it to the nearest data point on each side is maximized

Duality:

- Hard to minimize g(w) so we define $\mathcal{L}(\boldsymbol{w}) = \max G(\boldsymbol{w}, \boldsymbol{\alpha})$
- we use the property that $C[v_n]_+ = \max(0, Cv_n) = \max_{\alpha_n \in [0, C]} \alpha_n v_n$ - We can rewrite the problem as

$$\min_{\boldsymbol{w}} \max_{\alpha \in [0,C]^{N}} \sum_{n=1}^{N} \alpha_{n} (1 - y_{n} \boldsymbol{\phi}_{n}^{T} \boldsymbol{w}) + \frac{1}{2} \sum_{j=1}^{M} w_{j}^{2}$$

- This is differentiable, convex in \boldsymbol{w} and concave in α

- Minimax theorem:

 $\min_{\boldsymbol{\alpha}} \max_{\boldsymbol{\alpha}} G(\boldsymbol{w}, \boldsymbol{\alpha}) = \max_{\boldsymbol{\alpha}} \min_{\boldsymbol{\alpha}} G(\boldsymbol{w}, \boldsymbol{\alpha})$ because G is convex in w and concave in

- Derivative w.r.t. w:

$$\nabla_{\boldsymbol{w}}G(\boldsymbol{w}, \boldsymbol{\alpha}) = -\sum_{n=1}^{N} \alpha_{n}y_{n}\boldsymbol{x}_{n} + \lambda \boldsymbol{w}$$

 $\begin{array}{c}
\stackrel{\longleftarrow}{n=1} \\
- \text{ Equating this to 0, we get:}
\end{array}$

$$w(\alpha) = \frac{1}{\lambda} \sum_{n=1}^{N} \alpha_n y_n x_n = \frac{1}{\lambda} X Y \alpha$$

Y := diag(y)

 $\mathbf{r} := \operatorname{diag}(y)$ – Plugging \mathbf{w}^* back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,1]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2\lambda} \boldsymbol{\alpha}^T \mathbf{Y} \boldsymbol{X}^T \boldsymbol{X} \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{X}^T \mathbf{X}$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision boundary.

13 Kernel Ridge Regression

- The following is true for ridge regression

$$w^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y}$$
$$= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \boldsymbol{\alpha}^*$$

 Complexity of computing w: (1) $O(D^2N + D^3)$, (2) $O(DN^2 + N^3)$ - Thus we have

 $\boldsymbol{w}^* = \mathbf{X} \boldsymbol{\alpha}^*, \quad \text{with } \boldsymbol{w}^* \in \mathbb{R}^D \text{ and } \boldsymbol{\alpha}^* \in \mathbb{R}^N$

- The representer theorem allows us to write an equivalent optimization problem in terms

or
$$\alpha$$
.
 $\alpha = \underset{\alpha}{\operatorname{argmax}} \left(-\frac{1}{2} \alpha^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \alpha + \alpha^T \mathbf{y} \right)$
 $\mathbf{K} = \mathbf{X} \mathbf{X}^T$ is called the **kernel matrix** or

- Gram matrix.
- If K is positive definite, then it's called a Mercer Kernel.
- $\begin{array}{ll} & \mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j) \\ & \text{If the kernel is Mercer, then there exists a} \end{array}$ function $\phi(\mathbf{x})$ s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

- Kernel trick:
- compute dot-product in \mathbb{R}^m while remaining in \mathbb{R}^n Replace $\langle \mathbf{x}, \mathbf{x}' \rangle$ with $k(\mathbf{x}, \mathbf{x}')$.
- Common Kernel
- $\begin{array}{ll} & \text{Polynomial Kernel: } (\gamma \langle \mathbf{x}_i, \mathbf{x}_j \rangle + r)^d \\ & \text{Radial Basis function kernel (RBF)} \end{array}$

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Sigmoid Kernel: $tanh(\langle \mathbf{x}_i, \mathbf{x}_i \rangle + r)$
- Properties of kernels to ensure the existance of a corresponding ϕ :
 - symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$
- positive semi-definite.
- Thus we get

$$\mathbf{y} = \boldsymbol{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

14 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances outside the cluster.

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = \sum_{k=1}^{K} \sum_{n=1}^{N} z_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

such that $z_{nk} \in \{0,1\}$ and $\sum_{k=1}^K z_{nk} = 1$ K-means algorithm (Coordinate Descent): Initialize μ_k , then iterate

1. For all n, compute \mathbf{z}_n given $\boldsymbol{\mu}$

$$z_{nk} = \begin{cases} 1 & \text{if } k = \operatorname{argmin}_{j} ||\mathbf{x}_{n} - \boldsymbol{\mu}||_{2}^{2} \\ 0 & \text{otherwise} \end{cases}$$

2. For all
$$k$$
, compute μ_k given \mathbf{z}
$$\mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}}$$

- A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{z}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \mathbf{I}) \right]^{z_{nk}}$$

- K-means as a Matrix Factorization $\min \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{MZ}^T||_{\text{Frob}}^2$

Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

15 Gaussian Mixture Models

Clusters can be elliptical using a full covariance matrix instead of isotropic

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{z}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{z_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining z_n to be a random

$$p(z_n=k)=\pi_k \text{ where } \pi_k>0, \forall k, \sum_{k=1}^K \pi_k=1$$

Joint distribution of Gaussian mixture model

$$p(\mathbf{X}, \mathbf{z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^{N} p(\mathbf{x}_n | \mathbf{r}_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{z}_n | \boldsymbol{\pi})$$

$$=\prod_{n=1}^{N}\prod_{k=1}^{K}[(\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k},\boldsymbol{\Sigma}_{k}))^{z_{n}k}]\prod_{k=1}^{K}[\boldsymbol{\pi}]^{z_{n}k}$$

- z_n are called latent unobserved variables
- Unknown parameters are given by $\theta = \{\mu, \Sigma, \pi\}$
- We get the marginal likelihood by marginalizing z_n out from the likelihood

$$p(\mathbf{x}_n|\boldsymbol{\theta}) = \sum_{k=1}^K p(\mathbf{x}_n, z_n = k|\boldsymbol{\theta})$$

$$= \sum_{k=1}^K p(z_n = k|\boldsymbol{\theta})p(\mathbf{x}_n|z_n = k, \boldsymbol{\theta})$$

$$= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to $O(D^2K)$
- To get maximum likelihood estimate of θ , we

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

16 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\theta^{(t)}$ with equality when.

$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$

2. Maximization step: Update θ

$$\boldsymbol{\theta}^{(t+1)} = \operatorname*{argmax}_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$$

$$\boldsymbol{\mu}_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk}) \mathbf{x}_{n}}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\Sigma_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} q_{kn}^{(t)} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t+1)})^{T}}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\pi_k^{(t+1)} = \frac{1}{N} \sum_{n=1}^N q_{kn}^{(t)}$$

If covariance is diagonal → K-means.

17 Matrix factorization

- We have D movies and N users
- **X** is a matrix $D \times N$ with x_{dn} the rating of n'th user for d'th movie.
- We project data vectors x_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
- \mathbf{Z} a $N \times K$ matrix that gives features for the users
- W a D × K matrix that gives features for the movies

 $x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$ We can add a regularizer and minimize the $\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{n} \left[[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}]^2 \right]$

$$+\frac{\lambda_w}{2} \sum_{d=1}^{D} \mathbf{w}_d^T \mathbf{w}_d + \frac{\lambda_z}{2} \sum_{n=1}^{N} \mathbf{z}_n^T \mathbf{z}_n$$

We can use coordinate descent algorithm, by first minimizing w.r.t. **Z** given **W** and then minimizing W given Z. This is called

Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_{\mathbf{w}} \mathbf{I}_{\mathbf{w}})^{-1} \mathbf{W}^T \mathbf{X}$$

$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T\mathbf{Z} + \lambda_w \mathbf{I}_K)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- Complexity: $O(DNK^2 + NK^3) \rightarrow O(DNK^2)$ Probabilistic model

$$\prod_{(d,n)\in\Omega} \mathcal{N}(x_{dn}|\mathbf{w}_{d}^{T}\mathbf{z}_{n}, I) \times \prod_{n=1}^{N} \mathcal{N}(\mathbf{z}_{n}|0, \frac{1}{\lambda_{z}}I)$$
$$\times \prod_{d=1}^{D} \mathcal{N}(\mathbf{w}_{d}|0, \frac{1}{\lambda_{w}}I)$$

- Since many ratings are missing we cannot normalize the data. A solution is to add

$$\frac{1}{2} \sum_{(d,n) \in \Omega} (x_{dn} - \mathbf{w}_d^T \mathbf{z}_n - w_{0d} - z_{0n} - \mu)^2$$

18 Singular Value Decomposition

 Matrix factorization method $X = USV^T$

$$\mathbf{X} = \mathbf{USV}^{T}$$
- \mathbf{U} is a unitary $D \times D$ matrix

- V is a unitary N × N matrix - S is a non-negative diagonal matrix of size
- $D \times N$ which are called **singular values** appearing in a descending order.
- Columns of U and V are the left and right singular vectors respectively.

Assuming D < N we have

$$\mathbf{X} = \sum_{d=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular

values contain the high-frequency information. Dimensionality Reduction Take the matrix $S^{(K)}$ with the K first diagonal elements non zero. Then, rank-Kapprox:

$$\mathbf{X} \approx \mathbf{X}_K = \mathbf{U}\mathbf{S}^{(K)}\mathbf{V}^T$$

18.1 Principal Componement Analysis

- PCA is a dimensionality reduction method and a method to decorrelate the data $\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$ such that columns of \mathbf{W} are orthogonal.
- If the data is zero mean

$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$
$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

 Thus the columns of matrix U are called the principal components and they decorrelate the covariance matrix

 Using SVD, we can compute the matrices in the following way

$$\mathbf{W} = \mathbf{U}\mathbf{S}_D^{1/2}, \mathbf{Z}^T = \mathbf{S}^{1/2}\mathbf{V}^T$$

19 Neural Net

Basic structure: One input layer of size D, L hidden layers of size K, and one output layer. (feedforward network).

 $x_{i}^{(l)} = \phi \left(\sum_{i} w_{i,i}^{(l)} x_{i}^{(l-1)} + b_{i}^{(l)} \right)$

NN can represent the Rienmann sum with only two layers ⇒ It's powerful!

Cost function:

$$\frac{1}{N} \sum_{n=1}^{N} \left(y_n - f^{(L+1)} \circ \dots \circ f^{(1)}(\boldsymbol{x}_n^{(0)}) \right)^2$$
 We can use SGD to minimize the cost

Compact form:
$$\boldsymbol{W}_{i,j}^{(l)} = w_{i,j}^{(l)}$$

 $\boldsymbol{x}^{(l)} = f^{(l)}(\boldsymbol{x}^{(l-1)}) = \phi\left(\left(\boldsymbol{W}^{(l)}\right)^T \boldsymbol{x}^{(l-1)} + \boldsymbol{b}^{(l)}\right)$

19.1 Backpropagation Algorithm

- Forward pass: Compute

$$\mathbf{z}^{(l)} = (\mathbf{W}^{(l)})^T \mathbf{x}^{(l-1)} + \mathbf{b}^{(l)}$$
 with $\mathbf{x}^{(0)} = \mathbf{x}_n$ and $\mathbf{x}^{(l)} = \phi(\mathbf{z}^{(l)})$. Backward pass: Set

 $\delta^{(L+1)} = -2(y_n - \boldsymbol{x}^{(L+1)})\phi'(z^{(L+1)})$ (if squared loss). Then compute

$$\begin{split} \delta_j^{(l)} &= \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l+1)}} \frac{\partial z_k^{(l+1)}}{\partial z_j^{(l)}} \\ &= \sum_l \delta_k^{(l+1)} W_{j,k}^{(l+1)} \phi'(z_j^{(l)}) \end{split}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}} \\ &= \delta_j^{(l)} \mathbf{x}_i^{(l-1)} \end{split}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}} \\ &= \delta_z^{(l)} \cdot 1 = \delta_z^{(l)} \end{split}$$

19.2 Activation Functions

Sigmoid $\phi(x) = \frac{1}{1+e^{-x}}$ Positive, bounded

$$\phi'(x) \simeq 0$$
 for large $|x| \Rightarrow$ Learning slow.
Tanh $\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = \phi(2x) - 1/2$.

Balanced, bounded. Learning slow too.

ReLU $(x)_{+} = max 0, x Positive, unbounded$ Derivate = 1 if x > 0, 0 if x < 0Leaky ReLU $f(x) = \max \alpha x, x$ Remove 0

derivative

 $f(x) = max \mathbf{x}^T \mathbf{w}_1 + b_1, ..., \mathbf{x}^T \mathbf{w}_k + b_k$ (Generalization of ReLU)

19.3 Convolutional NN Sparse connections and weights sharing: reduce

on neighbours) 19.4 Reg, Data Augmentation and

Dropout – Regularization term: $\frac{1}{2} \sum_{l=1}^{L+1} \mu^{(l)} ||W^{(l)}||_F^2$

complexity. (e.g. pixels in pictures only depend

- Data Augm.: e.g. shift or rotation of pics Dropout: avoid overfit. Drop nodes

randomly. (Then average multiple drop-NN) 20 Bayes Net

- Graph example: p(x, y, z) = p(y|x)p(z|x)p(x) $: (y \leftarrow x \rightarrow z)$

D-Separation X and Y are D-separated by Z if every path from $x \in X$ to $y \in Y$ is blocked by Z.

Blocked Path contains a variable that

- is in Z and is head-to-tail or tail-to-tail

- the node is head-to-head and neither the node nor the descendant are in Z. Markov Blanket (which blocks node A from

- parents of A

- children of A

parents of children of A

the rest of the net) contains: