ML Cheat Sheet

1 Math Prerequisites

1.1 Derivatives

$$\begin{array}{ll} - & \partial(\mathbf{X}\mathbf{Y}) = (\partial\mathbf{X})\mathbf{Y} + \mathbf{X}(\partial\mathbf{Y}) \\ - & \frac{\partial\mathbf{f}(\mathbf{g}(\mathbf{u}(\mathbf{x})))}{\partial\mathbf{x}} = \frac{\partial\mathbf{u}(\mathbf{x})}{\partial\mathbf{x}} \frac{\partial\mathbf{g}(\mathbf{u})}{\partial\mathbf{u}} \frac{\partial\mathbf{f}(\mathbf{g})}{\partial\mathbf{g}} \end{array}$$

$$- \frac{\partial \mathbf{x}^T \mathbf{a}}{\partial \mathbf{x}} = \frac{\partial \mathbf{a}^T \mathbf{x}}{\partial \mathbf{x}} = \mathbf{a}$$

$$-\frac{\partial \mathbf{a}^T \mathbf{X} \mathbf{b}}{\partial \mathbf{X}} = \mathbf{a} \mathbf{b}^T$$
$$-\frac{\partial \mathbf{a}^T \mathbf{X}^T \mathbf{b}}{\partial \mathbf{X}} = \mathbf{b} \mathbf{a}^T$$

$$-\frac{\partial \mathbf{X}}{\partial \mathbf{X}_{ij}} = \mathbf{J}^{ij}, J^{ij} \text{ is the single entry matrix}$$

$$- \frac{\partial \mathbf{b}^T \mathbf{x}^T \mathbf{x} \mathbf{c}}{\partial \mathbf{x}} = \mathbf{X} \left(\mathbf{b} \mathbf{c}^T + \mathbf{c} \mathbf{b}^T \right)$$
$$- \frac{\partial \mathbf{x}^T \mathbf{B} \mathbf{x}}{\partial \mathbf{x}^T} = \left(\mathbf{B} + \mathbf{B}^T \right) \mathbf{x}$$

$$\frac{\partial \mathbf{x}}{\partial \mathbf{x}} = (\mathbf{B} + \mathbf{B})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s}) = \frac{\partial}{\partial \mathbf{x}} (\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W} (\mathbf{x} - \mathbf{A}\mathbf{s})$$

$$-\frac{\partial}{\partial \mathbf{x}}(\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s}) = 2\mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s})$$

$$- \frac{\partial}{\partial \mathbf{X}} \|\mathbf{X}\|_{\mathrm{F}}^{2} = \frac{\partial}{\partial \mathbf{X}} \operatorname{Tr} \left(\mathbf{X} \mathbf{X}^{H} \right) = 2\mathbf{X}$$

1.2 Linear Algebra

- positive definite (pd) if
$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

$$- (\mathbf{x} - \mathbf{b})^T (\mathbf{x} - \mathbf{b}) = \|\mathbf{x} - \mathbf{b}\|_2^2$$

-
$$\|\mathbf{X}\|_F = \|\mathbf{X}^T\|_F$$

- $det X = x_{11}x_{22} - x_{12}x_{21}$, if X is 2×2

Valid distribution p(x) > 0, $\forall x$ and $\sum p(x) = 1$ Model is identifiable iff $\theta_1 = \theta_2 \rightarrow P_{\theta_1} = P_{\theta_2}$

Gaussian (Not convex):

$$\begin{split} \mathcal{N}(x|\mu,\sigma^2) &= \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2}) \\ \mathcal{N}(x|\mu,\Sigma^2) &= \frac{\exp(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu))}{\sqrt{(2\pi)^D \det(\Sigma)}} \end{split}$$

- Poisson: P(k events in interval) = $e^{-\lambda} \frac{\lambda^k}{k!}$

- Bernoulli: $p(y|\mu) = \mu^{y}(1-\mu)^{1-y}$

1.4 Convexity

A function f(x) is convex if

- for any $\mathbf{x}_1, \mathbf{x}_2 \in \mathbf{X}$ and $0 \le \lambda \le 1$, we have : $f(\lambda \mathbf{x}_1 + (1 - \lambda)\mathbf{x}_2) \le \lambda f(\mathbf{x}_1) + (1 - \lambda)f(\mathbf{x}_2)$

it is a sum of convex functions

- composition of convex and linear functions

- f(x) = g(h(x)), g,h are convex, g increasing the Hessian H is positive semi-definite

2 Cost functions

Mean square error (MSE):

$$MSE(\boldsymbol{w}) = \frac{1}{N} \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

MSE is strictly convex thus it has only one global minumum value.

MSE is very prone to outliers.

Mean Absolute Error (MAE):

$$MAE = \frac{1}{N} \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

- MAE is more robust to outliers than MSE.

Huber loss

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \leq \delta \\ \delta |z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

 Huber loss is convex, differentiable, and also robust to outliers but hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Non-convex, non-diff., but robust to outliers. Hinge loss:

 $[1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$ Logistic loss: $\log(1 - \exp(y_n f(\mathbf{x}_n)))$

3 Optimization

- Local minimum:

 $L(w^*) \le L(w) \ \forall w : ||w - w^*|| < \epsilon$ Global minimum: $L(w^*) < L(w) \forall w$

3.1 Grid search

 Compute the cost over a grid of V points. Exponential complexity $\mathcal{O}(|V|^D)$. Hard to find a good value range. No guarantee to

3.2 GD - Gradient Descent (Batch)

- GD uses only first-order information

- Given cost function $\mathcal{L}(\mathbf{w})$ we want to find \mathbf{w} $\mathbf{w} = \arg\min \mathcal{L}(\mathbf{w})$

- Take steps in the opposite direction of gradient

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})$$

 With γ too big, method might diverge. With γ too small, convergence is slow.

 Very sensitive to ill-conditioning ⇒ always normalize features ⇒ allow different directions to converge at same speed.

3.3 SGD - Stochastic Gradient Descent

SGD update rule (only n-th training example):

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Idea: Cheap but unbiased estimate of grad.

$$\mathbb{E}[\nabla \mathcal{L}_n(\mathbf{w})] = \nabla L(\mathbf{w})$$

Robbins-Monroe condition: $\quad \quad - \quad \gamma^{(t)} : \sum_{t=1}^{\infty} \gamma^{(t)} = \infty; \sum_{t=1}^{\infty} (\gamma^{(t)})^2 < \infty$

- e.g. $\gamma^{(t)} = 1/(t+1)^r, r \in (0.5, 1)$

3.4 Mini-batch SGD

Update direction $(B \subset [N])$:

$$g^{(t)} := \frac{1}{|B|} \sum_{n \in B} \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Update rule : $\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \mathbf{g}^{(t)}$

3.5 Gradients for MSE

Define error e := y − Xw

- and MSE as follows:

$$\mathcal{L}(\mathbf{w}) = \frac{1}{2N} \sum_{n=1}^{N} (\mathbf{y}_n - \tilde{\mathbf{x}}_n^T \mathbf{w})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

Optimality condition

1. necessary: $\frac{d\mathcal{L}(\mathbf{w}^*)}{d\mathbf{w}} = -\frac{1}{N}\mathbf{X}^T\mathbf{e} = 0$ 2. sufficient: Hessian matrix is positive

definite: $\mathbf{H}(\mathbf{w}^*) = \frac{d^2 \mathcal{L}(\mathbf{w}^*)}{d\mathbf{w} d\mathbf{w}^T} = \frac{1}{N} X^T X$

3.6 Subgradients (Non-Smooth OPT)

A vector $\mathbf{g} \in \mathbb{R}^D$ s.t.

$$\mathcal{L}(\mathbf{u}) \ge \mathcal{L}(\mathbf{w}) + \mathbf{g}^T(\mathbf{u} - \mathbf{w}) \quad \forall \mathbf{u} \in \mathbb{R}^D$$

is the subgradient to \mathcal{L} at \mathbf{w} . If \mathcal{L} is differentiable at \mathbf{w} , we have $\mathbf{g} = \nabla \mathcal{L}(\mathbf{w})$

3.7 Constrained Optimization

Find solution min $\mathcal{L}(\mathbf{w})$ s.t. $\mathbf{w} \in \mathcal{C}$

 Add proj. onto C after each step: $P_{\mathcal{C}}(\mathbf{w}') = \arg\min |\mathbf{v} - \mathbf{w}'|, \mathbf{v} \in \mathcal{C}$

 $\mathbf{w}^{(t+1)} = P_{\mathcal{C}}[\mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})]$

- Use penalty functions

 $-\min \mathcal{L}(\mathbf{w}) + I_{\mathcal{C}}, I_{\mathcal{C}} = 0 \text{ if } \mathbf{w} \in \mathcal{C}, \text{ ow } + \infty$

 $-\min \mathcal{L}(\mathbf{w}) + \lambda |\mathbf{A}\mathbf{w} - \mathbf{b}|$ Stopping criteria when L(w) close to 0

3.8 Iteration complexities for MSE/MAE $GD = \mathcal{O}(ND)$

- MB-GD= $\mathcal{O}(BD)$ $- SGD = \mathcal{O}(D)$

4 Least Squares

- Use the first optimality conditions:

$$\nabla L(\mathbf{w}^*) = 0 \Rightarrow \mathbf{X}^T \mathbf{e} = \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}) = 0$$

When $\mathbf{X}^T\mathbf{X}$ is invertible, we have the closed-form expression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{v}$$

- thus we can predict values for a new \mathbf{x}_m

$$\mathbf{y}_m := \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} \mathbf{w}^* = \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

- The Gram matrix X^TX is pd and is also

invertible iff X has full column rank. Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$

 \mathbf{X} can be rank deficient when D > N or when the columns $\bar{\mathbf{x}}_d$ are nearly collinear. \Rightarrow matrix is ill-conditioned.

Can still solve using a linear system solver using normal equations:

$$\mathbf{X}^{\top}\mathbf{X}\mathbf{w} = \mathbf{X}^{\top}\mathbf{y}$$

5 Maximum Likelihood (MLE)

Let define the noise ε_n ~ N(0, σ²).

 $\rightarrow \mathbf{y}_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$ Another way of expressing this:

$$\begin{aligned} p(\mathbf{y}|\mathbf{X}, \mathbf{w}) &= \prod_{n=1}^{N} p(\mathbf{y}_{n}|\mathbf{x}_{n}, \mathbf{w}) \\ &= \prod_{n=1}^{N} \mathcal{N}(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}, \sigma^{2}) \end{aligned}$$

which defines the likelihood of observating y given \mathbf{X} and \boldsymbol{w}

 Define cost with log-likelihood $\mathcal{L}_{MLE}(\mathbf{w}) = \log p(\mathbf{y}|\mathbf{X}, \mathbf{w})$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (\mathbf{y}_n - \mathbf{x}_n^T \mathbf{w})^2 + cnst$$

 Maximum likelihood estimator (MLE) gives another way to design cost functions $\operatorname{argmin} \mathcal{L}_{MSE}(\mathbf{w}) = \operatorname{argmax} \mathcal{L}_{MLE}(\mathbf{w})$

MLE can also be interpreted as finding the model under which the observed data is most likely to have been generated from.

 $\mathbf{w}_{\mathrm{MLE}} \to \mathbf{w}_{\mathrm{true}}$ for large amount of data

6 Ridge Regression and LASSO

- Add regularization term

 $\min \mathcal{L}(\mathbf{w}) + \Omega(\mathbf{w})$

- L_2 -Reg. (Ridge): $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_2^2$ → small values of w_i, not sparse $-\rightarrow \mathbf{w}^* = (\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \text{ with } \lambda' = 2N\lambda$

 $- \rightarrow (\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1}$ exists (lifted eigenvalues)

- L_1 -Reg. (Lasso): $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_1$ → sparsity of weight vector

 → implicit model selection - Maximum-a-posteriori (MAP)

$$p(\mathbf{y}|\mathbf{X}\mathbf{w}) = \prod_{n=1}^{N} \mathcal{N}(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}, \sigma_{n}^{2})$$

 $p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|0, \sigma_0^2 \mathbf{I}_D)$ then $\rightarrow \mathbf{w}^* = \operatorname{argmax} p(\mathbf{y}|\mathbf{X}\mathbf{w}) \cdot p(\mathbf{w})$

$$\mathbf{w}^{\star} = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{\mathbf{w}}^{N} \frac{1}{2\sigma_{n}^{2}} (\mathbf{y}_{n} - \mathbf{x}^{T} \mathbf{w})^{2} + \frac{1}{2\sigma_{0}^{2}} ||\mathbf{w}||^{2}$$

7 Model Selection

 Generalisation error: L_D(f) = E[l(y, f(x))], but D normally unknown.

Instead approximate by $L_{Stest}(f_{train}) =$ $\frac{1}{|S_{test}|} \sum_{S_{test}} l(y_n, f_{S_{train}}(x_n))$

In expectation this equates the true error. - Worst case, if comparing K models:

 $-\mathbb{P}[\max_{k} |L_{D}(f_{k}) - L_{test}(f_{k})| >$ $\left| \frac{(bia)^2 \ln(2K/\delta)}{(bia)^2 \ln(2K/\delta)} \right| < \delta$

- Error decreases as $\mathbb{O}(1/\sqrt{|S_{test}|})$

Error only goes up by √ln(K) for testing K

use cross-validation for an efficient, unbiased estimate of generalisation error and variance.

8 Bias-Variance decomposition

 Simple (e.g. large λ) → large bias, but low variance

 Complex (e.g. small λ) → low bias, but large variance

The expected squared loss between true model and learned model is a sum of three non-negative terms:

 $\mathbb{E}_{S}[(f(x) + \epsilon - f_{S}(x))^{2}] = Var[\epsilon] + bias +$

- Bias = $(f(x) - \mathbb{E}_{S'}[f_{S'}(x)])^2$: Difference between actual value and expected

Variance = $\mathbb{E}_S[(\mathbb{E}_{S'}[f_{S'}(x)] - f_S(x)])^2]$: variance of predictions between training

All terms are lower bounds for the error.

Cannot do better than Var[ε].

9 Logistic Regression

 Binary classifier: use y ∈ {0, 1}. Can use least-squares to predict ŷ_{*}

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5\\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$

$$p(\mathbf{y}_n = 1 | \mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$

 $p(\mathbf{y}_n = 0 | \mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$ The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} \sigma(\mathbf{x}_{n}^{T} \mathbf{w})^{\mathbf{y}_{n}} (1 - \sigma(\mathbf{x}_{n}^{T} \mathbf{w}))^{1 - \mathbf{y}_{n}}$$

The negative log-likelihood (w.r.t. MLE):

The negative log-likelihood (w.r.t. MLE): related to the mean and the variance
$$\mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \mathbf{y}_n \ln \sigma(\mathbf{x}_n^T \mathbf{w}) + (1 - \mathbf{y}_n) \ln (1 - \sigma(\mathbf{x}_n^T \mathbf{w})) \quad \frac{dA(\eta)}{d\eta} = \mathbb{E}[\phi(\eta)], \quad \frac{d^2A(\eta)}{d\eta^2} = \operatorname{Var}[\phi(\eta)]$$

 $= \sum_{n=1}^{N} \ln[1 + \exp(\mathbf{x}_n^T \mathbf{w})] - \mathbf{y}_n \mathbf{x}_n^T \mathbf{w}$

$$n=1$$
- We can use the fact that
$$\frac{d}{dz}\ln(1+\exp(z))=\sigma(z)$$
- Gradient of the log-likelihood

$$\mathbf{g} = \nabla \mathcal{L}(\mathbf{w}) = \sum_{n=1}^{N} \mathbf{x}_n (\sigma(\mathbf{x}_n^T \mathbf{w}) - \mathbf{y}_n)$$

 $= \mathbf{X}^T [\sigma(\mathbf{X}\boldsymbol{w}) - \mathbf{y}]$

The neg. log-likelihood $-\mathcal{L}_{mle}(m{w})$ is convex Hessian of the neg. log-likelihood

- We know that

$$\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t))$$
- Hessian is the derivative of the gradient

$$\mathbf{H}(\mathbf{w}) = \frac{d\mathbf{g}(\mathbf{w})}{d\mathbf{w}^T} = \sum_{n=1}^{N} \frac{d}{d\mathbf{w}^T} \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w})$$
$$= \sum_{n=1}^{N} \mathbf{x}_n \mathbf{x}_n^T \sigma(\mathbf{x}_n^T \mathbf{w}) (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

where **S** is a $N \times N$ diagonal with

 $S_{nn} = \sigma(\mathbf{x}_n^T \mathbf{w})(1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$ - Newton's Method

- Uses second-order information and takes steps in the direction that minimizes a quadratic approximation (Taylor)

$$\mathcal{L}(\mathbf{w}) = \mathcal{L}(\mathbf{w}^{(k)}) + \nabla \mathcal{L}_k^T (\mathbf{w} - \mathbf{w}^{(k)})$$
$$+ (\mathbf{w} - \mathbf{w}^{(k)})^T \mathbf{H}_k (\mathbf{w} - \mathbf{w}^{(k)})$$

and it's minimum is at $\mathbf{w}^{k+1} = \mathbf{w}^{(k)} - \gamma_k \mathbf{H}_k^{-1} \nabla \mathcal{L}_k$

- Complexity: $O((ND^2 + D^3)I)$

- Regularized Logistic Regression

best weight vector ⇒ optimisation does not stop.

$$\underset{\mathbf{w}}{\operatorname{argmin}} - \sum_{n=1}^{N} \ln p(\mathbf{y}_{n} | \mathbf{x}_{n}^{T} \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|^{2}$$

Generalized Linear Model

$$p(\mathbf{y}|\boldsymbol{\eta}) = h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

For proper normalisation (\int p = 1):

$$(\boldsymbol{\eta}) = \ln[\int h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}))]$$

$$\rightarrow \exp(\log(\frac{\mu}{1-\mu})y + \log(1-\mu)))$$

(i) link function g relates η and μ

$$\eta = g(\mu) \Leftrightarrow \mu = g \quad (\eta)$$

$$\eta = \log(\frac{\mu}{1 + e^{\eta}}) \Leftrightarrow \mu = \frac{e^{\eta}}{1 + e^{\eta}}$$

(ii) Note that μ is the mean parameter of y

$$\eta = (\eta_1 = \mu/\sigma^2, \eta_2 = -1/(2\sigma^2))^T$$

 $\mu = -\eta_1/(2\eta_2)$; $\sigma^2 = -1/(2\eta_2)$ First and second derivatives of A(η) are

elated to the mean and the variance
$$\frac{dA(\eta)}{dt} = \mathbb{E}[\phi(\eta)], \quad \frac{d^2A(\eta)}{dt} = \operatorname{Var}[\phi(\eta)]$$

- The generalized maximum likelihood cost to

mize is
$$\min_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \log(p(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}))$$

where $p(\mathbf{y}_n | \mathbf{x}_n^T \mathbf{w})$ is an exponential family

$$\frac{d\mathcal{L}}{d\mathbf{w}} = \mathbf{X}^T [\mathbf{g}^{-1} (\mathbf{X} \mathbf{w}) - \phi(\mathbf{y})]$$

Performs best in low dimensions.

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} \mathbf{y}_n$$

 $f(\mathbf{x}) = modus\{x_n | \mathbf{x}_n \in nbh_k(\mathbf{x})\}\$

Small k → averaging over small area

a) Consider fixed fraction α of points and increase dimension rightarrow need to explore almost whole range in each

each other ⇒ choice of NN becomes

If data is linearly separable, there is no

→ use penalty term.

$$\underset{\mathbf{w}}{\operatorname{argmin}} - \sum_{n=1}^{N} \ln p(\mathbf{y}_{n} | \mathbf{x}_{n}^{T} \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|$$
10 Exponential family distribution &

Exponential family distribution

$$p(\mathbf{y}|\boldsymbol{\eta}) = h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$

$$A(\boldsymbol{\eta}) = \ln[\int_{\mathbb{R}} h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}))]$$

- Bernoulli distribution example

$$oldsymbol{\eta} = \mathbf{g}(oldsymbol{\mu}) \Leftrightarrow oldsymbol{\mu} = \mathbf{g}^{-1}(oldsymbol{\eta})$$

Gaussian distribution example
$$\exp((\frac{\mu}{\sigma^2}, \frac{-1}{2\sigma^2})(y, y^2)^T - \frac{\mu^2}{2\sigma^2} - \frac{1}{2}\ln(2\pi\sigma^2))$$

$$\eta = (\eta_1 = \mu/\sigma^2, \eta_2 = -1/(2\sigma^2))^T$$

related to the mean and the variance

 $-A(\eta)$ is convex

imize is
$$\min_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \log(p(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}))$$

$$\frac{d\mathcal{L}}{d\mathbf{w}} = \mathbf{X}^T [\mathbf{g}^{-1}(\mathbf{X}\mathbf{w}) - \phi($$

11 k-Nearest Neighbor (k-NN)

The k-NN classifier:

Large k → smoothing over large area

Curse of dimensionality:

dimension. b) In high dimensions, points are far from essentially random.

Need radius $r = \sqrt[d]{\left(1 - \frac{1}{N_{\overline{c}}}\right)}$

to have at least one data point in r^d rectangle with $p \geq \frac{1}{2}$.

- NN performance:

- Bayes classifier: $f_*(x) = \mathbb{1}\{\mathbb{P}[y=1|x] > \frac{1}{2}\}$
- $-\mathbb{E}_{S}[L(f_{S})] < 2L(f_{*}) + 4c\sqrt{d}N^{\frac{-1}{1+d}}$

12 Support Vector Machine

- Assume $y_n \in \{-1, 1\}$ and optimise

$$\mathcal{L}(\mathbf{w}) = \min_{\mathbf{w}} \sum_{n=1}^{N} [1 - \mathbf{y}_n x_n^T \mathbf{w}]_+ + \frac{\lambda}{2} ||\mathbf{w}||^2$$

- Can be optimised using subgradient descent.
- Case: Linear separability: We get a seperating hyperplane, no point in the margin and w, s.t margin is maximised (2/||w||).
- This is called hard-margin compared to soft-margin formulation.
- Duality:
- Hard to minimize $g(\mathbf{w})$ so we define $\mathcal{L}(\mathbf{w}) = \max_{\alpha} G(\mathbf{w}, \alpha)$
- we use the property that $[\mathbf{v}_n]_+ = \max(0, \mathbf{v}_n) = \max_{\alpha_n \in [0,1]} \alpha_n \mathbf{v}_n$
- We can rewrite the problem as

$$\min_{\mathbf{w}} \max_{\alpha} \sum_{n=1}^{N} \alpha_n (1 - \mathbf{y}_n \boldsymbol{\phi}_n^T \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|_2^2$$

- This is differentiable, convex in w and concave in α
- Minimax theorem:

 $\min_{\mathbf{w}} \max_{\boldsymbol{\alpha}} G(\mathbf{w}, \boldsymbol{\alpha}) = \max_{\boldsymbol{\alpha}} \min_{\mathbf{w}} G(\mathbf{w}, \boldsymbol{\alpha})$ because G is convex in \mathbf{w} and concave in

- Derivative w.r.t. w:

$$\nabla_{\mathbf{w}}G(\mathbf{w}, \alpha) = -\sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n + \lambda \mathbf{w}$$

$$\mathbf{w}(\boldsymbol{\alpha}) = \frac{1}{\lambda} \sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n = \frac{1}{\lambda} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

Plugging w* back in the dual problem

$$\max_{\boldsymbol{\alpha} \in [0,1]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2\lambda} \boldsymbol{\alpha}^T \mathbf{Y} \mathbf{X} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

- Data only enters as $\mathbf{K} = \mathbf{X}^T \mathbf{X}$.
- Non support vector: Example that lies on the correct side, outside margin $\alpha_n = 0$
- Essen. support vector: Example that lies on the margin $\alpha_n \in (0,1)$
- Bound support vector: Example that lies strictly inside the margin or wrong side
- Use Coordinates ascent to find α. Update one coordinate at a time.

13 Kernel Ridge Regression

- The following is true for ridge regression

$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{y} , (1)$$
$$= \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_N)^{-1} \mathbf{y} = \mathbf{X}^T \boldsymbol{\alpha}^* , (2)$$

- Complexity of computing w: (1)
- $O(D^2N + D^3)$, (2) $O(DN^2 + N^3)$ - Thus we have
- $\mathbf{w}^* = \mathbf{X}^T \boldsymbol{\alpha}^*, \text{ with } \mathbf{w}^* \in \mathbb{R}^D \text{ and } \boldsymbol{\alpha}^* \in \mathbb{R}^N$
- Following representer theorem write:
- $\alpha = \operatorname{argmax} \left(-\frac{1}{2} \alpha^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \alpha + \alpha^T \mathbf{y} \right)$

- K = XX^T is called the kernel matrix or Gram matrix.
- If K is positive definite and symmetric, then it's called a Mercer Kernel.
- $\begin{array}{l} \ \mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j) \\ \ \mathrm{If \ the \ kernel \ is \ Mercer, \ then \ there \ exists \ a} \end{array}$ function $\phi(\mathbf{x})$ s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

- Kernel trick:
 - compute dot-product in \mathbb{R}^m while remaining in \mathbb{R}^n
 - Replace $\langle \mathbf{x}, \mathbf{x}' \rangle$ with $k(\mathbf{x}, \mathbf{x}')$.

Common Kernel

- $-x \in \mathbb{R}, k(\mathbf{x}, \mathbf{x}') = (xx')^2 \Rightarrow \phi(x) = x^2$
- Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Thus we get

$$\mathbf{y} = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

- - $\begin{array}{l} \ \kappa(x,x') = a\kappa_1(x,x') + b\kappa_2(x,x') \\ \ \kappa(x,x') = \kappa_1(x,x')\kappa_2(x,x') \\ \ \kappa(x,x') = \kappa_1(f(x),f(x')) \end{array}$

14 K-means

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N z_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

such that $z_{nk} \in \{0,1\}$ and $\sum_{k=1}^{K} z_{nk} = 1$ – K-means algorithm (Coordinate Descent):

Initialize μ_k , then iterate

1. For all n, compute \mathbf{z}_n given $\boldsymbol{\mu}$ $z_{nk} = \begin{cases} 1 & \text{if } k = \operatorname{argmin}_{j} ||\mathbf{x}_{n} - \boldsymbol{\mu}||_{2}^{2} \\ 0 & \text{otherwise} \end{cases}$

2. For all k, compute μ_k given ${\bf z}$

$$\mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}}$$
 – A good initialization procedure is to choose

- the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{z}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{z_{nk}}$$

$$\begin{split} &-\log p(\mathbf{x}_n|\boldsymbol{\mu},z) = \sum^{N} \sum^{K}_{} \frac{1}{2} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 \mathbf{z}_{nk} + c' \\ &- \text{K-means as a Matrix Factorization} \end{split}$$

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{M}\mathbf{Z}^T||_{\text{Frob}}^2$$

- Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

15 Gaussian Mixture Models

Clusters can be elliptical using a full instead of isotropic covariance matrix.

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{z}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{z_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining z_n to be a random

$$p(z_n = k) = \pi_k \text{ where } \pi_k > 0, \forall k, \sum_{k=1}^K \pi_k = 1$$

- Joint distribution of Gaussian mixture model $p(\mathbf{X}, \mathbf{z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^{N} p(\mathbf{x}_{n} | \mathbf{r}_{n}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{z}_{n} | \boldsymbol{\pi})$

$$= \prod_{n=1}^{N} \prod_{k=1}^{K} [(\mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k))^{z_{nk}}] \prod_{k=1}^{K} [\pi_k]^{z_{nk}}$$

- z_n are called *latent* unobserved variables

- Unknown parameters are $\theta = \{\mu, \Sigma, \pi\}$

- We get the marginal likelihood by marginalizing z_n out from the likelihood

$$p(\mathbf{x}_n|\boldsymbol{\theta}) = \sum_{k=1}^K p(\mathbf{x}_n, z_n = k|\boldsymbol{\theta})$$
$$= \sum_{k=1}^K p(z_n = k|\boldsymbol{\theta})p(\mathbf{x}_n|z_n = k, \boldsymbol{\theta})$$
$$= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to $O(D^2K)$
- To get maximum likelihood estimate of θ , we

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

16 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\theta^{(t)}$ with equality when,

$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$
 2. Maximization step: Update $\boldsymbol{\theta}$

$$\theta^{(t+1)} = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$$
$$\mu_k^{(t+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk}) \mathbf{x}_n}{\sum_{n=1}^{N} q_{nn}^{(t)}}$$

$$\Sigma_k^{(t+1)} = \frac{\sum_{n=1}^{N} q_{kn}^{(t)} (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)})^T}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\pi_k^{(t+1)} = \frac{1}{N} \sum_{n=1}^{N} q_{kn}^{(t)}$$

- If covariance is diagonal → K-means.
- $-q_{nk}^{(t)}=p(z_n=k|x_n,\theta^{(t)})$ posterior of z_n

17 Matrix factorization

- Find X ≈ WZ^T
 - \mathbf{X} is $D \times N$ (e.g movies \times user)
 - **Z** is $N \times K$. **W** is $D \times K$ matrix

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{(d, n) \in \Omega} [x_{dn} - (\mathbf{W} \mathbf{Z}^T)_{dn}]^2$$

$$+\frac{\lambda_w}{2} \|\mathbf{W}\|_{\text{Frob}}^2 + \frac{\lambda_z}{2} \|\mathbf{Z}\|_{\text{Frob}}^2$$

SGD: For one fixed element (d, n) we derive entry (d', k) of **W** (if d = d' oth. 0):

$$\frac{\partial}{\partial w_{d',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] z_{nk}$$

And of **Z** (if n = n' oth. 0):

$$\frac{\partial}{\partial z_{n',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] w_{nk}$$

$$\mathbf{W}^{t+1} = \mathbf{W}^t - \gamma \nabla_w f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$
$$\mathbf{Z}^{t+1} = \mathbf{W}^t - \gamma \nabla_z f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$

We can use coordinate descent algorithm, by first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_K)^{-1} \mathbf{W}^T \mathbf{X}$$
$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_K)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- $\mathbb{O}(DNK + DK^2)$ and $\mathbb{O}(DNK + NK^2)$

18 Text Representation

- word2vec: map every word to a vector $w_i \in \mathbb{R}^K$. K large, that captures its semantics (efficient streaming).
- Topic model: Documents consist of collections of topics
- topic = probability distribution over
- use clustering to pick out respresentative topics

Word representations by matrix factorisation

- typically use log counts from co-occurance matrix
- $-\min_{w,z} L(w,z) =$
- $\frac{1}{2} \sum_{(d,n) \in \Omega} f_{dn} [x_{dn} (WZ^{\top})_{dn}]^2$
- \bar{f}_{dn} : importance of entry
- f_{dn} = 1 is okay, but better $f_{dn} = \min[1, (n_{dn}/N_{max}^{\alpha}], \alpha \in [0, 1], n_{dn}$ are
- this weighting is called GloVe (word2vec variant) and creates spatial analogies (nonstreaming)
- training with SGD or ALS
- Skip-Gram (original word2vec) uses binary classification to distinguish real from fake word pairs. Implicitly based on matrix factorisation.
- FastText: supervised sentence classification.
- Sentence as x_n bag-of-words representation, f is a linear classifier loss. $y_n \in \{0, 1\}$
- $-\min_{W,Z} L(W,Z) =$

$$\sum_{x_m} f(y_m, WZ^{\top}x_m), Wis1 \times K, Z|V| \times K$$

Singular Value Decomposition

- Matrix factorization method $X = USV^T$
- U orthonormal $D \times D$. V orthonormal $N \times N$
- S contains (non-negative) singular values in diagonal in descending order: $D \times N$
- Columns of U and V are the left and right singular vectors (eigenvectors of XX and $\mathbf{X}^{\top}\mathbf{X}$).
- Truncated SVD:

Take the matrix $\mathbf{S}^{(K)}$ with the K first diagonal elements non zero.

$$\mathbf{X} \approx \mathbf{X}_K = \mathbf{U}\mathbf{S}^{(K)}\mathbf{V}^T$$

20 Principal Component Analysis

- dimensionality reduction and decorrelation
- $\begin{aligned} &\|\mathbf{X} \hat{\mathbf{X}}\|_F^2 \geq \|\mathbf{X} \mathbf{U}_k \mathbf{U}_k^\top \mathbf{X}\|_F^2 = \sum_{i > K} s_i^2 \\ &- \text{ If the data has zero mean} \end{aligned}$

$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$

$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Columns of U are called principal
- components and decorrelate X's columns. Not invariant under scalings → normalize X
- Can compute U and S efficiently via $EVD(\mathbf{X}\mathbf{X}^{\top})$ or $EVD(\mathbf{X}^{\top}\mathbf{X})$

$$- x_j^{(l)} = \phi \left(\sum_i w_{i,j}^{(l)} x_i^{(l-1)} + b_j^{(l)} \right)$$

- NN with one hidden layer and sigmoid-like activation function can approximate any sufficiently smooth function on a bounded
- domain in average $(\leq \frac{(2Cr)^2}{n})$ and point-wise - Cost function:
- $\frac{1}{N} \sum_{n=1}^{N} \left(y_n f^{(L+1)} \circ \dots \circ f^{(1)}(\boldsymbol{x}_n^{(0)}) \right)^2$ We can use SGD to minimize the cost.

21.1 Backpropagation Algorithm

- Forward pass: Compute

$$\mathbf{z}^{(l)} = \left(\mathbf{W}^{(l)}\right)^T \mathbf{x}^{(l-1)} + \mathbf{b}^{(l)}$$
 with

- $\mathbf{x}^{(0)} = \mathbf{x}_n \text{ and } \mathbf{x}^{(l)} = \phi(\mathbf{z}^{(l)}).$ Backward pass: Set
- $\delta^{(L+1)} = -2(y_n \boldsymbol{x}^{(L+1)})\phi'(z^{(L+1)})$ (if squared loss). Then compute

$$\begin{split} \delta_j^{(l)} &= \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l+1)}} \frac{\partial z_k^{(l+1)}}{\partial z_j^{(l)}} \\ &= \sum_k \delta_k^{(l+1)} \boldsymbol{W}_{j,k}^{(l+1)} \phi'(z_j^{(l)}) \end{split}$$

$$\frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} = \sum_{k} \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}}$$
$$= \delta^{(l)} \mathbf{x}^{(l-1)}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}} \\ &= \delta_i^{(l)} \cdot 1 = \delta_i^{(l)} \end{split}$$

21.2 Activation Functions

Sigmoid $\phi(x) = \frac{1}{1+e^{-x}}$ Positive, bounded $\phi'(x) \simeq 0$ for large $|x| \Rightarrow$ Learning slow.

Tanh $\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = \phi(2x) - 1/2.$

Balanced, bounded. Learning slow too. ReLU $(x)_{+} = max 0, x Positive, unbounded.$ Derivate = 1 if x > 0, 0 if x < 0

Leaky ReLU $f(x) = \max \alpha x, x$ Remove 0 derivative

Maxout $f(x) = max \mathbf{x}^T \mathbf{w}_1 + b_1, ..., \mathbf{x}^T \mathbf{w}_k + b_k$ (Generalization of ReLU)

21.3 Convolutional NN

Sparse connections, weight sharing: reduce

21.4 Reg, Data Augmentation and

- Regularization term: $\frac{1}{2} \sum_{l=1}^{L+1} \mu^{(l)} ||W^{(l)}||_F^2$ - Weight decay is $\Theta[t](1-\eta\mu)$ in:
- $\Theta[t+1] = \Theta[t] + \eta(\nabla \mathcal{L} + \mu \Theta[t])$
- Data Augm.: e.g. shift or rotation of pics - Dropout: avoid overfit. Drop nodes randomly. (Then average multiple drop-NN

or divide by dropout rate.)

- 22 Bayes Net Graph example: p(x, y, z) = p(x)p(y|x)p(z|x)
- $(y \leftarrow x \rightarrow z)$ **D-Separation** X and Y are D-separated by
- Z if every path from $x \in X$ to $y \in Y$ is blocked by Z. (→ independent)
- Blocked Path contains a variable that - is in Z and is head-to-tail or tail-to-tail
- the node is head-to-head and neither the node nor any of its descendants are in Z.
- Markov Blanket (which blocks node A from the rest of the net) contains:
- parents of A
- children of A
- parents of children of A