ML Cheat Sheet

1 Math Prerequisites

1.1 Derivatives

$$\begin{array}{l} - \ \partial(\mathbf{XY}) = (\partial\mathbf{X})\mathbf{Y} + \mathbf{X}(\partial\mathbf{Y}) \\ - \ \frac{\partial \mathbf{f}(\mathbf{g}(\mathbf{u}(\mathbf{x})))}{\partial\mathbf{x}} = \frac{\partial \mathbf{u}(\mathbf{x})}{\partial\mathbf{x}} \frac{\partial \mathbf{g}(\mathbf{u})}{\partial\mathbf{u}} \frac{\partial \mathbf{f}(\mathbf{g})}{\partial\mathbf{g}} \\ \frac{\partial \mathbf{y}^T \mathbf{a}}{\partial\mathbf{x}^T \mathbf{a}} \frac{\partial \mathbf{a}^T \mathbf{x}}{\partial\mathbf{x}^T \mathbf{x}} \end{array}$$

$$-\frac{\partial \mathbf{x}^T \mathbf{a}}{\partial \mathbf{x}} = \frac{\partial \mathbf{a}^T \mathbf{x}}{\partial \mathbf{x}} = \mathbf{a}$$

$$- \frac{\partial \mathbf{a}^T \mathbf{X} \mathbf{b}}{\partial \mathbf{X}} = \mathbf{a} \mathbf{b}^T$$

$$- \frac{\partial \mathbf{a}^T \mathbf{X}^T \mathbf{b}}{\partial \mathbf{X}} = \mathbf{b} \mathbf{a}^T$$

 $-\frac{\partial \mathbf{X}}{\partial \mathbf{Y}} = \mathbf{J}^{ij}$, J^{ij} is the single entry matrix

$$-\frac{\partial \mathbf{b}^T \mathbf{X}^T \mathbf{X} \mathbf{c}}{\partial \mathbf{X}} = \mathbf{X} \left(\mathbf{b} \mathbf{c}^T + \mathbf{c} \mathbf{b}^T \right)$$
$$-\frac{\partial \mathbf{x}^T \mathbf{B} \mathbf{x}}{\partial \mathbf{x}} = \left(\mathbf{B} + \mathbf{B}^T \right) \mathbf{x}$$

$$-\frac{\partial}{\partial \mathbf{x}}(\mathbf{x} - \mathbf{A}\mathbf{s})^T \mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s}) = 2\mathbf{W}(\mathbf{x} - \mathbf{A}\mathbf{s})$$

$$- \frac{\partial}{\partial \mathbf{X}} \|\mathbf{X}\|_{\mathrm{F}}^2 = \frac{\partial}{\partial \mathbf{X}} \operatorname{Tr} \left(\mathbf{X} \mathbf{X}^H \right) = 2\mathbf{X}$$

1.2 Linear Algebra

- positive definite (pd) if
$$\mathbf{a}^T \mathbf{V} \mathbf{a} > 0$$

$$- (\mathbf{x} - \mathbf{b})^T (\mathbf{x} - \mathbf{b}) = \|\mathbf{x} - \mathbf{b}\|_2^2$$

$$- \|\mathbf{X}\|_{F} = \|\mathbf{X}^{T}\|_{F}$$

1.3 Distributions

Valid distribution p(x)>0, $\forall x$ and $\sum p(x)=1$ Model is identifiable iff $\theta_1=\theta_2\to P_{\theta_1}=P_{\theta_2}$

Gaussian (Not convex):

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$$

$$\mathcal{N}(x|\mu, \Sigma^2) = \frac{\exp(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu))}{\sqrt{(2\pi)^D \det(\Sigma)}}$$

Poisson: P(k events in interval) = e^{-λ λ κ}/_{t-1}

- Bernoulli: $p(y|\mu) = \mu^{y}(1-\mu)^{1-y}$

1.4 Convexity

A function f(x) is convex if

- for any $\mathbf{x}_1, \mathbf{x}_2 \in \mathbf{X}$ and $0 < \lambda < 1$, we have : $f(\lambda \mathbf{x}_1 + (1 - \lambda)\mathbf{x}_2) \le \lambda f(\mathbf{x}_1) + (1 - \lambda)f(\mathbf{x}_2)$

- it is a sum of convex functions

 composition of convex and linear functions - f(x) = g(h(x)), g,h are convex, g increasing

the Hessian H is positive semi-definite

1.5 Others

- Production of independent variables:

$$\operatorname{Var}(xy) = \mathbb{E}(x^2) \mathbb{E}(y^2) - [\mathbb{E}(x)]^2 [\mathbb{E}(y)]^2$$

Covariance matrix of a data vector x

$$\Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \mathbb{E}(\mathbf{x})) (\mathbf{x}_n - \mathbb{E}(\mathbf{x}))^T$$

Multi-class x

$$p(\mathbf{y}|\mathbf{X}, \beta) = \prod_{n=1}^{N} p(\mathbf{y}_{n}|\mathbf{x}_{n}, \beta)$$
$$= \prod_{n=1}^{K} \prod_{n=1}^{N} [p(\mathbf{y}_{n} = k|\mathbf{x}_{n}, \beta)]^{\tilde{y}_{n}k}$$

2 Cost functions

Cost functions are used to learn parameters that explain the data well.

 It is essential to make sure that a global minimum exist → lower bounded

Mean square error (MSE):

$$MSE(\boldsymbol{w}) = \frac{1}{N} \sum_{n=1}^{N} (y_n - f(\mathbf{x}_n))^2$$

- MSE is convex thus it has only one global minumum value

MSE is not good when outliers are present.

Mean Absolute Error (MAE):

$$MAE = \frac{1}{N} \sum_{n=1}^{N} |y_n - f(\mathbf{x}_n)|$$

Huber loss

$$Huber = \begin{cases} \frac{1}{2}z^2 &, |z| \le \delta \\ \delta|z| - \frac{1}{2}\delta^2 &, |z| > \delta \end{cases}$$

 Huber loss is convex, differentiable, and also robust to outliers but hard to set δ .

Tukey's bisquare loss

$$L(z) = \begin{cases} z(\delta^2 - z^2)^2 &, |z| < \delta \\ 0 &, |z| \ge \delta \end{cases}$$

Non-convex, non-diff., but robust to outliers. Hinge loss:

 $[1 - y_n f(\mathbf{x}_n)]_+ = \max(0, 1 - y_n f(\mathbf{x}_n))$ Logistic loss: $\log(1 - \exp(y_n f(\mathbf{x}_n)))$

3 Regression

- Model that assume linear relationship

$$\mathbf{y}_n \approx f(\mathbf{x}_n) := \mathbf{w}_0 + \mathbf{w}_1 \mathbf{x}_{n1} + \dots = \mathbf{w}_0 + \mathbf{x}_n^T \mathbf{w}$$

 $\approx \tilde{\mathbf{x}}_n^T \mathbf{w}$, where \tilde{x} contains offset comp.

- Prediction: predict the ouput for a new input vector.

- Interpretation: understand the effect of inputs on output.

4 Optimization

4.1 Grid search

Compute the cost over a grid of M points to find the minimum. Exponential Complexity. Hard to find a good range of values

4.2 GD - Gradient Descent (Batch)

- GD uses only first-order information and takes steps in the opposite direction of the gradient

Given cost function $\mathcal{L}(\mathbf{w})$ we want to find \mathbf{w} $\mathbf{w} = \arg\min \mathcal{L}(\mathbf{w})$

 Take steps in the opposite direction of the gradient

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})$$

With γ too big, method might diverge. With γ too small, convergence is slow.

4.3 SGD - Stochastic Gradient Descent

In ML, most cost functions are formulated as a sum over the training examples:

$$\mathcal{L}(\mathbf{w}) = \frac{1}{N} \sum_{n=1}^{N} \mathcal{L}_n(\mathbf{w})$$

⇒ SGD update rule (only n-th training exam.):

$$\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Idea: Cheap but unbiased estimate of grad.

$$\mathbb{E}[\nabla \mathcal{L}_n(\mathbf{w})] = \nabla(\mathbf{w})$$

4.4 Mini-batch SGD

Update direction $(B \subset [N])$:

$$g^{(t)} := \frac{1}{|B|} \sum_{n \in B} \nabla \mathcal{L}_n(\mathbf{w}^{(t)})$$

Update rule: $\mathbf{w}^{(t+1)} \leftarrow \mathbf{w}^{(t)} - \gamma \mathbf{g}^{(t)}$

4.5 Gradients for MSE

We define the error vector e:

$$\mathbf{e} := \mathbf{y} - \mathbf{A}\mathbf{w}$$

$$\mathcal{L}(\mathbf{w}) = \frac{1}{2N} \sum_{n=1}^{N} (\mathbf{y}_n - \tilde{\mathbf{x}}_n^T \mathbf{w})^2 = \frac{1}{2N} \mathbf{e}^T \mathbf{e}$$

then the gradient is given by

$$\nabla \mathcal{L}(\mathbf{w}) = -\frac{1}{N} \mathbf{X}^T \mathbf{e}$$

Optimality conditions:

1. necessary: gradient equal zero: $\frac{d\mathcal{L}(\mathbf{w}^*)}{d\mathbf{w}} = 0$

2. sufficient: Hessian matrix is positive

definite: $\mathbf{H}(\mathbf{w}^*) = \frac{d^2 \mathcal{L}(\mathbf{w}^*)}{dt}$ Very sensitive to illconditioning ⇒ always

Complexity: O(NDI) with I the number of

4.6 Subgradients (Non-Smooth OPT)

A vector $\mathbf{g} \in \mathbb{R}^D$ s.t.

normalize features.

$$\mathcal{L}(\mathbf{u}) \geq \mathcal{L}(\mathbf{w}) + \mathbf{g}^T(\mathbf{u} - \mathbf{w}) \quad \forall \mathbf{u} \in \mathbb{R}^D$$

is the subgradient to \mathcal{L} at \mathbf{w} . If \mathcal{L} is differentiable at \mathbf{w} , we have $\mathbf{g} = \nabla \mathcal{L}(\mathbf{w})$

4.7 Constrained Optimization

Find solution min $\mathcal{L}(\mathbf{w})$ s.t. $\mathbf{w} \in \mathcal{C}$ Add proj. onto C after each step:

$$P_{\mathcal{C}}(\mathbf{w}') = \arg\min|\mathbf{v} - \mathbf{w}'|, \ \mathbf{v} \in \mathcal{C}$$
$$\mathbf{w}^{(t+1)} = P_{\mathcal{C}}[\mathbf{w}^{(t)} - \gamma \nabla \mathcal{L}(\mathbf{w}^{(t)})]$$

Use penalty functions

- $\min \mathcal{L}(\mathbf{w}) + I_{\mathcal{C}}, I_{\mathcal{C}} = 0 \text{ if } \mathbf{w} \in \mathcal{C}, \text{ ow } + \infty$ $-\min \mathcal{L}(\mathbf{w}) + \lambda |\mathbf{A}\mathbf{w} - \mathbf{b}|$

Stopping criteria when L(w) close to 0

5 Least Squares

- Use the first optimality conditions:

$$\nabla L(\mathbf{w}^*) = 0 \Rightarrow \mathbf{X}^T \mathbf{e} = \mathbf{X}^T (\mathbf{y} - \mathbf{X} \mathbf{w}) = 0$$

When $\mathbf{X}^T\mathbf{X}$ is invertible, we have the closed-form expression

1-form expression
$$\mathbf{w}^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

 thus we can predict values for a new x_m $\mathbf{y}_m := \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} \mathbf{w}^* = \mathbf{x}_{\mathbf{m}}^{\mathbf{T}} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$

– The **Gram matrix** $\mathbf{X}^T \mathbf{X}$ is pd and is also invertible iff X has full column rank.

Complexity: $O(ND^2 + D^3) \equiv O(ND^2)$ **X** can be rank deficient when D > N or when the comlumns $\bar{\mathbf{x}}_d$ are nearly collinear. \Rightarrow matrix is ill-conditioned.

6 Maximum Likelihood (MLE)

Let define our mistakes ε_n ~ N(0, σ²).

$$\to \mathbf{y}_n = \mathbf{x}_n^T \mathbf{w} + \epsilon_n$$

- Another way of expressing this:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} p(\mathbf{y}_n | \mathbf{x}_n, \mathbf{w})$$
$$= \prod_{n=1}^{N} \mathcal{N}(\mathbf{y}_n | \mathbf{x}_n^T \mathbf{w}, \sigma^2)$$

which defines the likelihood of observating y given \mathbf{X} and \mathbf{w}

Define cost with log-likelihood $\mathcal{L}_{MLE}(\mathbf{w}) = \log p(\mathbf{y}|\mathbf{X}, \mathbf{w})$

$$= -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (\mathbf{y}_n - \mathbf{x}_n^T \mathbf{w})^2 + cnst$$

Maximum likelihood estimator (MLE) gives another way to design cost functions

 $\operatorname{argmin} \mathcal{L}_{MSE}(\mathbf{w}) = \operatorname{argmax} \mathcal{L}_{MLE}(\mathbf{w})$ - MLE can also be interpreted as finding the model under which the observed data is most

likely to have been generated from. - $\mathbf{w}_{\mathrm{MLE}} \to \mathbf{w}_{\mathrm{true}}$ for large amount of data

7 Ridge Regression (RR)

Linear models usually overfit. We can penalize them with a regularization term $\min \mathcal{L}(\mathbf{w}) + \Omega(\mathbf{w})$

- L_2 -Reg. (Ridge): $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_2^2$ $-\rightarrow$ small values of \mathbf{w}_i , not sparse

 $-\rightarrow \mathbf{w}^{\star} = (\mathbf{X}^T \mathbf{X} + \lambda' \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y} \text{ with } \lambda' = 2N\lambda$ - \rightarrow No ill cond., $(\mathbf{X}^T\mathbf{X} + \lambda'\mathbf{I})^{-1}$ exists

- L_1 -Reg. (Lasso): $\Omega(\mathbf{w}) = \lambda ||\mathbf{w}||_1$ $- \rightarrow$ large values of \mathbf{w}_i , sparse

 Maximum-a-posteriori (MAP)

$$\begin{split} p(\mathbf{y}|\mathbf{X}\mathbf{w}) &= \prod^{N} \mathcal{N}(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}, \sigma_{n}^{2}) \\ p(\mathbf{w}) &= \mathcal{N}(\mathbf{w}|0, \sigma_{0}^{2}\mathbf{I}_{D}) \\ \text{then} &\rightarrow \mathbf{w}^{\star} = \operatorname{argmax} p(\mathbf{y}|\mathbf{X}\mathbf{w}) \cdot p(\mathbf{w}) \end{split}$$

$$\mathbf{w}^{\star} = \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{n=1}^{N} \frac{1}{2\sigma^{2}} (\mathbf{y}_{n} - \mathbf{x}^{T} \mathbf{w})^{2} + \frac{1}{2\sigma^{2}} \|\mathbf{w}\|^{2}$$

8 Bias-Variance decomposition

- The expected test error can be expressed as the sum of two terms

- Squared bias: The average shift of the predictions

- Variance: measure how data points vary around their average. expected loss = $(bias)^2$ + variance + noise

Model bias and estimation bias are important

 RR increases estimation bias and reduces var Model more complex increases test error

- Small $\lambda \to \text{low bias but large variance}$

- Large $\lambda \rightarrow$ large bias but low variance Simple → large bias but low variance

 Complex → low bias but large variance $err = \sigma^2 + \mathbb{E}[f_{lse} - \mathbb{E}[f_{lse}]]^2 + [f_{true} - \mathbb{E}[f_{lse}]]^2$

9 Logistic Regression

- Classification relates input variables x to discrete output variable y Binary classifier: we use u = 0 for C₁ and

y = 1 for C_2 .

- Can use least-squares to predict
$$\hat{y}_*$$

$$\hat{y} = \begin{cases} \mathbf{C}_1 & \hat{y}_* < 0.5 \\ \mathbf{C}_2 & \hat{y}_* \ge 0.5 \end{cases}$$

- Logistic function

$$\sigma(x) = \frac{\exp(x)}{1 + \exp(x)}$$
$$p(\mathbf{y}_n = \mathbf{C}_1 | \mathbf{x}_n) = \sigma(\mathbf{x}^T \mathbf{w})$$

$$p(\mathbf{y}_n = \mathbf{C}_2 | \mathbf{x}_n) = 1 - \sigma(\mathbf{x}^T \mathbf{w})$$

- The probabilistic model:

$$p(\mathbf{y}|\mathbf{X}, \mathbf{w}) = \prod_{n=1}^{N} \sigma(\mathbf{x}_{n}^{T} \mathbf{w})^{\mathbf{y}_{n}} (1 - \sigma(\mathbf{x}_{n}^{T} \mathbf{w}))^{1 - \mathbf{y}_{n}}$$

- The log-likelihood (w.r.t. MLE):

$$\mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \mathbf{y}_n \ln \sigma(\mathbf{x}_n^T \mathbf{w}) + (1 - \mathbf{y}_n) \ln (1 - \sigma(\mathbf{x}_n^T \mathbf{w})) \text{ The generalized maximum likelihood cost to minimize is}$$

 $= \sum_{n=1}^{N} \ln[1 + \exp(\mathbf{x}_n^T \mathbf{w})] - \mathbf{y}_n \mathbf{x}_n^T \mathbf{w}$

n=1- We can use the fact that $\frac{d}{dx}\log(1+\exp(x))=\sigma(x)$ - Gradient of the log-likelihood

$$\mathbf{g} = \nabla \mathcal{L}(\mathbf{w}) = \sum_{n=1}^{N} \mathbf{x}_{n} (\sigma(\mathbf{x}_{n}^{T} \mathbf{w}) - \mathbf{y}_{n})$$

 The negative of the log-likelihood -L_{mle}(w) is convex

Hessian of the log-likelihood

- We know that $\frac{d\sigma(t)}{dt} = \sigma(t)(1 - \sigma(t))$ - Hessian is the derivative of the gradient

$$\mathbf{H}(\mathbf{w}) = -\frac{d\mathbf{g}(\mathbf{w})}{d\mathbf{w}^T} = \sum_{n=1}^{N} \frac{d}{d\mathbf{w}^T} \mathbf{x}_n \sigma(\mathbf{x}_n^T \mathbf{w})$$
$$= \sum_{n=1}^{N} \mathbf{x}_n \mathbf{x}_n^T \sigma(\mathbf{x}_n^T \mathbf{w}) (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

where S is a $N \times N$ diagonal matrix with diagonals

$$S_{nn} = \sigma(\mathbf{x}_n^T \mathbf{w}) (1 - \sigma(\mathbf{x}_n^T \mathbf{w}))$$

 The negative of the log-likelihood is not strictly convex.

Newton's Method

- Uses second-order information and takes steps in the direction that minimizes a quadratic approximation (Taylor)

$$\mathcal{L}(\mathbf{w}) = \mathcal{L}(\mathbf{w}^{(k)}) + \nabla \mathcal{L}_k^T (\mathbf{w} - \mathbf{w}^{(k)})$$

$$+(\mathbf{w} - \mathbf{w}^{(k)})^T \mathbf{H}_k (\mathbf{w} - \mathbf{w}^{(k)})$$

and it's minimum is at $\mathbf{w}^{k+1} = \mathbf{w}^{(k)} - \gamma_k \mathbf{H}_k^{-1} \nabla \mathcal{L}_k$ - Complexity: $O((ND^2 + D^3)I)$

$$\begin{aligned} & \textbf{Regularized Logistic Regression} \\ & \operatorname{argmin} - \sum_{\mathbf{w}}^{N} \ln p(\mathbf{y}_{n}|\mathbf{x}_{n}^{T}\mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|^{2} \end{aligned}$$

10 Exponential family distribution & Generalized Linear Model

Exponential family distribution

$$p(\mathbf{y}|\boldsymbol{\eta}) = h(y) \exp(\boldsymbol{\eta}^T \boldsymbol{\phi}(\mathbf{y}) - A(\boldsymbol{\eta}))$$
- **Bernoulli** distribution example
$$\rightarrow \exp(\log(\frac{\mu}{1-\mu})y + \log(1-\mu)))$$

(i) there is a relationship between η and μ through the link function

$$\eta = \log(\frac{\mu}{1-\mu}) \leftrightarrow \mu = \frac{e^{\eta}}{1+e^{\eta}}$$

(ii) Note that μ is the mean parameter of y

(iii) Relationship between the mean μ and η is defined using a link function a

using a link function
$$g$$

 $\eta = \mathbf{g}(\mu) \Leftrightarrow \mu = \mathbf{g}^{-1}(\eta)$

Gaussian distribution example

$$\exp((\frac{\mu}{\sigma^2}, \frac{-1}{2\sigma^2})(y, y^2)^T - \frac{\mu^2}{2\sigma^2} - \frac{1}{2}\ln(2\pi\sigma^2))$$
(i) link function

$$\eta = (\eta_1 = \mu/\sigma^2, \eta_2 = -1/(2\sigma^2))^T$$

 $\mu = -\eta_1/(2\eta_2)$; $\sigma^2 = -1/(2\eta_2)$ First and second derivatives of A(n) are

related to the mean and the variance
$$\frac{dA(\eta)}{d\eta} = \mathbb{E}[\phi(\eta)], \ \, \frac{d^2A(\eta)}{d\eta^2} = \mathrm{Var}[\phi(\eta)]$$

$$\min_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = -\sum_{n=1}^{N} \log(p(\mathbf{y}_n | \mathbf{x}_n^T \mathbf{w}))$$

where $p(\mathbf{y}_n|\mathbf{x}_n^T\mathbf{w})$ is an exponential family distribution

We obtain the solution

$$\frac{d\mathcal{L}}{d\mathbf{w}} = \mathbf{X}^T [\mathbf{g}^{-1}(\mathbf{X}\mathbf{w}) - \phi(\mathbf{y})]$$
11 k-Nearest Neighbor (k-NN)

The k-NN prediction for x is

$$f(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_n \in nbh_k(\mathbf{x})} \mathbf{y}_n$$

where $nbh_k(\mathbf{x})$ is the neightborhood of \mathbf{x} defined by the k closest points \mathbf{x}_n . Curse of dimensionality: Generalizing

correctly becomes exponentially harder as the

dimensionality grows.

Gathering more inputs variables may be bad

12 Support Vector Machine

- Combination of the kernel trick plus a modified loss function (Hinge loss)
- Solution to the dual problem is sparse and non-zero entries will be our support vectors
- Kernelised feature vector where μ_k are centroids

$$\phi(\mathbf{x}) = [k(\mathbf{x}, \boldsymbol{\mu}_1), ..., k(\mathbf{x}, \boldsymbol{\mu}_K)]$$

- In practice we'll take a subset of data points to be prototype -> sparse vector machine.
- Assume $y_n \in \{-1, 1\}$
- SVM optimizes the following cost

$$\mathcal{L}(\mathbf{w}) = \min_{\mathbf{w}} \sum_{n=1}^{N} [1 - \mathbf{y}_n \tilde{\boldsymbol{\phi}}_n^T \mathbf{w}]_+ + \frac{\lambda}{2} \|\mathbf{w}\|^2$$

- Minimum doesn't change with a rescaling of
- choose the hyperplane so that the distance from it to the nearest data point on each side is maximized

Duality:

- Hard to minimize $g(\mathbf{w})$ so we define $\mathcal{L}(\mathbf{w}) = \max_{\boldsymbol{\alpha}} G(\mathbf{w}, \boldsymbol{\alpha})$
- we use the property that

$$[\mathbf{v}_n]_+ = \max(0, \mathbf{v}_n) = \max_{\alpha_n \in [0, 1]} \alpha_n \mathbf{v}_n$$
- We can rewrite the problem as

We can rewrite the problem as
$$\min_{\mathbf{w}} \max_{\alpha} \sum_{n=1}^{N} \alpha_n (1 - \mathbf{y}_n \boldsymbol{\phi}_n^T \mathbf{w}) + \frac{\lambda}{2} \|\mathbf{w}\|_2^2$$

- This is differentiable, convex in \boldsymbol{w} and concave in a
- Minimax theorem:

 $\min_{\mathbf{w}} \max_{\mathbf{\alpha}} G(\mathbf{w}, \mathbf{\alpha}) = \max_{\mathbf{\alpha}} \min_{\mathbf{w}} G(\mathbf{w}, \mathbf{\alpha})$ because G is convex in \mathbf{w} and concave in

- Derivative w.r.t. w:

- Derivative w.r.t. **w**:
$$\nabla_{\mathbf{w}} G(\mathbf{w}, \boldsymbol{\alpha}) = -\sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n + \lambda \mathbf{w}$$
- Equating this to 0, we get:

$$\mathbf{w}(\boldsymbol{\alpha}) = \frac{1}{\lambda} \sum_{n=1}^{N} \alpha_n \mathbf{y}_n \mathbf{x}_n = \frac{1}{\lambda} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

$$\mathbf{Y} := \operatorname{diag}(\cdot)$$

Plugging w* back in the dual problem

Plugging
$$\mathbf{w}^*$$
 back in the dual problem
$$\max_{\boldsymbol{\alpha} \in [0,1]^N} \boldsymbol{\alpha}^T \mathbf{1} - \frac{1}{2\lambda} \boldsymbol{\alpha}^T \mathbf{Y} \mathbf{X} \mathbf{X}^T \mathbf{Y} \boldsymbol{\alpha}$$

- This is a differentiable least-squares problem. Optimization is easy using Sequential Minimal Optimization. It is also naturally kernelized with $\mathbf{K} = \mathbf{X}^T \mathbf{X}$
- The solution α is sparse and is non-zero only for the training examples that are instrumental in determining the decision
- α is the slope of lines that are lower bound to Hinge loss
- Non support vector: Example that lies on the correct side, outside margin $\alpha_n = 0$
- Essen. support vector: Example that lies on the margin $\alpha_n \in (0,1)$
- Bound support vector: Example that lies strictly inside the margin or wrong side
- Use Coordinates Descent to find α . Update one coordinate (argmin) at the time and others constant

13 Kernel Ridge Regression

- The following is true for ridge regression $\mathbf{w}^* = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_D)^{-1} \mathbf{X}^T \mathbf{v} . (1)$

$$= \mathbf{X}^{T} (\mathbf{X} \mathbf{X}^{T} + \lambda \mathbf{I}_{D})^{-1} \mathbf{y} = \mathbf{X}^{T} \boldsymbol{\alpha}^{*}, (2)$$

$$= \mathbf{X}^{T} (\mathbf{X} \mathbf{X}^{T} + \lambda \mathbf{I}_{N})^{-1} \mathbf{y} = \mathbf{X}^{T} \boldsymbol{\alpha}^{*}, (2)$$

Complexity of computing w: (1)

$$O(D^2N + D^3), (2) O(DN^2 + N^3)$$

Thus we have

 $\mathbf{w}^* = \mathbf{X}^T \boldsymbol{\alpha}^*$, with $\mathbf{w}^* \in \mathbb{R}^D$ and $\boldsymbol{\alpha}^* \in \mathbb{R}^N$

 The representer theorem allows us to write an equivalent optimization problem in terms

$$\alpha = \operatorname*{argmax}_{\boldsymbol{\alpha}} \left(-\frac{1}{2} \boldsymbol{\alpha}^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_N) \boldsymbol{\alpha} + \boldsymbol{\alpha}^T \mathbf{y} \right)$$

- $\mathbf{K} = \mathbf{X}\mathbf{X}^T$ is called the **kernel matrix** o Gram matrix.
- If K is positive definite, then it's called a Mercer Kernel.
- $\mathbf{K}_{i,j} = k(\mathbf{x}_i, \mathbf{x}_j)$ If the kernel is Mercer, then there exists a function $\phi(\mathbf{x})$ s.t.

$$k(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$$

Kernel trick:

- compute dot-product in \mathbb{R}^m while remaining in \mathbb{R}^n
- Replace (x, x') with k(x, x').

Common Kernel

- Polynomial Kernel: $(\gamma \langle \mathbf{x}_i, \mathbf{x}_j \rangle + r)^d$ Radial Basis function kernel (RBF)

$$k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2}(\mathbf{x} - \mathbf{x}')^T(\mathbf{x} - \mathbf{x}'))$$

- Sigmoid Kernel: $tanh(\langle \mathbf{x}_i, \mathbf{x}_i \rangle + r)$
- Properties of kernels to ensure the existance of a corresponding ϕ :
- symmetric: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$ - positive semi-definite.

$$\mathbf{y} = \mathbf{w}^T \mathbf{x} = \sum_{i=1}^K \alpha_i \mathbf{x}_i^T \mathbf{x} = \sum_{i=1}^K \alpha_i k(\mathbf{x}, \mathbf{x}_i)$$

14 K-means

- Unsupervised learning: Represent particular input patterns in a way that reflects the statistical structure of the overall collections of input partterns.
- Cluster are groups of points whose inter-point distances are small compared to the distances outside the cluster.

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = \sum_{k=1}^K \sum_{n=1}^N z_{nk} ||\mathbf{x}_n - \boldsymbol{\mu}_k||_2^2$$

- such that $z_{nk} \in \{0,1\}$ and $\sum_{k=1}^{K} z_{nk} = 1$ K-means algorithm (Coordinate Descent): Initialize μ_k , then iterate
 - 1. For all n, compute \mathbf{z}_n given $\boldsymbol{\mu}$

$$z_{nk} = \begin{cases} 1 & \text{if } k = \operatorname{argmin}_{j} ||\mathbf{x}_{n} - \boldsymbol{\mu}||_{2}^{2} \\ 0 & \text{otherwise} \end{cases}$$

2. For all k, compute μ_k given ${\bf z}$

$$u_k = \frac{\sum_{n=1}^{N} z_{nk} \mathbf{x}_n}{\sum_{n=1}^{N} z_{nk}}$$

- $\mu_k = \frac{\sum_{n=1}^N z_{nk} \mathbf{x}_n}{\sum_{n=1}^N z_{nk}}$ A good initialization procedure is to choose the prototypes to be equal to a random subset of K data points.
- Probabilistic model

$$p(\mathbf{z}, \boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \mathbf{I}) \right]^{z_{nk}}$$

$$\begin{aligned} -\log p(\mathbf{x}_n|\boldsymbol{\mu},z) &= \sum_{}^{N} \sum_{}^{K} \frac{1}{2} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 \mathbf{z}_{nk} + c' \\ &- \text{K-means as a Matrix Factorization} \\ &\min_{\mathbf{z},\boldsymbol{\mu}} \mathcal{L}(\mathbf{z},\boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{M} \mathbf{Z}^T||_{\text{Frob}}^2 \end{aligned}$$

$$\min_{\mathbf{z}, \boldsymbol{\mu}} \mathcal{L}(\mathbf{z}, \boldsymbol{\mu}) = ||\mathbf{X} - \mathbf{M}\mathbf{Z}^T||_{\text{Fro}}^2$$

Computation can be heavy, each example can belong to only on cluster and clusters have to be spherical.

15 Gaussian Mixture Models

- Clusters can be elliptical using a full covariance matrix instead of isotropic

$$p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{z}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\mathcal{N}(\mathbf{x}_{n}|\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right]^{z_{nk}}$$

Soft-clustering: Points can belong to several cluster by defining z_n to be a random

$$p(z_n = k) = \pi_k$$
 where $\pi_k > 0, \forall k, \sum_{k=1}^K \pi_k = 1$

Joint distribution of Gaussian mixture model

$$p(\mathbf{X}, \mathbf{z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^{N} p(\mathbf{x}_n | \mathbf{r}_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) p(\mathbf{z}_n | \boldsymbol{\pi})$$

$$= \prod_{n=1}^{N} \prod_{k=1}^{K} \left[\left(\mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}) \right)^{z_{nk}} \right] \prod_{k=1}^{K} \left[\pi_{k} \right]^{z_{nk}}$$
- z_{n} are called $latent$ unobserved variables

- Unknown parameters are given by
- $\theta = \{\mu, \Sigma, \pi\}$
- We get the marginal likelihood by marginalizing z_n out from the likelihood

$$\begin{split} p(\mathbf{x}_n|\boldsymbol{\theta}) &= \sum_{k=1}^K p(\mathbf{x}_n, z_n = k|\boldsymbol{\theta}) \\ &= \sum_{k=1}^K p(z_n = k|\boldsymbol{\theta}) p(\mathbf{x}_n|z_n = k, \boldsymbol{\theta}) \\ &= \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \end{split}$$

- Without a latent variable model, number of parameters grow at rate O(N)
- After marginalization, the growth is reduced to $O(D^2K)$
- To get maximum likelihood estimate of θ , we

$$\max_{\boldsymbol{\theta}} \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_{k} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

16 Expectation Maximization Algorithm

- [ALGORITHM] Start with $\theta^{(1)}$ and iterate
- 1. Expectation step: Compute a lower bound to the cost such that it is tight at the previous $\boldsymbol{\theta}^{(t)}$ with equality when,

$$q_{kn} = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}$$
 2. Maximization step: Update $\boldsymbol{\theta}$

$$\boldsymbol{\theta}^{(t+1)} = \operatorname*{argmax}_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)})$$

$$\boldsymbol{\mu}_{k}^{(t+1)} = \frac{\sum_{n=1}^{N} \gamma^{(i)}(r_{nk}) \mathbf{x}_{n}}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\Sigma_k^{(t+1)} = \frac{\sum_{n=1}^{N} q_{kn}^{(t)} (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{(t+1)})^T}{\sum_{n=1}^{N} q_{kn}^{(t)}}$$

$$\pi_k^{(t+1)} = \frac{1}{N} \sum_{n=1}^{N} q_{kn}^{(t)}$$

If covariance is diagonal → K-means.

17 Matrix factorization

- We have D movies and N users
- **X** is a matrix $D \times N$ with x_{dn} the rating of n'th user for d'th movie.
- We project data vectors \mathbf{x}_n to a smaller dimension $\mathbf{z}_n \in \mathbb{R}^M$
- We have now 2 latent variables:
- \mathbf{Z} a $N \times K$ matrix that gives features for the users
- \mathbf{W} a $D \times K$ matrix that gives features for the movies

 $x_{dn} \approx \mathbf{w}_d^T \mathbf{z}_n$ We can add a regularizer and minimize the

$$\mathcal{L}(\mathbf{W}, \mathbf{Z}) = \frac{1}{2} \sum_{(d,n) \in \Omega} [x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}]^2$$

$$+ \frac{\lambda_w}{2} \|\mathbf{W}\|_{\operatorname{Frob}}^2 + \frac{\lambda_z}{2} \|\mathbf{Z}\|_{\operatorname{Frob}}^2$$

SGD: For one fixed element (d, n) we derive entry (d', k) of **W** (if d = d' oth. 0):

$$\frac{\partial}{\partial w_{d',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] z_{nk}$$
And of **Z** (if $n = n'$ oth. 0):

$$\frac{\partial}{\partial z_{n',k}} f_{d,n}(\mathbf{W}, \mathbf{Z}) = -[x_{dn} - (\mathbf{W}\mathbf{Z}^T)_{dn}] w_{nk}$$

$$\mathbf{W}^{t+1} = \mathbf{W}^t - \gamma \nabla_w f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$

$$\mathbf{Z}^{t+1} = \mathbf{W}^t - \gamma \nabla_z f_{d,n}(\mathbf{W}^t, \mathbf{Z}^t)$$
 — We can use coordinate descent algorithm, by

first minimizing w.r.t. Z given W and then minimizing W given Z. This is called Alternating least-squares (ALS):

$$\mathbf{Z}^T \leftarrow (\mathbf{W}^T \mathbf{W} + \lambda_z \mathbf{I}_K)^{-1} \mathbf{W}^T \mathbf{X}$$

$$\mathbf{W}^T \leftarrow (\mathbf{Z}^T \mathbf{Z} + \lambda_w \mathbf{I}_K)^{-1} \mathbf{Z}^T \mathbf{X}^T$$

- Complexity: $O(DNK^2 + NK^3) \rightarrow O(DNK^2)$

18 Singular Value Decomposition

Matrix factorization method

$$\mathbf{X} = \mathbf{U}\mathbf{S}\mathbf{V}^T$$

- \mathbf{U} is a unitary $D \times D$ matrix
- V is a unitary N × N matrix
- S is a non-negative diagonal matrix of size $D \times N$ which are called **singular values** appearing in a descending order.
- Columns of U and V are the left and right singular vectors respectively.
- Assuming D < N we have

$$\mathbf{X} = \sum_{d=1}^{D} s_d \mathbf{u}_d \mathbf{v}_d^T$$

This tells you about the spectrum of X where higher singular vectors contain the low-frequency information and lower singular values contain the high-frequency information.

Truncated SVD: Take the matrix $\mathbf{S}^{(K)}$ with the K first diagonal elements non zero. Then, rank-Kapprox:

$$\mathbf{X} \approx \mathbf{X}_K = \mathbf{U}\mathbf{S}^{(K)}\mathbf{V}^T$$

19 Principal Component Analysis

- PCA is a dimensionality reduction method and a method $\underline{\mathbf{t}}$ o decorrelate the data $\mathbf{X} \approx \tilde{\mathbf{X}} = \mathbf{W}\mathbf{Z}^T$ such that columns of \mathbf{W} are
- orthogonal. If the data is zero mean

$$\Sigma = \frac{1}{N} \mathbf{X} \mathbf{X}^T \Rightarrow \mathbf{X} \mathbf{X}^T = \mathbf{U} \mathbf{S}^2 \mathbf{U}^T$$

$$\Rightarrow \mathbf{U}^T \mathbf{X} \mathbf{X}^T \mathbf{U} = \mathbf{U}^T \mathbf{U} \mathbf{S}^2 \mathbf{U}^T \mathbf{U} = \mathbf{S}^2$$

- Thus the columns of matrix U are called the principal components and they decorrelate the covariance matrix
- Using SVD, we can compute the matrices in the following way

wing way
$$\mathbf{W} = \mathbf{U}\mathbf{S}_D^{1/2}, \mathbf{Z}^T = \mathbf{S}^{1/2}\mathbf{V}^T$$

- Not invariant under scalings of the feature = arbitrariness, → normalize X

20 Neural Net

 Basic structure: One input layer of size D, L hidden layers of size K, and one output layer. (feedforward network).

$$\begin{aligned} x_j^{(l)} &= \phi\left(\sum_i w_{i,j}^{(l)} x_i^{(l-1)} + b_j^{(l)}\right). \\ &= \text{NN can represent the Rienmann sum with} \\ &\text{only two layers} \Rightarrow \text{It's powerful!} \end{aligned}$$

 $\frac{1}{N} \sum_{n=1}^{N} \left(y_n - f^{(L+1)} \circ \dots \circ f^{(1)}(\boldsymbol{x}_n^{(0)}) \right)^2$ We can use SGD to minimize the cost

20.1 Backpropagation Algorithm

Forward pass: Compute

Cost function:

$$\mathbf{z}^{(l)} = \left(\mathbf{W}^{(l)}\right)^T \mathbf{x}^{(l-1)} + \mathbf{b}^{(l)}$$
 with $\mathbf{x}^{(0)} = \mathbf{x}_n$ and $\mathbf{x}^{(l)} = \phi(\mathbf{z}^{(l)})$.

Backward pass: Set

 $\delta^{(L+1)} = -2(y_n - \boldsymbol{x}^{(L+1)})\phi'(z^{(L+1)})$ (if squared loss). Then compute

$$\begin{split} \boldsymbol{\delta}_{j}^{(l)} &= \frac{\partial \mathcal{L}_{n}}{\partial \boldsymbol{z}_{j}^{(l)}} = \sum_{k} \frac{\partial \mathcal{L}_{n}}{\partial \boldsymbol{z}_{k}^{(l+1)}} \frac{\partial \boldsymbol{z}_{k}^{(l+1)}}{\partial \boldsymbol{z}_{j}^{(l)}} \\ &= \sum_{l} \boldsymbol{\delta}_{k}^{(l+1)} \boldsymbol{W}_{j,k}^{(l+1)} \boldsymbol{\phi}'(\boldsymbol{z}_{j}^{(l)}) \end{split}$$

$$\frac{\partial \mathcal{L}_n}{\partial w_{i,j}^{(l)}} = \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}} = \frac{\delta^{(l)}}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial w_{i,j}^{(l)}}$$

$$\begin{split} \frac{\partial \mathcal{L}_n}{\partial b_j^{(l)}} &= \sum_k \frac{\partial \mathcal{L}_n}{\partial z_k^{(l)}} \frac{\partial z_k^{(l)}}{\partial b_j^{(l)}} = \frac{\partial \mathcal{L}_n}{\partial z_j^{(l)}} \frac{\partial z_j^{(l)}}{\partial b_j^{(l)}} \\ &= \delta_j^{(l)} \cdot 1 = \delta_j^{(l)} \end{split}$$

20.2 Activation Functions

Sigmoid $\phi(x) = \frac{1}{1+e^{-x}}$ Positive, bounded

 $\phi'(x) \simeq 0$ for large $|x| \Rightarrow$ Learning slow. Tanh $\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}} = \phi(2x) - 1/2$.

Balanced, bounded. Learning slow too. **ReLU** $(x)_{+} = \max 0, x$ Positive, unbounded

Derivate = 1 if x > 0, 0 if x < 0Leaky ReLU $f(x) = \max \alpha x, x$ Remove 0 derivative.

$f(x) = \max \mathbf{x}^T \mathbf{w}_1 + b_1, ..., \mathbf{x}^T \mathbf{w}_k + b_k$ (Generalization of ReLU)

20.3 Convolutional NN Sparse connections and weights sharing: reduce complexity. (e.g. pixels in pictures only depend

20.4 Reg, Data Augmentation and

- Dropout – Regularization term: $\frac{1}{2} \sum_{l=1}^{L+1} \mu^{(l)} ||W^{(l)}||_F^2$
- Weight decay is $\Theta[t](1-\eta\mu)$ in:
- $\Theta[t+1] = \Theta[t] + \eta(\nabla \mathcal{L} + \mu \Theta[t])$ Data Augm.: e.g. shift or rotation of pics
- Dropout: avoid overfit. Drop nodes randomly. (Then average multiple drop-NN)

on neighbours)

- 21 Bayes Net - Graph example: p(x, y, z) = p(y|x)p(z|x)p(x)
- $: (y \leftarrow x \rightarrow z)$ D-Separation X and Y are D-separated by Z if every path from $x \in X$ to $y \in Y$ is
- blocked by Z. (→ independent) Blocked Path contains a variable that
- is in Z and is head-to-tail or tail-to-tail - the node is head-to-head and neither the
- node nor any of its descendants are in Z. Markov Blanket (which blocks node A from
 - parents of A
 - children of A
 - parents of children of A

the rest of the net) contains: