

INTRODUCTION TO STATISTICS

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1. INTRODUCTION

Definition 1.0.1. Let $A \in \mathcal{B}(R^d)$ and $\Theta \neq \emptyset$. Suppose that $m(A) > 0$. We define

$$\mathcal{D}(A) = \{f \in L^1(A) : f \geq 0 \text{ and } \|f\|_1 = 1\}$$

and for $\theta \in \Theta$, we define

$$\mathcal{D}(A|\theta) = \{f : A \times \Theta \rightarrow \mathbb{R} : f(\cdot|\theta) \in \mathcal{D}(A)\}$$

2. SAMPLING

2.1. Inverse CDF Sampling.

2.2. Conditional Chain Sampling.

Definition 2.2.1. Let $A \subset \mathbb{R}^d$ be open, $a = (a_1, \dots, a_d) \in A$. Define $\tau_1 : \mathbb{R} \rightarrow \mathbb{R}^d$ and $A_1 \subset \mathbb{R}$ by

- $\tau_1(x) = (x, a_2, \dots, a_d)$
- $A_1 = \tau_1^{-1}(A)$

Choose $f_1 \in \mathcal{D}(A_1)$ and sample $b_1 \sim f_1$. For $j \in \{2, \dots, d\}$, define $\tau_j : \mathbb{R} \rightarrow \mathbb{R}^n$, A_j , choose f_j and sample b_j inductively by

- setting $\tau_j(x) = (b_1, \dots, b_{j-1}, x, a_{j+1}, \dots, a_d)$
- setting $A_j = \tau_j^{-1}(A)$
- choosing $f_j \in \mathcal{D}(A_j|b_1, \dots, b_{j-1})$ and setting $b_j \sim f_j$

Note that τ_j is continuous which implies that $A_j = \tau_j^{-1}(A)$ is open.

Exercise 2.2.2. Let $A \subset \mathbb{R}$ be open and $a = (a_1, \dots, a_d) \in A$. Define A_j , f_j and b_j as above and define $b \in A$ and $f : A \rightarrow \mathbb{R}$ by

$$b = (b_1, \dots, b_d)$$

and

$$f(x_1, \dots, x_d) = \prod_{j=1}^d f_j(x_j)$$

Then

- (1) $f \in \mathcal{D}(A)$
- (2) $b \sim f$

Proof. (1) Fubini's theorem implies that

$$\begin{aligned} \int_A f dm^d &= \int f_1(x_1) \left[\int f_2(x_2) \left[\dots \left[\int f_d(x_d) dm(x_d) \right] \dots \right] dm(x_2) \right] dm(x_1) \\ &= 1 \end{aligned}$$

(2) We observe that

$$\begin{aligned} [b] &= [b_d|b_{d-1}, \dots, b_1][b_{d-1}|b_{d-2}, \dots, b_1] \cdots [b_1] \\ &= f_d(b_d) \cdots f_1(b_1) \\ &= f(b) \end{aligned}$$

□

Exercise 2.2.3. Set $A = B^d(0, 1) \cap [0, 1]^d$ (the first orthant of the unit d -ball) and $a = 0$. Then $A_1 = (0, 1)$. Choose $f_1 = 1_{(0,1)}$ and sample $b_1 \sim f_1$. For each $j \in \{2, \dots, d\}$, set

$$s_j = \sqrt{1 - \sum_{k=1}^{j-1} b_k^2}$$

Then for each $j \in \{2, \dots, d\}$,

$$A_j = (0, s_j)$$

Proof. Clear. □

Exercise 2.2.4. Continuing from the previous problem, for each $j \in \{2, \dots, d\}$, choose $f_j = s_j^{-1} 1_{(0, s_j)}$. Then $f = \left(\prod_{j=2}^d s_j^{-1} \right) 1_{\prod_{j=2}^d (0, s_j)}$.

Proof. Clear. □

Definition 2.2.5. Now make $f_j \sim GP(\mu_j, c_j)$.

2.3. Importance Sampling.

2.4. Rejection Sampling.

Exercise 2.4.1. Let $f, g \in \mathcal{D}(\mathbb{R}^d)$ and $A \in \mathcal{B}(\mathbb{R}^d)$. Suppose that $m^d(A) > 0$. If $X \sim f$, then $X|X \in A \sim \|fI_A\|_1^{-1} fI_A$.

Proof. Let $C \in \mathcal{B}(\mathbb{R}^d)$. Then

$$\begin{aligned} P(X \in C|X \in A) &= P(X \in C \cap A)P(X \in A)^{-1} \\ &= \|fI_A\|_1^{-1} \int_C fI_A dm^d \end{aligned}$$

So $f_{X|X \in A} = \|fI_A\|_1^{-1} fI_A$. □

Exercise 2.4.2. Let $A, B \in \mathcal{B}(\mathbb{R}^d)$. Suppose that $A \subset B$ and $0 < m^d(A)$ and $m^d(B) < \infty$. If $X \sim \text{Uni}(B)$, then $X|X \in A \sim \text{Uni}(A)$.

Proof. Clear using the previous exercise with $f = I_B$. □

Exercise 2.4.3. (Fundamental Theorem of Simulation):

Let $f \in \mathcal{D}(\mathbb{R}^d)$ and $c > 0$. Define

$$G_c = \{(x, v) \in \mathbb{R}^{d+1} : 0 < v < cf(x)\}$$

- (1) If $X \sim f$ and $U \sim \text{Uni}(0, 1)$ are independent, then $(X, cUf(X)) \sim \text{Uni}(G_c)$.
- (2) If $(X, V) \sim \text{Uni}(G_c)$, then $X \sim f$.

Proof. First we note that $m^{d+1}(G_c) = c$.

- (1) Suppose that $X \sim f$ and $U \sim \text{Uni}(0, 1)$ are independent and put $Y = cUf(X)$. Then $Y|X = x \sim cUf(x) \sim \text{Uni}(0, cf(x))$ and we have that for each $x \in \text{supp } X$ and $y \in (0, cf(x))$,

$$\begin{aligned} f_{X,Y}(x, y) &= f_{Y|X}(y|x)f(x) \\ &= \frac{1}{cf(x)}f(x) \\ &= \frac{1}{c} \end{aligned}$$

So $(X, Y) \sim \text{Uni}(G_c)$

(2) Suppose that $(X, V) \sim \text{Uni}(G_c)$. Then $f_{X,V}(x, v) = \frac{1}{c} I_{G_c}(x, v)$. So

$$\begin{aligned} f_X(x) &= \int_{\mathbb{R}} \frac{1}{c} I_{G_c}(x, v) dm(v) \\ &= \int_0^{cf(x)} \frac{1}{c} dv \\ &= f(x) \end{aligned}$$

So $X \sim f$.

□

Exercise 2.4.4. Let $f, g \in \mathcal{D}(\mathbb{R}^d)$, $c_f, c_g > 0$ and $M > 0$. Put $\tilde{f} = c_f f$ and $\tilde{g} = c_g g$. Suppose that $\tilde{f} \leq M\tilde{g}$. If $Y \sim g$ and $U \sim \text{Uni}(0, 1)$ are independent, then $Y|U \leq \frac{\tilde{f}(Y)}{M\tilde{g}(Y)} \sim f$ and $P(U \leq \frac{\tilde{f}(Y)}{M\tilde{g}(Y)}) = \frac{c_f}{c_g M}$

Proof. Put

$$G_g = \{(y, v) \in \mathbb{R}^{d+1} : 0 < v < M\tilde{g}(y)\}$$

and

$$G_f = \{(y, v) \in \mathbb{R}^{d+1} : 0 < v < \tilde{f}(y)\}$$

Then $G_f \subset G_g$, $m^d(G_g) = c_g M$ and $m^d(G_f) = c_f$. By the first part of the fundamental theorem of simulation, we know that

$$(Y, MUc_g g(Y)) \sim \text{Uni}(G_g)$$

Since $\{(Y, MUc_g g(Y)) \in G_f\} = \{U \leq \frac{c_f f(Y)}{M c_g g(Y)}\}$, a previous exercise tells us that

$$(Y, MUc_g g(Y))|U \leq \frac{c_f f(Y)}{M c_g g(Y)} \sim \text{Uni}(G_f)$$

Then the second part of the fundamental theorem of simulation tells us that

$$Y|U \leq \frac{c_f f(Y)}{M c_g g(Y)} \sim f$$

Finally we have that

$$\begin{aligned} P\left(U \leq \frac{c_f f(Y)}{M c_g g(Y)}\right) &= P[(Y, MUc_g g(Y)) \in G_f] \\ &= \frac{c_f}{c_g M} \end{aligned}$$

□

Definition 2.4.5. (Rejection Sampling Algorithm):

Let $f, g \in \mathcal{D}(\mathbb{R}^d)$, $c_f, c_g > 0$ and $M > 0$. Put $\tilde{f} = c_f f$ and $\tilde{g} = c_g g$. Suppose that $\tilde{f} \leq M\tilde{g}$. We define the **rejection sampling algorithm** as follows:

- (1) sample $Y \sim g$ and $U \sim \text{Uni}(0, 1)$ independently
- (2) if $U \leq \frac{\tilde{f}(Y)}{M\tilde{g}(Y)}$, accept Y , else return to (1).

If we sample $(X_n)_{n \in \mathbb{N}}$ independently using the rejection sampler, then the previous exercises imply that $(X_n)_{n \in \mathbb{N}} \stackrel{iid}{\sim} f$ and the acceptance rate is $\frac{c_f}{c_g M}$.

Note 2.4.6. Phrasing the rejection sampler in terms of \tilde{f} and \tilde{g} instead of f and g is useful because we may not always be able to solve for the normalizing constants.

3. DECISION THEORY

3.1. Introduction.

Note 3.1.1. We employ the following notation and conventions:

- data space: a measurable space $(\mathcal{X}, \mathcal{F}_{\mathcal{X}})$
- parameter space: a measurable space $(\Theta, \mathcal{F}_{\Theta})$
- distribution family: $(P_{\theta})_{\theta \in \Theta} \subset \mathcal{P}(\mathcal{X}, \mathcal{F}_{\mathcal{X}})$
- estimation space: a measurable space $(\mathcal{E}, \mathcal{F}_{\mathcal{E}})$

Definition 3.1.2. Let $\eta : \Theta \rightarrow \mathcal{E}$. Then η is said to be an **estimand** if η is $(\mathcal{F}_{\Theta}, \mathcal{F}_{\mathcal{E}})$ -measurable.

Definition 3.1.3. Let $\eta : \Theta \rightarrow \mathcal{E}$ be an estimand and $\delta : \mathcal{X} \rightarrow \Theta$. Then δ is said to be an **estimator of η** if δ is $(\mathcal{F}_{\mathcal{X}}, \mathcal{F}_{\Theta})$ -measurable. We denote the set of estimators for η by Δ_{η} .

Definition 3.1.4. Let $\eta : \Theta \rightarrow \mathcal{E}$ be an estimand and $L : \Theta \times \mathcal{E} \rightarrow [0, \infty)$. Then L is said to be a **loss function for η** if

- (1) $L(\theta, \cdot)$ is $(\mathcal{F}_{\mathcal{E}}, \mathcal{B}(\mathbb{R}))$ -measurable
- (2) for each $\theta \in \Theta$, $L(\theta, \eta(\theta)) = 0$

Definition 3.1.5. Let $\eta : \Theta \rightarrow \mathcal{E}$ be an estimand and $L : \Theta \times \mathcal{E} \rightarrow [0, \infty)$ be a loss function for η . We define the **risk function associated to L** , denoted $R_L : \Theta \times \Delta_{\eta} \rightarrow [0, \infty)$, by

$$R_L(\theta, \delta) = \int_{\mathcal{X}} L(\theta, \delta(x)) dP_{\theta}(x)$$

Definition 3.1.6. Let $\eta : \Theta \rightarrow \mathcal{E}$ be an estimand, $L : \Theta \times \mathcal{E} \rightarrow [0, \infty)$ be a loss function for η and $\Pi \in \mathcal{P}(\Theta, \mathcal{F}_{\Theta})$.

4. POSTERIOR CONSISTENCY

4.1. Introduction.

Definition 4.1.1. Let $(\mathcal{X}, \mathcal{F})$ and Θ be