

# Pseudo-marginal MCMC

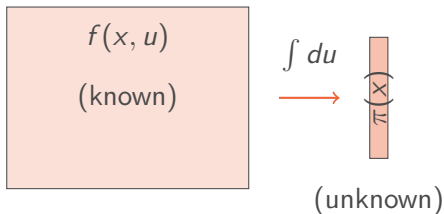
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# The Challenge: Intractable Marginals

## The Problem:

- ▶ Target:  $\pi(x) = \int f(x, u) du$
- ▶  $f(x, u)$  is known (complete data)
- ▶ Integral is **intractable**
- ▶ Standard MCMC requires exact  $\pi(x)$

**Key Insight:** We can **estimate**  $\pi(x)$  unbiasedly!



# The Pseudo-marginal Solution

## Key Prerequisites

For pseudo-marginal MCMC to be applicable, we need:

1. Ability to **evaluate**  $f(x, u)$  pointwise for any  $(x, u)$
2. Ability to **sample** from an importance distribution  $q_x(\cdot)$  over the  $u$ -space
3. The importance distribution must have appropriate support:  $q_x(u) > 0$  whenever  $f(x, u) > 0$

## Importance Sampling Estimator:

$$\hat{\pi}(x) = \frac{1}{N} \sum_{i=1}^N \frac{f(x, U_i)}{q_x(U_i)}, \quad U_i \sim q_x(\cdot)$$

**Key Property:**  $\mathbb{E}[\hat{\pi}(x)] = \pi(x)$   
(unbiased!)

**The Magic:** Replace  $\pi$  with  $\hat{\pi}$  in MH ratio!

$$\alpha = \min \left\{ 1, \frac{\hat{\pi}(y)q(y, x)}{\hat{\pi}(x)q(x, y)} \right\}$$

**Result:** Still targets correct  $\pi(x)$ !

# Why It Works: Extended Target

One can think of estimator (the “pseudo-marginal”) as the product of the true target and a random variable:

$$\hat{\pi}(x) = \pi(x)Z_x$$

where  $Z_x$  satisfies:

1. is non-negative:  $Z_x \geq 0$ ,
2. has density  $g_x(\cdot)$ :  $\int_0^\infty g_x(z)dz = 1$
3. has expectation 1:  $\mathbb{E}[Z_x] = \int_0^\infty zg_x(z)dz = 1$ .

# Why It Works: Extended Target

## Extended Target Construction:

$$\bar{\pi}(x, z) = \pi(x) \cdot z \cdot g_x(z)$$

where  $g_x(z)$  is the density of  $Z_x$

## Key Property:

$$\int \bar{\pi}(x, z) dz = \pi(x)$$

Now apply Metropolis–Hastings with proposal

$$\bar{q}((x, z), (y, w)) := q(x, y) \cdot g_y(w).$$

## Intuition:

- ▶ Run exact MCMC on  $(x, z)$  space
- ▶ Marginal in  $x$  gives correct target
- ▶  $z$  represents the "noise" in estimates

# Equivalence to MH on Extended Space

## Theorem (Equivalence)

*Metropolis-Hastings on the extended target  $\bar{\pi}$  with proposal  $\bar{q}$  is equivalent to the pseudo-marginal algorithm using estimates  $\hat{\pi}$ .*

Proof Sketch: The MH acceptance ratio on the extended space is:

$$\begin{aligned}\alpha_{\text{ext}} &= \min \left\{ 1, \frac{\bar{\pi}(y, w) \bar{q}((y, w), (x, z))}{\bar{\pi}(x, z) \bar{q}((x, z), (y, w))} \right\} \\ &= \min \left\{ 1, \frac{\pi(y) \cdot w \cdot g_y(w) \cdot q(y, x) \cdot g_x(z)}{\pi(x) \cdot z \cdot g_x(z) \cdot q(x, y) \cdot g_y(w)} \right\} \\ &= \min \left\{ 1, \frac{\pi(y) \cdot w \cdot q(y, x)}{\pi(x) \cdot z \cdot q(x, y)} \right\} = \min \left\{ 1, \frac{\hat{\pi}(y) q(y, x)}{\hat{\pi}(x) q(x, y)} \right\} = \alpha_{pm}\end{aligned}$$

In the last step, we used  $\hat{\pi}(x) = \pi(x)z$  and  $\hat{\pi}(y) = \pi(y)w$ , which is exactly the pseudo-marginal acceptance probability.

# Pseudo-marginal MCMC Algorithm

Given  $(X^{(t-1)}, \hat{\pi}^{(t-1)})$ :

1. **Propose:**  $Y \sim q(X^{(t-1)}, \cdot)$

2. **Estimate:**

- ▶ Sample  $U_i \sim q_Y(\cdot)$
- ▶  $\hat{\pi}(Y) = \frac{1}{N} \sum_i \frac{f(Y, U_i)}{q_Y(U_i)}$

3. **Accept with probability:**

$$\alpha = \min \left\{ 1, \frac{\hat{\pi}(Y) q(Y, X^{(t-1)})}{\hat{\pi}^{(t-1)} q(X^{(t-1)}, Y)} \right\}$$

4. **Update:**

- ▶ If accept:  $(X^{(t)}, \hat{\pi}^{(t)}) = (Y, \hat{\pi}(Y))$
- ▶ Else:  $(X^{(t)}, \hat{\pi}^{(t)}) = (X^{(t-1)}, \hat{\pi}^{(t-1)})$

## Critical Points:

- ▶ Store estimates with states! In the next iteration, use the stored  $\hat{\pi}(X^{(t-1)})$ .
- ▶ Fresh randomness for each proposal. Every time you propose a new state  $Y$ , you must generate a completely new, independent estimate  $\hat{\pi}(Y)$  using fresh random samples.
- ▶ Works with *any* MH proposal  $q$