

VALORE ATTESO e MATRICE DI COVARIANZA

$$(X, Y), \begin{pmatrix} X \\ Y \end{pmatrix}$$

$$\begin{aligned} \text{VALORE ATTESO : } & (\mathbb{E}[X], \mathbb{E}[Y]) \\ & \mathbb{E}[(X, Y)] \end{aligned}$$

FUNZIONE DI UN VETTORE ALEATORIO

$$h: \mathbb{R}^2 \rightarrow \mathbb{R}$$

$$Z = h(X, Y)$$

Teorema

(X, Y) vettore aleatorio discreto con densità di probabilità
congiunta $P(X, Y)$,

$$\mathbb{E}[Z] = \mathbb{E}[h(X, Y)] = \sum_{i,j} h(x_i, y_j) P_{(X,Y)}(x_i, y_j)$$

COROLARIO

(X, Y) vettore aleatorio discreto (cioè, X e Y sono v.a. discrete) .

$$X \perp\!\!\!\perp Y \quad \cancel{\text{def}} \quad \mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

N.B.

$$\begin{aligned} X \perp\!\!\!\perp Y : \quad & \mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B), \quad A, B \subset \mathbb{R} \\ X \perp\!\!\!\perp Y \iff & P_{(X,Y)}(x_i, y_j) = P_X(x_i)P_Y(y_j), \quad \forall i, \forall j \end{aligned}$$

DIM

$$\begin{aligned} \mathbb{E}[XY] &\stackrel{\text{def}}{=} \sum_{i,j} x_i y_j P_{(X,Y)}(x_i, y_j) \stackrel{X \perp\!\!\!\perp Y}{=} \sum_{i,j} x_i y_j P_X(x_i)P_Y(y_j) \\ &= \left(\sum_i x_i P_X(x_i) \right) \left(\sum_j y_j P_Y(y_j) \right) = \mathbb{E}[X]\mathbb{E}[Y]. \end{aligned}$$

OSS.

$$X \perp\!\!\!\perp Y \iff f(X) \perp\!\!\!\perp g(Y)$$

Si può dimostrare che

$$X \perp\!\!\!\perp Y \iff$$

$$\mathbb{E}[f(X)g(Y)] = \mathbb{E}[f(X)]\mathbb{E}[g(Y)]$$

MATRICE DI COVARIANZA

$$\mathbb{E}((X, Y)) = (\mathbb{E}[X], \mathbb{E}[Y])$$

$(X, Y) :$

$$\begin{pmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{pmatrix} = \begin{pmatrix} \text{Var}(X) & \text{Cov}(X, Y) \\ \text{Cov}(Y, X) & \text{Var}(Y) \end{pmatrix}$$

$$(X, Y) = (X_1, X_2)$$

$$\begin{pmatrix} \text{Cov}(X_1, X_1) & \text{Cov}(X_1, X_2) \\ \text{Cov}(X_2, X_1) & \text{Cov}(X_2, X_2) \end{pmatrix}$$

$$a_{ij} = \text{Cov}(X_i, X_j)$$

$$1) \quad \text{Cov}(X_1, X_1) = \text{Var}(X_1) \quad | \quad 2) \quad \text{Cov}(X, Y) = \text{Cov}(Y, X)$$

Definizione
 X e Y sono v.a. discrete. La covarianza di X e Y
 è data da

$$\begin{aligned}\text{Cov}(X, Y) &= \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] \\ &= \sum_{i,j} (x_i - \mathbb{E}[X])(y_j - \mathbb{E}[Y]) P_{(X,Y)}(x_i, y_j)\end{aligned}$$

Se $\text{Cov}(X, Y) = 0$, le v.a. X e Y si dicono
 SCORRELATE.

Se $\text{Var}(X) > 0$ e $\text{Var}(Y) > 0$, definiamo il COEFFICIENTE
 DI CORRELAZIONE:

$$\rho_{X,Y} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}}.$$

OSS.

1)

$x_i - \mathbb{E}[X] =$ scarto dalla media di x_i

$\text{Cov}(X, Y) =$ somma del prodotto degli scarti
dalla media

2)

$$\text{Cov}(X, X) = \text{Var}(X)$$

3) $\text{Cov}(X, Y) = \text{Cov}(Y, X)$

Teorema 3.2 del Cap. "Stat. descr.
e Teoremi limiti"

4)

$$-1 \leq \rho_{X,Y} \leq 1$$

Inoltre :

$$\rho_{X,Y} = \pm 1 \iff Y = aX + b$$

$$\rho_{X,Y} = \pm 1$$

Ex. 6

$$|\text{Cov}(X, Y)| \leq \sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}$$

5) La covarianza ci dice se c'è approssimativamente
DIPENDENZA LINEARE.

$\rho_{X,Y} = \pm 1$, allora c'è veramente dipendenza lineare.

$\rho_{X,Y} = 0$, allora non è detto che X e Y siano
indipendenti

$$X \perp\!\!\!\perp Y \quad \cancel{\Rightarrow} \quad \text{Cov}(X, Y) = 0 \quad (X \text{ e } Y \text{ correlate})$$

Teorema

$$X \perp\!\!\!\perp Y \Rightarrow \text{Cov}(X, Y) = 0$$

Lemme

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y]$$

Dimm (Lemme)

$$\begin{aligned} \text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] = \\ &= E[XY - XE[Y] - YE[X] + E[X]E[Y]] \\ &= E[XY] - E[Y]E[X] - \cancel{E[X]E[Y]} + \cancel{E[X]E[Y]} \end{aligned}$$

N.B.

$$E[aX + b] = aE[X] + b$$

Domin (Teorema)

$$X \perp\!\!\!\perp Y \implies E[XY] = E[X]E[Y]$$

Quindi

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] =$$

$$= 0$$

↑
 $X \perp\!\!\!\perp Y$

$$Y = \underbrace{\log(1 + \sin X)}_Z$$

Ex. 2.3

X e Y v.a. discrete :

$$X \perp\!\!\!\perp Y, \quad X \sim B\left(\frac{1}{2}\right), \quad Y \sim B\left(\frac{1}{2}\right)$$

Siano $U = X + Y$ e $V = |X - Y|$

- Determinare congiunta e marginali di U e V .
- $P(V < U) = ?$
- Calcolare $\text{Var}(U)$, $\text{Var}(V)$, $\text{Cov}(U, V)$.
- U e V sono indipendenti?

a)

X \ Y	0	1	P_X
0	$\frac{1}{14}$	$\frac{1}{14}$	$\frac{1}{2}$
1	$\frac{1}{14}$	$\frac{1}{14}$	$\frac{1}{2}$
P_Y	$\frac{1}{2}$	$\frac{1}{2}$	1

(X, Y)	(U, V)
$(0, 0)$	$(0, 0)$
$(0, 1)$	$(1, 1)$
$(1, 0)$	$(1, 1)$
$(1, 1)$	$(2, 0)$

$$S_U = \{0, 1, 2\} \quad \text{e} \quad S_V = \{0, 1\}$$

U \ V	0	1	P_{UV}
0	$\frac{1}{14}$	0	$\frac{1}{14}$
1	0	$\frac{1}{12}$	$\frac{1}{12}$
2	$\frac{1}{14}$	0	$\frac{1}{14}$
P_V	$\frac{1}{12}$	$\frac{1}{2}$	1

$$V \sim B\left(\frac{1}{2}\right)$$

$$U \sim \text{Bin}\left(2, \frac{1}{2}\right)$$

$$(U = V \bmod 2)$$

$$b) \quad P(V < U) = \sum_{\substack{i,j: \\ V_j < U_i}} P(U, V)(u_i, v_j)$$

$$= P(U, V)(2, 0) + P(U, V)(2, 1) = P_U(2) = \frac{1}{4}$$

$$c) \quad E[U] = 2 \cdot \frac{1}{2} = 1 \quad E[V] = \frac{1}{2}$$
$$\text{Var}(U) = 2 \cdot \frac{1}{4} = \frac{1}{2} \quad \text{Var}(V) = \frac{1}{4}$$

$$E[UV] = \sum_{i,j} u_i v_j P(U, V)(u_i, v_j) = 1 \cdot 1 \cdot P(U, V)(1, 1) = \frac{1}{2}$$

$$\text{Cov}(U, V) = E[UV] - E[U]E[V] = \frac{1}{2} - \frac{1}{2} = 0$$

d) $\cup \perp\!\!\!\perp V$?

$$\cup \perp\!\!\!\perp V \Leftrightarrow p_{(\cup, V)}^{(u_i, v_j)} = p_U^{(u_i)} p_V^{(v_j)}$$

No, infatti ad esempio

$$p_{(\cup, V)}^{(0, 0)} = \frac{1}{4} \neq \underset{\uparrow}{p_U^{(0)}} \underset{\uparrow}{p_V^{(0)}}$$
$$\frac{1}{4} \quad \frac{1}{2}$$