- Reporting Bias, not External Focus: A Robust Bayesian Meta-Analysis and
- Systematic Review of the Attentional Focus Literature
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16 Abstract

Evidence has ostensibly been accumulating over the past two decades suggesting that an 17 external focus on the intended movement effect (e.g., on the golf club during a swing) is 18 superior to an internal focus on body movements (e.g., on your arms during a swing) for 19 skill acquisition. Seven previous meta-studies have all reported evidence of external focus 20 superiority. The most comprehensive of these concluded an external focus enhances motor 21 skill retention, transfer, and performance, leads to reduced eletromyographic activity during performance, and that more distal external foci are superior to proximal external 23 foci for performance. Here, we re-analyzed these data using robust Bayesian meta-analyses that included several plausible models of publication bias. We found moderate to strong 25 evidence of publication bias for all analyses. After correcting for publication bias, estimated mean effects were negligible: g = 0.01 (performance), g = 0.15 (retention), g = 0.050.09 (transfer), g = 0.06 (electromyography), and g = -0.01 (distance effect). Bayes factors indicated data favored the null for each analysis, ranging from $BF_{01} = 1.3$ (retention) to 29 5.75 (performance). We found clear evidence of heterogeneity in each analysis, suggesting 30 the impact of attentional focus depends on yet unknown contextual factors. Our results 31 contradict the existing consensus that an external focus is always more effective than an 32 internal focus. Instead, focus of attention appears to have a variety of effects that we 33 cannot account for, and on average those effects are small to nil. These results parallel 34 previous metascience suggesting publication bias has obfuscated the motor learning 35 literature. 36 Keywords: Motor learning, Motor skills, Attentional focus, OPTIMAL theory, Skill

37 Keywords: Motor learning, Motor skills, Attentional focus, OPTIMAL theory, Skill acquisition, Metascience

- ³⁹ Public Significance Statement A robust Bayesian meta-analysis showed that directing
- learners to focus their attention on their intended movement effects—often called an
- external focus—may have little-to-no effect on motor performance and learning on average.
- 42 Although the consensus among researchers and practitioners has been that an external
- focus is superior to focusing on one's own body during practice, the present results suggest
- 44 this may depend on unknown factors and our current understanding has been distorted by
- ⁴⁵ publication bias. These results highlight that a more cautious approach is necessary when
- recommending the use of external foci in applied settings until a more reliable body of
- 47 literature can be established using preregistration, Registered Reports, and well-powered
- designs through multi-site collaborations.

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Reporting Bias, not External Focus: A Robust Bayesian Meta-Analysis and Systematic Review of the Attentional Focus Literature

Where should you focus when performing and/or learning a motor skill? The most 51 basic of questions for a novice learner and an experienced performer alike. Is it better to 52 focus on what you are doing: where your body is in space and how it is behaving? Or is it 53 better to focus on what you intend to do: the end effect you are trying to achieve independent of how your body achieves it? This question has been the topic of decades of research comparing an internal focus of attention (i.e., focusing on your own body) to an external focus of attention (i.e., focusing on the intended effect of the action). Gabriele Wulf pioneered this area of inquiry in 1998, publishing a two-experiment article illustrating the benefits of adopting an external focus (Wulf et al., 1998). In the experiments, instructing learners to focus on the wheels of a ski simulator (Experiment 1) or the markers on a balance platform (Experiment 2) led to improved motor learning compared to focusing on one's feet. Dozens of studies have since replicated these initial findings (see Wulf, 2007, 2013 for reviews). 63 Previous reviews have argued that research shows benefits of an external focus in 64 four main areas: (a) effectiveness at accuracy and balance tasks, (b) efficiency in electromyographic activity, force production, speed, and endurance tasks, (c) promoting 66 automaticity, and (d) enhancing movement form (Chua et al., 2021; Wulf, 2007, 2013; Wulf 67 & Lewthwaite, 2016). A leading explanation for the mechanism causing these benefits is goal-action coupling: a process proposed in Wulf and Lewthwaite's (2016) OPTIMAL theory involving a shift at the neural level that simultaneously directs action toward success and stifles deleterious self-focused cognition. Although focus of attention is 71 fundamental to the OPTIMAL theory, various perspectives in motor behavior have offered 72 complementary accounts for external focus benefits. For example, it has been argued from 73 the constraints-based approach that an external focus promotes the search of the task during practice and provides a constraint on emerging actions (Davids et al., 2003). It has

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also been argued that actions and perceptions share a common (cognitive) code; therefore,
   focusing on the intended (perceptual) effect of an action is consistent with its underlying
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   neural coding (Hommel et al., 2001; Prinz, 1990; Wulf & Prinz, 2001). While research
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   continues to explore the putative mechanisms, there is consensus in the motor learning
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   community that adopting an external focus of attention can improve motor performance,
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   retention, transfer, and movement efficiency—at least most of the time (Chua et al., 2021;
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   Grgic et al., 2021; Grgic & Mikulic, 2022; T. Kim et al., 2017; Lee & Carnahan, 2021; Li
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   et al., 2022; Makaruk et al., 2020; Nicklas et al., 2022).<sup>1</sup>
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          Buttressed by the largely positive results in the research literature, external focus of
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   attention is now widely recommended outside of academia, including by sport coaches
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   (skating: Smale, 2021; golf: T. Neumann, 2017; tennis: Kuzdub, 2022; baseball: Peterson,
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   2019), fitness coaches (Kompf, 2015; N. Winkelman, 2015), and therapists (Lo, 2019;
   Magne & Edge, 2017). Researchers continue to study the use of externally focused
   instructions and feedback in clinical settings (Johnson et al., 2023) and are currently
   developing strategies for increasing awareness of the research among rehabilitation
   professionals (Hussien et al., 2023a, 2023b; Hussien & Ste-Marie, 2023). As external focus
   becomes evermore mainstream, recent concerns that much of the motor learning literature
   may be exaggerated by reporting bias (e.g., Lohse et al., 2016; McKay, Hussien, et al.,
   2022; McKay, Yantha, et al., 2022; Mesquida et al., 2022; Twomey et al., 2021) underlines
   the need for careful assessment of the evidence. The external focus literature may be
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   especially at risk because substantial reporting bias has been found in the motor learning
   literature investigating the other factors within OPTIMAL theory (Bacelar et al., 2022;
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   McKay et al., 2023; McKay, Yantha, et al., 2022). Note that reporting bias encompasses
   various forms of selection bias that limit the availability of data. Potential reporting bias
   <sup>1</sup> Despite acknowledging this as the general consensus in the field, it is important to note that there are
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¹ Despite acknowledging this as the general consensus in the field, it is important to note that there are mixed findings and alternative discussions in this area of research (e.g., Bernier et al., 2016; Brick et al., 2014; Canning, 2005; Collins et al., 2016; Emanuel et al., 2008; Lawrence et al., 2011; Maurer & Munzert, 2013; Peh et al., 2011; Perkins-Ceccato et al., 2003; Schorer et al., 2012; Zentgraf & Munzert, 2009).

mechanisms can be modeled, though models cannot determine the specific reason for censorship within a literature.

Previous Meta-Analyses

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There have been seven meta-analyses comparing the effects of internal and external 103 focus instructions on motor outcomes. Five have focused on specific task-types: balance 104 (T. Kim et al., 2017), jumping (Makaruk et al., 2020), sprinting (Li et al., 2022), strength 105 (Grgic et al., 2021), and endurance (Grgic & Mikulic, 2022). A sixth included all motor 106 tasks and focused specifically on the immediate effect on performance (Nicklas et al., 2022). 107 Chua and colleagues (2021) conducted the most comprehensive meta-analysis of the seven, 108 including all task-types and estimating effects on performance, retention, transfer, electromyography activity, and the distance effect. All seven studies reported the results of random effects meta-analyses as the primary estimates for the effect of focus of attention. 111 Although there was some variance in point estimates and confidence intervals, each of the 112 studies reported evidence that an external focus is superior to an internal focus. 113 Importantly, a random effects model assumes no reporting bias and has been shown 114 to be quite biased in the presence of selective reporting for statistical significance (Barto, 115 Maier, Shanks, et al., 2023; Bom & Rachinger, 2019; Carter et al., 2015; Carter et al., 116 2019; Kvarven et al., 2020; Stanley et al., 2017, 2022). Two of the seven previous studies 117 (Chua et al., 2021; T. Kim et al., 2017) reported evidence of funnel plot asymmetry, which 118 is consistent with selective reporting of significant results. Two studies did not find 119 evidence of funnel plot asymmetry (Li et al., 2022; Nicklas et al., 2022), and the other 120 three did not investigate reporting bias at all (Grgic et al., 2021; Grgic & Mikulic, 2022; 121 Makaruk et al., 2020). Both studies that observed evidence of reporting bias conducted a 122 fail-safe-style sensitivity analysis, but did not correct the primary estimates for the 123 presence of bias. The meta-analysis by Chua et al. (2021) did calculate worst-case-scenario 124 estimates based on a random effects meta-analysis of the non-significant results. Thus, 125 although reporting bias may be prevalent in the field of motor learning (Lohse et al., 2016), 126

and two previous meta-analyses have found evidence of reporting bias in the attentional focus literature (Chua et al., 2021; T. Kim et al., 2017), the primary estimates from all previous meta-analyses assume bias is absent.

Consistent with the other studies, Chua et al. (2021) reported moderate benefits of 130 an external focus for learning measures (g = 0.58) and small benefits for performance 131 measures (g = 0.26) and the distance effect (g = 0.22). Chua and colleagues also reported 132 a large effect on electromyography activity (g = 0.83). In lieu of bias-corrected estimates, 133 worst-case scenario estimates were calculated to evaluate how sensitive the primary 134 estimates were to an assumed model of reporting bias. Under the assumed model, 135 significant results in the predicted direction are published without censorship, and all 136 non-significant results and significant results in the opposite direction are censored at the 137 same rate. The worst-case scenario is simply the random effects estimate of all the non-preferred outcomes as a preference for significant results in the predicted direction 139 cannot upwardly bias an estimate if significant results are removed. If the worst-case 140 scenario is positive, then one can conclude that no amount of reporting bias could 141 attenuate the point estimate to the null value. However, this conclusion is only merited if 142 censorship is entirely captured by the assumed model. If other plausible mechanisms of 143 censorship are present, then the assumed model does not hold, and the worst-case scenario 144 estimates can no longer be considered as such. 145

Although Chua et al. (2021) concluded that no amount of reporting bias could 146 attenuate the effect to the null value for any measure (performance, retention, transfer, 147 electromyography, and the distance effect), there are several plausible censorship 148 mechanisms that were unexplored. For example, it is plausible that nearly significant 149 results, often called non-significant trends (Otte et al., 2022), were censored less than other 150 non-significant trends. It is also possible that point estimates favoring an internal focus 151 were the least preferred result. If these plausible alternative censorship mechanisms were 152 active in the attentional focus literature, then the random effects estimate of "all 153

non-significant in the predicted direction" results would be positively biased. Although
Chua et al. (2021) concluded that external focus superiority is not sensitive to reporting
bias, it remains unknown if that conclusion is sensitive to the form of reporting bias that
was assumed.

Seven previous meta-analyses provide primary estimates of the potential benefit of

158 Present Study

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an external focus of attention while assuming reporting bias is absent. Given the evidence 160 of reporting bias reported in two of those studies (Chua et al., 2021; T. Kim et al., 2017), 161 along with evidence of extensive bias in related literatures (e.g., Lohse et al., 2016; McKay 162 et al., 2023), bias-corrected estimates are needed. There are several plausible mechanisms of reporting bias, and the true model is unknowable. Therefore, using a robust Bayesian approach to meta-analysis (Barto, Maier, Wagenmakers, et al., 2023), we leveraged 165 Bayesian model-averaging to fit several plausible models of reporting bias to the attentional 166 focus literature examined by Chua et al. (2021). Greater weight was given to the models 167 that best accounted for the results and less weight was given to poorly performing models. 168 This approach allowed us to calculate reporting-bias-adjusted estimates for the effect of 169 attentional focus on motor learning, performance, electromyography activity, and for the 170 distance effect. Our approach naturally allowed us to evaluate Chua and colleagues' (2021) 171 claims that no amount of reporting bias could attenuate the effect to the null value. 172 In addition to censorship mechanisms, we also explored the role of post hoc outcome 173 selection leading to potentially exaggerated estimates. The previous seven meta-analyses 174 either did not specify exactly how outcomes were selected for analysis (Grgic et al., 2021; 175 Grgic & Mikulic, 2022; T. Kim et al., 2017; Li et al., 2022; Makaruk et al., 2020), excluded 176 studies that had more than one performance measurement unless the measures could be 177 ranked and a primary measure could be selected (Nicklas et al., 2022), or selected the 178 outcome positioned as primary in the original research article (Chua et al., 2021). The 179 external focus literature has not made use of preregistration or Registered Reports, so it is 180

possible that the most impressive results have sometimes been positioned as primary
because they were the most impressive. If this sort of *post hoc* selection is present, then
selecting outcomes based on their status in the original article may lead to biased
estimates. To evaluate the possibility of *post hoc* selection bias, we extracted effect size
estimates for the retention test outcomes that were not selected by Chua et al. (2021), but
could have been, and compared them to the selected "primary" outcomes.

In the present study, we addressed the following questions: (a) What is the 187 reporting-bias-adjusted estimate for the effect of attentional focus on learning, 188 performance, electromyography activity, and the distance effect? (b) How sensitive are 189 random effects estimates to the assumption that reporting bias is absent? (c) How sensitive 190 are Chua and colleagues' (2021) conclusions that no amount of reporting bias could 191 attenuate the effect to the null value to the specific model of censorship that was 192 evaluated? and (d) How influential was post hoc selection bias on the estimated benefits of 193 an external focus of attention on retention performance? 194

195 Methods

Transparency and Openness

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We adhered to the Meta-Analysis Journal Reporting Standards (MARS) guidelines 197 for meta-analytic reporting (Appelbaum et al., 2018). The data, code, and preregistration 198 for this study are publicly available at https://osf.io/vfmx2 (McKay & Carter, 2024).² The 190 data for each primary outcome measure were collected and reported by Chua et al. (2021). 200 Our re-analysis of those data was not preregistered as we were already aware of Chua and 201 colleagues' (2021) primary conclusions and had seen the data visualizations in their study. 202 Data for up to three additional outcomes from each experiment that examined retention 203 test performance were collected and analyzed according to our preregistered protocol. 204 There were no deviations from our preregistration. 205

Statistical analyses were conducted using R (Version 4.3.2; R Core Team, 2023) and

² Data and scripts are also available at https://github.com/cartermaclab/proj_foa-optimal-theory.

the R-packages compute.es (Version 0.2.5; Re, 2013), daff (Version 0.3.5; Fitzpatrick et al., 207 2019), extrafont (Version 0.19; Chang, 2023), faux (Version 1.2.1; DeBruine, 2023), qqdist 208 (Version 3.2.1; Kay, 2023), qt (Version 0.9.0; Iannone et al., 2023), magick (Version 2.7.4; 209 Ooms, 2023), metafor (Version 4.0.0; Viechtbauer, 2010), patchwork (Version 1.2.2; 210 Pedersen, 2022), plotly (Version 4.10.2; Sievert, 2020), Publication Bias (Version 2.3.0; 211 Braginsky et al., 2023), renv (Version 0.17.2; Ushey, 2023), RoBMA (Version 2.3.2; Barto & 212 Maier, 2020), stringi (Version 1.7.12; Gagolewski, 2022), and tidyverse (Version 2.0.0; 213 Wickham et al., 2019) were used in this project. 214

215 Eligibility

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Our analysis was restricted to the studies included in the study by Chua et al.

(2021), meaning our study inherits the inclusion criteria imposed in their study: (a)

published in English between February 1998 and April 2019, (b) in a peer-reviewed journal,

(c) compared internal and external foci of attention, or at least two types of external focus,

(d) measured motor learning or performance, (e) used a within-participant design to

measure performance and a between participants design to measure learning, (f) included

sufficient data to calculate effect sizes, and (g) were experiments.

223 Data Collection Process

The data reported by Chua et al. (2021) were extracted directly from the published article. Additionally, up to three outcomes were extracted from each experiment included in the meta-analysis of retention test performance.³ Data were extracted in duplicate by a team of six researchers working independently (AC, JS, CDF, HH, KA, FA). The lead author evaluated each pair of extractions using the R package daff (Fitzpatrick et al., 2019) for consensus and resolved all conflicts.

Outcome measures were selected for extraction based on our preregistered priority

³ We chose to focus on retention effects because the performance estimates were already small. The retention estimates were substantial, and retention tests are often the focal learning measure in an experiment. Almost all transfer tests were from studies that also included a retention test, so focusing on retention outcomes was the simplest way to test our research question.

Table 1

Priority list for extracting outcome measures.

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Priority	Measure	Priority	Measure
1	Absolute error	6	Relative timing error
2	Root mean squared error / Total error	7	Absolute constant error
3	Accuracy points	8	Movement time
4	Variable error	9	Movement form (Expert raters)
5	Absolute timing error	10	Other

list (see Table 1). A priority list achieved two goals. First, it prevented selection bias when

several outcomes were reported in a study by establishing which outcomes to select a 232 priori. Second, the list prioritized outcomes most connected to the goal of the task over 233 outcomes only correlated with success. This ensured the dependent variables most 234 indicative of goal-action coupling were selected from each study. 235 The sample sizes, direction of effect, means, and standard deviations were extracted 236 for each measure when available. If standard deviations were not reported, data were 237 extracted in the following order of priority: means and standard errors, F-values, then 238 t-values. If the required data were not reported in the text of the article, but were 239 presented in figures with error bars, then the mean and standard deviation were extracted by digitizing the plots (Rohatgi, 2022). Data from six studies were digitized. If data could 241 not be extracted with plot digitization, then the authors were emailed, and the data were 242 requested. If the authors did not respond, a follow up email was sent. Emails were sent to authors requesting data for five effects, and one author responded with the requested data.

compute.es (Re, 2013). Risk of bias from methodological weaknesses was well probed by

Hedges' q for the newly extracted outcomes was calculated using the R package

Chua et al. (2021) and was not revisited in this study.

48 Synthesis Methods

249 Influential Cases

We screened the data for influential cases using the R package metafor (Viechtbauer & Cheung, 2010). After fitting univariate random effects models for each meta-analysis, externally standardized residuals and Cook's distances were calculated. Studies identified as extreme by both measures were considered influential and a sensitivity analysis was conducted with the studies removed.⁴

55 Reporting Bias

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We implemented a robust Bayesian approach (Barto, Maier, Wagenmakers, et al., 256 2023) to reanalyze the five meta-analyses reported by Chua et al. (2021). We used neutral 257 default priors for the presence of an effect (Normal(M = 0, SD = 1), p = 0.5), the presence 258 of heterogeneity (InvGamma(1, 0.15), p = 0.5), and the presence of reporting bias (p =259 0.5). Reporting bias was probed using selection models and funnel plot regression models. 260 In the selection model class, six different weight-function models were fit to model 261 censorship based on specific p-value thresholds. For example, one selection model captures 262 the possibility that significant results in the predicted direction are more likely to survive 263 to be published than both null results and significant results in the unpredicted direction. 264 Another selection model captures the possibility that results in the unpredicted direction 265 are the least likely to survive censorship, while non-significant trends are more likely than 266 other null results, but not as likely as significant results to survive. 267

A total of six selection models capturing different plausible censorship scenarios are assigned half of the prior probability that reporting bias exists. The other half of the prior probability is allocated to funnel plot regression models. The precision-effect test (PET) and precision-effect estimate with standard errors (PEESE) respectively model a linear and

⁴ Our approach to influential case screening differed from the approach employed by Chua et al. (2021) and we therefore arrived at a different number of outliers for each analysis (see Supplementary A for more details).

quadratic relationship between standard error and effect size. If the data were censored 272 such that lower p-values had a higher probability of surviving, a correlation would emerge 273 between two otherwise independent causes of p-values: effect sizes and standard errors. 274 The PET method fits a linear relationship between effect size and standard error, modeling 275 a consistent level of censorship across studies. The PEESE method fits a quadratic 276 relationship, reflecting the possibility that studies with small standard errors, and thus 277 large samples, are likely to be reported regardless of the results, and small studies with 278 large standard errors require increasingly impressive results to garner publication.⁵ 279 A total of 36 models were fit to the data with every combination of the eight 280

281 reporting bias models, models assuming an effect, no effect, heterogeneity, no heterogeneity,
282 and no reporting bias (see Supplementary B for more details). The estimates of each model
283 were combined using Bayesian model-averaging, where model estimates are weighted based
284 on how well the model fit the data. A single posterior distribution was generated for the
285 average effect of an external focus and the average value of tau—the estimated
286 heterogeneity. Further, Bayes Factors were calculated measuring the evidence in favor of an
287 effect, the presence of heterogeneity, and reporting bias.

Post-Hoc Selection Bias

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A multi-level mixed effects model with outcomes nested in study, and with cluster-robust standard errors compared the outcomes selected by Chua et al. (2021) to the additional outcomes that might have been selected instead. Profile analysis was conducted to ensure the model converged on unique solutions for estimates of Mu (the mean effect of external focus in the population) and tau.

⁵ Priors for the six selection models were: ω [two-sided: 0.05] \sim CumDirichlet(1, 1), ω [two-sided: 0.1, 0.05] \sim CumDirichlet(1, 1, 1), ω [one-sided: 0.05] \sim CumDirichlet(1, 1, 1), ω [one-sided: 0.05, 0.025] \sim CumDirichlet(1, 1, 1), ω [one-sided: 0.5, 0.05] \sim CumDirichlet(1, 1, 1, 1). Priors for the two regression models were: PET \sim Cauchy(0, 1)[0, Inf], PEESE \sim Cauchy(0, 5)[0, Inf].

294 Results

Model-averaged posterior distributions for each analysis with and without outliers are presented in Figure 1.6

Performance

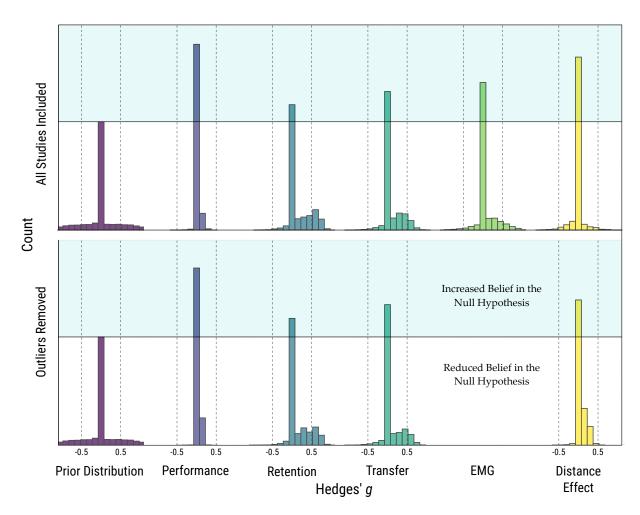
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Influence analyses revealed four studies (Marchant, Greig, & Scott, 2009; Nadzalan 298 et al., 2015; Porter et al., 2010; Sherwood et al., 2014 Experiments 1 and 2) could be 299 considered outliers in the performance meta-analysis. We report the results with all studies 300 included first, then with outliers removed. The mean of the model-averaged posterior distribution for the difference between external and internal foci of attention on motor skill 302 performance was g = 0.01, 95% credible interval: 0, 0.17. The data were over 5 times more 303 compatible with the null hypothesis than the alternative, $BF_{10} = 0.17$. There was clear 304 evidence of heterogeneity, $\tau = 0.40$, BF $_{\rm rf} =$ Infinite. There was also clear evidence of 305 publication bias, $BF_{pb} = 162,651.73$. Removing influential cases did not substantively 306 change the conclusions: g = 0.02, 95% credible interval: 0, 0.16, BF₁₀ = 0.26; $\tau = 0.25$, 307 $BF_{rf} = 602774614$; $BF_{pb} = 97,268.05$.

⁶ Model convergence diagnostics were conducted for all RoBMA (Barto & Maier, 2020) analyses. In each case, Rhat convergence values were less than 1.05 and effect sampling sizes were a few hundred or more.

Figure 1

Posterior plots of the standardized mean difference with and without outliers.



Note. The effect size estimates (g) of each meta-analysis with all studies included (top row) and with outliers removed (bottom row). The histograms in the first column reflect the prior distribution, with 50% of the probability density concentrated on zero effect (the null hypothesis) and 50% of the density normally distributed ($M=0,\ SD=1$). The model-averaged posterior distributions for performance, retention, transfer, electromyography, and the distance effect are presented in the second through sixth columns, respectively. Increased belief in the null hypothesis is visible for each analysis, illustrated by the increased height of the spike at g=0 in all posteriors relative to the prior distribution.

309 Retention

Two studies (Ahmad et al., 2013; Tse, 2019) were identified as possible outliers in 310 the retention test meta-analysis. Again, the results with all studies included are reported 311 first, then with outliers removed. The mean of the model-averaged posterior distribution 312 for the effect of focus of attention on retention was g = 0.15, 95% credible interval: -0.17, 313 0.74. The data were somewhat more consistent with the null hypothesis than the 314 alternative, $BF_{10} = 0.75$. There was clear evidence of heterogeneity, $\tau = 0.65$, $BF_{rf} =$ 315 Infinite. The data were 5.9 times more compatible with models assuming publication bias 316 than without, $BF_{pb} = 5.92$. Removing two influential cases did not substantively change 317 the conclusions: $g=0.14,\,95\%$ credible interval: -0.18, 0.73, BF₁₀ = 0.73; $\tau=0.50,\,$ BF_{rf} = 318 1,688,117,430.52; BF_{pb} = 7.62. 319

320 Transfer

One possible outlier (Tse, 2019) was identified in the transfer test meta-analysis. 321 The mean of the model-averaged posterior distribution of all transfer outcomes was q =322 0.09, 95\% credible interval: -0.21, 0.62. The results were somewhat more likely under the 323 null hypothesis than the alternative, $BF_{10} = 0.57$. There was clear evidence of 324 heterogeneity, $\tau = 0.56$, BF_{rf} = Infinite. The data were more than 6.4 times more likely 325 under models assuming publication bias, $BF_{pb} = 6.45$. Removing one influential case did 326 not substantively change the conclusions: $g=0.09,\,95\%$ credible interval: -0.23, 0.63, BF₁₀ 327 = 0.55; τ = 0.45, BF_{rf} = 101,220.12; BF_{pb} = 9.06. 328

Electromyography

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There were no outliers identified in the electromyography meta-analysis. The mean of the model-averaged posterior distribution for the effect of attentional focus on electromyography activity was g=0.06, 95% credible interval: -0.35, 0.69. The data were twice as likely under the null hypothesis as the alternative, $BF_{10}=0.47$. There was clear evidence of heterogeneity, $\tau=0.49$, $BF_{rf}=$ Infinite. There was very strong evidence of publication bias, $BF_{pb}=26.40$.

336 Distance Effect

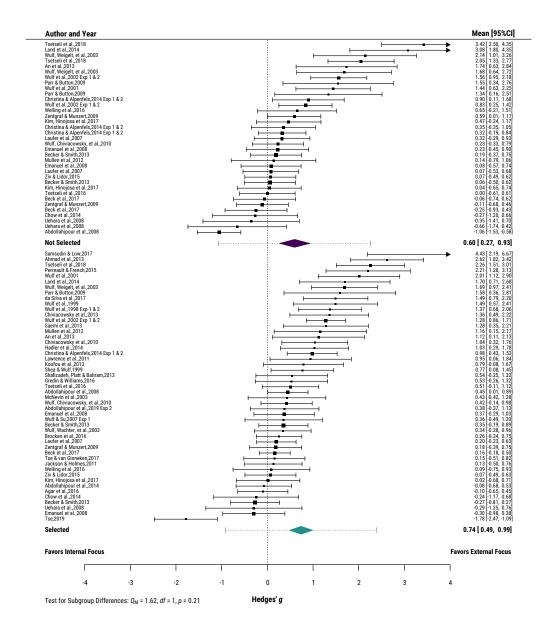
One possible outlier (Lohse et al., 2014) was identified in the distance effect meta-analysis. The mean of the model-averaged posterior distribution for the difference between distal and proximal external foci was g = -0.01, 95% credible interval: -0.38, 0.30. The results were over 3.8 times more likely under the null hypothesis than the alternative, $BF_{10} = 0.26$. There was clear evidence of heterogeneity, $\tau = 0.42$, $BF_{rf} = 25.58$. There was overwhelming evidence of publication bias, $BF_{pb} = 31.18$. Removing the influential case did not substantively change the conclusions: g = 0.06, 95% credible interval: 0, 0.32, $BF_{10} = 0.52$; $\tau = 0.42$, $BF_{rf} = 2.38$; $BF_{pb} = 2.97$.

345 Selection Moderator

Outcomes selected for inclusion in Chua and colleagues' (2021) meta-analysis of retention performance were somewhat larger (g = 0.74, 95% confidence interval: 0.49, 0.99) than the additional outcomes that could have been extracted but were not (g = 0.60, 95% confidence interval: 0.27, 0.93; see Figure 2). However, the difference between selected and not-selected outcomes was not statistically significant, F(1, 45) = 1.62, p = 0.21.

Figure 2

Forest plot of retention outcomes separated by "selected" moderator.



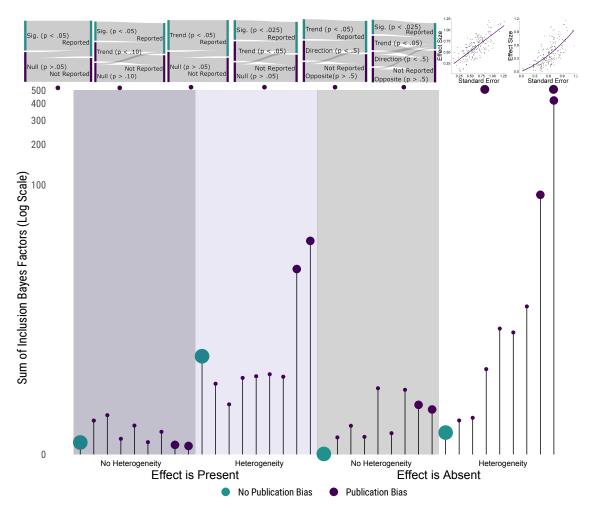
Note. Standardized mean difference (g) and 95% confidence intervals for each study included in the meta-analysis of retention outcomes. The green polygon represents the mean and 95% confidence interval for outcomes that Chua et al. (2021) selected for analysis. The purple polygon represents the estimate for outcomes reported in the original experiments but were not selected by Chua et al. (2021). The error bars extending from both polygons reflect the 95% prediction interval, illustrating the range of outcomes we would expect to observe in 95% of studies randomly sampled from the same population of studies included in this analysis. The prediction intervals account for the substantial unexplained heterogeneity present in these data, showing that even without correcting for publication bias we would expect outcomes across the entire plausible range of effects.

Individual Model Fit

As implied by the results of each analysis, the best performing models overall
assumed heterogeneity, publication bias, and zero effect (see Figure 3). The best fitting
publication bias models were the PET and PEESE funnel plot regression models, as well as
the selection models that assumed directional hypotheses, particularly those that modeled
censorship based on the direction of the point estimate. This pattern of findings suggests
complex, results-based selection mechanisms linked to more than just statistical
significance.

Figure 3

Total Inclusion Bayes Factor for each model relative to the ensemble, summed across each of the five analyses with and without outliers.



Note. Higher Inclusion Bayes Factors indicate better agreement with the data than the average of the ensemble. The green circles represent naïve fixed and random effects models that assume no publication bias. The purple circles represent six selection models and two regression models, each modeling publication bias in a different way. A figure illustrating each of the publication bias models is displayed below the lollipop plot, shown in the same left-to-right order they follow in the plot above. The size of each circle reflects the prior probability assigned to the model (p = 0.125 for naïve models, p = 0.031 for regression models, and p = 0.01 for selection models). The naïve and publication bias models were fit testing four scenarios: (a) an effect is present, no heterogeneity, (b) an effect is present, heterogeneity is present, (c) an effect is absent, no heterogeneity, and (d) an effect is absent, heterogeneity is present. The PEESE model, presented on the far right in each scenario, dominated the other models when assuming an effect is absent and heterogeneity is present. To better illustrate the performance of each model in the ensemble, Inclusion Bayes Factors are shown on a log scale on the y-axis.

359 Discussion

We re-evaluated the evidence in support of an external focus benefit for learning, 360 performance, muscular efficiency, and the distance effect. Seven previous meta-analyses 361 have relied on the results of naïve random effects models that assume zero reporting bias in 362 the primary estimates. However, it has become clear that such an assumption may not be 363 appropriate for motor learning research (McKay et al., 2023; McKay, Yantha, et al., 2022). 364 Each of the previous seven meta-analyses concluded that an external focus is superior to an 365 internal focus. Kim et al. (2017) reported the benefits applied to balance learning, 366 performance, and transfer. Makaruk et al. (2020) found the same for jump performance 367 and Li et al. (2022) reported similar results for sprint performance. Grgic and colleagues 368 reported external focus benefits for both muscular strength and endurance (Grgic et al., 2021; Grgic & Mikulic, 2022). Nicklas et al. (2022) reported the advantage of an external 370 focus over an internal focus applied to immediate performance in general. The most 371 comprehensive of the meta-analyses, and the study whose data we reanalyzed, was 372 conducted by Chua et al. (2021). They estimated small to moderate benefits for each 373 specific effect and concluded that no amount of publication bias could attenuate the 374 observed effects to zero. 375

Our results differ from previous meta-analyses as reporting bias was unaccounted for 376 in their primary estimates. This is a serious limitation of the previous external focus 377 meta-analyses as simulation studies have clearly demonstrated that random-effects models 378 result in large biases and high rates of false positives in the presence of publication-selection 379 bias (Barto, Maier, Shanks, et al., 2023; Bom & Rachinger, 2019; Stanley et al., 2017, 380 2022), which have been further supported when random-effects are compared with 381 preregistered multilab replications (Kvarven et al., 2020). We therefore explicitly modeled 382 bias; and consequently, estimated trivially small effects in each analysis. Although Chua 383 et al. (2021) concluded that no amount of publication bias could reduce the effects to the 384 null, our models suggest the data favor the null hypothesis for each analysis. If the only 385

type of reporting bias in the literature is one-sided selection at p=0.05, then Chua and 386 colleagues' conclusions were justified. However, if there were other considerations, such as 387 sample size, trends, and direction of point estimates, the assumptions of their model were 388 violated. Our analysis suggests this is the case for the focus of attention literature. Thus, 389 similar to previous simulation studies our findings illustrate that reporting bias can cause 390 random effects models to produce even large effect estimates when the true model is null. 391 The random effects estimates reported by Chua et al. (2021) ranged from small to large 392 and our corrected estimates range from essentially nil to trivial at most. 393

Although we observed somewhat larger estimates among effects selected by Chua et al. (2021) than among alternative outcomes that could have been selected, the difference was small and easily attributable to chance. The stronger signal for selection came from censorship prior to appearing in the published sample. Thus, the average reader of this literature would not have been inoculated against bias by having access to the complete results of each article. The biasing influence of censorship would have already affected the sample of information readers could access.

These findings underscore uncertainty about external focus benefits. Adding to this 401 uncertainty, we observed significant unexplained heterogeneity in effects. This 402 heterogeneity could imply that focus of attention has a range of effects that depend on 403 situational factors. If so, our results suggest that an internal focus may be superior to an 404 external focus in nearly as many situations as the reverse. Alternatively, this heterogeneity 405 may be due to methodological idiosyncrasies, unmodeled selection, or poor data curation at 406 any level. As with censorship mechanisms, we have no way to know which potential sources 407 of heterogeneity were at play. 408

Unfortunately, the present results add to a growing body of metascience questioning
the extant support for the predictions in Wulf and Lewthwaite's (2016) OPTIMAL theory
of motor learning (see McKay et al., 2023 for a recent meta-analysis on the other two
pillars in the theory). In addition to predicting external focus benefits for learning and

performance, OPTIMAL theory also predicts beneficial effects for autonomy and enhanced 413 expectancies via similar underlying mechanisms (Wulf & Lewthwaite, 2016). The primary 414 corpus of evidence supporting motor learning benefits from autonomy is the self-controlled 415 practice literature. Self-controlled practice involves asking learners to choose an aspect of 416 their practice environment and the published literature suggests this will confer noticeable 417 benefits to performance and learning (for a review see Ste-Marie et al., 2020). However, 418 like the external focus literature, the self-controlled practice research shows substantial 419 evidence of reporting bias and more support for the null hypothesis (McKay, Yantha, et al., 420 2022). Approximately the same pattern emerges for the enhanced expectancies research 421 (Bacelar et al., 2022). Although the published literature appears to unequivocally 422 demonstrate the predicted motor benefits of enhancing a learner's expectancy for success, 423 accounting for reporting bias suggests uncertainty and heterogeneity (McKay et al., 2023). Taken together, this meta-evidence suggests the underlying mechanism common to all 425 three factors of the tripartite OPTIMAL theory may be censorship. The mechanisms forwarded in OPTIMAL theory are made no less valid by this conclusion; it is the evidence 427 rather than the theory that has been impugned by this body of work.⁷ 428

429 Limitations

By reanalyzing the dataset from Chua et al. (2021), we inheritted the limitations of
their original study, including restricting the literature search to studies published in
English. Our results highlight the impact of reporting bias, which may be a systemic issue
with contributions from authors, reviewers, editors, journals, institutions, and funders. By
excluding articles not published in English, this dataset omits studies that have potentially
been reported from outside the Western cultural milieu, where the impact of systemic
pressures may be expressed differently. Further, because studies that have not been

⁷ This conclusion also applies to other theories and perspectives (e.g., ecological dynamics) that have been forwarded based on the extant attentional focus literature to account for a supposed external focus advantage (e.g., Davids et al., 2003; Gottwald et al., 2023; Hommel et al., 2001; Prinz, 1990; Wulf & Prinz, 2001).

published were omitted, and these studies may or may not have been subjected to the 437 scrutiny of peer-review and editorial discretion, this dataset does not represent a total 438 account of all studies ever conducted on external focus of attention and the impact of 439 systemic factors on reporting bias. Although the dataset is restricted to published studies 440 that have been peer-reviewed, and Chua et al. (2021) thoroughly probed the risk of bias 441 among those studies, such efforts are no guarantee on the accuracy of the underlying data 442 (see McKay & Carter, 2023). Caution is always warranted when interpreting the results of 443 a retrospective meta-analysis of non-Registered Reports, especially without individual participant data. 445

Overall, the evidence in the review contains small sample sizes and small to 446 moderate risk of bias according to Chua and colleagues (2021). None of the studies were preregistered. There were 20 studies missing due to insufficient information to calculate effect sizes in the original data set and another four missing effects from our extraction of secondary outcomes. We excluded as many as four outliers based on leverage statistics whereas Chua et al. (2021) removed as many as 18 outliers using a different criteria. We 451 recommend that, moving forward, researchers conducting meta-analyses in motor learning 452 and related areas (e.g., psychology, neuroscience, sport and exercise science) adopt the use 453 of leverage statistics as the default approach for identifying outliers (see Deeks et al., 2023; 454 Viechtbauer & Cheung, 2010 for discussions). 455

Lastly, we did not explore whether manipulation checks verified that the instructed attentional focus was adopted during performance. OPTIMAL theory predicts that when learners focus on their intended effect on the environment, they facilitate goal-action coupling, benefiting performance and learning. Our analysis only investigated whether instructions or feedback impacted performance. Perhaps a missing moderator in our analysis was the extent to which focus instructions were followed in each experiment. We chose not to explore this possibility because there are no validated manipulation checks.

Recommendations and Conclusions

The potential benefit of adopting an external focus of attention is among the most 464 important contributions of academic motor learning research. It fits with numerous 465 theoretical perspectives in the scientific literature and has been widely promoted in an 466 array of applied settings, including sports, rehabilitation, and education. Our findings 467 impugn the evidential basis for the superiority of an external focus of attention. However, 468 rather than establishing nil or trivial benefits from focusing externally, uncertainty 469 remains. The posteriors include interesting effects, there may be important moderators, 470 and our estimates may have overcorrected for bias. We simply do not know if an external 471 focus provides meaningful benefits to motor learning and performance or not. 472 Building knowledge about external focus effects can be accelerated by adoption of 473 the Registered Report publication format (Chambers, 2019). Registered Reports prevent 474 publication bias (Scheel et al., 2021), and when they include preregistration of analysis 475 plans, they prevent p-hacking (Simmons et al., 2011) and HARKing (Kerr, 1998) as well. 476 Limited resources may prevent individual laboratories from collecting sufficient sample 477 sizes for a well-powered Registered Report, so researchers are encouraged to collaborate 478 extensively to achieve the sample sizes necessary to make progress.

480 Author contributions (CRediT Taxonomy)

481 Conceptualization: BM, MJC

482 Data curation: BM, MJC

483 Formal analysis: BM

484 Funding acquisition: MJC

Investigation: BM, AEC, JS, CSDF, HH, KA, FCA, MJC

486 Methodology: BM, MJC

487 Project administration: BM, MJC

488 Software: BM, MJC

489 Supervision: MJC

Validation: BM, MJC

Visualization: BM, MJC

492 Writing – original draft: BM, AEC, JS, CSDF, HH, KA, FCA, MJC

Writing – review & editing: BM, AEC, JS, CSDF, HH, KA, FCA, MJC

Data availability statement

The data and scripts can be accessed using either of the following links:

https://osf.io/vfmx2/ or https://github.com/cartermaclab/proj_foa-optimal-theory

Declaration of interest statement

498 All authors declare no conflicts of interest.

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