

## Fisher-Neyman Factorization Lemma

$T$  is sufficient for  $\theta$  iff there are functions

$$g \text{ & } h \text{ such that } L(\theta | x_1, \dots, x_n) = g(T(x_1, \dots, x_n), \theta) \cdot h(x_1, \dots, x_n)$$

Example: let  $n$  iid  $X_i \sim \text{Poisson}(\lambda)$  be our sample

$$\begin{aligned} f_{X_i}(x) &= \frac{e^{-\lambda} \lambda^x}{x!}, \quad L(\lambda | x_1, \dots, x_n) = \frac{e^{-\lambda} \lambda^{x_1}}{x_1!} \cdot \dots \cdot \frac{e^{-\lambda} \lambda^{x_n}}{x_n!} \\ &= \frac{e^{-n\lambda} \lambda^{x_1 + \dots + x_n}}{x_1! \cdot \dots \cdot x_n!} = \underbrace{e^{-n\lambda} \lambda^{T}}_{g(T, \lambda)} \cdot \underbrace{\frac{1}{x_1! \cdot \dots \cdot x_n!}}_{h(x_1, \dots, x_n)} \end{aligned}$$

thus  $T = x_1 + \dots + x_n$  is a sufficient statistic for  $\lambda$

Example: let  $n$  iid  $X_i \sim N(\mu, \sigma^2=1)$  be our sample

\* We want to find a sufficient statistic for  $\mu$ .

$$\begin{aligned} L(\mu | x_1, \dots, x_n) &= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x_1 - \mu)^2} \cdot \dots \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x_n - \mu)^2} \\ &= \frac{1}{(2\pi)^{\frac{n}{2}}} e^{-\frac{1}{2} \sum_{i=1}^n (x_i - \mu)^2} = \frac{1}{(2\pi)^{\frac{n}{2}}} e^{-\frac{1}{2} \sum_{i=1}^n x_i^2 - 2\sum_{i=1}^n x_i \mu + n\mu^2} \\ &= \frac{1}{(2\pi)^{\frac{n}{2}}} \cdot e^{-\frac{1}{2} \left( \sum_{i=1}^n x_i^2 - 2\mu \sum_{i=1}^n x_i + n\mu^2 \right)} \\ &= \frac{1}{(2\pi)^{\frac{n}{2}}} \cdot e^{-\frac{1}{2} \left( n\mu^2 - 2\mu \sum_{i=1}^n x_i \right)} \cdot e^{-\frac{1}{2} \sum_{i=1}^n x_i^2} \end{aligned}$$

$$= \underbrace{\frac{1}{(2\pi)^{\frac{n}{2}}} \cdot e^{-\frac{1}{2} \sum_{i=1}^n x_i^2}}_{g(T, \mu)} \cdot e^{-\frac{1}{2} \sum_{i=1}^n x_i^2} h(x_1, \dots, x_n)$$

So  $T = \sum_{i=1}^n x_i$  is a

sufficient statistic for  $\mu$ .

We can also say that

$\bar{X} = \frac{\sum_{i=1}^n x_i}{n}$  is a sufficient statistic for  $\mu$ .

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Example: Let  $n$  iid  $X_i \sim N(\mu=0, \sigma^2)$  be our random sample.

Find a sufficient statistic for  $\sigma^2$

$$\begin{aligned} L(\sigma^2 | X_1, \dots, X_n) &= \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2\sigma^2} X_1^2} \cdot \dots \cdot \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2\sigma^2} X_n^2} \\ &= \underbrace{\frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} e^{-\frac{1}{2\sigma^2} (X_1^2 + \dots + X_n^2)}}_{g(T, \sigma^2)} \cdot \underbrace{1}_{h(X_1, \dots, X_n)} \end{aligned}$$

so  $T = \sum_{i=1}^n X_i^2$  is a sufficient statistic for  $\sigma^2$

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Example: Let  $n$  iid  $X_i \sim N(\mu, \sigma^2)$  be our sample

We want to find a sufficient statistic for  $(\mu, \sigma^2)$

$$L(\mu, \sigma^2 | X_1, \dots, X_n) = \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{1}{2\sigma^2} (X_1 - \mu)^2} \cdot \dots \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{1}{2\sigma^2} (X_n - \mu)^2}$$

$$\begin{aligned}
 L((\mu, \sigma^2) | X_1, \dots, X_n) &= \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{-1}{2\sigma^2}(X_i - \mu)^2} \cdot \dots \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{-1}{2\sigma^2}(X_n - \mu)^2} \\
 &= \frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} \cdot e^{\frac{-1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2} = \frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} \cdot e^{\frac{-1}{2\sigma^2} \sum_{i=1}^n (X_i^2 - 2X_i\mu + \mu^2)} \\
 &= \underbrace{\frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} \cdot e^{\frac{-1}{2\sigma^2} \left( \sum_{i=1}^n X_i^2 - 2\mu \sum_{i=1}^n X_i + n\mu^2 \right)}}_{g(T_1, T_2) ; (\mu, \sigma^2)} \cdot \underbrace{e^{\frac{-1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2}}_{h(X_1, \dots, X_n)}
 \end{aligned}$$

So  $(T_1, T_2) = \left( \sum_{i=1}^n X_i, \sum_{i=1}^n X_i^2 \right)$  is

jointly sufficient for  $(\mu, \sigma^2)$

We also could've solved this problem from a 2<sup>nd</sup> approach

$$\begin{aligned}
 L((\mu, \sigma^2) | X_1, \dots, X_n) &= \frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} \cdot e^{\frac{-1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2} \\
 &= \frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} \cdot e^{\frac{-1}{2\sigma^2} \sum_{i=1}^n ((X_i - \bar{X}) + (\bar{X} - \mu))^2}
 \end{aligned}$$

$$\sum_{i=1}^n (X_i - \bar{X})^2 + \underbrace{\sum_{i=1}^n 2(X_i - \bar{X})(\bar{X} - \mu)}_{=0} + \sum_{i=1}^n (\bar{X} - \mu)^2$$

$$\text{Therefore } \frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} \cdot e^{\frac{-1}{2\sigma^2} \left( \sum_{i=1}^n (X_i - \bar{X})^2 + \sum_{i=1}^n (\bar{X} - \mu)^2 \right)}$$

$$L((\mu, \sigma^2) | X_1, \dots, X_n) = \underbrace{\frac{1}{\sigma^n (2\pi)^{\frac{n}{2}}} \cdot e^{\frac{-1}{2\sigma^2} \left( \sum_{i=1}^n (X_i - \bar{X})^2 + \sum_{i=1}^n (X_i - \mu)^2 \right)}}_{g(T_3, T_4, (\mu, \sigma^2))} \cdot h(X_1, \dots, X_n)$$

note:  $\sum_{i=1}^n (\bar{X} - \mu)^2 = n(\bar{X} - \mu)^2$

So the pair  $(T_3 = \bar{X}, T_4 = \sum_{i=1}^n (X_i - \bar{X})^2)$  are also jointly sufficient for  $(\mu, \sigma^2)$

Note here  $T_4 = \frac{s^2}{n-1}$  so we can also say

that  $(T_5 = \bar{X}, T_6 = s^2)$  are also jointly sufficient for  $(\mu, \sigma^2)$

Note: we can usually derive one pair from another

Note: Sufficiency is about information, and when we have enough information to find our parameters

Example

Let  $n$  iid  $U_i \sim \text{uniform}[0, \vartheta]$  be our sample

find a sufficient statistic for  $\vartheta$   $\left\{ f(x, \vartheta) = \frac{1}{\vartheta} \text{ for } 0 \leq x \leq \vartheta \right\}$

$$L(\vartheta | U_1, \dots, U_n) = \frac{1}{\vartheta^n} \text{ for } 0 \leq U_i \leq \vartheta$$

$$g(T, \vartheta) = \frac{1}{\vartheta^n} \quad \ell \quad h(X_1, \dots, X_n) = 1$$

$$g(\tau, \theta) = \bar{\theta}^n \notin h(x_1, \dots, x_n) = 1$$

So  $T = \max\{u_1, \dots, u_n\}$  is a sufficient statistic for  $\theta$

## What makes the Fisher-Neyman Lemma True?

$$\text{Suppose } L(\theta | x_1, \dots, x_n) = g(s(x_1, \dots, x_n), \theta) \cdot h(x_1, \dots, x_n)$$

$$\begin{aligned} & f_{(x_1, \dots, x_n | s)}(x_1, \dots, x_n | S = s(x_1, \dots, x_n)) \\ &= \frac{f_{(x_1, \dots, x_n)}(x_1, \dots, x_n)}{f_s(s)} = \frac{\underset{\text{constant}}{\boxed{g(s(x_1, \dots, x_n))}} \cdot h(x_1, \dots, x_n)}{\int g(s(x_1, \dots, x_n)) \cdot h(x_1, \dots, x_n) dx_1 \dots dx_n} \end{aligned}$$

So if we can  
factor it implies  
sufficiency

$$s(x_1, \dots, x_n) = S$$

$$h(x_1, \dots, x_n) \text{ doesn't depend on } \theta$$

Why does sufficiency imply factorization?

$f_{x_1, \dots, x_n | T}(x_1, \dots, x_n | T(x_1, \dots, x_n))$  does not depend on  $T$  by the definition

$$= \frac{f_{x_1, \dots, x_n}(x_1, \dots, x_n)}{f_T(T(x_1, \dots, x_n))} = h(x_1, \dots, x_n)$$

$$L(\vartheta | x_1, \dots, x_n) = \underbrace{f_T(T(x_1, \dots, x_n))}_{g(T, \vartheta)} \cdot \underbrace{h(x_1, \dots, x_n)}_{h(x_1, \dots, x_n)}$$