Appendix S3. Compute model averaged parameter estimates via LOOIC()

2020 - 2021 Skagit River steelhead forecast.

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This is version 0.20.12.09.

1 Background

This appendix shows how generate model averaged parameter estimates and generate figures relevant to the 2020-2021 wild Skagit River steelhead forecast. All analyses require the R software (v3.5 or later), as well as a few packages that are not included with the base installation of R.

```
if(!require("readr")) {
   install.packages("readr")
   library("readr")
}
if(!require("captioner")) {
   devtools::install_github("adletaw/captioner")
   library("captioner")
}
if(!require("coda")) {
```

```
install.packages("coda")
  library("coda")
}
if(!require("here")) {
  install.packages("here")
  library("here")
}
if(!require("gsl")) {
  install.packages("gsl")
  library("gsl")
}
if(!require("loo")) {
  install.packages("loo")
  library("loo")
}
## set default caption delimter
fig_cap <- captioner(infix = ".")</pre>
## set directory locations
datadir <- here("data")</pre>
analdir <- here("analysis")</pre>
savedir <- here("analysis/cache")</pre>
## better round/floor/ceiling
around <- function(x, func = "round", prec = 1) {</pre>
  ## 'func' can be "round", "floor", or "ceiling"
  ## 'prec' is desired precision (eg, 0.1 is to nearest tenth)
  if(!is.double(x)) {
    stop("'x' must be a real number")
  if(!(func %in% c("round", "floor", "ceiling"))) {
    stop("'func' must be \"round\", \"floor\", or \"ceiling\"")
  }
  if(prec <= 0) {
    stop("'prec' cannot be less than or equal to 0")
 do.call(func, list(x / prec)) * prec
}
```

2 Load the information

Here we load in the estimated parameters and states from the selected model, as well as the covariates and harvest data and escapement data.

```
fit_bh_cov_MA1_AR1 <- readRDS(file.path(savedir, "fit_bh_cov_MA1_AR1.rds"))
fit_bh_cov_AR1 <- readRDS(file.path(savedir, "fit_bh_cov_AR.rds"))</pre>
```

```
## covariate(s)
dat_cvrs <- read_csv(file.path(datadir, "skagit_sthd_covars.csv"))
## total number of covariates
n_cov <- dim(dat_cvrs)[2] - 1

## escapement
dat_esc <- read_csv(file.path(datadir, "skagit_sthd_esc.csv"))
## log of escapement
ln_dat_esc <- c(log(dat_esc$escapement), rep(NA, n_fore))

## harvest
dat_harv <- read_csv(file.path(datadir, "skagit_sthd_catch.csv"))
## drop year col & first age_max rows
dat_harv <- c(dat_harv$catch, rep(0, n_fore))</pre>
```

3 Main results

Call models and compute LOOIC Via loo() and compare() with full table of results. Note that elpd_diff will be negative (positive) if the expected predictive accuracy for the first (second) model is higher. We estimate pseudo Bayesian model weights for the purposes of model averaging. We also need to convert the mcmc.list output into a more user-friendly form for plotting, etc.

```
## results
n_{mods} < -2
mod_res_MA1_AR1 <- do.call("rbind", fit_bh_cov_MA1_AR1)</pre>
mod_res_AR1 <- do.call("rbind", fit_bh_cov_AR1)</pre>
mod_fits <- list(fit_bh_cov_MA1_AR1,fit_bh_cov_AR1)</pre>
LOOIC <- vector("list", n_mods)
## extract log densities from JAGS objects
for(i in 1:n_mods) {
  ## convert mcmc.list to matrix
  tmp_lp <- as.matrix(mod_fits[[i]])</pre>
  ## extract pointwise likelihoods
  tmp_lp <- tmp_lp[,grepl("lp_", colnames(tmp_lp))]</pre>
  ## if numerical underflows, convert -Inf to 5% less than min(likelihood)
  if(any(is.infinite(tmp_lp))) {
    tmp_lp[is.infinite(tmp_lp)] <- NA</pre>
    tmp_min <- min(tmp_lp, na.rm = TRUE)</pre>
    tmp_lp[is.na(tmp_lp)] <- tmp_min * 1.05</pre>
  }
  ## calculate LOOIC
  LOOIC[[i]] <- loo(tmp_lp)
}
```

```
## compute pseudo weights
model_weights <- loo_model_weights(LOOIC, method = "pseudobma",optim_method = "BFGS", optim_co
## LOOIC for all data
tbl_LOOIC <- round(loo_compare(x = LOOIC), 2)
rownames(tbl_L00IC) <- sub("model", "", rownames(tbl_L00IC))</pre>
tbl_LOOIC <- tbl_LOOIC[order(as.numeric(rownames(tbl_LOOIC))), ]</pre>
tbl_LOOIC \leftarrow cbind(model = c("B-H", "B-H"),
                   error = c("MA1\_AR1","AR1"),
                   as.data.frame(tbl_L00IC),pseudo_bma_weight = as.matrix(model_weights))
tbl_LOOIC[order(tbl_LOOIC[,"looic"]), ]
##
     model
             error elpd_diff se_diff elpd_loo se_elpd_loo p_loo se_p_loo looic se_looic
       В-Н
               AR1
                         0.00
## 2
                                 0.00 -401.95
                                                      49.27 146.41
                                                                       11.48 803.90
                                                                                        98.54
## 1
       B-H MA1_AR1
                       -16.07
                                10.16 -418.02
                                                      51.57 161.98
                                                                       18.23 836.04
                                                                                       103.15
     pseudo_bma_weight
##
## 2
           0.992673609
## 1
           0.007326391
```

3.1 Total population size

#yp_min <- min(p_dat)</pre>

Here is our estimate of the total run size (i.e., catch + escapement) over time. The black points are the data, the blue line is the median posterior estimate, and the shaded region is the 95% credible interval. Note that the y-axis is on a log scale.

```
clr <- rgb(0, 0, 255, alpha = 50, maxColorValue = 255)
## estimated terminal abundance forecast
p_dat <- mod_res_MA1_AR1[,grep("Sp", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_AR
p_dat <- apply(p_dat, 2, quantile, CI_vec)
p_dat <- p_dat + matrix(dat_harv, length(CI_vec), n_yrs+n_fore, byrow = TRUE)

# p_dat_AR1MA1 <- mod_res_MA1_AR1[,grep("Sp", colnames(mod_res_MA1_AR1))]

# p_dat_AR1MA1 <- apply(p_dat_AR1MA1, 2, quantile, CI_vec)
# p_dat_AR1MA1 <- p_dat_AR1MA1 + matrix(dat_harv, length(CI_vec), n_yrs+n_fore, byrow = TRUE)

# p_dat_AR1 <- mod_res_AR1[,grep("Sp", colnames(mod_res_AR1))]

# p_dat_AR1 <- apply(p_dat_AR1, 2, quantile, CI_vec)
# p_dat_AR1 <- apply(p_dat_AR1, 2, quantile, CI_vec)
# p_dat_AR1 <- p_dat_AR1 + matrix(dat_harv, length(CI_vec), n_yrs+n_fore, byrow = TRUE)

## time seq
t_idx_f <- seq(yr_frst, length.out = n_yrs+n_fore)
## plot</pre>
```

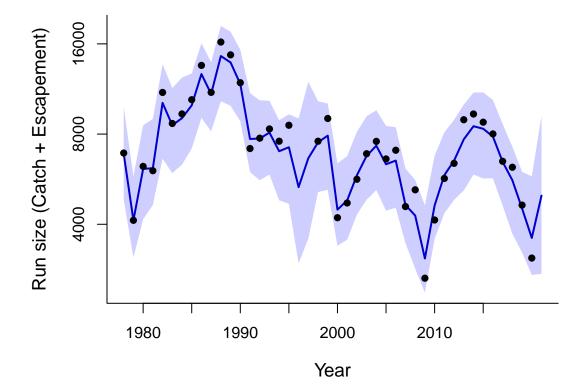


Figure 1: Time series of the estimated total population size (catch plus the adults that escaped to spawn). The observed data are the points; the solid line is the median estimate and the shaded region indicates the 95% credible interval.

3.2 2021 terminal run size forecast

Here are several percentiles for the 2021 forecast for the total run size (i.e., catch + escapement).

```
## forecast
## 2.5% 2846
## 25% 3859
## 50% 4985
## 75% 7101
## 97.5% 9005
```

3.3 Spawner-recruit relationship

xM <- around(max(s_dat), "ceiling", dd)</pre>

par(mai = c(0.8, 0.8, 0.1, 0.1), omi = c(0, 0, 0, 0))

Here is the relationship between spawner and subsequent recruits (a), assuming mean values for all covariates. Gray lines show the median relationship for each of the 43 years based on a_t . Note that for plotting purposes only in (b) and (c), the density in the largest bin for each parameter contains counts for all values greater or equal to that. Vertical arrows under the x-axes in (b) and (c) indicate the 2.5^{th} , 50^{th} , and 97.5^{th} percentiles.

```
layout(matrix(c(1,1,2,3),2,2),c(3,2),c(1,1))
xoffSet <- 0.05
yoffSet <- 0.03
## colors for plotting
clr \leftarrow rgb(100, 0, 200,
                                  alpha = seq(200, 100,
                                                                        length.out = age_max-age_min+n_fore),
                                  maxColorValue = 255)
## posterior of spawners
s_dat <- mod_res_MA1_AR1[,grep("Sp", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_AR
s_dat <- apply(s_dat, 2, quantile, CI_vec)</pre>
s_dat <- s_dat[, 1:(n_yrs-age_min+n_fore)]</pre>
## posterior of recruits
r_dat <- mod_res_MA1_AR1[,grep("tot_ln_Rec", colnames(mod_res_MA1_AR1))]*model_weights[1] + model_weights[1] + model_weights[1]
r_dat <- exp(apply(r_dat, 2, quantile, CI_vec))</pre>
## median values for a & b
aa <- mod_res_MA1_AR1[,grep("ln_BH_a", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_.
aa <- apply(aa, 2, median)
bb <- mod_res_MA1_AR1[,grep("beta", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_AR1
bb <- median(bb)</pre>
## empty plot space for spawner-recruit relationships
dd <- 3000
yM <- around(max(r_dat), "ceiling", dd)</pre>
```

```
plot(s_dat[2,], r_dat[2,], xlim = c(0,xM), ylim = c(0,yM), type = "n",
     xaxs = "i", yaxs = "i", cex.lab = 1.2,
     xlab = expression(Spawners~(10^3)),
     ylab = expression(Recruits~(10^3)),
     xaxt = "n", yaxt = "n", bty="L")
axis(1, at = seq(0,xM,dd*2), labels = seq(0,xM,dd*2)/1000)
axis(2, at = seq(0,yM,dd*2), labels = seq(0,yM,dd*2)/1000, las=1)
for(i in 1:length(aa)) {
  lines(exp(aa[i]) * seq(0,xM) / (1 + bb * seq(0,xM)),
        col = "darkgray")
abline(a = 0,b = 1,lty = "dashed")
## add S-R estimates and medians
nCB <- n_yrs-age_max
## years with complete returns
points(s_dat[2, 1:nCB], r_dat[2, 1:nCB],
       xlim = c(0,xM), ylim = c(0,yM),
       pch = 16, col = "blue3")
segments(s_dat[2, 1:nCB], r_dat[1, 1:nCB],
         s_dat[2, 1:nCB], r_dat[3, 1:nCB],
         col = "blue3")
segments(s_dat[1, 1:nCB], r_dat[2, 1:nCB],
         s_dat[3, 1:nCB], r_dat[2, 1:nCB],
         col = "blue3")
nTB <- dim(s_dat)[2]
## years with incomplete returns
segments(s_dat[2, (nCB+1):nTB], r_dat[1, (nCB+1):nTB],
         s_dat[2, (nCB+1):nTB], r_dat[3, (nCB+1):nTB],
         col = clr)
segments(s_dat[1, (nCB+1):nTB], r_dat[2, (nCB+1):nTB],
         s_dat[3, (nCB+1):nTB], r_dat[2, (nCB+1):nTB],
         col = clr)
points(s_dat[2, (nCB+1):nTB],r_dat[2, (nCB+1):nTB],
       xlim = c(0,xM), ylim = c(0,yM),
       pch = 16, col = clr)
text(x = par()$usr[1] + diff(par()$usr[1:2]) * xoffSet,
     y = par() usr[4] - diff(par() usr[3:4]) * yoffSet,
     "(a)")
## posterior for alpha
clr \leftarrow rgb(0, 0, 255, alpha = 50, maxColorValue = 255)
a_thresh <- 59
par(mai = c(0.8, 0.4, 0.3, 0.1))
## B-H alpha
R_alpha_est <- mod_res_MA1_AR1[,grep("alpha", colnames(mod_res_MA1_AR1))]*model_weights[1] + m
alphaCI <- quantile(R_alpha_est, CI_vec)</pre>
R_alpha_est[R_alpha_est > a_thresh] <- a_thresh</pre>
```

```
hist(R_alpha_est, freq = FALSE, breaks = seq(0, a_thresh+1, 2),
     col = clr, border = "blue3",
     xlab = "", ylab = "", main = "", cex.lab = 1.2, yaxt = "n")
aHt <- (par()$usr[4]-par()$usr[3])/12
arrows(alphaCI, par()$usr[3], alphaCI,par()$usr[3]-aHt,
       code = 1, length = 0.05, xpd = NA, col = "blue3", lwd = 1.5)
mtext(expression(Instrinsic~productivity~(alpha)), 1, line = 3, cex = 1)
text(x = par() usr[1],
     y = par() usr[4] * 1.05,
     "(b)", xpd=NA)
## posterior for K
par(mai = c(0.8, 0.4, 0.3, 0.1))
aa <- mod_res_MA1_AR1[,grep("alpha", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_AR
bb <- mod_res_MA1_AR1[,grep("beta", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_AR1
## K in 1000s
R_b_est <- (aa-1) / bb / 1000
R_b_est <- R_b_est[R_b_est > 0]
R_b_CI <- quantile(R_b_est, CI_vec)</pre>
## pile into last ban for plotting
R_b_{est}[R_b_{est} > 13] <- 13
brks <- seq(around(min(R_b_est), "floor"),</pre>
            around(max(R_b_est), "ceiling"),
            length.out = length(seq(0, a_thresh, 2)))
hist(R_b_est, freq = FALSE, breaks = brks, col = clr, border = "blue3",
     xlab = "", xaxt = "n", yaxt = "n",
     main = "", ylab = "", cex.lab = 1.2)
axis(1, at = seq(around(min(R_b_est), "floor"),
                 around(max(R_b_est), "ceiling"),
                 2))
aHt <- (par()$usr[4] - par()$usr[3]) / 12
arrows(R_b_CI, par()$usr[3], R_b_CI,par()$usr[3]-aHt,
       code = 1, length = 0.05, xpd = NA, col = "blue3", lwd = 1.5)
mtext(expression(paste("Carrying capacity (",italic(K),", ",10^3,")")),
      side = 1, line = 3, cex = 1)
text(x = par() usr[1],
     y = par() usr[4] * 1.05,
     "(c)", xpd=NA)
```

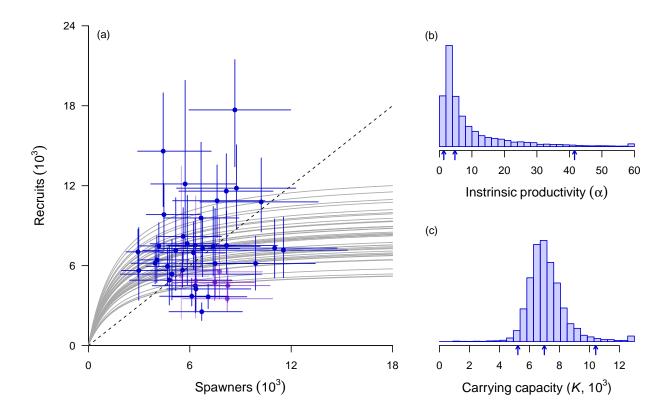


Figure 2: Relationship between the number of spawning adults and their subsequent surviving offspring (recruits), assuming mean values for all covariates (a); and the estimated posterior distributions for the intrinsic productivity (b) and carrying capacity (c). Points in (a) are medians of the posterior estimates; error bars indicate the 95% credible intervals. Blue points are for estimates with complete broods; purple points are for the most recent years with incomplete broods. Gray lines show the median relationship for each of the 41 years in the time series based on annual model estimates of productivity. Note that for plotting purposes only in (b) and (c), the density in the largest bin for each parameter contains counts for all values greater than or equal to it. Vertical arrows under the x-axes in (b) and (c) indicate the 2.5th, 50th, and 97.5th percentiles.

Here are summaries of the posterior distributions for α and K.

```
## intrinsic productivity
round(alphaCI, 2)

## 2.5% 50% 97.5%

## 1.35 4.78 41.59

## carrying capacity
round(R_b_CI, 2)

## 2.5% 50% 97.5%

## 5.23 7.00 10.42
```

3.4 Covariate effects

Here are time series plots of the covariates (a-c) and histograms of their effects on productivity (d-f).

```
clr \leftarrow rgb(0, 0, 255, alpha = 50, maxColorValue = 255)
xoffSet <- 0.04
yoffSet <- 0.03
par(mfrow=c(n_cov, 2), mai=c(0.4, 0.2, 0.1, 0.1), omi=c(0.2, 0.5, 0, 0))
c_est <- mod_res_MA1_AR1[,grep("gamma", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_</pre>
ylN \leftarrow floor(min(c_est)*10)/10
ylM <- ceiling(max(c_est)*10)/10</pre>
brks <- seq(ylN,ylM,length.out=diff(c(ylN,ylM))*40+1)
t_idx <- seq(yr_frst,length.out=n_yrs-age_min+n_fore)</pre>
dat_cvrs <- as.matrix(dat_cvrs[seq(length(t_idx)),])</pre>
for(i in 1:n_cov) {
  if(i==4) {
    dat_cvrs[,i+1] <- dat_cvrs[,i+1]/1000</pre>
  ## plot covar ts
 plot(dat_cvrs[, "year"], dat_cvrs[, i+1],
       pch = 16, col = "blue3", type = "o",
       xlab = "", ylab = "", main = "", bty = "L",
       cex.axis = 1.2)
  text(x = par()$usr[1] + diff(par()$usr[1:2]) * xoffSet,
       y = par() usr[4] - diff(par() usr[3:4]) * yoffSet,
       paste0("(",letters[i],")"),
       cex = 1.2
 mtext(side = 2, cov_names[i], line = 3, cex = 1.2)
  if(i == n_cov) {
    mtext(side = 1, "Brood year", line = 3)
  }
  ## plot covar effect
 hist(c_est[,i],
       freq = FALSE, breaks = brks, col = clr, border = blue3,
       xlab = "", yaxt = "n", main = "", ylab = "", cex.axis = 1.2)
  c_CI <- quantile(c_est[,i],CI_vec)</pre>
  aHt <- (par()$usr[4]-par()$usr[3])/20
  arrows(c_CI, par()$usr[3]-0.005, c_CI, par()$usr[3] - aHt,
         code = 1, length = 0.05, xpd = NA, col = "blue3", lwd = 1.5)
  abline(v = 0, lty = "dashed")
  text(x = par()$usr[1] + diff(par()$usr[1:2]) * xoffSet,
       y = par() usr[4] - diff(par() usr[3:4]) * yoffSet,
       paste0("(",letters[i+n_cov],")"),
```

```
cex = 1.2)
  if(i == n_cov) { mtext(side = 1, "Effect size", line = 3) }
}
      4000
                                                                     (d)
         _ (a)
Max flow (m^3 s^{-1})
      3000
      2000
      1000
                                                                                1 1
-0.2
                                                                                         0.0
                                     2000
                                                                        -0.4
                         1990
                                                 2010
                                                                                                  0.2
             1980
                                                                                                          0.4
                                                                                                                  0.6
           (b)
                                                                     (e)
      220
      200
Min \ flow \ (m^3 \ s^{-1})
      180
      160
      140
                                                                        -0.4
                                                                                -0.2
                                                                                         0.0
                                                                                               0.2
                                                 2010
             1980
                         1990
                                     2000
                                                                                                          0.4
                                                                                                                  0.6
                                                                     (f)
         (c)
      0
 NPGO
      0
      7
      7
                                                                        -0.4
                                                                                                                  0.6
                                                                                                  0.2
                                                                                                          0.4
                                                                                         0.0
                                     2000
                                                 2010
                                                                                -0.2
             1980
                         1990
                                                                                         Effect size
                               Brood year
```

Figure 3: Time series of the environmental covariates used in the model (a-d), and their estimated effects on population productivity (e-g). Small arrows under histograms denote 2.5th, 50th, and 97.5th percentiles of the posterior distribution.

Here is a summary of the covariate effect sizes

```
gamma_CI <- apply(c_est, 2, quantile, c(2.5, 5, 50, 95, 97.5)/100)
t(round(gamma_CI, 2))

## 2.5% 5% 50% 95% 97.5%

## gamma[1] -0.23 -0.21 -0.10 0.00 0.02

## gamma[2] -0.15 -0.13 -0.01 0.10 0.12

## gamma[3] -0.01 0.02 0.19 0.35 0.38
```

3.5 Process errors

Here is the time series of the residuals from the process model. They represent the population's productivity after accounting for the effects of density dependence and environmental covariates.

```
## time sequence
t_idx_a <- seq(yr_frst, length.out = n_yrs-age_min+n_fore)</pre>
## plot data
p_dat <- mod_res_MA1_AR1[,grep("res_ln_Rec", colnames(mod_res_MA1_AR1))]*model_weights[1] + model_weights[1] + model_weights[1]
p_dat <- apply(p_dat, 2, quantile, CI_vec)</pre>
yp_min <- min(p_dat)</pre>
yp_max <- max(p_dat)</pre>
## plot
par(mai = c(0.8, 0.8, 0.1, 0.1), omi = c(0, 0.2, 0.1, 0.2))
plot(t_idx_a, p_dat[3,],
                    type = "n", bty = "L",
                    ylim = c(yp_min,yp_max),
                    xlab = "Brood year", ylab = "Process error", main = "",
                    cex.lab = 1.2
abline(h = 0, lty = "dashed")
polygon(c(t_idx_a, rev(t_idx_a)), c(p_dat[3,], rev(p_dat[1,])),
                                 col = clr, border = NA)
lines(t_idx_a, p_dat[2,], col = "blue3", lwd = 2)
```

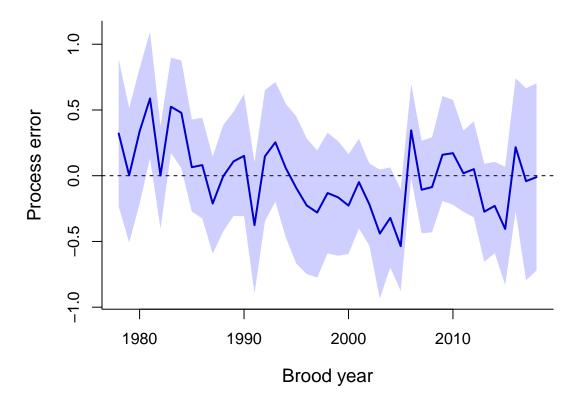


Figure 4: Time series of the estimated process errors, which represent the population's productivity after accounting for the effects of density dependence and environmental covariates. The solid line is the median estimate and the shaded region indicates the 95% credible interval.

3.6 Management reference points

colnames(ref_pts) <- ref_names</pre>

spawner series for optimal yield profile

Here are a number of management reference points.

```
## abbreviations for ref points
ref_names <- c("MSY", "Smsy", "Umsy", "Umax")
## proportions of MSY to consider
yld_prop <- c(0.75, 0.85, 0.95)
## median values for a & b
aa <- mod_res_MA1_AR1[,grep("E_BH_a", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_Ai
alpha <- exp(aa)
mcmc <- length(aa)
beta <- mod_res_MA1_AR1[,grep("beta", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod_res_Ai
## empty matrix for ref pts
ref_pts <- matrix(NA, mcmc, length(ref_names))</pre>
```

```
SS \leftarrow seq(100, 1e4, 100)
## empty matrix for optimal yield profiles
OYP <- matrix(0, length(SS), length(yld_prop))</pre>
for(i in 1:mcmc) {
    ## spawners at MSY
    ref_pts[i, "Smsy"] <- (alpha[i] / beta[i]) * sqrt(1 / alpha[i]) - (1 / beta[i])</pre>
    ref_pts[i, "MSY"] <- (ref_pts[i, "Smsy"] * alpha[i]) /</pre>
                              (1 + beta[i] * ref_pts[i, "Smsy"]) - ref_pts[i, "Smsy"]
    ## harvest rate at MSY
    ref_pts[i, "Umsy"] <- 1 - sqrt(1 / alpha[i])</pre>
    ## max harvest rate
    ref_pts[i, "Umax"] <- 1 - 1/alpha[i]</pre>
    ## yield over varying S
    yield <- ((SS * alpha[i]) / (1 + beta[i] * SS)) - SS</pre>
    for(j in 1:length(yld_prop)) {
        OYP[,j] <- OYP[,j] + 1*(yield > yld_prop[j] * ref_pts[i, "MSY"])
}
OYP <- OYP/mcmc
## Prob of overfishing
hh <- seq(100)
Pr_over <- cbind(hh,hh,hh)</pre>
colnames(Pr_over) <- c("Umsy75","Umsy","Umax")</pre>
for(i in hh) {
  Pr_over[i,"Umsy75"] <- sum(ref_pts[,"Umsy"] * 0.75 < i/100)/mcmc</pre>
 Pr_over[i,"Umsy"] <- sum(ref_pts[,"Umsy"] < i/100)/mcmc</pre>
 Pr_over[i,"Umax"] <- sum(ref_pts[,"Umax"] < i/100)/mcmc</pre>
}
## posterior exploitation rate & spawner abundance
aer <- Sp_ts <- mod_res_MA1_AR1[,grep("Sp", colnames(mod_res_MA1_AR1))]*model_weights[1] + mod
aer <- aer[,1:n yrs]</pre>
for(i in 1:n_yrs) {
    aer[,i] <- dat_harv[i] / (dat_harv[i] + Sp_ts[,i])</pre>
}
layout(matrix(c(2, 1, 4, 3), 2, 2), heights = c(1, 5))
yoffSet <- 0.10
yoffSet <- 0.05
## (a) Optimal yield profile
par(mai=c(0.9, 0.9, 0, 0), omi=c(0, 0, 0.1, 0.1))
x_lp <- yld_prop</pre>
for(i in 1:length(x_lp)) {
    x_{p[i]} \leftarrow SS[max(which(OYP[,i] == max(OYP[,i])
```

```
| abs(OYP[,i] - (yld_prop[i]-0.3)) \le 0.05))]
}
matplot(SS, OYP, type="l", lty="solid", ylim=c(0,1),
        col=c("slateblue","blue","darkblue"), lwd=2,
        xlab = "Spawners", ylab = "Probability of X% of MSY", main = "",
        las=1, cex.lab=1.2)
points(x = x_1p, y = yld_prop-0.3,
       pch = 21, cex = 3.5,
       col = "white", bg = "white")
text(x = x_lp, y = yld_prop-0.3, paste0(yld_prop*100, "%"),
     col=c("slateblue","blue","darkblue"), cex=0.7)
text(x = par()$usr[1] + xoffSet * diff(par()$usr[1:2]),
     y = par() susr[4] - yoffSet * diff(par() susr[3:4]),
     "(a)")
## marginal histogram of posterior spawner abundances
par(mai=c(0, 0.9, 0.05, 0))
hist(Sp_ts[Sp_ts<1e4], breaks = 40,</pre>
     col = clr, border = "blue3",
     yaxs = "i", xaxt = "n", yaxt = "n",
     main = "", ylab = "")
## (b) Probability of overfishing
par(mai=c(0.9, 0.9, 0, 0))
matplot(Pr_over, type = "1", lwd = 2, lty = "solid",
        col = c("slateblue","blue","darkblue"),
        ylab="Probability of overfishing",
        xlab="Harvest rate", xaxt="n",
        las = 1, cex.lab = 1.2)
axis(1, seq(0,100,20), seq(0,100,20)/100)
x_{p} < c(0, 0, 0)
for(i in 1:length(x_lp)) {
  x_{p[i]} \leftarrow \max(\text{which}(abs(Pr_over[,i] - 0.5) \leftarrow 0.05))
points(x = x_1p, y = rep(0.5, 3), pch = 21, cex = 4,
       col = "white", bg = "white")
text(x = x_1p, y = 0.5, expression(U[M75], U[MSY], U[Max]),
     col = c("slateblue", "blue", "darkblue"), cex = 0.8)
text(x = par()$usr[1] + xoffSet * diff(par()$usr[1:2]),
     y = par() susr[4] - yoffSet * diff(par() susr[3:4]),
     "(b)")
## marginal histogram of posterior harvest rates
par(mai = c(0, 0.9, 0.05, 0))
hist(aer, breaks = seq(0, 40)/40,
     col = clr, border = "blue3",
     yaxs = "i", xaxt = "n", yaxt = "n",
     main = "", ylab = "")
```

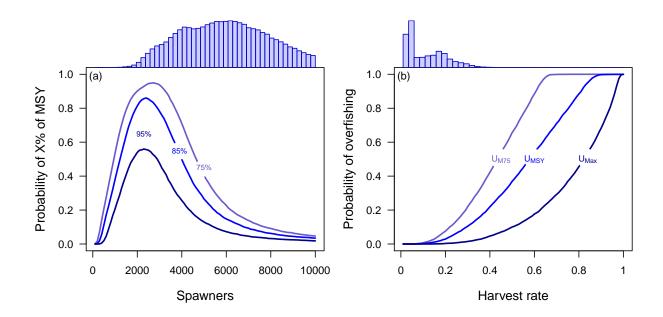


Figure 5: Plots of (a) the probability that a given number of spawners produces average yields achieving 95%, 85%, or 75% of the estimated maximum sustainable yield (MSY); and (b) the cumulative probability of overfishing the population, based on harvest rates equal to those at 75% of MSY, at MSY, and at the maximum per recruit. The histograms above (a) and (b) are distributions of the posterior estimates for the number of spawners and harvest rates, respectively; the histogram in (a) has been truncated at 10⁴.