R2. Model fitting and evaluation

2023 - 2024 Skagit River steelhead forecast.

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This is version 0.23.12.19.	

Requirements

All analyses require the R software (v3.4.3) for data retrieval, data processing, and summarizing model results, and the JAGS software (v4.2.0) for Markov chain Monte Carlo (MCMC) simulation. Please note that some of the R code below may not work with older versions of JAGS due to some changes in the ways that arrays are handled.

We also need a few packages that are not included with the base installation of R, so we begin by installing them (if necessary) and then loading them.

```
if(!require("here")) {
  install.packages("here")
  library("here")
if(!require("readr")) {
  install.packages("readr")
  library("readr")
if(!require("rjags")) {
  install.packages("rjags")
  library("rjags")
if(!require("loo")) {
  install.packages("loo")
  library("loo")
if(!require("ggplot2")) {
  install.packages("ggplot2")
  library("ggplot2")
if(!require("coda")) {
  install.packages("coda")
  library("coda")
if(!require("shinystan")) {
  install.packages("shinystan")
  library("shinystan")
if(!require("R2jags")) {
  install.packages("R2jags")
  library("R2jags")
if(!require("dclone")) {
  install.packages("dclone")
  library("dclone")
}
if(!require("snow")) {
  install.packages("snow")
  library("snow")
if(!require("rstan")) {
  install.packages("rstan")
  library("rstan")
}
```

```
if(!require("RColorBrewer")) {
   install.packages("RColorBrewer")
   library("RColorBrewer")
}
if(!require("gtools")) {
   install.packages("gtools")
   library("gtools")
}

## set directory locations
datadir <- here("data")
jagsdir <- here("jags")
analdir <- here("analysis")
savedir <- here("analysis/cache")</pre>
```

We also need a couple of helper functions.

```
## better round
Re2prec <- function(x, map = "round", prec = 1) {</pre>
  ## 'map' can be "round", "floor", or "ceiling"
  ## 'prec' is nearest value (eg, 0.1 means to nearest tenth; 1 gives normal behavior)
 if(prec<=0) { stop("\"prec\" cannot be less than or equal to 0") }</pre>
 do.call(map,list(x/prec))*prec
}
## wrapper function to fit JAGS models & rearrange output
fit_jags <- function(model, data, params, inits, ctrl, dir = jagsdir) {</pre>
  jm <- jags.model(file.path(jagsdir, model),</pre>
                    data,
                    inits,
                    ctrl$chains,
                    ctrl$burn.
                    quiet = TRUE)
 return(coda.samples(jm, params, ctrl$length, ctrl$thin))
}
#alternative wrapper to fit model in parallel; one chain per core
fit_jags2<-function(model,data,params,inits,ctrl,dir=jagsdir){</pre>
  cl <- makeCluster(3, type = "SOCK")</pre>
  inits2 <- jags.fit(data=data,</pre>
                      params=params,
                      model=file.path(jagsdir, model),
                      inits=inits,
                      n.chains=ctrl$chains,
                      n.adapt = 0,
                      n.update = 0,
                      n.iter = 0)$state(internal = TRUE)
  jm <- jags.parfit(cl=cl,</pre>
                     data = data,
                     params = params,
                     model = file.path(jagsdir, model),
                     inits = inits2,
                     n.adapt = ctrl$burn*0.5,
```

```
n.update = ctrl$burn*0.5,
                     n.iter = ctrl$length-ctrl$burn,
                     thin = ctrl$thin,
                     n.chains = ctrl$chains
  stopCluster(cl)
  return(jm)
#generate summary stats file from MCMC object
sum_stats<-function(mcmclist){</pre>
  ESS<-apply(as.matrix(mcmclist),2,ess_bulk)</pre>
  Rhat<-apply(as.matrix(mcmclist),2,Rhat)</pre>
  summary_stats<-summary(mcmclist)</pre>
  summary_stats<-data.frame(summary_stats$statistics,summary_stats$quantiles,ESS,Rhat)</pre>
}
# functions for approximate LFO
# many functions modified from:
# https://qithub.com/paul-buerkner/LFO-CV-paper/blob/master/case-study-LFO-CV.Rmd
#load complete model fits & model refits with subset data
loadmodfits<-function(modelnames){</pre>
  mod_fits<-list(NULL)</pre>
  for(i in 1:length(modelnames)){
    mod_fits[[i]] <- readRDS(file.path(savedir,paste0(modelnames[i],"_y",n_forecasts+1,".rds")))</pre>
    \#mod_fits[[i]] \leftarrow readRDS(file.path(savedir,paste0("fit_",modelnames[i],".rds")))
  }
  return(mod_fits)
#refits
loadrefits<-function(refitname,N,L){</pre>
  numrefits < -N-L+1
  re_fits<-list()
  for(i in 1:numrefits){
     re_fits[[i]] <- readRDS(file.path(savedir,paste0(refitname,"_y",i,".rds")))</pre>
  return(re_fits)
}
# more stable than log(sum(exp(x)))
log_sum_exp <- function(x) {</pre>
  max x <- max(x)
  \max_{x} + \log(\sup(\exp(x - \max_{x})))
# more stable than log(mean(exp(x)))
log_mean_exp <- function(x) {</pre>
  log_sum_exp(x) - log(length(x))
}
# compute log of raw importance ratios
```

```
# sums over observations *not* over posterior samples
sum_log_ratios <- function(ll, ids = NULL) {</pre>
  if (!is.null(ids)) ll <- ll[, ids, drop = FALSE]</pre>
  - rowSums(11)
# for printing comparisons later
rbind_print <- function(...) {</pre>
  round(rbind(...), digits = 2)
}
#function to extract log likelihood from fitted model
extract_log_lik<-function(m,esc_only,N,mod_fits){</pre>
  #extract pontwise log likelihoods
  tmp_lp <- as.matrix(mod_fits[[m]])</pre>
  ## extract pointwise likelihoods
  tmp_lp <- tmp_lp[,grepl("lp_", colnames(tmp_lp))]</pre>
  \#\# if numerical underflows, convert -Inf to 5% less than min(likelihood)
  if(any(is.infinite(tmp_lp))) {
    tmp_lp[is.infinite(tmp_lp)] <- NA</pre>
    tmp_min <- min(tmp_lp, na.rm = TRUE)</pre>
    tmp_lp[is.na(tmp_lp)] <- tmp_min * 1.05</pre>
  if(esc_only =="Yes"){
    tmp_lp<-tmp_lp[,grepl("esc", colnames(tmp_lp))]</pre>
  #qet yrs assoc
  names_loglik<-data.frame(strsplit(colnames(tmp_lp),"\\[|\\]"))</pre>
  yrnames<-as.numeric(names_loglik[2,])</pre>
  loglik <- matrix(NA,ncol=N,nrow=dim(tmp_lp)[1])</pre>
  for(i in 1:N){
    if(!is.null(ncol(tmp_lp[,yrnames==i]))){
      loglik[,i] = apply(tmp_lp[,yrnames==i],1,sum)
    }else(loglik[,i] = tmp_lp[,yrnames==i])
  return(loglik)
approx_LFO<-function(N=N,L,m=m,esc_only,mod_fits,userefits,refitname,thres){
  loglik = extract_log_lik(m=m, esc_only = esc_only,N=N,mod_fits = mod_fits)
  ## look at Pareto k's
  k_LOOIC<-pareto_k_values(loo(loglik))[(L+1):N]</pre>
  if(userefits=="Yes"){
    re_fits =loadrefits(refitname=refitname, N=N, L=L)
  i_refit <- L
  refits <- L
  ks <- NULL
  approx_elpds_1sap <- rep(NA, N)
  for (i in (N - 1):L) {
    logratio <- sum_log_ratios(loglik, (i + 1):N)</pre>
    psis_obj <- suppressWarnings(psis(logratio))</pre>
```

```
k<-pareto_k_values(psis_obj)
    ks \leftarrow c(ks, k)
    if(k>thres & userefits=="Yes"){
    #use_refit of model based on the first[i] observations
      i_refit <- i
      refits <- c(refits, i)</pre>
      loglik = extract_log_lik(m=(i+1)-L+1, esc_only = esc_only, N=N, mod_fits = re_fits)
      approx elpds 1sap[i + 1] <- log mean exp(loglik[, i + 1])
    }else{
      lw <- weights(psis_obj, normalize = TRUE)[, 1]</pre>
      approx_elpds_1sap[i + 1] <- log_sum_exp(lw + loglik[, i + 1])</pre>
    }
  }
  results<-list(approx_elpds_1sap,ks,k_L00IC)</pre>
  names(results)<-c("LFO","ks","k_LOOIC")</pre>
  return(results)
}
plot_ks <- function(ks, thres = 0.7,N,L) {
  ids = N:(L + 1)
  dat_ks <- data.frame(ks = ks, ids = ids)</pre>
  ggplot(dat_ks, aes(x = ids, y = ks)) +
    geom_point(aes(color = ks > thres), shape = 3, show.legend = FALSE) +
    geom_hline(yintercept = thres, linetype = 2, color = "red2") +
    scale color manual(values = c("cornflowerblue", "darkblue")) +
    labs(x = "Data point", y = "Pareto k") +
    ylim(-0.5, max(dat ks$ks))
#function for printing out a read text file
processFile = function(filepath) {
  con = file(filepath, "r")
  while ( TRUE ) {
    line = readLines(con, n = 1)
    if ( length(line) == 0 ) {
      break
    }
    cat(paste0(noquote(line)),"\n")
  close(con)
#calculate stacking weights
find_stack_weights<-function(tau,metric,n,initial_weights,preds,obs){</pre>
  tweights <- initial weights
  preds<-as.matrix(preds)</pre>
  obs<-obs
  tau=tau
  skill_list<-c(NULL)</pre>
  metric=metric
  for(i in 1:n){
    pred_trs_ensemble<- preds %*% as.vector(tweights)</pre>
    Error <- pred_trs_ensemble - obs</pre>
    SE <- Error<sup>2</sup>
    PE <- Error/obs_trs
```

```
APE <- abs(PE)
   LAR <- log(obs_trs/pred_trs_ensemble)
   RMSE <- apply(SE,2,function(x){sqrt(mean(x))})</pre>
   MPE <- apply(PE,2,function(x){mean(x)})</pre>
   MAPE <- apply(APE,2,function(x){mean(x)})
   MSA \leftarrow apply(LAR, 2, function(x) \{100*(exp(mean(abs(x)))-1)\})
   if(i==1){
      skill=get(metric)
      weights=tweights
   if(get(metric)<skill){</pre>
      skill=get(metric)
      weights=tweights
   }
   skill_list<-c(skill_list,min(get(metric),skill))</pre>
   keep<-rbinom(1,prob=skill/get(metric),1)</pre>
   if(keep==1){tweights=tweights }else(tweights=weights)
   tweights = rdirichlet(n=1,alpha = tweights*tau+0.001)
 results<-list(weights,skill,skill_list)</pre>
 return(results)
}
#===========
#function to fit or load modelfits
fit_load_mods<-function(models){</pre>
  ## empty list for fits
 mod_fits <- vector("list", n_mods*(n_forecasts+1))</pre>
  ## counter to index fitted jags models (33 in total: 3 models x 11 1 year ahead forecasts including u
  ## return year)
  t <- 1
  for(n in 1:n_mods){
    ## counter to index data to feed model for year specific forecasts
    ## first forecast will be for 10 years prior to the most recent return year;
    ## last forecast will be current forecast for the upcoming return year
   c <- 0
    #n < -1
   model <- models[n]</pre>
   for(i in 1:(n_forecasts+1)){
      #i <-1
      if(file.exists(file.path(savedir,paste(model,"_","y",i,".rds",sep = "")))) {
       mod_fits[[t]] <- readRDS(file.path(savedir,paste(model,"_","y",i,".rds",sep = "")))</pre>
        c < - c + 1
        t < -t + 1
   } else { ## else, fit & save
          ## cnt & time stamp
          cat("Count =", t, "; Time =", round(((proc.time()-timer_start)/60)["elapsed"], 1), "\n",
              file="cnt_time.txt", append=TRUE)
```

```
#range of years. Last year in range
dat_yrs <- seq(yr_frst,(yr_last - n_forecasts + c),1)</pre>
## number of years of data
n yrs <- length(dat yrs)</pre>
## get first & last years
yr frst forecast <- min(dat yrs)</pre>
yr_last_forecast <- max(dat_yrs)</pre>
## get escapement data
dat_esc_forecast <- dat_esc[which(dat_esc$year %in% dat_yrs),]</pre>
## log of escapement
ln_dat_esc <- c(log(dat_esc_forecast$escapement),rep(NA,n_fore))</pre>
## get age data
dat_age_forecast <- dat_age[which(dat_age$year %in% dat_yrs),]</pre>
## drop year col & first age min+age skip rows
dat_age_forecast <- dat_age_forecast[-(1:(age_min+age_skip)),-1]</pre>
## add row(s) of NA's for forecast years
if(n_fore > 0) {
  dat_age_forecast <- rbind(dat_age_forecast,</pre>
                              matrix(0, n_fore, A,
                                     dimnames = list(n_yrs+seq(n_fore),colnames(dat_age_forecas
## total num of age obs by cal yr
dat_age_forecast[,"sum"] <- apply(dat_age_forecast, 1, sum)</pre>
## row indices for any years with no obs age comp
idx_NA_yrs <- which(dat_age_forecast$sum<A, TRUE)</pre>
## replace 0's in yrs w/o any obs with NA's
dat_age_forecast[idx_NA_yrs,(1:A)] <- NA</pre>
## change total in yrs w/o any obs from 0 to A to help dmulti()
dat_age_forecast[idx_NA_yrs,"sum"] <- A</pre>
## convert class
dat_age_forecast <- as.matrix(dat_age_forecast)</pre>
## get harvest data
dat_harv_forecast <- dat_harv[which(dat_harv$year %in% dat_yrs),]</pre>
## drop year col & first age_max rows
dat_harv_forecast <- c(dat_harv_forecast$catch,rep(NA,n_fore))</pre>
## get covariate data
dat_cvrs_forecast <- dat_cvrs[which(dat_cvrs$year <= yr_last + n_fore - age_min),1:4]</pre>
## drop year col
dat_cvrs_forecast <- dat_cvrs_forecast[,-1]</pre>
## transform the covariates to z-scores
scl_cvrs_forecast <- scale(dat_cvrs_forecast)</pre>
## total number of covariates
n_cov <- dim(dat_cvrs_forecast)[2]</pre>
```

```
## ----jags_setup----
## 1. Data to pass to JAGS
dat_jags <- list(dat_age = dat_age_forecast,</pre>
                 ln dat esc = ln dat esc,
                  dat_harv = dat_harv_forecast,
                  A = A
                  age_min = age_min,
                  age max = age max,
                  age_skip = age_skip,
                  n_yrs = n_yrs,
                 n_fore = n_fore)
## 2. Model params/states for JAGS to return
      These are specific to the process model,
##
      so we define them in 'par_jags' below.
if(model == "IPM_BH_cov_AR" | model == "IPM_BH_cov_AR_resid"){
  init_vals_cov <- function() {</pre>
  list(alpha = 5,
       beta_inv = exp(mean(ln_dat_esc, na.rm = TRUE)),
       gamma = rep(0, 3),
       pi_tau = 10,
       pi_eta = rep(1,A),
       pi_vec = matrix(c(0.01,0.35,0.47,0.15,0.01,0.01),
                        n_yrs-age_min+n_fore, A,
                        byrow = TRUE),
       Rec_mu = log(1000),
       Rec_sig = 0.1,
       tot_ln_Rec = rep(log(1000), n_yrs - age_min + n_fore),
       phi_prior = 0.5,
       innov_1 = 0)
  }
  ## params/states to return
  par_jags<- c("alpha", "E_BH_a", "ln_BH_a",</pre>
                 "beta",
                 "gamma",
                 "Sp", "Rec", "tot ln Rec", "ln RS",
                 "pi_eta", "pi_tau",
                 "sigma_r", "sigma_s", "w", "res_ln_Rec",
                 "lp_age", "lp_esc", "phi", "Run"
}else{
  init_vals_cov <- function() {</pre>
    list(alpha = 5,
         beta_inv = exp(mean(ln_dat_esc, na.rm = TRUE)),
         gamma = rep(0, 3),
         pi_tau = 10,
         pi_eta = rep(1,A),
         \# pi\_vec = matrix(c(0.01, 0.35, 0.47, 0.15, 0.01, 0.01),
         #
                            n_yrs-age_min+n_fore, A,
                            byrow = TRUE),
```

```
Rec_mu = log(1000),
                    Rec_sig = 0.1,
                    tot_ln_Rec = rep(log(1000), n_yrs - age_min + n_fore),
                    phi_prior = 0.5, theta_res_prior = 0.5,
                    innov_1 = 0)
               }
             ## params/states to return
             par_jags <- c("alpha", "E_BH_a", "ln_BH_a",</pre>
                            "beta",
                            "gamma",
                            "Sp", "Rec", "tot_ln_Rec", "ln_RS", "tot_Run",
                            "pi_eta", "pi_tau",
                            "sigma_r", "sigma_s", "res_ln_Rec", "w", "theta_res", "phi",
                            "lp_age", "lp_esc", "Run"
          }#endif
          ## set of multi-covariate models
          cset <- colnames(scl_cvrs_forecast)</pre>
          dat_jags$n_cov <- length(cset)</pre>
          dat_jags$mod_cvrs <- scl_cvrs_forecast[1:(n_yrs-age_min+1), cset]</pre>
          ## fit model & save it
          \# mod\_fits[[t]] \leftarrow fit\_jags(paste(model,".txt",sep = ""), dat\_jags, par\_jags,
                                         init_vals_cov, mcmc_ctrl)
          mod_fits[[t]]<-fit_jags2(model=paste(model,".txt",sep = ""),</pre>
                               data=dat_jags,
                               params=par_jags,
                               inits=init_vals_cov,
                               ctrl=mcmc_ctrl
          saveRDS(mod_fits[[t]], file.path(savedir,paste(model,"_","y",i,".rds",sep = "")))
          summary_stats<-NULL
          summary_stats<-sum_stats(mcmclist= mod_fits[[t]])</pre>
          write.csv(summary_stats,file.path(savedir, paste(model,"_","y",i,"_summary_stats.csv",sep = "
          c < - c + 1
          t < -t + 1
      }## end if
    }##next forecast year(i)
  }## next model(n)
  return(mod_fits)
}
```

User inputs

We begin by supplying values for the following parameters, which we need for model fitting and evaluation.

```
## first & last years of fish data
yr frst <- 1978
yr_last <- 2023
## min & max adult age classes
age_min <- 3
age_max <- 8
## years (if any) of age-comp to skip; see below
age_skip <- 0
## number of years ahead for run forecasts from the most recent year of data
n_fore <- 1
## number of recent year forecasts
n_forecasts <- 10
## first year of 1 step ahead forecast
yr begin <- 2014
## last year of 1 step ahead forecast
yr_end <- 2023</pre>
## upper threshold for Gelman & Rubin's potential scale reduction factor (Rhat).
Rhat_thresh <- 1.1
```

Next we specify the names of three necessary data files containing the following information:

- 1. observed total number of adult spawners (escapement) by year;
- 2. observed age composition of adult spawners by year;
- 3. observed total harvest by year;

```
## 1. file with escapement data
## [n_yrs x 2] matrix of obs counts; 1st col is calendar yr
fn_esc <- "skagit_sthd_esc.csv"

## 2. file with age comp data
## [n_yrs x (1+A)]; 1st col is calendar yr
fn_age <- "skagit_sthd_age.csv"

## 3. file with harvest data
## [n_yrs x 2] matrix of obs catch; 1st col is calendar yr
fn_harv <- "skagit_sthd_catch.csv"</pre>
```

Loading the fish data

Here we load in the first three data files and do some simple calculations and manipulations. First the spawner data:

```
## escapement
dat_esc <- read_csv(file.path(datadir, fn_esc))
## years of data
dat_yrs <- dat_esc$year

## number of years of data
n_yrs <- length(dat_yrs)

## log of escapement
ln_dat_esc <- c(log(dat_esc$escapement),rep(NA,n_fore))</pre>
```

Next the age composition data:

```
## age comp data
dat_age <- read_csv(file.path(datadir, fn_age))</pre>
## num of age classes
A <- age_max - age_min + 1
# ## drop year col & first age_min+age_skip rows
# dat_age <- dat_age[-(1:(age_min+age_skip)),-1]</pre>
# ## add row(s) of NA's for forecast years
# if(n_fore > 0) {
  dat_age <- rbind(dat_age,</pre>
                      matrix(0, n_fore, A,
#
                              dimnames = list(n_yrs+seq(n_fore),
#
                                              colnames(dat_age))))
# }
# ## total num of age obs by cal yr
# dat_age[,"sum"] <- apply(dat_age, 1, sum)</pre>
# ## row indices for any years with no obs age comp
# idx_NA_yrs <- which(dat_age$sum<A, TRUE)</pre>
# ## replace 0's in yrs w/o any obs with NA's
\# dat_age[idx_NA_yrs,(1:A)] \leftarrow NA
# ## change total in yrs w/o any obs from 0 to A to help dmulti()
# dat_age[idx_NA_yrs,"sum"] <- A</pre>
# ## convert class
# dat_age <- as.matrix(dat_age)</pre>
```

And then the harvest data:

```
## harvest
dat_harv <- read_csv(file.path(datadir, fn_harv))
## drop year col & first age_max rows
#dat_harv <- c(dat_harv$catch,rep(NA,n_fore))</pre>
```

Loading the covariates

Our analysis investigates 5 covariates as possible drivers of the population's instrinic growth rate:

1. Maximum river discharge in winter;

- 2. Minimum river discharge in summer;
- 3. North Pacific Gyre Oscillation;

All of the covariates are contained in the file /data/skagit_sthd_covars.csv. We will load and then standardize them to have zero-mean and unit-variance.

```
dat_cvrs <- read_csv(file.path(datadir, "skagit_sthd_covars.csv"))
## drop year col
# dat_cvrs <- dat_cvrs[,-1]
# ## transform the covariates to z-scores
# scl_cvrs <- as.matrix(scale(dat_cvrs))
# ## total number of covariates
# n_cov <- dim(dat_cvrs)[2]</pre>
```

Specifying models in JAGS

Now we can specify the model in JAGS. We evaluated a total of six different models, which we outline below, based on a beverton holt process model with covariates. Specifically, we evaluated three different methods to model recruitment residuals including a moving average process lagged 1 year (MA1), an auto-regressive moving average process lagged 1 year (AR1MA1), and an auto-regressive process lagged 1 year (AR1). In addition, we evaluated two methods for modeling maturation including a mean reverting process, and a random walk process.

Beverton-Holt with covars and AR1 process errors (MA1 recruitment residuals). Here we will print out the model (contained in a separate text file)

```
processFile(file.path(jagsdir, "IPM_BH_cov_AR.txt"))
```

```
##
##
       model {
##
##
       ##----
       ## PRIORS
##
##
       ##----
##
       ## alpha = intrinsic productivity
##
       alpha ~ dnorm(0,0.001) T(0,);
##
       mu_BH_a <- log(alpha);</pre>
       E_BH_a \leftarrow mu_BH_a + sigma_r/(2 - 2*phi^2);
##
##
##
       ## strength of dens depend
##
       beta inv ~ dnorm(0, 1e-9) T(0,);
##
       beta <- 1/beta_inv;</pre>
##
       ## covariate effects
##
##
       for(i in 1:n_cov) { gamma[i] ~ dnorm(0,0.01) }
##
##
       ## AR(1) coef for proc errors
       #phi ~ dunif(-0.999,0.999);
##
##
       #phi <- 0;
```

```
##
       phi_prior ~ dbeta(2,2);
##
       phi <- phi_prior*2-1;</pre>
##
       #phi ~ dunif(0,0.999);
##
##
       ## innovation in first year
##
       innov_1 ~ dnorm(0,tau_r*(1-phi*phi));
##
##
       ## process variance for recruits model
##
       sigma_r ~ dnorm(0, 2e-2) T(0,);
##
       tau_r <- 1/sigma_r;</pre>
##
##
       ## obs variance for spawners
       tau_s <- 1/sigma_s;</pre>
##
##
       sigma_s ~ dnorm(0, 0.001) T(0,);
##
##
       ## unprojectable early recruits;
##
       ## hyper mean across all popns
       Rec mu ~ dnorm(0,0.001);
##
##
       ## hyper SD across all popns
       Rec_sig ~ dunif(0,100);
##
##
       ## precision across all popns
##
       Rec_tau <- pow(Rec_sig,-2);</pre>
       ## multipliers for unobservable total runs
##
##
       ttl_run_mu ~ dunif(1,5);
##
       ttl_run_tau ~ dunif(1,20);
##
##
       ## get total cal yr returns for first age_min yrs
##
       for(i in 1:(age_min+age_skip)) {
##
       ln_tot_Run[i] ~ dnorm(ttl_run_mu*Rec_mu,Rec_tau/ttl_run_tau);
##
       tot_Run[i] <- exp(ln_tot_Run[i]);</pre>
##
       }
##
##
       ## maturity schedule
##
       ## unif vec for Dirch prior
       theta <-c(1,10,10,5,1,1)
##
##
       ## hyper-mean for maturity
##
       pi_eta ~ ddirch(theta);
##
       ## hyper-prec for maturity
##
       pi_tau ~ dnorm(0, 0.01) T(0,);
##
       for(t in 1:(n_yrs-age_min+n_fore)) { pi_vec[t,1:A] ~ ddirch(pi_eta*pi_tau) }
##
##
       ## estimated harvest rate
       for(t in 1:(n_yrs+n_fore)) { h_rate[t] ~ dunif(0,1) }
##
##
       ##-----
##
       ## LIKELIHOOD
##
       ##-----
       ## predicted recruits in BY t
##
##
       covar[1] <- inprod(gamma, mod_cvrs[1,]);</pre>
       ln_BH_a[1] <- mu_BH_a + covar[1];</pre>
##
##
       E_{n_{ec}[1]} \leftarrow n_{BH_a[1]} + n_{Sp[1]} - \log(1 + beta*Sp[1]) + phi*innov_1;
##
       tot_ln_Rec[1] ~ dnorm(E_ln_Rec[1],tau_r);
##
       res_ln_Rec[1] <- tot_ln_Rec[1] - E_ln_Rec[1];
##
       w[1] <- phi * innov_1 + res_ln_Rec[1];
##
```

```
##
       ## median of total recruits
##
       tot_Rec[1] <- exp(tot_ln_Rec[1]);</pre>
##
##
       ## R/S
##
       ln_RS[1] <- tot_ln_Rec[1] - ln_Sp[1];</pre>
##
##
       ## brood-yr recruits by age
       for(a in 1:A) {
##
##
       Rec[1,a] <- tot_Rec[1] * pi_vec[1,a];</pre>
##
       }
##
##
       ## brood years 2:(n_yrs-age_min)
       for(t in 2:(n_yrs-age_min+n_fore)) {
##
##
       ## predicted recruits in BY t
##
       covar[t] <- inprod(gamma, mod_cvrs[t,]);</pre>
##
       ln_BH_a[t] <- mu_BH_a + covar[t];</pre>
##
       E_{n_{ec}[t]} < n_{BH_a[t]} + n_{Sp[t]} - \log(1 + beta*Sp[t]) + phi*res_ln_{Rec}[t-1];
##
       tot_ln_Rec[t] ~ dnorm(E_ln_Rec[t],tau_r);
##
       res_ln_Rec[t] <- tot_ln_Rec[t] - E_ln_Rec[t];</pre>
       w[t] <- phi * res_ln_Rec[t-1] + res_ln_Rec[t];</pre>
##
##
##
       ## median of total recruits
##
       tot_Rec[t] <- exp(tot_ln_Rec[t]);</pre>
##
##
       ## R/S
##
       ln_RS[t] <- tot_ln_Rec[t] - ln_Sp[t];</pre>
##
       ## brood-yr recruits by age
##
##
       for(a in 1:A) {
##
       Rec[t,a] <- tot_Rec[t] * pi_vec[t,a];</pre>
##
       } ## end t loop over year
##
##
##
       ## get predicted calendar year returns by age
       ## matrix Run has dim [(n yrs-age min) x A]
##
##
       ## step 1: incomplete early broods
##
       ## first cal yr of this grp is first brood yr + age_min + age_skip
##
       for(i in 1:(age_max-age_min-age_skip)) {
##
       ## projected recruits
##
##
       for(a in 1:(i+age_skip)) {
       Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
##
##
##
       ## imputed recruits
##
       for(a in (i+1+age_skip):A) {
       lnRec[i,a] ~ dnorm(Rec_mu,Rec_tau);
##
##
       Run[i,a] <- exp(lnRec[i,a]);</pre>
##
       }
##
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
##
       ## predicted age-prop vec for multinom
```

```
##
       for(a in 1:A) {
##
       age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
##
       ## multinomial for age comp
##
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
       lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
##
##
       ## step 2: info from complete broods
##
##
       ## first cal yr of this grp is first brood yr + age_max
       for(i in (A-age_skip):(n_yrs-age_min-age_skip+n_fore)) {
##
       for(a in 1:A) {
##
##
       Run[i,a] \leftarrow Rec[(age_skip+i)-a+1,a];
##
##
##
       ## total run size
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
##
##
       ## predicted age-prop vec for multinom
##
       for(a in 1:A) {
##
       age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
       }
##
##
       ## multinomial for age comp
##
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
       lp_age[i] <- ifelse(i < n_yrs-age_min-age_skip+n_fore, logdensity.multi(dat_age[i,1:A],age_v[i,1</pre>
##
##
       ## get predicted calendar year spawners
##
##
       ## first cal yr is first brood yr
##
       for(t in 1:(n_yrs+n_fore)) {
       ## obs model for spawners
##
       #Sp[t] <- max(10,tot_Run[t] - dat_harv[t]);
##
       est_harv[t] = ifelse(t > n_yrs,1,h_rate[t] * tot_Run[t]);
##
       dat_harv[t] ~ dlnorm(log(est_harv[t]), 20);
##
##
       Sp[t] = tot_Run[t] - est_harv[t];
##
       ln_Sp[t] \leftarrow log(Sp[t]);
##
       ln_dat_esc[t] ~ dnorm(ln_Sp[t], tau_s);
##
##
       lp_esc[t] <- ifelse(t < n_yrs + 1,logdensity.norm(ln_dat_esc[t],ln_Sp[t], tau_s),0);</pre>
##
##
     } ## end model description
##
## Warning in readLines(con, n = 1): incomplete final line found on 'C:/Users/cruff/Documents/projects/
## RWorkflow/Skagit-River-Steelhead-Forecast/jags/IPM_BH_cov_AR.txt'
```

Beverton-Holt with covars and AR1MA1 recruitment residuals

```
processFile(file.path(jagsdir, "IPM_BH_cov_MA1_AR1.txt"))
```

```
##
## model {
##
##
     ##----
##
     ## PRIORS
##
     ##----
##
     ## alpha = intrinsic productivity
##
     alpha ~ dnorm(0,0.001) T(0,);
##
     mu_BH_a <- log(alpha);</pre>
##
     E_BH_a \leftarrow mu_BH_a + sigma_r/(2 - 2*phi^2);
##
##
     ## strength of dens depend
     beta_inv ~ dnorm(0, 1e-9) T(0,);
##
##
     beta <- 1/beta_inv;</pre>
##
##
     ## covariate effects
##
     for(i in 1:n_cov) { gamma[i] ~ dnorm(0,0.01) }
##
     ## AR(1) coef for recruitment residual
##
     #phi ~ dunif(-0.999,0.999);
##
##
     #phi <- 0;</pre>
##
     phi_prior ~ dbeta(2,2);
     phi <- phi_prior*2-1;</pre>
##
##
     #phi ~ dunif(0,0.999);
##
##
     ## MA(1) coef recruitment residual
##
     theta_res_prior ~ dbeta(2,2);
     theta_res <- theta_res_prior*2-1;</pre>
##
     #theta_res ~ dunif(0,0.999);
##
##
##
     ## innovation in first year
##
     #innov_1 ~ dnorm(0,tau_r*(1-phi*phi));#AR1
     innov_1 \sim dnorm(0,(1-phi^2)/((1+2*phi*theta_res+theta_res^2)*sigma_r^2));#AR1MA1
##
##
##
     ## process variance for recruits model
##
     sigma_r ~ dnorm(0, 2e-2) T(0,);
##
     tau_r <- 1/sigma_r;</pre>
##
##
     ## obs variance for spawners
     tau_s <- 1/sigma_s;</pre>
##
##
     sigma s \sim dnorm(0, 0.001) T(0,);
##
     ## unprojectable early recruits;
##
##
     ## hyper mean across all popns
     Rec_mu ~ dnorm(0,0.001);
##
     ## hyper SD across all popns
##
     Rec_sig ~ dunif(0,100);
##
##
     ## precision across all popns
##
     Rec_tau <- pow(Rec_sig,-2);</pre>
     ## multipliers for unobservable total runs
##
## ttl_run_mu ~ dunif(1,5);
## ttl_run_tau ~ dunif(1,20);
##
##
     ## get total cal yr returns for first age_min yrs
```

```
##
     for(i in 1:(age_min+age_skip)) {
##
        ln_tot_Run[i] ~ dnorm(ttl_run_mu*Rec_mu,Rec_tau/ttl_run_tau);
##
        tot_Run[i] <- exp(ln_tot_Run[i]);</pre>
##
     }
##
##
     ## maturity schedule
     ## unif vec for Dirch prior
##
     theta \leftarrow c(1,10,10,5,1,1)
##
##
     ## hyper-mean for maturity
     pi_eta ~ ddirch(theta);
##
##
     ## hyper-prec for maturity
     pi_tau ~ dnorm(0, 0.01) T(0,);
##
     for(t in 1:(n_yrs-age_min+n_fore)) { pi_vec[t,1:A] ~ ddirch(pi_eta*pi_tau) }
##
##
##
     ## estimated harvest rate
##
     for(t in 1:(n_yrs+n_fore)) { h_rate[t] ~ dunif(0,1) }
##
##
     ##-----
##
     ## LIKELIHOOD
##
     ##-----
##
     ## predicted recruits in BY t
##
     covar[1] <- inprod(gamma,mod_cvrs[1,]);</pre>
##
     ln_BH_a[1] <- mu_BH_a + covar[1];</pre>
     E_{\ln_{ec}[1]} \leftarrow \ln_{ec}[1] + \ln_{ec}[1] - \log(1 + \beta_{ec}[1]) + phi * innov_1 + theta_{ec}[1]
##
     tot_ln_Rec[1] ~ dnorm(E_ln_Rec[1], tau_r);
##
     res_ln_Rec[1] <- tot_ln_Rec[1] - E_ln_Rec[1];</pre>
##
##
     w[1] <- phi * innov_1 + theta_res * 0 + res_ln_Rec[1]</pre>
##
##
     ## median of total recruits
##
     tot_Rec[1] <- exp(tot_ln_Rec[1]);</pre>
##
##
     ## R/S
##
     ln_RS[1] <- tot_ln_Rec[1] - ln_Sp[1];</pre>
##
##
     ## brood-yr recruits by age
##
     for(a in 1:A) {
##
       Rec[1,a] <- tot_Rec[1] * pi_vec[1,a];</pre>
##
##
##
     ## brood years 2:(n_yrs-age_min)
     for(t in 2:(n_yrs-age_min+n_fore)) {
##
##
       ## predicted recruits in BY t
       covar[t] <- inprod(gamma, mod_cvrs[t,]);</pre>
##
##
       ln_BH_a[t] <- mu_BH_a + covar[t];</pre>
##
##
##
       #version 4; more similar to AR1 original model
##
       ##
       E_{n_Rec[t]} \leftarrow n_BH_a[t] + n_Sp[t] - \log(1 + beta*Sp[t]) + phi * w[t-1] + theta_res * res_ln_Rec[t]
       tot_ln_Rec[t] ~ dnorm(E_ln_Rec[t], tau_r);
##
##
       res_ln_Rec[t] <- tot_ln_Rec[t] - E_ln_Rec[t];</pre>
##
       w[t] <- phi * w[t-1] + theta_res * res_ln_Rec[t-1] + res_ln_Rec[t];
##
##
```

```
##
       ## median of total recruits
##
       tot_Rec[t] <- exp(tot_ln_Rec[t]);</pre>
##
       ## R/S
##
       ln_RS[t] <- tot_ln_Rec[t] - ln_Sp[t];</pre>
##
       ## brood-yr recruits by age
##
       for(a in 1:A) {
##
         Rec[t,a] <- tot_Rec[t] * pi_vec[t,a];</pre>
##
       }
##
     } ## end t loop over year
##
##
     ## get predicted calendar year returns by age
     ## matrix Run has dim [(n_yrs-age_min) x A]
##
     ## step 1: incomplete early broods
##
     ## first cal yr of this grp is first brood yr + age_min + age_skip
##
##
     for(i in 1:(age_max-age_min-age_skip)) {
##
       ## projected recruits
##
       for(a in 1:(i+age_skip)) {
         Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
##
##
       ## imputed recruits
##
       for(a in (i+1+age_skip):A) {
##
         lnRec[i,a] ~ dnorm(Rec_mu,Rec_tau);
##
         Run[i,a] <- exp(lnRec[i,a]);</pre>
##
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
       ## predicted age-prop vec for multinom
       for(a in 1:A) {
##
##
         age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
       ## multinomial for age comp
##
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
       lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
##
##
##
##
     ## step 2: info from complete broods
##
     ## first cal yr of this grp is first brood yr + age max
##
     for(i in (A-age_skip):(n_yrs-age_min-age_skip+n_fore)) {
##
       for(a in 1:A) {
         Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
##
##
       ## total run size
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
##
       ## predicted age-prop vec for multinom
##
       for(a in 1:A) {
##
         age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
       ## multinomial for age comp
##
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
       #lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
##
##
       lp_age[i] <- ifelse(i < n_yrs-age_min-age_skip+n_fore,</pre>
##
       logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]),0)
##
     }
##
```

```
##
     ## get predicted calendar year spawners
##
     ## first cal yr is first brood yr
     for(t in 1:(n_yrs+n_fore)) {
##
       ## obs model for spawners
##
       # Sp[t] <- max(10,tot_Run[t] - dat_harv[t]);
##
       est_harv[t] = ifelse(t > n_yrs,1,h_rate[t] * tot_Run[t]);
##
       dat_harv[t] ~ dlnorm(log(est_harv[t]), 20);
##
       Sp[t] = tot_Run[t] - est_harv[t];
##
##
       ln_Sp[t] <- log(Sp[t]);</pre>
##
       ln_dat_esc[t] ~ dnorm(ln_Sp[t], tau_s);
##
       lp_esc[t] <- ifelse(t < n_yrs + 1,logdensity.norm(ln_dat_esc[t],ln_Sp[t], tau_s),0);</pre>
##
##
## } ## end model description
##
## Warning in readLines(con, n = 1): incomplete final line found on 'C:/Users/cruff/Documents/projects/
## RWorkflow/Skagit-River-Steelhead-Forecast/jags/IPM_BH_cov_MA1_AR1.txt'
```

Beverton-Holt with covars and AR1 recruitment residuals

```
processFile(file.path(jagsdir, "IPM_BH_cov_AR_resid.txt"))
```

```
##
## model {
##
##
     ##----
##
     ## PRIORS
##
     ##----
     ## alpha = intrinsic productivity
##
##
     alpha ~ dnorm(0,0.001) T(0,);
##
     mu_BH_a <- log(alpha);</pre>
     E_BH_a \leftarrow mu_BH_a + sigma_r/(2 - 2*phi^2);
##
##
##
     ## strength of dens depend
     beta_inv ~ dnorm(0, 1e-9) T(0,);
##
##
     beta <- 1/beta_inv;</pre>
##
##
     ## covariate effects
##
     for(i in 1:n_cov) { gamma[i] ~ dnorm(0,0.01) }
##
##
     ## AR(1) coef for recruitment residual
     #phi ~ dunif(-0.999,0.999);
##
     #phi <- 0;</pre>
##
##
     phi_prior ~ dbeta(2,2);
     phi <- phi_prior*2-1;</pre>
##
##
     #phi ~ dunif(0,0.999);
##
##
     ## innovation in first year
     innov_1 ~ dnorm(0,tau_r*(1-phi*phi));#AR1
##
##
```

```
##
     ## process variance for recruits model
##
     sigma_r ~ dnorm(0, 2e-2) T(0,);
##
     tau_r <- 1/sigma_r;</pre>
##
##
     ## obs variance for spawners
##
     tau s <- 1/sigma s;
##
     sigma s \sim dnorm(0, 0.001) T(0,);
##
##
     ## unprojectable early recruits;
##
     ## hyper mean across all popns
##
     Rec_mu ~ dnorm(0,0.001);
     ## hyper SD across all popns
##
     Rec_sig ~ dunif(0,100);
##
##
     ## precision across all popns
##
    Rec_tau <- pow(Rec_sig,-2);</pre>
##
    ## multipliers for unobservable total runs
## ttl_run_mu ~ dunif(1,5);
## ttl_run_tau ~ dunif(1,20);
##
##
     ## get total cal yr returns for first age_min yrs
##
     for(i in 1:(age_min+age_skip)) {
##
        ln_tot_Run[i] ~ dnorm(ttl_run_mu*Rec_mu,Rec_tau/ttl_run_tau);
##
        tot_Run[i] <- exp(ln_tot_Run[i]);</pre>
##
##
##
     ## maturity schedule
##
     ## unif vec for Dirch prior
     theta <-c(1,10,10,5,1,1)
##
     ## hyper-mean for maturity
##
     pi_eta ~ ddirch(theta);
##
##
     ## hyper-prec for maturity
##
     pi_tau ~ dnorm(0, 0.01) T(0,);
     for(t in 1:(n_yrs-age_min+n_fore)) { pi_vec[t,1:A] ~ ddirch(pi_eta*pi_tau) }
##
##
##
     ## estimated harvest rate
##
     for(t in 1:(n_yrs+n_fore)) { h_rate[t] ~ dunif(0,1) }
##
##
     ##----
##
     ## LIKELIHOOD
     ##-----
##
##
     ## predicted recruits in BY t
##
     covar[1] <- inprod(gamma,mod_cvrs[1,]);</pre>
     ln_BH_a[1] <- mu_BH_a + covar[1];</pre>
##
##
     E_{n_{ec}[1]} \leftarrow n_{BH_a[1]} + n_{Sp[1]} - \log(1 + beta*Sp[1]) + phi * innov_1;
##
     tot_ln_Rec[1] ~ dnorm(E_ln_Rec[1], tau_r);
##
     res_ln_Rec[1] <- tot_ln_Rec[1] - E_ln_Rec[1];</pre>
##
     w[1] <- phi * innov_1 + res_ln_Rec[1];
##
##
     ## median of total recruits
     tot_Rec[1] <- exp(tot_ln_Rec[1]);</pre>
##
##
##
     ## R/S
##
     ln_RS[1] <- tot_ln_Rec[1] - ln_Sp[1];</pre>
##
```

```
##
     ## brood-yr recruits by age
##
     for(a in 1:A) {
       Rec[1,a] <- tot_Rec[1] * pi_vec[1,a];</pre>
##
##
##
##
     ## brood years 2:(n_yrs-age_min)
##
     for(t in 2:(n yrs-age min+n fore)) {
       ## predicted recruits in BY t
##
##
       covar[t] <- inprod(gamma, mod_cvrs[t,]);</pre>
       ln_BH_a[t] <- mu_BH_a + covar[t];</pre>
##
##
       E_{n_{ec}[t]} < n_{BH_a[t]} + n_{Sp[t]} - \log(1 + beta*Sp[t]) + phi * w[t-1];
##
       tot_ln_Rec[t] ~ dnorm(E_ln_Rec[t], tau_r);
       res_ln_Rec[t] <- tot_ln_Rec[t] - E_ln_Rec[t];</pre>
##
##
       w[t] <- phi * w[t-1] + res_ln_Rec[t];
##
##
       ## median of total recruits
##
       tot_Rec[t] <- exp(tot_ln_Rec[t]);</pre>
       ## R/S
##
##
       ln_RS[t] <- tot_ln_Rec[t] - ln_Sp[t];</pre>
##
       ## brood-yr recruits by age
##
       for(a in 1:A) {
##
         Rec[t,a] <- tot_Rec[t] * pi_vec[t,a];</pre>
##
       7
##
     } ## end t loop over year
##
##
     ## get predicted calendar year returns by age
##
     ## matrix Run has dim [(n_yrs-age_min) x A]
     ## step 1: incomplete early broods
##
     ## first cal yr of this grp is first brood yr + age_min + age_skip
##
     for(i in 1:(age_max-age_min-age_skip)) {
##
##
       ## projected recruits
##
       for(a in 1:(i+age_skip)) {
##
         Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
       }
##
       ## imputed recruits
##
       for(a in (i+1+age_skip):A) {
##
         lnRec[i,a] ~ dnorm(Rec mu, Rec tau);
##
         Run[i,a] <- exp(lnRec[i,a]);</pre>
       }
##
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
       ## predicted age-prop vec for multinom
       for(a in 1:A) {
##
##
         age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
       ## multinomial for age comp
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
##
       lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
     }
##
##
##
     ## step 2: info from complete broods
     ## first cal yr of this grp is first brood yr + age_max
##
##
     for(i in (A-age_skip):(n_yrs-age_min-age_skip+n_fore)) {
       for(a in 1:A) {
##
```

```
##
         Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
       }
       ## total run size
##
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
       ## predicted age-prop vec for multinom
##
       for(a in 1:A) {
##
         age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
       }
##
##
       ## multinomial for age comp
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
##
       #lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
       lp_age[i] <- ifelse(i < n_yrs-age_min-age_skip+n_fore,</pre>
##
##
       logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]),0)
##
     }
##
##
     ## get predicted calendar year spawners
     ## first cal yr is first brood yr
##
##
     for(t in 1:(n_yrs+n_fore)) {
##
       ## obs model for spawners
##
       # Sp[t] <- max(10,tot_Run[t] - dat_harv[t]);
##
       est_harv[t] = ifelse(t > n_yrs,1,h_rate[t] * tot_Run[t]);
       dat_harv[t] ~ dlnorm(log(est_harv[t]), 20);
##
##
       Sp[t] = tot_Run[t] - est_harv[t];
       ln_Sp[t] \leftarrow log(Sp[t]);
##
##
       ln_dat_esc[t] ~ dnorm(ln_Sp[t], tau_s);
       lp_esc[t] <- ifelse(t < n_yrs + 1,logdensity.norm(ln_dat_esc[t],ln_Sp[t], tau_s),0);</pre>
##
##
##
## } ## end model description
##
## Warning in readLines(con, n = 1): incomplete final line found on 'C:/Users/cruff/Documents/projects/
## RWorkflow/Skagit-River-Steelhead-Forecast/jags/IPM_BH_cov_AR_resid.txt'
```

Fitting the models and generating the one year ahead forecasts

For the most recent 10 years (2014 - 2023), fit the model to data through year t-1 and generate a forecast for year t.

Before fitting the model in JAGS, we need to specify the MCMC control parameters.

```
## 1. MCMC control params
mcmc_ctrl <- list(
   chains = 4,
   length = 200000, #5e5,
   burn = 100000, #2e5,
   thin = 100#400
)

## total number of MCMC samples after burnin
mcmc_samp <- mcmc_ctrl$length*mcmc_ctrl$thin</pre>
```

Model selection

We evaluated model performance following three separate approaches outlined below.

Model selection via root mean squared error (RMSE) and mean absolute percent error (MAPE) model performance statistics.

```
tot_mods <- n_forecasts*n_mods
# qet escapement data
dat_esc_forecast <- dat_esc[which(dat_esc$year %in% seq(yr_begin,yr_end,1)),]</pre>
## get harvest data
dat_harv_forecast <- dat_harv[which(dat_harv$year %in% seq(yr_begin,yr_end,1)),]</pre>
## observed terminal run size
obs_trs <- dat_esc_forecast$escapement + dat_harv_forecast$catch</pre>
pred trs <- NULL
pred_trs_post <- vector("list", n_forecasts)</pre>
for(n in 1:n_mods){
  #n <- 1
  pred_esc <- NULL</pre>
  for(i in 1:(n_forecasts)){
    #i <- 1
    mod_res<-NULL
    mod_res<-as.matrix(readRDS(file.path(savedir,paste0(models[n],"_y",i,".rds"))))</pre>
    p_dat_post <- mod_res[,grep("Sp", colnames(mod_res))]</pre>
    p_dat <- round(median(p_dat_post[,dim(p_dat_post)[2]]))</pre>
    p_dat_post <- p_dat_post[,dim(p_dat_post)[2]]</pre>
    pred_esc[i] <- p_dat</pre>
    pred_trs_post[[i]] <- cbind(pred_trs_post[[i]],p_dat_post)</pre>
  pred_trs_mod <- pred_esc + 1#+ dat_harv_forecast$catch #you don't need to add catch in because it is
```

```
pred_trs <- cbind(pred_trs,pred_trs_mod)</pre>
  #names(pred_trs) <- paste(models[n], "_ ", "pred_trs", sep = "")</pre>
colnames(pred_trs) <- models</pre>
## compute model performance statistics
Error <- pred trs - obs trs
SE <- Error<sup>2</sup>
PE <- Error/obs trs
APE <- abs(PE)
LAR <- log(obs_trs/pred_trs)</pre>
RMSE <- apply(SE,2,function(x){sqrt(mean(x))})</pre>
MPE <- apply(PE,2,function(x){mean(x)})</pre>
MAPE <- apply(APE,2,function(x){mean(x)})
MSA \leftarrow apply(LAR, 2, function(x) \{100*(exp(median(abs(x)))-1)\})
model_selection <- data.frame(RMSE,MPE,MAPE,MSA)</pre>
weights<-apply(model_selection[,!colnames(model_selection)=="MPE"], 2,function(x) (1/x)/sum(1/x))</pre>
colnames(weights)<-paste0(colnames(weights), "_weight")</pre>
model_selection <-data.frame (model_selection, weights)
```

Model Selection Via Approximate Leave-Future-Out Cross Validation

Detailed methods described here: **link**. This approach is better suited to evaluating the predictive performance of Bayesian time series models than the more traditional model performance metrics such as RMSE and MAPE.

```
N=yr_last-yr_frst+1
L=N-n_forecasts
thres=0.1
esc_only="No"
userefits="Yes"
mod_fits<-loadmodfits(modelnames=models)

LF01<-approx_LF0(N=N,L=L,m=1,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[1]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot_ks(LF01$ks,N=N,L=L,thres=thres)
# plot_ks(LF01$ks_LOOIC,N=N,L=L,thres=thres)</pre>
```

LF02<-approx_LF0(N=N,L=L,m=2,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[2]

```
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot_ks(LFO2$ks, N=N,L=L,thres=thres)
# plot_ks(LFO2$k_LOOIC, N=N,L=L,thres=thres)
LF03<-approx_LF0(N=N,L=L,m=3,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[3]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot_ks(LFO3$ks, N=N,L=L,thres=thres)
# plot_ks(LF03$k_L00IC, N=N,L=L,thres=thres)
LF04<-approx_LF0(N=N,L=L,m=4,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[4]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot ks(LFO3$ks, N=N,L=L,thres=thres)
# plot_ks(LF03$k_L00IC, N=N,L=L,thres=thres)
LF05<-approx_LF0(N=N,L=L,m=5,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[5]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot ks(LF03$ks, N=N,L=L,thres=thres)
# plot_ks(LF03$k_L00IC, N=N,L=L,thres=thres)
LF06<-approx_LF0(N=N,L=L,m=6,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[6]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
```

Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.

Calculating Stacking Weights (finding model averaging weights based on one-step-ahead performance of weighted average models). This is an experimental calculation of stacking weights based on linear comminations of model forecasts to optimize one step ahead RMSE, MSA, or MAPE...it is not ready for prime time yet. An alternative is to figure out how to use methods similar to "stacking weights" here for LFOIC: link

Model Selection Via loo() and compare() with full table of results.

Note that elpd_diff will be negative (positive) if the expected predictive accuracy for the first (second) model is higher.

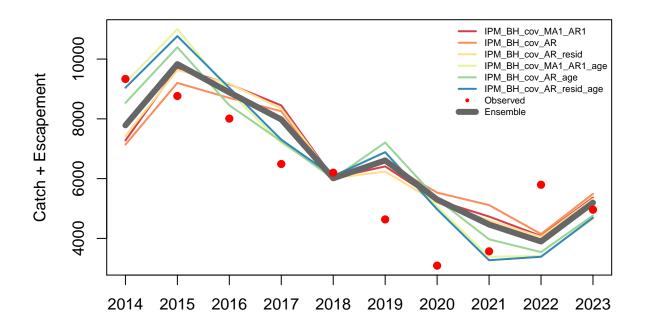
```
LOOIC <- vector("list", n_mods)
## extract log densities from JAGS objects
for(i in 1:n_mods) {
    #i <- 1
    ## convert mcmc.list to matrix
    tmp_lp <- as.matrix(readRDS(file.path(savedir,pasteO(models[i],"_y",11,".rds"))))
    ## extract pointwise likelihoods
    tmp_lp <- tmp_lp[,grepl("lp_", colnames(tmp_lp))]
    ## if numerical underflows, convert -Inf to 5% less than min(likelihood)
    if(any(is.infinite(tmp_lp))) {
        tmp_lp[is.infinite(tmp_lp)] <- NA
        tmp_min <- min(tmp_lp, na.rm = TRUE)
        tmp_lp[is.na(tmp_lp)] <- tmp_min * 1.05</pre>
```

Model Averaging and 2023 - 2024 forecast

Here is weighted average forecast for 2023 - 2024 based on the model LFOIC weights. We also compare predictions for return years 2014 - 2023 generated from each of the six models evaluated including an ensemble (weighted) to observed run size.

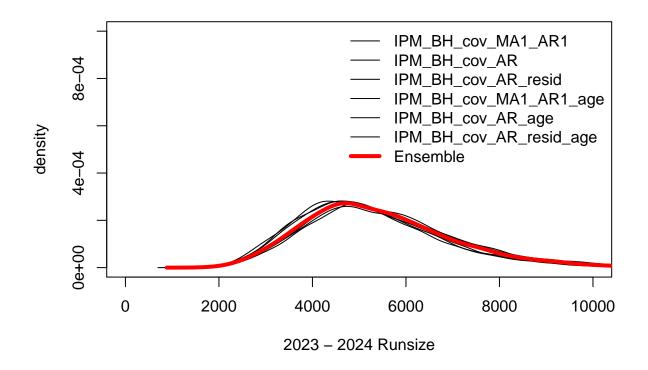
```
## extract median 2023 forecast from each model
f_dat<-data.frame(</pre>
  sort(unlist(mod_fits[[1]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[2]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[3]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[4]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[5]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[6]][,paste0("Sp","[",n_yrs+n_fore,"]")]))
colnames(f dat)<-models</pre>
model_selection[,"2024_forecast"] <- apply(f_dat,2,median)</pre>
weighted_forecast_dist <-(</pre>
  as.matrix(f_dat) %*% (as.vector(model_selection[,"LFOIC_weight"]))
ensemble_forecast_posterior_retrospective <- cbind(as.matrix(pred_trs_post[[5]]) %*% (as.vector(model_s
colnames(ensemble_forecast_posterior_retrospective)<-c("ensemble_forecast_posterior_2018",</pre>
                                                         "ensemble_forecast_posterior_2020",
                                                         "ensemble_forecast_posterior_2021",
                                                         "ensemble forecast posterior 2023",
                                                         "ensemble_forecast_posterior_2024")
```

```
ensemble_forecast_posterior_retrospective <- data.frame(ensemble_forecast_posterior_retrospective)</pre>
ensemble_forecast_posterior<-data.frame(weighted_forecast_dist)</pre>
colnames(ensemble_forecast_posterior)<-c("ensemble_forecast_posterior")</pre>
write.csv(ensemble_forecast_posterior_retrospective,file.path(savedir, "ensemble_forecast_posterior_retr
write.csv(ensemble_forecast_posterior,file.path(savedir,"ensemble_forecast_posterior.csv"),row.names = 1
weighted_forecast_quantiles<-quantile(weighted_forecast_dist,c(0.025,0.25,0.50,0.75,0.975))</pre>
weighted_forecast<-weighted_forecast_quantiles[3]</pre>
print(model selection)
                               RMSE
##
                                            MPE
                                                     MAPE
                                                                MSA RMSE_weight MAPE_weight MSA_weight
                           1506.449 0.15060990 0.2595526 29.22246
                                                                                  0.1528328 0.09459257
## IPM_BH_cov_MA1_AR1
                                                                      0.1622170
## IPM_BH_cov_AR
                           1537.450 0.15921507 0.2707490 28.95770
                                                                     0.1589460
                                                                                  0.1465127 0.09545743
## IPM_BH_cov_AR_resid
                           1431.184 0.13865942 0.2457822 27.14594
                                                                      0.1707479
                                                                                  0.1613956 0.10182839
## IPM_BH_cov_MA1_AR1_age 1452.131 0.10448264 0.2152409 11.06835
                                                                                  0.1842966 0.24974165
                                                                      0.1682848
## IPM_BH_cov_AR_age
                           1448.173 0.12019277 0.2293002 11.48066
                                                                      0.1687448
                                                                                  0.1729966 0.24077255
## IPM BH cov AR resid age 1428.577 0.09751065 0.2179979 12.70282
                                                                      0.1710595
                                                                                  0.1819658 0.21760741
                                        LFOIC delta_LFOIC LFOIC_weight stacking_weights LOOIC
## IPM_BH_cov_MA1_AR1
                           6.842469 -13.68494
                                                  1.991150
                                                             0.13967569
                                                                                   0.0000 857.07
                           7.838044 -15.67609
## IPM_BH_cov_AR
                                                  0.000000
                                                             0.37800153
                                                                                   0.0000 871.63
## IPM_BH_cov_AR_resid
                           7.004595 -14.00919
                                                  1.666898
                                                             0.16425977
                                                                                   0.1476 837.13
## IPM_BH_cov_MA1_AR1_age
                                                                                   0.8524 907.32
                           6.495261 -12.99052
                                                  2.685566
                                                             0.09870299
## IPM_BH_cov_AR_age
                           6.382780 -12.76556
                                                  2.910528
                                                             0.08820242
                                                                                   0.0000 915.12
## IPM_BH_cov_AR_resid_age 6.779545 -13.55909
                                                  2.116997
                                                             0.13115760
                                                                                   0.0000 904.58
                           delta_LOOIC LOOIC_weight 2024_forecast
## IPM_BH_cov_MA1_AR1
                                  19.94
                                              0.0000
                                                          5374.166
                                  34.50
                                              0.0000
                                                          5269.556
## IPM_BH_cov_AR
                                  0.00
## IPM_BH_cov_AR_resid
                                              0.9758
                                                          5380.146
## IPM_BH_cov_MA1_AR1_age
                                 70.19
                                              0.0000
                                                          5010.022
## IPM_BH_cov_AR_age
                                  77.99
                                              0.0000
                                                          4939.139
## IPM_BH_cov_AR_resid_age
                                  67.45
                                              0.0241
                                                          5023.906
print("The model-averaged forecast is:")
## [1] "The model-averaged forecast is:"
print(weighted_forecast_quantiles)
##
       2.5%
                 25%
                          50%
                                    75%
                                           97.5%
## 2869.318 4286.212 5215.354 6369.499 9275.038
ensemble_median<-as.matrix(pred_trs)%*%(as.vector(model_selection[,"LFOIC_weight"]))
cols<-brewer.pal(length(models), "Spectral")</pre>
matplot(as.matrix(data.frame(pred_trs,ensemble_median)),type="1",lty=1,col=c(cols,"grey40"),lwd=c(rep(2
axis(1,1:n_forecasts,(yr_last-n_forecasts+1):(yr_last))
points(x=1:n_forecasts,y=obs_trs,cex=1.5,pch=20,col="red")
legend("topright",legend=c(models,"Observed","Ensemble"),lty=c(rep(1,length(models)),NA),col=c(cols,"re-
```



```
#density plot for final forecasts and ensemble
res<-apply(f_dat,2,function(x) density(x))
plot(x=1,y=1,ylim=c(0,0.001),xlim=c(0,10000),ylab="density",xlab="2023 - 2024 Runsize")
lapply(res,function(x) lines(x$y~x$x))</pre>
```

```
## $IPM_BH_cov_MA1_AR1
## NULL
## $IPM_BH_cov_AR
## NULL
## $IPM_BH_cov_AR_resid
## NULL
##
## $IPM_BH_cov_MA1_AR1_age
## NULL
##
## $IPM_BH_cov_AR_age
## NULL
##
## $IPM_BH_cov_AR_resid_age
## NULL
lines(density(weighted_forecast_dist)$y~density(weighted_forecast_dist)$x,lwd=4,col="red")
legend("topright",legend=c(models,"Ensemble"),lwd=c(rep(1,length(models)),4),col=c(rep("black",length(models)))
```

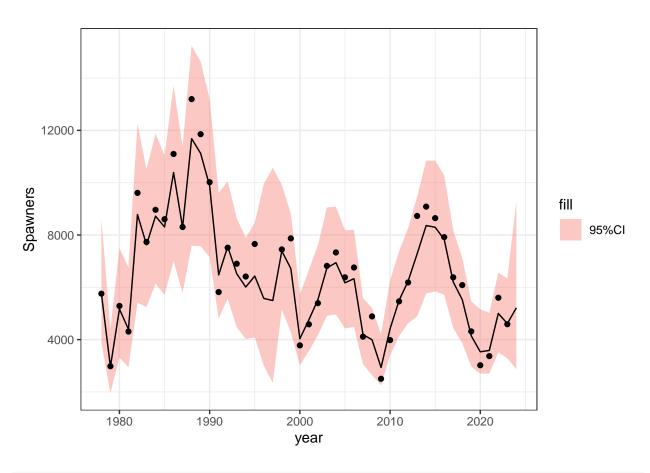


Now lets graph model estimated spawners, recruits, and recruits per spawner:

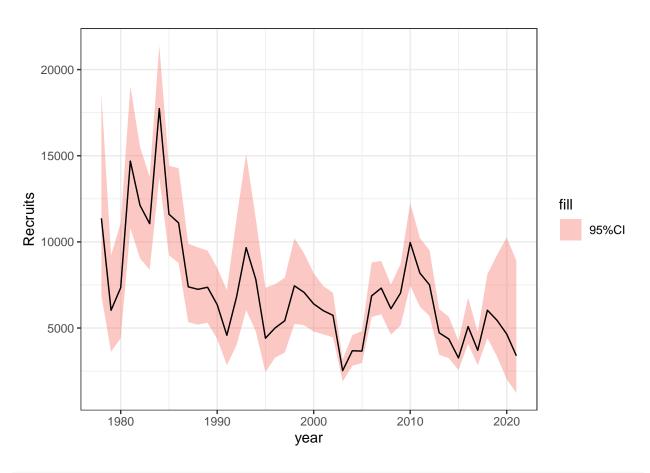
```
RSdat<-function(var,yrs){
  sims<-length(sort(unlist(mod_fits[[1]][,paste0(var,"[",i,"]")])))</pre>
  R_mat<-matrix(NA,ncol=length(yrs),nrow=sims)</pre>
  for(i in 1:length(yrs)){
    Rdat<-data.frame(
      sort(unlist(mod_fits[[1]][,paste0(var,"[",min(yrs)+i-1,"]")])),
      sort(unlist(mod_fits[[2]][,paste0(var,"[",min(yrs)+i-1,"]")])),
      sort(unlist(mod_fits[[3]][,paste0(var,"[",min(yrs)+i-1,"]")])),
      sort(unlist(mod_fits[[4]][,paste0(var,"[",min(yrs)+i-1,"]")])),
      sort(unlist(mod_fits[[5]][,paste0(var,"[",min(yrs)+i-1,"]")])),
      sort(unlist(mod_fits[[6]][,paste0(var,"[",min(yrs)+i-1,"]")]))
    )
    dim(Rdat)
    weighted R <-(
      as.matrix(Rdat) %*% (as.vector(model selection[,"LFOIC weight"]))
    )
    R_mat[,i]<-weighted_R</pre>
    colnames(R_mat)<-yrs</pre>
  }
  return(R_mat)
Sdat<-RSdat(var="Sp",yrs=c(1:(n_yrs+n_fore)))</pre>
Rdat<-exp(RSdat(var="tot_ln_Rec",yrs=c(1:(n_yrs+n_fore-3))))</pre>
RSdat <-exp(RSdat(var="ln_RS",yrs=c(1:(n_yrs+n_fore-3))))
```

```
library(tidyverse)
## Warning: package 'tidyverse' was built under R version 4.0.5
## Warning: package 'tibble' was built under R version 4.0.5
## Warning: package 'tidyr' was built under R version 4.0.5
## Warning: package 'dplyr' was built under R version 4.0.5
## Warning: package 'forcats' was built under R version 4.0.5
Sdat<-as.data.frame(Sdat)%>%
 pivot_longer(cols=everything())%>%
 rename(year=name)%>%
  group_by(year) %>%
  summarise(SpawnerAbundance = quantile(value, c(0.025, 0.5, 0.975)), q = c(0.025, 0.5, 0.975))%%
  mutate(year=as.numeric(year)+yr_frst-1)%>%
  pivot_wider(names_from = q, values_from = SpawnerAbundance)
Rdat<-as.data.frame(Rdat)%>%
  pivot_longer(cols=everything())%>%
  rename(year=name)%>%
  group_by(year) %>%
  summarise(RecruitAbundance = quantile(value, c(0.025, 0.5, 0.975)), q = c(0.025, 0.5, 0.975))%%
  mutate(year=as.numeric(year)+yr_frst-1)%>%
  pivot_wider(names_from = q, values_from = RecruitAbundance)
RSdat<-as.data.frame(RSdat)%>%
  pivot_longer(cols=everything())%>%
  rename(year=name)%>%
  group_by(year) %>%
  summarise(RS = quantile(value, c(0.025, 0.5, 0.975)), q = c(0.025, 0.5, 0.975))%>%
  mutate(year=as.numeric(year)+yr_frst-1)%>%
  pivot_wider(names_from = q, values_from = RS)
ggplot(Sdat,aes(x=year,y='0.5'))+
  geom_ribbon(aes(ymin = '0.025', ymax = '0.975',fill="95%CI"),alpha=0.4)+
  geom line()+
  geom_point(dat_esc,mapping=aes(y=escapement,x=year))+
 theme_bw()+
 ylab("Spawners")
```

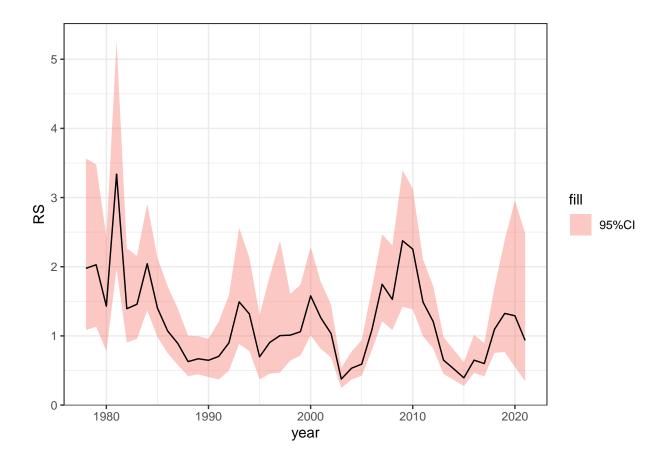
Warning: Removed 2 rows containing missing values (geom_point).



```
ggplot(Rdat,aes(x=year,y='0.5'))+
  geom_ribbon(aes(ymin = '0.025', ymax = '0.975',fill="95%CI"),alpha=0.4)+
  geom_line()+
  theme_bw()+
  ylab("Recruits")
```



```
ggplot(RSdat,aes(x=year,y='0.5'))+
  geom_ribbon(aes(ymin = '0.025', ymax = '0.975',fill="95%CI"),alpha=0.4)+
  geom_line()+
  theme_bw()+
  ylab("RS")
```



Compare sib adjusted forecasts with unadjusted

Currently implemented for only two of the six model including the AR1MA1 recruitment residuals with random walk maturation and the AR1 recruitment residuals with random walk maturation.

```
## observed terminal run size
obs_trs <- dat_esc_forecast$escapement + dat_harv_forecast$catch</pre>
pred trs <- NULL
pred_trs_adj <- NULL</pre>
for(n in 1:n_mods){
  #n <- 1
 pred_esc <- NULL</pre>
 pred_esc_adj <- NULL</pre>
 c <- 0
  f < -2
  adj<- NULL
  for(i in 1:(n_forecasts)){
    #i <- 1
    #range of years. Last year in range
    dat_yrs <- seq(yr_frst,(yr_last - n_forecasts + c),1)</pre>
    ## number of years of data
    n_yrs <- length(dat_yrs)</pre>
  mod res<-NULL
  #model refit and one step ahead forecast
  mod res pred<-as.matrix(readRDS(file.path(savedir,paste0(models[n]," y",i,".rds"))))</pre>
  #model refit for year i+1 to extract "observed" state
  mod_res_obs<-as.matrix(readRDS(file.path(savedir,paste0(models[n],"_y",i + 1,".rds"))))</pre>
  p_dat_pred <- mod_res_pred[,grep("Sp", colnames(mod_res_pred))]</pre>
  p_dat_pred <- round(median(p_dat_pred[,n_yrs + 1]))</pre>
 p_dat_obs <- mod_res_obs[,grep("Sp", colnames(mod_res_obs))]</pre>
 p_dat_obs <- round(median(p_dat_obs[,n_yrs + 2]))</pre>
 pred_esc[i] <- p_dat_pred</pre>
  #take same model and implement sibling adjustment
  ## forecast for year t
  p_dat_pred_adj <- cbind(((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",1,"]")])),</pre>
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",2,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",3,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",4,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",5,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",6,"]")])))
  p_dat_pred_adj_sum <- apply(p_dat_pred_adj,1,FUN = "sum")</pre>
  median(p_dat_pred_adj_sum)
```

```
## observation for year t+1
  p_dat_obs_adj <- cbind(((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",1,"]")])),</pre>
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",2,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",3,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",4,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",5,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",6,"]")])))
  ## forecast for year t+1 to be adjusted
  p_dat_pred_t <- cbind(((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",1,"]")])),</pre>
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",2,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",3,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",4,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",5,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",6,"]")])))
  \#median(p\_dat\_obs\_adj[,a-1])/median(p\_dat\_pred\_adj[,a-1])
    p_dat_pred_adj_a <- NULL</pre>
    \#adj a \leftarrow NULL
    for (a in 2:A){
      \#adj_a[a] \leftarrow median(as.vector(p_dat_obs_adj[,a-1])/as.vector(p_dat_pred_adj[,a-1]))
      p_dat_pred_adj_a_temp <- as.vector(p_dat_pred_t[,a])*(as.vector(p_dat_obs_adj[,a-1])/as.vector(p_
      p_dat_pred_adj_a <- cbind(p_dat_pred_adj_a,p_dat_pred_adj_a_temp)</pre>
    }
    \#adj \leftarrow rbind(adj, t(adj_a))
    p_dat_pred_adj_a <- apply(p_dat_pred_adj_a,1,FUN = sum)</pre>
    pred_esc_adj[f] <- round(median(p_dat_pred_adj_a))</pre>
    #pred_esc_adj[f] <- median(apply(p_dat_pred_t,1,FUN = "sum"))</pre>
    f \leftarrow f+1
    c <- c+1
  }#next forecast
  pred_trs_mod <- pred_esc + 1#+ dat_harv_forecast$catch #you don't need to add catch in because it is
  pred_trs_mod_adj <- pred_esc_adj #not adjusted because adjusted prediction comes from "Run" rather th
  pred_trs <- cbind(pred_trs,pred_trs_mod)</pre>
  pred_trs_adj <- cbind(pred_trs_adj,pred_trs_mod_adj)</pre>
  #names(pred_trs) <- paste(models[n], "_ ", "pred_trs", sep = "")</pre>
}
colnames(pred_trs) <- models</pre>
colnames(pred_trs_adj) <- paste(models,"_","sib_adjust",sep = "")</pre>
pred_trs_adj <- pred_trs_adj[-11,]</pre>
```

```
pred_trs<-data.frame(pred_trs,pred_trs_adj)</pre>
pred_trs<-pred_trs[-1,]</pre>
## compute model performance statistics
Error <- pred_trs - obs_trs[-1]</pre>
SE <- Error<sup>2</sup>
PE <- Error/obs_trs[-1]
APE <- abs(PE)
LAR <- log(obs_trs/pred_trs)
RMSE <- apply(SE,2,function(x){sqrt(mean(x))})</pre>
MPE <- apply(PE,2,function(x){mean(x)})</pre>
MAPE <- apply(APE,2,function(x){mean(x)})</pre>
MSA \leftarrow apply(LAR, 2, function(x) \{100*(exp(mean(abs(x))-1))\})
model_selection <- data.frame(RMSE,MPE,MAPE,MSA)</pre>
weights<-apply(model_selection[,!colnames(model_selection)=="MPE"], 2,function(x) (1/x)/sum(1/x))</pre>
colnames(weights)<-paste0(colnames(weights),"_weight")</pre>
model_selection<-data.frame(model_selection,weights)</pre>
print(model_selection)
##
                                            RMSE
                                                        MPE
                                                                  MAPE
                                                                            MSA RMSE_weight MAPE_weight
## IPM_BH_cov_MA1_AR1_age
                                        1530.140 0.1175441 0.2377043 40.65169
                                                                                   0.3113602
                                                                                                0.2936797
## IPM_BH_cov_AR_resid_age
                                        1502.703 0.1118211 0.2387440 47.27748
                                                                                   0.3170452
                                                                                                0.2924008
## IPM_BH_cov_MA1_AR1_age_sib_adjust
                                        2541.719 0.2355240 0.3378344 59.71886
                                                                                   0.1874419
                                                                                                0.2066365
## IPM_BH_cov_AR_resid_age_sib_adjust 2587.118 0.2320653 0.3367808 64.53250
                                                                                                0.2072830
                                                                                   0.1841527
                                        MSA_weight
## IPM_BH_cov_MA1_AR1_age
                                         0.3154065
## IPM_BH_cov_AR_resid_age
                                         0.2712032
## IPM_BH_cov_MA1_AR1_age_sib_adjust
                                         0.2147028
## IPM_BH_cov_AR_resid_age_sib_adjust 0.1986876
#sib adjust plot
cols<-brewer.pal(length(models)*2, "Spectral")</pre>
matplot(as.matrix(data.frame(pred_trs)),type="l",lty=1,col=c(cols,"grey40"),lwd=c(rep(2,length(models)*
axis(1,1:(n_forecasts-1),(yr_last-n_forecasts+2):(yr_last))
points(x=1:(n_forecasts-1),y=obs_trs[-1],cex=1.5,pch=20,col="red")
legend("topright",legend=c(rownames(model_selection),"Observed"),lty=c(rep(1,length(models)*2),NA),col=
```

