R2. Model fitting and evaluation

2020 - 2021 Skagit River steelhead forecast.

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| This is version 0.21.01.29. | |

Requirements

All analyses require the R software (v3.4.3) for data retrieval, data processing, and summarizing model results, and the JAGS software (v4.2.0) for Markov chain Monte Carlo (MCMC) simulation. Please note that some of the R code below may not work with older versions of JAGS due to some changes in the ways that arrays are handled.

We also need a few packages that are not included with the base installation of R, so we begin by installing them (if necessary) and then loading them.

```
if(!require("here")) {
  install.packages("here")
  library("here")
}
if(!require("readr")) {
```

```
install.packages("readr")
  library("readr")
if(!require("rjags")) {
  install.packages("rjags")
  library("rjags")
if(!require("loo")) {
  install.packages("loo")
  library("loo")
}
if(!require("ggplot2")) {
  install.packages("ggplot2")
  library("ggplot2")
if(!require("coda")) {
  install.packages("coda")
  library("coda")
if(!require("shinystan")) {
  install.packages("shinystan")
  library("shinystan")
if(!require("R2jags")) {
  install.packages("R2jags")
  library("R2jags")
if(!require("dclone")) {
  install.packages("dclone")
  library("dclone")
if(!require("snow")) {
  install.packages("snow")
  library("snow")
}
if(!require("rstan")) {
  install.packages("rstan")
  library("rstan")
}
if(!require("RColorBrewer")) {
  install.packages("RColorBrewer")
  library("RColorBrewer")
if(!require("gtools")) {
  install.packages("gtools")
  library("gtools")
}
## set directory locations
datadir <- here("data")</pre>
jagsdir <- here("jags")</pre>
analdir <- here("analysis")</pre>
savedir <- here("analysis/cache")</pre>
```

We also need a couple of helper functions.

```
## better round
Re2prec <- function(x, map = "round", prec = 1) {
  ## 'map' can be "round", "floor", or "ceiling"
  ## 'prec' is nearest value (eg, 0.1 means to nearest tenth; 1 gives normal behavior)
 if(prec<=0) { stop("\"prec\" cannot be less than or equal to 0") }</pre>
 do.call(map,list(x/prec))*prec
## wrapper function to fit JAGS models & rearrange output
fit_jags <- function(model, data, params, inits, ctrl, dir = jagsdir) {</pre>
  jm <- jags.model(file.path(jagsdir, model),</pre>
                    data,
                    inits,
                    ctrl$chains,
                    ctrl$burn,
                    quiet = TRUE)
 return(coda.samples(jm, params, ctrl$length, ctrl$thin))
}
#alternative wrapper to fit model in parallel; one chain per core
fit_jags2<-function(model,data,params,inits,ctrl,dir=jagsdir){</pre>
  cl <- makeCluster(3, type = "SOCK")</pre>
  inits2 <- jags.fit(data=data,</pre>
                      params=params,
                      model=file.path(jagsdir, model),
                      inits=inits,
                      n.chains=ctrl$chains,
                      n.adapt = 0,
                      n.update = 0,
                      n.iter = 0)$state(internal = TRUE)
  jm <- jags.parfit(cl=cl,</pre>
                     data = data,
                     params = params,
                     model = file.path(jagsdir, model),
                     inits = inits2,
                     n.adapt = ctrl$burn*0.5,
                     n.update = ctrl$burn*0.5,
                     n.iter = ctrl$length-ctrl$burn,
                     thin = ctrl$thin,
                     n.chains = ctrl$chains
  stopCluster(cl)
  return(jm)
#generate summary stats file from MCMC object
sum_stats<-function(mcmclist){</pre>
 ESS<-apply(as.matrix(mcmclist),2,ess_bulk)
  Rhat<-apply(as.matrix(mcmclist),2,Rhat)</pre>
  summary_stats<-summary(mcmclist)</pre>
  summary_stats<-data.frame(summary_stats$statistics,summary_stats$quantiles,ESS,Rhat)
```

```
# functions for approximate LFO
# many functions modified from:
# https://github.com/paul-buerkner/LFO-CV-paper/blob/master/case-study-LFO-CV.Rmd
#load complete model fits & model refits with subset data
loadmodfits<-function(modelnames){</pre>
  mod_fits<-list(NULL)</pre>
  for(i in 1:length(modelnames)){
    mod_fits[[i]] <- readRDS(file.path(savedir,paste0(modelnames[i],"_y",n_forecasts+1,".rds")))</pre>
    \#mod_fits[[i]] \leftarrow readRDS(file.path(savedir,pasteO("fit_",modelnames[i],".rds")))
  return(mod_fits)
#refits
loadrefits<-function(refitname,N,L){</pre>
  numrefits<-N-L+1
  re fits<-list()
  for(i in 1:numrefits){
     re_fits[[i]] <- readRDS(file.path(savedir,paste0(refitname,"_y",i,".rds")))</pre>
  return(re_fits)
}
# more stable than log(sum(exp(x)))
log_sum_exp <- function(x) {</pre>
  \max x < -\max(x)
  \max_{x} + \log(\sup(\exp(x - \max_{x})))
# more stable than log(mean(exp(x)))
log_mean_exp <- function(x) {</pre>
  log_sum_exp(x) - log(length(x))
# compute log of raw importance ratios
# sums over observations *not* over posterior samples
sum log ratios <- function(ll, ids = NULL) {</pre>
  if (!is.null(ids)) ll <- ll[, ids, drop = FALSE]</pre>
  - rowSums(11)
# for printing comparisons later
rbind_print <- function(...) {</pre>
  round(rbind(...), digits = 2)
}
#function to extract log likelihood from fitted model
extract_log_lik<-function(m,esc_only,N,mod_fits){</pre>
  #extract pontwise log likelihoods
  tmp_lp <- as.matrix(mod_fits[[m]])</pre>
  ## extract pointwise likelihoods
  tmp_lp <- tmp_lp[,grepl("lp_", colnames(tmp_lp))]</pre>
  ## if numerical underflows, convert -Inf to 5% less than min(likelihood)
```

```
if(any(is.infinite(tmp_lp))) {
    tmp_lp[is.infinite(tmp_lp)] <- NA</pre>
    tmp_min <- min(tmp_lp, na.rm = TRUE)</pre>
    tmp_lp[is.na(tmp_lp)] \leftarrow tmp_min * 1.05
  if(esc only =="Yes"){
    tmp_lp<-tmp_lp[,grepl("esc", colnames(tmp_lp))]</pre>
  #qet yrs assoc
  names_loglik<-data.frame(strsplit(colnames(tmp_lp),"\\[|\\]"))</pre>
  yrnames<-as.numeric(names_loglik[2,])</pre>
  loglik <- matrix(NA,ncol=N,nrow=dim(tmp_lp)[1])</pre>
  for(i in 1:N){
    if(!is.null(ncol(tmp_lp[,yrnames==i]))){
      loglik[,i] = apply(tmp_lp[,yrnames==i],1,sum)
    }else(loglik[,i] = tmp_lp[,yrnames==i])
  return(loglik)
}
approx_LFO<-function(N=N,L,m=m,esc_only,mod_fits,userefits,refitname,thres){
  loglik = extract_log_lik(m=m, esc_only = esc_only, N=N, mod_fits = mod_fits)
  ## look at Pareto k's
  k_LOOIC <- pareto_k_values(loo(loglik))[(L+1):N]
  if(userefits=="Yes"){
    re_fits =loadrefits(refitname=refitname, N=N, L=L)
  i refit <- L
  refits <- L
  ks <- NULL
  approx_elpds_1sap <- rep(NA, N)
  for (i in (N - 1):L) {
    logratio <- sum_log_ratios(loglik, (i + 1):N)</pre>
    psis_obj <- suppressWarnings(psis(logratio))</pre>
    k<-pareto_k_values(psis_obj)
    ks \leftarrow c(ks, k)
    if(k>thres & userefits=="Yes"){
    #use_refit of model based on the first[i] observations
      i_refit <- i
      refits <- c(refits, i)
      loglik = extract_log_lik(m=(i+1)-L+1, esc_only = esc_only,N=N,mod_fits = re_fits)
      approx_elpds_1sap[i + 1] <- log_mean_exp(loglik[, i + 1])
    }else{
      lw <- weights(psis_obj, normalize = TRUE)[, 1]</pre>
      approx_elpds_1sap[i + 1] <- log_sum_exp(lw + loglik[, i + 1])</pre>
  }
  results<-list(approx_elpds_1sap,ks,k_L00IC)</pre>
  names(results)<-c("LFO","ks","k_LOOIC")</pre>
  return(results)
```

```
plot_ks <- function(ks, thres = 0.7,N,L) {</pre>
  ids = N:(L + 1)
  dat ks <- data.frame(ks = ks, ids = ids)
  ggplot(dat ks, aes(x = ids, y = ks)) +
    geom_point(aes(color = ks > thres), shape = 3, show.legend = FALSE) +
    geom_hline(yintercept = thres, linetype = 2, color = "red2") +
    scale_color_manual(values = c("cornflowerblue", "darkblue")) +
    labs(x = "Data point", y = "Pareto k") +
    ylim(-0.5, max(dat_ks$ks))
#function for printing out a read text file
processFile = function(filepath) {
  con = file(filepath, "r")
  while ( TRUE ) {
    line = readLines(con, n = 1)
    if ( length(line) == 0 ) {
      break
    cat(paste0(noquote(line)),"\n")
  }
  close(con)
}
#calculate stacking weights
find_stack_weights<-function(tau,metric,n,initial_weights,preds,obs){</pre>
  tweights <- initial_weights
  preds<-as.matrix(preds)</pre>
  obs<-obs
  tau=tau
  skill_list<-c(NULL)</pre>
  metric=metric
  for(i in 1:n){
    pred_trs_ensemble<- preds %*% as.vector(tweights)</pre>
    Error <- pred_trs_ensemble - obs</pre>
    SE <- Error<sup>2</sup>
    PE <- Error/obs_trs
    APE <- abs(PE)
    LAR <- log(obs_trs/pred_trs_ensemble)
    RMSE <- apply(SE,2,function(x){sqrt(mean(x))})</pre>
    MPE <- apply(PE,2,function(x){mean(x)})</pre>
    MAPE <- apply(APE,2,function(x){mean(x)})
    MSA \leftarrow apply(LAR, 2, function(x) \{100*(exp(mean(abs(x))-1))\})
    if(i==1){
      skill=get(metric)
      weights=tweights
    }
    if(get(metric)<skill){</pre>
      skill=get(metric)
      weights=tweights
    }
    skill_list<-c(skill_list,min(get(metric),skill))</pre>
    keep<-rbinom(1,prob=skill/get(metric),1)</pre>
```

```
if(keep==1){tweights=tweights }else(tweights=weights)
    tweights = rdirichlet(n=1,alpha = tweights*tau+0.001)
  results <-list(weights, skill, skill_list)
  return(results)
#function to fit or load modelfits
fit_load_mods<-function(models){</pre>
  ## empty list for fits
  mod_fits <- vector("list", n_mods*(n_forecasts+1))</pre>
  ## counter to index fitted jags models (33 in total: 3 models x 11 1 year ahead forecasts including u
  ## return year)
  t <- 1
  for(n in 1:n_mods){
    ## counter to index data to feed model for year specific forecasts
    ## first forecast will be for 10 years prior to the most recent return year;
    ## last forecast will be current forecast for the upcoming return year
    c <- 0
    #n <-2
    model <- models[n]
    for(i in 1:(n forecasts+1)){
      if(file.exists(file.path(savedir,paste(model,"_","y",i,".rds",sep = "")))) {
        mod_fits[[t]] <- readRDS(file.path(savedir,paste(model,"_","y",i,".rds",sep = "")))</pre>
        c < - c + 1
        t < -t + 1
    } else { ## else, fit & save
          ## cnt & time stamp
          cat("Count =", t, "; Time =", round(((proc.time()-timer_start)/60)["elapsed"], 1), "\n",
              file="cnt_time.txt", append=TRUE)
          #range of years. Last year in range
          dat_yrs <- seq(yr_frst,(yr_last - n_forecasts + c),1)</pre>
          ## number of years of data
          n_yrs <- length(dat_yrs)</pre>
          ## get first & last years
          yr_frst_forecast <- min(dat_yrs)</pre>
          yr_last_forecast <- max(dat_yrs)</pre>
          ## get escapement data
          dat_esc_forecast <- dat_esc[which(dat_esc$year %in% dat_yrs),]</pre>
          ## log of escapement
          ln_dat_esc <- c(log(dat_esc_forecast$escapement),rep(NA,n_fore))</pre>
          ## get age data
```

```
dat_age_forecast <- dat_age[which(dat_age$year %in% dat_yrs),]</pre>
## drop year col & first age_min+age_skip rows
dat_age_forecast <- dat_age_forecast[-(1:(age_min+age_skip)),-1]</pre>
## add row(s) of NA's for forecast years
if(n fore > 0) {
  dat_age_forecast <- rbind(dat_age_forecast,</pre>
                             matrix(0, n fore, A,
                                     dimnames = list(n_yrs+seq(n_fore), colnames(dat_age_forecas)
## total num of age obs by cal yr
dat_age_forecast[,"sum"] <- apply(dat_age_forecast, 1, sum)</pre>
## row indices for any years with no obs age comp
idx_NA_yrs <- which(dat_age_forecast$sum<A, TRUE)</pre>
## replace 0's in yrs w/o any obs with NA's
dat_age_forecast[idx_NA_yrs,(1:A)] <- NA</pre>
## change total in yrs w/o any obs from 0 to A to help dmulti()
dat_age_forecast[idx_NA_yrs,"sum"] <- A</pre>
## convert class
dat_age_forecast <- as.matrix(dat_age_forecast)</pre>
## get harvest data
dat_harv_forecast <- dat_harv[which(dat_harv$year %in% dat_yrs),]</pre>
## drop year col & first age_max rows
dat_harv_forecast <- c(dat_harv_forecast$catch,rep(NA,n_fore))</pre>
## get covariate data
dat_cvrs_forecast <- dat_cvrs[which(dat_cvrs$year <= yr_last + n_fore - age_min),1:4]
## drop year col
dat_cvrs_forecast <- dat_cvrs_forecast[,-1]</pre>
## transform the covariates to z-scores
scl_cvrs_forecast <- scale(dat_cvrs_forecast)</pre>
## total number of covariates
n_cov <- dim(dat_cvrs_forecast)[2]</pre>
## ----jaqs_setup-----
## 1. Data to pass to JAGS
dat_jags <- list(dat_age = dat_age_forecast,</pre>
                  ln_dat_esc = ln_dat_esc,
                  dat_harv = dat_harv_forecast,
                  A = A,
                  age_min = age_min,
                  age_max = age_max,
                  age_skip = age_skip,
                  n_yrs = n_yrs,
                  n_fore = n_fore)
## 2. Model params/states for JAGS to return
      These are specific to the process model,
      so we define them in 'par_jags' below.
if(model == "IPM_BH_cov_AR"|model == "IPM_BH_cov_AR_resid"){
  init_vals_cov <- function() {</pre>
```

```
list(alpha = 5,
       beta_inv = exp(mean(ln_dat_esc, na.rm = TRUE)),
       gamma = rep(0, 3),
       pi_tau = 10,
       pi_eta = rep(1,A),
       pi_vec = matrix(c(0.01, 0.35, 0.47, 0.15, 0.01, 0.01),
                        n_yrs-age_min+n_fore, A,
                        byrow = TRUE),
       Rec_mu = log(1000),
       Rec_sig = 0.1,
       tot_ln_Rec = rep(log(1000), n_yrs - age_min + n_fore),
       phi_prior = 0.5,
       innov_1 = 0)
  }
  ## params/states to return
  par_jags<- c("alpha", "E_BH_a", "ln_BH_a",</pre>
                 "beta",
                 "gamma",
                 "Sp", "Rec", "tot_ln_Rec", "ln_RS",
                 "pi_eta", "pi_tau",
                 "sigma_r", "sigma_s", "w", "res_ln_Rec",
                 "lp_age","lp_esc","phi","Run"
                 )
}else{
  init_vals_cov <- function() {</pre>
    list(alpha = 5,
         beta_inv = exp(mean(ln_dat_esc, na.rm = TRUE)),
         gamma = rep(0, 3),
         pi_tau = 10,
         pi_eta = rep(1,A),
         \# pi\_vec = matrix(c(0.01, 0.35, 0.47, 0.15, 0.01, 0.01),
                             n_yrs-age_min+n_fore, A,
                             byrow = TRUE),
         Rec_mu = log(1000),
         Rec_sig = 0.1,
         tot_ln_Rec = rep(log(1000), n_yrs - age_min + n_fore),
         phi_prior = 0.5, theta_res_prior = 0.5,
         innov 1 = 0)
    }
  ## params/states to return
  par_jags <- c("alpha", "E_BH_a", "ln_BH_a",</pre>
                 "beta",
                 "gamma",
                 "Sp", "Rec", "tot_ln_Rec", "ln_RS", "tot_Run",
                 "pi_eta", "pi_tau",
                 "sigma_r", "sigma_s", "res_ln_Rec", "w", "theta_res", "phi",
                 "lp_age","lp_esc","Run"
```

```
}#endif
        ## set of multi-covariate models
        cset <- colnames(scl_cvrs_forecast)</pre>
        dat_jags$n_cov <- length(cset)</pre>
        dat_jags$mod_cvrs <- scl_cvrs_forecast[1:(n_yrs-age_min+1), cset]</pre>
        ## fit model & save it
        \# mod\_fits[[t]] \leftarrow fit\_jags(paste(model,".txt",sep = ""), dat\_jags, par\_jags,
                                       init_vals_cov, mcmc_ctrl)
        mod_fits[[t]]<-fit_jags2(model=paste(model,".txt",sep = ""),</pre>
                             data=dat_jags,
                             params=par_jags,
                             inits=init_vals_cov,
                             ctrl=mcmc_ctrl
        saveRDS(mod_fits[[t]], file.path(savedir,paste(model,"_","y",i,".rds",sep = "")))
        summary_stats<-NULL</pre>
        summary_stats<-sum_stats(mcmclist= mod_fits[[t]])</pre>
        write.csv(summary_stats,file.path(savedir, paste(model,"_","y",i,"_summary_stats.csv",sep = "
        c < - c + 1
        t < -t + 1
    }## end if
  }##next forecast year(i)
}## next model(n)
return(mod_fits)
```

User inputs

We begin by supplying values for the following parameters, which we need for model fitting and evaluation.

```
## first & last years of fish data
yr_frst <- 1978
yr_last <- 2020

## min & max adult age classes
age_min <- 3
age_max <- 8

## years (if any) of age-comp to skip; see below
age_skip <- 0

## number of years ahead for run forecasts from the most recent year of data
n_fore <- 1

## number of recent year forecasts</pre>
```

```
n_forecasts <- 10

## first year of 1 step ahead forecast
yr_begin <- 2011

## last year of 1 step ahead forecast
yr_end <- 2020

## upper threshold for Gelman & Rubin's potential scale reduction factor (Rhat).
Rhat_thresh <- 1.1</pre>
```

Next we specify the names of three necessary data files containing the following information:

- 1. observed total number of adult spawners (escapement) by year;
- 2. observed age composition of adult spawners by year;
- 3. observed total harvest by year;

```
## 1. file with escapement data
## [n_yrs x 2] matrix of obs counts; 1st col is calendar yr
fn_esc <- "skagit_sthd_esc.csv"

## 2. file with age comp data
## [n_yrs x (1+A)]; 1st col is calendar yr
fn_age <- "skagit_sthd_age.csv"

## 3. file with harvest data
## [n_yrs x 2] matrix of obs catch; 1st col is calendar yr
fn_harv <- "skagit_sthd_catch.csv"</pre>
```

Loading the fish data

Here we load in the first three data files and do some simple calculations and manipulations. First the spawner data:

```
## escapement
dat_esc <- read_csv(file.path(datadir, fn_esc))
## years of data
dat_yrs <- dat_esc$year

## number of years of data
n_yrs <- length(dat_yrs)

## log of escapement
ln_dat_esc <- c(log(dat_esc$escapement),rep(NA,n_fore))</pre>
```

Next the age composition data:

```
## age comp data
dat_age <- read_csv(file.path(datadir, fn_age))
## num of age classes
A <- age_max - age_min + 1

# ## drop year col & first age_min+age_skip rows
# dat_age <- dat_age[-(1:(age_min+age_skip)),-1]
#</pre>
```

```
# ## add row(s) of NA's for forecast years
# if(n_fore > 0) {
   dat_age <- rbind(dat_age,</pre>
                      matrix(0, n_fore, A,
#
                              dimnames = list(n_yrs+seq(n_fore),
#
                                              colnames(dat age))))
# }
# ## total num of age obs by cal yr
# dat_age[,"sum"] <- apply(dat_age, 1, sum)</pre>
# ## row indices for any years with no obs age comp
# idx_NA_yrs <- which(dat_age$sum<A, TRUE)</pre>
# ## replace 0's in yrs w/o any obs with NA's
# dat age[idx NA yrs, (1:A)] <- NA
# ## change total in yrs w/o any obs from 0 to A to help dmulti()
\# dat_aqe[idx_NA_yrs,"sum"] \leftarrow A
# ## convert class
# dat_age <- as.matrix(dat_age)</pre>
```

And then the harvest data:

```
## harvest
dat_harv <- read_csv(file.path(datadir, fn_harv))
## drop year col & first age_max rows
#dat_harv <- c(dat_harv$catch,rep(NA,n_fore))</pre>
```

Loading the covariates

Our analysis investigates 5 covariates as possible drivers of the population's instrinic growth rate:

- 1. Maximum river discharge in winter;
- 2. Minimum river discharge in summer;
- 3. North Pacific Gyre Oscillation;

All of the covariates are contained in the file /data/skagit_sthd_covars.csv. We will load and then standardize them to have zero-mean and unit-variance.

```
dat_cvrs <- read_csv(file.path(datadir, "skagit_sthd_covars.csv"))
## drop year col
# dat_cvrs <- dat_cvrs[,-1]
# ## transform the covariates to z-scores
# scl_cvrs <- as.matrix(scale(dat_cvrs))
# ## total number of covariates
# n_cov <- dim(dat_cvrs)[2]</pre>
```

Specifying models in JAGS

Now we can specify the model in JAGS. We fit a total one model, which we outline below, based on a beverton holt process model with covariates.

Beverton-Holt with covars and AR1 process errors (MA1 recruitment residuals). Here we will print out the model (contained in a separate text file)

```
processFile(file.path(jagsdir, "IPM_BH_cov_AR.txt"))
```

```
##
##
       model {
##
##
       ##----
##
       ## PRIORS
##
       ##----
##
       ## alpha = intrinsic productivity
       alpha ~ dnorm(0,0.001) T(0,);
##
##
       mu_BH_a <- log(alpha);</pre>
##
       E_BH_a \leftarrow mu_BH_a + sigma_r/(2 - 2*phi^2);
##
##
       ## strength of dens depend
       beta_inv ~ dnorm(0, 1e-9) T(0,);
##
##
       beta <- 1/beta_inv;</pre>
##
##
       ## covariate effects
##
       for(i in 1:n_cov) { gamma[i] ~ dnorm(0,0.01) }
##
##
       ## AR(1) coef for proc errors
       #phi ~ dunif(-0.999,0.999);
##
##
       #phi <- 0;</pre>
##
       phi_prior ~ dbeta(2,2);
       phi <- phi_prior*2-1;</pre>
##
##
       #phi ~ dunif(0,0.999);
##
##
       ## innovation in first year
##
       innov_1 ~ dnorm(0,tau_r*(1-phi*phi));
##
##
       ## process variance for recruits model
##
       sigma_r ~ dnorm(0, 2e-2) T(0,);
##
       tau_r <- 1/sigma_r;</pre>
##
##
       ## obs variance for spawners
##
       tau_s <- 1/sigma_s;</pre>
##
       sigma_s ~ dnorm(0, 0.001) T(0,);
##
       ## unprojectable early recruits;
##
##
       ## hyper mean across all popns
##
       Rec_mu ~ dnorm(0,0.001);
       ## hyper SD across all popns
##
##
       Rec_sig ~ dunif(0,100);
##
       ## precision across all popns
       Rec_tau <- pow(Rec_sig,-2);</pre>
##
##
       ## multipliers for unobservable total runs
##
       ttl_run_mu ~ dunif(1,5);
##
       ttl_run_tau ~ dunif(1,20);
##
##
       ## get total cal yr returns for first age_min yrs
##
       for(i in 1:(age_min+age_skip)) {
       ln_tot_Run[i] ~ dnorm(ttl_run_mu*Rec_mu,Rec_tau/ttl_run_tau);
##
##
       tot_Run[i] <- exp(ln_tot_Run[i]);</pre>
##
##
##
       ## maturity schedule
```

```
##
       ## unif vec for Dirch prior
##
       theta <-c(1,10,10,5,1,1)
##
       ## hyper-mean for maturity
##
       pi_eta ~ ddirch(theta);
##
       ## hyper-prec for maturity
       pi_tau ~ dnorm(0, 0.01) T(0,);
##
##
       for(t in 1:(n_yrs-age_min+n_fore)) { pi_vec[t,1:A] ~ ddirch(pi_eta*pi_tau) }
##
##
       ## estimated harvest rate
##
       for(t in 1:(n_yrs+n_fore)) { h_rate[t] ~ dunif(0,1) }
##
##
       ## LIKELIHOOD
       ##-----
##
##
       ## predicted recruits in BY t
##
       covar[1] <- inprod(gamma,mod_cvrs[1,]);</pre>
##
       ln_BH_a[1] <- mu_BH_a + covar[1];</pre>
##
       E_{n_{ec}[1]} < n_{BH_a[1]} + n_{Sp[1]} - \log(1 + \beta_{Sp[1]}) + phi*innov_1;
##
       tot_ln_Rec[1] ~ dnorm(E_ln_Rec[1],tau_r);
##
       res_ln_Rec[1] <- tot_ln_Rec[1] - E_ln_Rec[1];
       w[1] <- phi * innov_1 + res_ln_Rec[1];
##
##
##
       ## median of total recruits
##
       tot_Rec[1] <- exp(tot_ln_Rec[1]);</pre>
##
##
       ## R/S
##
       ln_RS[1] <- tot_ln_Rec[1] - ln_Sp[1];</pre>
##
       ## brood-yr recruits by age
##
##
       for(a in 1:A) {
##
       Rec[1,a] <- tot_Rec[1] * pi_vec[1,a];</pre>
##
##
       ## brood years 2:(n_yrs-age_min)
##
##
       for(t in 2:(n_yrs-age_min+n_fore)) {
       ## predicted recruits in BY t
##
##
       covar[t] <- inprod(gamma, mod_cvrs[t,]);</pre>
##
       ln_BH_a[t] <- mu_BH_a + covar[t];</pre>
##
       E_{n_{ec}[t]} \leftarrow n_BH_a[t] + n_Sp[t] - \log(1 + beta*Sp[t]) + phi*res_ln_Rec[t-1];
##
       tot_ln_Rec[t] ~ dnorm(E_ln_Rec[t],tau_r);
       res_ln_Rec[t] <- tot_ln_Rec[t] - E_ln_Rec[t];</pre>
##
##
       w[t] <- phi * res_ln_Rec[t-1] + res_ln_Rec[t];</pre>
##
       ## median of total recruits
##
##
       tot_Rec[t] <- exp(tot_ln_Rec[t]);</pre>
##
##
       ## R/S
       ln_RS[t] <- tot_ln_Rec[t] - ln_Sp[t];</pre>
##
##
##
       ## brood-yr recruits by age
       for(a in 1:A) {
##
##
       Rec[t,a] <- tot_Rec[t] * pi_vec[t,a];</pre>
##
##
       } ## end t loop over year
##
```

```
##
       ## get predicted calendar year returns by age
       ## matrix Run has dim [(n_yrs-age_min) x A]
##
       ## step 1: incomplete early broods
##
##
       ## first cal yr of this grp is first brood yr + age_min + age_skip
##
##
       for(i in 1:(age_max-age_min-age_skip)) {
##
       ## projected recruits
       for(a in 1:(i+age_skip)) {
##
##
       Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
       }
##
##
       ## imputed recruits
       for(a in (i+1+age_skip):A) {
##
       lnRec[i,a] ~ dnorm(Rec_mu,Rec_tau);
##
##
       Run[i,a] <- exp(lnRec[i,a]);</pre>
##
       }
##
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
##
       ## predicted age-prop vec for multinom
##
       for(a in 1:A) {
##
       age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
       }
##
       ## multinomial for age comp
##
##
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
       lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
##
##
##
       ## step 2: info from complete broods
##
       ## first cal yr of this grp is first brood yr + age_max
       for(i in (A-age_skip):(n_yrs-age_min-age_skip+n_fore)) {
##
##
       for(a in 1:A) {
       Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
##
##
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
##
       ## predicted age-prop vec for multinom
       for(a in 1:A) {
##
       age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
##
##
       ## multinomial for age comp
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
       lp_age[i] <- ifelse(i < n_yrs-age_min-age_skip+n_fore, logdensity.multi(dat_age[i,1:A],age_v[i,1</pre>
##
##
       }
##
##
       ## get predicted calendar year spawners
##
       ## first cal yr is first brood yr
       for(t in 1:(n_yrs+n_fore)) {
##
       ## obs model for spawners
##
```

```
#Sp[t] <- max(10,tot_Run[t] - dat_harv[t]);
##
##
       est_harv[t] = ifelse(t > n_yrs,1,h_rate[t] * tot_Run[t]);
       dat_harv[t] ~ dlnorm(log(est_harv[t]), 20);
##
       Sp[t] = tot_Run[t] - est_harv[t];
##
##
       ln_Sp[t] <- log(Sp[t]);</pre>
##
       ln_dat_esc[t] ~ dnorm(ln_Sp[t], tau_s);
##
##
       lp_esc[t] <- ifelse(t < n_yrs + 1,logdensity.norm(ln_dat_esc[t],ln_Sp[t], tau_s),0);</pre>
##
##
     } ## end model description
##
## Warning in readLines(con, n = 1): incomplete final line found on 'C:/Users/buehrtwb/OneDrive -
## Washington State Executive Branch Agencies/Documents/Scripts/Skagit-River-Steelhead-Forecast/jags/
## IPM_BH_cov_AR.txt'
```

Beverton-Holt with covars and AR1MA1 recruitment residuals

```
processFile(file.path(jagsdir, "IPM_BH_cov_MA1_AR1.txt"))
```

```
##
## model {
##
##
     ##----
##
     ## PRIORS
     ##----
##
##
     ## alpha = intrinsic productivity
##
     alpha ~ dnorm(0,0.001) T(0,);
     mu BH a <- log(alpha);</pre>
     E_BH_a <- mu_BH_a + sigma_r/(2 - 2*phi^2);</pre>
##
##
##
     ## strength of dens depend
##
     beta_inv ~ dnorm(0, 1e-9) T(0,);
##
     beta <- 1/beta_inv;</pre>
##
##
     ## covariate effects
##
     for(i in 1:n_cov) { gamma[i] ~ dnorm(0,0.01) }
##
##
     ## AR(1) coef for recruitment residual
     #phi ~ dunif(-0.999,0.999);
##
##
     #phi <- 0;</pre>
##
     phi_prior ~ dbeta(2,2);
##
     phi <- phi_prior*2-1;</pre>
##
     #phi ~ dunif(0,0.999);
##
     ## MA(1) coef recruitment residual
##
     theta_res_prior ~ dbeta(2,2);
##
##
     theta_res <- theta_res_prior*2-1;</pre>
##
     #theta_res ~ dunif(0,0.999);
##
##
     ## innovation in first year
     #innov_1 ~ dnorm(0,tau_r*(1-phi*phi));#AR1
##
##
     innov_1 ~ dnorm(0,(1-phi^2)/((1+2*phi*theta_res+theta_res^2)*sigma_r^2));#AR1MA1
##
```

```
##
     ## process variance for recruits model
##
     sigma_r ~ dnorm(0, 2e-2) T(0,);
##
     tau_r <- 1/sigma_r;</pre>
##
##
     ## obs variance for spawners
##
     tau s <- 1/sigma s;
##
     sigma s \sim dnorm(0, 0.001) T(0,);
##
##
     ## unprojectable early recruits;
##
     ## hyper mean across all popns
##
     Rec_mu ~ dnorm(0,0.001);
     ## hyper SD across all popns
##
     Rec_sig ~ dunif(0,100);
##
##
     ## precision across all popns
##
    Rec_tau <- pow(Rec_sig,-2);</pre>
##
    ## multipliers for unobservable total runs
## ttl_run_mu ~ dunif(1,5);
## ttl_run_tau ~ dunif(1,20);
##
##
     ## get total cal yr returns for first age_min yrs
##
     for(i in 1:(age_min+age_skip)) {
##
        ln_tot_Run[i] ~ dnorm(ttl_run_mu*Rec_mu,Rec_tau/ttl_run_tau);
##
        tot_Run[i] <- exp(ln_tot_Run[i]);</pre>
##
##
##
     ## maturity schedule
##
     ## unif vec for Dirch prior
     theta <-c(1,10,10,5,1,1)
##
     ## hyper-mean for maturity
##
     pi_eta ~ ddirch(theta);
##
##
     ## hyper-prec for maturity
##
     pi_tau ~ dnorm(0, 0.01) T(0,);
     for(t in 1:(n_yrs-age_min+n_fore)) { pi_vec[t,1:A] ~ ddirch(pi_eta*pi_tau) }
##
##
##
     ## estimated harvest rate
##
     for(t in 1:(n_yrs+n_fore)) { h_rate[t] ~ dunif(0,1) }
##
##
     ##----
##
     ## LIKELIHOOD
     ##-----
##
##
     ## predicted recruits in BY t
##
     covar[1] <- inprod(gamma,mod_cvrs[1,]);</pre>
     ln_BH_a[1] <- mu_BH_a + covar[1];</pre>
##
##
     E_{n_{ec}[1]} \leftarrow n_{BH_a[1]} + n_{Sp[1]} - \log(1 + beta*Sp[1]) + phi * innov_1 + theta_res * 0;
     tot_ln_Rec[1] ~ dnorm(E_ln_Rec[1], tau_r);
##
     res_ln_Rec[1] <- tot_ln_Rec[1] - E_ln_Rec[1];</pre>
     w[1] <- phi * innov_1 + theta_res * 0 + res_ln_Rec[1]</pre>
##
##
##
     ## median of total recruits
     tot_Rec[1] <- exp(tot_ln_Rec[1]);</pre>
##
##
##
     ## R/S
##
     ln_RS[1] <- tot_ln_Rec[1] - ln_Sp[1];</pre>
##
```

```
##
    ## brood-yr recruits by age
##
    for(a in 1:A) {
##
      Rec[1,a] <- tot_Rec[1] * pi_vec[1,a];</pre>
##
##
##
    ## brood years 2:(n_yrs-age_min)
    for(t in 2:(n_yrs-age_min+n_fore)) {
##
      ## predicted recruits in BY t
##
##
      covar[t] <- inprod(gamma, mod_cvrs[t,]);</pre>
##
      ln_BH_a[t] <- mu_BH_a + covar[t];</pre>
##
##
##
      #version 4; more similar to AR1 original model
##
      ##
      ##
      tot_ln_Rec[t] ~ dnorm(E_ln_Rec[t], tau_r);
##
      res_ln_Rec[t] <- tot_ln_Rec[t] - E_ln_Rec[t];</pre>
      w[t] <- phi * w[t-1] + theta_res * res_ln_Rec[t-1] + res_ln_Rec[t];
##
##
##
##
      ## median of total recruits
##
      tot_Rec[t] <- exp(tot_ln_Rec[t]);</pre>
##
      ## R/S
##
      ln_RS[t] <- tot_ln_Rec[t] - ln_Sp[t];</pre>
##
      ## brood-yr recruits by age
##
      for(a in 1:A) {
##
        Rec[t,a] <- tot_Rec[t] * pi_vec[t,a];</pre>
##
##
    } ## end t loop over year
##
##
    ## get predicted calendar year returns by age
##
    ## matrix Run has dim [(n_yrs-age_min) x A]
    ## step 1: incomplete early broods
##
##
    ## first cal yr of this grp is first brood yr + age_min + age_skip
    for(i in 1:(age_max-age_min-age_skip)) {
##
##
      ## projected recruits
##
      for(a in 1:(i+age_skip)) {
##
        Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
##
      ## imputed recruits
##
      for(a in (i+1+age_skip):A) {
##
        lnRec[i,a] ~ dnorm(Rec_mu,Rec_tau);
        Run[i,a] <- exp(lnRec[i,a]);</pre>
##
##
##
      ## total run size
      tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
      ## predicted age-prop vec for multinom
##
##
      for(a in 1:A) {
##
        age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
      ## multinomial for age comp
##
      dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
      lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
    }
##
```

```
##
##
     ## step 2: info from complete broods
##
     ## first cal yr of this grp is first brood yr + age max
     for(i in (A-age_skip):(n_yrs-age_min-age_skip+n_fore)) {
##
##
       for(a in 1:A) {
##
         Run[i,a] <- Rec[(age skip+i)-a+1,a];</pre>
##
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
       ## predicted age-prop vec for multinom
##
##
       for(a in 1:A) {
##
         age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
       ## multinomial for age comp
##
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
       #lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
##
       lp_age[i] <- ifelse(i < n_yrs-age_min-age_skip+n_fore,</pre>
##
       logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]),0)
##
     }
##
##
     ## get predicted calendar year spawners
##
     ## first cal yr is first brood yr
     for(t in 1:(n_yrs+n_fore)) {
##
       ## obs model for spawners
##
       # Sp[t] <- max(10,tot_Run[t] - dat_harv[t]);
##
##
       est_harv[t] = ifelse(t > n_yrs,1,h_rate[t] * tot_Run[t]);
##
       dat_harv[t] ~ dlnorm(log(est_harv[t]), 20);
       Sp[t] = tot_Run[t] - est_harv[t];
##
##
       ln_{Sp}[t] \leftarrow log(Sp[t]);
##
       ln_dat_esc[t] ~ dnorm(ln_Sp[t], tau_s);
##
       lp_esc[t] <- ifelse(t < n_yrs + 1,logdensity.norm(ln_dat_esc[t],ln_Sp[t], tau_s),0);</pre>
##
##
## } ## end model description
## Warning in readLines(con, n = 1): incomplete final line found on 'C:/Users/buehrtwb/OneDrive -
## Washington State Executive Branch Agencies/Documents/Scripts/Skagit-River-Steelhead-Forecast/jags/
## IPM_BH_cov_MA1_AR1.txt'
```

Beverton-Holt with covars and AR1 recruitment residuals

##

mu_BH_a <- log(alpha);</pre>

 $E_BH_a \leftarrow mu_BH_a + sigma_r/(2 - 2*phi^2);$

```
processFile(file.path(jagsdir, "IPM_BH_cov_AR_resid.txt"))

##
## model {
##
## ##-----
## ## PRIORS
## ##-----
## ## alpha = intrinsic productivity
## alpha ~ dnorm(0,0.001) T(0,);
```

```
##
##
     ## strength of dens depend
     beta_inv ~ dnorm(0, 1e-9) T(0,);
##
     beta <- 1/beta_inv;</pre>
##
##
##
     ## covariate effects
##
     for(i in 1:n_cov) { gamma[i] ~ dnorm(0,0.01) }
##
##
     ## AR(1) coef for recruitment residual
##
     #phi ~ dunif(-0.999,0.999);
##
     #phi <- 0;</pre>
##
     phi_prior ~ dbeta(2,2);
##
     phi <- phi_prior*2-1;</pre>
##
     #phi ~ dunif(0,0.999);
##
##
     ## innovation in first year
##
     innov_1 ~ dnorm(0,tau_r*(1-phi*phi));#AR1
##
##
     ## process variance for recruits model
##
     sigma_r ~ dnorm(0, 2e-2) T(0,);
##
     tau_r <- 1/sigma_r;</pre>
##
##
     ## obs variance for spawners
##
     tau_s <- 1/sigma_s;</pre>
##
     sigma_s ~ dnorm(0, 0.001) T(0,);
##
##
     ## unprojectable early recruits;
     ## hyper mean across all popns
##
     Rec_mu ~ dnorm(0,0.001);
##
     ## hyper SD across all popns
##
##
     Rec_sig ~ dunif(0,100);
##
     ## precision across all popns
##
     Rec_tau <- pow(Rec_sig,-2);</pre>
##
     ## multipliers for unobservable total runs
## ttl_run_mu ~ dunif(1,5);
## ttl_run_tau ~ dunif(1,20);
##
##
     ## get total cal yr returns for first age_min yrs
##
     for(i in 1:(age_min+age_skip)) {
        ln_tot_Run[i] ~ dnorm(ttl_run_mu*Rec_mu,Rec_tau/ttl_run_tau);
##
##
        tot_Run[i] <- exp(ln_tot_Run[i]);</pre>
##
     }
##
##
     ## maturity schedule
     ## unif vec for Dirch prior
##
     theta \leftarrow c(1,10,10,5,1,1)
##
##
     ## hyper-mean for maturity
##
     pi_eta ~ ddirch(theta);
##
     ## hyper-prec for maturity
##
     pi_tau ~ dnorm(0, 0.01) T(0,);
##
     for(t in 1:(n_yrs-age_min+n_fore)) { pi_vec[t,1:A] ~ ddirch(pi_eta*pi_tau) }
##
##
     ## estimated harvest rate
     for(t in 1:(n_yrs+n_fore)) { h_rate[t] ~ dunif(0,1) }
##
```

```
##
##
     ##----
     ## LIKELIHOOD
##
     ##-----
##
##
     ## predicted recruits in BY t
     covar[1] <- inprod(gamma, mod cvrs[1,]);</pre>
##
     ln_BH_a[1] <- mu_BH_a + covar[1];</pre>
##
     E_{\ln Rec[1]} < \ln_BH_a[1] + \ln_Sp[1] - \log(1 + beta*Sp[1]) + phi * innov_1;
##
##
     tot_ln_Rec[1] ~ dnorm(E_ln_Rec[1], tau_r);
##
     res_ln_Rec[1] <- tot_ln_Rec[1] - E_ln_Rec[1];</pre>
##
     w[1] <- phi * innov_1 + res_ln_Rec[1];
##
     ## median of total recruits
##
##
     tot_Rec[1] <- exp(tot_ln_Rec[1]);</pre>
##
##
     ## R/S
##
     ln_RS[1] <- tot_ln_Rec[1] - ln_Sp[1];</pre>
##
##
     ## brood-yr recruits by age
##
     for(a in 1:A) {
##
       Rec[1,a] <- tot_Rec[1] * pi_vec[1,a];</pre>
##
##
##
     ## brood years 2:(n_yrs-age_min)
##
     for(t in 2:(n_yrs-age_min+n_fore)) {
##
       ## predicted recruits in BY t
##
       covar[t] <- inprod(gamma, mod_cvrs[t,]);</pre>
       ln_BH_a[t] <- mu_BH_a + covar[t];</pre>
##
##
       E_{n_{ec}[t]} < ln_{BH_a[t]} + ln_{Sp[t]} - log(1 + beta*Sp[t]) + phi * w[t-1];
##
       tot_ln_Rec[t] ~ dnorm(E_ln_Rec[t], tau_r);
##
       res_ln_Rec[t] <- tot_ln_Rec[t] - E_ln_Rec[t];</pre>
##
       w[t] <- phi * w[t-1] + res_ln_Rec[t];
##
##
       ## median of total recruits
##
       tot_Rec[t] <- exp(tot_ln_Rec[t]);</pre>
##
       ## R/S
##
       ln_RS[t] \leftarrow tot_ln_Rec[t] - ln_Sp[t];
##
       ## brood-yr recruits by age
       for(a in 1:A) {
##
##
         Rec[t,a] <- tot_Rec[t] * pi_vec[t,a];</pre>
##
##
     } ## end t loop over year
##
##
     ## get predicted calendar year returns by age
     ## matrix Run has dim [(n_yrs-age_min) x A]
##
     ## step 1: incomplete early broods
##
     ## first cal yr of this grp is first brood yr + age_min + age_skip
##
##
     for(i in 1:(age_max-age_min-age_skip)) {
##
       ## projected recruits
##
       for(a in 1:(i+age_skip)) {
##
         Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
##
       ## imputed recruits
##
       for(a in (i+1+age_skip):A) {
```

```
##
         lnRec[i,a] ~ dnorm(Rec_mu,Rec_tau);
##
         Run[i,a] <- exp(lnRec[i,a]);</pre>
##
       }
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
##
       ## predicted age-prop vec for multinom
##
       for(a in 1:A) {
##
         age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
       ## multinomial for age comp
##
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
       lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
##
##
##
     ## step 2: info from complete broods
##
     ## first cal yr of this grp is first brood yr + age_max
##
     for(i in (A-age_skip):(n_yrs-age_min-age_skip+n_fore)) {
##
       for(a in 1:A) {
##
         Run[i,a] <- Rec[(age_skip+i)-a+1,a];</pre>
##
##
       ## total run size
##
       tot_Run[i+age_min+age_skip] <- sum(Run[i,1:A]);</pre>
       ## predicted age-prop vec for multinom
##
       for(a in 1:A) {
##
##
         age_v[i,a] <- Run[i,a] / tot_Run[i+age_min];</pre>
##
##
       ## multinomial for age comp
       dat_age[i,1:A] ~ dmulti(age_v[i,1:A],dat_age[i,A+1]);
##
       #lp_age[i] <- logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]);</pre>
##
       lp_age[i] <- ifelse(i < n_yrs-age_min-age_skip+n_fore,</pre>
##
##
       logdensity.multi(dat_age[i,1:A],age_v[i,1:A],dat_age[i,A+1]),0)
##
     }
##
##
     ## get predicted calendar year spawners
##
     ## first cal vr is first brood vr
##
     for(t in 1:(n_yrs+n_fore)) {
##
       ## obs model for spawners
##
       # Sp[t] <- max(10,tot_Run[t] - dat_harv[t]);
##
       est_harv[t] = ifelse(t > n_yrs,1,h_rate[t] * tot_Run[t]);
       dat_harv[t] ~ dlnorm(log(est_harv[t]), 20);
##
       Sp[t] = tot Run[t] - est harv[t];
##
##
       ln_Sp[t] \leftarrow log(Sp[t]);
##
       ln_dat_esc[t] ~ dnorm(ln_Sp[t], tau_s);
##
       lp_esc[t] <- ifelse(t < n_yrs + 1,logdensity.norm(ln_dat_esc[t],ln_Sp[t], tau_s),0);</pre>
##
##
## } ## end model description
## Warning in readLines(con, n = 1): incomplete final line found on 'C:/Users/buehrtwb/OneDrive -
## Washington State Executive Branch Agencies/Documents/Scripts/Skagit-River-Steelhead-Forecast/jags/
## IPM_BH_cov_AR_resid.txt'
```

Fitting the models and generating the one year ahead forecasts

Before fitting the model in JAGS, we need to specify the MCMC control parameters.

```
## 1. MCMC control params
mcmc_ctrl <- list(</pre>
  chains = 4,
  length = 200000, #5e5,
  burn = 100000, #2e5,
  thin = 100 #400
## total number of MCMC samples after burnin
mcmc_samp <- mcmc_ctrl$length*mcmc_ctrl$chains/mcmc_ctrl$thin</pre>
## fit or load models
models=c("IPM_BH_cov_MA1_AR1",
             "IPM_BH_cov_AR",
             "IPM_BH_cov_AR_resid",
             "IPM_BH_cov_MA1_AR1_age",
             "IPM_BH_cov_AR_age",
             "IPM_BH_cov_AR_resid_age"
n mods<-length(models)</pre>
mod_fits <- fit_load_mods(models=models)</pre>
```

Model selection

```
tot_mods <- n_forecasts*n_mods
# get escapement data
dat_esc_forecast <- dat_esc[which(dat_esc$year %in% seq(yr_begin,yr_end,1)),]</pre>
## get harvest data
dat_harv_forecast <- dat_harv[which(dat_harv$year %in% seq(yr_begin,yr_end,1)),]</pre>
## observed terminal run size
obs_trs <- dat_esc_forecast$escapement + dat_harv_forecast$catch</pre>
pred_trs <- NULL</pre>
for(n in 1:n_mods){
  #n <- 1
  pred_esc <- NULL</pre>
  for(i in 1:(n_forecasts)){
    #i <- 1
    mod res<-NULL
    mod_res<-as.matrix(readRDS(file.path(savedir,paste0(models[n],"_y",i,".rds"))))</pre>
    p_dat <- mod_res[,grep("Sp", colnames(mod_res))]</pre>
    p_dat <- round(median(p_dat[,dim(p_dat)[2]]))</pre>
    pred_esc[i] <- p_dat</pre>
```

```
pred_trs_mod <- pred_esc + 1#+ dat_harv_forecast$catch #you don't need to add catch in because it is
  pred_trs <- cbind(pred_trs,pred_trs_mod)</pre>
  #names(pred trs) <- paste(models[n], " ", "pred trs", sep = "")</pre>
colnames(pred_trs) <- models</pre>
## compute model performance statistics
Error <- pred_trs - obs_trs</pre>
SE <- Error<sup>2</sup>
PE <- Error/obs_trs
APE <- abs(PE)
LAR <- log(obs_trs/pred_trs)
RMSE <- apply(SE,2,function(x){sqrt(mean(x))})</pre>
MPE <- apply(PE,2,function(x){mean(x)})</pre>
MAPE <- apply(APE,2,function(x){mean(x)})</pre>
MSA \leftarrow apply(LAR, 2, function(x) \{100*(exp(mean(abs(x))-1))\})
model_selection <- data.frame(RMSE,MPE,MAPE,MSA)</pre>
weights<-apply(model_selection[,!colnames(model_selection)=="MPE"], 2,function(x) (1/x)/sum(1/x))</pre>
colnames(weights)<-paste0(colnames(weights), "_weight")</pre>
model_selection<-data.frame(model_selection,weights)</pre>
```

Model Selection Via Approximate Leave-Future-Out Cross Validation following the methods here: link.

```
N=yr_last-yr_frst+1
L=N-n_forecasts
thres=0.1
esc_only="No"
userefits="Yes"
mod_fits<-loadmodfits(modelnames=models)</pre>
LF01<-approx_LF0(N=N,L=L,m=1,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[1]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot ks(LF01$ks,N=N,L=L,thres=thres)
# plot_ks(LF01$k_L00IC, N=N, L=L, thres=thres)
LF02<-approx_LF0(N=N,L=L,m=2,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[2]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
```

```
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot_ks(LF02$ks, N=N,L=L,thres=thres)
# plot_ks(LF02$k_L00IC, N=N,L=L,thres=thres)
LF03<-approx_LF0(N=N,L=L,m=3,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[3]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot_ks(LF03$ks, N=N,L=L,thres=thres)
# plot ks(LF03$k LOOIC, N=N,L=L,thres=thres)
LF04<-approx_LF0(N=N,L=L,m=4,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[4]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot_ks(LF03$ks, N=N,L=L,thres=thres)
# plot_ks(LF03$k_L00IC, N=N,L=L,thres=thres)
LF05<-approx_LF0(N=N,L=L,m=5,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[5]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot_ks(LFO3$ks, N=N,L=L,thres=thres)
# plot ks(LF03$k LOOIC, N=N,L=L,thres=thres)
LF06<-approx_LF0(N=N,L=L,m=6,esc_only=esc_only,mod_fits=mod_fits,userefits=userefits,refitname=models[6]
## Warning: Relative effective sample sizes ('r_eff' argument) not specified.
## For models fit with MCMC, the reported PSIS effective sample sizes and
## MCSE estimates will be over-optimistic.
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
# plot ks(LFO3$ks, N=N,L=L,thres=thres)
# plot_ks(LF03$k_L00IC, N=N,L=L,thres=thres)
ELPD < -c(sum(LF01\$LF0,na.rm=T),
        sum(LF02$LF0,na.rm=T),
        sum(LF03$LF0,na.rm=T),
        sum(LF04$LF0,na.rm=T),
        sum(LF05$LF0,na.rm=T),
        sum(LF06$LF0,na.rm=T)
```

Calculating Stacking Weights (finding model averaging weights based on one-step-ahead performance of weighted average models). This is an experimental calculation of stacking weights based on linear comminations of model forecasts to optimize one step ahead RMSE, MSA, or MAPE...it is not ready for prime time yet. An alternative is to figure out how to use methods similar to "stacking weights" here for LFOIC: link

Model Selection Via loo() and compare() with full table of results. Note that elpd_diff will be negative (positive) if the expected predictive accuracy for the first (second) model is higher.

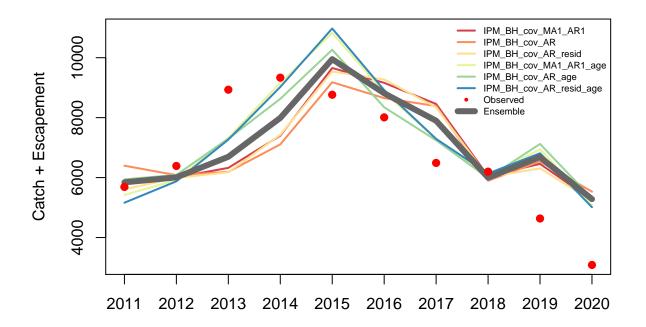
```
LOOIC <- vector("list", n_mods)
## extract log densities from JAGS objects
for(i in 1:n_mods) {
     #i <- 1
     ## convert mcmc.list to matrix
     tmp_lp <- as.matrix(readRDS(file.path(savedir,paste0(models[i],"_y",11,".rds"))))</pre>
     ## extract pointwise likelihoods
     tmp_lp <- tmp_lp[,grepl("lp_", colnames(tmp_lp))]</pre>
      ## if numerical underflows, convert -Inf to 5% less than min(likelihood)
     if(any(is.infinite(tmp_lp))) {
           tmp_lp[is.infinite(tmp_lp)] <- NA</pre>
           tmp_min <- min(tmp_lp, na.rm = TRUE)</pre>
           tmp_lp[is.na(tmp_lp)] <- tmp_min * 1.05</pre>
     ## calculate LOOIC
     LOOIC[[i]] <- loo(tmp_lp)
}
## compute pseudo weights
\#model\_weights < -loo\_model\_weights(LOOIC, method = "pseudobma", optim\_method = "BFGS", optim\_control = (loo_model_weights) | loo_model_weights(looIC, method = "pseudobma", optim_method = (loo_model_weights) | loo_model_weights(looIC, method = "pseudobma", optim_method = (loo_model_weights) | loo_model_weights(looIC, method = "pseudobma", optim_method = (loo_model_weights) | loo_model_weights(looIC, method = (loo_model_weights) | loo_model_weights(loom_weights) | loo_model_weights(loom_weights) | loo_model_weights(loom_weights) | loo_model_weights(loom_weights) | loo_model_weights(loom_weights) | loo_model_weights(loom_weights(loom_weights) | loo_model_weights(loom_weights(loom_weights)) | loo_model_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loom_weights(loo
model_weights <- loo_model_weights(LOOIC, method = "stacking", optim_method = "BFGS", optim_control = li
## LOOIC for all data
tbl_LOOIC <- round(loo_compare(x = LOOIC), 2)
rownames(tbl_L00IC) <- sub("model", "", rownames(tbl_L00IC))</pre>
tbl_LOOIC <- tbl_LOOIC[order(as.numeric(rownames(tbl_LOOIC))), ]</pre>
tbl_LOOIC <- cbind(model = models,</pre>
                                                      as.data.frame(tbl_LOOIC),LOOIC_weight = as.matrix(model_weights))
```

```
tbl_L00IC$delta_L00IC<-tbl_L00IC$looic-min(tbl_L00IC$looic)
model_selection$L00IC<-tbl_L00IC$looic
model_selection$delta_L00IC<-tbl_L00IC$delta_L00IC
model_selection$L00IC_weight<-round(tbl_L00IC$L00IC_weight,4)</pre>
```

Model Averaging and 2020 forecast

```
## extract median 2020 forecast from each model
f_dat<-data.frame(</pre>
  sort(unlist(mod fits[[1]][,paste0("Sp","[",n yrs+n fore,"]")])),
  sort(unlist(mod_fits[[2]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[3]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[4]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[5]][,paste0("Sp","[",n_yrs+n_fore,"]")])),
  sort(unlist(mod_fits[[6]][,paste0("Sp","[",n_yrs+n_fore,"]")]))
colnames(f_dat)<-models</pre>
model_selection[,"2020_forecast"] <- apply(f_dat,2,median)</pre>
weighted_forecast_dist <-(</pre>
  as.matrix(f_dat) %*% (as.vector(model_selection[,"LFOIC_weight"]))
)
ensemble_forecast_posterior<-data.frame(weighted_forecast_dist)</pre>
colnames(ensemble_forecast_posterior)<-c("ensemble_forecast_posterior")</pre>
write.csv(ensemble_forecast_posterior,file.path(savedir,"ensemble_forecast_posterior.csv"),row.names = :
weighted forecast quantiles <- quantile (weighted forecast dist, c(0.025, 0.25, 0.50, 0.75, 0.975))
weighted_forecast<-weighted_forecast_quantiles[3]</pre>
print(model selection)
##
                                                                MSA RMSE_weight MAPE_weight MSA_weight
                                RMSF.
                                            MPF.
                                                     MAPE
                            1573.605 0.10955940 0.2261345 45.18838
## IPM_BH_cov_MA1_AR1
                                                                      0.1537808
                                                                                   0.1558265 0.1639863
## IPM_BH_cov_AR
                            1641.418 0.11051231 0.2389946 45.67331
                                                                      0.1474275
                                                                                   0.1474416 0.1622451
## IPM_BH_cov_AR_resid
                            1563.364 0.09683464 0.2205361 45.03837
                                                                      0.1547881
                                                                                   0.1597823 0.1645324
## IPM_BH_cov_MA1_AR1_age
                           1334.804 0.12616270 0.1951553 43.59182
                                                                      0.1812926
                                                                                  0.1805626 0.1699923
## IPM_BH_cov_AR_age
                            1319.431 0.12904289 0.1969258 43.55082
                                                                      0.1834050
                                                                                   0.1789392 0.1701523
## IPM_BH_cov_AR_resid_age 1349.594 0.11806776 0.1985809 43.82403
                                                                      0.1793059
                                                                                   0.1774479 0.1690916
##
                                ELPD
                                         LFOIC delta_LFOIC LFOIC_weight stacking_weights LOOIC
## IPM_BH_cov_MA1_AR1
                            7.062682 -14.12536
                                                  2.086366
                                                               0.1126168
                                                                                    0.0000 826.65
                            8.105865 -16.21173
## IPM_BH_cov_AR
                                                  0.000000
                                                               0.3196331
                                                                                    0.0000 803.97
## IPM_BH_cov_AR_resid
                            7.164528 -14.32906
                                                  1.882673
                                                               0.1246908
                                                                                    0.0000 782.48
## IPM_BH_cov_MA1_AR1_age 7.274766 -14.54953
                                                  1.662197
                                                               0.1392228
                                                                                    0.4857 826.69
## IPM_BH_cov_AR_age
                            7.309776 -14.61955
                                                  1.592178
                                                               0.1441832
                                                                                    0.5143 836.05
                                                                                   0.0000 825.47
## IPM_BH_cov_AR_resid_age 7.411696 -14.82339
                                                  1.388337
                                                               0.1596534
                            delta LOOIC LOOIC weight 2020 forecast
## IPM_BH_cov_MA1_AR1
                                  44.17
                                              0.0000
                                                           4717.728
## IPM_BH_cov_AR
                                  21.49
                                              0.0953
                                                           5118.656
## IPM_BH_cov_AR_resid
                                   0.00
                                              0.6241
                                                           4555.983
## IPM_BH_cov_MA1_AR1_age
                                  44.21
                                              0.0000
                                                           3380.080
## IPM_BH_cov_AR_age
                                  53.57
                                              0.0000
                                                           3916.261
## IPM_BH_cov_AR_resid_age
                                  42.99
                                              0.2806
                                                           3297.029
```

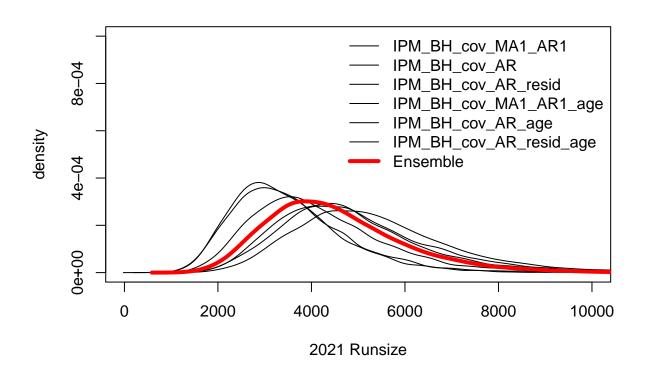
```
print("The model-averaged forecast is:")
## [1] "The model-averaged forecast is:"
print(weighted_forecast_quantiles)
                                   75%
                                           97.5%
##
       2.5%
                 25%
                          50%
## 2283.296 3470.260 4297.101 5322.188 8165.904
ensemble_median<-as.matrix(pred_trs)%*%(as.vector(model_selection[,"LFOIC_weight"]))
cols<-brewer.pal(length(models), "Spectral")</pre>
matplot(as.matrix(data.frame(pred_trs,ensemble_median)),type="1",lty=1,col=c(cols,"grey40"),lwd=c(rep(2
axis(1,1:n_forecasts,(yr_last-n_forecasts+1):(yr_last))
points(x=1:n_forecasts,y=obs_trs,cex=1.5,pch=20,col="red")
legend("topright",legend=c(models,"Observed","Ensemble"),lty=c(rep(1,length(models)),NA),col=c(cols,"re-
```



```
#density plot for final forecasts and ensemble
res<-apply(f_dat,2,function(x) density(x))
plot(x=1,y=1,ylim=c(0,0.001),xlim=c(0,10000),ylab="density",xlab="2021 Runsize")
lapply(res,function(x) lines(x$y~x$x))

## $IPM_BH_cov_MA1_AR1
## NULL
##
## $IPM_BH_cov_AR
## NULL</pre>
```

```
## $IPM_BH_cov_AR_resid
## NULL
##
## $IPM_BH_cov_MA1_AR1_age
## NULL
##
## $IPM_BH_cov_AR_age
## NULL
##
## $IPM_BH_cov_AR_resid_age
## NULL
##
## $IPM_BH_cov_AR_resid_age
## NULL
lines(density(weighted_forecast_dist)$y~density(weighted_forecast_dist)$x,lwd=4,col="red")
legend("topright",legend=c(models,"Ensemble"),lwd=c(rep(1,length(models)),4),col=c(rep("black",length(models)))
```



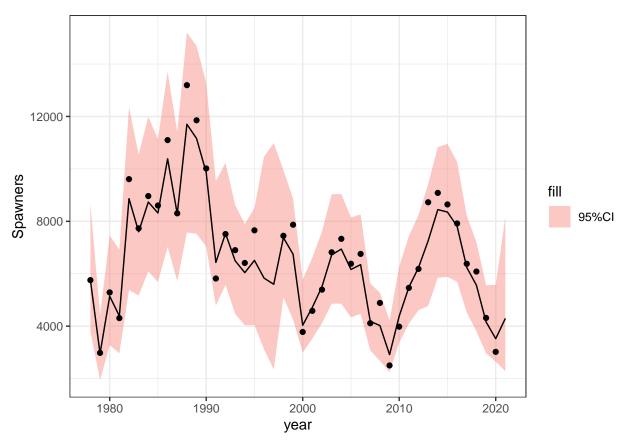
Now lets graph model estimated spawners, recruits, and recruits per spawner:

##

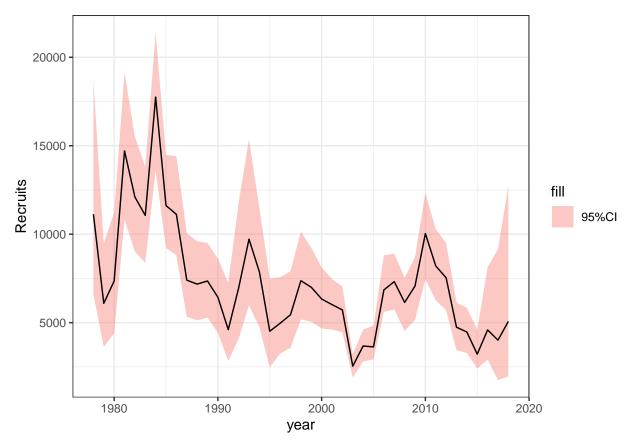
```
RSdat<-function(var,yrs){
    sims<-length(sort(unlist(mod_fits[[1]][,paste0(var,"[",i,"]")])))
    R_mat<-matrix(NA,ncol=length(yrs),nrow=sims)
    for(i in 1:length(yrs)){
        Rdat<-data.frame(
            sort(unlist(mod_fits[[1]][,paste0(var,"[",min(yrs)+i-1,"]")])),
            sort(unlist(mod_fits[[2]][,paste0(var,"[",min(yrs)+i-1,"]")])),
        sort(unlist(mod_fits[[3]][,paste0(var,"[",min(yrs)+i-1,"]")])),
        sort(unlist(mod_fits[[4]][,paste0(var,"[",min(yrs)+i-1,"]")])),</pre>
```

```
sort(unlist(mod_fits[[5]][,paste0(var,"[",min(yrs)+i-1,"]")])),
      sort(unlist(mod_fits[[6]][,paste0(var,"[",min(yrs)+i-1,"]")]))
   )
   dim(Rdat)
   weighted_R <-(</pre>
      as.matrix(Rdat) %*% (as.vector(model_selection[,"LFOIC_weight"]))
   R mat[,i] <-weighted R
    colnames(R_mat)<-yrs</pre>
 return(R_mat)
}
Sdat<-RSdat(var="Sp",yrs=c(1:(n_yrs+n_fore)))</pre>
Rdat<-exp(RSdat(var="tot_ln_Rec",yrs=c(1:(n_yrs+n_fore-3))))</pre>
RSdat<-exp(RSdat(var="ln_RS",yrs=c(1:(n_yrs+n_fore-3))))
library(tidyverse)
Sdat<-as.data.frame(Sdat)%>%
  pivot_longer(cols=everything())%>%
  rename(year=name)%>%
  group_by(year) %>%
  summarise(SpawnerAbundance = quantile(value, c(0.025, 0.5, 0.975)), q = c(0.025, 0.5, 0.975))%%
  mutate(year=as.numeric(year)+yr_frst-1)%>%
  pivot_wider(names_from = q, values_from = SpawnerAbundance)
Rdat<-as.data.frame(Rdat)%>%
  pivot_longer(cols=everything())%>%
  rename(year=name)%>%
  group_by(year) %>%
  summarise(RecruitAbundance = quantile(value, c(0.025, 0.5, 0.975)), q = c(0.025, 0.5, 0.975))%%
  mutate(year=as.numeric(year)+yr_frst-1)%>%
  pivot_wider(names_from = q, values_from = RecruitAbundance)
RSdat<-as.data.frame(RSdat)%>%
  pivot_longer(cols=everything())%>%
  rename(year=name)%>%
  group_by(year) %>%
  summarise(RS = quantile(value, c(0.025, 0.5, 0.975)), q = c(0.025, 0.5, 0.975))%%
  mutate(year=as.numeric(year)+yr_frst-1)%>%
  pivot_wider(names_from = q, values_from = RS)
ggplot(Sdat,aes(x=year,y=`0.5`))+
  geom_ribbon(aes(ymin = `0.025`, ymax = `0.975`,fill="95%CI"),alpha=0.4)+
  geom line()+
  geom_point(dat_esc,mapping=aes(y=escapement,x=year))+
  theme_bw()+
 ylab("Spawners")
```

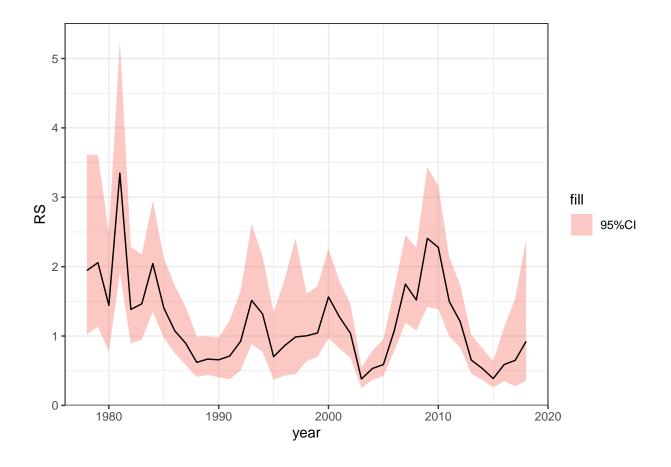
Warning: Removed 2 rows containing missing values (geom_point).



```
ggplot(Rdat,aes(x=year,y=`0.5`))+
  geom_ribbon(aes(ymin = `0.025`, ymax = `0.975`,fill="95%CI"),alpha=0.4)+
  geom_line()+
  theme_bw()+
  ylab("Recruits")
```



```
ggplot(RSdat,aes(x=year,y=`0.5`))+
  geom_ribbon(aes(ymin = `0.025`, ymax = `0.975`,fill="95%CI"),alpha=0.4)+
  geom_line()+
  theme_bw()+
  ylab("RS")
```



Compare sib adjusted forecasts with unadjusted

```
models=c(
             #"IPM_BH_cov_MA1_AR1",
             #"IPM_BH_cov_AR",
             #"IPM_BH_cov_AR_resid",
             "IPM_BH_cov_MA1_AR1_age",
             #"IPM_BH_cov_AR_age",
             "IPM_BH_cov_AR_resid_age"
n_mods<-length(models)</pre>
mod_fits <- fit_load_mods(models=models</pre>
####new version 1.7
tot_mods <- n_forecasts*n_mods
# get escapement data
dat_esc_forecast <- dat_esc[which(dat_esc$year %in% seq(yr_begin,yr_end,1)),]</pre>
## get harvest data
dat_harv_forecast <- dat_harv[which(dat_harv$year %in% seq(yr_begin,yr_end,1)),]</pre>
## observed terminal run size
obs_trs <- dat_esc_forecast$escapement + dat_harv_forecast$catch</pre>
pred_trs <- NULL</pre>
```

```
pred_trs_adj <- NULL</pre>
for(n in 1:n_mods){
  \#n < -1
 pred esc <- NULL
 pred_esc_adj <- NULL</pre>
 c <- 0
 f < -2
  adj<- NULL
  for(i in 1:(n_forecasts)){
    \#i <- 1
    #range of years. Last year in range
    dat_yrs <- seq(yr_frst,(yr_last - n_forecasts + c),1)</pre>
    ## number of years of data
    n_yrs <- length(dat_yrs)</pre>
 mod res<-NULL
  #model refit and one step ahead forecast
  mod_res_pred<-as.matrix(readRDS(file.path(savedir,paste0(models[n],"_y",i,".rds"))))</pre>
  #model refit for year i+1 to extract "observed" state
  mod_res_obs<-as.matrix(readRDS(file.path(savedir,paste0(models[n],"_y",i + 1,".rds"))))</pre>
  p_dat_pred <- mod_res_pred[,grep("Sp", colnames(mod_res_pred))]</pre>
  p_dat_pred <- round(median(p_dat_pred[,n_yrs + 1]))</pre>
  p_dat_obs <- mod_res_obs[,grep("Sp", colnames(mod_res_obs))]</pre>
  p_dat_obs <- round(median(p_dat_obs[,n_yrs + 2]))</pre>
 pred_esc[i] <- p_dat_pred</pre>
  #take same model and implement sibling adjustment
  ## forecast for year t
  p_dat_pred_adj <- cbind(((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",1,"]")])),</pre>
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",2,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",3,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",4,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",5,"]")])),
                  ((mod_res_pred[,paste0("Run","[",n_yrs - age_min + 1,",",6,"]")])))
  p_dat_pred_adj_sum <- apply(p_dat_pred_adj,1,FUN = "sum")</pre>
  median(p_dat_pred_adj_sum)
  ## observation for year t+1
  p_dat_obs_adj <- cbind(((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",1,"]")])),</pre>
```

```
((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",2,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",3,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",4,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",5,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 1,",",6,"]")])))
  ## forecast for year t+1 to be adjusted
  p_dat_pred_t <- cbind(((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",1,"]")])),</pre>
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",2,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",3,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",4,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",5,"]")])),
                  ((mod_res_obs[,paste0("Run","[",n_yrs - age_min + 2,",",6,"]")])))
  \#median(p\_dat\_obs\_adj[,a-1])/median(p\_dat\_pred\_adj[,a-1])
    p_dat_pred_adj_a <- NULL</pre>
    \#adj_a \leftarrow NULL
    for (a in 2:A){
      #a <- 3
      \#adj_a[a] \leftarrow median(as.vector(p_dat_obs_adj[,a-1])/as.vector(p_dat_pred_adj[,a-1]))
      p_dat_pred_adj_a_temp <- as.vector(p_dat_pred_t[,a])*(as.vector(p_dat_obs_adj[,a-1])/as.vector(p_
      p_dat_pred_adj_a <- cbind(p_dat_pred_adj_a,p_dat_pred_adj_a_temp)</pre>
    }
    \#adj \leftarrow rbind(adj, t(adj_a))
    p_dat_pred_adj_a <- apply(p_dat_pred_adj_a,1,FUN = sum)</pre>
    pred_esc_adj[f] <- round(median(p_dat_pred_adj_a))</pre>
    #pred_esc_adj[f] <- median(apply(p_dat_pred_t,1,FUN = "sum"))</pre>
    f \leftarrow f+1
    c <- c+1
  } #next forecast
  pred_trs_mod <- pred_esc + 1#+ dat_harv_forecast$catch #you don't need to add catch in because it is
  pred_trs_mod_adj <- pred_esc_adj #not adjusted because adjusted prediction comes from "Run" rather th
  pred_trs <- cbind(pred_trs,pred_trs_mod)</pre>
  pred_trs_adj <- cbind(pred_trs_adj,pred_trs_mod_adj)</pre>
  #names(pred_trs) <- paste(models[n],"_","pred_trs",sep = "")</pre>
colnames(pred_trs) <- models</pre>
colnames(pred_trs_adj) <- paste(models,"_","sib_adjust",sep = "")</pre>
pred_trs_adj <- pred_trs_adj[-11,]</pre>
pred_trs<-data.frame(pred_trs,pred_trs_adj)</pre>
pred_trs<-pred_trs[-1,]</pre>
```

```
## compute model performance statistics
Error <- pred_trs - obs_trs[-1]</pre>
SE <- Error<sup>2</sup>
PE <- Error/obs_trs[-1]
APE <- abs(PE)
LAR <- log(obs_trs/pred_trs)</pre>
RMSE <- apply(SE,2,function(x){sqrt(mean(x))})</pre>
MPE <- apply(PE,2,function(x){mean(x)})</pre>
MAPE <- apply(APE,2,function(x){mean(x)})
MSA \leftarrow apply(LAR, 2, function(x) \{100*(exp(mean(abs(x))-1))\})
model_selection <- data.frame(RMSE,MPE,MAPE,MSA)</pre>
weights<-apply(model_selection[,!colnames(model_selection)=="MPE"], 2,function(x) (1/x)/sum(1/x))</pre>
colnames(weights)<-paste0(colnames(weights),"_weight")</pre>
model_selection <- data.frame (model_selection, weights)
print(model_selection)
##
                                            RMSE
                                                        MPE
                                                                 MAPE
                                                                            MSA RMSE_weight MAPE_weight
## IPM_BH_cov_MA1_AR1_age
                                        1404.147 0.1454355 0.2115845 39.90132
                                                                                  0.3368552
                                                                                               0.3047165
## IPM_BH_cov_AR_resid_age
                                        1411.751 0.1414614 0.2103704 46.96026
                                                                                  0.3350409
                                                                                               0.3064751
## IPM_BH_cov_MA1_AR1_age_sib_adjust 2967.527 0.3304070 0.3435246 60.49714 0.1593900
                                                                                               0.1876817
## IPM_BH_cov_AR_resid_age_sib_adjust 2803.530 0.3169427 0.3205606 63.56399 0.1687139
                                                                                               0.2011267
                                        MSA_weight
##
## IPM_BH_cov_MA1_AR1_age
                                         0.3187785
## IPM_BH_cov_AR_resid_age
                                         0.2708606
## IPM_BH_cov_MA1_AR1_age_sib_adjust
                                         0.2102526
## IPM_BH_cov_AR_resid_age_sib_adjust 0.2001083
#sib adjust plot
cols<-brewer.pal(length(models)*2, "Spectral")</pre>
matplot(as.matrix(data.frame(pred_trs)),type="l",lty=1,col=c(cols,"grey40"),lwd=c(rep(2,length(models)*
axis(1,1:(n_forecasts-1),(yr_last-n_forecasts+2):(yr_last))
points(x=1:(n_forecasts-1),y=obs_trs[-1],cex=1.5,pch=20,col="red")
legend("topright",legend=c(rownames(model_selection),"Observed"),lty=c(rep(1,length(models)*2),NA),col=
```

