Econometrics CheatSheet

Basic concepts

Definition of econometrics

Econometrics - is a social science discipline with the objective of quantify the relationships between economic agents, contrast economic theories and evalue and implement government and business policies.

Econometric model - is a simplificated representation of the reality to explain economic phenomena.

Data types

1. Cross section: data taken at a given moment in time, an static "photo". Order does not matter.
2. Temporal series: observation of one/many vairable/s across time. Order does matter.
3. Panel data: consist of a temporal serie for each observation of a cross section.
4. Pooled cross sections: combines cross sections from different temporal periods.

Phases of an econometric model

1. Specification
2. Estimation
3. Validation
4. Utilization

Assumptions of the econometric model

Under this assumptions the estimators of the parameters will present "good properties". GAUSS MARKOV ASSUMPTIONS (EXTENDED)

- Parameters linearity.
- The sample of the poblation is random. Caracteristics:
  - Independence: independence, that guarantees that all the covariances between independents are zero.
  - Identical distribution: that guarantees that the n expected values and variances of the observations are the same.
- $E(u/X_1, X_2, ..., X_k) = 0$, guarantees that the estimations are unbiased, that have some implications:
  - $E(u) = 0$ there are none systematic errors.
  - $Cov(u, X_1) = Cov(u, X_2) = ... = Cov(u, X_k) = 0$ there are no relevant variables not included in the model.
  - $E(Y/X_1, X_2, ..., X_k) = \beta_0 + \beta_1 X_1 + \beta_k X_k$ the lineal relation between Y and $X_1, ..., X_k$ is fulfilled, at least in average.
- Homocedasticity: $Var(u_i/X_{1i}, X_{2i}, ..., X_{ki}) = \sigma^2$, the variability of the error is the same for all levels of x. Guarantees that the estimations are efficient. Implies that: $Var(Y_i/X_{1i}, X_{2i}, ..., X_{ki}) = \sigma^2$, the variability of the dependent variable is the same for all levels of x.
- No autocorrelation: $Cov(u_i, u_j) = 0 \rightarrow Cov(Y_i Y_j / X) = 0$ for every i different from j. The errors do not contain information about other errors.
- The distribution of the errors is normal (is not always necessary).
- No multicolineality: none of the independent variables is constant nor exist an exact (or aproximate) linear relation between them, they are linearly independents.
- The number of available data is greater than $k+1$ (β parameters to estimate).

The homocedasticity and no autocorrelation asumptions can also be written in matrix form: $Var(u/X) = \sigma^2 I_n$

Interpretation of the coefficients

Model	Dependent	Independent	Interpretation β_1
Level-level	y	x	$\Delta y = \beta_1 \Delta x$
Level-log	y	$log(x)$	$\Delta y = (\beta_1/100)[1\%\Delta x]$
Log-level	$log(y)$	x	$\%\Delta y = (100\beta_1)\Delta x$
Log-log	$log(y)$	$log(x)$	$\%\Delta y = \beta_1 \%\Delta x$
Quadratic	y	$x + x^2$	$\Delta y = (\beta_1 + 2\beta_2 x)\Delta x$

OLS estimation of the model

Simple regression model

$Y_i = \beta_0 + \beta_1 X_{1i} + u_i, i = 1, ..., n$

Definitions

$\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$

$\hat{u}_i = Y_i - \hat{Y}_i = Y_i - (\hat{\beta}_0 + \hat{\beta}_1 X_i)$

Objective is minimize the square sum of resid:

$Min \sum_{i=1}^n \hat{u}_i^2 = Min \sum_{i=1}^n [Y_i - (\hat{\beta}_0 + \hat{\beta}_1 X_i)]^2$

With

$\hat{\beta}_0 = \overline{Y} - \hat{\beta}_1 \overline{X}$

$\hat{\beta}_1 = \frac{Cov(Y,X)}{Var(X)}$

Multiple regression model

$Y_i = \beta_0 + \beta_1 X_{1i} + ... + \beta_k X_{ki} + u_i, i = 1, ..., n$

$\hat{u}_i = Y_i - \hat{Y}_i = Y_i - (\hat{\beta}_0 + \hat{\beta}_1 X_i + ... + \hat{\beta}_k X_{ki})$

Objective:

$Min \sum_{i=1}^n \hat{u}_i^2$

Then

$\hat{\beta}_0 = \overline{Y} - \hat{\beta}_1 \overline{X}_1 - ... - \hat{\beta}_k \overline{X}_k$

$\hat{\beta}_j = \frac{Cov(Y, resid(X_j))}{Var(resid(X_j))}$

Properties of OLS

- Lineality in Y.
- Normality: $Y/X\ N(\beta_0 + \beta_1 X, \sigma^2)$

- Expected value of the estimator: $E(\hat{\beta}_1/X_i) = \beta_1$, then $\hat{\beta}_1$ is an unbiased estimator of β_1
- Variance of the estimator: $Var(\hat{\beta}_1/X_i) = \frac{\sigma^2}{nVar(X_i)}$

Eficiency of OLS estimators, Gauss-Markov Theorem. In the context of the simple or multiple linear regression model, the OLS estimators of the parameters are those with the lowest variance between the lineal and unbiased estimators

Central Limit Theorem

Under the CLT, $\hat{\beta}_j$ is a consistent estimator of the poblational parameter β_i.

$plim\hat{\beta}_i = \beta_i$

Regression Analysis

Study and predict the mean value of a variable regarding the base of fixed values of other variables. We usually use Ordinary Least Squares (OLS).

Correlation Analysis

The correlation analysis not distinguish between dependent and independent variables. **Simple Correlation** Measure the grade of lineal association between two variables.

Utilization

Interpretation of the model

Heterocedasticity

The residuals u_i of the poblational regression function don't have the same variance σ^2:

$Var(u_i \mid x_i) = \sigma_i^2; i = 1, ..., n$

Consequences

Under the Gauss-Markov Theorem asumptions, OLS estimators are not efficient. The estimations of the variance of the estimators are biased. The hyphotesis contrast and the confidence intervals are not reliable.

Detection

Plots (look for structures in plots with the square residuals) and contrasts: Park test, Goldfield-Quandt, Bartlett, Breush-Pagan, CUSUMQ, Spearman, White. White's null hypothesis:

$H_0 = HOMOCEDASTICITY$

Correction

- When the variance structure is known, use weighted least squares.
- When the variance structure is not known: make asumptions of the possible structure and apply weighted least squares
- Supossing that σ_i^2 is proportional to x_i^2, divide by x_i

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