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# Abstract

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# Methods

## Participants

Participants were drawn from the 2004, 2006, 2008, 2010, 2012, and 2014 waves of the Health and Retirement Study (HRS). The HRS is a nationally representative longitudinal panel study of individuals over the age of 50 and their spouses of any age in the United States of America. Specifically, the RAND HRS data files (RAND Center for the Study of Aging, 2008) were used as many variables have been pre-cleaned and the files are more user-friendly. The HRS is supported by the National Institute on Aging (NIA U01AG009740) and the Social Security Administration. In 2004, the HRS piloted a self-completed psychosocial questionnaire administered to a random sample of respondents (n=3,262). Beginning in 2006, the HRS began enhanced face-to-face interviews on a rotating basis of 50% of the core panel. The respondent psychosocial questionnaire was left with participants to complete and return by mail at the end of this interview, every 2 years, resulting in longitudinal data for the same participants every 4 years. In the present analysis, participants were excluded if they were younger than 65 at their first wave included assessment wave, if they did not have at least one wave of response data for all cognitive and social variables of interest, and if they reported ever having received a diagnosis of “memory-related disease”, Alzheimer’s disease, or dementia.

## Measures

## Analytical Strategy

A series of models were estimated, first to examine social and cognitive variables of interest individually, and then as bivariate models. Based on Bollen and Curran’s (2004, 2006) recommendations, models were estimated in a progressive series i) autoregressive model, ii) latent growth model, iii) the full ALT model, iv) a latent growth model nested within the ALT model, v) the ALT model without a slope, vi) the ALT model with the slope variance constrained to 0, and vii) the ALT model with autoregressive parameters constrained to equality over time. For ALT models, the first measurement point for all processes was included in the model as predetermined. These were run as univariate models for each variable of interest to understand each variable individually. Model fit was estimated with multiple fit indices: the chi-squared likelihood ratio test, the Comparative Fit Index (CFI), the Tucker-Lewis Index (TLI), the Standardized Root Mean Square Residual (SRMR) and the Root Mean Square Error of Approximation (RMSEA). Values greater than .95 indicate good fit for CFI with greater than .90 indicating adequate fit (Bollen, 2989; Hu & Bentler, 1999). For RMSEA values smaller than .08 or .06 are considered acceptable and good, respectively. For SRMR .10 and .08 are considered acceptable and good fit respectively (Bollen, 2989; Hu & Bentler, 1999). Nested models were compared using the chi-square difference test (chi-square difference; Bollen, 1989). The estimation of bivariate models followed a similar strategy with bivariate models estimated for each social-cognitive variable pair. For bivariate models the full ALT model was specified with correlations between the first measurement points of the cognitive and social variable, with autoregressive, cross-lagged parameters, and with time-specific correlations between the two processes that were unrestricted across time. In the bivariate model series restrictions were progressively added to the full ALT: i) fixing the slope variance to zero; (ii) excluding the slope; (iii) excluding the time-specific uniqunesses’ correlations; (iv) constraining the time-specific uniquenesses’ correlations to equality over time; (iv) constraining the autoregressive parameters to equality over time; (v) constraining the cross-lagged parameters to equality over time. The constraints were added first to the cognitive processes and then to the social process one at a time. All models were estimated with the MLR option for maximum likelihood estimation with robust standard errors in Mplus 7.4 (Muthen & Muthen, 2012-2015). Maximum likelihood estimation can handle even large proportions of missing data assuming missing at random, by using all available information from all cases (Muthen, 1998-2004). Once a final model was identified for each bivariate combination, covariates were added to the model. The intercept and slope parameters of each processes (the cognitive and social) and the first measurement point (estimated as predetermined) were regressed on each covariate.

# Results

The descriptive statistics for all included variables are presented in Table 1. Model fit indices are presented for each model series in Tables 2 though X. Each processes was investigating separately as a univariate model and then each cognitive-social combination was estimated. Results for each process are presented followed by bivariate relations. For both univariate and bivariate

## Immediate Word Recall

For immediate word recall, the full ALT model (model 4) evidenced the best model fit (see Table 2). The intercept of immediate word recall score was ( = 5.44). Overall, immediate word recall performance decreased over time ( = -0.08 , *p* = 0) with significant variance in the slope ( = 0.00, *p* = 0). The ability of immediate word recall preformance to predict immediate word recall performance two years later, once overall trajectories were taken into account, was small and not significant ( = 0.01, *p* = 0.29; = 0.01, *p* = 0.28; = -0.02, *p* = 0.03; = -0.03, *p* = 0.09; = -0.03, *p* = 0.08) with the exception of immediate recall performance at time three significantly predicting performance at time four ( = -0.02, *p* = 0.03).

To evaluate the impact of objective and subjective social factors on immediate word recall performance, four separate bivariate model series were estimated. In all four model series parameter estimates indicated that immediate word recall had a significant negative slope with a small but significant variance, consistent with the univariate immediate word recall results (social network: = -0.09, *p* = 0, = 0.00, *p* = 0; social support: = -0.12, *p* = 0, = 0.01, *p* = 0; social contact: = -0.06, *p* = 0 = 0.00, *p* = 0; loneliness: = -0.08, *p* = 0, = 0.00, *p* = 0). However, in the bivariate models, allowing the autoregressive parameters to vary over time did not significantly improve model fit compared to the less complex models with the autoregressive parameters constrained to equality over time (see tables X through Y). When the autoregressive parameters were constrained to equality over time, the ability of immediate recall performance to predict future immediate word recall performance was not significant for three of the four models (it was significant in the immediate word recall - social network model). The autoregressive parameter values were similar in the full bivariate ALTs but reached significance in several cases: in the immediate word recall - social network model immediate recall performance the 5th wave significantly predicted immediate recall performance at the 6th wave ( = -0.05, *p* = 0.04). In the bivariate immediate word recall - social support ALT model ( = -0.03, *p* = 0.05) and bivariate immediate word recall - loneliness immediate word recall performance at time 3 significantly predicted immediate word recall preformance at time 4 ( = -0.03, *p* = 0.04).

## Delayed Word Recall

For delayed word recall, univariate model results indicate that the full ALT model (model 4) is the best model by fit indices (see Table 3). Intercept ( = 4.38, *p* = 0). Overall, delayed word recall performance decreased over time ( = -0.09 , *p* = 0) with significant variance in the slope ( = 0.00, *p* = 0). The ability of delayed word recall performance to predict later delayed word recall performance was not significant for the first to second wave ( = 0.01, *p* = 0.30). Previous delayed word recall performance significantly predicted future performance for the next three waves = 0.02, *p* = 0.04; = -0.02, *p* = 0.05; = -0.03, *p* = 0.04) but delayed word recall performance at time 5 (2012) did not significantly predict performance at time 6 (2014) = -0.04, *p* = 0.06 over and above the overall trajectory of change. A plot of the predicted mean trajectory is shown in Figure 2.

When delayed word recall was examined in relation to the four social factors, the slope of delayed recall performance was consistently significant and negative with a small but significance variance parameter (social network: = -0.08, *p* = 0, = 0.01, *p* = 0; social support: = -0.13, *p* = 0, = 0.01, *p* = 0; social contact: = -0.09, *p* = 0 = 0.01, *p* = 0; loneliness: = -0.10, *p* = 0, = 0.01, *p* = 0). Across all four bivariate models with delayed word recall performance, allowing the autoregressive parameters of delayed word recall to vary over time did not signficantly improve the model fit over the bivariate models with the equality constraint (social network: = -0.01, *p* = 0.45, social support: = -0.01, *p* = 0.15, social contact: = -0.01, *p* = 0.54 loneliness: = -0.01, *p* = 0.37). In two of the four bivariate models, social contact and social network, with delayed word recall, previous delayed word recall performance did not significantly predict performance two years later at any occasion and thus constraining the autogressive parameter did not change the substantive interpretation. In the other two bivariate models, loneliness and social support, delayed word recall at time two significantly predicted delayed recall performance at time three when the autoregressive parameters were allowed to vary over time (social support: = -0.03, *p* = 0.03, loneliness: = -0.02, *p* = 0.04). However the negative significance indicates that better performance in 2006 actually predicted worse performance in 2008.

## Mental Status

For the univariate mental status models, the full ALT model (model 4) shows the best model fit according to fit indicies (see Table 4). The estimate mental status intercept is high ( = 8.52, *p* = 0). With significant variability. Overall, mental status decreased over time ( = -0.25 , *p* = 0) with significant variance in the slope ( = 0.00, *p* = 0). The ability of earlier mental status to predict later mental status, over and above the overall trajectory of change was consistently significant( = 0.06, *p* = 0; = 0.11, *p* = 0; = 0.12, *p* = 0; = 0.19, *p* = 0; = 0.23, *p* = 0). A plot of the predicted mean trajectory of mental status performance is shown in Figure 3.

When mental status was examined in relation to the four social factors, the slope of mental status was consistently significant and negative with a small but significance variance parameter (social network: = -0.27, *p* = 0, = 0.00, *p* = 0; social support: = -0.23, *p* = 0, = 0.00, *p* = 0; social contact: = -0.23, *p* = 0 = 0.00, *p* = 0; loneliness: = -0.26, *p* = 0, = 0.00, *p* = 0). Mental status significantly predicted mental status two years later, over and above the overall trajectory of change, with the autoregressive parameters getting larger over time (social network: = 0.04, *p* = 0, = 0.10, *p* = 0; = 0.12, *p* = 0; = 0.19, *p* = 0; = 0.24, *p* = 0; social support: = 0.01, *p* = 0.19 = 0.06, *p* = 0; = 0.11, *p* = 0; = 0.20, *p* = 0; = 0.25, *p* = 0; social contact: = 0.01, *p* = 0.19, = 0.06, *p* = 0; = 0.11, *p* = 0; = 0.20, *p* = 0; = 0.25, *p* = 0; loneliness: = 0.05, *p* = 0, = 0.10, *p* = 0; = 0.12, *p* = 0; = 0.18, *p* = 0; = 0.23, *p* = 0).

## Loneliness

For the univariate loneliness models, the ALT model with the autoregressive parameters constrained to equality over time was the best fitting most parsimonious model according to fit indices (see Table 5). ( = 1.38, *p* = 0). With significant variability. Overall, loneliness increased slightly but significantly over time ( = 0.00 , *p* = ) with significant variance in the slope ( = 0.00, *p* = 0.00). Previous loneliness scores significantly predicted later loneliness, over and above the overall trajectory of loneliness ( = 0.04, *p* = 0.00).

When loneliness was included in bivariate models to examine the relations between loneliness and cogitive function, as in the univariate models, the slope term was small and significant in the immediate word recall - loneliness model ( = 0.00, *p* = 0.03, = 0.00, *p* = 0.01). However, in the delayed word recall - loneliness and mental status - loneliness models the mean slope term of loneliness was not significant but the variance of the loneliness slope was small but significant (delayed word recall: = 0, *p* = 0.92, = 0.00, *p* = 0.01; mental status: = 0.00, *p* = 0.10, = 0, *p* = 0.04). In the bivariate models constraining the autoregressive parameters of loneliness on loneliness two years later to be stable over time did not significantly decrease model fit and was the more parsimonious model. However, in these models there was a consistent significant effect, over and above the estimated trajectory, of loneliness on loneliness two years later (immediate word recall: = 0.04, *p* = 0.05; delayed word recall: = 0.05, *p* = 0.03; mental status: = 0.10, *p* = 0.03).

There were some substantive differences when the autoregressive parameters for loneliness were constrained over time. In some of the bivariate models, when the autoregressive parameter of loneliness was allowed to vary over time there were occasions when the value was negative however, when constrained across time the loneliness autoregressive parameter was significantly positive as was the case in the unvariate models (see tables). This instability may be due to the fact that loneliness scores were only collected every four years as opposed to every two years like the cognitive data. Therefore, models with the autoregressive parameter constrained to equality were retained as the most parsimonious models with which the effects of covariates were investigated.

## Social Contact

For social contact, the full ALT model was the best fitting model according to model fit indices (see Table 6). The estimated intercept was ( = 1.38, *p* = 0) with significant variability. Social contact increased over time ( = 1.28 , *p* = 0) and the variance in the slope was not significant ( = 0.18, *p* = 0). The ability of earlier social contact to predict later social contact, over and above the overall trajectory of change was consistently significant but negative ( = -0.11, *p* = 0; = -0.20, *p* = 0; = -0.30, *p* = 0; = -0.41, *p* = 0; = -0.53, *p* = 0). When the autoregressive parameters are fixed over time, the slope becomes significant and negative, this is consistent with the univariate LGM as well. This suggests that when the autoregressive parameters are allowed to vary over time the decline in social contact over time is accounted for by the autoregressive parameters rather than in in the linear slope term, however, examing trajectory plots reveals that in both cases the mean predicted trajectory of social contact shows a decrease over time in social contact (see Figure 5).

For two of the three bivariate models with social contact (mental status - social contact, delayed word recall - social contact) the full ALT model was not able to be estimated due to convergence problems. Additional constraints on social contact were needed for models to converge without issue. There was no significant decrease in model fit when autoregressive parameters of social contact were constrained over time. In bivariate models with the autoregressive parameters of social contact constrained over time, the slope of social contact was significant and negative indicating that overall there was a decrease in social contact over time in the delayed word recall - social contact model only (delayed word recall: = -0.21, *p* = 0, = 0.15, *p* = 0.01). Whereas in the bivariate models with immediate word recall and mental status, the slope term for social contact was not significant (immediate word recall: = -0.08, *p* = 0.40, = 0.17, *p* = 0.00; mental status: = -0.00, *p* = 0.84, = 0.01, *p* = 0). In all three models there was significant variance in the slope term. Social contact did not significantly predict social contact two years later in any of the bivariate models (delayed word recall: = 0.00, *p* = 0.76; immediate word recall: = -0.02, *p* = 0.42; mental status: = 0.01, *p* = 0.49).

## Social Support

For univariate social support models, comparing model results showed that autoregressive parameters to equality over time did not result in significantly poorer model fit compared to the full ALT model and so is retained as the more parsimonious model. The intercept is ( = 9.78, *p* = 0) with significant variability. Social support did not show a significant mean trend over time ( = 0.00 , *p* = 0.73) but there was significant variance in the slope ( = 0.01, *p* = 0). Previous social support scores significantly predicted later social support, over and above the overall trajectory of social support ( = -0.02, *p* = 0). See Figure 7 for a plot of the mean predicted trajectory.

In the bivariate models investigating the relations between cognitive performance and social support, consistent with the univariate model, the slope of social support was consistently not significant but the variance of the slope term was significant (immediate word recall: = 0.00, *p* = 0.52, = 0.01, *p* = 0; delayed word recall: = 0.00, *p* = 0.50, = 0.01, *p* = 0; mental status: =-0.00, *p* = 0.84 = 0.01, *p* = 0). As in the univariate models, model fit was not significantly improved by allowing the autoregressive parameters of social support to vary over time in any of the bivariate models. For the mental status - social support bivariate model ( = 0.01, *p* = 0.49) autoregressive parameters were not significant. However, the immediate word recall - social support model ( = -0.03, *p* = 0) and delayed word recall - social support bivariate models ( = -0.03, *p* = 0) the autoregressive parameters were significant and negative.

## Social Network

Among the univariate social network models, the full ALT model showed significantly better fit than all models except the model with autoregressive parameters constrained across time which is superior as the more parsimonious model. The intercept was ( = 3.38, *p* = 0) with significant variability. Social network size significantly declined over time ( = -0.05 , *p* = 0) and there was significant variance in these trajectories ( = 0.00, *p* = 0). Previous social network scores significantly predicted later social network size, over and above the overall trajectory of social network ( = 0.04, *p* = 0).

In the three bivariate models investigating the relations between cognitive performance and social network, there was no significant mean slope in the bivariate immediate word recall - social network however the variance was significant ( = 0.07, *p* = 0.30, = 0.00, *p* = 0). In the bivariate models with delayed word recall and mental status there was a significant negative mean slope with significant variance (delayed word recall: = -0.05, *p* = 0, = 0.00, *p* = 0; mental status: =-0.06, *p* = 0.01 = 0.00, *p* = 0). For delayed word recall - social network bivariate model ( = 0.05, *p* = 0) the autoregressive parameters were significant and positive. However, for the mental status - social support bivariate model ( = 0.06, *p* = 0.15) and the immediate word recall - social network model ( = -0.05, *p* = 0.17) the autoregressive parameters were not significant. However, for all three of the bivariate models with cognitive performance where the social network autoregressive parameters were not constrained to equality over time the autoregressive parameters were consistently not significant.

## Bivariate model results

### *Immediate word recall - Loneliness*

The results of the bivariate immediate word recall - loneliness model indicate that the time specific correlations can be removed (model 9) without significantly changing the overall fit of the model. As described above autoregressive parameters for both immediate word recall and loneliness can also be constrained to equality over time without significantly changing the overall fit of the model (model 13), and the cross-lagged regressions of loneliness-on-immediate word recall (model 15) can also be constrained to be stationary over time. Thus model 15 is identified as the model with the best fit according to all fit indices examined and shows at least acceptable model fit across all indices. Looking at the results of this model in more detail, the correlation between the first measurement points was not significant (corr. = 0.01, *p* = 0.43). The correlation between the intercept factors was signficant (corr. = 0, *p* = 0.70). Note that in these models the intercept represents the portion of the time 2 variable remaining unexplained by the time 1 variable, thus this unexplained portion of immdediate word recall performance and loneliness were significantly related but initial levels of each were not. Further, that the time-specific correlations could be removed without changing the overall fit of the model indicates that the time-specific uniquenesses, after accounting for the slope, of loneliness and immediate word recall are not related. The cross-lagged regressions of loneliness-on-immediate word recall were consistently not signficant as indicated by the non-significant fixed cross-lagged regression parameters. However, the immediate-word-recall-on-loneliness cross-lagged regressions showed a small but significant effect that decreased to non-signficance over time. More specifically, the ability of state-like deviations in loneliness at time 1 significantly predicted state-like deviations in immediate word recall performance at time 2, and state-like deviations in loneliness at time 2 signficantly predicted state-like deviations in immediate word recall performance at time 3 but this was not true of state-like deviations in loneliness at time 3, 4, or 5.

### immediate word recall - social contact

The results of the bivariate models of immediate word recall and social contact indicate that although the bivariate LGM has acceptable model fit by a number of indices a direct comparison of the LGM nested within the ALT showed that the ALT is superior in model fit. Further it was revealed that the slope variance of social contact can be constrained to 0 without significantly affecting model fit but the variance was retained to maintain consistency with other models including social contact (CHECK THIS). For the bivariate ALT model the time-specific correlations of immediate word recall and social contact could be excluded without significantly decreasing model fit (model 9), further the autoregressive parameters for both immediate word recall and social contact could be constrained to equality over time without significantly decreasing model fit. However, the cross-lagged parameters cannot be constrained to equality without signficantly decreasing model fit. Thus, model 13 was identifed as the best fitting model according to the chi-square difference test. This model had excellent model fit according to all fit indices. Examining the results of this model in more detail, the correlation between inital immediate word recall performance and social contact was not significant (corr. = -0.43, *p* = 0.18). The correlation between the intercept factors was signficant (corr. = 0.02, *p* = 0). Estimating time-specific correlations between immediate word recall and social contact did not significantly affect model fit indicating that there was no significant relationship between the time-specific uniqueness of each process after accounting for overall slope. The autoregressive parameters were constrained across time and not significant indicating that deviations from the linear slope trajectory were not significantly related to deviations from that trajectory at the next occasion for either immediate word recall performance or social contact. The cross-lagged parameters were allowed to vary over time. The social contact-on-immediate word recall regressions were consistently not significant. However, the immediate word recall-on-social contact cross-lagged regressions became significant over time. More specifically, state-like deviations in time 1 social contact was not related to state-like deviations in time 2 immediate word recall nor were state-like deviations in time 2 social contact related to time 3 state like deviations in immediate word recall. However, state-like deviations in social contact at time 3 were significantly related to state-like deviations in immediate word recall at time 4. The relation between state-like deviations in social contact at time 4 and state-like deviations in immediate recall at time 5 just reached significance and this was maintained in the relation between time 5 social contact and time 6 immediate word recall. this indicates that the ability of deviations from the linear trajectory of social contact increase in their ability to predict deviations in the linear trajectory of immediate word recall performance over time from not being predictive to having a small but significant relation. However, this relation is interestingly negative suggesting that those with greater than expected social contact show worse than expected immediate word recall performance two years later. The relation between the linear slope terms of immediate word recall performance and social contact was significant and positive, indicating that overall trajectories of social contact and immediate word recall performance trend together but that deviations from these trajectories have different relations.

### *Immediate word recall - social support*

The model fit results of the bivariate ALT model of immediate word recall performance and self-reported social support are reported in Table X. The bivariate LGM had adequate model fit by the CFI, TLI, RMSEA, and SRMR. However the ALT model had superior model fit by all indices and was superior to the ALT-LGM according the ∆χ2. Estimating bivariate models with a series of constraints revealed that excluding the time-specific correlations between uniquenesses of immediate word recall performance and social support did not significantly decrease model fit meaning occasion specific uniquenesses, after accounting for overall linear slope, of these two processes were not significantly related (model 9). Further, as described above, the autoregressive parameters of both immediate word recall and social support can be constrained to equality over time without significantly changing model fit (model 13). Model 13 is the best fitting model according to all fit indices. Examining the results of this model more closely, the correlation between inital immediate word recall performance and social contact was not significant (corr. = 0.04, *p* = 0.36). The correlation between the intercept factors was also not signficant (corr. = , *p* = ). There was no significant association between the linear slopes of immediate word recall and social support (corr. = -0.00, *p* = 0.47). The cross-lagged immediate word recall-on-social support regressions could not be constrained to equality over time without a significant decrease in model fit (model 14). Examination of these parameters shows that time-specific deviations in social support initally predicted time-specific deviations in immediate word recall but that this effect decreases to non-significance over time. The association is positive indicating that, after accounting for linear trajectory, higher reported social support was associated with better immediate word recall two years later.

### *Immmediate word recall - social network*

Model fit indices and comparisons are presented in Table X. Examining model fit results, it appears the full ALT model had superior fit to both the autoregressive and LGM models. Further, examining the the ∆χ2 between the LGM nested within the ALT model and the full ALT showed that estimating the full ALT significantly improved model fit. However, excluding time-specific correlations between the uniquenesses of immediate word recall and social network did not significantly decrease model fit and so they were excluded in favour of the more parsimonious model (model 9). Further, as described above the autoregressive parameters were constrained to be equal over time as this constraint also did not signficantly affect model fit (model 11). Thus, model 11 was retained as the model with the best fit according to all fit indices examined except for SRMR which showed very slightly better fit for the model with the social network autoregressions also fixed. Model 11 showed good model fit according to all indices. Examining the results of this model more closely revealed that the correlation between inital immediate word recall performance and social network was not significant (corr. = 0.01, *p* = 0.70). However, correlation between the intercept factors was signficant (corr. = , *p* = ). Constraining cross-lagged immediate word recall-on-social network were contrained to equality over time resulted in poorer model fit and so the parameters were allowed to vary over time, but still remained consistentently not significant. However, the social network-on-immediate word recall cross lagged regressions showed a variable affect over time. Deviations from the linear slope in immediate word recall performance did not predict deviations from the linear slope in social network size two years later for the first measurement occasions but immediate word recall at time 3 showed a small but significantly negative relationship with social network size at time 4. Time specific deviatios in immediate recall performance at time 4 did not predict deviations in social network size at time 5 however, deviations in immediate recall performance at time 5 significantly and positively predicted deviations in social network size at time 6. There was no significant relationship between the linear slope trajectories of immediate word recall performance and social network (0, *p* = 0.42).

### delayed word recall - loneliness

The model fit results of the bivariate ALT model of delayed word recall performance and self-reported loneliness are reported in Table 2. The bivariate LGM had adequate model fit by the CFI, TLI, RMSEA, and SRMR. However the ALT model had superior model fit by several indices and was superior to the ALT-LGM according the ∆χ2. Model fit results show that the time specific correlations between delayed word recall performance and loneliness can be removed, and the autoregressive parameters for both processes can be fixed to equality over time without significanty decreasing model fit. Constraining the cross-lagged regressions to equality resulted in significantly poorer model fit. Although the model fit decrease when the variance of the slope term for loneliness was fixed to 0, the slope term and variance was retained because the although there was no significant overall slope for loneliness there was a small but significant variance in slope. Thus, the the ALT model without time-specific correlations between delayed word recall and loneliness, and fixed autoregressions for both processes was retained. Results reveal that deviations from the predicted trajectory in delayed word recall performance did not predict later delayed word recall performance. Deviations from the the predicted trajectory of loneliness did significantly predict later deviations ( = , *p* = ). Examining the results of this model more closely revealed that the correlation between inital delayed word recall performance and social network was not significant (corr. = -0.01, *p* = 0.51). However, correlation between the intercept factors was signficant (corr. = , *p* = ). There was no significant relation between the trajectory of delayed recall performance and self-reported loneliness (0, *p* = 0.63). The cross-lagged regression terms both delayed word recall - on - loneliness and loneliness - on - delayed word recall could not be contrained over time without a significant decrease in model fit. The delayed word recall - on - loneliness cross-lagged regressions were largely not significant with the exception of deviations in loneliness at time 2 significantly predicting deviations in delayed word recall performance at time 3. None of the loneliness - on - delayed word recall cross-lagged regressions were significant indicating that time-specific deviations, after accounting for linear change, in delayed recall performance was not significantly related to any later deviations in self-reported loneliness, after accounting for overall level and linear change in loneliness. The covariates were then added to this model.

### *delayed word recall - social contact*

The model fit results of the bivariate ALT model of delayed word recall performance and social contact are reported in Table X. The bivariate LGM had adequate model fit by the CFI, TLI, RMSEA, and SRMR. However the ALT model had superior model fit by several indices and was superior to the LGM nested within the ALT according the ∆χ2. The correlations of time-specific uniquenesses can be excluded without a significant reduction in model fit (model 9). The autoregressive parameters of delayed word recall and the autoregressive parameters of social contact can both be constrained to equality over time without significantly decreasing model fit. However, constraining the cross-lagged parameters of delayed word recall - on - social contact and social contact - on - cognitive resulted in significantly poorer model fit. Model 12 is the best fitting model according the ∆χ2. The results of this model show that the correlation between intital delayed word recall performance and intital social contact was not significant (corr. = -0.09, *p* = 0.81). The correlation between the intercept terms of delayed word recall and social contact was significant (corr. = , *p* = ). There was a significant relationship between the trajectory of delayed recall performance and social contact (0.02, *p* = 0.00). Although they were allowed to vary over time all cross-lagged regression parameters were not significant indiating that there was no significant relationship between time-specific deviations from their linear trajectories were not related.

### *delayed word recall - social support*

The model fit results of the bivariate models of delayed word recall and social supported are presented in Table X. The bivariate LGM showed adequate model fit by three of the indices the CFI, TLI, and RMSEA. However, the model fit of the full ALT model appeared superior by examining these indices and the LGM nested within the ALT showed significantly poorer model fit according to the ∆χ2. The progressive addition of further constraints showed that the correlations of time-specific uniquenesses of delayed word recall performance and social support could be excluded without a significant reduction in model fit (model 9). As described above the autoregressions of both delayed word recall and social support could be constrained to equality over time without a significant reduction in model fit (model 13). Constraining the cross-lagged regressions of delayed word recall - on - social support resulted in a significant loss in model fit (model 14) but the cross-lagged regressions of social support on delayed word recall did not result in a significant reduction in model fit (model 15). Thus model 15 is the best fitting model according to most fit indices and the ∆χ2 test. The results of this model show that the correlation between inital delayed recall performance and initial self reported social support was not signicant 0.01, *p* = 0.93). The correlation between the intercept terms of delayed word recall and social contact was also not significant (corr. = , *p* = ). There was no significant relationship between the trajectories of delayed word recall performance and social support (-0.00, *p* = 0.07). The social support - on - delayed word recall cross-lagged regressions were not significant. The cross-lagged regression terms of delayed word recall performance - on - social support showed a consistently positive relationship that ranged from 0.018 to 0.034 and although the parameter estimate was consistently within this range only the first two waves were significant.

### *delayed word recall - social network*

The model fit results of the bivariate models of delayed word recall are presented in Table . The bivariate autoregressive model had inadequate model fit. The unconditional LGM showed good model fit according to all fit indicies however the model fit of the full ALT appeared better according to the CFI, TLI, RMSEA and SRMR. Further, the LGM nested within the ALT model showed a significant reduction in model fit from the full ALT model. Unlike the other bivariate ALT models, excluding the time specific uniquenesses signficantly decreased model fit. However, examining the time-specific uniqueness' correlations between delayed word recall and social network did not reveal a decernible pattern. There was a significant correlation between time-specific uniquenesses at times 2 and 3 but the direction of the correlations were opposite of one another with a positive relationship at time 2 and a negative relationship at time 3. As described above the autoregressive parameters for delayed word recall could be constrained to equality over time but the autoregressive parameters for social network could not be constrained without a significant reduction in model fit. The delayed word recall - on - social network parameter could not be constrained to equality without a significant reduction in model fit (model 14) but the social network on delayed word recall regressions could be constrained (model 15). However, none of either the social network - on - delayed word recall or delayed word recall - on - social network cross-lagged regressions were significant. In this model, the correlation between inital delayed recall performance and initial social network was not signicant 0.04, *p* = 0.25). However, the correlation between the intercept terms of delayed word recall and social network was significant (corr. = , *p* = ). There relationship between the trajectories of delayed word recall performance and social network was trending towards significance (0.00, *p* = 0.06). Examining the final model (model 15) compared to the full ALT, there were no changes in model interpretation except for the two significant time-specific correlations.

### *mental status - loneliness*

The results of the bivariate models with mental status and loneliness are presented in Table . The bivariate autoregressive model had inadequate model fit. The unconditional bivariate LGM showed adequate model fit but the model fit of the full ALT was superior and in the direct comparison between the LGM nested within the ALT and the full ALT model the nested LGM showed signficantly poorer model fit according to the ∆χ2. Progessively investigating model constraints showed that the following constraints could be added without a significant reduction in model fit: excluding the loneliness slope, excluding the time specific correlations between loneliness and mental status, and constraining the loneliness autoregressive parameters to equality over time. Thus, the model that is the best fitting according to the model fit indices is model 15. Examining the parameter estimates of this model reveals that the inital values of mental status and loneliness are not significantly correlated, (corr. = 0.00, *p* = 0.71 however the intercept terms showed a significant relationship (corr. = , *p* = ). As described above mental status showed a significant mean linear decline over time with significant autoregressive parameters, meaning that previous mental status performance significantly predicted mental status two years later after accounting for the overall decline over time. The autoregressive parameters of loneliness were also positive and significant while the linear slope term could be excluded indicating that loneliness was best explained by previous loneliness performance. The cross-lagged regressions were not significant for mental status - on - loneliness were constrained to equality over time and were not significant indicating that loneliness at one occasion did not predict mental status two years later (mental status - on - loneliness = 0.11, *p* = 0.01). However, the loneliness - on - mental status cross-lagged regressions were also stationary over time but significant and negative indicating that poorer mental status predicted greater loneliness two years later (loneliness-on-mental status = -0.02, *p* = 0.04). The parameter estimates, and thus conclusions drawn from Model 15 differed from the full ALT model is several important aspects. First, in full ALT model, consistent with the univariate model of loneliness and the other bivariate models including loneliness, the slope of loneliness was significant and positive indicating that overall loneliness increases over time ( = 0.08, *p* = 0.01) but there is also significant variance in the trajectories of loneliness indicating that not all individuals show a pattern of increasing loneliness ( = 0.00, *p* = 0.02). Second, in the full ALT model, loneliness at time 1 significantly and positively predicted mental status at time 2, after which loneliness was not predictive of mental status two years later. Lastly in the full ALT model although the parameter etimates of loneliness - on - mental status are similar (-0.01to -0.02) to loneliness - on - mental status cross-lagged regression in model 15 (-0.01, *p* = 0.11) except in the full ALT model they do not reach the level of significance except for the first loneliness - on - mental status regression (*p* = 0.04).

### *mental status - social contact*

The results of the bivariate mental status and social contact are presented in Table Y. The bivariate autoregressive model had inadequate model fit. The bivariate LGM had adequate model fit however, the model fit indices for the full ALT model appeared superior except by the SRMR. The mental status - social contact full ALT model did not properly converge, nor did the ALT model with mental status slope variance constrained to 0 or the ALT model with the mental status slope excluded. Subsequent models (model 7 through model 14) with additional constraints did converge without issue. Although the full ALT model did not converge properly, the LGM nested in ALT showed significantly poorer model fit according to the ∆χ2 than the ALT models. Progressivley constraining processes showed that the following constraints can be added without a significant reduction in model fit: constraining the variance of the social contact slope 0 and excluding time-specific correlations. However, examining the model parameters independently revealed that there were some differences between different models that affect substantive interpretation.

### *mental status - social support*

The results of the bivariate mental status and social support are presented in Table Y. The bivariate autoregressive model had inaedquate model fit. The bivariate LGM showed adequate model fit by all indices except SRMR which indicated inadequate model fit. The full ALT model showed better model fit across all indices and the LGM nested within the ALT had significantly poorer model fit according the ∆χ2. Progessively investigating model constraints showed that the following constraints could be added without a significant reduction in model fit: time-specific correlations can be excluded, the autoregressive parameters of social support can be constrained to equality over time, social support - on - mental status cross-lagged parameters can be constrained to equality over time. Thus, model 14 is the best fitting model according to χ2 difference testing. In this model there was no significant relationship between the first occasions of mental status and social support (corr. = 0.00, *p* = 0.95 or between intercept terms (corr. = , *p* = ). There was a significant decline in mental status over time with significant variance in the slope. Social support showed no signficant mean trend over time but there was significant variance in the slope term indicating that some individuals may have trajectories of increasing social support while others show decreasing social support over time. Examining the relations between change in mental status and change in social support, there was no significant relationship between the linear slope of mental status and the linear slope of social support  
(corr. = 0.00, *p* = 0.10). The cross-lagged regressions of mental status - on - social support could not be constrained to equality over time without a significant reduction in model fit. Social support at time 1 significantly predicted mental status at time 2 (0.03, *p* = 0.00) and social support at time 2 significantly predicted mental status at time 3 (0.03, *p* = 0.00) however, time-specific uniquenesses in social support did not predict time-specific uniquenesses in mental status at other occasion. The cross-lagged regressions of social support - on - mental status were constrained to equality over time and were not significant (-0.04, *p* = 0.10). Notably though, there are differences in parameter significance between the full ALT model and model 15, best supported by fit indices. Specifically, in the full ALT model the social support linear slope term was positive and significant (-0.14, *p* = 0.65) and the social support autoregressive parameters were significant and negative from time 3 onwards. Thus, it appears that the shape of the trajectory of social support is explained differently in the ALT model with and without constraints. However, the relations between mental status and social support remained consistent across the full ALT model and model 15 with constraints.

### *mental status - social network* I

The results of the bivariate mental status and social network are presented in Table Z. The bivariate autoregressive model showed inadequate model fit. The bivariate LGM showed adequate model fit but the full ALT model was superior by all indices except for SRMR. Further, the LGM nested within the ALT showed significantly poorer fit than the full ALT model according the ∆χ2. Conducting a model fit comparison of models with progressively added model contraints showed that the following constraints could be added without a significant decrease in model fit: the variance of the social network slope could be constrained to 0, the correlations of time-specific uniquenesses can be excluded, the autoregressive parameters for social network can be constrained to equality over time, and both the social network - on - mental status and the mental status - on - social network cross-lagged regression parameters can be constrained to equality over time. Thus, the best fitting model according the the ∆χ2 test comparisons was model 15. This model also showed good model fit by all other fit indices. In this model there was no significant correlation between the first included mental status and social network scores (corr. = 0.00, *p* = 0.79). However there was a significant correlation between intercept terms (corr. = , *p* = ). There was no significant relationship between the linear slopes of mental status and social network (0, *p* = 0.86). Social network significantly predicted mental status two years later (0.05, *p* = 0.01) such that higher time specific uniquenesses in social network predicted greater time specific uniquenesses in mental status. Mental status did not significantly predict later social network (-0.01, *p* = 0.51). It should be noted that in the full ALT model although the parameter estimates were similiar and often higher, ranging from (0.05 to 0.11) the regression of time-specific uniquenesses in mental status on time-specific uniquenesses in social network reached statistical significance at only time 2 and time 3.

## Bivariate models with covariates

# Discussion

# References

Table 1 *Descriptive statistics by year*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
|  | 2004 | 2006 | 2008 | 2010 | 2012 | 2014 |
|  | M (SD) | M (SD) | M (SD) | M (SD) | M (SD) | M (SD) |
|  | n = 5531 | n = 5720 | n = 5810 | n = 5698 | n = 5165 | n = 4454 |
| Women (%) | 59.68 | 51.21 | 50 | 50 | 50 | 50 |
| Age | 72.13 (5.82) | 74.07 (5.83) | 76.01 (5.85) | 78.31 (5.77) | 79.62 (5.47) | 81.15 (5.26) |
| Yrs Education | 12.38 (3.1) | 12.38 (3.1) | 12.38 (3.1) | 12.38 (3.1) | 12.38 (3.1) | 12.38 (3.1) |
| Health Conditions | 1.96 (1.19) | 2.13 (1.22) | 2.28 (1.23) | 2.47 (1.25) | 2.53 (1.26) | 2.58 (1.26) |
| Mental status | 8.53 (0.79) | 8.51 (0.81) | 8.44 (0.88) | 8.07 (1.1) | 8.09 (1.16) | 7.97 (1.32) |
| Word recall immediate | 5.45 (1.5) | 5.31 (1.53) | 5.18 (1.54) | 4.86 (1.64) | 4.74 (1.63) | 4.63 (1.65) |
| Word recall delayed | 4.4 (1.84) | 4.23 (1.89) | 4.12 (1.88) | 3.75 (1.95) | 3.61 (1.97) | 3.49 (1.96) |
| Psychosocial Variables | n = 1061 | n = 2787 | n = 2737 | n = 2646 | n = 2235 | n = 2031 |
| Loneliness | 1.35 (0.47) | 1.43 (0.51) | 1.44 (0.51) | 1.43 (0.51) | 1.46 (0.5) | 1.43 (0.51) |
| Social contact | 30.56 (8.37) | 29.6 (8.14) | 29.61 (8.6) | 29.41 (8.49) | 28.99 (8.78) | 28.65 (8.75) |
| Social support | 9.81 (1.53) | 9.58 (1.52) | 9.56 (1.6) | 9.58 (1.56) | 9.61 (1.59) | 9.58 (1.61) |
| Depression | 1.16 (1.72) | 1.23 (1.75) | 1.23 (1.73) | 1.3 (1.78) | 1.34 (1.82) | 1.38 (1.86) |
| Social network | 3.44 (0.76) | 3.41 (0.72) | 3.31 (0.77) | 3.25 (0.79) | 3.12 (0.84) | 3.04 (0.86) |

Table 2 *Model Fit Indices for Immediate Word Recall*

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Model |  | df | CM |  | df | CFI | TLI | RMSEA | SRMR |
| Autoregressive, univariate | 2803.192 | 10 | - | NA | NA | 0.631 | 0.446 | 0.217 | 0.231 |
| LGM | 112.905 | 16 | - | NA | NA | 0.987 | 0.988 | 0.032 | 0.014 |
| LGM, quadratic | 102.364 | 15 | - | NA | NA | 0.988 | 0.988 | 0.031 | 0.014 |
| ALT, full model | 64.897 | 11 | 4 | NaN | 0 | 0.993 | 0.990 | 0.029 | 0.015 |
| LGM, nested in ALT | 112.905 | 16 | 4 | 48.03 | 5 | 0.987 | 0.988 | 0.032 | 0.014 |
| ALT, no slope variance | 115.864 | 13 | 4 | 49.04 | 2 | 0.986 | 0.984 | 0.036 | 0.021 |
| ALT, no slope | 646.224 | 14 | 4 | 538.75 | 3 | 0.916 | 0.910 | 0.087 | 0.058 |
| ALT, fixed regressions | 99.613 | 15 | 4 | 34.67 | 4 | 0.989 | 0.989 | 0.031 | 0.014 |

Table 3 *Model Fit Indices for Mental Status*

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Model |  | df | CM |  | df | CFI | TLI | RMSEA | SRMR |
| Autoregressive, univariate | 1189.141 | 10 | - | NA | NA | 0.811 | 0.717 | 0.141 | 0.181 |
| LGM | 489.443 | 16 | - | NA | NA | 0.924 | 0.929 | 0.071 | 0.037 |
| LGM, quadratic | 401.335 | 15 | - | NA | NA | 0.938 | 0.938 | 0.066 | 0.038 |
| ALT, full model | 92.170 | 11 | 4 | NaN | 0 | 0.987 | 0.982 | 0.035 | 0.029 |
| LGM, nested in ALT | 489.443 | 16 | 4 | 495.56 | 5 | 0.924 | 0.929 | 0.071 | 0.037 |
| ALT, no slope variance | 156.623 | 13 | 4 | 67.76 | 2 | 0.977 | 0.973 | 0.043 | 0.034 |
| ALT, no slope | 1132.965 | 14 | 4 | 1006.75 | 3 | 0.821 | 0.808 | 0.116 | 0.132 |
| ALT, fixed regressions | 330.187 | 15 | 4 | 295.85 | 4 | 0.950 | 0.950 | 0.059 | 0.037 |

Table 4 *Model Fit Indices for Loneliness*

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Model |  | df | CM |  | df | CFI | TLI | RMSEA | SRMR |
| Autoregressive, univariate | 267.307 | 10 | - | NA | NA | 0.880 | 0.819 | 0.066 | 0.254 |
| LGM | 33.187 | 16 | - | NA | NA | 0.992 | 0.992 | 0.013 | 0.172 |
| LGM, quadratic | 29.866 | 15 | - | NA | NA | 0.993 | 0.993 | 0.013 | 0.172 |
| ALT, full model | 12.203 | 11 | 4 | NaN | 0 | 0.999 | 0.999 | 0.004 | 0.115 |
| LGM, nested in ALT | 33.187 | 16 | 4 | 19.19 | 5 | 0.992 | 0.992 | 0.013 | 0.172 |
| ALT, no slope variance | 30.746 | 13 | 4 | 17.04 | 2 | 0.992 | 0.990 | 0.015 | 0.182 |
| ALT, no slope | 31.193 | 14 | 4 | 16.03 | 3 | 0.992 | 0.991 | 0.014 | 0.175 |
| ALT, fixed regressions | 21.152 | 15 | 4 | 8.42 | 4 | 0.997 | 0.997 | 0.008 | 0.177 |

Table 5 *Model Fit Indices for Social Contact*

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Model |  | df | CM |  | df | CFI | TLI | RMSEA | SRMR |
| Autoregressive, univariate | 248.593 | 10 | - | NA | NA | 0.916 | 0.874 | 0.063 | 0.170 |
| LGM | 34.332 | 16 | - | NA | NA | 0.994 | 0.994 | 0.014 | 0.076 |
| LGM, quadratic | 29.462 | 15 | - | NA | NA | 0.995 | 0.995 | 0.013 | 0.076 |
| ALT, full model | 8.520 | 11 | 4 | NaN | 0 | 1.000 | 1.001 | 0.000 | 0.064 |
| LGM, nested in ALT | 34.332 | 16 | 4 | 24.91 | 5 | 0.994 | 0.994 | 0.014 | 0.076 |
| ALT, no slope variance | 24.980 | 13 | 4 | 14.01 | 2 | 0.996 | 0.995 | 0.012 | 0.085 |
| ALT, no slope | 64.647 | 14 | 4 | 49.63 | 3 | 0.982 | 0.981 | 0.025 | 0.063 |
| ALT, fixed regressions | 34.307 | 15 | 4 | 25.19 | 4 | 0.993 | 0.993 | 0.015 | 0.075 |

Table 6 *Model Fit Indices for Social Support*

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Model |  | df | CM |  | df | CFI | TLI | RMSEA | SRMR |
| Autoregressive, univariate | 326.229 | 10 | - | NA | NA | 0.897 | 0.845 | 0.073 | 0.270 |
| LGM | 73.949 | 16 | - | NA | NA | 0.981 | 0.982 | 0.025 | 0.213 |
| LGM, quadratic | 60.922 | 15 | - | NA | NA | 0.985 | 0.985 | 0.023 | 0.213 |
| ALT, full model | 39.097 | 11 | 4 | NaN | 0 | 0.991 | 0.987 | 0.021 | 0.179 |
| LGM, nested in ALT | 73.949 | 16 | 4 | 34.24 | 5 | 0.981 | 0.982 | 0.025 | 0.213 |
| ALT, no slope variance | 74.973 | 13 | 4 | 35.27 | 2 | 0.980 | 0.977 | 0.028 | 0.208 |
| ALT, no slope | 72.677 | 14 | 4 | 29.70 | 3 | 0.981 | 0.979 | 0.027 | 0.201 |
| ALT, fixed regressions | 44.013 | 15 | 4 | 4.99 | 4 | 0.991 | 0.991 | 0.018 | 0.210 |

Table 7 *Model Fit Indices for Social Network*

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Model |  | df | CM |  | df | CFI | TLI | RMSEA | SRMR |
| Autoregressive, univariate | 140.513 | 10 | - | NA | NA | 0.940 | 0.909 | 0.047 | 0.179 |
| LGM | 56.197 | 16 | - | NA | NA | 0.981 | 0.983 | 0.021 | 0.106 |
| LGM, quadratic | 23.042 | 15 | - | NA | NA | 0.996 | 0.996 | 0.009 | 0.107 |
| ALT, full model | 8.966 | 11 | 4 | NaN | 0 | 1.000 | 1.001 | 0.000 | 0.121 |
| LGM, nested in ALT | 56.197 | 16 | 4 | 46.82 | 5 | 0.981 | 0.983 | 0.021 | 0.106 |
| ALT, no slope variance | 35.653 | 13 | 4 | 20.77 | 2 | 0.990 | 0.988 | 0.017 | 0.117 |
| ALT, no slope | 49.113 | 14 | 4 | 32.70 | 3 | 0.984 | 0.983 | 0.021 | 0.122 |
| ALT, fixed regressions | 18.739 | 15 | 4 | 9.70 | 4 | 0.998 | 0.998 | 0.006 | 0.111 |

Table 8 *Model Fit Indices for Delayed Word Recall and Loneliness*

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Model |  | df | CM |  | df | CFI | TLI | RMSEA | SRMR |
| LGM, bivariate unconditional | 203.454 | 64 | - | NA | NA | 0.988 | 0.988 | 0.019 | 0.088 |
| Autoregressive model, bivariate | 3400.978 | 40 | - | NA | NA | 0.711 | 0.523 | 0.119 | 0.175 |
| ALT, full model | 99.183 | 38 | - | NA | NA | 0.995 | 0.991 | 0.016 | 0.059 |
| ALT, nested LGM | 203.482 | 63 | 3 | 103.09 | 25 | 0.988 | 0.987 | 0.019 | 0.088 |
| ALT, no cognitive slope variance | 170.430 | 42 | 3 | 65.11 | 4 | 0.989 | 0.983 | 0.023 | 0.058 |
| ALT, no cognitive slope | 372.684 | 43 | 3 | 253.46 | 5 | 0.972 | 0.957 | 0.036 | 0.067 |
| ALT, no social slope variance | 112.802 | 42 | 3 | 13.31 | 4 | 0.994 | 0.990 | 0.017 | 0.093 |
| ALT, no social slope | 115.320 | 43 | 3 | 15.69 | 5 | 0.994 | 0.990 | 0.017 | 0.088 |
| ALT, no time specific correlations | 102.667 | 43 | 3 | 4.15 | 5 | 0.995 | 0.992 | 0.015 | 0.056 |
| 10-ALT-09 + fixed autoregressions for cognitive | 109.974 | 47 | 9 | 7.33 | 4 | 0.995 | 0.992 | 0.015 | 0.056 |
| 11-ALT-09 + fixed autoregressions for social | 109.270 | 47 | 9 | 6.90 | 4 | 0.995 | 0.992 | 0.015 | 0.091 |
| ALT-12-covariates | 126.953 | 61 | 9 | 24.25 | 18 | 0.995 | 0.993 | 0.013 | 0.080 |
| ALT-12-covariates | 147.015 | 73 | 17 | -21.15 | 20 | 0.995 | 0.992 | 0.013 | 0.071 |
| ALT-9, plus autoregressions fixed across time | 116.862 | 51 | 12 | -9.35 | -10 | 0.994 | 0.993 | 0.015 | 0.091 |
| 14 - ALT-cognitive quadratic, no social slope | 183.715 | 59 | 17 | 21.88 | 6 | 0.989 | 0.988 | 0.019 | 0.089 |
| ALT-12 fixed autoregressions and fixed cross-lagged | 170.688 | 59 | 17 | 9.37 | 6 | 0.990 | 0.989 | 0.018 | 0.090 |
| 15 - ALT-cognitive quadratic | 161.463 | 53 | 3 | 61.32 | 15 | 0.991 | 0.988 | 0.019 | 0.090 |
| 16 - ALT-cognitive quadratic, cog fixed | 94.127 | 41 | 3 | -4.81 | 3 | 0.995 | 0.993 | 0.015 | 0.062 |
| 14 - ALT-cognitive quadratic, full model | 87.463 | 37 | 3 | -9.64 | -1 | 0.996 | 0.992 | 0.015 | 0.063 |

Figure 1

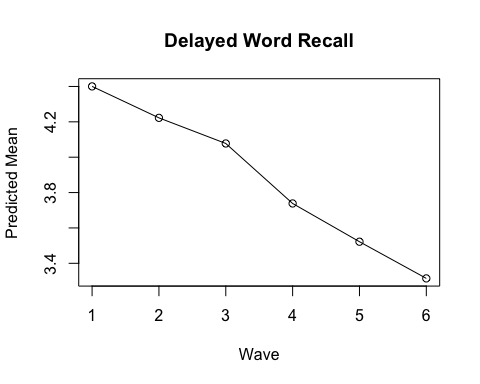
Figure 2 

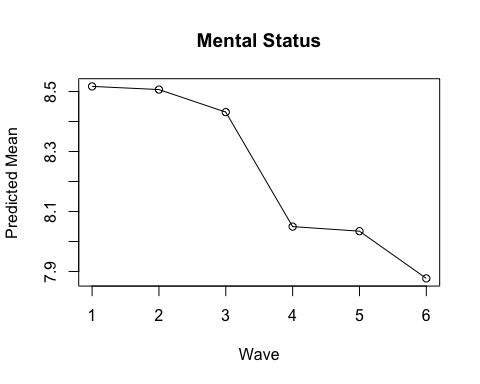
Figure 3 

Figure 4

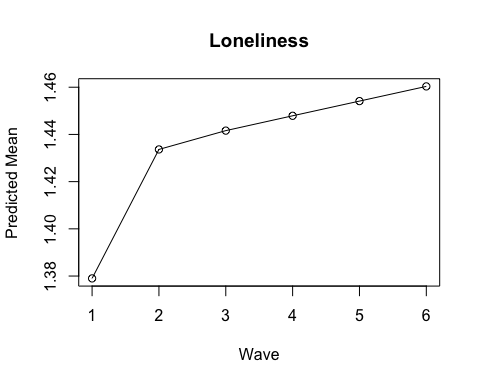


Figure 5

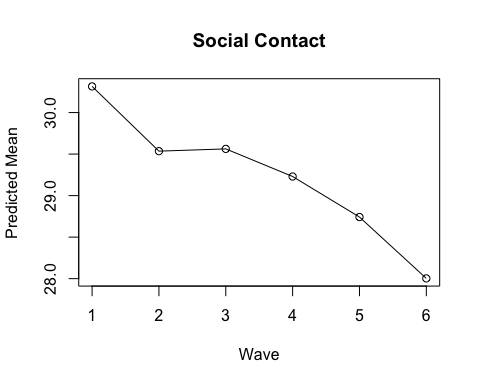


Figure 6

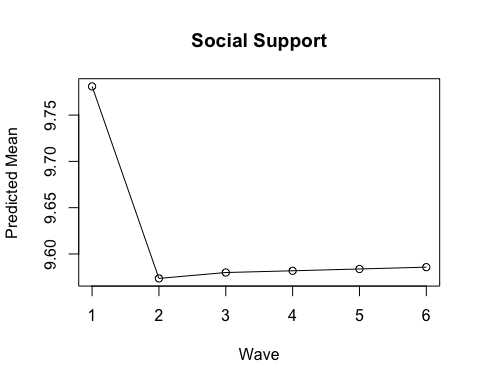


Figure 7

