1

2 Abstract

- Prior research suggests that across a wide range of cognitive, educational, and health-based
- 4 measures, first-born children outperform their later-born peers. Expanding on this literature
- 5 using naturalistic home-recorded data and parental vocabulary report, we find that early
- 6 language outcomes vary by number of siblings in a sample of 43 English-learning U.S.
- 7 children from mid-to-high socioeconomic status homes. More specifically, we find that
- s children in our sample with two or more but not one older siblings had smaller productive
- 9 vocabularies at 18 months, and heard less input from caregivers across several measures than
- their peers with less than two siblings. We discuss implications regarding what infants
- experience and learn across a range of family sizes in infancy.
- 12 Keywords: Siblings, Lexical Development, Input Effects, Language Acquisition

13

Analysing the effect of sibling number on input and output in the first 18 months

A common simplifying assumption in research on language development is that there is 14 a theoretical "optimum" environment for early language, whereby the input is tailored to a 15 single infant's needs, changing over time as language capacity grows (e.g. Soderstrom, 2007; 16 Stern, Spieker, Barnett, & MacKain, 1983). However, for many infants and for many reasons, 17 language acquisition occurs across diverse social contexts that can influence the learning environment, including the presence of older siblings in the home (Fenson et al., 1994). According to the United States Census Bureau (2010), around one third of children are born into households with at least one other infant present, and one in every five infants is acquiring language in a household shared with two or more other children. Similar statistics are reported for British infants (Office for National Statistics, 2018), where the average household has 1.75 children, and 15% of households have three children or more. More globally, in most parts of the world, few children grow up without siblings (United-Nations, 25 2017). In this paper, we consider the role of siblings in the early language environment of English-learning infants. We use naturalistic home-recorded data to measure input in earlierand later-born infants, alongside their productive vocabulary over the first 18 months of life.

Prior research suggests that infants born to households with older children may be
slower to learn language. Fenson and colleagues (1994) found that by 30 months of age,
children with older siblings performed worse than those with no siblings across
parent-reported measures of productive vocabulary, use of word combinations, and mean
length of utterance. This 'sibling effect' may be the result of differences in input between
first- and later-born children: some research finds that infants with older siblings hear less
speech aimed specifically at them, and what they do hear is understood to be linguistically
less supportive of early language development (Hoff-Ginsberg, 1998; Oshima-Takane &
Robbins, 2003). In contrast, some studies have noted linguistic advantages for later-borns,
who may have stronger social-communicative skills (Hoff, 2006), better understanding of

what aspects of language are being measured.

- pronouns (Oshima-Takane, Goodz, & Derevensky, 1996), and better conversational abilities
 (Dunn & Shatz, 1989). Overall, while the particulars differ across studies, prior work
 suggests that the presence of siblings in the home leads to differences in infants' early
 linguistic experiences and skills, though the direction of these differences varies depending on
- Numerous studies have attempted to better understand how siblings affect the language development trajectory, with comparisons of language acquisition across first- and later-borns. Here again, findings are mixed, but overall two general conclusions can be drawn. First, analyses consistently show that infants with older siblings generally have slower vocabulary development (Berglund, Eriksson, & Westerlund, 2005; Fenson et al., 1994; Pine, 1995; Zambrana, Ystrom, & Pons, 2012), and this effect increases with number of older siblings (Gurgand et al., 2022; Karwath, Relikowski, & Schmitt, 2014; Peyre et al., 2016). Furthermore, this finding is consistent across cultures (e.g. European French (Gurgand et al., 51 2022; Havron et al., 2019); Singaporean (Havron et al., 2022); Kenyan (Jakiela, Ozier, Fernald, & Knauer, 2020); and German (Karwath et al., 2014)). However, this finding is not as clear-cut as has been previously assumed: Hoff-Ginsberg (1998) shows first-borns to have better lexical and syntactic skills up until 2:5, but later-born infants had better conversational abilities during the same time-period. Recent studies have also identified effects for age gap between the target child and their siblings (whereby larger age gaps correlate with lower vocabulary scores, Gurgand et al., 2022; Havron et al., 2022) and sibling sex (whereby older brothers have a negative effect on vocabulary outcomes, but not older sisters, Havron et al., 2019; Jakiela et al., 2020) though neither of these effects are found consistently across datasets; Havron et al. (2022) and Gurgand et al. (2022) find no effect for sibling sex, whereas Havron et al. (2019) find no effect for age gap. Some of these differences across studies may relate to insufficient power to detect relatively small effects, perhaps leading to under- or over-estimation of effect sizes, or simultaneous contributing factors that are difficult to disentangle.

The second general finding pertains to sibling-related differences in the early linguistic 66 environment: infants with no siblings receive more input overall, and this more closely 67 reflects what is typically considered to be 'high quality' input in the extant literature 68 (i.e. more input in an infant-directed speech style (Ramírez-Esparza, García-Sierra, & Kuhl, 69 2014); longer utterance length (Barnes, Gutfreund, Satterly, & Wells, 1983); higher lexical diversity (Rowe & Snow, 2020)). Indeed, the very presence of a sibling in the linguistic 71 environment changes the way language is used. When siblings are present (i.e. triadic interactions), mothers' input has been found to be more focused on regulating behaviour, as opposed to the language-focused speech that is common in dyadic contexts (Oshima-Takane & Robbins, 2003). Reports show that the mean length of utterance is longer in the input of first-born infants (Hoff-Ginsberg, 1998; but see also Oshima-Takane & Robbins, 2003 for a comparison of dyadic and triadic contexts), who also hear more questions directed at them than later-borns. Both Jones and Adamson (1987) and Oshima-Takane and Robbins (2003) report no difference between the overall number of word types produced by mothers in dyadic and triadic settings, but the proportion of speech directed at the target infant is drastically reduced when input is shared with siblings.

As Hoff (2006) explains, infants with siblings have less experience of speech directed at them, but they do have an advantage over their first-born peers in that they are subject to more overheard speech. This may be an important source of input for infants with one or more older siblings. Akhtar, Jipson and Callanan (2001) show that, by age 2;6, children can learn both novel object labels and novel verbs through overhearing. Slightly younger children (aged 1;11-2;2) were also able to learn the novel object labels, but not verb labels.

Two-year-old infants can even learn novel object labels while doing activities that distract them from the language input, and when the novel words are produced non-saliently (Akhtar, 2005). This suggests that, while the learning environment for later-borns might differ from that of first-born infants, there may be ample opportunity for them to learn from the speech that surrounds them; namely overheard speech directed at their older sibling(s). Evidence is

mainly drawn from work testing infants aged 2 and above (e.g. Akhtar, 2005; Fitch,
Liberman, Luyster, & Arunachalam, 2020; Foushee, Srinivasan, & Xu, 2021), and generally
relies on experimental work rather than observations of the home environment. However,
Floor and Akhtar (2006) tested younger infants to find that the capacity to learn from
overheard speech is available from as early as 16 months, at least in an experimental setting.

There thus may be a trade-off, even in early development, between highly supportive 98 one-to-one input from a caregiver (cf. Ramírez-Esparza et al., 2014) and the potential 99 benefits drawn from communicating with (or overhearing communication with) a sibling. In 100 the present study, we test the extent to which having more versus fewer siblings in the home 101 environment may affect the linguistic environment in ways that could lead to differences in 102 vocabulary development over the course of the first 18 months of life. In analyzing infants' 103 growing productive vocabulary and linguistic environment in relation to the presence of older 104 siblings in their household, the present work expands on the extant literature in two key 105 ways. First, much of the existing literature identifying links between sibling number and 106 vocabulary outcomes draws on large-scale questionnaire data, rather than naturalistic 107 day-to-day interactions in the home. In contrast, we analyze an existing corpus of home 108 recordings in concert with vocabulary checklists, in order to capture the reality of the early 109 linguistic environment and how this is affected by sibling number. Second, we consider the 110 opportunities that overheard speech might present in the infant's linguistic environment. We examine the effect of sibling number on overall amount of input produced in our naturalistic 112 recordings, as well as, crucially, the extent to which parents label objects being attended to 113 by the infant (object presence¹). The analysis of object presence will allow us to gain insight 114 into the kinds of learning opportunities being presented to infants in the early input, based 115 on the previous research showing that object labeling - even when not directed specifically at 116

¹ We've retained the term object presence for continuity with prior work using this variable but note that what this variable captures isn't merely whether the object was present but rather whether it was present when the word for it was said aloud

the target child - can be a valuable source for acquiring linguistic knowledge. Based on work summarized above, we expect that both the language environment and infants' early productive vocabulary will vary as a function of how many older siblings they have.

20 Hypotheses

Synthesising the work above in broad strokes, given prior research showing that early lexical development is more advanced among first-born infants (e.g. Hoff-Ginsberg, 1998), we predict that children with more siblings will have lower productive vocabularies than their peers with fewer siblings. However, we have no *a priori* predictions about how these differences will manifest gradiently (e.g. linear decrease for each additional sibling, a threshold effect where we see a drop after a certain sibship size, etc.).

With regard to the infants' linguistic environment, we hypothesise that infants with 127 more siblings will experience lower prevalence of two aspects of the language input previously 128 shown to support language development: **amount of input** and **amount of object** 129 **presence.** Just as for productive vocabulary size, we do not make a priori predictions 130 regarding the shape of these effects, beyond predicting a decrease with sibling number. 131 Regarding input specifically, following previous studies that show infants with siblings to 132 receive less speech directed at them (Jones & Adamson, 1987; Oshima-Takane & Robbins, 133 2003), we expect to see the same effect in our sample. In terms of object presence (by which, 134 as noted above, we mean word and object co-occurrence, e.g. mother saying "cat" when the 135 child is looking at a cat), we predict a decrease as sibling number increases. This is because, by hypothesis, as caregivers' attention is drawn away from one-to-one interactions with the 137 infant, there is likely less opportunity for contingent talk and joint attention. Prior research suggests links between object presence and early word learning (Bergelson & Aslin, 2017; 139 Cartmill et al., 2013), though to our knowledge this has not been examined in relation to 140 sibling number. 141

142 Methods

We analyze data from the SEEDLingS corpus, a longitudinal set of data incorporating 143 home recordings, parental reports and experimental studies from the ages of 0:6 to 1:6. See 144 Bergelson, Amatuni, Dailey, Koorathota, & Tor (2019) for further details on the full set of 145 home-recorded data and its annotations. The present study draws on the parental report 146 data to index child vocabulary size, and annotations of hour-long home video recordings, 147 taken on a monthly basis during data collection, to index input.² We note at the outset that 148 with such a multidimensional dataset there are always alternative ways of conducting 149 analyses of input and output; due to limited power in our sample, we are unable to consider 150 all potential contributing variables (e.g. the same dataset was analysed in a previous study 151 and found that mothers' work schedules were associated with vocabulary knowledge at 17 152 months (Laing & Bergelson, 2019); we do not analyse that variable further here due to our 153 limited sample size, though we note that in a preliminary analysis number of siblings and 154 maternal work schedule were unrelated; see S4). Our goal here is to make motivated decisions 155 that we clearly describe, provide some alternative analytic choices in the supplementals, and 156 to share the data with readers such that they are free to evaluate alternative approaches. 157

58 Participants

Forty-four families in New York State completed the year-long study. Infants (21 females) were from largely middle-class households; 33 mothers had attained a B.A. degree or higher. Based on parental report, no infants had speech- or hearing-relevant diagnoses; none were low birth weight (all >2,500g); 42 were white, two were from multi-racial backgrounds. All infants heard >75% English on a regular basis and lived in two parent

² We also ran our input analysis using data sub-sampled from day-long audio recordings taken on a different day from the video data reported below; results were consistent with those outlined below for most analyses (see Supplementary Materials, S1).

homes. Two participants were dizygotic twins; we retain one twin in the current sample, considering the other only as a sibling³. Thus our final sample size was 43 infants.

Sibling Details. Sibling number was computed based on parental report in the 166 demographics questionnaires completed at 0;6 (Sibling number range: 0-4). Siblings were on 167 average 4.11 years older than the infants in this study (SD: 4.01 years, R: 0-17 years).⁴ All 168 siblings lived in the household with the infant full time, apart from one infant who had two 169 older half siblings (and no other full siblings) who lived with their other parent part of the 170 time. Both older siblings were present for at least some of the monthly recordings. One 171 family had a foster child live in the home for 2 months of data collection, who is not 172 accounted for in our data; the target infant had one sibling. All siblings were older than or of 173 the same age as the infant in question. 174

Materials

Parental report data. To index each child's language abilities, we draw on data 176 from vocabulary checklists (MacArthur-Bates Communicative Development Inventory, hereafter CDI, Fenson et al. 1994), administered monthly from 0;6 to 1;6, along with a demographics questionnaire; each month's CDI survey came pre-populated with the previous 179 month's answers to save on reduplicated effort. Because the majority of infants did not 180 produce their first word until around 0;11 according to CDI reports (M=10.67, SD=2.23), we 181 use CDI data from 0;10 onwards in our analysis. CDI production data for each month is 182 taken as a measure of the infants' lexical development. CDI data for production has been 183 well validated by prior work, including work in this sample (Frank, Braginsky, Yurovsky, & 184 Marchman, 2021; Moore, Dailey, Garrison, Amatuni, & Bergelson, 2019). Of the intended 13 185

³ Results were consistent when both twins were removed from the dataset, see S2.

⁴ For six infants, siblings' exact birthdates were not provided, and so age difference was estimated by subtracting the infant's age (6 months) from the sibling's age in years, as listed on the questionnaire (e.g. if a sibling was 5 years old, they were classed as being 4.5 years older than the infant).

CDIs collected for each of the 43 infants, 26 were missing across 11 infants (leaving 559 CDIs in total). 4 infants had 4 CDI data-points missing, while the majority (n = 5) had only one missing data-point.

Home-recorded video data. Every month between 0.6 and 1.5, infants were 189 video-recorded for one hour in their home, capturing a naturalistic representation of each 190 infant's day-to-day input. Corresponding to our CDI measures above, here we draw on 191 recordings taken between 0;10 and 1;5. We did not ask families to ensure certain family 192 members were or were not present; our video recordings capture whoever was home at the 193 time families opted to schedule. Here we draw on data from the two caregivers who 194 produced the most words in each recording; in 86% of cases this was the mother, and 10% of 195 cases the father. Fathers produced the second highest number of words in 48% of cases (see 196 S3 for a full breakdown of speakers classed as caregivers in the dataset; note that the two 197 main caregivers might differ for a given infant across sessions). At the child level, the modal 198 caretaker across the 8 videos was the mother for 37 infants, father for 4 infants, and 199 grandmother for the remaining 1 infant. One infant had an equal number of sessions (4 each) 200 with the mother and babysitter as the most talkative caregiver. Infants were a hat with two small Looxcie video cameras attached, one pointed slightly up, and one pointed slightly down; this captured the scene from the infants' perspective. In the event that infants refused to wear the hats, caregivers were the same kind of camera on a headband. Additionally, a 204 camcorder on a tripod was set up in the room where infants and caretakers were interacting 205 to capture a broader view; families were asked to move this camcorder if they changed rooms. 206 The dataset includes 8 videos for each child, one for each month that we analyzed. 207

Object words (i.e. concrete nouns) deemed to be said to, by, or loudly and clearly near
the target child were annotated by trained coders for several properties of interest to the
broader project on noun learning. Here we examine annotations for speaker, i.e. who
produced each noun, and object presence, i.e. whether the noun's referent was present and
attended to by the infant (see "Derived Input Measures" below).

Derived Input measures. Two input measures were derived based on the 213 individual word level annotations of concrete nouns directed to or near the target child in 214 this corpus, each pertaining to an aspect of the input that is established as important in 215 early language learning: **overall household input** (how many concrete nouns does each 216 infant hear? Note that this measure only includes speech produced directly to or close by the 217 target child; see example below) and **object presence** (what proportion of this input is 218 referentially transparent?), detailed below. The original dataset coded for synthesised speech 219 from toys/electronics and speech from speakers on screens or radio; these were excluded here, 220 alongside speech from experimenters (from equipment setup/takedown; 4714 tokens excluded 221 from the video data taken between 10-17 months), leaving 69741 tokens in the input analysis 222 in total. 223

Neither of household input or object presence are, in our view, interpretable as "pure" 224 quality or quantity input measures; we hold that quality and quantity are inextricably linked 225 in general, and specifically we include (by design) only object words that the recordings 226 suggest were possible learning instances for the infants who heard them, wherein quantity 227 and quality are conflated. This included only concrete, imageable nouns that were addressed 228 directly to the child (e.g. "Have you got your toy bear?"), or sufficiently loud and proximal that they were clearly audible to the child (e.g. "Can you pass me the toy bear?", directed at the sibling while mother, infant, and sibling play on the rug). As mentioned above, only speech produced in the infant's immediate surroundings (i.e. speech that would have been 232 clearly heard by the target infant) was coded. 233

Household Input reflects how many nouns infants heard in the recordings from their
two main caregivers (operationalized as the two adults who produced the most nouns in each
recording; see above) and (where relevant) siblings. Input from speakers (adults or children)
other than these two caregivers (and siblings) was relatively rare during video recordings,
accounting for 0.61% of input overall (SD=3.81%), and is excluded from our analysis. This
measure of the early language environment is based on evidence showing strong links

between the amount of speech heard in the early input and later vocabulary size (Anderson, Graham, Prime, Jenkins, & Madigan, 2021). We specifically consider only nouns produced 241 by speakers in the child's environment, directed to or produced clearly near the child, as 242 nouns are what was annotated in the broader SEEDLingS project from which these data are 243 taken; concrete nouns are acquired earlier in development in English and cross-linguistically 244 (Braginsky, Yurovsky, Marchman, & Frank, 2019). As in any sample of naturalistic 245 interaction, the number of nouns correlates highly with the number of words overall 246 (e.g. based on automated analyses of adult word counts vs. manual noun-only annotations, 247 Bulgarelli & Bergelson, 2020). Thus, noun count in the monthly hour of video data serves as 248 our household input proxy. 240

Object Presence was coded as "yes", "no", or "unsure" for each object word annotated 250 in the home recordings, as produced by the two main caregivers detailed above, based on 251 trained annotators' assessment of whether the referent of the word (i.e. the object) was 252 present and attended to or touched by the child or the caregiver. For example, if the 253 caregiver was pointing at a ball while the saying the word ball, this was coded as "yes". If 254 the infant was holding (but not looking at) a bottle while the caregiver said bottle, this 255 would also be coded as "yes". On the other hand, if the caregiver refers to shoes that need to 256 get put away in the other room, that would be coded as "no", as it was not present during 257 object labeling. In the video data, 145 instances (0.28\% on average per infant) of object 258 presence were marked as unsure; these instances were not included in this analysis. 250

260 Data analysis

While we set out to test the hypotheses outlined above, aspects of our analysis were
exploratory in nature. In respect of this, and on the advice of a helpful anonymous reviewer,
we focus on descriptive and confirmatory measures of analysis through data visualization and
effect size reporting alongside significance testing. For each key variable tested, we present
these three avenues for understanding the data, alongside any further follow-up exploratory

266 analyses, where appropriate.

All reported models were generated in R (R Core Team, 2019) using the *lmerTest* 267 package to run linear mixed-effects regression models when needed (Kuznetsova, Brockhoff, 268 & Christensen, 2017). P-values were generated by likelihood ratio tests resulting from nested 269 model comparison. All models include infant as a random effect. Since the raw data were 270 highly skewed, log-transformed data and/or proportions were used for the reported models 271 and model comparisons (1 was added to the raw infant production data of all infants before 272 log-transformation to retain infants with vocabularies of 0); this brought our data closer to 273 normality, though note that the model comparisons run here are not overly sensitive to 274 skewed data. That said, given that all of our variables of interest (CDI score, household 275 input and object presence) did differ significantly from normality by Shapiro tests, we opted 276 to run non-parametric tests (two-sample two-tailed Wilcoxon Tests) on non-transformed 277 data for all post-hoc comparisons, where divergences from normality are more likely to have 278 an outsize effect. Where multiple post-hoc comparisons are run on the same dataset, 279 Bonferroni corrections are applied (e.g. with an adjusted p-value threshold of .025 for 2 between-group comparisons). Unless otherwise specified, all figures display non-transformed 281 data for interpretive ease.

While we have a substantial amount of data for each participant, our limited n means we are under-powered to consider multiple demographic variables simultaneously given the data distribution (e.g. sibling number and sex, see Table 1; as luck would have it both infants with 3 siblings were girls and both with 4 were boys). There were no correlations between sibling number or child word production and maternal age/education/work hours. See S4 for further details.

Results

Our analyses consider infants' total productive vocabulary⁵ alongside our two input
measures – nouns in household input and extent of object presence in the input – as a
function of sibling number.

Vocabulary development was highly variable across the 43 infants, according to the 293 CDI data we had available. By 18 months, 2 infants produced no words (taken from 36 294 available CDIs at this time-point), while mean productive vocabulary size was 60.28 words 295 (SD=78.31, Mdn=30.50). Three infants had substantially larger-than-average (3SDs above 296 the monthly mean) vocabularies at certain time-points in the data; we counted one of these 297 infants as an outlier and remove this child's data from the CDI analysis given that their 298 vocabulary was higher for multiple consecutive months (1;1-1;6). The other two infants had higher vocabularies at 10-11 months only (when variance was quite limited, see Figure 1), and were retained to maximize data inclusion. This left 42 infants (19 females) in the analysis of vocabulary size. Infants had one sibling on average (M=0.86, Mdn=1, SD=1.10). See Table 1. 303

Effect of siblings on infants' productive vocabulary

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We first modeled the effect of siblings on reported productive vocabulary. We explored three possible variations on how to represent the sibling effect: a binary variable (0 vs. >0 siblings), aggregated groups (None vs. One vs. 2+ siblings), and discrete sibling number (0 to 4 siblings), comparing the following nested model structures, where (1) is the baseline model and (2) includes siblings as the variable of interest.

1. vocabulary size (log-transformed) \sim age (months) + (1|subject)

⁵ While in principle we could have just used noun productive vocabulary, in practice noun and total vocabulary is correlated >.95 in this age range; we opted to retain the overall total vocabulary, as lexical class is not a straightforward notion in the early lexicon.

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Table 1
Sibling number by female and male
infants (n=42). One child was an
outlier, and was removed from the
CDI analysis and this table; see
text for details.

n Siblings	Female	Male	Total
0	9	12	21
1	6	6	12
2	2	3	5
3	2	0	2
4	0	2	2
Total	19	23	42

2. vocabulary size (log-transformed) \sim siblings [binary, group or discrete] + age (months) + (1|subject)

In our sample, simply having siblings (i.e. as a binary variable) did not predict CDI productive vocabulary size, while both discrete sibling number and sibling group did. See Table 2.6

⁶ While our sample size and distribution leaves it statistically questionable to consider both sex and sibling number, for completeness we did also run a model that included sex in addition to age and sibling number (our primary variable of interest). Sex did not improve model fit over and above the effect of siblings in any of the three comparisons (ps all >0.54).

Table 2

Output from likelihood ratio tests comparing regression models that predict the effects of sibling number (binary, grouped and discrete variables) on vocabulary size. Month was included in each model as a fixed effect; subject was included as a random effect. R^2 values are included to reflect model goodness-of-fit, though we note that utility and interpretability of this metric for this model type is debated (see

https://bbolker.github.io/mixed models-misc/glmmFAQ.html).

Model	Df	Chisq	p value	R^2
0 vs. >0 siblings	1	2.13	0.14	0.81
Sibling group	2	8.00	0.02	0.81
Sibling number	1	6.08	0.01	0.81

Table 3

Full model output from linear mixed effects regression models comparing language

development over time in relation to sibling group. Age in months was included as a fixed effect; subject was included as a random effect.

Effect	Estimate	Std. Error	df	t value	p
Intercept	-2.69	0.26	156.59	-10.27	< 0.001
SibGroupOne	-0.01	0.30	42.08	-0.05	0.963
SibGroup2+	-0.94	0.33	42.84	-2.81	0.007
month	0.34	0.01	315.13	25.19	< 0.001

Having more siblings was associated with a smaller vocabulary size over the course of early development. This is consistent with previous findings (Hoff-Ginsberg, 1998; Pine, 1995). We find that for each additional sibling, infants were reported to have produced 24% fewer words on average⁷. The 'sibling effect' is thus present in our data.

In terms of our grouped sibling variable (i.e. 0 vs. 1 vs. 2+ siblings), infants with one 320 sibling produced 64 words on average at 18 months, which is, on average, 5 words more than 321 their firstborn peers. Consistent with the model results, infants with two or more siblings 322 produced substantially fewer words at 18 months than those with one or 2+ siblings (based on the raw data: None vs. 2+: 59 vs. 13; One vs. 2+: 64 vs. 13). Post-hoc Wilcoxon Rank Sum tests comparing reported productive vocabulary at 18 months (where there's the widest 325 vocabulary range) revealed significantly larger vocabularies for infants with one sibling 326 compared to those with two or more siblings (W=5, p=.004, CI=[-72.00,-12.00]), but no 327 difference between infants with one sibling and those with no siblings (W=79.50, p=.631, CI=[-34.00,34.00]). See Table 4.

⁷ Average difference in vocabulary size between 0 vs. 1, 1 vs. 2, 2 vs. 3 and 3 vs. 4 siblings at 18 months, as a percentage of firstborn vocabulary size.

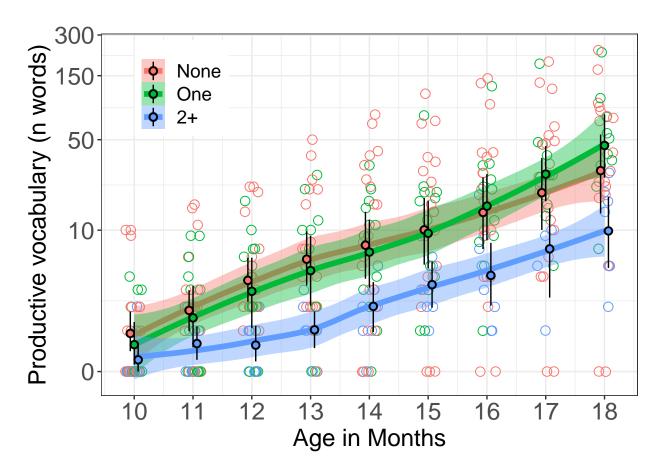


Figure 1. Reported productive vocabulary acquisition (CDI) over time (n=42; one child was an outlier, and was removed from the CDI analysis and this figure; see text for details). Colors denote sibling group; line with colored confidence band reflects local estimator (loess) fit over individual infants' vocabulary at each month. Filled circles indicate mean with bootstrapped CIs computed over each month's data. Open circles (jittered horizontally) show individual infants' vocabulary size at each month. Y-axis utilizes log-transformed vertical spacing for visual clarity.

Effect of siblings on infants' input

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Having established that infants' productive vocabulary varied as a function of sibling 331 number in all but the binary version of the measure (0 vs. >0 siblings), we turn to our input 332 measures to test whether *input* varied by a child's sibling status. For these analyses we 333 report here the group sibling division (0 vs. 1 vs. 2+) as this lets us keep relatively similar 334 Ns across groups, thus making variance more comparable for post-hoc comparisons (the 335 discrete sibling number (0-4) version is reported for completeness in S5; results hold for both 336 input variables). We also now include the child who was a multi-month vocabulary outlier above, given that input and vocabulary are not tested in the same model. One infant of the full sample of 43 infants was an outlier in that they heard substantially more input words and words with object presence than all the other infants in the sample in four of their recording sessions. Given that these sessions were not contiguous, we opted to keep this 341 infant in the analyses reported below, though all results hold when they are removed from 342 our sample (see S6). 343

While we didn't have strong a priori expectations about how overall input or object
presence would vary by age or sex, these were included in initial model comparisons to see if
they improved fit alongside a random effect of infant. Both variables improved fit for the
input model, and only age did for the object presence model. Thus our baseline models
include these sets of control variables, respectively. See Table 5 for final model estimates.

Caregiver Input. We tested overall quantity of input (aggregated across the two main caregivers, as outlined above, and siblings) in our model alongside age, sex and subject, as noted above, and a significant effect was found for the effect of sibling group ($\chi^2(2)$ = 9.09, p=.011, R^2 =0.59). Averaging across infants, those with one sibling heard on average 8% fewer words than those with no siblings in any given hour-long recording, while infants with two or more siblings heard 45% fewer words than those with no siblings.

We then ran post-hoc tests to compare mean amount of input across sibling groups;

these showed a significant difference in average input received between infants with one sibling versus those with two or more siblings (W=11, p=.002, CI=[-120.87,-39.87];

Bonferroni-corrected p-threshold = .025 for all reported Wilcoxon tests) while amount of input did not differ between infants with no siblings and those with one sibling (W=147, p=.723, CI=[-40.50,57.50]).

While we operationalized caregiver input in our models as input speech from the two 361 adults who produced the most words in any given session, in 86.59% of cases this was the 362 mother or father. Considering mothers and fathers specifically, maternal input accounted for 363 75% of object words in the data overall (M=195.56 words, Mdn=162.31, SD=109.02)⁸. 364 Fathers accounted for an average of 18% (M=58.42, Mdn=32.50, SD=64.88), while infants 365 with siblings received around 12% of their input from their brothers and sisters (M=22.97, 366 Mdn=18, SD=18.49). See Figure 2, showing the raw values of input from mothers, fathers 367 and siblings, which are consistent with the group trends reported above. As well as the overall input being greater for firstborns, compared with infants with one or 2+ siblings, note also that the variance is greater for this group, and decreases as sibling number increases. 370 This is shown in the SDs reported in Table 4, and in the data points visualized in Figure 2. 371

Overall, for infants who had siblings, at least one other child was present in 76% of video recordings (n = 133 recordings, SD = 24%). Wilcoxon Rank Sum tests comparing mean monthly input showed no difference between the amount of sibling input received by infants with one sibling compared with those with two or more siblings (W=31, p=.071, CI=[-14.50,2]). Looking at mothers and fathers individually, infants with two or more siblings heard significantly less input from their mothers than those with one sibling (W=5, p<.001, CI=[-124.88,-41.92]), while there was no difference between those with one vs. no siblings (W=125, p = .985, CI=[-48.23,51.50]). Finally, amount of paternal input did not

⁸ One family in our sample had two mothers; rather than artificially assigning one parent to another category, we averaged both mothers' input for this child; we acknowledge that this is an imperfect solution but found it better than the alternatives.

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Table 4

Data summary of our two input measures and reported vocabulary size at 18 months. Input measures represent input from the two adults who produced the most words in any given session, plus siblings.

	No siblings		1 sibling		2+ siblings	
Variable	Mean	SD	Mean	SD	Mean	SD
Productive Vocabulary 18m (CDI)	58.89	60.76	64.10	61.97	13.00	9.49
N Input utterances, 10-17 months	213.33	101.51	196.27	53.96	117.36	26.46
% object presence in input, 10-17 months	68.61	10.93	55.90	9.79	46.38	6.17

differ between groups (one vs. none: W=108, p=.388, CI=[-12.37,57.62]; one vs. 2+: W=21, p=.945, CI=[-74.33,56.54]).

Object presence. On average, 60% of annotated utterances included a referent that was present and attended to by the infant (Mdn=0.61, SD=0.12). See Table 4. Consistent with our hypothesis that infants with more siblings would hear fewer words in referentially transparent conditions (i.e. they would experience lower object presence) than those with fewer siblings, our models reveal a significant effect for sibling group on object presence $(\chi^2(2) = 27.52, p < .001, R^2=0.55)$.

Descriptively, infants with no siblings experienced on average 32% more object presence in their input than those with two or more siblings, and 19% more than those with one sibling. Post-hoc comparisons revealed significant between-group differences: infants with no siblings experienced significantly more object presence than those with one sibling (W=240, p < .001, CI=[0.07,0.20]; Bonferroni-corrected p-threshold = .025). Likewise, infants with one sibling experienced significantly more object presence those with two or more siblings (W=25, p=.025, CI=[-0.18,-0.01]). See Table 5 and Figure 3.

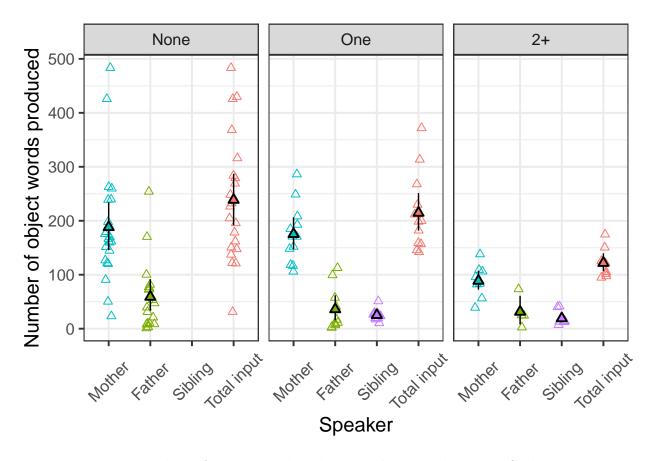


Figure 2. Mean number of words produced by Mothers, Fathers and Siblings, as well as total family input (mean input from mother + father + sibling(s)), across sessions recorded between 10-17 months. Open triangles represent values for individual infants; filled triangles show group means. In the case where the infant had two mothers, mean maternal input is shown. See S7 for a month-by-month visualisation of caregiver input.

Table 5

Full model output from linear mixed effects regression models comparing our two input measures (object words produced in caregiver input and object presence) over time in relation to sibling group. Age in months was included as a fixed effect in both models, sex was included in the caregiver input model only; subject was included as a random effect.

Variable	Effect	Estimate	Std. Error	df	t value	p value
Caregiver input	Intercept	4.88	0.18	185.99	27.46	< 0.001

	${\bf Sib Group One}$	0.01	0.15	43.00	0.05	0.960
	SibGroup2+	-0.49	0.17	43.00	-2.98	0.005
	month	0.03	0.01	301.00	2.95	0.003
	sexM	-0.18	0.13	43.00	-1.41	0.164
Object presence	Intercept	0.57	0.04	321.44	12.73	< 0.001
	SibGroupOne	-0.13	0.03	43.00	-3.81	< 0.001
	SibGroup2+	-0.22	0.04	43.00	-5.90	< 0.001
	month	0.01	0.00	301.00	2.93	0.004

Sibling presence. So far, our analysis takes into account the differences in input 395 based on whether or not the target child has a sibling, but does not directly consider 396 whether sibling presence in the recordings affected these variables. That is, if it is the active 397 presence of the sibling that affects how the caretaker interacts with the target child, then we 398 would expect to see a difference in our input measures when the sibling is present vs. absent. 390 On the other hand, if the very fact of having a sibling changes the way that a caregiver 400 interacts with the infant regardless of whether any sibling is actually present on the scene, 401 then no difference should be observed. While sibling presence in each recording was not 402 coded directly in the dataset, for this exploratory analysis we can get at this with an 403 admittedly imperfect proxy: did the sibling produce nouns in the recording? If yes, we can 404 safely assume they are present; if not we (less safely, but reasonably for initial exploratory 405 purposes) assume they are not. As reported above, by this measure, at least one sibling was present in 76% of recordings for the infants who had a sibling.

Since the presence of a sibling in any given infant's data changed month-on-month (i.e. sometimes the sibling was present and sometimes they were not), and since our measure of sibling presence is imperfect, we opt here to describe the pattern of data without drawing any strong conclusions from statistical models. Descriptively, the presence of a sibling

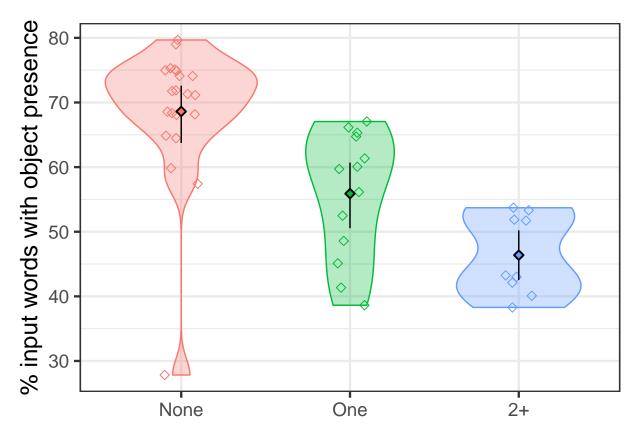


Figure 3. Percentage of input words produced with object presence across sibling groups. Error bars and filled diamonds show 95% CIs and mean proportion of object presence across sibling groups. Open shapes indicate mean proportion of object presence per infant, collapsing across age and jittered horizontally for visual clarity. See S7 for a month-by-month visualisation of object presence.

affected the amount of object presence in the data, but not the amount of input. See Table 6
and Figures 4 and 5. Overall, the presence of a sibling negatively affected object presence,
and this was consistent over time; when a sibling was present, infants in both groups heard
less object presence. This effect was stronger for infants with two or more siblings (though
note that it is unclear from our measure how many siblings were present, and it is possible
that only one sibling was present in the recording), and overall this was true regardless of
whether the sibling was present, or whether the infant was alone with the caregiver. The
picture is less clear for caregiver input, where the presence of a sibling has a more variable

Table 6

Data summary of our two input measures according to presence or absence of siblings during the recording. Input measures represent input from the two adults who produced the most words in any given session, plus siblings.

		1 sibling		2+ siblings	
Variable	Sibling presence	Mean	SD	Mean	SD
N Input utterances, 10-17 months	Sibling not present	139.36	73.23	88.67	32.39
	Sibling present	126.92	38.69	75.01	26.29
% object presence in input, 10-17 months	Sibling not present	70.26	14.39	64.60	12.10
	Sibling present	53.91	8.42	39.25	8.99

effect on the number of object words produced by caregivers (see Figure 4), particularly for infants with only one sibling. This supports our findings above, suggesting that the presence of one additional child does not have any negative effects on the amount of input that caregiver provide. However, input was consistently lower for the group with two or more siblings, and this was true regardless of whether or not a sibling was present; again, this is consistent with the findings reported above for this input measure.

426 Discussion

We investigated the nature of infant language development in relation to number of
children in the household. Previous research found a delay in lexical acquisition for
later-born infants (Fenson et al., 1994; Hoff, 2006), with differences in input across birth
order reported as a root cause. Our results add several new dimensions to this, by testing for
differences across more vs. fewer older siblings, and by looking at input during child-centered
home recordings. Infants with more siblings were reported to say fewer words by 18 months,
heard fewer nouns from their parents, and were less likely to be attending to an object when

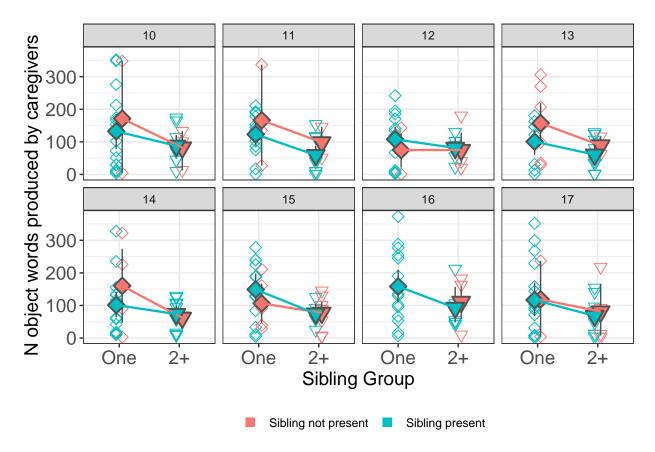


Figure 4. Difference in number of input words in infants' input according to whether or not a sibling or siblings were present during the time of recording, for each month of data. Infants with no siblings were not included in the plots for visual ease. Open diamonds represent individual infants in the data; filled diamonds represent means and 95%CIs, colours represent presence or absence of siblings during the recording session.

434 hearing its label.

Importantly, and in contrast with some previous research (Hoff-Ginsberg, 1998;
Oshima-Takane & Robbins, 2003), infants with one sibling showed no decrement in lexical
production and minimal reduction in input in comparison to first-born infants. That is, our
results suggest that simply having a sibling does not contribute to input or vocabulary
differences across children (as measured here), while having more than one sibling seems to
do so. Indeed, infants with zero and one sibling had similar results for productive vocabulary,
and parental noun input overall (though not object presence). Moreover, parental input was

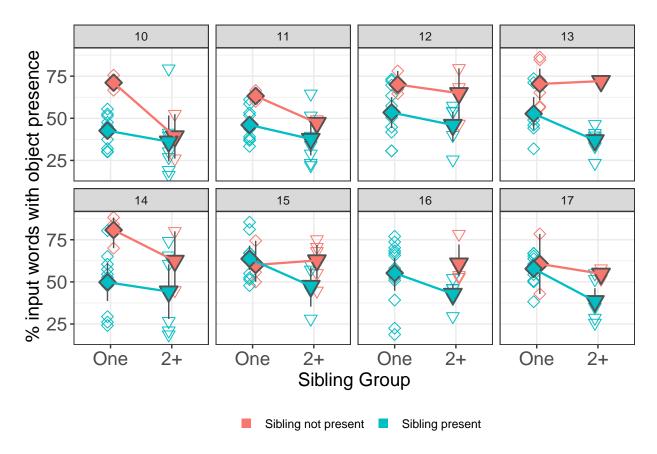


Figure 5. Difference in % of object presence in infants' input according to whether or not a sibling or siblings were present during the time of recording, for each month of data. Infants with no siblings were not included in the plots for visual ease. Open shapes represent individual infants in the data; filled shapes represent means and 95%CIs, colours represent presence or absence of siblings during the recording session.

not affected by the presence or absence of the sibling in the room. In contrast, infants with two or more siblings said fewer words, and also heard fewer input words overall.

With regards to object presence, having siblings made it increasingly less likely to hear
an object label when attending to it, and this effect increased with sibship size (i.e. children
with more siblings heard input containing a lower proportion of object presence). Unlike for
total parental noun input (which was reduced for 2+ siblings but not modulated based on
whether siblings were present in the recordings), reduced object presence for children with
more siblings was particularly notable in recordings with siblings present.

The sibling effect. When we considered the effect of sibling status – that is, 450 whether or not infants had any siblings, disregarding specific sibling number – our findings 451 showed that having siblings made no difference to infants' lexical production capacities. This 452 contrasts with Hoff-Ginsberg (1998), who found that, by 18 months, laterborns exhibit lower 453 language skills. However, Oshima-Takane and colleagues (1996) found no overall differences 454 between first- and second-born children across a range of language measures taken at 21 455 months. Our results suggest that considering sibling quantity may be a more sensitive way to 456 reveal demographic effects than their (coarser-grained) first- vs. later-born status. We find 457 that the more older siblings a child had, the lower their reported productive vocabulary at 18 458 months. This adds to findings from Fenson and colleagues (1994), who found a weak but 459 significant negative correlation between birth order and production of both words and 460 gestures. Controlling for age, our model showed that infants with 2 or more siblings produced 46 fewer words than the average 59 words produced by firstborns in our data by 18 months. 462

While infants with more siblings heard less input speech overall, having one sibling did 463 not significantly reduce the number of nouns in an infant's input. This is in direct contrast 464 with reports from the literature; Hoff (2006) states that "when a sibling is present, each child 465 receives less speech directed solely at...her because mothers produce the same amount of 466 speech whether interacting with one or two children" (p.67, italics added). While this does 467 not appear to be the case in the present dataset, it may be due to the circumstances of the 468 home-recorded data: while siblings were present in many of the recordings (76% of 460 recordings in which the target child had a sibling), given the focus of the data collection, 470 parents may have had a tendency to direct their attention - and consequently their linguistic 471 input - more towards the target child; our samples also differed in other ways 472 (e.g. sociocultural context) that may have influenced the results as well. Alternatively, our results may diverge from those of Hoff (2006) due to the nature of our input measure, which only took nouns into account. That said, we find this alternative explanation unlikely given 475 work by Bulgarelli and Bergelson (2020) showing that nouns are a reliable proxy for overall

input in this dataset, suggesting that this measure provides an appropriate representation of overall input directed at the target child.

In contrast to the other results, our analysis of object presence showed a more linear 479 'sibling effect'. In this case, even having one sibling led to fewer word-object pairs presented in the input. This was true regardless of whether or not other siblings were present, but 481 object presence was further negatively affected by the presence of a sibling in the room. 482 Presence of a labeled object with congruent input speech has been found to support early 483 word learning across several studies. For instance, Bergelson and Aslin (2017) combined 484 analysis of this home-recorded data at six months with an experimental study to show that 485 word-object presence in naturalistic caregiver input correlated with comprehension of nouns 486 (tested using eye-tracking). Relatedly, Gogate and colleagues (2000) propose that contingent 487 word production supports the learning of novel word-object combinations, with "multimodal 488 motherese" - whereby a target object word is produced in movement or touch-based 480 synchrony with its referent - supporting word learning. More broadly, lower rates of 490 referential transparency for common non-nouns like hi and uh-oh have been proposed to 491 potentially explain why these words are learned later than common concrete nouns 492 (Bergelson & Swingley, 2013). While the present results on object presence don't speak 493 directly to word learning, they do suggest that this potentially helpful learning support is less available for children with more siblings.

Siblingese as a learning opportunity? We also found that infants with siblings
did not hear much speech from their older brothers and sisters. Similar findings are reported
in a lab-based interaction study by Oshima-Takane and Robbins (2003), who found that
older siblings rarely talked directly to the target child; instead, most input from siblings was
overheard speech from sibling-mother interactions. One possibility raised by these results is
that perhaps parents are able to compensate or provide relatively similar input and learning
support for one or two children, but once children outnumber parents, this balancing act of
attention, care, and time becomes unwieldy. While the current sample is relatively limited

and homogeneous in the family structures and demographics it includes, future work could fruitfully investigate this possibility by considering whether (controlling for other potential contributors like SES, Hoff-Ginsberg, 1998) the presence of more caregivers (whether parents, relatives, or other adults) helps foster language development.

Alternatively, second-borns might 'even out' with children with no siblings due to a 508 trade-off between direct attention from the caregiver and the possibility of more 509 sophisticated social-communicative interactions. For these infants there is still ample 510 opportunity to engage with the mother in one-to-one interactions, allowing a higher share of 511 her attention than is available to third- or later-borns. Furthermore, triadic interactions can 512 benefit the development of a number of linguistic and communication skills (Barton & 513 Tomasello, 1991; Dunn & Shatz, 1989). Second-borns may also benefit from overheard 514 speech in their input, supporting the acquisition of nouns and even more complex lexical 515 categories (Floor & Akhtar, 2006; Oshima-Takane et al., 1996). For infants with one sibling, 516 the benefits of observing/overhearing interactions between sibling and caregiver, as well as 517 the possibility for partaking in such interactions, may outweigh the decrease in some aspects 518 of the input (i.e., in our data, only observed in object presence). Having more than one 519 sibling may throw this off-balance, such that the possibilities for both supportive one-to-one input and more sophisticated interactions are simultaneously diminished. 521

Importantly, the present results make no claims about eventual outcomes for these children: generally speaking, regardless of sibling number, all typically-developing infants reach full and fluent language use. Indeed, some research suggests that sibling effects, while they may be clear in early development, are not always sustained into childhood; e.g. twins are known to experience a delay in language development into the third year, but are quick to catch up thereafter (Dales, 1969; Tomasello, Mannle, & Kruger, 1986). This demonstrates the cognitive adaptability of early development, which brings about the acquisition of language across varying and allegedly 'imperfect' learning environments. Infants' capacity to develop linguistic skills from the resources that are available to them – whether that is

infant-directed object labels or overheard abstract concepts – highlights the dynamic and adaptable nature of early cognitive development, and a system that is sufficiently robust to bring about the same outcome across populations.

Of course, the 'success' of early language development is defined by Limitations. 534 how success is measured. Here we chose word production as our measure of linguistic 535 capability; we did not consider other equally valid measures such as language comprehension 536 or early social-interaction skills. Similarly, our input measures focused on nouns; other 537 lexical classes may reveal different effects, though they are generally far sparser in 538 production until toddlerhood. Our analysis of vocabulary relied on parental report data; this 539 method could have biased our first-born sample towards more accurate or larger vocabulary 540 reports owing to their parents having more time and attention to spend observing their 541 vocabulary development (see Kartushina et al., 2022 for a discussion of this possibility in 542 light of the COVID-19 pandemic, though note the present data were collected in 2014-16). 543 In the supplementary materials, we provide validation data for the CDI relative to children's own productions by running correlation tests between reported (CDI) vocabulary and the number of word types produced by each infant in the audio and video data. In short, for the 0 and 1 sibling groups there is a strong and significant positive correlation between the CDI at 18 months and child word types; for 2+ siblings this correlation is weaker and does not reach significance, though this is likely due to the small n in this group relative to the others 549 (due to missing CDIs at 18 months, this includes data from n = 18/21, 11/12, and 7/9550 children with 0, 1, and 2+ siblings, respectively); see S8. 551

There is also some imbalance in group sizes across our data; our sample was not pre-selected for sibling number, and so group sizes are unmatched across the analysis. Including a larger number of infants with 2+ siblings may have revealed a different pattern of results. We might also expect that age of older siblings would affect the nature of the early linguistic environment, given that larger age difference is found to be a predictor of lower vocabulary size in the current literature (Gurgand et al., 2022; Havron et al., 2022);

our sample did not allow us to link sibling age to number of words produced by that sibling, 558 but future work may wish to take this into account. Finally, more work across wider and 559 larger populations is necessary to unpack the generalizability of the present results. Our 560 sample is reflective of average household sizes in middle-class families across North America 561 and Western Europe (Office for National Statistics, 2018; United States Census Bureau, 562 2010), but it is not unusual in some communities and parts of the world for households to 563 include between three and six children on average (Institute for Family Studies & Wheatley 564 Institution, 2019). Adding to this, it is also necessary to consider cross-cultural differences in 565 the way children are addressed by their parents, other caretakers, and other children (Bunce 566 et al., 2020; Casillas, Brown, & Levinson, 2019; Shneidman & Goldin-Meadow, 2012). For 567 instance, Bunce and colleagues (2020) find relatively similar rates of target child directed 568 speech across US, Canadian, Argentinian, UK, Papuan and Mayan samples, some differences in who the input comes from, and large effects of number of talkers present. These results 570 suggest that caution is advisable before generalizing the current results to any other 571 socio-cultural contexts, but also pose exciting open questions regarding what variability in 572 experiences do – or don't – change about early language interaction and development. 573

574 Conclusion

Our results with English-learning infants in the US support prior findings from the 575 literature showing that later-born infants have slower lexical acquisition than their first-born 576 peers. However, we highlight an important difference from previous findings, namely that in 577 the present sample, second-born infants show no such effect, while infants with more than two siblings have significantly smaller productive vocabularies at age 18 months. We also 579 identified similar group differences in overall noun input and object presence. While we did not test these corresponding vocabulary and input measures directly, our results suggest that 581 having more siblings affects a child's early language environment, which in turn may lead to 582 slower vocabulary growth in the first 18 months of life. We look forward to future studies 583

considering the granularity of more versus fewer siblings, and how this relates to language

 $_{585}$ abilities over the course of development.

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