1

2 Abstract

- ³ Prior research suggests that across a wide range of cognitive, educational, and health-based
- 4 measures, first-born children outperform their later-born peers. Expanding on this
- 5 literature using naturalistic home-recorded data and parental vocabulary report, we find
- 6 that early language outcomes vary by number of siblings in a sample of 43 English-learning
- U.S. children from mid-to-high socioeconomic status homes. More specifically, we find that
- 8 children in our sample with two or more but not one older siblings had smaller
- 9 productive vocabularies at 18 months, and heard less input from caregivers across several
- measures than their peers with less than two siblings. We discuss implications regarding
- what infants experience and learn across a range of family sizes in infancy.
- 12 Keywords: Siblings, Lexical Development, Input Effects, Language Acquisition

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Analysing the effect of sibling number on input and output in the first 18 months

A common simplifying assumption in research on language development is that there 14 is a theoretical "optimum" environment for early language, whereby the input is tailored to 15 a single infant's needs, changing over time as language capacity grows (e.g. Soderstrom, 16 2007; Stern, Spieker, Barnett, & MacKain, 1983). However, for many infants and for many 17 reasons, language acquisition occurs across diverse social contexts that can influence the learning environment, including the presence of older siblings in the home (Fenson et al., 1994). According to the United States Census Bureau (2010), around one third of children are born into households with at least one other infant present, and one in every five infants is acquiring language in a household shared with two or more other children. Similar statistics are reported for British infants (Office for National Statistics, 2018), where the average household has 1.75 children, and 15% of households have three children or more. More globally, in most parts of the world, few children grow up without siblings 25 (United-Nations, 2017). In this paper, we consider the role of siblings in the early language 26 environment of English-learning infants. We use naturalistic home-recorded data to 27 measure input in earlier- and later-born infants, alongside their productive vocabulary over the first 18 months of life. 29

Prior research suggests that infants born to households with older children may be
slower to learn language. Fenson and colleagues (1994) found that by 30 months of age,
children with older siblings performed worse than those with no siblings across
parent-reported measures of productive vocabulary, use of word combinations, and mean
length of utterance. This 'sibling effect' may be the result of differences in input between
first- and later-born children: some research finds that infants with older siblings hear less
speech aimed specifically at them, and what they do hear is understood to be linguistically
less supportive of early language development (Hoff-Ginsberg, 1998; Oshima-Takane &
Robbins, 2003). In contrast, some studies have noted linguistic advantages for later-borns,

- who may have stronger social-communicative skills (Hoff, 2006), better understanding of pronouns (Oshima-Takane, Goodz, & Derevensky, 1996), and better conversational abilities (Dunn & Shatz, 1989). Overall, while the particulars differ across studies, prior work suggests that the presence of siblings in the home leads to differences in infants' early linguistic experiences and skills, though the direction of these differences varies depending on what aspects of language are being measured.
- Numerous studies have attempted to better understand how siblings affect the 45 language development trajectory, with comparisons of language acquisition across first- and later-borns. Here again, findings are mixed, but overall two general conclusions can be drawn. First, analyses consistently show that infants with older siblings generally have slower vocabulary development (Berglund, Eriksson, & Westerlund, 2005; Fenson et al., 1994; Pine, 1995; Zambrana, Ystrom, & Pons, 2012), and this effect increases with number of older siblings (Gurgand et al., 2022; Karwath, Relikowski, & Schmitt, 2014; Peyre et al., 51 2016). Furthermore, this finding is consistent across cultures (e.g. European French 52 (Gurgand et al., 2022; Havron et al., 2019); Singaporean (Havron et al., 2022); Kenyan (Jakiela, Ozier, Fernald, & Knauer, 2020); and German (Karwath et al., 2014)). However, this finding is not as clear-cut as has been previously assumed: Hoff-Ginsberg (1998) shows first-borns to have better lexical and syntactic skills up until 2;5, but later-born infants had better conversational abilities during the same time-period. Recent studies have also identified effects for age gap between the target child and their siblings (whereby larger age gaps correlate with lower vocabulary scores, Gurgand et al., 2022; Havron et al., 2022) and sibling sex (whereby older brothers have a negative effect on vocabulary outcomes, but not older sisters, Havron et al., 2019; Jakiela et al., 2020) though neither of these effects are found consistently across datasets; Havron et al. (2022) and Gurgand et al. (2022) find no effect for sibling sex, whereas Havron et al. (2019) find no effect for age gap. Some of these differences across studies may relate to insufficient power to detect relatively small effects, perhaps leading to under- or over-estimation of effect sizes, or simultaneous contributing

66 factors that are difficult to disentangle.

The second general finding pertains to sibling-related differences in the early 67 linguistic environment: infants with no siblings receive more input overall, and this more closely reflects what is typically considered to be 'high quality' input in the extant literature (i.e. more input in an infant-directed speech style (Ramírez-Esparza, García-Sierra, & Kuhl, 2014); longer utterance length (Barnes, Gutfreund, Satterly, & 71 Wells, 1983); higher lexical diversity (Rowe & Snow, 2020)). Indeed, the very presence of a sibling in the linguistic environment changes the way language is used. When siblings are present (i.e. triadic interactions), mothers' input has been found to be more focused on regulating behaviour, as opposed to the language-focused speech that is common in dyadic 75 contexts (Oshima-Takane & Robbins, 2003). Reports show that the mean length of utterance is longer in the input of first-born infants (Hoff-Ginsberg, 1998; but see also 77 Oshima-Takane & Robbins, 2003 for a comparison of dyadic and triadic contexts), who also hear more questions directed at them than later-borns. Both Jones and Adamson (1987) and Oshima-Takane and Robbins (2003) report no difference between the overall number of word types produced by mothers in dyadic and triadic settings, but the 81 proportion of speech directed at the target infant is drastically reduced when input is shared with siblings.

As Hoff (2006) explains, infants with siblings have less experience of speech directed at them, but they do have an advantage over their first-born peers in that they are subject to more overheard speech. This may be an important source of input for infants with one or more older siblings. Akhtar, Jipson and Callanan (2001) show that, by age 2;6, children can learn both novel object labels and novel verbs through overhearing. Slightly younger children (aged 1;11-2;2) were also able to learn the novel object labels, but not verb labels. Two-year-old infants can even learn novel object labels while doing activities that distract them from the language input, and when the novel words are produced non-saliently (Akhtar, 2005). This suggests that, while the learning environment for later-borns might

differ from that of first-born infants, there may be ample opportunity for them to learn from the speech that surrounds them; namely overheard speech directed at their older sibling(s). Evidence is mainly drawn from work testing infants aged 2 and above (e.g. Akhtar, 2005; Fitch, Liberman, Luyster, & Arunachalam, 2020; Foushee, Srinivasan, & Xu, 2021), and generally relies on experimental work rather than observations of the home environment. However, Floor and Akhtar (2006) tested younger infants to find that the capacity to learn from overheard speech is available from as early as 16 months, at least in an experimental setting.

There thus may be a trade-off, even in early development, between highly supportive 101 one-to-one input from a caregiver (cf. Ramírez-Esparza et al., 2014) and the potential 102 benefits drawn from communicating with (or overhearing communication with) a sibling. 103 In the present study, we test the extent to which having more versus fewer siblings in the home environment may affect the linguistic environment in ways that could lead to 105 differences in vocabulary development over the course of the first 18 months of life. In 106 analyzing infants' growing productive vocabulary and linguistic environment in relation to 107 the presence of older siblings in their household, the present work expands on the extant 108 literature in two key ways. First, much of the existing literature identifying links between 109 sibling number and vocabulary outcomes draws on large-scale questionnaire data, rather 110 than naturalistic day-to-day interactions in the home. In contrast, we analyze an existing 111 corpus of home recordings in concert with vocabulary checklists, in order to capture the 112 reality of the early linguistic environment and how this is affected by sibling number. 113 Second, we consider the opportunities that overheard speech might present in the infant's 114 linguistic environment. We examine the effect of sibling number on overall amount of input 115 produced in our naturalistic recordings, as well as, crucially, the extent to which parents 116 label objects being attended to by the infant (object presence¹). The analysis of object 117

¹ We've retained the term object presence for continuity with prior work using this variable but note that what this variable captures isn't merely whether the object was present but rather whether it was present

presence will allow us to gain insight into the kinds of learning opportunities being
presented to infants in the early input, based on the previous research showing that object
labeling - even when not directed specifically at the target child - can be a valuable source
for acquiring linguistic knowledge. Based on work summarized above, we expect that both
the language environment and infants' early productive vocabulary will vary as a function
of how many older siblings they have.

124 Hypotheses

Synthesising the work above in broad strokes, given prior research showing that early lexical development is more advanced among first-born infants (e.g. Hoff-Ginsberg, 1998), we predict that children with more siblings will have lower productive vocabularies than their peers with fewer siblings. However, we have no *a priori* predictions about how these differences will manifest gradiently (e.g. linear decrease for each additional sibling, a threshold effect where we see a drop after a certain sibship size, etc.).

With regard to the infants' linguistic environment, we hypothesise that infants with 131 more siblings will experience lower prevalence of two aspects of the language input 132 previously shown to support language development: amount of input and amount of 133 **object presence.** Just as for productive vocabulary size, we do not make a priori 134 predictions regarding the shape of these effects, beyond predicting a decrease with sibling 135 number. Regarding input specifically, following previous studies that show infants with 136 siblings to receive less speech directed at them (Jones & Adamson, 1987; Oshima-Takane & 137 Robbins, 2003), we expect to see the same effect in our sample. In terms of object presence 138 (by which, as noted above, we mean word and object co-occurrence, e.g. mother saying 139 "cat" when the child is looking at a cat), we predict a decrease as sibling number increases. This is because, by hypothesis, as caregivers' attention is drawn away from one-to-one 141

when the word for it was said aloud

interactions with the infant, there is likely less opportunity for contingent talk and joint attention. Prior research suggests links between object presence and early word learning (Bergelson & Aslin, 2017; Cartmill et al., 2013), though to our knowledge this has not been examined in relation to sibling number.

146 Methods

We analyze data from the SEEDLingS corpus, a longitudinal set of data 147 incorporating home recordings, parental reports and experimental studies from the ages of 148 0:6 to 1:6. See Bergelson, Amatuni, Dailey, Koorathota, & Tor (2019) for further details on 149 the full set of home-recorded data and its annotations. The present study draws on the 150 parental report data to index child vocabulary size, and annotations of hour-long home 151 video recordings, taken on a monthly basis during data collection, to index input.² We note 152 at the outset that with such a multidimensional dataset there are always alternative ways 153 of conducting analyses of input and output; due to limited power in our sample, we are 154 unable to consider all potential contributing variables (e.g. the same dataset was analysed 155 in a previous study and found that mothers' work schedules were associated with 156 vocabulary knowledge at 17 months (Laing & Bergelson, 2019); we do not analyse that 157 variable further here due to our limited sample size, though we note that in a preliminary 158 analysis number of siblings and maternal work schedule were unrelated; see S4). Our goal 159 here is to make motivated decisions that we clearly describe, provide some alternative analytic choices in the supplementals, and to share the data with readers such that they are free to evaluate alternative approaches.

² We also ran our input analysis using data sub-sampled from day-long audio recordings taken on a different day from the video data reported below; results were consistent with those outlined below for most analyses (see Supplementary Materials, S1).

Participants

Forty-four families in New York State completed the year-long study. Infants (21 females) were from largely middle-class households; 33 mothers had attained a B.A. degree or higher. Based on parental report, no infants had speech- or hearing-relevant diagnoses; none were low birth weight (all >2,500g); 42 were white, two were from multi-racial backgrounds. All infants heard >75% English on a regular basis and lived in two parent homes. Two participants were dizygotic twins; we retain one twin in the current sample, considering the other only as a sibling³. Thus our final sample size was 43 infants.

Sibling Details. Sibling number was computed based on parental report in the 171 demographics questionnaires completed at 0;6 (Sibling number range: 0-4). Siblings were 172 on average 4.11 years older than the infants in this study (SD: 4.01 years, R: 0-17 years).⁴ 173 All siblings lived in the household with the infant full time, apart from one infant who had 174 two older half siblings (and no other full siblings) who lived with their other parent part of 175 the time. Both older siblings were present for at least some of the monthly recordings. One 176 family had a foster child live in the home for 2 months of data collection, who is not 177 accounted for in our data; the target infant had one sibling. All siblings were older than or 178 of the same age as the infant in question. 179

180 Materials

Parental report data. To index each child's language abilities, we draw on data from vocabulary checklists (MacArthur-Bates Communicative Development Inventory, hereafter CDI, Fenson et al. 1994), administered monthly from 0;6 to 1;6, along with a

 $^{^3}$ Results were consistent when both twins were removed from the dataset, see S2.

⁴ For six infants, siblings' exact birthdates were not provided, and so age difference was estimated by subtracting the infant's age (6 months) from the sibling's age in years, as listed on the questionnaire (e.g. if a sibling was 5 years old, they were classed as being 4.5 years older than the infant).

demographics questionnaire; each month's CDI survey came pre-populated with the 184 previous month's answers to save on reduplicated effort. Because the majority of infants 185 did not produce their first word until around 0:11 according to CDI reports (M=10.67, 186 SD=2.23), we use CDI data from 0:10 onwards in our analysis. CDI production data for 187 each month is taken as a measure of the infants' lexical development. CDI data for 188 production has been well validated by prior work, including work in this sample (Frank, 189 Braginsky, Yurovsky, & Marchman, 2021; Moore, Dailey, Garrison, Amatuni, & Bergelson, 190 2019). Of the intended 13 CDIs collected for each of the 43 infants, 26 were missing across 191 11 infants (leaving 559 CDIs in total). 4 infants had 4 CDI data-points missing, while the 192 majority (n = 5) had only one missing data-point. 193

Home-recorded video data. Every month between 0.6 and 1.5, infants were 194 video-recorded for one hour in their home, capturing a naturalistic representation of each 195 infant's day-to-day input. Corresponding to our CDI measures above, here we draw on 196 recordings taken between 0:10 and 1:5. We did not ask families to ensure certain family 197 members were or were not present; our video recordings capture whoever was home at the 198 time families opted to schedule. Here we draw on data from the two caregivers who 199 produced the most words in each recording; in 86% of cases this was the mother, and 10% 200 of cases the father. Fathers produced the second highest number of words in 48% of cases 201 (see S3 for a full breakdown of speakers classed as caregivers in the dataset; note that the 202 two main caregivers might differ for a given infant across sessions). At the child level, the 203 modal caretaker across the 8 videos was the mother for 37 infants, father for 4 infants, and 204 grandmother for the remaining 1 infant. One infant had an equal number of sessions (4 each) with the mother and babysitter as the most talkative caregiver. Infants were a hat with two small Looxcie video cameras attached, one pointed slightly up, and one pointed slightly down; this captured the scene from the infants' perspective. In the event that 208 infants refused to wear the hats, caregivers were the same kind of camera on a headband. 209 Additionally, a camcorder on a tripod was set up in the room where infants and caretakers 210

were interacting to capture a broader view; families were asked to move this camcorder if
they changed rooms. The dataset includes 8 videos for each child, one for each month that
we analyzed.

Object words (i.e. concrete nouns) deemed to be said to, by, or loudly and clearly
near the target child were annotated by trained coders for several properties of interest to
the broader project on noun learning. Here we examine annotations for speaker, i.e. who
produced each noun, and object presence, i.e. whether the noun's referent was present and
attended to by the infant (see "Derived Input Measures" below).

Derived Input measures. Two input measures were derived based on the 219 individual word level annotations of concrete nouns directed to or near the target child in 220 this corpus, each pertaining to an aspect of the input that is established as important in 221 early language learning: **overall household input** (how many concrete nouns does each 222 infant hear? Note that this measure only includes speech produced directly to or close by 223 the target child; see example below) and **object presence** (what proportion of this input 224 is referentially transparent?), detailed below. The original dataset coded for synthesised 225 speech from toys/electronics and speech from speakers on screens or radio; these were 226 excluded here, alongside speech from experimenters (from equipment setup/takedown; 4714 227 tokens excluded from the video data taken between 10-17 months), leaving 69741 tokens in 228 the input analysis in total. 229

Neither of household input or object presence are, in our view, interpretable as

"pure" quality or quantity input measures; we hold that quality and quantity are

inextricably linked in general, and specifically we include (by design) only object words

that the recordings suggest were possible learning instances for the infants who heard

them, wherein quantity and quality are conflated. This included only concrete, imageable

nouns that were addressed directly to the child (e.g. "Have you got your toy bear?"), or

sufficiently loud and proximal that they were clearly audible to the child (e.g. "Can you

pass me the toy bear?", directed at the sibling while mother, infant, and sibling play on the

rug). As mentioned above, only speech produced in the infant's immediate surroundings
(i.e. speech that would have been clearly heard by the target infant) was coded.

Household Input reflects how many nouns infants heard in the recordings from their 240 two main caregivers (operationalized as the two adults who produced the most nouns in 241 each recording; see above) and (where relevant) siblings. Input from speakers (adults or 242 children) other than these two caregivers (and siblings) was relatively rare during video 243 recordings, accounting for 0.61% of input overall (SD=3.81%), and is excluded from our 244 analysis. This measure of the early language environment is based on evidence showing 245 strong links between the amount of speech heard in the early input and later vocabulary 246 size (Anderson, Graham, Prime, Jenkins, & Madigan, 2021). We specifically consider only 247 nouns produced by speakers in the child's environment, directed to or produced clearly 248 near the child, as nouns are what was annotated in the broader SEEDLingS project from 249 which these data are taken; concrete nouns are acquired earlier in development in English 250 and cross-linguistically (Braginsky, Yurovsky, Marchman, & Frank, 2019). As in any 251 sample of naturalistic interaction, the number of nouns correlates highly with the number 252 of words overall (e.g. based on automated analyses of adult word counts vs. manual 253 noun-only annotations, Bulgarelli & Bergelson, 2020). Thus, noun count in the monthly hour of video data serves as our household input proxy. 255

Object Presence was coded as "yes", "no", or "unsure" for each object word
annotated in the home recordings, as produced by the two main caregivers detailed above,
based on trained annotators' assessment of whether the referent of the word (i.e. the
object) was present and attended to or touched by the child or the caregiver. For example,
if the caregiver was pointing at a ball while the saying the word ball, this was coded as
"yes". If the infant was holding (but not looking at) a bottle while the caregiver said bottle,
this would also be coded as "yes". On the other hand, if the caregiver refers to shoes that
need to get put away in the other room, that would be coded as "no", as it was not present
during object labeling. In the video data, 145 instances (0.28% on average per infant) of

object presence were marked as unsure; these instances were not included in this analysis.

66 Data analysis

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While we set out to test the hypotheses outlined above, aspects of our analysis were
exploratory in nature. In respect of this, and on the advice of a helpful anonymous
reviewer, we focus on descriptive and confirmatory measures of analysis through data
visualization and effect size reporting alongside significance testing. For each key variable
tested, we present these three avenues for understanding the data, alongside any further
follow-up exploratory analyses, where appropriate.

All reported models were generated in R (R Core Team, 2019) using the *lmerTest* 273 package to run linear mixed-effects regression models when needed (Kuznetsova, Brockhoff, 274 & Christensen, 2017). P-values were generated by likelihood ratio tests resulting from 275 nested model comparison. All models include infant as a random effect. Since the raw data 276 were highly skewed, log-transformed data and/or proportions were used for the reported 277 models and model comparisons (1 was added to the raw infant production data of all 278 infants before log-transformation to retain infants with vocabularies of 0); this brought our 279 data closer to normality, though note that the model comparisons run here are not overly 280 sensitive to skewed data. That said, given that all of our variables of interest (CDI score, 281 household input and object presence) did differ significantly from normality by Shapiro 282 tests, we opted to run non-parametric tests (two-sample two-tailed Wilcoxon Tests) on 283 non-transformed data for all post-hoc comparisons, where divergences from normality are 284 more likely to have an outsize effect. Where multiple post-hoc comparisons are run on the 285 same dataset, Bonferroni corrections are applied (e.g. with an adjusted p-value threshold of 286 .025 for 2 between-group comparisons). Unless otherwise specified, all figures display non-transformed data for interpretive ease.

While we have a substantial amount of data for each participant, our limited n means

we are under-powered to consider multiple demographic variables simultaneously given the
data distribution (e.g. sibling number and sex, see Table 1; as luck would have it both
infants with 3 siblings were girls and both with 4 were boys). There were no correlations
between sibling number or child word production and maternal age/education/work hours.
See S4 for further details.

295 Results

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Our analyses consider infants' total productive vocabulary⁵ alongside our two input measures – nouns in household input and extent of object presence in the input – as a function of sibling number.

Vocabulary development was highly variable across the 43 infants, according to the 299 CDI data we had available. By 18 months, 2 infants produced no words (taken from 36 300 available CDIs at this time-point), while mean productive vocabulary size was 60.28 words 301 (SD=78.31, Mdn=30.50). Three infants had substantially larger-than-average (3SDs above 302 the monthly mean) vocabularies at certain time-points in the data; we counted one of these 303 infants as an outlier and remove this child's data from the CDI analysis given that their 304 vocabulary was higher for multiple consecutive months (1;1-1;6). The other two infants had 305 higher vocabularies at 10-11 months only (when variance was quite limited, see Figure 1), 306 and were retained to maximize data inclusion. This left 42 infants (19 females) in the 307 analysis of vocabulary size. Infants had one sibling on average (M=0.86, Mdn=1, SD=1.10). See Table 1.

⁵ While in principle we could have just used noun productive vocabulary, in practice noun and total vocabulary is correlated >.95 in this age range; we opted to retain the overall total vocabulary, as lexical class is not a straightforward notion in the early lexicon.

Table 1
Sibling number by female and
male infants (n=42). One child
was an outlier, and was removed
from the CDI analysis and this
table; see text for details.

n Siblings	Female	Male	Total
0	9	12	21
1	6	6	12
2	2	3	5
3	2	0	2
4	0	2	2
Total	19	23	42

Effect of siblings on infants' productive vocabulary

We first modeled the effect of siblings on reported productive vocabulary. We
explored three possible variations on how to represent the sibling effect: a binary variable
(0 vs. >0 siblings), aggregated groups (None vs. One vs. 2+ siblings), and discrete sibling
number (0 to 4 siblings), comparing the following nested model structures, where (1) is the
baseline model and (2) includes siblings as the variable of interest.

- 1. vocabulary size (log-transformed) \sim age (months) + (1|subject)
- 2. vocabulary size (log-transformed) \sim siblings [binary, group or discrete] + age (months) + (1|subject)

In our sample, simply having siblings (i.e. as a binary variable) did not predict CDI productive vocabulary size, while both discrete sibling number and sibling group did. See

Table 2.6

Table 2

Output from likelihood ratio tests comparing regression models that predict the effects of sibling number (binary, grouped and discrete variables) on vocabulary size. Month was included in each model as a fixed effect; subject was included as a random effect. R^2 values are included to reflect model goodness-of-fit, though we note that utility and interpretability of this metric for this model type is debated (see

https://bbolker.github.io/mixed models-misc/glmmFAQ.html).

Model	Df	Chisq	p value	R^2
0 vs. >0 siblings	1	2.13	0.14	0.81
Sibling group	2	8.00	0.02	0.81
Sibling number	1	6.08	0.01	0.81

⁶ While our sample size and distribution leaves it statistically questionable to consider both sex and sibling number, for completeness we did also run a model that included sex in addition to age and sibling number (our primary variable of interest). Sex did not improve model fit over and above the effect of siblings in any of the three comparisons (ps all >0.54).

Table 3

Full model output from linear mixed effects regression models comparing language

development over time in relation to sibling group. Age in months was included as a fixed

effect; subject was included as a random effect.

Effect	Estimate	Std. Error	df	t value	p
Intercept	-2.69	0.26	156.59	-10.27	< 0.001
SibGroupOne	-0.01	0.30	42.08	-0.05	0.963
SibGroup2+	-0.94	0.33	42.84	-2.81	0.007
month	0.34	0.01	315.13	25.19	< 0.001

Having more siblings was associated with a smaller vocabulary size over the course of early development. This is consistent with previous findings (Hoff-Ginsberg, 1998; Pine, 1995). We find that for each additional sibling, infants were reported to have produced 24% fewer words on average. The 'sibling effect' is thus present in our data.

In terms of our grouped sibling variable (i.e. 0 vs. 1 vs. 2+ siblings), infants with one 326 sibling produced 64 words on average at 18 months, which is, on average, 5 words more 327 than their firstborn peers. Consistent with the model results, infants with two or more 328 siblings produced substantially fewer words at 18 months than those with one or 2+ siblings (based on the raw data: None vs. 2+: 59 vs. 13; One vs. 2+: 64 vs. 13). Post-hoc Wilcoxon Rank Sum tests comparing reported productive vocabulary at 18 months (where 331 there's the widest vocabulary range) revealed significantly larger vocabularies for infants 332 with one sibling compared to those with two or more siblings (W=5, p=.004, 333 CI=[-72.00,-12.00]), but no difference between infants with one sibling and those with no 334 siblings (W=79.50, p=.631, CI=[-34.00,34.00]). See Table 4. 335

⁷ Average difference in vocabulary size between 0 vs. 1, 1 vs. 2, 2 vs. 3 and 3 vs. 4 siblings at 18 months, as a percentage of firstborn vocabulary size.

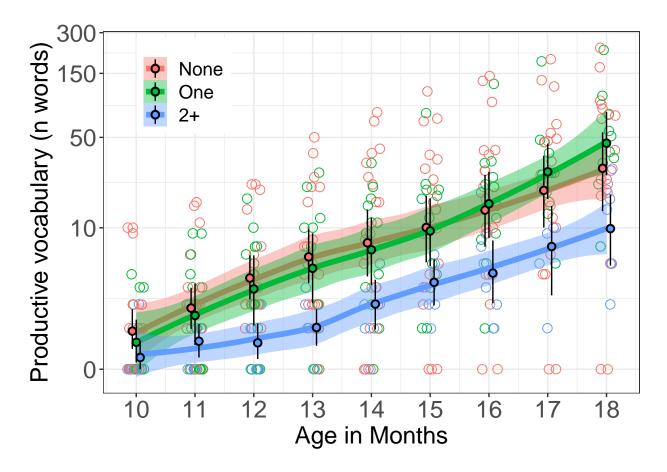


Figure 1. Reported productive vocabulary acquisition (CDI) over time (n=42; one child was an outlier, and was removed from the CDI analysis and this figure; see text for details). Colors denote sibling group; line with colored confidence band reflects local estimator (loess) fit over individual infants' vocabulary at each month. Filled circles indicate mean with bootstrapped CIs computed over each month's data. Open circles (jittered horizontally) show individual infants' vocabulary size at each month. Y-axis utilizes log-transformed vertical spacing for visual clarity.

Effect of siblings on infants' input

Having established that infants' productive vocabulary varied as a function of sibling 337 number in all but the binary version of the measure (0 vs. >0 siblings), we turn to our 338 input measures to test whether *input* varied by a child's sibling status. For these analyses 339 we report here the group sibling division (0 vs. 1 vs. 2+) as this lets us keep relatively 340 similar Ns across groups, thus making variance more comparable for post-hoc comparisons 341 (the discrete sibling number (0-4) version is reported for completeness in S5; results hold 342 for both input variables). We also now include the child who was a multi-month 343 vocabulary outlier above, given that input and vocabulary are not tested in the same 344 model. One infant of the full sample of 43 infants was an outlier in that they heard 345 substantially more input words and words with object presence than all the other infants in 346 the sample in four of their recording sessions. Given that these sessions were not 347 contiguous, we opted to keep this infant in the analyses reported below, though all results 348 hold when they are removed from our sample (see S6). 340

While we didn't have strong a priori expectations about how overall input or object
presence would vary by age or sex, these were included in initial model comparisons to see
if they improved fit alongside a random effect of infant. Both variables improved fit for the
input model, and only age did for the object presence model. Thus our baseline models
include these sets of control variables, respectively. See Table 5 for final model estimates.

Caregiver Input. We tested overall quantity of input (aggregated across the two main caregivers in each session, as outlined above, and siblings) in our model alongside age, sex and subject, as noted above, and a significant effect was found for the effect of sibling group ($\chi^2(2) = 9.09$, p=.011, $R^2=0.59$). Averaging across infants, those with one sibling heard on average 8% fewer words than those with no siblings in any given hour-long recording, while infants with two or more siblings heard 45% fewer words than those with no siblings.

We then ran post-hoc tests to compare mean amount of input across sibling groups; these showed a significant difference in average input received between infants with one sibling versus those with two or more siblings (W=11, p=.002, CI=[-120.87,-39.87]; Bonferroni-corrected p-threshold = .025 for all reported Wilcoxon tests) while amount of input did not differ between infants with no siblings and those with one sibling (W=147, p=.723, CI=[-40.50,57.50]).

While we operationalized caregiver input in our models as input speech from the two 368 adults who produced the most words in any given session, in 87% of cases this was the 369 mother or father. Considering mothers and fathers specifically, maternal input accounted 370 for 75% of object words in the data overall (M=163.14 words, Mdn=151.69, SD=90.73)⁸. 371 Fathers accounted for an average of 18% (M=58.42, Mdn=32.50, SD=64.88), while infants 372 with siblings received around 12% of their input from their brothers and sisters (M=22.97, 373 Mdn=18, SD=18.49). See Figure 2, showing the raw values of input from mothers, fathers 374 and siblings, which are consistent with the group trends reported above. As well as the 375 overall input being greater for firstborns, compared with infants with one or 2+ siblings, 376 note also that the variance is greater for this group, and decreases as sibling number 377 increases. This is shown in the SDs reported in Table 4, and in the data points visualized in Figure 2.

Overall, for infants who had siblings, at least one other child was present in 76% of video recordings (n = 133 recordings, SD = 24%). Wilcoxon Rank Sum tests comparing mean monthly input showed no difference between the amount of sibling input received by infants with one sibling compared with those with two or more siblings (W=31, p=.071, CI=[-14.50,2]). Looking at mothers and fathers individually, infants with two or more siblings heard significantly less input from their mothers than those with one sibling

⁸ One family in our sample had two mothers; rather than artificially assigning one parent to another category, we averaged both mothers' input for this child; we acknowledge that this is an imperfect solution but found it better than the alternatives.

Table 4

Data summary of our two input measures and reported vocabulary size at 18 months. Input measures represent input from the two adults who produced the most words in any given session, plus siblings.

	No siblings		1 sibling		2+ siblings	
Variable	Mean	SD	Mean	SD	Mean	SD
Productive Vocabulary 18m (CDI)	58.89	60.76	64.10	61.97	13.00	9.49
N Input utterances, 10-17 months	213.33	101.51	196.27	53.96	117.36	26.46
% object presence in input, 10-17 months	68.61	10.93	55.90	9.79	46.38	6.17

(W=5, p < .001, CI=[-124.88,-41.92]), while there was no difference between those with one vs. no siblings (W=125, p = .985, CI=[-48.23,51.50]). Finally, amount of paternal input did not differ between groups (one vs. none: W=108, p=.388, CI=[-12.37,57.62]; one vs. 2+: W=21, p=.945, CI=[-74.33,56.54]).

Object presence. On average, 60% of annotated utterances included a referent that was present and attended to by the infant (Mdn=0.61, SD=0.12). See Table 4. Consistent with our hypothesis that infants with more siblings would hear fewer words in referentially transparent conditions (i.e. they would experience lower object presence) than those with fewer siblings, our models reveal a significant effect for sibling group on object presence ($\chi^2(2) = 27.52$, p < .001, $R^2 = 0.55$).

Descriptively, infants with no siblings experienced on average 32% more object presence in their input than those with two or more siblings, and 19% more than those with one sibling. Post-hoc comparisons revealed significant between-group differences: infants with no siblings experienced significantly more object presence than those with one sibling (W=240, p<.001, CI=[0.07,0.20]; Bonferroni-corrected p-threshold = .025). Likewise, infants with one sibling experienced significantly more object presence those with

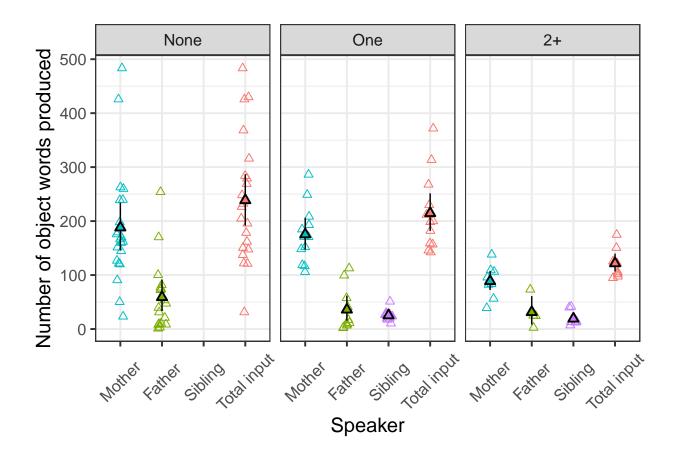


Figure 2. Mean number of words produced by Mothers, Fathers and Siblings, as well as total family input (mean input from mother + father + sibling(s)), across sessions recorded between 10-17 months. Open triangles represent values for individual infants; filled triangles show group means. In the case where the infant had two mothers, mean maternal input is shown. See S7 for a month-by-month visualisation of caregiver input.

two or more siblings (W=25, p=.025, CI=[-0.18,-0.01]). See Table 5 and Figure 3. Table 5

Full model output from linear mixed effects regression models comparing our two input measures (object words produced in caregiver input and object presence) over time in relation to sibling group. Age in months was included as a fixed effect in both models, sex was included in the caregiver input model only; subject was included as a random effect.

Variable	Effect	Estimate	Std. Error	df	t value	p value
Caregiver input	Intercept	4.88	0.18	185.99	27.46	< 0.001
	SibGroupOne	0.01	0.15	43.00	0.05	0.960
	SibGroup2+	-0.49	0.17	43.00	-2.98	0.005
	month	0.03	0.01	301.00	2.95	0.003
	sexM	-0.18	0.13	43.00	-1.41	0.164
Object presence	Intercept	0.57	0.04	321.44	12.73	< 0.001
	SibGroupOne	-0.13	0.03	43.00	-3.81	< 0.001
	SibGroup2+	-0.22	0.04	43.00	-5.90	< 0.001
	month	0.01	0.00	301.00	2.93	0.004

Sibling presence. So far, our analysis takes into account the differences in input 403 based on whether or not the target child has a sibling, but does not directly consider 404 whether sibling presence in the recordings affected these variables. That is, if it is the 405 active presence of the sibling that affects how the caretaker interacts with the target child, 406 then we would expect to see a difference in our input measures when the sibling is present 407 vs. absent. On the other hand, if the very fact of having a sibling changes the way that a 408 caregiver interacts with the infant regardless of whether any sibling is actually present on 409 the scene, then no difference should be observed. While sibling presence in each recording 410 was not coded directly in the dataset, for this exploratory analysis we can get at this with 411 an admittedly imperfect proxy: did the sibling produce nouns in the recording? If yes, we 412 can safely assume they are present; if not we (less safely, but reasonably for initial 413 exploratory purposes) assume they are not. As reported above, by this measure, at least 414 one sibling was present in 76% of recordings for the infants who had a sibling. 415

Since the presence of a sibling in any given infant's data changed month-on-month (i.e. sometimes the sibling was present and sometimes they were not), and since our

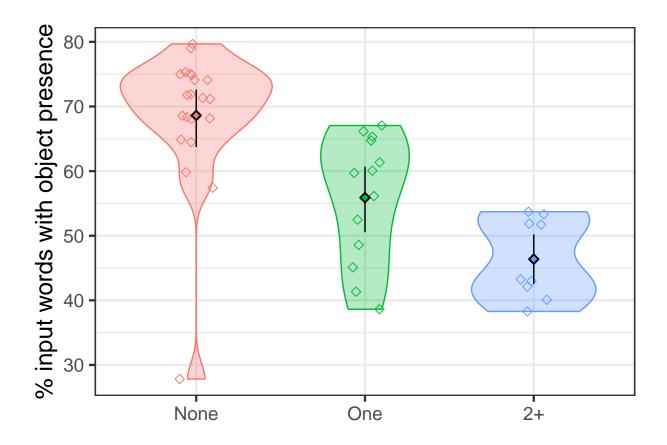


Figure 3. Percentage of input words produced with object presence across sibling groups. Error bars and filled diamonds show 95% CIs and mean proportion of object presence across sibling groups. Open shapes indicate mean proportion of object presence per infant, collapsing across age and jittered horizontally for visual clarity. See S7 for a month-by-month visualisation of object presence.

measure of sibling presence is imperfect, we opt here to describe the pattern of data
without drawing any strong conclusions from statistical models. Descriptively, the presence
of a sibling affected the amount of object presence in the data, but not the amount of
input. See Table 6 and Figures 4 and 5. Overall, the presence of a sibling negatively
affected object presence, and this was consistent over time; when a sibling was present,
infants in both groups heard less object presence. This effect was stronger for infants with
two or more siblings (though note that it is unclear from our measure how many siblings

Table 6

Data summary of our two input measures according to presence or absence of siblings

during the recording. Input measures represent input from the two adults who produced the

most words in any given session, plus siblings.

		1 sibling		2+ siblings	
Variable	Sibling presence	Mean	SD	Mean	SD
N Input utterances, 10-17 months	Sibling not present	139.36	73.23	88.67	32.39
	Sibling present	126.92	38.69	75.01	26.29
% object presence in input, 10-17 months	Sibling not present	70.26	14.39	64.60	12.10
	Sibling present	53.91	8.42	39.25	8.99

were present, and it is possible that only one sibling was present in the recording), and 425 overall this was true regardless of whether the sibling was present, or whether the infant was alone with the caregiver. The picture is less clear for caregiver input, where the presence of a sibling has a more variable effect on the number of object words produced by 428 caregivers (see Figure 4), particularly for infants with only one sibling. This supports our findings above, suggesting that the presence of one additional child does not have any 430 negative effects on the amount of input that caregiver provide. However, input was 431 consistently lower for the group with two or more siblings, and this was true regardless of 432 whether or not a sibling was present; again, this is consistent with the findings reported 433 above for this input measure. 434

435 Discussion

We investigated the nature of infant language development in relation to number of children in the household. Previous research found a delay in lexical acquisition for later-born infants (Fenson et al., 1994; Hoff, 2006), with differences in input across birth

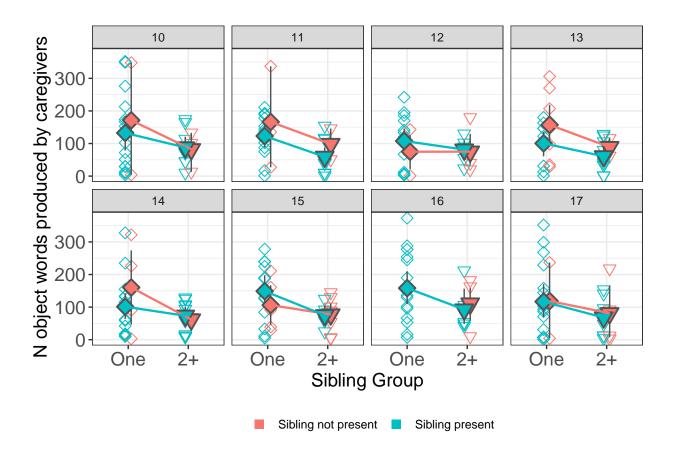


Figure 4. Difference in number of input words in infants' input according to whether or not a sibling or siblings were present during the time of recording, for each month of data. Infants with no siblings were not included in the plots for visual ease. Open diamonds represent individual infants in the data; filled diamonds represent means and 95%CIs, colours represent presence or absence of siblings during the recording session.

order reported as a root cause. Our results add several new dimensions to this, by testing
for differences across more vs. fewer older siblings, and by looking at input during
child-centered home recordings. Infants with more siblings were reported to say fewer
words by 18 months, heard fewer nouns from their parents, and were less likely to be
attending to an object when hearing its label.

Importantly, and in contrast with some previous research (Hoff-Ginsberg, 1998;
Oshima-Takane & Robbins, 2003), infants with one sibling showed no decrement in lexical

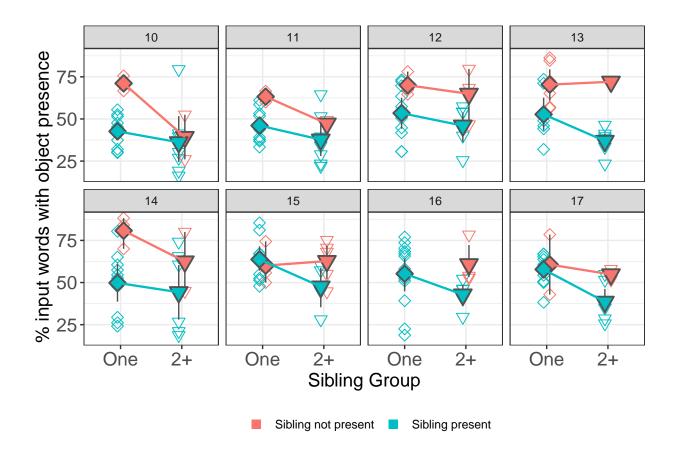


Figure 5. Difference in % of object presence in infants' input according to whether or not a sibling or siblings were present during the time of recording, for each month of data. Infants with no siblings were not included in the plots for visual ease. Open shapes represent individual infants in the data; filled shapes represent means and 95%CIs, colours represent presence or absence of siblings during the recording session.

production and minimal reduction in input in comparison to first-born infants. That is, our results suggest that simply having a sibling does not contribute to input or vocabulary differences across children (as measured here), while having more than one sibling seems to do so. Indeed, infants with zero and one sibling had similar results for productive vocabulary, and parental noun input overall (though not object presence). Moreover, parental input was not affected by the presence or absence of the sibling in the room. In contrast, infants with two or more siblings said fewer words, and also heard fewer input

words overall.

With regards to object presence, having siblings made it increasingly less likely to
hear an object label when attending to it, and this effect increased with sibship size

(i.e. children with more siblings heard input containing a lower proportion of object

presence). Unlike for total parental noun input (which was reduced for 2+ siblings but not

modulated based on whether siblings were present in the recordings), reduced object

presence for children with more siblings was particularly notable in recordings with siblings

present.

When we considered the effect of sibling status – that is, The sibling effect. 461 whether or not infants had any siblings, disregarding specific sibling number – our findings 462 showed that having siblings made no difference to infants' lexical production capacities. 463 This contrasts with Hoff-Ginsberg (1998), who found that, by 18 months, laterborns 464 exhibit lower language skills. However, Oshima-Takane and colleagues (1996) found no 465 overall differences between first- and second-born children across a range of language 466 measures taken at 21 months. Our results suggest that considering sibling quantity may be 467 a more sensitive way to reveal demographic effects than their (coarser-grained) first-468 vs. later-born status. We find that the more older siblings a child had, the lower their 469 reported productive vocabulary at 18 months. This adds to findings from Fenson and 470 colleagues (1994), who found a weak but significant negative correlation between birth 471 order and production of both words and gestures. Controlling for age, our model showed that infants with 2 or more siblings produced 46 fewer words than the average 59 words 473 produced by firstborns in our data by 18 months.

While infants with more siblings heard less input speech overall, having one sibling
did not significantly reduce the number of nouns in an infant's input. This is in direct
contrast with reports from the literature; Hoff (2006) states that "when a sibling is present,
each child receives less speech directed solely at...her because mothers produce the same
amount of speech whether interacting with one or two children" (p.67, italics added). While

this does not appear to be the case in the present dataset, it may be due to the circumstances of the home-recorded data: while siblings were present in many of the 481 recordings (76\% of recordings in which the target child had a sibling), given the focus of 482 the data collection, parents may have had a tendency to direct their attention - and 483 consequently their linguistic input - more towards the target child; our samples also 484 differed in other ways (e.g. sociocultural context) that may have influenced the results as 485 well. Alternatively, our results may diverge from those of Hoff (2006) due to the nature of 486 our input measure, which only took nouns into account. That said, we find this alternative 487 explanation unlikely given work by Bulgarelli and Bergelson (2020) showing that nouns are 488 a reliable proxy for overall input in this dataset, suggesting that this measure provides an 489 appropriate representation of overall input directed at the target child. 490

In contrast to the other results, our analysis of object presence showed a more linear 491 'sibling effect'. In this case, even having one sibling led to fewer word-object pairs presented 492 in the input. This was true regardless of whether or not other siblings were present, but 493 object presence was further negatively affected by the presence of a sibling in the room. 494 Presence of a labeled object with congruent input speech has been found to support early 495 word learning across several studies. For instance, Bergelson and Aslin (2017) combined 496 analysis of this home-recorded data at six months with an experimental study to show that 497 word-object presence in naturalistic caregiver input correlated with comprehension of 498 nouns (tested using eye-tracking). Relatedly, Gogate and colleagues (2000) propose that 499 contingent word production supports the learning of novel word-object combinations, with 500 "multimodal motherese" - whereby a target object word is produced in movement or 501 touch-based synchrony with its referent - supporting word learning. More broadly, lower rates of referential transparency for common non-nouns like hi and uh-oh have been proposed to potentially explain why these words are learned later than common concrete 504 nouns (Bergelson & Swingley, 2013). While the present results on object presence don't 505 speak directly to word learning, they do suggest that this potentially helpful learning 506

support is less available for children with more siblings.

Siblingese as a learning opportunity? We also found that infants with siblings 508 did not hear much speech from their older brothers and sisters. Similar findings are 509 reported in a lab-based interaction study by Oshima-Takane and Robbins (2003), who 510 found that older siblings rarely talked directly to the target child; instead, most input from 511 siblings was overheard speech from sibling-mother interactions. One possibility raised by 512 these results is that perhaps parents are able to compensate or provide relatively similar 513 input and learning support for one or two children, but once children outnumber parents, 514 this balancing act of attention, care, and time becomes unwieldy. While the current sample 515 is relatively limited and homogeneous in the family structures and demographics it includes, 516 future work could fruitfully investigate this possibility by considering whether (controlling 517 for other potential contributors like SES, Hoff-Ginsberg, 1998) the presence of more 518 caregivers (whether parents, relatives, or other adults) helps foster language development. 519

Alternatively, second-borns might 'even out' with children with no siblings due to a 520 trade-off between direct attention from the caregiver and the possibility of more 521 sophisticated social-communicative interactions. For these infants there is still ample 522 opportunity to engage with the mother in one-to-one interactions, allowing a higher share 523 of her attention than is available to third- or later-borns. Furthermore, triadic interactions 524 can benefit the development of a number of linguistic and communication skills (Barton & 525 Tomasello, 1991; Dunn & Shatz, 1989). Second-borns may also benefit from overheard 526 speech in their input, supporting the acquisition of nouns and even more complex lexical 527 categories (Floor & Akhtar, 2006; Oshima-Takane et al., 1996). For infants with one sibling, the benefits of observing/overhearing interactions between sibling and caregiver, as well as the possibility for partaking in such interactions, may outweigh the decrease in 530 some aspects of the input (i.e., in our data, only observed in object presence). Having more 531 than one sibling may throw this off-balance, such that the possibilities for both supportive 532 one-to-one input and more sophisticated interactions are simultaneously diminished. 533

Importantly, the present results make no claims about eventual outcomes for these 534 children: generally speaking, regardless of sibling number, all typically-developing infants 535 reach full and fluent language use. Indeed, some research suggests that sibling effects, while 536 they may be clear in early development, are not always sustained into childhood; e.g. twins 537 are known to experience a delay in language development into the third year, but are quick 538 to catch up thereafter (Dales, 1969; Tomasello, Mannle, & Kruger, 1986). This 539 demonstrates the cognitive adaptability of early development, which brings about the 540 acquisition of language across varying and allegedly 'imperfect' learning environments. Infants' capacity to develop linguistic skills from the resources that are available to them – 542 whether that is infant-directed object labels or overheard abstract concepts – highlights the 543 dynamic and adaptable nature of early cognitive development, and a system that is sufficiently robust to bring about the same outcome across populations.

Of course, the 'success' of early language development is defined by Limitations. 546 how success is measured. Here we chose word production as our measure of linguistic 547 capability; we did not consider other equally valid measures such as language 548 comprehension or early social-interaction skills. Similarly, our input measures focused on 549 nouns; other lexical classes may reveal different effects, though they are generally far 550 sparser in production until toddlerhood. Our analysis of vocabulary relied on parental 551 report data; this method could have biased our first-born sample towards more accurate or 552 larger vocabulary reports owing to their parents having more time and attention to spend 553 observing their vocabulary development (see Kartushina et al., 2022 for a discussion of this 554 possibility in light of the COVID-19 pandemic, though note the present data were collected in 2014-16). In the supplementary materials, we provide validation data for the CDI relative to children's own productions by running correlation tests between reported (CDI) 557 vocabulary and the number of word types produced by each infant in the audio and video 558 data. In short, for the 0 and 1 sibling groups there is a strong and significant positive 559 correlation between the CDI at 18 months and child word types; for 2+ siblings this 560

correlation is weaker and does not reach significance, though this is likely due to the small n in this group relative to the others (due to missing CDIs at 18 months, this includes data from n = 18/21, 11/12, and 7/9 children with 0, 1, and 2+ siblings, respectively); see S8.

There is also some imbalance in group sizes across our data; our sample was not 564 pre-selected for sibling number, and so group sizes are unmatched across the analysis. 565 Including a larger number of infants with 2+ siblings may have revealed a different pattern 566 of results. We might also expect that age of older siblings would affect the nature of the 567 early linguistic environment, given that larger age difference is found to be a predictor of 568 lower vocabulary size in the current literature (Gurgand et al., 2022; Havron et al., 2022); 569 our sample did not allow us to link sibling age to number of words produced by that 570 sibling, but future work may wish to take this into account. Finally, more work across 571 wider and larger populations is necessary to unpack the generalizability of the present 572 results. Our sample is reflective of average household sizes in middle-class families across 573 North America and Western Europe (Office for National Statistics, 2018; United States 574 Census Bureau, 2010), but it is not unusual in some communities and parts of the world for 575 households to include between three and six children on average (Institute for Family 576 Studies & Wheatley Institution, 2019). Adding to this, it is also necessary to consider 577 cross-cultural differences in the way children are addressed by their parents, other 578 caretakers, and other children (Bunce et al., 2020; Casillas, Brown, & Levinson, 2019; 579 Shneidman & Goldin-Meadow, 2012). For instance, Bunce and colleagues (2020) find 580 relatively similar rates of target child directed speech across US, Canadian, Argentinian, UK, Papuan and Mayan samples, some differences in who the input comes from, and large effects of number of talkers present. These results suggest that caution is advisable before generalizing the current results to any other socio-cultural contexts, but also pose exciting 584 open questions regarding what variability in experiences do – or don't – change about early 585 language interaction and development. 586

7 Conclusion

Our results with English-learning infants in the US support prior findings from the 588 literature showing that later-born infants have slower lexical acquisition than their 589 first-born peers. However, we highlight an important difference from previous findings, 590 namely that in the present sample, second-born infants show no such effect, while infants 591 with more than two siblings have significantly smaller productive vocabularies at age 18 592 months. We also identified similar group differences in overall noun input and object 593 presence. While we did not test these corresponding vocabulary and input measures 594 directly, our results suggest that having more siblings affects a child's early language environment, which in turn may lead to slower vocabulary growth in the first 18 months of life. We look forward to future studies considering the granularity of more versus fewer siblings, and how this relates to language abilities over the course of development.

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