- Analysing the effect of sibling number on input and output in the first 18 months
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9 Abstract

Prior research suggests that across a wide range of cognitive, educational, and health-based measures, first-born children outperform their later-born peers. Expanding on this literature using naturalistic home-recorded data and parental vocabulary report, we find that early language outcomes vary by number of siblings in a sample of 43 English-learning U.S. children from mid-to-high socioeconomic status homes. More specifically, we find that children in our sample with two or more - but not one - older siblings had smaller productive vocabularies at 18 months, and heard less input from caregivers across several measures than their peers with two or more siblings. We discuss implications regarding what infants experience and learn across a range of family sizes in infancy.

19 Keywords: Siblings, Lexical Development, Input Effects, Language Acquisition

20 Word count: X

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Analysing the effect of sibling number on input and output in the first 18 months

A common simplifying assumption in research on language development is that there is 22 a theoretical "optimum" environment for early language, whereby the input is tailored to a 23 single infant's needs, changing over time as language capacity grows (e.g. Soderstrom, 2007; 24 Stern, Spieker, Barnett, & MacKain, 1983). However, for many infants and for many reasons, language acquisition occurs across diverse social contexts that can influence the learning environment, including the presence of older siblings in the home (Fenson et al., 1994). According to the United States Census Bureau (2010), around one third of children are born into households with at least one other infant present, and one in every five infants is acquiring language in a household shared with two or more other children. Similar statistics are reported for British infants (Office for National Statistics, 2018), where the average 31 household has 1.75 children, and 15% of households have three children or more. In this paper, we consider the role of siblings in the early language environment of English-learning infants. We use naturalistic home-recorded data to measure input in earlier- and later-born infants in relation to their productive vocabulary over the first 18 months of life.

Prior research suggests that infants born to households with older children may be
slower to learn language. Fenson and colleagues (1994) found that by 30 months of age,
children with older siblings performed worse than those with no siblings across
parent-reported measures of productive vocabulary, use of word combinations, and mean
length of utterance. This 'sibling effect' may be the result of differences in input between
first- and later-born children: some research finds that infants with older siblings hear less
speech aimed specifically at them, and what they do hear is understood to be linguistically
less supportive of early language development (Hoff-Ginsberg, 1998; Oshima-Takane &
Robbins, 2003). In contrast, some studies have noted linguistic advantages for later-borns,
who may have stronger social-communicative skills (Hoff, 2006), better understanding of
pronouns (Oshima-Takane, Goodz, & Derevensky, 1996), and better conversational abilities

- (Dunn & Shatz, 1989). Overall, while the particulars differ across studies, prior work suggests that the presence of siblings in the home leads to differences in infants' early linguistic experiences and skills, though the direction of these differences varies depending on what aspects of language are being measured.
- Numerous studies have attempted to better understand how siblings affect the 51 language development trajectory, with comparisons of language acquisition across first- and later-borns, and analyses of mothers' input in dyadic (infant + mother) and triadic (infant + mother + older sibling) situations. Here again, findings are mixed, but overall two general conclusions can be drawn. First, analyses consistently show that infants with older siblings generally have slower vocabulary development (Berglund, Eriksson, & Westerlund, 2005; Fenson et al., 1994; Pine, 1995; Zambrana, Ystrom, & Pons, 2012), though effect sizes tend to be small, with significant differences more typically found in the earliest stages of language learning. Hoff-Ginsberg (1998) shows first-borns to have better lexical and syntactic skills up until 2;5, but later-born infants had better conversational abilities during the same time-period. Relatedly, using a large longitudinal dataset of French-learning 2-5 61 year olds, Havron and colleagues (2019) find no effect of age gap between siblings, but lower standardized language scores in children with older brothers (but not sisters) relative to those without siblings, based on parental report and direct battery assessments. Some of these differences across studies may relate to insufficient power to detect relatively small effects or simultaneous contributing factors that are difficult to disentangle.
- The second general finding pertains to sibling-related differences in the early linguistic environment: infants with no siblings receive more input overall, and this more closely reflects what is typically considered to be 'high quality' input in the extant literature. Indeed, the very presence of a sibling in the linguistic environment changes the way language is used. When siblings are present (i.e. triadic interactions), mothers' input has been found to be more focused on regulating behaviour, as opposed to the language-focused speech that is

common in dyadic contexts (Oshima-Takane & Robbins, 2003). Reports show that the mean length of utterance is longer in the input of first-born infants (Hoff-Ginsberg, 1998; but see also Oshima-Takane & Robbins, 2003 for a comparison of dyadic and triadic contexts), who also hear more questions directed at them than later-borns. Both Jones and Adamson (1987) and Oshima-Takane and Robbins (2003) report no difference between the overall number of word types produced by mothers in dyadic and triadic settings, but the proportion of speech directed at the target infant is drastically reduced when input is shared with siblings.

As Hoff (2006) explains, infants with siblings have less experience of speech directed at them, but they do have an advantage over their first-born peers in that they are subject to more overheard speech. Barton and Tomasello (1991) show that by as early as 19 months, infants with siblings are already able to take part in triadic conversations, which were almost three times longer than dyadic conversations. The authors suggest that the presence of siblings may shift the learning context, and facilitate infants' participation in communicative interactions: infants are under less pressure to participate in a triadic interaction, meaning the conversation can continue even when the infant is unable to respond. As a result, the infants in Barton and Tomasello's study took more conversational turns in triadic interactions than dyadic ones.

There thus may be a trade-off in development between highly supportive one-to-one input from a caregiver (cf. Ramírez-Esparza, García-Sierra, & Kuhl, 2014) and the potential benefits drawn from communicating with (or overhearing communication with) a sibling. In the present study, we test the extent to which having more versus fewer siblings in the home environment may lead to differences in vocabulary development and the early linguistic environment over the course of the first 18 months of life. In analyzing infants' growing productive vocabulary in relation to the presence of older siblings in their household, the present work expands on the extant literature in two key ways. First, prior work generally considered birth order as a binomial factor (i.e. comparing first-born infants with

second-borns, e.g. Oshima-Takane & Robbins, 2003), or 'later-borns' (e.g. Hoff-Ginsberg, 1998), potentially missing graded effects. Instead of this approach, we consider how having 100 more versus fewer siblings is linked to an infant's lexical development and their early 101 linguistic environment. Second, much of the existing literature in this area is drawn from 102 questionnaire data or brief interactions recorded in the lab (but see Dunn & Shatz, 1989 for 103 a study of naturalistic home-recorded data), rather than naturalistic day-to-day interactions 104 in the home. In contrast, we analyze an existing corpus of home recordings in concert with 105 vocabulary checklists. Based on work summarized above, we expect that both the language 106 environment and infants' early productive vocabulary will vary as a function of how many 107 older siblings they have. 108

109 Hypotheses

In broad strokes, given prior research showing that early lexical development is more 110 advanced among first-born infants (e.g. Hoff-Ginsberg, 1998), our prediction regarding 111 infants' productive vocabulary is that if this effect is gradient, then children with more 112 siblings will have lower productive vocabularies than their peers with fewer siblings. With 113 regard to the infants' linguistic environment, we hypothesize that infants' with more siblings 114 will experience lower prevalence of two aspects of the language input previously shown to 115 support language development: amount of input and amount of object presence. To 116 unpack each of these input aspects, following previous studies that show infants with siblings 117 to receive less speech directed at them (Jones & Adamson, 1987; Oshima-Takane & Robbins, 2003), we expect to see the same effect in our sample. By object presence we mean word and object co-occurrence, e.g. mother saying "cat" when the child is looking at a cat. We predict object presence will decrease as sibling number increases, because as caregivers' attention is 121 drawn away from one-to-one interactions with the infant, there is likely less opportunity for 122 contingent talk and joint attention. Prior research suggests links between object presence 123

and early word learning (Bergelson & Aslin, 2017; Cartmill et al., 2013), though to our knowledge this has not been examined in relation to sibling number.

126 Methods

We analyze data from the SEEDLingS corpus (Bergelson, Amatuni, Dailey, 127 Koorathota, & Tor, 2019), a longitudinal set of data incorporating home recordings, parental 128 reports and experimental studies from the ages of 0;6 to 1;6. The present study draws on the 129 parental report data to index child vocabulary size, and annotations of hour-long home video 130 recordings, taken on a monthly basis during data collection, to index input. We also ran our 131 input analysis using data sub-sampled from day-long audio recordings taken on a different 132 day from the video data reported below; all results were consistent with those outlined below 133 (see Supplementary Materials). We note at the outset that with such a multidimensional 134 dataset there are always alternative ways of conducting analyses of input and output; our 135 goal here is to make motivated decisions that we clearly describe, provide some alternative 136 analytic choices in the supplementals, and to share the data with readers such that they are 137 free to evaluate alternative approaches. 138

139 Participants

Forty-four families in New York State completed the year-long study. Infants (21 females) were from largely middle-class households; 33 mothers had attained a B.A. degree or higher. Based on parental report, no infants had speech- or hearing-relevant diagnoses; none were low birth weight (all >2,500g); 42 were white, two were from multi-racial backgrounds. All infants heard >75% English on a regular basis and lived in two parent homes. Two participants were dizygotic twins; we retain one twin in the current sample, considering the other only as a sibling. Thus our final sample size was 43 infants.

Sibling Details. Sibling number was computed based on parental report in the 147 demographics questionnaires completed at 0:6 (R: 0-4). Siblings were on average 4.11 years 148 older than the infants in this study (Mdn days: 1,498.50, SD: 1465, R: 0-17 years). All 149 siblings lived in the household with the infant full time, apart from one infant who had two 150 older half siblings (and no other full siblings) who lived with their other parent part of the 151 time. Both older siblings were present for at least some of the monthly recordings. One 152 family had a foster child live in the home for 2 months of data collection, who is not 153 accounted for in our data; the target infant had one sibling. All siblings were older than or of 154 the same age as the infant in question. 155

156 Materials

Parental report data. To index each child's language abilities, we draw on data 157 from vocabulary checklists (MacArthur-Bates Communicative Development Inventory, 158 hereafter CDI; Fenson et al., 1994), administered monthly from 0:6 to 1:6, along with a 159 demographics questionnaire; each month's CDI survey came pre-populated with the previous 160 month's answers to save on reduplicated effort. Because the majority of infants did not 161 produce their first word until around 0:11 according to CDI reports (M=10.67, SD=2.23), we 162 use CDI data from 0;10 onwards in our analysis. CDI production data for each month is 163 taken as a measure of the infants' lexical development. CDI data for production has been 164 well validated by prior work, including work in this sample (Frank, Braginsky, Yurovsky, & 165 Marchman, 2021; Moore, Dailey, Garrison, Amatuni, & Bergelson, 2019). Of the intended 13 166 CDIs collected for each of the 43 infants, 26 were missing across 11 infants (leaving 559 CDIs 167 in total). 4 infants had 4 CDI data-points missing, while the majority (n = 5) had only one 168 missing data-point. 169

¹For six infants, siblings' exact birthdates were not provided, and so age difference was estimated by subtracting the infant's age (6 months) from the sibling's age in years, as listed on the questionnaire (e.g. if a sibling was 5 years old, they were classed as being 4.5 years older than the infant).

Home-recorded video data. Every month between 0.6 and 1.5, infants were 170 video-recorded for one hour in their home, capturing a naturalistic representation of each 171 infant's day-to-day input. We did not ask families to ensure certain family members were or 172 were not present; our video recordings capture whoever was home at the time families opted 173 to schedule. Infants were a hat with two small Looxcie video cameras attached, one pointed 174 slightly up, and one pointed slightly down; this captured the scene from the infants' 175 perspective. In the event that infants refused to wear the hats, caregivers were the same 176 kind of camera on a headband. Additionally, a camcorder on a tripod was set up in the room 177 where infants and caretakers were interacting to capture a broader view; families were asked 178 to move this camcorder if they changed rooms. The dataset includes 12 videos for each child 179 except one, who had 11. 180

Object words (i.e. concrete nouns) deemed to be said to, by, or loudly and clearly near the target child were annotated by trained coders for several properties of interest to the broader project on noun learning. Here we examine annotations for speaker, i.e. who produced each noun, and object presence, i.e. whether the noun's referent was present and attended to by the infant (see "Derived Input Measures" below). See Bergelson et al. (2019) for further details on the full set of home-recorded data and its annotations.

Derived Input measures. Two input measures were derived based on the 187 annotations of concrete nouns in this corpus, each pertaining to an aspect of the input that 188 is established as important in early language learning: overall household input (how 189 many concrete nouns does each infant hear?) and object presence (how much of this input is referentially transparent?), detailed below. We note that neither of these measures are, in 191 our view, interpretable as "pure" quality or quantity input measures; we hold that quality 192 and quantity are inextricably linked in general, and specifically we include (by design) only 193 object words that the recordings suggest were possible learning instances for the infants who 194 heard them, wherein quantity and quality are conflated.

Household Input reflects how many nouns infants heard in the recordings from their 196 mother, father and (where relevant) siblings. Other speakers' input was relatively rare 197 during video recordings, accounting for 11.42% of input overall (SD=22.82%), and is 198 excluded from our analysis. This measure of the early language environment is based on 199 evidence showing strong links between the amount of speech heard in the early input and 200 later vocabulary size (Anderson, Graham, Prime, Jenkins, & Madigan, 2021). This analysis 201 considers only nouns produced by speakers in the child's environment (which is what was 202 annotated in the broader SEEDLingS project); concrete nouns are acquired earlier in 203 development in English and cross-linguistically (Braginsky, Yurovsky, Marchman, & Frank, 204 2019). In this corpus (as in any sample of naturalistic interaction) the number of nouns 205 correlates highly with the number of words overall (e.g. based on automated analyses of 206 adult word counts vs. manual noun-only annotations; Bulgarelli & Bergelson, 2020). Thus, 207 noun count in the monthly hour of video data serves as our household input proxy. 208

Object Presence was coded as "yes," "no," or "unsure" for each object word annotated in the home recordings based on trained annotators' assessment of whether the referent of the word (i.e. the object) was present and attended to by the child. In the video data, 169 instances (0.25% on average per infant) of object presence were marked as unsure; these instances were not included in this analysis.

In the following analyses, we consider infants' total productive vocabulary² alongside
these two input measures – nouns in household input and extent of object presence in the
input – as a function of sibling number. Since the raw data are highly skewed,
log-transformed data and/or proportions are used for statistical analysis (1 was added to the
raw infant production data of all infants before log-transformation to retain infants with

²While in principle we could have just used noun productive vocabulary, in practice noun and total vocabulary is correlated >.95 in this age range; we opted to retain the overall total vocabulary, as lexical class is not a straightforward notion in the early lexicon.

Table 1
Sibling number by female and male infants.

n Siblings	Female	Male	Total
0	9	12	21
1	7	6	13
2	2	3	5
3	2	0	2
4	0	2	2
Total	20	23	43

vocabularies of 0.) All figures display non-transformed data for interpretive ease.

Results

Vocabulary development was highly variable across the 43 infants, according to the 221 CDI data we had available. By 18 months, 2 infants produced no words (taken from 36 222 available CDIs at this time-point), while mean productive vocabulary size was 60.28 words 223 (SD=78.31, Mdn=30.50). Three infants had substantially larger-than-average (3SDs above 224 the mean monthly vocabulary score) vocabularies at certain time-points in the data; we 225 opted to count one of these infants as an outlier and remove this child's data from the CDI analysis given that their vocabulary was higher for multiple consecutive months (1:1-1:6). The other two infants had higher vocabularies at 10-11 months only (when variance was quite limited, see Figure 1), and were retained to maximize data inclusion. This left 42 229 infants (19 females) in the analysis of vocabulary size. Infants had one sibling on average 230 (M=0.86, Mdn=1, SD=1.10). See Table 1.

232 Model structure for fixed and random effects

All reported models were generated in R (R Core Team, 2019) using the *lmerTest* 233 package to run linear mixed-effects regression models (Kuznetsova, Brockhoff, & Christensen, 234 2017). P-values were generated by likelihood ratio tests resulting from nested model 235 comparison. All models include infant as a random effect. All post-hoc tests are two-sample, two-tailed Wilcoxon Tests; given that all of our variables of interest (CDI score, household 237 input and object presence) differed significantly from normal by Shapiro tests, we opted to run non-parametric tests for all post-hoc comparisons. Where multiple post-hoc comparisons are run on the same dataset, Bonferroni corrections are applied (e.g. with an adjusted 240 p-value threshold of .025 for 2 between-group comparisons are done). While we have a 241 substantial amount of data for each participant, our limited n means we are under-powered 242 to consider multiple demographic variables simultaneously given the data distribution 243 (e.g. sibling number and sex, see Table 1; as luck would have it both infants with 3 siblings 244 were girls and both with 4 were boys). See Supplementary Data for correlations between 245 sibling number and maternal age/education. 246

247 Effect of siblings on infants' productive vocabulary

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We first modeled the effect of siblings on reported productive vocabulary. We explored
three possible variations on how to represent the sibling effect: a binary variable (0 vs. >0
siblings), aggregated groups (None vs. One vs. 2+ siblings), and discrete sibling number (0
to 4 siblings), comparing the following nested model structures, where (1) is the baseline
model and (2) includes siblings as the variable of interest.

- 1. vocabulary size (log-transformed) \sim age (months) + (1|subject)
- 25. vocabulary size (log-transformed) ~ siblings [binary, group or discrete] + age (months)
 + (1|subject)

In our sample, simply having siblings (i.e. as a binary variable) did not predict CDI productive vocabulary size, while both discrete sibling number and sibling group did. See Table 2.3

Table 2

Output from likelihood ratio tests

comparing regression models that predict

the effects of sibling number (binary,

grouped and discrete variables) on

vocabulary size. Month was included in

each model as a fixed effect; subject was

included as a random effect.

Model	Df	Chisq	p value
0 vs. >0 siblings	1.00	2.04	0.15
Sibling group	2.00	7.03	0.03
Sibling number	1.00	5.26	0.02

Table 3

Full model output from linear mixed effects regression model comparing language development over time in relation to sibling group. Age in months was included as a fixed effect; subject was included as a random effect.

Effect	Estimate	Std. Error	df	t value	p
Intercept	-1.92	0.17	95.77	-11.38	< 0.001
SibGroupOne	-0.03	0.23	42.11	-0.11	0.911

³While our sample size and distribution leaves it statistically questionable to consider both sex and sibling number, for completeness we did also run a model that included sex in addition to age and sibling number (our primary variable of interest). Sex did not improve model fit over and above the effect of siblings in any of the three comparisons (ps all >0.68).

SibGroup2+	-0.66	0.25	42.90	-2.64	0.011
month	0.29	0.01	479.07	34.37	< 0.001

Having more siblings was associated with a smaller vocabulary size over the course of early development. This is consistent with previous findings (Hoff-Ginsberg, 1998; Pine, 1995). We find that for each additional sibling, infants were reported to have produced 21.14% fewer words. The 'sibling effect' is thus present in our data.

In terms of our grouped sibling variable (i.e. 0 vs. 1 vs. 2+ siblings) infants with one 263 sibling acquire only 3% fewer words than firstborns over the course of our analysis, while 264 infants with two or more siblings produce 66% fewer words. See Table 3 and and Figure 1. 265 Post-hoc Wilcoxon Rank Sum tests comparing reported productive vocabulary at 18 months 266 (where there's the widest vocabulary range) revealed significantly larger vocabularies for 267 infants with one sibling compared to those with two or more siblings (W=5, p=.004), but 268 no difference between infants with one sibling and those with no siblings (W=79.50, p=269 .631). See 4. 270

Effect of siblings on infants' input

Having established that infants' productive vocabulary varied as a function of how
many siblings in all but the binary version of the measure (0 vs. >0 siblings), we turn to our
input measures to test whether *input* varied by a child's sibling status. For these analyses we
report here the group sibling division (0 vs. 1 vs. 2+) for as this lets us keep relatively
similar Ns across groups, thus making variance more comparable for post-hoc comparisons
(the discrete sibling number (0-4) version is reported for completeness in the supplementals;
results hold for object presence but diverged somewhat for overall household input). We also
now include the child whose who was a multi-month vocabulary outlier above given that

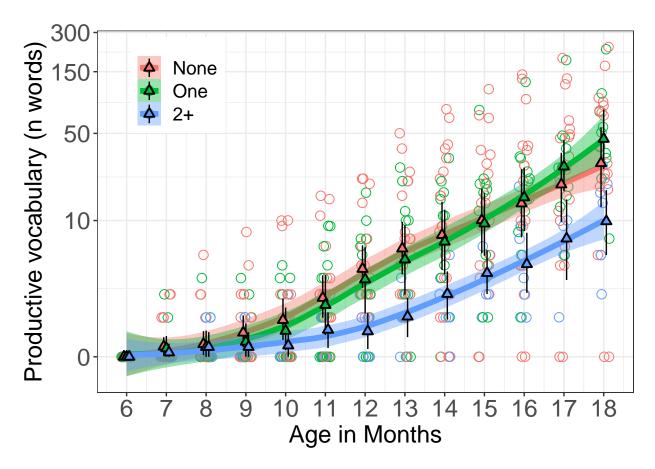


Figure 1. Reported productive vocabulary acquisition (CDI) over time. Colors denote sibling group; line with colored confidence band reflects local estimator (loess) fit over individual infants' vocabulary at each month. Triangles indicate mean with bootstrapped CIs computed over each month's data. Points (jittered horizontally) show individual infants' vocabulary size at each month. Y-axis utilizes log-transformed vertical spacing for visual clarity.

input and vocabulary are not tested in the same model.

We first tested for the effect of infant sex and age on our two input measures. Models of nouns in the input were a better fit when both age and sex were included alongside a random effect of infant, relative to a model with only this random effect (ps < .042), while age (p = .002) but not sex (p = .172) improved model fit for object presence. Significant variables will be added to the relevant models below. See Supplementary Data for correlations between input and maternal education/age.

Table 4

Data summary of our two input measures and reported vocabulary size at 18 months.

	No siblings		1 sibling		2+ siblings	
Variable	Mean	SD	Mean	SD	Mean	SD
% object presence in input, 10-17 months	0.67	0.15	0.56	0.15	0.46	0.18
N Input utterances, 10-17 months	180.63	124.85	184.43	84.76	100.19	52.80
Productive Vocabulary 18m (CDI)	58.89	60.76	92.64	111.42	13.00	9.49

Parental Input. Mothers provided the largest proportion of the infants' overall input across the sample (81%, M=136.71 object words, Mdn=115, SD=112.16). Fathers accounted for an average of 13% (M=18.81, Mdn=0, SD=43.37), while infants with siblings received around 6% of their input from their brothers and sisters (M=15.84, Mdn=10, SD=19.19). See Table 4 and Figure 2. We tested overall quantity of input (aggregated across mothers, fathers, and siblings) in our model alongside age, sex and subject, as reported above, and a significant effect was found for the effect of sibling group ($\chi^2(2) = 14.66$, p=.001).

We then ran post-hoc tests to compare mean amount of input across sibling groups; 294 these showed a significant difference in average input received between infants with one 295 sibling versus those with two or more siblings (W=4, p<.001; Bonferroni-corrected 296 p-threshold = .025 for all reported Wilcoxon tests) while amount of input did not differ 297 between infants with no siblings and those with one sibling (W=117, p=.506). On average, 298 in any given hour-long recording, infants with no siblings heard 0.50 fewer object words in 299 their input than those with one sibling, and 80 more than those with two or more siblings. 300 Infants with one sibling heard 80 more object words than those with two or more siblings. 301 See Table 4 for overall group differences (M and SD) in amount of input.

Next, we tested how much of that input came from siblings (for infants who had them).

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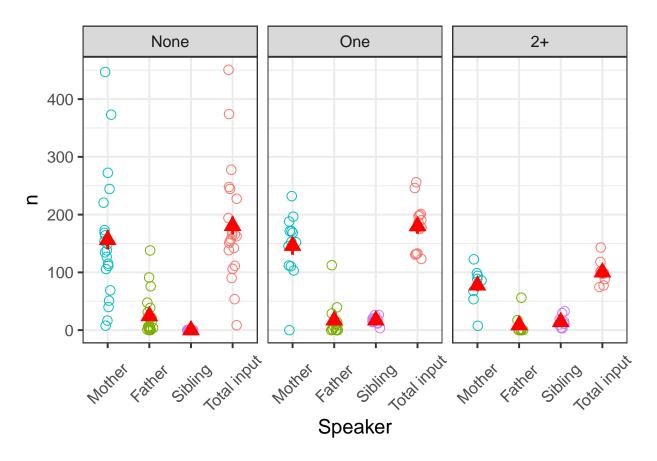


Figure 2. Mean number of words produced by Mothers, Fathers and Siblings, as well as total family input (mother + father + sibling(s)), across sessions recorded between 10-17 months. Circles represent values for individual infants; red triangles show group means.

Overall, for infants who had siblings, at least one other child was present in 70.72% of recordings (n = 263). Wilcoxon Rank Sum tests showed no difference between the amount of sibling input received by infants with one sibling compared with those with two or more siblings (W=43, p=.316). Looking at caregivers individually, infants with two or more siblings heard significantly less input from their mothers than those with one sibling (W=12, p=.001), while there was no difference between those with one vs. no siblings (W=126, p=.727). Finally, amount of paternal input did not differ between groups (one vs. none: W=161, p=.395; one vs. 2+: W=45, p=.382).

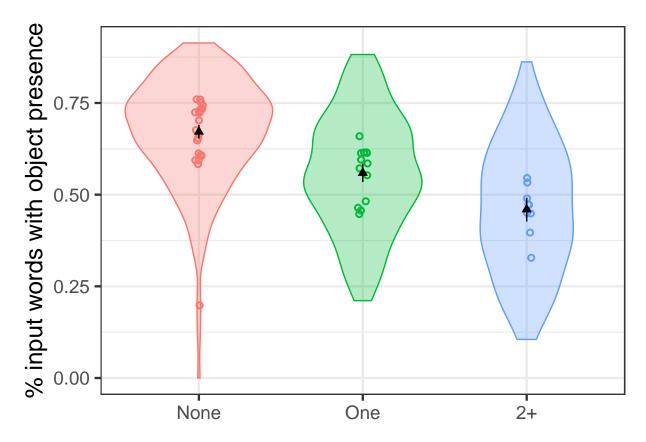


Figure 3. Proportion of input words produced with object presence across sibling groups. Error bars and black triangles show 95% CIs and mean proportion of object presence across sibling groups. Dots indicate mean proportion of object presence per infant, collapsing across age and jittered horizontally for visual clarity.

Object presence. On average, 59% of annotated utterances included a referent that 312 was present and attended to by the infant (Mdn=0.59, SD=0.14). See Table 4. Consistent 313 with our hypothesis that infants with more siblings would hear fewer words in referentially 314 transparent conditions (i.e. they would experience lower object presence) than those with fewer siblings, our models reveal a significant effect for sibling group on object presence 316 $(\chi^2(2) = 25.17, p < .001)$. See Figure 3. Infants with no siblings experienced 20% more object presence in their input than those with two or more siblings, and 10% more than 318 those with one sibling. Post-hoc comparisons revealed significant between-group differences: 319 infants with no siblings experienced significantly more object presence than those with one 320

sibling (W=224, p=.001; Bonferroni-corrected p-threshold = .025). Likewise, infants with one sibling experienced significantly more object presence those with two or more siblings (W=20, p=.009).

324 Discussion

We investigated the nature of infant language development in relation to number of 325 children in the household. Previous research found a delay in lexical acquisition for 326 later-born infants (Fenson et al., 1994; Hoff, 2006), with differences in input across birth 327 order reported as a root cause. Our results add several new dimensions to this, by testing for 328 differences across more vs. fewer older siblings, and by looking at input during child-centered 320 home recordings. Infants with more siblings were reported to say fewer words by 18 months, 330 heard fewer nouns from their parents, and were less likely to be attending to an object when 331 hearing its label. 332

Importantly, and in contrast with some previous research (Hoff-Ginsberg, 1998;
Oshima-Takane & Robbins, 2003), infants with one sibling showed no decrement in lexical
production and minimal reduction in input in comparison to first-born infants. That is, our
results suggest that simply having a sibling does not contribute to input or vocabulary
differences across children (as measured here), while having more than one siblings seems to
do so. Indeed, infants with zero and one sibling had similar results for productive vocabulary,
and parental noun input overall (though not object presence). In contrast, infants with two
or more siblings said fewer words, and also heard fewer input words overall, with
proportionally less object co-presence, compared with their peers.

The sibling effect. When we considered the effect of sibling status – that is,
whether or not infants had any siblings, disregarding specific sibling number – our findings
showed that having siblings made no difference to infants' lexical production capacities. This

contrasts with Hoff-Ginsberg (1998), who found that, by 18 months, laterborns exhibit lower 345 language skills. However, Oshima-Takane and colleagues (1996) found no overall differences 346 between first- and second-born children across a range of language measures taken at 21 347 months. Our results suggest that considering sibling quantity may be a more sensitive way to 348 reveal demographic effects than their (coarser-grained) first- vs. later-born status. We find 349 that more older siblings a child had, the lower their reported productive vocabulary at 18 350 months. This adds to findings from Fenson and colleagues (1994), who found a weak but 351 significant negative correlation between birth order and production of both words and 352 gestures. Controlling for age, our model showed that for each additional older sibling, infants 353 produced more than 30% fewer words by 18 months. 354

While infants with more siblings heard less input speech overall, having one sibling did 355 not significantly reduce the number of nouns in an infant's input. This is in direct contrast 356 with reports from the literature; Hoff (2006) states that "when a sibling is present, each child 357 receives less speech directed solely at...her because mothers produce the same amount of 358 speech whether interacting with one or two children" (p.67, italics added). While this does 359 not appear to be the case in the present dataset, it may be due to the circumstances of the 360 home-recorded data: while siblings were present in many of the recordings (70.72% of 361 recordings in which the target child had a sibling), given the focus of the data collection, 362 parents may have had a tendency to direct their attention - and consequently their linguistic 363 input - more towards the target child; of course our samples also differed in other ways 364 (e.g. sociocultural context) that may have influenced the results as well. Alternatively, our results may diverge from those of Hoff (2006) due to the nature of our input measure, which only took nouns into account. That said, we find this alternative explanation unlikely given work by Bulgarelli and Bergelson (2020) showing that nouns are a reliable proxy for overall 368 input in this dataset, suggesting that this measure provides an appropriate representation of 369 overall input directed at the target child. 370

In contrast to the other results, our analysis of object presence showed a more linear 371 'sibling effect.' In this case, even having one sibling led to fewer word-object pairs presented 372 in the input. Presence of a labeled object with congruent input speech has been found to 373 support early word learning across several studies. For instance, Bergelson and Aslin (2017) 374 combined analysis of this home-recorded data at six months with an experimental study to 375 show that word-object co-presence in naturalistic caregiver input correlated with 376 comprehension of nouns (tested using eve-tracking.) Relatedly, Gogate and colleagues (2000) 377 propose that contingent word production supports the learning of novel word-object 378 combinations, with "multimodal motherese" - whereby a target object word is produced in 379 movement or touch-based synchrony with its referent - supporting word learning. More 380 broadly, lower rates of referential transparency for common non-nouns like hi and uh-oh have 381 been proposed to potentially explain why these words are learned later than common concrete nouns (Bergelson & Swingley, 2013). While the present results on object presence 383 don't speak directly to word learning, they do suggest that this potentially helpful learning support is less available for children with more siblings. 385

Siblingese as a learning opportunity? We also found that infants with siblings 386 did not hear much speech from their older brothers and sisters. Similar findings are reported 387 in a lab-based interaction study by Oshima-Takane and Robbins (2003), who found that 388 older siblings rarely talked directly to the target child; instead, most input from siblings was 380 overheard speech from sibling-mother interactions. One possibility raised by these results is 390 that perhaps parents are able to compensate or provide relatively similar input and learning 391 support for one or two children, but once children outnumber parents, this balancing act of attention, care, and time, becomes unwieldy. While the current sample is relatively limited and homogeneous in the family structures and demographics it includes, future work could fruitfully investigate this possibility by considering whether (controlling for other potential 395 contributors like SES, Hoff-Ginsberg, 1998) the presence of more caregivers (whether parents, 396 relatives, or other adults) helps foster language development. 397

Alternatively, second-borns might 'even out' with children with no siblings due to a 398 trade-off between direct attention from the caregiver and the possibility of more 399 sophisticated social-communicative interactions. For these infants there is still ample 400 opportunity to engage with the mother in one-to-one interactions, allowing a higher share of 401 her attention than is available to third- or later-borns. Furthermore, triadic interactions can 402 benefit the development of a number of linguistic and communication skills (Barton & 403 Tomasello, 1991; Dunn & Shatz, 1989). Second-borns may also benefit from overheard 404 speech in their input, supporting the acquisition of nouns and even more complex lexical 405 categories (Floor & Akhtar, 2006; Oshima-Takane et al., 1996). For infants with one sibling, 406 the benefits of observing/overhearing interactions between sibling and caregiver, as well as 407 the possibility for partaking in such interactions, may outweigh the decrease in some aspects 408 of the input (i.e., in our data, only observed in object presence). Having more than one sibling may throw this off-balance. 410

Importantly, the present results make no claims about eventual outcomes for these 411 children: generally speaking, regardless of sibling number, all typically-developing infants 412 reach full and fluent language use. Indeed, some research suggests that sibling effects, while 413 they may be clear in early development, are not always sustained into childhood; e.g. twins 414 are known to experience a delay in language development into the third year, but are quick 415 to catch up thereafter (Dales, 1969; Tomasello, Mannle, & Kruger, 1986). This demonstrates 416 the cognitive adaptability of early development, which brings about the acquisition of 417 language across varying and allegedly 'imperfect' learning environments. Infants' capacity to 418 develop linguistic skills from the resources that are available to them – whether that is 419 infant-directed object labels or overheard abstract concepts – highlights the dynamic and adaptable nature of early cognitive development, and a system that is sufficiently robust to 421 bring about the same outcome across populations.

Of course, the 'success' of early language development is defined by Limitations. 423 how success is measured. Here we chose word production as our measure of linguistic 424 capability; we did not consider other equally valid measures such as language comprehension 425 or early social-interaction skills. Similarly, our input measures focused on nouns; other lexical 426 classes may reveal different effects, though they are generally far sparser in production until 427 toddlerhood. There is also some imbalance in group sizes across our data; our sample was 428 not pre-selected for sibling number, and so group sizes are unmatched across the analysis. 429 Including a larger number of infants with 2+ siblings may have revealed a different pattern 430 of results. Finally, more work across wider and larger populations is necessary to unpack the 431 generalizability of the present results. Our sample is reflective of average household sizes in 432 middle-class families across North America and Western Europe (Office for National 433 Statistics, 2018; United States Census Bureau, 2010), but it is not unusual in some communities and parts of the world for households to include between three and six children 435 on average (Institute for Family Studies & Wheatley Institution, 2019). Adding to this, it is also necessary to consider cross-cultural differences in the way children are addressed by 437 their parents, other caretakers, and other children [Casillas, Brown, & Levinson (2019); 438 (schneidman_language_2012?); bunce_cross-cultural_2020]. For instance, Bunce and colleagues [-bunce_cross-cultural_2020] find relatively similar rates of target child directed 440 speech across US, Canadian, Argentinian, UK, Papuan and Mayan samples, some differences 441 in who the input comes from, and large effects of number of talkers present. These results 442 suggest that caution is advisable before generalizing the current results to any other 443 socio-cultural contexts, but also leave open exciting open question regarding what variability 444 in experiences do-or don't-change about early language interaction and development. 445

Conclusion. In conclusion, our results with English-learning infants in the US
support prior findings from the literature showing that later-born infants have slower lexical
acquisition than their first-born peers. However, we highlight an important difference from
previous findings, namely that in the present sample, second-born infants show no such

effect, while infants with more than two siblings have significantly smaller productive
vocabularies at age 18 months. We related this directly to the infants' input over a period of
one year, showing differences as a function of sibling for both overall noun input and object
presence. We look forward to future studies considering the granularity of more versus fewer
siblings, and how this relates to language abilities over the course of development.

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