- Analysing the effect of sibling number on input and output in the first 18 months
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5 Abstract

- The 'sibling effect' has been widely reported in studies examining a breadth of topics in
- the academic literature, suggesting firstborn children are advantaged across a range of
- 8 cognitive, educational and health-based measures compared with their later-born peers.
- 9 Expanding on this literature using naturalistic home-recorded data and parental
- vocabulary report, we find that early language outcomes vary by number of siblings.
- Specifically, we find that children with two or more older siblings but not one had
- smaller vocabularies at 18 months, and heard less input from caregivers across several
- measures. We discuss implications regarding what infants experience and learn across a
- 14 range of family sizes in infancy.
- 15 Keywords: Siblings, Lexical Development, Input Effects, Language Acquisition
- Word count: X

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Analysing the effect of sibling number on input and output in the first 18 months

Many studies assume a theoretical "optimum" environment for early language 18 development, whereby the input is tailored to a single infant's needs, changing over time as 19 language capacity develops (e.g. Soderstrom, 2007; Stern, Spieker, Barnett, & MacKain, 20 1983). However, for many infants and for many reasons, language acquisition occurs amid 21 various domestic and social factors that can influence the learning environment, including the presence of older siblings in the home (Fenson et al., 1994). According to the United 23 States Census Bureau (2010), around one third of children are born into households with at least one other infant present, and one in every five infants is acquiring language in a household shared with two or more other children. Similar statistics are reported for British infants (Office for National Statistics, 2018), where the average household has 1.75 children, and 15% of households have three children or more. In this paper, we consider the role of siblings in the early language environment of English-learning infants. We use naturalistic home-recorded data to measure input in earlier- and later-born infants in relation to their lexical development over the first 18 months of life. 31

Prior research suggests that infants born to households with older children may be
slower to learn language. Fenson and colleagues (1994) found that by 30 months of age,
children with older siblings performed worse than those with no siblings across measures of
productive vocabulary, use of word combinations, and mean length of utterance (MLU).
This 'sibling effect' may be manifested in input differences between first- and later-born
children: some research finds that infants with older siblings hear less speech aimed
specifically at them, and what they do hear is understood to be linguistically less
supportive of early language development (Hoff-Ginsberg, 1998; Oshima-Takane &
Robbins, 2003). Contrastingly, some studies have noted linguistic advantages for
later-borns, who may have stronger social-communicative skills (Hoff, 2006), better
understanding of pronouns (Oshima-Takane, Goodz, & Derevensky, 1996), and better

conversational abilities (Dunn & Shatz, 1989). Overall, while the particulars differ across studies, prior work suggests that the presence of siblings in the home leads to differences in infants' early linguistic experience.

Numerous studies have attempted to better understand how siblings affect the 46 language development trajectory, with comparisons of language acquisition across first- and later-borns, and analyses of mothers' input in dyadic (infant + mother) and triadic (infant + mother + older sibling) situations. Findings are mixed, but overall two general conclusions can be drawn. First, analyses consistently show that infants with older siblings 50 generally have slower vocabulary development (Berglund, Eriksson, & Westerlund, 2005; 51 Fenson et al., 1994; Pine, 1995; Zambrana, Ystrom, & Pons, 2012), but this is often marginal, and typical of the earliest stages of language learning. Hoff-Ginsberg (1998) shows first-borns to have better lexical and syntactic skills up until 2:5, but later-born infants had better conversational abilities during the same time-period. Some of these differences may relate to insufficient power to detect relatively small effects or simultaneous contributing factors. For instance, using a large longitudinal dataset of french-learning 2-5 year olds, Havron and colleagues (2019) find no effect of age gap between siblings but lower standardized language scores in children with older brothers (but not sisters) relative to those without siblings, based on parental report and direct battery assessments.

The second general finding pertains to sibling-related differences in the early linguistic environment: infants with no siblings receive more input overall, and this more closely reflects what is typically considered to be 'high quality' input in the extant literature. Indeed, the very presence of a sibling in the linguistic environment changes the way language is used. When siblings are present (i.e. triadic interactions), mothers' input is more focused on regulating behaviour, as opposed to the language-focused speech that is common in dyadic contexts (Oshima-Takane & Robbins, 2003). Reports show that MLU is longer in the input of first-born infants (Hoff-Ginsberg, 1998; but see also Oshima-Takane & Robbins, 2003 for a comparison of dyadic and triadic contexts), who also hear more

questions directed at them than later-borns. Both Jones and Adamson (1987) and
Oshima-Takane and Robbins (2003) report no difference between the overall number of
word types produced by mothers in dyadic and triadic settings, but the proportion of
speech directed at the target infant is drastically reduced when input is shared with
siblings.

As Hoff (2006) explains, infants with siblings have less experience of speech directed at them, but they do have an advantage over their first-born peers in that they are subject to more overheard speech. Indeed, the input of first-borns may be explicitly tailored to their needs, but equally this means it might be less varied. Barton and Tomasello (1991) show that by as early as 19 months, infants with siblings are already able to take part in triadic conversations, which were almost three times longer than dyadic conversations. The authors suggest that the presence of siblings may shift the learning context, and facilitate infants' participation in communicative interactions: infants are under less pressure to participate in a triadic interaction, meaning the conversation can continue even when the infant is unable to respond. As a result, the infants in Barton and Tomasello's study took more conversational turns in triadic interactions than dyadic ones.

There thus may be a trade-off in development between highly supportive one-to-one input from a caregiver (cf. Ramírez-Esparza, García-Sierra, & Kuhl, 2014) and the potential benefits drawn from communicating with a sibling. In the present study, we test the extent to which having more versus fewer siblings in the home environment may lead to differences in vocabulary development and the early linguistic environment over the course of the first 18 months of life.

In analyzing infants' lexical development in relation to the presence of older siblings in their household, the present work expands on the extant literature in two key ways.

First, prior work generally considered birth order as a binomial factor (i.e. comparing first-born infants with second-borns, e.g. Oshima-Takane & Robbins, 2003), or

'later-borns' (e.g. Hoff-Ginsberg, 1998), potentially missing graded effects. Instead of this approach, we consider the number of siblings, i.e. how having more versus fewer siblings is 97 linked to an infant's lexical development and their early linguistic environment. Second, much of the existing literature in this area is drawn from questionnaire data or brief interactions recorded in the lab (but see Dunn & Shatz, 1989 for a study of naturalistic 100 home-recorded data), rather than naturalistic day-to-day interactions in the home. In 101 contrast, we analyze an existing corpus of daylong audio- and hour-long video-recordings in 102 concert with vocabulary checklists. Based on work summarized above, we expect that both 103 the language environment and infants' early vocabulary will vary as a function of how 104 many older siblings they have. 105

106 Hypotheses

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Research has already shown that early lexical development is more advanced among first-born infants (e.g. Hoff-Ginsberg, 1998). We expect to see the same effect in our data, but we hypothesize that the closer granularity of this analysis will show a gradient decline in infants' lexical abilities in relation to an increasing number of siblings. With regard to the infants' linguistic environment, we hypothesize that increased sibling number will have a negative effect on factors of the early input that are known to support language development. To test this, we adopt two input measures, established in the literature as being important for early language learning. The measures and our specific predictions are as follows:

1) Amount of input will be lower for children with more siblings. Following previous studies that show infants with siblings to receive less speech directed at them (Jones & Adamson, 1987; Oshima-Takane & Robbins, 2003), we expect to see the same effect in our sample. Given the link between amount of one-to-one input from the caregiver and vocabulary size (Ramírez-Esparza et al., 2014), we expect that infants

- who hear less input overall (i.e. by our predictions, those with more siblings) will have smaller productive vocabularies.
- 2) Amount of object presence (word + object co-occurrence, e.g. mother says 'cat' when there is a cat in the room) will decrease as sibling number increases. As caregivers' attention is drawn away from one-to-one interactions with the infant, we expect there to be less opportunity for contingent talk and joint attention. The co-occurrence of words alongside their referents is thought to contribute to the earlier learning of nouns over verbs (Bergelson & Swingley, 2013), as it supports the word learning process through the concrete mapping of word to referent (Gleitman, Cassidy, Nappa, Papafragou, & Trueswell, 2005). We thus expect that increased sibling number will be associated with less object presence in caregiver speech, and subsequently a smaller productive vocabulary.

133 Methods

We analyze data from the SEEDLingS corpus (Bergelson, Amatuni, Dailey,
Koorathota, & Tor, 2019), a longitudinal set of data incorporating home recordings,
parental reports and experimental studies from the ages of 0;6 to 1;6. The present study
draws on the parental report data to index child vocabulary size, and annotations of
hour-long home video recordings, taken on a monthly basis during data collection, to index
input. We also ran our input analysis using day-long audio recordings taken on a different
day from the video data reported below; all results were consistent with those outlined
below (see Supplementary Materials).

42 Participants

Forty-four families in New York State completed the year-long study. Infants (21 females) were from largely middle-class households; 33 mothers had attained a B.A. degree

or higher. All infants had normal birthweight with no reported speech- or hearing-relevant diagnoses. Forty-two infants were Caucasian; two were from multi-racial backgrounds.

147 Materials

Parental report data To index each child's language abilities, we draw on data 148 from vocabulary checklists [Macarthur-Bates Communicative Development Inventory, 149 hereafter CDI; Fenson et al. (1994), administered monthly from 0;6 to 1;6, along with a 150 demographics questionnaire. Because the majority of infants did not produce their first 151 word until around 0;11 according to CDI reports (M=10.70, SD=2.22)¹, we use CDI data 152 from 0:10 on-wards in our analysis. CDI production data for each month is taken as a 153 measure of the infants' lexical development over the course of the analysis period. Note that noun production in the home-recorded data correlates significantly with reported noun 155 production in the CDI checklists from the age of 12 months, as reported in a previous study by Moore and colleagues (2019). 157

Home-recorded video and audio data Every month between 0:6 and 1:5, infants 158 were video-recorded for one hour in their home, capturing a naturalistic representation of 159 each infant's day-to-day input. Infants were a hat with two small Looxcie video cameras 160 attached, one pointed slightly up, and one pointed slightly down; this allowed us to record 161 the scene from the infants' perspective. In the event that infants refused to wear the hats, 162 caregivers were the same kind of camera on a headband. Additionally, a camcorder was set 163 up in the home. On a different day in the same month, the infants were recorded for upto 164 16 hours using a LENA recorder ("LENA Research Foundation," 2018) hidden in a small 165 vest worn by the infant. Object words (i.e. concrete nouns) deemed to be directed to or 166 attended by the child were annotated by trained coders. Here we examine annotations for

 $^{^{1}}$ Note that reported word production began earlier than observed word production (i.e. in the video recordings) but this difference was not statistically significant (see Moore, Dailey, Garrison, Amatuni, & Bergelson, 2019).

speaker, i.e. who produced a word, and object presence, i.e. whether the word's referent was present and attended to by the infant.

70 Procedure

We analyzed number of siblings based on parental report in the demographics questionnaires completed at 0;6 (R: 0-4). Siblings were on average 4.05 years older than the infants in this study (Mdn days: 1477, SD: 1477, R: 0-17 years). All siblings lived in the household with the infant; one infant had XX older siblings who lived in the household XX% of the time. One family had a foster child live in the home for XX months of data collection. All siblings were older than or of the same age as the infant in question.

177 Input measures

Two input measures were derived based on the annotations of concrete nouns in this corpus, each pertaining to an aspect of the input that is established as important in early language learning: **overall household input** (how many concrete nouns does each infant hear?) and **object presence** (how much of this input is referentially transparent?). Each is described in further detail below.

Household Input reflects how many nouns infants heard in the recordings from their mother, father and (where relevant) siblings. Other speakers' input was relatively rare during video recordings, accounting for 11.42% of input overall (SD=22.82%), and is excluded from our analysis. This measure of the early language environment is based on evidence showing strong links between the amount of speech heard in the early input and

² For six infants, siblings' exact birthdates were not provided, and so age difference was estimated by subtracting the infant's age (6 months) from the sibling's age in years, as listed on the questionnaire (e.g. if a sibling was 5 years old, they were classed as being 4.5 years older than the infant).

³ Two infants in the dataset were dizygotic twins; our pattern of results holds when one of these infants is removed from the dataset.

later vocabulary size (Anderson, Graham, Prime, Jenkins, & Madigan, 2021). This analysis
considers only nouns produced by speakers in the child's environment; concrete nouns are
acquired earlier in development in English and cross-linguistically (Braginsky, Yurovsky,
Marchman, & Frank, 2019), and for this corpus of data, noun production correlates
strongly with automated adult word count estimates (Bulgarelli & Bergelson, 2020).
Higher noun count in this data thus indicates higher input across the board.

Object Presence was coded for each object word in the home recordings based on 194 whether or not the annotator determined the object in question as present and attended to 195 by the child. This is a metric of referential transparency, which has been suggested to aid in 196 learning (Bergelson & Swingley, 2013). Bergelson and Aslin (2017) found that word-object 197 co-presence in the home correlated with infants' ability to recognise the same words in an 198 eve-tracking experiment, suggesting an advantage for object labeling in word learning. This 199 is consistent with findings from Cartmill and colleagues (2013), who found that more 200 referentially-transparent interactions with the caregiver (as judged by adult speakers 201 observing videos where target words were blanked out) predicted larger vocabulary size at 202 54 months. Indeed, presence of the labelled object decreases the ambiguity of the learning 203 environment (Yurovsky, Smith, & Yu, 2013), and may be a crucial component of 204 supportive contingent talk (McGillion, Pine, Herbert, & Matthews, 2017). 205

In the following analyses, we consider infants' productive vocabulary alongside our
two input measures – amount of household input and extent of object presence in the input
– as a function of sibling number. These measures index both input quality and quantity
(though we make no distinction between quality/quantity of input here), and will be
analysed in relation to infants' productive vocabulary (all word types included) as our
dependent measure. Since the raw data are highly skewed, log-transformed data⁴ and/or
proportions are used for statistical analysis. All figures display non-transformed data for

⁴ 1 was added to the raw infant production data of all infants before log-transformation to retain infants with vocabularies of 0.

Table 1
Sibling number by female and male infants.

n Siblings	Female	Male	Total
0	9	12	21
1	7	6	13
2	2	3	5
3	2	0	2
4	0	2	2
Total	20	23	43

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214 Results

Vocabulary development was highly variable across the 44 infants. By 18 months, 2 infants produced no words, while mean productive vocabulary size was 60.28 words (SD=78.31, Mdn=30.50). One female infant had a substantially larger reported vocabulary (3SDs above the mean monthly vocabulary score) between 1;1 and 1;6 and was classed as an outlier. We removed her from our data, leaving 43 infants (20 females) in the present analysis. Infants had one sibling on average (M=0.86, Mdn=1, SD=1.09). See Table 1.

221 Model structure for fixed and random effects

All reported models were generated in R (R Core Team, 2019) using the *lmerTest*package to run linear mixed-effects regression models (Kuznetsova, Brockhoff, &
Christensen, 2017). P-values were generated by likelihood ratio tests resulting from nested
model comparison. All models include infant as a random effect. All post-hoc tests are

two-sample, two-tailed Wilcoxon Tests; given that all of our variables differed significantly 226 from normal by Shapiro tests, we opted to run non-parametric tests for all post-hoc 227 comparisons. Where multiple post-hoc comparisons are run on the same dataset, 228 Bonferroni corrections are applied with an adjusted threshold of a=0.03, accounting for two 229 between-group comparisons (no siblings vs. one sibling, one sibling vs. multiple siblings; see 230 below). Our dataset included 12 hours of home-recorded video data and 12 vocabulary 231 questionnaires per child (as well as 12 days of audio recordings). While this data provides a 232 substantial representation of each child's early language environment and development, we 233 acknowledge that it is nevertheless limited in terms of group size, and means we cannot 234 effectively combine analyses of sibling number alongside other measures such as sex. 235

Before considering sibling status, we first modeled infants' productive vocabulary at 236 18 months (taken from CDI questionnaires) as a function of age, sex, and mother's 237 education. There was no effect of sex $(p=.632)^5$, and no correlation with mothers' 238 education level (across five categories from High School to Doctorate; r = -0.01, p = .139). 239 As expected, age had a significant effect on productive vocabulary (p < .001), and so we 240 include age as a fixed effect in our models. Because we expected that maternal age and 241 education might have an effect on both sibling number and infant productive vocabulary, 242 we ran further correlations to test these variables. There was no correlation between 243 mother's education and number of siblings (r = -0.01, p = .928), and a marginal positive 244 correlation between mother's age and number of siblings (Spearman's r = 0.28, p = .069); 245 older mothers tended to have more children. However, no correlation was found between 246 mothers' age and productive vocabulary at 18 months (r = -0.04, p = .822). 247

⁵ Though note that Dailey and Bergelson (under review) find a sex difference in the number of word types produced in the home-recorded data between 6 and 17 months.

Effect of siblings on infants' productive vocabulary

We next modeled the effect of siblings on reported productive vocabulary. We tested
three variables representing the sibling effect: a binary variable (0 vs. >0 siblings),
aggregated groups (None vs. One vs. 2+), and discrete sibling number (0 vs. 1 vs. 2 vs. 3
vs. 4 siblings) using the following nested model structure, where (1) is the baseline model,
and (2) includes siblings as the variable of interest:

1. vocabulary size (log-transformed) \sim age (months) + (1|subj)

Table 2

255 2. vocabulary size (log-transformed) \sim siblings [binary, group or discrete] + age (months) + (1|subj)

In our sample, the simple fact of having siblings (i.e. as a binary variable) did not affect reported CDI vocabulary size, while both discrete sibling number and sibling group did. See Table 2.

Output from likelihood ratio tests
comparing regression models that predict
the effects of sibling number (binary,
grouped and discrete variables) on
vocabulary size. Month was included in
each model as a fixed effect; subject was

included as a random effect.

Model	Df	Chisq	p value
0 vs. >0 siblings	1.00	2.27	0.13
Sibling group	2.00	7.96	0.02
Sibling number	1.00	6.24	0.01

Table 3

Full model output from linear mixed effects regression model comparing language

development over time in relation to sibling group. Age in months was included as a fixed

effect; subject was included as a random effect.

Effect	Estimate	Std. Error	df	t value	p
Intercept	0.70	0.20	68.30	3.43	0.001
SibGroupOne	-0.05	0.29	43.10	-0.17	0.863
SibGroup2+	-0.94	0.33	43.88	-2.83	0.007
month11	0.38	0.14	321.34	2.69	0.007
month12	0.77	0.14	321.64	5.39	< 0.001
month13	1.07	0.14	321.58	7.60	< 0.001
month14	1.39	0.14	321.45	9.93	< 0.001
month15	1.69	0.14	321.54	11.96	< 0.001
month16	2.03	0.14	321.54	14.40	< 0.001
month17	2.45	0.14	321.79	16.96	< 0.001
month18	2.82	0.15	322.13	19.07	< 0.001

Having more siblings was associated with a smaller vocabulary size over the course of
early development. This is consistent with previous findings (Hoff-Ginsberg, 1998; Pine,
1995). Moreover, for each additional sibling, infants were reported to have acquired 30.59%
fewer words. The 'sibling effect' is thus present in our data. The grouped sibling variable
(0 vs. 1 vs. 2+) was selected as our measure of siblings as it allows analysis across more
equal group sizes. However, note that all regression reported below were consistent when
discrete sibling number was analysed (see Supplementary Data).

According to CDI reports, infants with one sibling acquire only 5.07% fewer words
than firstborns over the course of our analysis, while infants with two or more siblings

produce 93.73% fewer words. See Tables 3 and 4, and Figure 1. Post-hoc Wilcoxon Rank
Sum tests comparing reported productive vocabulary at 18 months revealed significantly
larger vocabularies for infants with one sibling compared to those with two or more siblings (W=5, p=.004), but no difference between infants with one sibling and those with no
siblings (W=79.50, p=.631).

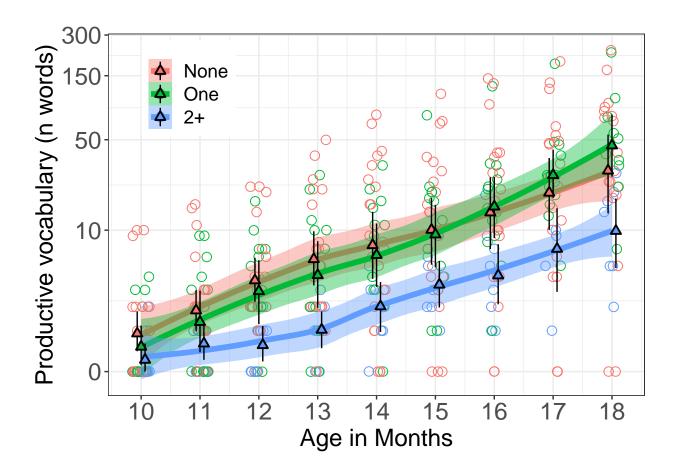


Figure 1. Reported productive vocabulary acquisition (CDI) over time. Colors denote sibling group; line with colored confidence band reflects local estimator (loess) fit over individual infants' vocabulary at each month. Triangles indicate mean with bootstrapped CIs computed over each month's data. Points (jittered horizontally) show individual infants' vocabulary size at each month. Y-axis utilizes log-transformed vertical spacing for visual clarity.

Effect of siblings on infants' input

Having established that infants' productive vocabulary varied as a function of how many siblings they had, we turn to our input measures to test whether input varied by a child's sibling status. To keep relatively similar Ns across groups we used the 0 vs. 1 vs. 2+ siblings division. That said, with the exception of household input (see Supplementary Data) all reported model outcomes hold if we model discrete sibling number as a fixed effect instead.

As with our previous analysis, we first modeled infants' input (maternal input only)
as a function of age, sex and maternal education. This time, there was no effect for age,
nor sex or maternal education (ps all>.260) on the amount of input produced by mothers.
We therefore excluded all three variables from our models.

Table 4

Data summary of all three input variables and reported vocabulary size at 18 months.

	No siblings		1 sibling		2+ siblings	
Variable	Mean	SD	Mean	SD	Mean	SD
% object presence in input	0.67	0.15	0.56	0.15	0.46	0.18
N Input utterances, 10-17 months	180.63	124.85	184.43	84.76	100.19	52.80
Productive Vocabulary 18m (CDI)	58.89	60.76	92.64	111.42	13.00	9.49

Parental Input. Mothers provided the largest proportion of the infants' overall input across the sample (80%, M=146.10 object words, Mdn=125, SD=119.97). Fathers accounted for an average of 14% (M=22.13, Mdn=0, SD=48.31), while infants with siblings received around 6% of their input from their brothers and sisters (M=16.18, Mdn=11, SD=18.51). See Table 4 and Figure 2. We tested overall quantity of input (aggregated across mothers, fathers, and siblings) in our model, and a significant effect was found ($\chi^2(2) = 18.48, p < .001$). We then ran post-hoc tests to compare mean amount of input

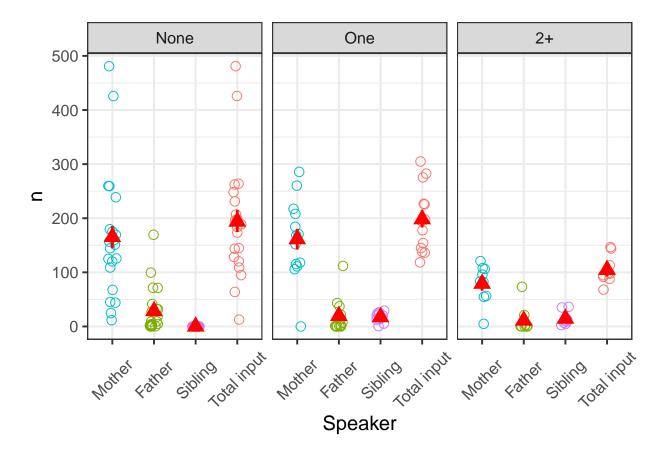


Figure 2. Mean number of words produced by Mothers, Fathers and Siblings, as well as total family input (mother + father + sibling(s)), across sessions recorded between 10-17 months. Circles represent values for individual infants; red triangles show group means.

across sibling groups; these showed a significant difference in average input received between infants with one sibling versus those with two or more siblings (W=7, p<.001, a=0.03) while amount of input did not differ between infants with no siblings and those with one sibling (W=120, p=.576). On average, in any given hour-long recording, infants with one siblings heard 4 **more** object words in their input than those with no sibling, and 89 more than those with two or more siblings. Infants with one sibling heard 94 more object words than those with two or more siblings.

Next, we tested how much of that input came from siblings (for infants who had them). Overall, for infants who had siblings, at least one other child was present in 72.16%

of recordings (n = 176). Wilcoxon Rank Sum tests showed no difference between the 301 amount of sibling input received by infants with one sibling compared with those with two 302 or more siblings (W=40, p=.235), contrasting with predictions set out in our first 303 hypothesis. Looking at caregivers individually, infants with two or more siblings heard 304 significantly less input from their mothers than those with one sibling (W=15, p=.003), 305 while there was no difference between those with one vs. no siblings (W=126, p=.727). 306 Finally, amount of paternal input did not differ between groups (one vs. none: W=152, p 307 = .606; one vs. 2+: W=42, p=.296). 308

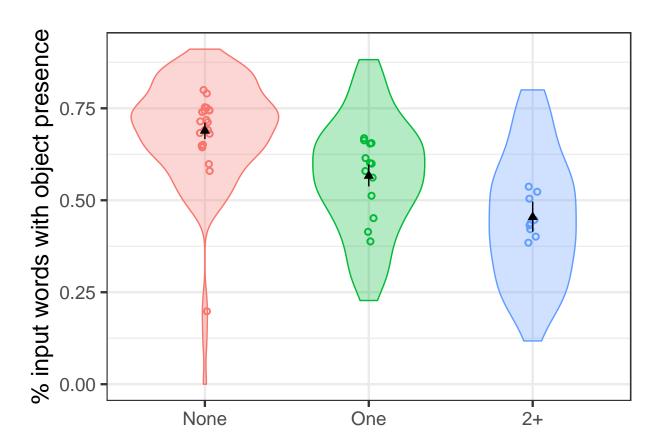


Figure 3. Proportion of input words produced with object presence across sibling groups. Error bars and black triangles show 95% CIs and mean proportion of object presence across sibling groups. Dots indicate mean proportion of object presence per infant, collapsing across age and jittered horizontally for visual clarity.

Object presence. On average, 60% of utterances were produced in the presence of 309 the relevant object (Mdn=0.60, SD=0.12). We hypothesized that infants with more 310 siblings would hear fewer words in referentially transparent conditions (i.e. they would 311 experience lower object presence) than those with fewer siblings. Indeed, modelling the 312 quantity of object present tokens that infants heard, we find a significant effect for sibling 313 group on object presence ($\chi^2(2) = 26.09$, p < .001). See Figure 3. Infants with no siblings 314 experienced 23% more object presence in their input than those with two or more siblings, 315 and 12% more than those with one sibling. Post-hoc comparisons revealed significant 316 between-group differences: infants with no siblings experienced significantly more object 317 presence than those with one sibling (W=234, p<.001, a=0.03). Likewise, infants with 318 one sibling experienced significantly more object presence those with two or more siblings 319 (W=20, p=.009). See Table 4.

321 Discussion

We investigated the nature of infant language development in relation to number of 322 children in the household. Previous research found a delay in lexical acquisition for 323 later-born infants (Fenson et al., 1994; Hoff, 2006), with differences in input across birth 324 order reported as a root cause. Our results add several new dimensions to this, by testing 325 for differences across more vs. fewer older siblings, and by looking at input during 326 child-centered home recordings. Infants with more siblings were reported to say fewer 327 words by 18 months, heard fewer nouns from their parents and siblings, and experienced 328 less 'object co-presence' when hearing them. 329

Importantly, and in contrast with some previous research (Hoff-Ginsberg, 1998;
Oshima-Takane & Robbins, 2003), infants with one sibling showed no delay in lexical
production and minimal reduction in input in comparison to first-born infants. That is, our
results suggest that simply having a sibling does not contribute to input or vocabulary
differences across children (as measured here), while having more than one siblings seems

to do so. Indeed, infants with zero and one sibling had similar results for productive vocabulary, and parental noun input overall, but not object presence, which we return to below. In contrast, infants with two or more siblings said fewer words, and also heard fewer input words with proportionally less object co-presence, compared with their peers.

When we considered the effect of sibling status – that is, whether or not infants had 339 any siblings, disregarding specific sibling number – our findings showed that having siblings made no difference to infants' lexical production capacities. This contrasts with Hoff-Ginsberg (1998), who found that, by 18 months, laterborns exhibit lower language skills. However, Oshima-Takane and colleagues (1996) found no overall differences between 343 first- and second-born children across a range of language measures taken at 21 months. 344 Our finer-grained results suggest a greater role for *sibling quantity* over first- vs. later-born 345 status. The more older siblings a child had, the lower their reported productive vocabulary 346 at 18 months. This adds to findings from Fenson and colleagues (1994), who found a weak 347 but significant negative correlation between birth order and production of both words and 348 gestures. Controlling for age, our model showed that for each additional older sibling, 340 infants produced more than 30% fewer words by 18 months. While infants with more 350 siblings heard less input speech overall, having one sibling did not significantly reduce the 351 number of nouns in an infant's input. This is in direct contrast with reports from the 352 literature; Hoff (2006) states that "when a sibling is present, each child receives less speech 353 directed solely at...her because mothers produce the same amount of speech whether 354 interacting with one or two children" (p.67, italics added). While this does not appear to 355 be the case in the present dataset, it may be due to the circumstances of the home-recorded data: while siblings were present in many of the recordings (72.16% of recordings in which the target child had a sibling), given the focus of the data collection, parents may have had a tendency to direct their attention - and consequently their 359 linguistic input - more towards the target child. Alternatively, our results may diverge from 360 those of Hoff (2006) due to the nature of our input measure, which only took nouns into 361

account. However, Bulgarelli and Bergelson (2020) show that nouns are a reliable proxy for overall input in this dataset, thus suggesting that this measure provides an appropriate representation of overall input directed at the target child.

Moreover, infants with siblings did not hear much speech from their older brothers 365 and sisters. This is contrary to our hypothesis, as having more siblings did not predict 366 more sibling input. Similar findings are reported in a lab-based interaction study by 367 Oshima-Takane and Robbins (2003), who found that older siblings rarely talked directly to 368 the target child; instead, most input from siblings was overheard speech from 369 sibling-mother interactions. However, results from Havron and colleagues (2019) indirectly 370 suggest that speech from siblings may affect language development, and not necessarily in 371 a negative direction. They found that children with older brothers had lower verbal skills 372 than children with no siblings; children with older sisters did not show this effect. The 373 authors propose that this differential effect could be due sisters having positive effects on 374 language development (i.e., the effect is derived from supportive sibling input), or perhaps 375 due to brothers' additional demands on caregiver time and attention, thus directing 376 caregiver attention away from the target child (i.e., the effect is derived from changes to 377 caregiver input). We did not analyse sibling sex in our data, but future analyses could consider input speech in relation to sibling sex. 379

The 'sibling effect' was most marked in our analysis of object presence. In this case,
even having one sibling led to fewer word-object pairs presented in the input. Presence of a
labeled object with congruent input speech is known to be supportive in early word
learning: Bergelson and Aslin (2017) combined analysis of this home-recorded data with an
experimental study to show that word-object co-presence in naturalistic caregiver input
supported comprehension of nouns when tested using eye-tracking. Gogate and colleagues
(2000) state that contingent word production supports the learning of novel word-object
combinations, as "multimodal motherese" - whereby a target word is produced in
synchrony with its referent, often involving movement or touch of the object - supports

word learning by demonstrating novel word-object combinations. Indeed, lower rates of referential transparency in children's input have also been proposed to explain why common non-nouns like *hi* and *uh-oh* are learned later than concrete nouns (Bergelson & Swingley, 2013).

Object presence varied more linearly across sibling quantity, suggesting it may be a 393 less critical driver of early word production. Given that infants with one sibling heard 394 approximately the same number of object words in the input than those with no siblings, 395 input, rather than object presence, may be a more crucial factor in predicting a child's 396 vocabulary size by 18 months. Alternatively, the reduced object presence for children with 397 one sibling may have been compensated for in other ways we did not measure here, which 398 in turn resulted in the indistinguishable vocabulary difference in the 0 and 1 sibling 399 children at 18 months. 400

More generally, one possibility raised by these results is that perhaps parents are able to compensate or provide relatively similar input and learning support for one or two children, but once children outnumber parents, this balancing act of attention, care, and time, becomes unwieldy. While the current sample is relatively limited and homogenous in the family structures and demographics it includes, future work could fruitfully investigate this possibility by considering whether (controlling for other potential contributors like SES, Hoff-Ginsberg, 1998) the presence of more caregivers (whether parents, relatives, or other adults) helps foster language development.

Alternatively, second-borns might 'even out' with children with no siblings due to a
trade-off between direct attention from the caregiver and the possibility of more
sophisticated social-communicative interactions. For these infants there is still ample
opportunity to engage with the mother in one-to-one interactions, allowing a higher share
of her attention than is available to third- or later-borns. Furthermore, triadic interactions
can benefit the development of a number of linguistic and communication skills (Barton &

Tomasello, 1991; Dunn & Shatz, 1989). Second-borns may also benefit from overheard
speech in their input, supporting the acquisition of nouns and even more complex lexical
categories (Floor & Akhtar, 2006; Oshima-Takane et al., 1996). For infants with one
sibling, the benefits of observing/overhearing interactions between sibling and caregiver, as
well as the possibility for partaking in such interactions, may outweigh the decrease in
some aspects of the input (i.e., in our data, only observed in object presence). Having more
than one sibling may throw this off-balance.

Importantly, the present results make no claims about eventual outcomes for these 422 children: generally speaking, regardless of sibling number, all typically-developing infants 423 reach full and fluent language use. Indeed, some research suggests that sibling effects, while 424 they may be clear in early development, are not always sustained into childhood; e.g. twins 425 are known to experience a delay in language development into the third year, but are quick 426 to catch up thereafter (Dales, 1969; Tomasello, Mannle, & Kruger, 1986). This 427 demonstrates the cognitive adaptability of early development, which brings about the 428 acquisition of language across varying and allegedly 'imperfect' learning environments. 429 Infants' capacity to develop linguistic skills from the resources that are available to them – 430 whether that is infant-directed object labels or overheard abstract concepts – highlights the dynamic and adaptable nature of early cognitive development, and a system that is 432 sufficiently robust to bring about the same outcome across populations.

Of course, the 'success' of early language development is defined by how success is
measured. Here we chose word production as our measure of linguistic capability; we did
not consider other equally valid measures such as language comprehension or early
social-interaction skills. Similarly, our input measures focused on nouns; other lexical
classes may reveal different effects, though they are generally sparser until toddlerhood.
There is also some imbalance in group sizes across our data; our sample was not
pre-selected for sibling number, and so group sizes are unmatched across the analysis.

Including a larger number of infants with 2+ siblings may have revealed a different pattern

of results. Finally, more work across wider and larger populations is necessary to unpack 442 the generalizability of the present results. Our sample is refelective of average household 443 sizes in middle-class families across North America and Western Europe (Office for 444 National Statistics, 2018; United States Census Bureau, 2010), but it is not unusual in 445 some communities and parts of the world for households to include between three and six 446 children on average (Institute for Family Studies & Wheatley Institution, 2019). Adding to 447 this, it is also necessary to consider cross-cultural differences in the way children are 448 addressed by their parents. Casillas, Brown and Levinson (2019) found that almost all of 449 Tseltal Mayan children's input came from speech directed at other people (21 minutes per 450 hour, compared with just under 4 minutes/hour of specifically child-directed input), while 451 Shneidman and Goldin-Meadow (2012) found that 69% of speech directed at Mayan 452 children came from siblings (in comparison with 10% for children in the USA).

In conclusion, our results support the general findings from the literature showing 454 that later-born infants have slower lexical acquisition than their first-born peers. However, 455 we highlight an important difference from previous findings, namely that in the present 456 sample, second-born infants show no such effect, while infants with more than two siblings 457 have significantly smaller vocabularies at age 18 months. We related this directly to the 458 infants' input over a period of one year. Future studies should consider the granularity of 459 more versus fewer siblings, and how this relates to language abilities over the course of 460 development. 461

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