Expressing model uncertainty in Bayesian variable selection using credible sets

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Introduction

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There are a number of classical approaches: subset selection, stepwise selection, penalized maximum likelihood (Lasso, elastic net, MCP, etc.).



Bayesian variable selection

Assume a parametric model $y \sim f(x^{\gamma}, \theta)$ where x^{γ} is a subset of included variables indexed by γ ($\gamma_i = 1$ if the *i*-th variable is included and 0 otherwise).

Put a prior on γ . For example, $\gamma_i \sim \text{Bernoulli}(\pi)$, $\pi \sim \text{Be}(a, b)$ then a and b can be chosen to encourage sparsity.

This leads to a posterior distribution $p(\gamma \mid \text{data})$, which expresses our uncertainty about γ .



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Recent work on high-dimensional problems

- MCMC methods Importance Tempering (Zanella and Roberts, 2019), ASI (Griffin et al., 2021), PARNI (Liang et al., 2022)
- Stochastic search SVEN (Li et al., 2023)



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To understand the relative importance of different variables, there are summaries

- Posterior inclusion probabilities (PIPs): $p(\gamma_i \mid \mathsf{Data})$
- Maximum a posterior (MAP) model: the mode of γ | Data.
- Median model $\hat{\gamma}$ where $\hat{\gamma}_i = I(p(\gamma_i \mid \mathsf{Data}) > 0.5)$.

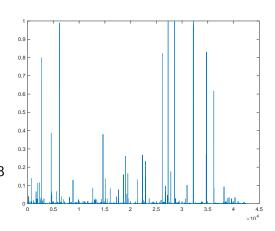


GWAS example: Systemic Lupus Erythematosus (case-control study)

chromosome 3 n = 10,995 Cases: 4,036 Controls: 6,959

$$p = 42,430$$

Median model has 13 variables





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Under an independent prior on γ , uncorrelated variables in linear models \iff independence of γ_i 's.

These relationships are due to multi-collinearity (*i.e.* correlation between variables). For example, due to linkage disequilibrium in GWAS.



Simulated example (George and McCulloch, 1997)

Linear regression example with n = 180 and p = 15.

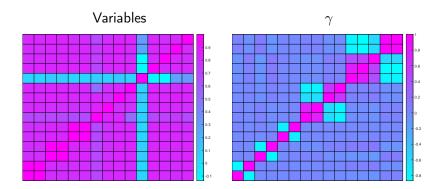
Non-zero regression coefficients are 1, 3, 5, 7, 8, 11, 12, 13.

Strong multicollinearity between variables:

- 1 and 2
- 3 and 4
- 5 and 6
- **7**, 8, 9, 10
- **11**, 12 13, 14, 15



Simulated example - Correlation





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Bayesian variable selection leads to a posterior distribution on a high-dimensional discrete space. The same is true of a lot of Bayesian nonparametric methods (e.g. clustering, feature allocation).



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There are summarisations methods for some problems (particularly clustering). How to represent uncertainty?



Credible sets

Let Γ be the set of all possible combination of variables then $A \subset \Gamma$ is a $100\alpha\%$ credible set (CS) if $p(A \mid Data) \ge \alpha$.



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Let Γ be the set of all possible combination of variables then $A \subset \Gamma$ is a $100\alpha\%$ credible set (CS) if $p(A \mid Data) \ge \alpha$.

The smallest $100\alpha\%$ CS can be found by

1. Rank the models by decreasing probability,

$$p(\gamma^{(1)} \mid \mathsf{Data}) \geq p(\gamma^{(2)} \mid \mathsf{Data}) \geq p(\gamma^{(3)} \mid \mathsf{Data}) \geq \dots \geq p(\gamma^{(2^p)} \mid \mathsf{Data})$$

2. Find the smallest K such that $\sum_{k=1}^{K} p(\gamma^{(k)} \mid \text{Data}) \ge \alpha$ then $\{\gamma^{(1)}, \gamma^{(2)}, \dots, \gamma^{(K)}\}$ is the smallest 100% CS.



Simulated example (smallest 50% credible set)

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	Prob
1	0	0	1	1	0	0	0	1	1	0	0	0	1	1	0.0849
1	0	1	0	1	0	0	0	1	1	0	0	0	1	1	0.0817
1	0	0	1	1	0	0	0	1	1	1	1	1	0	0	0.0424
1	0	0	1	1	0	0	0	1	1	0	0	1	1	1	0.0396
0	1	1	0	1	0	0	0	1	1	0	0	0	1	1	0.0345
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0	1	0	1	1	0	0	0	1	1	0	0	1	1	1	0.0248
1	0	1	0	1	0	0	0	1	1	1	1	1	0	0	0.0218
1	0	1	0	0	1	0	0	1	1	0	0	0	1	1	0.0218
1	0	0	1	0	1	0	0	1	1	0	0	0	1	1	0.0202
0	1	1	0	1	0	0	0	1	1	0	0	1	1	1	0.0183
1	0	1	1	1	0	0	0	1	1	0	0	0	1	1	0.0176
0	1	1	0	1	0	0	0	1	1	1	1	1	0	0	0.0142



Simulated example (smallest 50% credible set)

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	Prob
1	0	0	1	1	0	0	0	1	1	0	0	0	1	1	0.0849
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Strategy

Other $100\alpha\%$ CS may be easier to understand and calculable using MCMC output.

The strategy is

- Remove variables with low PIPs.
- Partition remaining variables into approximately uncorrelated blocks
- Approximate the distribution in each block.
- Construct the credible sets from the approximation.



Estimating the correlation structure

Calculate the correlation $\rho_{ij} = \text{Correlation}(\gamma_i, \gamma_j)$ under the posterior distribution.



Estimating the correlation structure

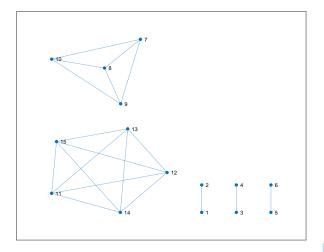
Calculate the correlation $\rho_{ij} = \text{Correlation}(\gamma_i, \gamma_j)$ under the posterior distribution.

Define the matrix A by $A_{ij} = I(|\rho_{ij}| > \tau)$ for some user-chosen threshold τ .

Find the components of the graph defined by the adjacency matrix *A*.



Simulated example





Choosing τ

Smaller au leads to

- Larger components
- Smaller credible sets
- Harder to understand and compute the approximation



Multivariate Bernoulli distribution (Dai et al., 2013)

Let \mathcal{D} be the set of non-empty subsets of $\{1,2,\ldots,K\}$, *i.e.* $\mathcal{D}=\{\{1\},\{2\},\ldots,\{1,2,\ldots,K\}\}.$

The K-dimensional multivariate Bernoulli distribution with parameters $\mathbf{f} = (f^{\epsilon} \in \mathbb{R} \mid \epsilon \in \mathcal{D})^T$ has the log probability mass function

$$\log p(y) = \sum_{r=1}^{K} \left(\sum_{1 \le j_1 < j_2 < \dots < j_r \le K} f^{j_1 j_2 \dots j_r} B^{j_1 j_2 \dots j_r} \right) - b(\mathbf{f})$$

where $B^{j_1j_2...j_r}(y) = y_{j_1}y_{j_2}...y_{j_r}$ and $b(\mathbf{f})$ is the log normalizing constant.



Properties

- The multivariate Bernoulli distribution is a member of the exponential family and f are the natural parameters.
- These natural parameters can be linked to the general parameters using the relationship

$$= \frac{p\left(\begin{array}{c} \text{even } \# \text{ zeros among } j_1, j_2, \dots, j_r \text{ components} \\ \text{and other components are all zero} \end{array}\right)}{p\left(\begin{array}{c} \text{odd } \# \text{ zeros among } j_1, j_2, \dots, j_r \text{ components} \\ \text{and other components are all zero} \end{array}\right)}.$$



Properties

For random vector $Y = (Y_1, \ldots, Y_K)$ following the multivariate Bernoulli distribution, suppose there are two blocks of nodes $Y' = (Y_1, Y_2, \ldots, Y_r)$ and $Y'' = (Y_{r+1}, Y_{r+2}, \ldots, Y_s)$, and denote index set $\tau_1 = \{1, 2, \ldots, r\}$ and $\tau_2 = \{r+1, r+2, \ldots, s\}$. Then Y' and Y'' are independent if and only if

$$f^{\tau} = 0, \forall \tau \cap \tau_1 = \emptyset \text{ and } \tau \cap \tau_2 = \emptyset.$$



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Restricting the model to only first and second order terms, (*i.e.* $f^{j_1j_2...j_r} = 0$ for all $j_1j_2...j_r$ with r > 2) leads the quadratic exponential binary model (Cox and Weimurth, 1994).



Approximation

Suppose there are r blocks $\gamma_{m_1}, \ldots, \gamma_{m_r}$ and $q(\gamma|\mathbf{f})$ is the approximating multivariate Bernoulli distribution.

$$egin{aligned} \mathsf{KL} &= \sum p(\gamma \mid \mathsf{Data}) \log p(\gamma) - \sum p(\gamma \mid \mathsf{Data}) \log q(\gamma) \ &= C - \sum_{i=1}^q p(\gamma \mid \mathsf{Data}) \log q(\gamma \mid \mathbf{f}) \ &= C - \sum_{i=1}^q p(\gamma_{m_j} \mid \mathsf{Data}) \log q(\gamma_{m_j} \mid \mathbf{f}) \end{aligned}$$



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If there is a sample $\gamma^{(1)}, \gamma^{(2)}, \dots, \gamma^{(N)} \sim p(\gamma \mid \mathsf{Data})$ then the a Monte Carlo approximation to the KL divergence is used

$$-\frac{1}{N}\sum_{i=1}^{q}\sum_{j=1}^{N}\log q\left(\left.\gamma_{m_{j}}^{(i)}\right|\mathbf{f}\right)$$



Finding the credible set

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$$p(S \mid \mathsf{Data}) = \prod_{i=1}^r p(Q_i \mid \mathsf{Data}).$$



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This allows the derivation of algorithms which control $p(S \mid Data)$ by changing the elements of S_1, \ldots, S_r .



Example (3 variables / 2 blocks)

$$\Gamma_1 = \{0,1\}, \Gamma_2 = \{(0,0), (0,1), (1,0), (1,1)\}$$

	Bloo	ck 1	Block 2							
	0	1	(0, 0)	(0, 0) $(0, 1)$ $(1, 0)$						
Prob	0.9 0.1		0	0.5	0.5	0				



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$$S_{1} = \{0, 1\} \Rightarrow p(S_{1}) = 1,$$

$$S_{2} = \{(0, 0), (0, 1), (1, 0), (1, 1)\} \Rightarrow p(S_{2}) = 1,$$

$$S = \{(0, 0, 0), (0, 0, 1), (0, 1, 0), (0, 1, 1), (1, 0, 0), (1, 0, 1), (1, 1, 0),$$

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$$S = \{(0, 0, 0), (0, 0, 1), (0, 1, 0), (0, 1, 1), (1, 0, 0), (1, 0, 1), (1, 1, 0),$$

$$(1, 1, 1)\} \Rightarrow p(S) = 1.$$

$$S_1 = \{0\} \Rightarrow p(S_1) = 0.9, S_2 = \{(0,1),(1,0)\} \Rightarrow p(S_2) = 1$$

 $S = \{(0,0,1),(0,1,0)\} \Rightarrow p(S) = 0.9.$



Algorithms

1. Calculate the probability of all possible credible sets (there are $2^{\#\Gamma_i}$). Find the smallest set with probability above the desired level.

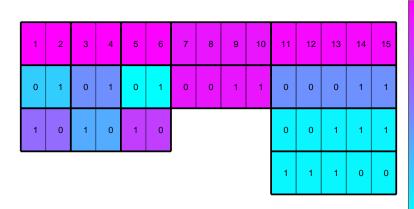


Algorithms

- 1. Calculate the probability of all possible credible sets (there are $2^{\#\Gamma_i}$). Find the smallest set with probability above the desired level.
- 2. Let Δ_i be the smallest change in $p(S_i)$ by removing an element from S_i . Choose $k = \arg\min(\Delta_1, \ldots, \Delta_r)$ and remove the corresponding element from S_k . Continue until removing any element leads to a probability below the desired level.



Simulated example (50 % credible set)



0.9 0.8 0.7 0.6 0.5 0.4 0.3

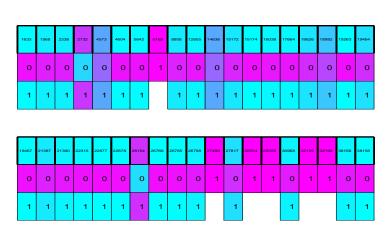
0.2

Simulated example (smallest 50% credible set)

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	Prob
1	0	0	1	1	0	0	0	1	1	0	0	0	1	1	0.0849
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1	0	1	0	0	1	0	0	1	1	0	0	0	1	1	0.0218
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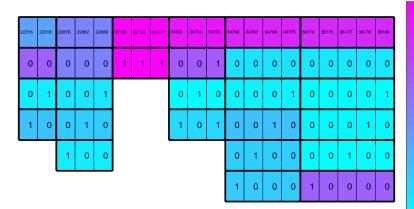


GWAS example



0.9 0.8 0.7 0.6 0.5 0.4 0.3 0.2 0.1

GWAS example



0.9

0.8

0.7

0.6

0.5

0.4

0.3

0.2

- One of 34766, 34767, 34768 and 34775 is included with probability 0.84 (individual PIPs are 0.10, 0.27, 0.22, 0.24)
- One of 22876, 22882, and 22889 is included with probability 0.50 (individual PIPs are 0.14, 0.23, 0.12)



Discussion

- Credible sets are useful way to explore uncertainty in the posterior distribution in Bayesian variable selection
- The method can identify blocks of highly correlated variables which can dilute marginal posterior inclusion probabilities
- The methods work with MCMC but could be easily extended to other inference frameworks (e.g. variational Bayes)
- These approaches could be extended to other discrete structures by finding a representation of the posterior with independence structure (e.g. factor models, etc.)



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