



Flexible Bayesian Modeling for Longitudinal Binary and Ordinal Responses

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Dec 7th, 2023

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- Focus on longitudinal studies with binary outcome, then extend the method to deal with longitudinal studies with ordinal outcome;

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- The main quantities of interests in such a study are:
 - the probability response curve;
 - the lead-lag correlations among repeated measurements;
- **Flexible** modeling made **easy**:
 - **Flexible**: allow general forms for the probability response curves and the temporal correlation structure;
 - **Easy**: clear prior specification strategy, efficient posterior inference algorithm, easy to work with expressions for quantities of interest.

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- Objective: analyzing the change of valence and arousal responses to evaluate students' affects as the term progresses.

A taxonomic review of models

- Marginal models: Molenberghs and Verbeke (2006);
- Conditional models: Di Lucca et al. (2013), DeYoreo and Kottas (2018);
- Subject-specific models:
 - Continuous: Ghosh and Hanson (2010); Quintana et al. (2016);
 - Binary: Jara et al. (2007); Tang and Duan (2012);
 - Mixed-scale: Kunihama et al. (2019);
- Functional data analysis tools: functional principal components analysis Van Der Linde (2009); Matuk et al. (2022).

Proposed Methodology

Hierarchical formulation of the proposed model

- Adopt a **functional data perspective**, treating each observed data vector \mathbf{Y}_i as the evaluation of trajectory $Y_i(\tau)$ on grid $\boldsymbol{\tau}_i = (\tau_{i1}, \dots, \tau_{iT_i})^\top$, for $i = 1, \dots, n$;

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- At the observed data level, we assume:

$$Y_i(\tau_{it}) \mid Z_i(\tau_{it}), \epsilon_{it} \stackrel{\text{ind.}}{\sim} \text{Bin}(1, \varphi(Z_i(\tau_{it}) + \epsilon_{it})), \quad t = 1, \dots, T_i, \quad i = 1, \dots, n,$$

where $\varphi(x) = \exp(x)/\{1 + \exp(x)\}$, and the error terms $\epsilon_{it} \mid \sigma_\epsilon^2 \stackrel{\text{i.i.d.}}{\sim} N(0, \sigma_\epsilon^2)$;

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- The main building block for the model construction is a **hierarchical Gaussian process prior** for $Z_i(\cdot)$, which we termed the signal process.

$$Z_i \mid \mu, \Sigma \stackrel{i.i.d.}{\sim} GP(\mu, \Sigma), \quad \mu \mid \Sigma \sim GP(\mu_0, \Sigma/\kappa), \quad \Sigma \sim IWP(\nu, \Psi_\phi).$$

Specifically we set $\kappa = (\nu - 3)^{-1}$;

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- We use an **Inverse-Wishart process (IWP) prior** for the covariance kernel. It is defined such that, on any finite grid $\boldsymbol{\tau} = (\tau_1, \dots, \tau_T)$ the projection $\Sigma(\boldsymbol{\tau}, \boldsymbol{\tau})$ follows $IW(\nu, \Psi_\phi(\boldsymbol{\tau}, \boldsymbol{\tau}))$. Here, $\Psi_\phi(\cdot, \cdot)$ is a non-negative definite function with parameters ϕ .

Proposition

Under the proposed model formulation, the signal process $Z(\tau)$ follows marginally a student-t process (TP). That is, for a generic grid vector $\tau = (\tau_1, \dots, \tau_T)^\top$,

$$Z_\tau = Z(\tau) \sim MVT(\nu, \mu_{0\tau}, \Psi_{\tau,\tau}), \text{ where } \mu_{0\tau} = \mu_0(\tau), \text{ and } \Psi_{\tau,\tau} = \Psi_\phi(\tau, \tau);$$

- TP is closed under marginalization. We can utilize the analytical form of the TP predictive distribution to develop a predictive inference scheme that resembles that of GP-based models. It is particularly useful in posterior inference;
- We can study the local behavior, such as smoothness, of the signal process trajectories by modeling them as TP;
- Modeling as TP facilitates the interpretation of the degrees of freedom parameter ν . It controls how heavy tailed the process is.

Highlights of the MCMC algorithm

- Recall that under unbalanced setting, the grid vectors for each subject τ_i are different.
We consider pooled grid $\tau = \cup_{i=1}^n \tau_i$;
- Let $\tilde{\mathbf{Z}}_i = Z_i(\tau)$, $\mathbf{Z}_i = Z_i(\tau_i)$, and $\mathbf{Z}_i^* = \tilde{\mathbf{Z}}_i \setminus \mathbf{Z}_i$;
- Factorizing the prior of $\tilde{\mathbf{Z}}_i$ as $p(\tilde{\mathbf{Z}}_i | \mu, \Sigma) = p(\mathbf{Z}_i^* | \mathbf{Z}_i, \mu, \Sigma)p(\mathbf{Z}_i | \mu, \Sigma)$. In a MCMC iteration, we first sample \mathbf{Z}_i , then conditioning on \mathbf{Z}_i to sample \mathbf{Z}_i^* (GP-based predictive sampling);
- Binary response to continuous latent process with errors, $Y_i(\tau_{it}) | Z_i(\tau_{it}), \epsilon_{it} \stackrel{ind.}{\sim} Bin(1, \varphi(Z_i(\tau_{it}) + \epsilon_{it}))$, reminds us the Pólya-Gamma technique;
- All model parameters can be sampled via Gibbs sampling, with **standard full condition distributions**.

Prediction and uncertainty

- Consider predicting $\tilde{\mathbf{Z}}_i^+ = Z_i(\tau^+)$, where $\tau^+ \supset \tau$ is a finer grid. Let $\check{\tau} = \tau^+ \setminus \tau$ and $\check{\mathbf{Z}}_i = Z_i(\check{\tau})$;
- We have the joint distribution:

$$\begin{pmatrix} \tilde{\mathbf{Z}}_i \\ \check{\mathbf{Z}}_i \end{pmatrix} \sim MVT \left(\nu, \begin{pmatrix} \mu_{0\tau} \\ \mu_{0\check{\tau}} \end{pmatrix}, \begin{pmatrix} \Psi_{\tau,\tau} & \Psi_{\tau,\check{\tau}} \\ \Psi_{\check{\tau},\tau} & \Psi_{\check{\tau},\check{\tau}} \end{pmatrix} \right),$$

and the prediction for $\check{\mathbf{Z}}_i$ are made based on the conditional distribution:

$$\check{\mathbf{Z}}_i | \tilde{\mathbf{Z}}_i \sim MVT \left(\nu + |\tau|, \check{\boldsymbol{\mu}}_{i\check{\tau}}, \frac{\nu + S_{i\tau} - 2}{\nu + |\tau| - 2} \check{\boldsymbol{\Psi}}_{\check{\tau},\check{\tau}} \right);$$

- For an in-sample subject, we first predict $Z_i(\tau_i^*)$ conditioning on $Z_i(\tau_i)$ by the GP predictive distribution, and next predict $Z_i(\check{\tau})$ conditioning on $Z_i(\tau_i)$ and $Z_i(\tau_i^*)$ by the TP predictive distribution;
- TP is scaling the predictive covariance by a factor that is related to the prediction error on observed grid, which can adjust the predictive covariance at unobserved grid points.

Model extension to deal with longitudinal ordinal responses

- Suppose the observation on subject i at time τ_{it} , denoted by Y_{it} , takes C possible categories;
- We encode the response as a vector with binary entries $\mathbf{Y}_{it} = (Y_{i1t}, \dots, Y_{iCt})$, such that $Y_{it} = j$ is equivalent to $Y_{ijt} = 1$ and $Y_{ikt} = 0$ for any $k \neq j$;
- We assume a multinomial response distribution for \mathbf{Y}_{it} , factorized in terms of binomial distributions (continuation-ratio logits),

$$Mult(\mathbf{Y}_{it} | m_{it}, \omega_{i1t}, \dots, \omega_{iCt}) = \prod_{j=1}^{C-1} Bin(Y_{ijt} | m_{ijt}, \varphi(Z_{ijt} + \epsilon_{ijt}))$$

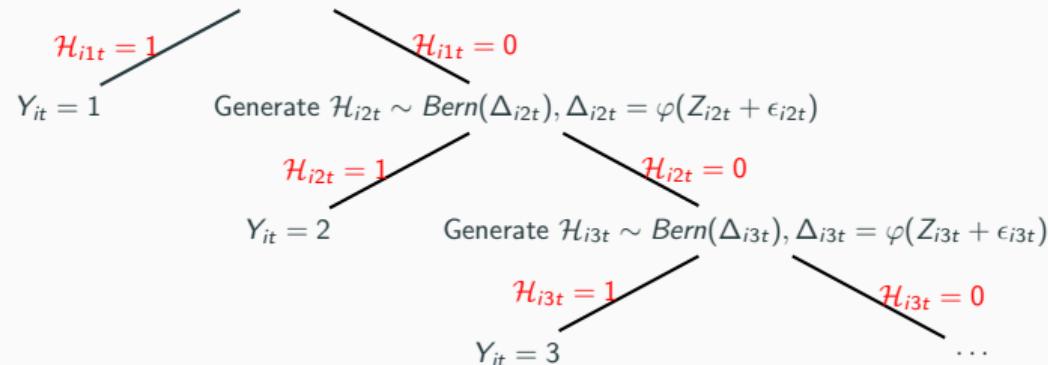
where $m_{it} = \sum_{j=1}^C Y_{ijt} \equiv 1$, $m_{i1t} = m_{it}$, and $m_{ijt} = m_{it} - \sum_{k=1}^{j-1} Y_{ikt}$;

- We adopt the proposed hierarchical GP-IWP modeling framework on $\{Z_{ijt}\}$ separately.

Sequential treatment of ordinal response and its practical implication

- The continuation-ratio logits structure offers a **sequential mechanism** to allocate the ordinal response Y_{it} .

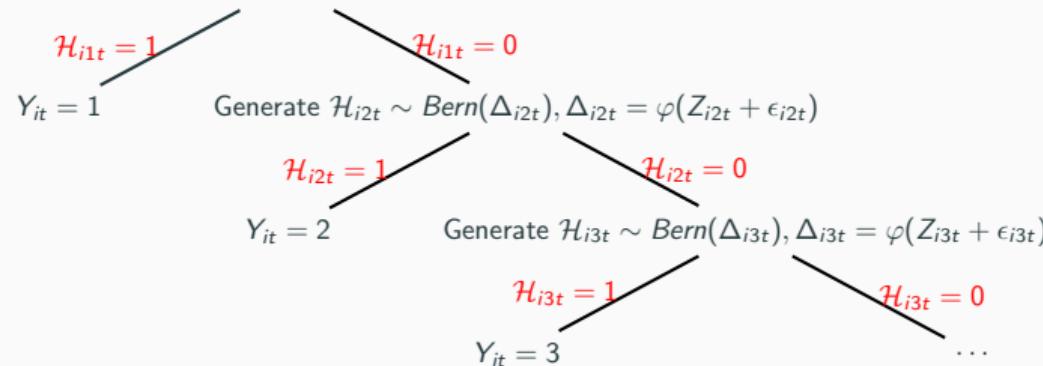
Generate $\mathcal{H}_{i1t} \sim Bern(\Delta_{i1t})$, $\Delta_{i1t} = \varphi(Z_{i1t} + \epsilon_{i1t})$



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- We can re-organize the original data set containing longitudinal ordinal responses to create $C - 1$ data sets with longitudinal binary outcomes. Then, fit the proposed model for binary responses **parallelly** on the $C - 1$ re-organized data sets.

Synthetic Data Examples

General settings

- In both experiments, we simulate longitudinal binary responses from:

$$Y_i(\tau_i) \mid \mathcal{Z}_i(\tau_i) \stackrel{\text{ind.}}{\sim} \text{Bin}(1, \eta(\mathcal{Z}_i(\tau_i))), \quad \tau_i = (\tau_{i1}, \dots, \tau_{iT_i}), \quad i = 1, \dots, n,$$

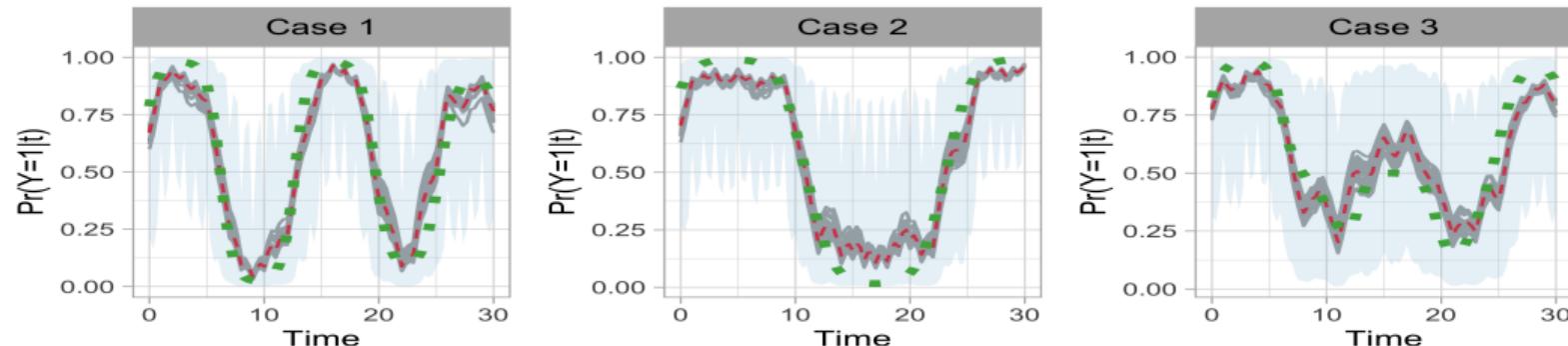
$$\mathcal{Z}_i(\tau_i) = f(\tau_i) + \omega_i + \epsilon_i \quad \epsilon_i \stackrel{\text{i.i.d.}}{\sim} N(\mathbf{0}, \sigma_\epsilon^2 \mathbf{I}),$$

a generic data generating process with:

- $\eta(\cdot)$: a generic link function mapping \mathbb{R} to $(0, 1)$;
- $f(\tau)$: a generic signal function of time;
- ω_i : a realization from a mean 0 continuous process that depicts the temporal covariance within the i -th subject.

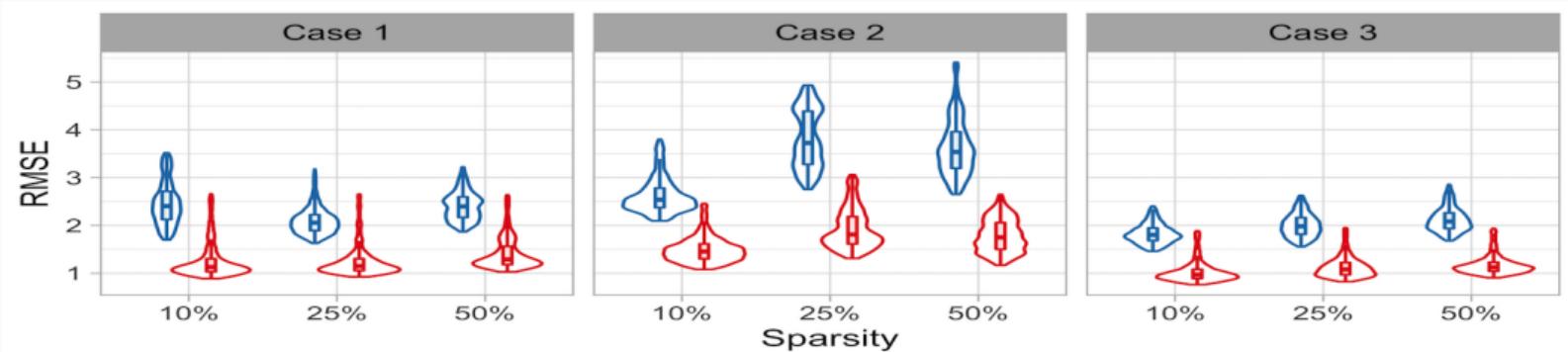
First set of experiments: result

- Focusing on the performance in recovering the fluctuation of the temporal trend;
- We simulate data with different link function, signal function, and temporal covariance structure combinations;
- To enforce an unbalanced study design, we randomly drop out a proportion of the simulated data. We consider different choices of drop out proportions.



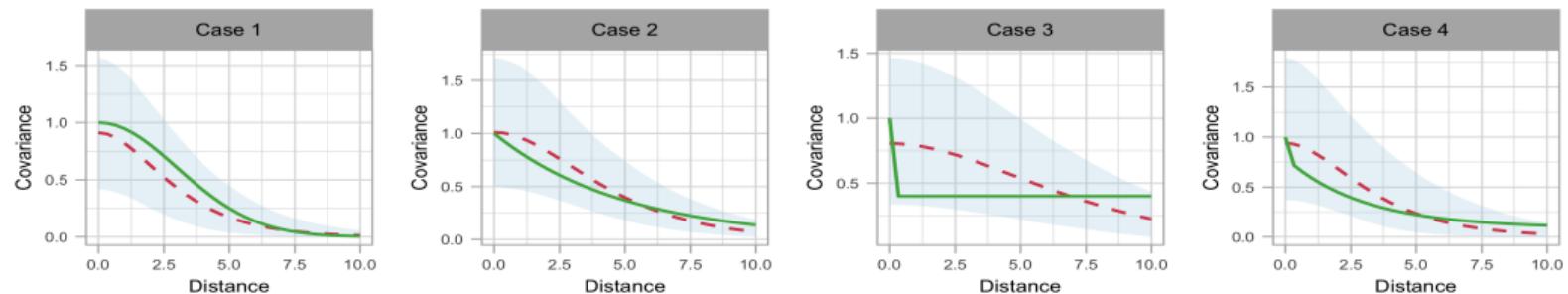
First set of experiments: comparison

- We compare the proposed model with its simplified backbone;
- Instead of modeling the mean function μ through a GP, we consider modeling it parametrically by $\mu(\tau) \equiv \mu_0$, and $\mu_0 \sim N(a_\mu, b_\mu)$;
- For criterion, we use the rooted mean square error (RMSE) between the model estimated signal process and the truth.



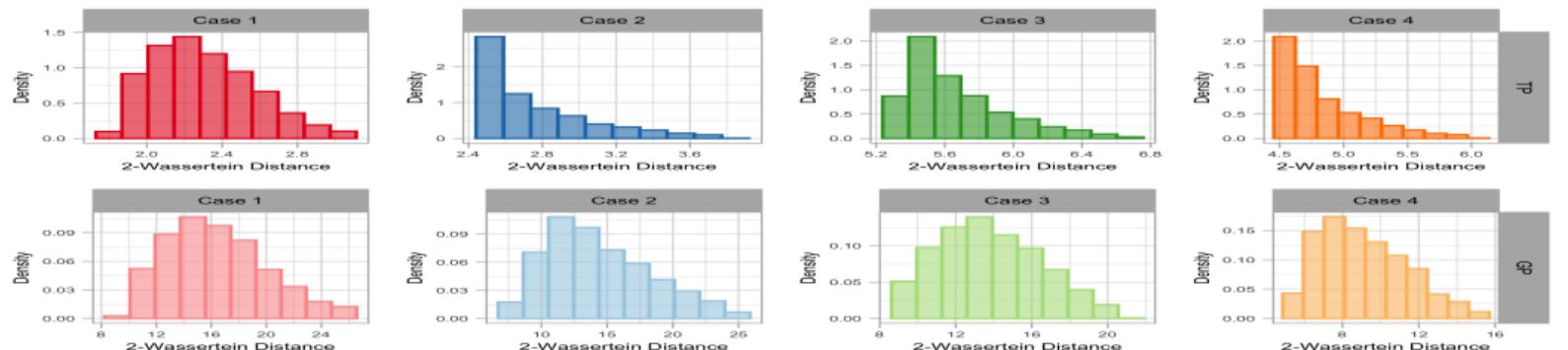
Second set of experiments: result

- Focusing on the performance in **the within subject covariance structure**;
- We simulate data with a number of possible choices for ω_i ;
- None of these choices imply covariance structures that are in the same form as the covariance kernel used in the proposed model.



Second set of experiments: comparison

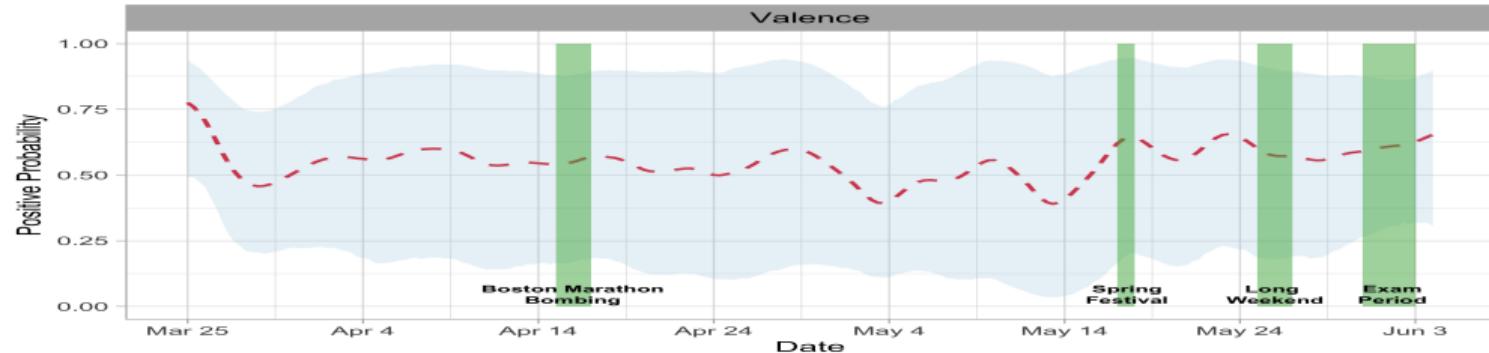
- We consider an alternative, simplified modeling approach, instead of modeling the covariance function nonparametrically, we assume a covariance kernel of certain parametric form;
- Specifically, $Z_i \stackrel{i.i.d.}{\sim} GP(\mu, \Psi_\phi)$, $\mu \sim GP(\mu_0, \Psi_\phi/\kappa)$, with parametric Ψ_ϕ ;
- We compute the 2-Wasserstein distance between the model estimated distribution of ω_i and the truth.



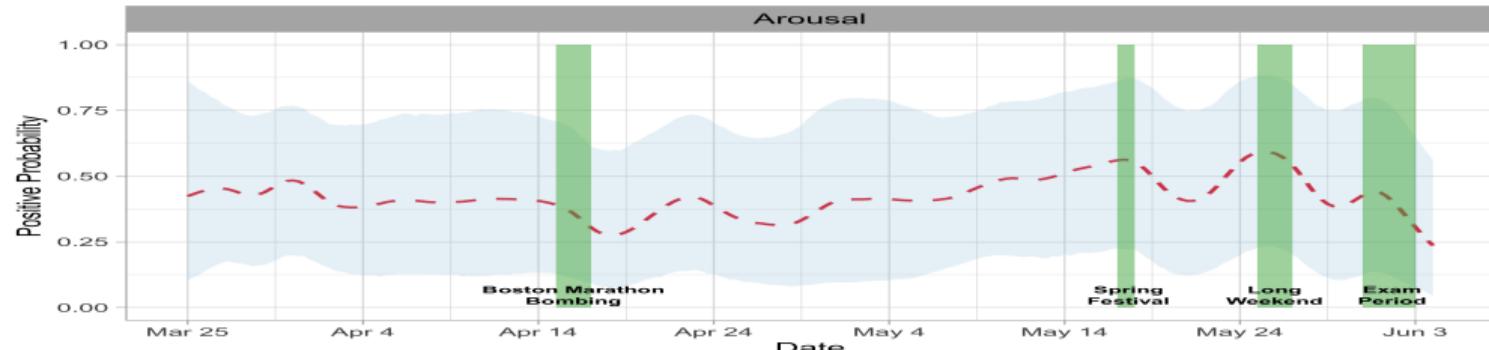
Real Data Application (Studentlife study)

Binary response case: result

- Valence:



- Arousal:



The mood coordinate space

Distress



Arousal +



Excitement

Valence +

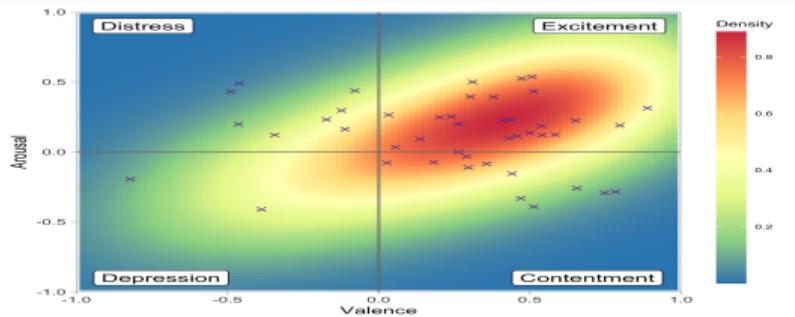


Depression

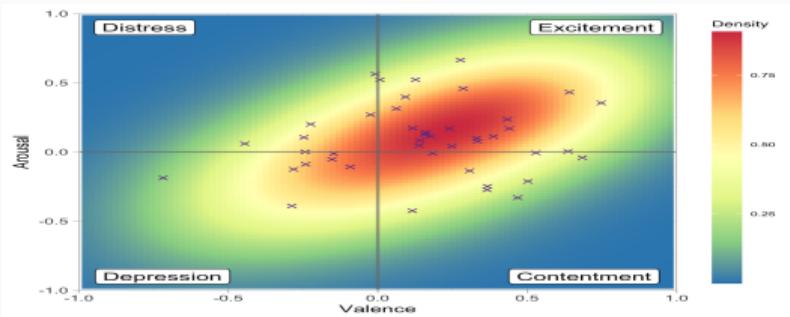


Contentment

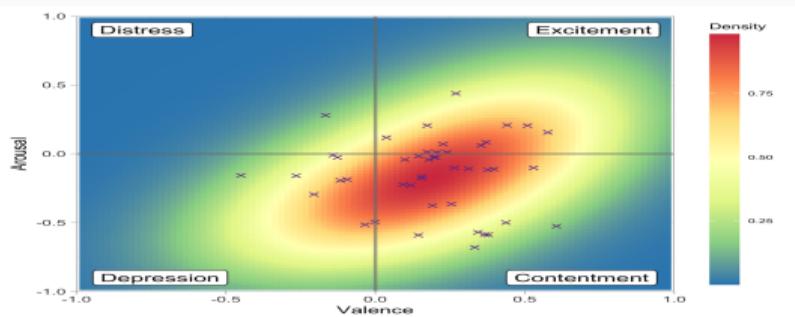
Categorizing emotional status



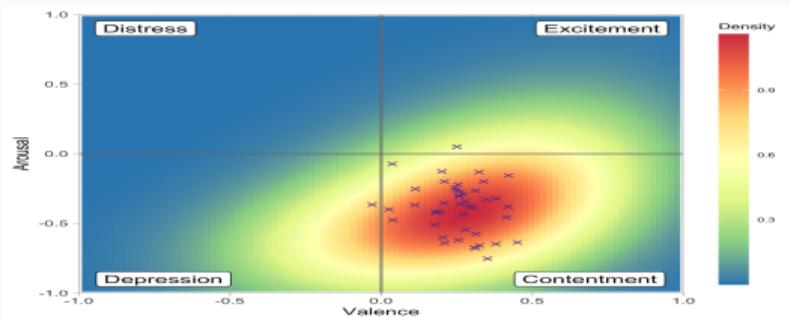
Green Key



Memorial Day

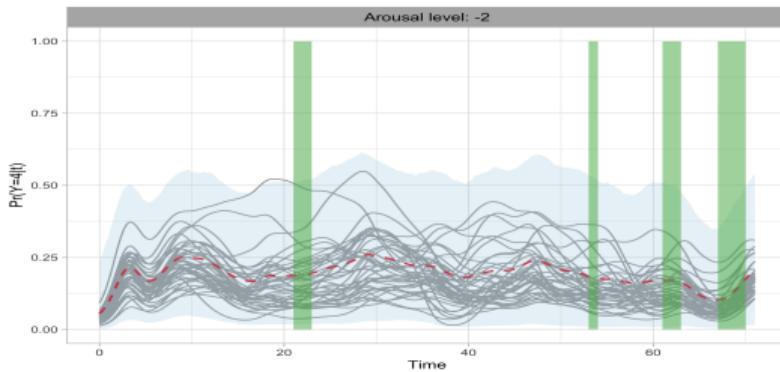
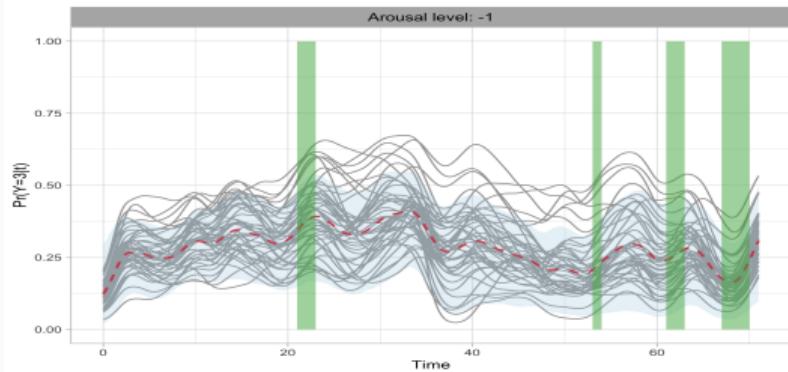
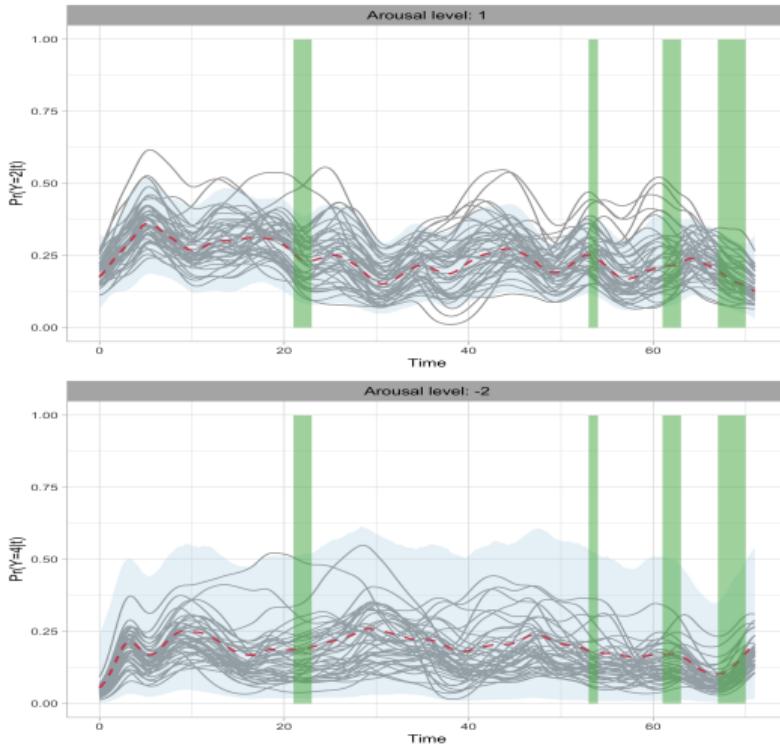
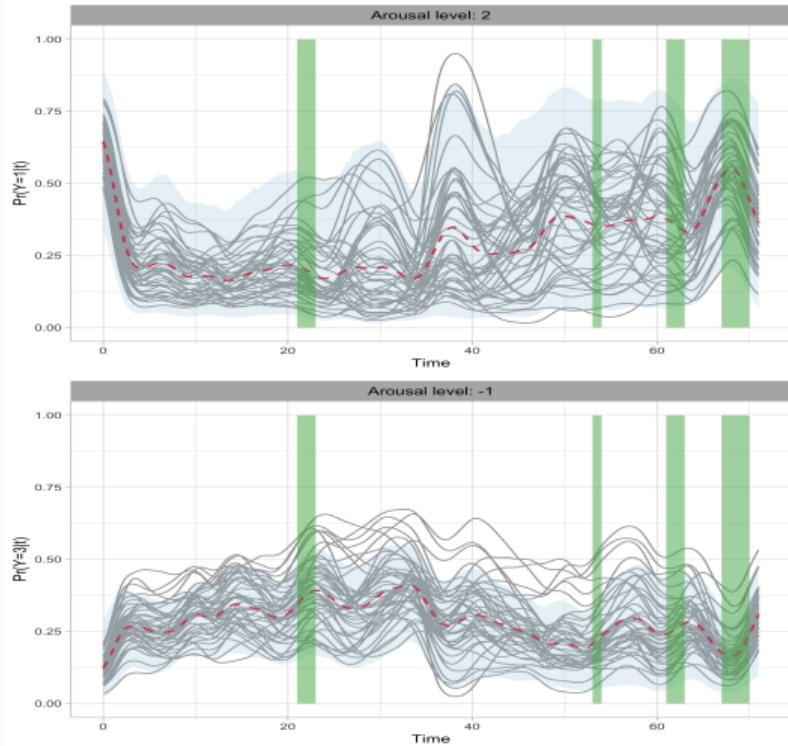


Final Exams Begin



Final Exams End

Four levels ordinal arousal score data



Concluding Remarks

Summary of contributions

- We model the mean and covariance **jointly and nonparametrically**, avoiding potential biases caused by a pre-specified model structure;
- The model **unifies the toolbox** for balanced and unbalanced longitudinal studies;
- The model encourages **borrowing of strength**, preserving systematic patterns that are common across all subject responses;
- We develop a **computationally efficient** posterior simulation method by taking advantage of conditional conjugacy.
- The model can be extended to deal with **ordinal responses** with a moderate to large number of categories.

MANY THANKS!

I am happy to answer any questions.

Jizhou Kang and Athanasios Kottas (2023+), "Flexible Bayesian Modeling for Longitudinal Binary and Ordinal Responses", arXiv: 2307.00224.

