

# Model Based Approach to Network Interference

Cornell STSCI / INFO / ILRST 3900

Fall 2025

[causal3900.github.io](https://causal3900.github.io)

Nov 20 2025

# Learning goals for today

At the end of class, you will be able to

- ▶ Use model based regression to estimate global average treatment effect under interference
- ▶ Use inverse probability weighting (IPW) to estimate global average treatment effect with a given exposure mapping
- ▶ Explain the implications of the choice of randomized design on the variance of the estimator

# Logistics

- ▶ Project Check-ins due Nov 25
- ▶ PSET 6 due Nov 25; Quiz 6 Dec 2

# Review of Network Interference

- ▶ Under interference the potential outcome of indiv  $i$  can depend on the treatments of others as well
- ▶ Requires a change in notation to indicate the additional dependence, e.g.  $Y_i^{\mathbf{a}}$  where  $\mathbf{a} = (a_1, a_2, \dots, a_n)$
- ▶ Assume potential outcome  $Y_i^{\mathbf{a}}$  depends only on  $\mathbf{a}$  only through treatment  $a_i$  and exposure level  $e_i$  as given by exposure mapping  $e_i = f_i(\mathbf{a})$ , e.g. neighborhood interference, anonymous interference

# Basic Solutions

- ▶ We will focus on estimating the Global Average Treatment Effect from randomized control trials under the neighborhood interference assumption

$$\text{GATE} = \frac{1}{n} \sum_{i=1}^n \left( Y_i^{(1,1)} - Y_i^{(0,0)} \right)$$

- ▶ Two methods for estimation under exchangeability:
  - ▶ Standardization & parametric g-formula with outcome model
  - ▶ Inverse treatment probability weighted estimator
- ▶ Earliest solutions for interference modify these approaches to estimate means under desired treatment and exposure levels

# Recap of using outcome modeling

- ▶ Learn a parametric model to predict expected outcome  $Y$  given treatment and covariates



- ▶ Estimate  $Y_i^a$  using the learned model,  $\hat{E}(Y \mid L = \ell_i, A = a)$
- ▶ Average estimates over all units

$$\hat{E}(Y^a) = \frac{1}{n} \sum_i \hat{E}(Y \mid L = \ell_i, A = a)$$

- ▶ Need  $L$  to be a sufficient adjustment set so that we have conditional exchangeability
- ▶ Under RCT, don't even need to condition on  $L$

# Outcome Modeling under Network Interference

- ▶ Key Idea: Learn a parametric model to predict expected outcome  $Y_i^{(a,e)}$  given treatment and exposure level
- ▶ Fit model to data  $\{(A_i, E_i, Y_i)\}_{i \in [n]}$
- ▶ Typically requires anonymous interference where exposure level is number or fraction of treated neighbors treated
- ▶ For every unit  $i$ , use the learned model to predict the outcome under treatment  $a$  and exposure level  $e$ , denoted  $\hat{Y}_i^{(a,e)}$
- ▶ Average over all units,

$$\hat{E}(Y^{a,e}) = \frac{1}{n} \sum_i \hat{Y}_i^{(a,e)}$$

# Outcome Modeling under Network Interference

- ▶ Linear models are most common, e.g.  $Y_i = \alpha A_i + \beta E_i + \gamma$ , where  $E_i$  is fraction of treated neighbors
- ▶ Global Average Treatment Effect

$$\begin{aligned}\widehat{\text{GATE}} &= \frac{1}{n} \sum_i \hat{Y}_i^{(1,1)} - \frac{1}{n} \sum_i \hat{Y}_i^{(0,0)} \\ &= \frac{1}{n} \sum_i (\hat{\alpha} + \hat{\beta} + \hat{\gamma}) - \frac{1}{n} \sum_i \hat{\gamma} = \hat{\alpha} + \hat{\beta}\end{aligned}$$

- ▶ Direct Average Treatment Effect

$$\begin{aligned}\widehat{\text{DATE}} &= \frac{1}{n} \sum_i \hat{Y}_i^{(1,0)} - \frac{1}{n} \sum_i \hat{Y}_i^{(0,0)} \\ &= \frac{1}{n} \sum_i (\hat{\alpha} + \hat{\gamma}) - \frac{1}{n} \sum_i \hat{\gamma} = \hat{\alpha}\end{aligned}$$

- ▶ Indirect Average Treatment Effect

$$\begin{aligned}\widehat{\text{IATE}} &= \frac{1}{n} \sum_i \hat{Y}_i^{(0,1)} - \frac{1}{n} \sum_i \hat{Y}_i^{(0,0)} \\ &= \frac{1}{n} \sum_i (\hat{\beta} + \hat{\gamma}) - \frac{1}{n} \sum_i \hat{\gamma} = \hat{\beta}\end{aligned}$$

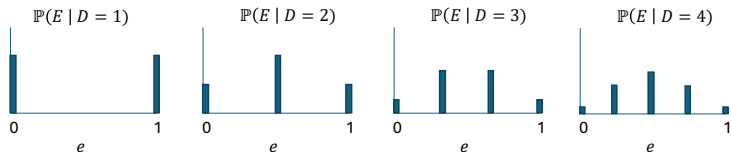


# Outcome Modeling under Network Interference

- ▶ What assumptions are needed?
- ▶ What is the relevant causal graph in the network setting?
- ▶ When do we need to think about adjusting for confounders?

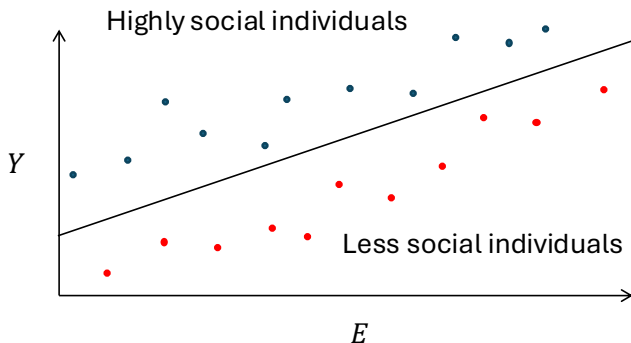
# Outcome Modeling under Network Interference

- ▶ When could we still need to condition on covariates to get a sufficient adjustment set even if treatments are randomized?
- ▶ E.g. let exposure level be fraction of treated neighbors, then distribution of  $E_i$  depends on number of neighbors  $D_i$



# Outcome Modeling under Network Interference

- ▶ When could we still need to condition on covariates to get a sufficient adjustment set even if treatments are randomized?
- ▶ E.g. let exposure level be fraction of treated neighbors, then distribution of  $E_i$  depends on number of neighbors  $D_i$
- ▶ Number of neighbors  $D_i$  affects outcome even when conditioned on exposure level



## Recap of Inverse probability of treatment weighting

- Estimate means by averaging the outcomes of units with treatment  $A_i = a$  multiplied by the inverse of probability of the treatment conditioned on covariates  $\mathbb{P}(A_i = a \mid L_i)$

$$\hat{E}(Y^a) = \frac{1}{n} \sum_{i:A_i=a} \frac{Y_i}{\mathbb{P}(A_i = a \mid L_i)}$$

- $\pi_i$  denotes probability that  $i$  is treated conditioned on its covariates, s.t.  $\mathbb{P}(A_i = 1 \mid L_i) = \pi_i$  and  $\mathbb{P}(A_i = 0 \mid L_i) = 1 - \pi_i$
- Take difference of estimates for treated and control

$$\hat{E}(Y^1) - \hat{E}(Y^0) = \frac{1}{n} \left( \sum_i \frac{A_i Y_i}{\hat{\pi}_i} - \sum_i \frac{(1 - A_i) Y_i}{1 - \hat{\pi}_i} \right)$$

- Requires conditional exchangeability

# IPW under network interference

- Modify to use exposure mapping (assume RCT)

$$\hat{E}(Y^{(a,e)}) = \frac{1}{N} \sum_{i:A_i=a, E_i=e} \frac{Y_i}{\mathbb{P}(A_i = a, E_i = e)}$$
$$\widehat{\text{GATE}} = \hat{E}(Y^{(1,1)}) - \hat{E}(Y^{(0,0)})$$

- Does not require anonymous interference, can use any exposure mapping
- Variance will depend on the exposure probabilities  $\mathbb{P}(A_i = 1, E_i = \mathbf{1})$  and  $\mathbb{P}(A_i = 0, E_i = \mathbf{0})$
- Observational studies are significantly more complex as we now need to care about the joint treatment probability distribution as it relates to the exposure levels

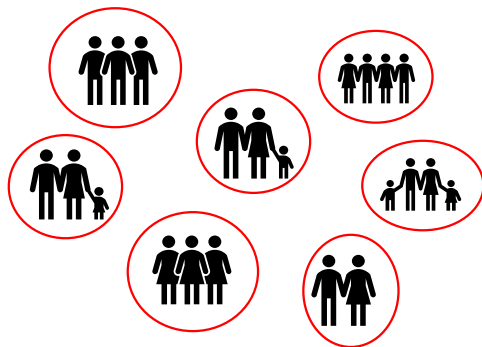
## Variance of IPW estimator

$$\hat{E}(Y^{(a,e)}) = \frac{1}{N} \sum_{i:A_i=a, E_i=e} \frac{Y_i}{\mathbb{P}(A_i = a, E_i = e)}$$

- ▶ If exposure probabilities  $\mathbb{P}(A_i = 1, E_i = \mathbf{1})$  and  $\mathbb{P}(A_i = 0, E_i = \mathbf{0})$  are small, then any measurement noise in the outcomes will be amplified, leading to high variance
- ▶ Let  $D_i$  denote the number of neighbors (including  $i$  itself)
- ▶ Under independent treatment w/prob 0.5, exposure probability is exponential in  $D_i$ , i.e.  $\mathbb{P}(A_i = 1, E_i = \mathbf{1}) = (0.5)^{D_i}$
- ▶ For  $D_i = 5$ , exposure prob is 0.000976, s.t. in a network of 1000 nodes, likely no units observed under full treatment
- ▶ Sophisticated clustered treatment assignments reduce variance by increasing probability of full treatment / control

# Cluster Randomized Designs

- ▶ Initially motivated by networks consisting of many tightly connected households
- ▶ No interference edges across households
- ▶ Assign treatments to each household jointly
- ▶ If household treatment probability is 0.5, then full exposure probability is 0.5



# Cluster Randomized Designs

- ▶ Can use clustering algorithms on general graphs
- ▶ Assign treatments to each cluster jointly
- ▶ If cluster treatment probability is 0.5, then full exposure probability is  $0.5^{(\# \text{ neighboring clusters})}$





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