ISyE 6644 - Spring 2019 - Test #2 Solutions (Revised 7/5/19)

This test is 120 minutes. You're allowed *two* cheat sheets (4 sides total).

This test requires a proctor. All questions are 3 points, except #33, which is 4 points.

1. Consider the following joint p.m.f.

$$\begin{array}{c|cccc} f(x,y) & X=1 & X=2 \\ \hline Y=0 & 0.06 & 0.24 \\ Y=1 & 0.14 & 0.56 \\ \end{array}$$

YES or NO? Are X and Y independent?

Solution: Let's re-do the table to include the marginals:

| f(x,y) | X = 1 | X = 2 | $f_Y(y)$ |
|----------|-------|-------|----------|
| Y = 0 | 0.06 | 0.24 | 0.3 |
| Y = 1 | 0.14 | 0.56 | 0.7 |
| $f_X(x)$ | 0.2 | 0.8 | 1 |

Note that $f(x,y) = f_X(x)f_Y(y)$ for all four combinations of x and y, so the answer is YES. \square

2. TRUE or FALSE? Most discrete-event simulations proceed by moving the simulation clock to the most-imminent event on the future events list; executing that event (including any adds, deletes, or swaps to the FEL); and then repeating this cycle.

Solution: TRUE. □

3. TRUE or FALSE? In Arena, it is *not possible* to Seize and Release a *specific* member of a resource set, and then Seize that same guy a *second* time.

Solution: FALSE. \square

| 4. | In Arena, where do we specify the maximum buffer size for a certain queue? |
|----|--|
| | (a) In the Queue spreadsheet found in the Basic Process template. |
| | (b) In the Queue module found in the Advanced Process template. |
| | (c) In the Queue block found in the Blocks template. |
| | (d) In the Process module found in the Basic Process template. |
| | (e) Both (c) and (d). |
| | Solution: Just (c). \Box |
| 5. | TRUE or FALSE? The Arena resource that we've named ${\tt Barber}$ can belong to 3 different resource sets. |
| | Solution: TRUE. |
| 6. | TRUE or FALSE? Arena resources within a particular resource set can each have different schedules. |
| | Solution: TRUE. |
| 7. | What is TNOW in Arena? |
| | (a) An attribute that keeps track of the current time. |
| | (b) A variable that keeps track of the current time. |
| | (c) An attribute that keeps track of the current customer's time in the system. |
| | (d) A variable that keeps track of the current customer's time in the system. |
| | Solution: (b). \Box |
| 8. | Which statement best describes Arena arrivals? |
| | (a) Arrivals can be generated in batches of random size. |
| | (b) Arrivals can be generated via arbitrary interarrival distributions. |
| | (c) Arrivals can follow a nonhomogeneous Poisson process. |
| | |

| (d) EXPO arrivals constitute a Poisson process. |
|--|
| (e) All of the above. |
| Solution: (e). \Box |
| 9. What is the variance of the Arena expression (0 == 0)*NORM(0,4)? |
| (a) 0 |
| (b) 1 |
| (c) 4 |
| (d) 16 |
| (e) 64 |
| Solution: In Arena, a $NORM(a, b)$ random variable has variance b^2 . Therefore, |
| $Var\Big((\mathtt{0} == \mathtt{0}) * \mathtt{NORM}(\mathtt{10}, \mathtt{4})\Big) \; = \; Var\Big(1 * \mathtt{NORM}(\mathtt{10}, \mathtt{4})\Big) \; = \; 16.$ |
| So the answer is (d). \Box |
| 10. TRUE or FALSE? In Arena, the expression NORM(0,1)+NORM(0,1) has the same distribution as the expression TRIA(0,1,2). |
| Solution: FALSE. In fact, the $Tria(0,1,2)$ is the sum of two $Unif(0,1)$'s, not the sum of two $Nor(0,1)$'s. \square |
| 11. Which statement below describing Arena arrivals is false? |
| (a) Arrivals can be generated in batches of random size. |
| (b) Arrivals can be generated via arbitrary interarrival distributions. |
| (c) Arrivals can follow a nonhomogeneous Poisson process. |
| (d) EXPO arrivals constitute a Poisson process. |
| (e) You can use NORM random variables to generate negative interarrival times. |
| Solution: (e) . \Box |

| 12. | Consider the generator $X_{i+1} = (3X_i + 5) \mod(8)$. Using $X_0 = 0$, calculate the PRN U_{2001} . |
|-----|---|
| | (a) 0 |
| | (b) 0.125 |
| | (c) 0.5 |
| | (d) 0.625 |
| | (e) 1 |
| | (f) 3 |
| | Solution: If $X_0 = 0$, then we have $X_1 = 5$, $X_2 = 4$, $X_3 = 1$, and $X_4 = 0$, so it cycles after 4 iterations. Thus, $X_{2001} = X_{1997} = \cdots = X_1 = 5$, so that $U_{2001} = 5/8$. So the answer is (d). \square |
| 13. | Consider our desert island generator $X_{i+1}=16807X_i\mathrm{mod}(2^{31}-1)$. If $X_0=3456789$, find the value of X_1 . |
| | (a) 0 |
| | (b) 0.61359 |
| | (c) 25,152,971 |
| | (d) 116,194,254 |
| | (e) 1,993,604,588 |
| | (f) $2^{31} - 582$ |
| | Solution: By hook or by crook (e.g., by the algorithm given in the notes), we find that $X_1 = 116,194,254$, so that the answer is (d). \Box |
| 14. | TRUE or FALSE? There are perfectly good PRN generators out there having periods greater than $2^{19900}!$ |
| | Solution: Crazy, but TRUE. \square |
| | |

15. Consider the following 20 PRN's.

How many runs up and down do you get from this sequence?

- (a) 8
- (b) 14
- (c) 15
- (d) 20
- (e) 33

Solution: Using the usual notation from class, we have the following sequence of +'s and -'s:

This corresponds to A = 14 runs, i.e., choice (b). \Box

16. Suppose that we have a sequence of n=100 PRN's, and we observe 56 runs up and down. Using $\alpha=0.05$, do we ACCEPT or REJECT the null hypothesis of independence?

Solution: By class notes, under the null hypothesis of independence, $A \approx \text{Nor}(\frac{2n-1}{3}, \frac{16n-29}{90}) \sim \text{Nor}(66.33, 17.46)$. Thus, the test statistic is

$$Z_0 = \frac{A - \mathsf{E}[A]}{\sqrt{\mathsf{Var}(A)}} = \frac{56 - 66.33}{\sqrt{17.46}} = -2.47.$$

Since $|Z_0| > z_{\alpha/2} = 1.96$, we REJECT H_0 . Thus, we conclude that they're dependent. \square

17. Suppose we sample 1000 PRN's and we wish to conduct a χ^2 goodness-of-fit test at level $\alpha = 0.10$ of the hypothesis that the numbers are Unif(0,1). Here are the results, divided into 5 intervals.

| interval | O_i |
|------------|-------|
| [0.0, 0.2] | 250 |
| (0.2, 0.4] | 120 |
| (0.4, 0.6] | 110 |
| (0.6, 0.8] | 270 |
| (0.8, 1.0] | 250 |

Use the tabled results to find the value of the g-o-f statistic, χ_0^2 .

- (a) 1.22
- (b) 12.2
- (c) 122
- (d) 1220
- (e) 24400

Solution: Since we have k = 5 equiprobable intervals, we obtain $E_i = n/k = 200$ for all i. Thus,

$$\chi_0^2 = \sum_{i=1}^k \frac{(O_i - E_i)^2}{E_i} = \sum_{i=1}^5 \frac{(O_i - 200)^2}{200} = 122,$$

and so the answer is (c). \Box

18. Similar to Problem 17, let's do a χ^2 test for uniformity with n=1000 PRN's, k=5 equi-probability intervals, and level $\alpha=0.10$. But now suppose it turns out that $\chi_0^2=56$ (instead of whatever answer you got before). Do we ACCEPT or REJECT the null hypothesis of uniformity?

Solution: First of all, the appropriate test quantile is $\chi^2_{\alpha,k-1} = \chi^2_{0.10,4} = 7.78$. Since $\chi^2_0 \gg \chi^2_{\alpha,k-1}$, we easily REJECT uniformity. \square

- 19. Suppose the random variable X has p.d.f. $f(x) = 3(x-1)^2$ for $1 \le x \le 2$. Find the inverse of its c.d.f., i.e., $F^{-1}(U)$, where U is a PRN.
 - (a) $3(X-1)^2$
 - (b) $3(U-1)^2$

- (c) $(X-1)^3$
- (d) $(U-1)^3$
- (e) $U^{1/3} + 1$

Solution: We have

$$F(x) = \int_{1}^{x} 3(t-1)^{2} dt = (x-1)^{3}, \text{ for } 1 \le x \le 2.$$

Set $F(X) = (X-1)^3 = U$. Solving, we have $X = F^{-1}(U) = U^{1/3} + 1$. So the answer is (e). \square

- 20. Use the inverse transform method with U=0.27 to generate a realization of $Z\sim \operatorname{Nor}(0,1)$. (You'll need the normal tables for this.) And then generate $X\sim \operatorname{Nor}(3,4)$ by making the appropriate linear transformation $X=\mu+\sigma Z$. What's your final value for X?
 - (a) 0.73
 - (b) 1.77
 - (c) 2.73
 - (d) 3.27
 - (e) 5.27

Solution: We take

$$X = \mu + \sigma Z = \mu + \sigma \Phi^{-1}(U) = 3 + 2\Phi^{-1}(0.27) = 3 + 2(-0.6128) = 1.77.$$

Thus, the answer is (b). \Box

- 21. Suppose that X has the Gamma(α, λ) distribution with (messy) c.d.f. $F(x) = \frac{\lambda^{\alpha}}{\Gamma(\alpha)} \int_{0}^{x} t^{\alpha-1} e^{-\lambda t} dt$. What is the distribution of the random variable 3F(X) + 2?
 - (a) Unif(0,1)
 - (b) Unif(0,3)
 - (c) Unif(2,5)
 - (d) $Gamma(\alpha, \lambda)$

(e) $3\operatorname{Gamma}(\alpha,\lambda) + 2$

Solution: By the Inverse Transform Theorem, $F(X) \sim \text{Unif}(0,1)$. Thus,

$$3F(X) + 2 \sim \text{Unif}(0,3) + 2 \sim \text{Unif}(2,5).$$

So the answer is (c). \square

- 22. The number of tries until a UGA student gets one of a series of multiple choice questions correct is $X \sim \text{Geom}(0.3)$. Use the PRN U = 0.50 to generate X via inverse transform.
 - (a) 0
 - (b) 1
 - (c) 2
 - (d) 8
 - (e) 15
 - (f) 22

Solution: We have

$$X = \left\lceil \frac{\ell \ln(1-U)}{\ell \ln(1-p)} \right\rceil = \left\lceil \frac{\ell \ln(0.50)}{\ell \ln(0.7)} \right\rceil = 2.$$

So the answer is (c). \Box

- 23. Suppose that $U_1 = 0.70$ and $U_2 = 0.41$ are realizations of two i.i.d. Unif(0,1)'s. Use the Box–Muller method to generate two i.i.d. standard normals, Z_1 and Z_2 . [Note that there are a number of ways to do this problem, but I expect that you will use the standard method; and that answer is reflected in the choices below.]
 - (a) -0.713 and 0.453
 - (b) 0 and 1.96
 - (c) -0.844 and 0.038
 - (d) 0.844 and 0.038

(e) 1.788 and 0.076

Solution: Being very careful to use radians (not degrees), we have

$$Z_1 = \sqrt{-2\ell n(U_1)}\cos(2\pi U_2) = -0.713$$

 $Z_2 = \sqrt{-2\ell n(U_1)}\sin(2\pi U_2) = 0.453.$

This is answer (a). \Box

- 24. If Z_1 and Z_2 are i.i.d. Nor(0,1), then we say that $(Z_1 + \delta)/Z_2$ has the noncentral t distribution with 1 degree of freedom and noncentrality parameter δ . If $Z_1 = -0.843$ and $Z_2 = 0.053$, generate a noncentral t random variate with one degree of freedom and noncentrality parameter 2. [Note that there are a number of ways to do this problem, but I expect that you will simply follow the directions given in the problem; it is that answer that is reflected in the choices below.]
 - (a) -40.70
 - (b) 0
 - (c) 1.96
 - (d) 21.83
 - (e) 40.70

Solution: (-0.843 + 2)/0.053 = 21.83, so that the answer is (d).

25. TRUE or FALSE? If Z_1 and Z_2 are i.i.d. Nor(0,1)'s, e.g., resulting from the Box–Muller method, then $Z_1^2 + Z_2^2 \sim \text{Exp}(1/2)$.

Solution: As explained in our class notes,

$$Z_2^2 + Z_2^2 = -2\ell n(U_1)[\cos^2(2\pi U_2) + \sin^2(2\pi U_2)] = -2\ell n(U_1) \sim \text{Exp}(1/2),$$

where the distributional result follows by inverse transform (which we've used a million times). Thus, the assertion is TRUE. \Box

26. Consider the following 20 PRN's.

Use the sum of these PRN's to generate a single approximately Nor(0,1) random variate via our "desert island" technique.

- (a) -0.527
- (b) 0
- (c) 0.217
- (d) 0.527
- (e) 1.96

Solution: From our class notes, we have

$$Z = \frac{\sum_{i=1}^{n} U_i - \frac{n}{2}}{\sqrt{\frac{n}{12}}} = \frac{10.28 - 10}{\sqrt{20/12}} = 0.217.$$

Thus, the answer is (c). \square

- 27. Suppose that $U_1 = 0.65$, $U_2 = 0.45$, $U_3 = 0.82$, $U_4 = 0.09$, and $U_5 = 0.26$. Use our acceptance-rejection technique from class to generate a Pois($\lambda = 1.7$) random variate. (You may not need to use all of the uniforms.)
 - (a) 0
 - (b) 1
 - (c) 2
 - (d) 3
 - (e) 4

Solution: Define $p_n \equiv \prod_{i=1}^{n+1} U_i$. We'll stop as soon as $p_n < e^{-1.7} = 0.1827$. Let's make the following convenient table.

| n | U_{n+1} | p_n | Stop? |
|---|-----------|--------|-------|
| 0 | 0.65 | 0.65 | nope |
| 1 | 0.45 | 0.2925 | nope |
| 2 | 0.82 | 0.2399 | nope |
| 3 | 0.09 | 0.0216 | yup |

So we take N = 3, i.e., answer (d). \square

- 28. Suppose U_1 and U_2 are PRN's and $\Phi(x)$ is the standard normal c.d.f. Find $\Pr\left(\Phi^{-1}(U_1) + \Phi^{-1}(U_2) \leq \sqrt{2}\right)$.
 - (a) 0
 - (b) 0.159
 - (c) 0.240
 - (d) 0.760
 - (e) 0.841

Solution: By Inverse Transform, $\Phi^{-1}(U_1)$ and $\Phi^{-1}(U_2)$ are i.i.d. Nor(0,1). Thus,

$$\Pr\left(\Phi^{-1}(U_1) + \Phi^{-1}(U_2) \le \sqrt{2}\right) = \Pr\left(\operatorname{Nor}(0, 2) \le \sqrt{2}\right)$$

= $\Pr(\operatorname{Nor}(0, 1) \le 1)$
= $\Phi(1) = 0.8413.$

This is (e). \Box

- 29. Suppose Z_1 and Z_2 are i.i.d. standard normal with c.d.f. $\Phi(x)$. Find the value of $\Pr(\Phi(Z_1) + \Phi(Z_2) \leq 2)$.
 - (a) 0
 - (b) 0.25
 - (c) 0.5
 - (d) 0.75
 - (e) 1

Solution: By Inverse Transform, $\Phi(Z_1)$ and $\Phi(Z_2)$ are i.i.d. Unif(0,1). Thus,

$$\Pr(\Phi(Z_1) + \Phi(Z_2) \le 2) = \Pr(\text{Tria}(0, 1, 2) \le 2) = 1.$$

So the answer is (e). \Box

- 30. Consider a RV X having p.d.f. $f(x) = cx^3(1-x)^2e^{-x}$, for 0 < x < 1, where c is the constant that makes this monster integrate to 1 (and which I'm too lazy to calculate right now). Which method would you most likely use to generate X?
 - (a) inversion
 - (b) convolution
 - (c) Box–Muller
 - (d) acceptance-rejection
 - (e) composition
 - (f) nonhomogeneous Poisson

Solution: This is an obvious case for A-R, choice (d). \Box

- 31. Suppose that X_1, X_2, X_3 are i.i.d. with p.d.f. f(x) = 2x for 0 < x < 1. Give an expression involving a *single* PRN U that you can use to generate a realization of $\max\{X_1, X_2, X_3\}$.
 - (a) $\max\{U_1^{1/2}, U_2^{1/2}, U_3^{1/2}\}$
 - (b) $\max\{U_1^2, U_2^2, U_3^2\}$
 - (c) $U^{1/2}$
 - (d) $U^{1/6}$
 - (e) $U^{1/8}$
 - (f) U^{8}

Solution: I'll first do a general derivation. Consider X_1, X_2, \ldots, X_n i.i.d. with c.d.f. F(x) and let $Y \equiv \max\{X_1, X_2, \ldots, X_n\}$. Then (as explained in class) the c.d.f. of Y is

$$G(y) = \Pr(Y \le y)$$

$$= \Pr(\max\{X_1, X_2, \dots, X_n\} \le y)$$

$$= \Pr(X_1 \le y, X_2 \le y, \dots, X_n \le y)$$

$$= [\Pr(X_1 \le y)]^n \text{ (since the } X_i\text{'s are i.i.d.})$$

$$= [F(y)]^n.$$

Now, by Inverse Transform, $G(Y) = U \sim \text{Unif}(0,1)$, and so $[F(Y)]^n = U$. Since the c.d.f. of X is clearly $F(x) = x^2$ for 0 < x < 1, and since n = 3 here, we have

$$U = G(Y) = [F(Y)]^n = (Y^2)^n = Y^6,$$

so that $Y = U^{1/6}$. This is choice (d). \square

- 32. Consider a 2×2 covariance matrix $\Sigma = \begin{pmatrix} 3 & 2 \\ 2 & 3 \end{pmatrix}$. Find a lower-triangular matrix C such that $CC' = \Sigma$, and tell me what the entry c_{22} is.
 - (a) 0
 - (b) 5/3
 - (c) $\sqrt{5/3}$
 - (d) $2/\sqrt{3}$
 - (e) 4/3

Solution: By class notes, we saw that for k=2,

$$C = \begin{pmatrix} \sqrt{\sigma_{11}} & 0 \\ \frac{\sigma_{12}}{\sqrt{\sigma_{11}}} & \sqrt{\sigma_{22} - \frac{\sigma_{12}^2}{\sigma_{11}}} \end{pmatrix} = \begin{pmatrix} \sqrt{3} & 0 \\ \frac{2}{\sqrt{3}} & \sqrt{3 - \frac{2^2}{3}} \end{pmatrix} = \begin{pmatrix} \sqrt{3} & 0 \\ \frac{2}{\sqrt{3}} & \sqrt{\frac{5}{3}} \end{pmatrix}.$$

(As a check, you'll see that $CC' = \Sigma$, as desired.) Thus, $c_{22} = \sqrt{5/3} = 1.291$, answer (c). \square

- 33. BONUS: How many trombones are there in the big parade?
 - (a) 0
 - (b) 1 (since it's the loneliest number)
 - (c) 2 (which can be as sad as 1)
 - (d) 13 (the unluckiest number)
 - (e) 38 (the average UGA IQ)
 - (f) 76

Solution: 76 (with 110 cornets close at hand), answer (f).