

# Efficient General Option Pricing

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## Abstract

The Black-Scholes/Merton model has been Nobel prize winning successful, but parameterizing models that fit market data has been a remarkably recalcitrant problem. This note explores using cumulants of the log of the the risk-neutral distribution of the underlying to achieve this.

Instead of working with the difference of the cumulants from a log-normal distribution as in Jarrow-Rudd [6] we work with the difference of the cumulants from a normal distribution. This is more natural since all but the first two the cumulants of a normal distribution are zero.

## 1 Outline

The forward value of an European put option having strike  $k$  and forward  $F$  at expiration is

$$E \max\{k - F, 0\} = E(k - F)1(F \leq k) = kP(F \leq k) - fP^*(F \leq k)$$

where  $f = EF$  and  $dP^*/dP = F/f$ . We derive formulas for  $P(F \leq k)$  and  $P^*(F \leq k)$  that can be computed efficiently when  $\log F$  is approximately normal.

In section 2 we recall some basic fact about cumulants. Section 3 uses an Edgeworth expansion to give an explicit formula for the cumulative distribution function of the perturbed distribution. Option valuation is considered in section 4. We show how the Esscher transform can be calculated using the same method as in section 3. The last section collects some general remarks.

## 2 Cumulants

The *cumulants*,  $(\kappa_n)$ , of a random variable  $X$  are defined by

$$\kappa(u) = \log Ee^{uX} = \sum_{n=0}^{\infty} \kappa_n \frac{u^n}{n!}$$

Taking  $u = 0$  yields  $\kappa_0 = 0$ . Since  $(d/du)^n \kappa(u)|_{u=0} = \kappa_n$  it is easy to work out that  $\kappa_1 = EX$  and  $\kappa_2 = \text{Var}(X)$ . Higher order cumulants are less intuitive.

Cumulants have some handy properties. The cumulants of a random variable plus a constant are the same except the first cumulant is increased by the constant. More generally, the cumulants of the sum of two independent random variables are the sums of their cumulants. They scale homogeneously, the  $n$ -th cumulant of a constant times a random variable is  $\kappa_n(cX) = c^n \kappa_n(X)$ .

Another nice property of cumulants is that they are more likely to exist than moments,  $m_n = EX^n$ . The relationship between cumulants and moments involves Bell polynomials[3].

$$Ee^{uX} = \sum_{n=0}^{\infty} m_n \frac{u^n}{n!} = \exp\left(\sum_{n=1}^{\infty} \kappa_n \frac{u^n}{n!}\right) = \sum_{n=0}^{\infty} B_n(\kappa_1, \dots, \kappa_n) \frac{u^n}{n!}$$

where  $B_n(\kappa_1, \dots, \kappa_n)$  is the  $n$ -th complete Bell polynomial. This is just a special case of the Faà di Bruno formula first proved by Louis François Antoine Arbogast in 1800[2]. Bell polynomials satisfy the recurrence [4]  $B_0 = 1$  and

$$B_{n+1}(x_1, \dots, x_{n+1}) = \sum_{k=0}^n \binom{n}{k} B_{n-k}(x_1, \dots, x_{n-k}) x_{k+1}.$$

### 2.1 Examples

(Normal). If  $X$  is normal then  $Ee^X = e^{EX + \text{Var}(X)/2}$  so  $Ee^{uX} = e^{uEX + u^2 \text{Var}(X)/2}$  showing the third and higher order cumulants vanish. If the cumulants of a random variable vanish after some point, then it must be normal[8] (Theorem 7.3.5). Something to keep in mind with computer implementations.

(Poisson). If  $X$  is Poisson with mean  $\mu$  then  $\kappa_n = \mu$  for all  $n$ .

(Exponential). If  $X$  is exponential with mean  $\mu$  then  $\kappa_n = (n-1)! \mu^n$ .

(Gamma). If  $X$  is gamma with mean  $\alpha/\beta$  and variance  $\alpha/\beta^2$  then  $\kappa_n = (n-1)! \alpha/\beta^n$ . Note gamma is exponential when  $\alpha = 1$  and  $\beta = 1/\mu$ .

### 3 Edgeworth Expansion

Given random variables  $X$  and  $Y$ , we have

$$\log Ee^{iuY} - \log Ee^{iuX} = \sum_{n=1}^{\infty} \Delta\kappa_n(iu)^n/n!$$

where  $(\Delta\kappa_n)$  are the differences of the cumulants of  $X$  and  $Y$ , so

$$Ee^{iuY} = Ee^{iuX} \sum_{n=0}^{\infty} B_n(\Delta\kappa_1, \dots, \Delta\kappa_n)(iu)^n/n!.$$

Let  $F$  and  $G$  be the cumulative distribution functions of  $X$  and  $Y$  respectively. Since the Fourier transform of  $F'$  is  $-iu\hat{F}(u)$ , the Fourier transform of the  $n$ -th derivative  $F^{(n)}$  is  $(-iu)^n\hat{F}(u)$ . Taking the inverse Fourier transform shows

$$G(x) = \sum_{n=0}^{\infty} (-1)^n \frac{B_n}{n!} F^{(n)}(x)$$

#### 3.1 Hermite Polynomials

The derivatives the standard normal cumulative distribution can be computed using Hermite polynomials[1] pp. 793–801. Recall the Hermite polynomials are the result of the Gram-Schmidt process applied to the basis  $(x^n)_{n \geq 0}$  on the Hilbert space having inner product  $(f, g) = \int_{-\infty}^{\infty} f(x)g(x) e^{-x^2/2} dx$ . The standard results are

$$H_n(x) = (-1)^n e^{x^2/2} \frac{d^n}{dx^n} e^{-x^2/2}$$

and they satisfy the recurrence  $H_0(x) = 1$ ,  $H_1(x) = x$  and

$$H_{n+1}(x) = xH_n(x) - nH_{n-1}(x).$$

Note some authors use  $He_n(x)$  instead of  $H_n(x)$ .

### 4 Option Valuation

The forward value of a put on a stock  $S$  with strike  $k$  is

$$E \max\{k - S, 0\} = kP(S \leq k) - ES P^*(S \leq k)$$

where  $dP^*/dP = S/ES$  is the *Esscher transform*[5]. This is the analog of  $E \max\{k - S, 0\} = kN(-d_2) - sN(-d_1)$  when  $S$  is lognormal.

## 4.1 Black Model

The Black model is  $S = S_t = se^{-\sigma^2 t/2 + \sigma B_t}$  where  $(B_t)$  is standard Brownian motion and  $t$  is time in years. Note  $ES = s$  and  $S \leq k$  if and only if  $B_t/\sqrt{t} \leq z$  where  $z = (\log k/s + \sigma^2 t/2)/\sigma\sqrt{t}$ . In the usual notation,  $z = -d_2$ .

Letting  $X = B_t/\sqrt{t}$  and so  $F = N$  is the standard normal cumulative distributing function we have  $P(X \leq z) = F(z)$  and  $P^*(X \leq z) = P(X + \sigma\sqrt{t} \leq z) = F(z - \sigma\sqrt{t})$  where we use the fact

$$Ee^U f(V) = Ee^U Ef(V + \text{Cov}(U, V))$$

if  $U$  and  $V$  are jointly normal. In the standard notation  $z - \sigma\sqrt{t} = -d_1$ .

## 4.2 Generalized Black Model

The generalized Black model is  $S = se^{-\kappa(t, \sigma) + \sigma X_t}$  where  $(X_t)$  is any stochastic process and  $\kappa(t, \sigma) = \log E \exp(\sigma X_t)$ . Note  $ES = s$  and  $S \leq k$  if and only if  $X_t/\sqrt{t} \leq (\log k/s + \kappa(t, \sigma))/\sigma\sqrt{t}$ .

Letting  $Y = X_t/\sqrt{t}$  and  $G$  be the corresponding cumulative distribution function,  $G(z) = P(Y \leq z)$ , we need to compute

$$G^*(z) = P^*(Y \leq z) = Ee^{-\kappa(t, \sigma) + \sigma\sqrt{t}Y} 1(Y \leq z).$$

Letting  $\gamma = \sigma\sqrt{t}$ , the cumulants of  $G^*$  can be found from

$$\begin{aligned} \log Ee^{uY^*} &= \log Ee^{-\kappa(t, \gamma) + \gamma Y} e^{uY} \\ &= -\kappa(t, \gamma) + \log Ee^{(\gamma+u)Y} \\ &= -\sum_{n=1}^{\infty} \kappa_n \frac{\gamma^n}{n!} + \kappa_n \frac{(u+\gamma)^n}{n!} \\ &= \sum_{n=1}^{\infty} \sum_{k=0}^{n-1} \kappa_n \binom{n}{k} \frac{\gamma^{n-k} u^k}{n!} \\ &= \sum_{k=1}^{\infty} \sum_{n=k+1}^{\infty} \kappa_n \binom{n}{k} \frac{\gamma^{n-k} u^k}{n!} \\ &= \sum_{k=1}^{\infty} \sum_{n=k+1}^{\infty} \kappa_n \frac{\gamma^{n-k}}{(n-k)!} \frac{u^k}{k!} \\ &= \sum_{k=1}^{\infty} \left( \sum_{n=1}^{\infty} \kappa_{n+k} \frac{\gamma^n}{n!} \right) \frac{u^k}{k!} \end{aligned}$$

This shows the cumulants of  $Y^*$  are  $\kappa_k^* = \sum_{n=1}^{\infty} \kappa_{n+k} \gamma^n / n!$  so we can also compute  $P(Y^* \leq z) = G^*(z)$  using the same technique as we did for  $G$ .

If  $G$  is standard normal then all cumulants vanish except  $\kappa_2 = 1$ . The only nonzero cumulants of  $G^*$  are  $\kappa_1 = \gamma$  and  $\kappa_2 = 1$ . The corresponding reduced Bell polynomials are  $B_n = \gamma^n / n!$  and we see the above formula is just the Taylor series expansion of  $G(z - \gamma)$ .

### 4.3 Lévy Processes

Kolmogorov's version of the Lévy-Khintchine theorem[7] states that if a random variable  $X$  is infinitely divisible there exists a number  $\gamma$  and a non-decreasing function  $G$  defined on the real line such that

$$\kappa(u) = \log Ee^{uX} = u\gamma + \int_{-\infty}^{\infty} K_u(x) dG(x),$$

where  $K_u(x) = (e^{ux} - 1 - ux)/x^2$ . Note the first cumulant of  $X$  is  $\gamma$  and for  $n \geq 2$ ,  $\kappa_n = \int_{-\infty}^{\infty} x^{n-2} dG(x)$ . In particular the variance of  $X$  is  $\int_{-\infty}^{\infty} dG(x) = G(\infty) - G(-\infty)$ .

If  $(X_t)$  is a Lévy process then  $X_1$  is infinitely divisible and  $\log Ee^{uX_t} = t\kappa(u)$ .

## 5 Remarks

The Gram-Charlier A series expands the quotients of cumulative distribution functions  $G/F$  using Hermite polynomials, but does not have asymptotic convergence, whereas the Edgeworth expansion involves the quotient of characteristic functions  $\hat{G}/\hat{F}$  in terms of cumulants and does have asymptotic convergence, ignoring some dainty facts [9].

A software implementation in C++ is available at <https://fmsgjr.codeplex.com> and Excel add-ins at <https://xllgjr.codeplex.com>.

## References

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