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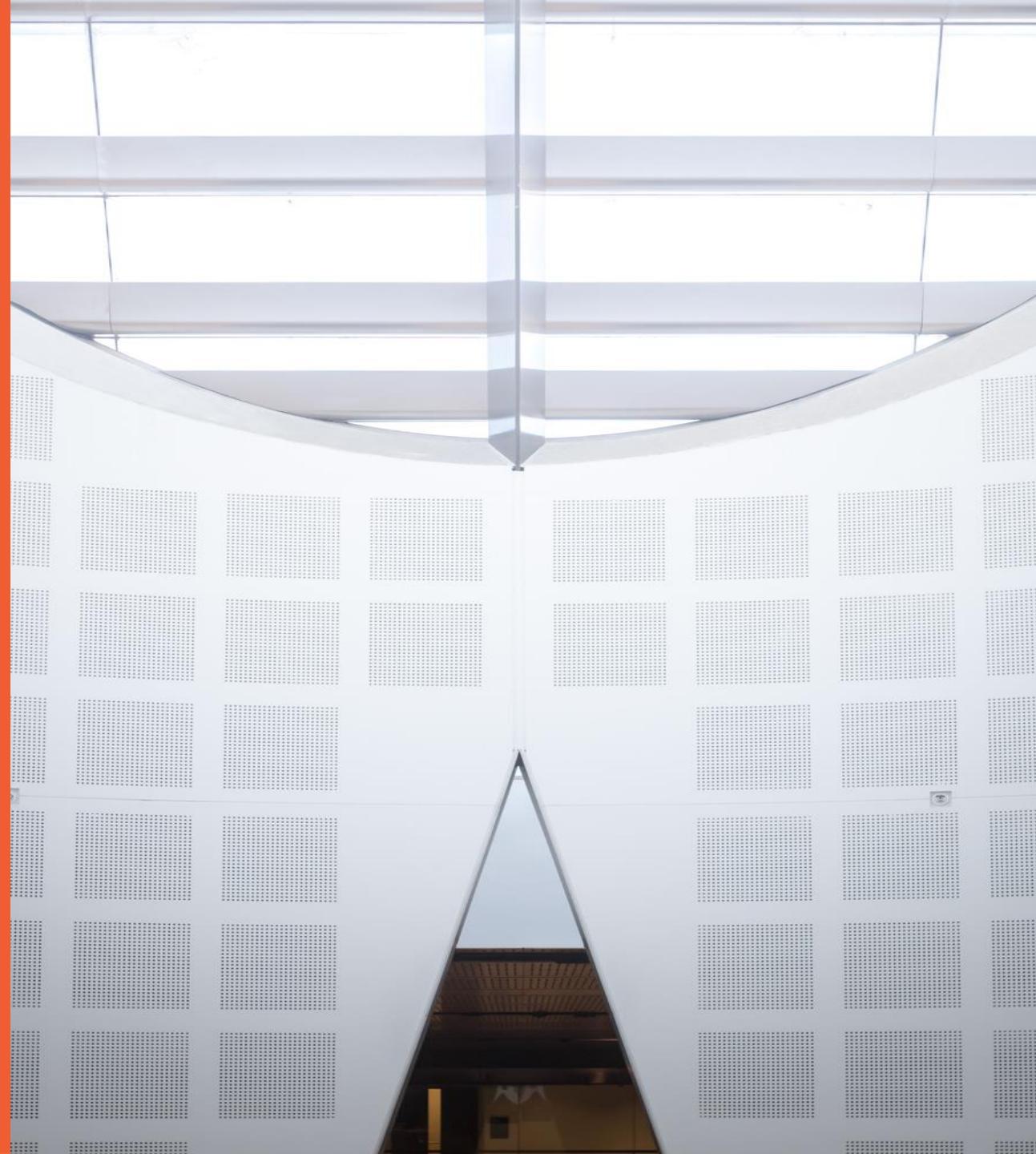
# COMPx270: Randomised and Advanced Algorithms

## Lecture 11: Learning and testing probability distributions

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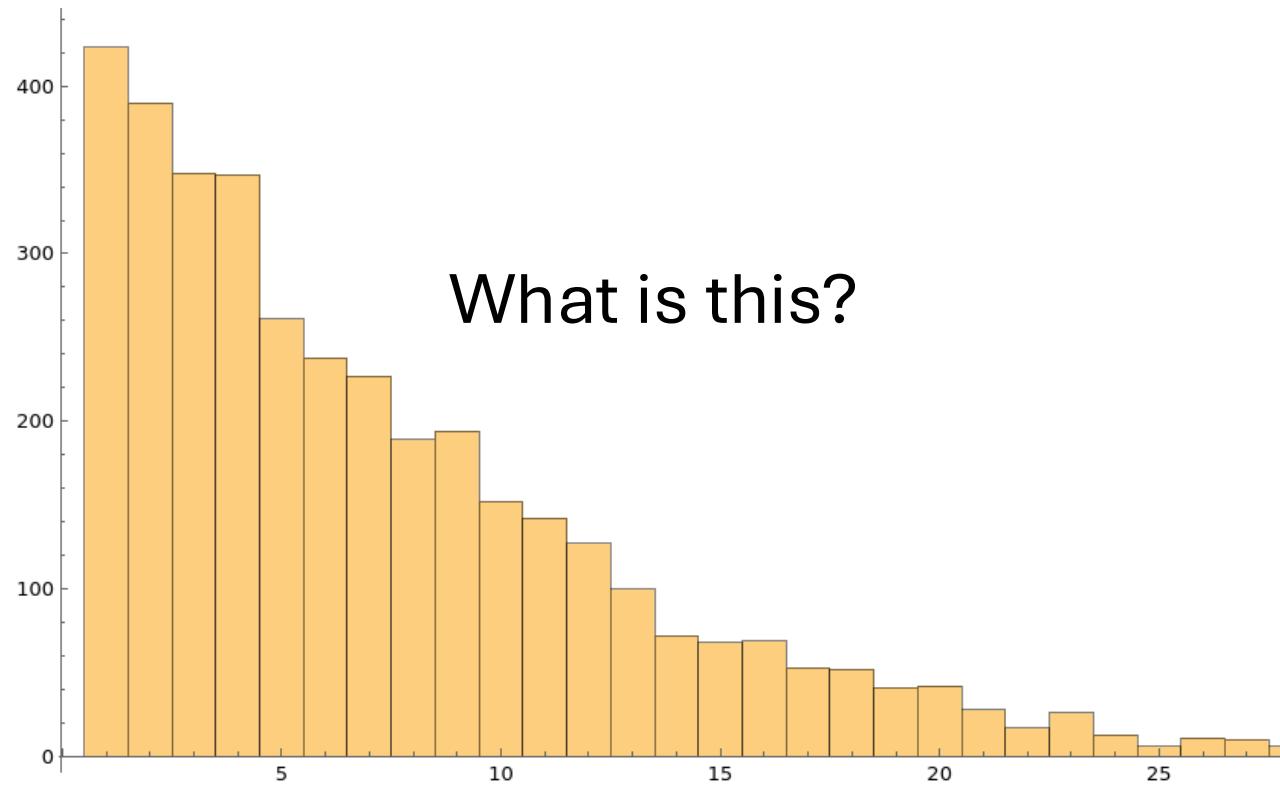
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## Some housekeeping

- A2 still being marked: deepest apologies (my fault)
- A3 (after Simple Extension) due tomorrow
- Don't forget the "participation" assignment (Oct 18)
- Sample exam is out, will be the topic of Week 13
- Feedback welcome: <https://forms.office.com/r/DymMcfn47n>
- Final exam on Tues, Nov 12 (9am) → what is allowed?

# A question



# Learning and testing (discrete) probability distributions \*

$p$  over  $X$  (discrete, finite domain) of size  $K$

- Learn  $p \rightsquigarrow$  get  $\hat{p}$  (distri over  $X$ ) st  $d(p, \hat{p}) \leq \varepsilon$   
( $K$  numbers)  
↑  
TBD
- Learn a parameter of  $p$  (estimation) : a number
- Learn a bit (test if  $p$  satisfies some property)

---

" $T^{BD}$ " :  $d(p, q) =$  total variation distance  
(quantifies how far  $p, q$  are)

Access  $x_1, \dots, x_n \stackrel{i.i.d}{\sim} p$  (algo only sees  $x_1, \dots, x_n$ )

\*  $p: X \rightarrow [0, 1]$

$$\sum_{i \in X} p(i) = 1$$

## Preliminaries on probability distributions

TV distance

$p, q$  over  $\mathcal{X}$

A metric would be nice

Bounded (in  $[0,1]$ ? ) also?

Meaning?

$$TV(p, q) = \sup_{S \subseteq \mathcal{X}} (p(S) - q(S)) = \sup_{S \subseteq \mathcal{X}} |p(S) - q(S)|$$

$$p(S) = \sum_{i \in S} p(i)$$

$$= \sup_{\substack{\text{any algo} \\ A: \mathcal{X} \rightarrow \{0,1\}}} \left( \Pr_{x \sim p}[A(x)=1] - \Pr_{x \sim q}[A(x)=1] \right)$$

## Preliminaries on probability distributions

Fact

$$TV(p, q) = \frac{1}{2} \sum_{i \in X} |p(i) - q(i)| = \frac{1}{2} \|p - q\|_1,$$

Pf

Take  $S^* = \{x : p(x) > q(x)\}$

$$TV(p, q) \geq p(S^*) - q(S^*) = \sum_{i \in S^*} (p(i) - q(i)) \geq \sum_{i \in S^*} |p(i) - q(i)|$$

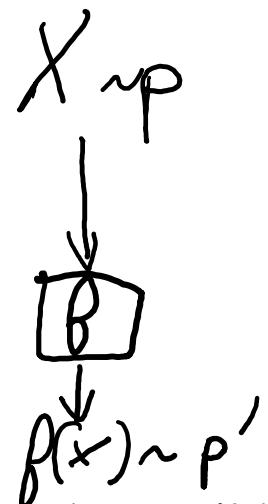
But  $\sum_{i \in X} |p(i) - q(i)| = \sum_{i \in S^*} (p(i) - q(i)) + \underbrace{\sum_{i \notin S^*} (q(i) - p(i))}_{\begin{aligned} &= 2 \sum_{i \in S^*} (p(i) - q(i)) \\ &= 2(p(S^*) - q(S^*)) \end{aligned}} + \underbrace{\sum_{i \notin S^*} q(i) - \sum_{i \notin S^*} p(i)}_{= 1 - q(S^*) - (1 - p(S^*))}$

## Preliminaries on probability distributions

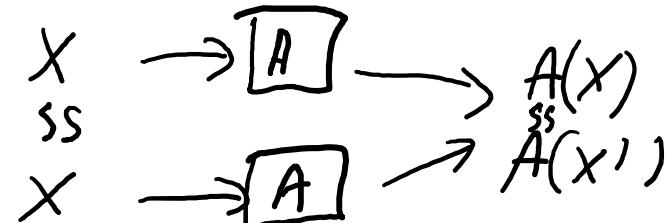
DPI (Data Processing Inequality)

Take any  $f: X \rightarrow Y$

If  $X \sim p$  let  $p'$  be the distn of  $f(X)$   
 $X' \sim q$  let  $q'$  —————  $f(X')$



$$TV(p', q') \leq TV(p, q)$$



## A view of TV distance

Alice and Bob play a game, where they both know two probability distributions  $\mathbf{p}, \mathbf{q}$ . Alice starts by tossing a fair coin, and does not show the outcome to Bob: if it is Heads, then she draws  $x \sim \mathbf{p}$ ; if it is Tails, she draws  $x \sim \mathbf{q}$ . Then she shows the value of  $x$  to Bob, who must guess if the coin toss was Heads. Clearly, just by random guessing, Bob can win the game with probability  $1/2$ . What the lemma says is that he can do better: there is a strategy for him to win with probability

$$\Pr[\text{Bob wins}] = \frac{1}{2} + \frac{d_{\text{TV}}(\mathbf{p}, \mathbf{q})}{2}$$

and, moreover, this is the best possible.

## The case of a coin □

How many times  $n$  do you need to flip the coin to learn its true bias  $p$  to accuracy  $\pm \varepsilon$ , and be correct with probability at least  $1 - \delta$ ?

## The case of a coin □

**Theorem 50.** Suppose we are promised that the true bias  $p$  of the coin satisfies  $0 \leq p < q \leq \frac{1}{2}$ , for some known value  $q$ . Then estimating the bias of the coin to an additive  $\varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{q}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)

# The case of a coin □

**Theorem 50.** Suppose we are promised that the true bias  $p$  of the coin satisfies  $0 \leq p < q \leq \frac{1}{2}$ , for some known value  $q$ . Then estimating the bias of the coin to an additive  $\varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{q}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)

**Corollary 50.1.** Estimating the bias of a coin to an additive  $\varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)

## The case of a coin $\square$

**Theorem 50.** Suppose we are promised that the true bias  $p$  of the coin satisfies  $0 \leq p < q \leq \frac{1}{2}$ , for some known value  $q$ . Then estimating the bias of the coin to an additive  $\varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{q}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)

$$\hat{p} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\Pr[|\hat{p} - p| > \varepsilon] \underset{\text{Hoeffding}}{\leq} 2e^{-2\varepsilon^2 n} \underset{\text{want}}{\leq} \delta$$

$$\text{take } n = \left\lceil \frac{1}{2\varepsilon^2} \ln \frac{2}{\delta} \right\rceil$$

$$\underset{\text{Chernoff}}{\leq} 2e^{-\frac{\varepsilon^2 q}{3} n} \underset{\text{want}}{\leq} \delta$$

(if  $q > \varepsilon$   
(if not, easy...))

$$n = \left\lceil \frac{3q}{\varepsilon^2} \right\rceil \ln \frac{2}{\delta}$$

## The case of a coin: what about testing? □

Is my coin a fair coin?

## The case of a coin: what about testing? $\square$

**Theorem 51.** Testing whether the bias of a coin is  $1/2$  or at least  $1/2 + \varepsilon$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples. (Moreover, this is optimal.)

Not better than learning  
 $p$ .

## The case of a coin: what about testing? $\square$

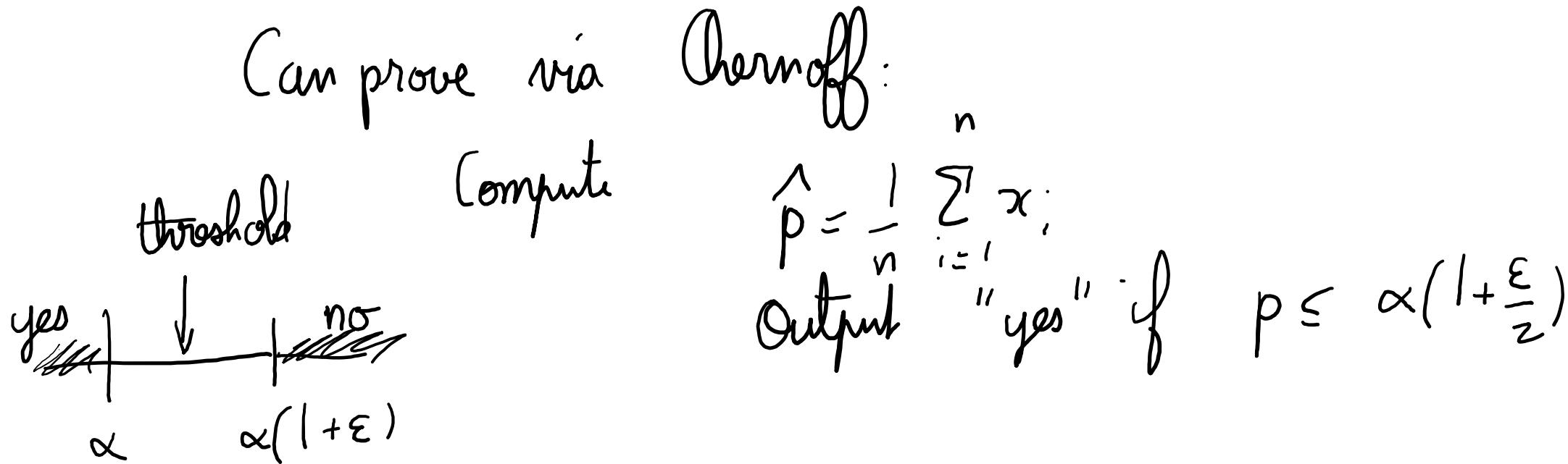
**Theorem 52.** For any  $0 < \alpha \leq 1/2$  and  $\varepsilon \in (0, 1]$ , testing whether the bias of a coin is at most  $\alpha$  or at least  $\alpha(1 + \varepsilon)$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\alpha\varepsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples.

$$\alpha = \frac{1}{2} \rightarrow \text{previous theorem}$$

$$\alpha \text{ small}, \quad \varepsilon = 1 \quad n \leq \frac{\alpha}{\delta} \geq 2\alpha : O\left(\frac{1}{\alpha} \log \frac{1}{\delta}\right)$$

# The case of a coin: what about testing? $\square$

**Theorem 52.** For any  $0 < \alpha \leq 1/2$  and  $\epsilon \in (0, 1]$ , testing whether the bias of a coin is at most  $\alpha$  or at least  $\alpha(1 + \epsilon)$ , with probability at least  $1 - \delta$ , can be done with  $n = O\left(\frac{1}{\alpha\epsilon^2} \log \frac{1}{\delta}\right)$  i.i.d. samples.



## Beyond coins: $k$ is large

Domain sizes grow quite fast, and in most settings  $k$  is huge.

## Learning in TV distance

Unknown  $p$  over  $\mathcal{X}$  ( $|\mathcal{X}| = k$ )

Parameter  $\varepsilon, \delta$

Get  $x_1, \dots, x_n \stackrel{\text{i.i.d}}{\sim} p$  for  $n$  to be chosen

Goal: output  $\hat{p}$  (distrib. over  $\mathcal{X}$ )  
such that

$$\Pr[\text{TV}(p, \hat{p}) > \varepsilon] \leq \delta$$

Minimise  $\frac{n}{\zeta}$  (sample complexity)  
 $\hookrightarrow n(k, \varepsilon, \delta)$

## Learning in TV distance: first attempt

$$p = (p_1, p_2, \dots, p_k)$$

Want:  $\hat{p} \in [0,1]^k$  st

$$\Delta \quad \|\hat{p} - p\|_1 \leq 2\epsilon \quad (\omega/p \geq 1 - \delta)$$

(If  $\|\hat{p}\|_1 \neq 1$ ,  
can normalize)

\* implies

$$|\|\hat{p}\|_1 - \underbrace{\|p\|_1}_{=1}| \leq 2\epsilon$$

$$\sum_{i=1}^k |\hat{p}_i - p_i|$$

Enough to estimate each  $p_i$  by  $\hat{p}_i$  to  $\pm \frac{2\epsilon}{k}$

w/p  $1 - \frac{\delta}{k}$  each

for instances of "estimate bias  $p_i$  of a coin"

Gives us

$$n = O\left(\frac{1}{(\epsilon/k)^2} \log \frac{1}{\delta/k}\right) = O\left(\frac{k^2}{\epsilon^2} \log \frac{k}{\delta}\right)$$

## Learning in TV distance: second attempt

What if instead of  $\hat{p}_i \approx p_i \pm \frac{2\epsilon}{k}$ , we try  $\hat{p}_i \approx (1 \pm 2\epsilon)p_i$ .

Then  $\|\hat{p} - p\|_1 = \sum_{i=1}^k |p_i - \hat{p}_i| \leq \sum_{i=1}^k 2\epsilon p_i = 2\epsilon$

$\nexists p \min_{1 \leq i \leq k} p_i \geq \frac{\epsilon}{k}$ , we can (if no lower bound on  $p_i$ , can't)

Chernoff bound + union bound ( $\delta' = \frac{\delta}{k}$ )

$$\left( \leq e^{-\epsilon^2 \cdot \frac{\delta}{k} \cdot n} \right)$$

when using Chernoff

$$n = O\left(\frac{k}{\epsilon^3} \log \frac{k}{\delta}\right)$$

A replace  $p$   
 by  
 $p' = (1 - \frac{\epsilon}{2})p + \frac{\epsilon}{2}u_k$   
 Learn  $p'$  to TV  $\frac{\epsilon}{2}$

## Learning in TV distance: third attempt

**Theorem 53.** Learning an unknown distribution  $\mathbf{p} \in \Delta(\textcolor{teal}{k})$  to total variation distance  $\textcolor{violet}{\epsilon}$  (with success probability  $1 - \delta$ ) can be done with

$$\textcolor{red}{n} = O\left(\frac{\textcolor{teal}{k} + \log \frac{1}{\delta}}{\textcolor{violet}{\epsilon}^2}\right)$$

i.i.d. samples. (Moreover, this is optimal.)

The empirical estimator works  
 $\hat{p}_i = \frac{N_i}{n}$   $\leftarrow \begin{array}{l} \# \text{ of times } i \text{ is seen} \\ \text{in the } n \text{ samples} \end{array}\right.$

$$n = O\left(\frac{k + \log \frac{1}{\delta}}{\epsilon^2}\right)$$

i.i.d. samples. (Moreover, this is optimal.)

## Learning in TV distance: third attempt

Want  $TV(p, \hat{p}) \leq \epsilon$

$$TV(p, \hat{p}) > \epsilon \Leftrightarrow \exists S \subseteq \mathcal{X}, |p(S) - \hat{p}(S)| > \epsilon$$

Fix any  $S \subseteq \mathcal{X}$ .

$$\Pr[|p(S) - \hat{p}(S)| > \epsilon] \leq 2e^{-2\epsilon^2 n} \xrightarrow{\text{want}} \frac{\delta}{2^k}$$

this ↑  
is Hoeffding.  
Bias of a coin!

$$p(S) = \sum_{i \in S} p_i$$

$$\begin{aligned} \hat{p}(S) &= \sum_{i \in S} \hat{p}_i \\ &= \underbrace{\# \text{samples falling in } S}_{n} \end{aligned}$$

$$\begin{aligned} n &= \lceil \frac{\ln(\frac{2}{\delta})}{2\epsilon^2} \rceil \\ &= O\left(\frac{k + \log(\frac{1}{\delta})}{\epsilon^2}\right) \end{aligned}$$

+ union bound over  $2^k$  subsets.

**Theorem 53.** Learning an unknown distribution  $\mathbf{p} \in \Delta(k)$  to total variation distance  $\epsilon$  (with success probability  $1 - \delta$ ) can be done with

$$n = O\left(\frac{k + \log \frac{1}{\delta}}{\epsilon^2}\right)$$

i.i.d. samples. (Moreover, this is optimal.)

## Learning in TV distance: second third attempt

$$\begin{aligned} \mathbb{E}[\text{TV}(\mathbf{p}, \hat{\mathbf{p}})] &= \frac{1}{2} \mathbb{E}[\|\mathbf{p} - \hat{\mathbf{p}}\|_1] = \frac{1}{2} \sum_{i=1}^k \mathbb{E}[|p_i - \hat{p}_i|] \\ &\stackrel{\text{Jensen}}{\leq} \frac{1}{2} \sum_{i=1}^k \sqrt{\mathbb{E}[(p_i - \hat{p}_i)^2]} = \frac{1}{2n} \sum_{i=1}^k \sqrt{\underbrace{\text{Var}(n\hat{p}_i)}_{np_i(1-p_i) \leq np_i}} \\ &\stackrel{f(x)=x^2}{\leq} \frac{1}{2\sqrt{n}} \sum_{i=1}^k \sqrt{p_i} \stackrel{\text{Jensen}}{\leq} \frac{\sqrt{k}}{2\sqrt{n}} \end{aligned}$$

$$\hat{p}_i = p_i$$

$$n\hat{p}_i \sim \text{Bin}(n, p_i)$$

$$\stackrel{\text{want}}{\leq} \frac{\epsilon}{10}$$

$$\text{True for } n \geq \frac{25k}{\epsilon^2} + \text{Markov.}$$

$$\begin{aligned} b. \frac{1}{n} \sum_{i=1}^k \sqrt{p_i} &= k \underbrace{\frac{1}{k} \sum_{i=1}^k \sqrt{p_i}}_{\text{inu}} \\ &\leq k \sqrt{\underbrace{\frac{1}{k} \sum_{i=1}^k p_i}_{\frac{1}{n}}} = \sqrt{k} \end{aligned}$$

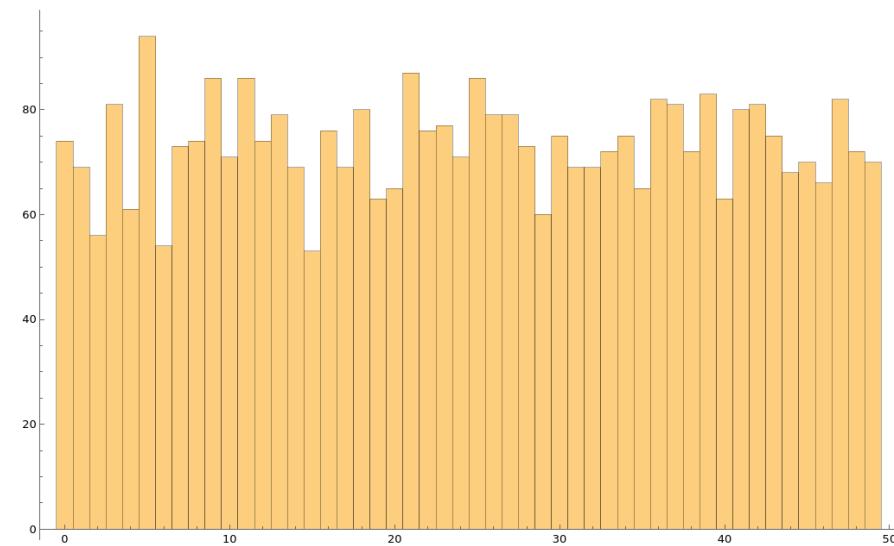
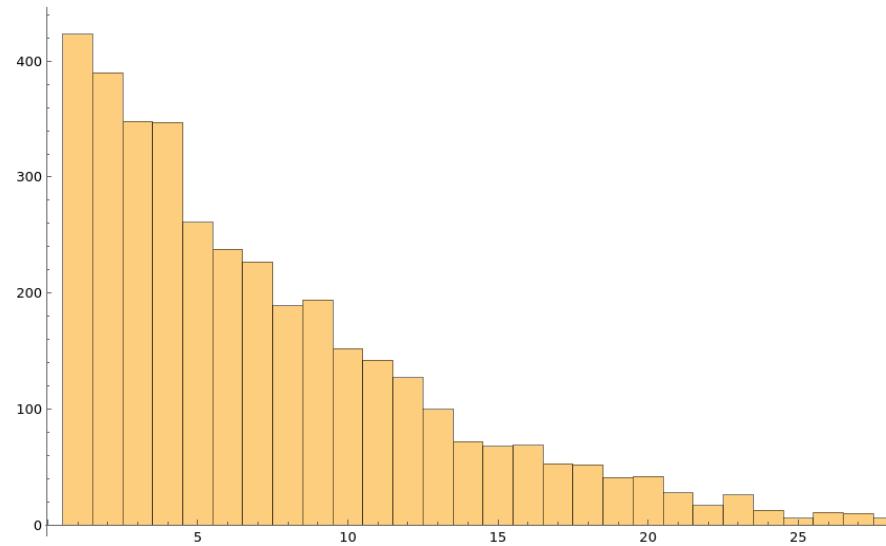
# Learning in TV distance: second third attempt

**Theorem 53.** Learning an unknown distribution  $\mathbf{p} \in \Delta(\textcolor{teal}{k})$  to total variation distance  $\varepsilon$  (with success probability  $1 - \delta$ ) can be done with

$$\textcolor{red}{n} = O\left(\frac{\textcolor{teal}{k} + \log \frac{1}{\delta}}{\varepsilon^2}\right)$$

i.i.d. samples. (Moreover, this is optimal.)

# Testing in TV distance



# Testing in TV distance: identity testing

Give an algorithm  $A$  which takes parameters  $\varepsilon, \delta \in (0, 1]$  and  $n$  samples from  $p$ , and:

- If  $p = q$ , then  $\Pr[A \text{ outputs yes}] \geq 1 - \delta$ ;
- If  $d_{\text{TV}}(p, q) > \varepsilon$ , then  $\Pr[A \text{ outputs no}] \geq 1 - \delta$

(if  $0 < d_{\text{TV}}(p, q) \leq \varepsilon$ , then  $A$  is off the hook and can output whatever).



## Testing in TV distance: identity testing via learning

$$n = O\left(\frac{k + \log(1/\delta)}{\epsilon^2}\right) \text{ is an upper bound}$$

$p \rightarrow \hat{p}$  st  $\text{TV}(p, \hat{p}) \leq \frac{\epsilon}{2}$   $\rightarrow$  check if  $\text{TV}(\hat{p}, q) \leq \frac{\epsilon}{2}$   $\rightarrow$  if not, say "no"

Key:  $\text{TV}(p, q) \leq \text{TV}(p, \hat{p}) + \text{TV}(\hat{p}, q)$

# Testing in TV distance: uniformity is all you need

**Theorem 54** (Identity to uniformity reduction). Suppose there is an algorithm  $A$  for uniformity testing, which takes  $n = n(k, \varepsilon, \delta)$  i.i.d. samples from the unknown distribution. Then there is an algorithm  $A'$  for identity testing over a domain of size  $k$  to any fixed  $\mathbf{q} \in \Delta(k)$ , which takes  $n = n(4k, \varepsilon/4, \delta)$  i.i.d. samples from the unknown distribution. Moreover,  $A'$  is efficient if  $A$  is.

Upshot: "wlog,  $p = u_h$  vs  $u$ "  
 $TV(p, u_h) > \varepsilon$   
(no need to worry about other  $q$ )

## Testing in TV distance: uniformity testing

Uniform is simpler

(Focus on  $S = \frac{1}{3}$ )

Birthday paradox:

If  $p$  is uniform on  $\frac{k}{2}$  elements  
vs  $p$  is the uniform dist (over  $k$  elements)

$$n = \Omega(\sqrt{k})$$

## Testing in TV distance: uniformity testing

**Theorem 55.** Testing uniformity of an unknown distribution  $\mathbf{p} \in \Delta(\textcolor{blue}{k})$  to total variation distance  $\textcolor{green}{\varepsilon}$  (with success probability  $2/3$ ) can be done with

$$\textcolor{red}{n} = O\left(\frac{\sqrt{\textcolor{blue}{k}}}{\textcolor{green}{\varepsilon}^2}\right)$$

i.i.d. samples, using Algorithm 21. (Moreover, this is optimal for constant success probability.)

## Testing in TV distance: uniformity testing, key ideas

$$TV \rightarrow \ell_2$$

$$\|p - u_h\|_1 \leq \sqrt{k} \|p - u_h\|_2$$

so if  $p = u_h$ ,  $\|p - u_h\|_2 = 0$

if  $TV(p, u_h) > \varepsilon$ ,  $\|p - u_h\|_2 > \frac{2\varepsilon}{\sqrt{k}}$

## Testing in TV distance: uniformity testing, key ideas

$$\textcircled{2} \quad \|p - u_h\|_2^2 = \|p\|_2^2 - \frac{1}{k}$$

$$\sum_i (p_i - \frac{1}{k})^2 = \sum_i p_i^2 - \frac{2}{k} \left( \sum_i p_i \right) + \frac{1}{k} = \|p\|_2^2 - \frac{1}{k}$$

- $\|p\|_2^2 = \frac{1}{k}$  if  $p = u_h$
- $\|p\|_2^2 > \frac{1 + 4\epsilon^2}{k} \iff TV(p, u_h) > \epsilon$

$$\|p\|_2^2 = \Pr_{x, y \sim p} [x = y]$$

$$\underbrace{\sum_i \Pr_{x, y \sim p} [x = y = i]}_{P_i \cdot P_i}$$

## Testing in TV distance: uniformity testing, algorithm

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**Input:** Multiset of  $n$  i.i.d. samples  $x_1, \dots, x_n \in \mathcal{X}$ , parameters  $\varepsilon \in (0, 1]$  and  $k = |\mathcal{X}|$

1: Set  $\tau \leftarrow \frac{1+2\varepsilon^2}{k}$

2: Compute  $\triangleright O(n)$  time if  $\mathcal{X}$  is known

$$Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}} = \frac{1}{\binom{n}{2}} \sum_{j \in \mathcal{X}} \binom{N_j}{2}$$

where  $N_j \leftarrow \sum_{t=1}^n \mathbb{1}_{\{x_t=j\}}$ .

3: **if**  $Z \geq \tau$  **then return** no  $\triangleright$  Not uniform

4: **else return** yes  $\triangleright$  Uniform

---

## Testing in TV distance: uniformity testing

$$\textcircled{1} \quad \mathbb{E} Z = \|p\|_2^2 \quad (\text{easy}) \quad \left. \right\}$$

$$\textcircled{2} \quad \text{Var } Z < \text{small} \quad (\text{really hard}) + \text{Chebyshev} \quad \left. \right\}$$

"Easy": can prove a bound on  $\text{Var } Z$

giving  
 $n = O\left(\frac{k}{\varepsilon^4}\right)$

Getting  $O\left(\frac{\sqrt{k}}{\varepsilon^2}\right)$  is hard

<b>Input:</b> Multiset of $n$ i.i.d. samples $x_1, \dots, x_n \in \mathcal{X}$ , parameters $\varepsilon \in (0, 1]$ and $k =  \mathcal{X} $
1: Set $\tau \leftarrow \frac{1+2\varepsilon^2}{k}$
2: Compute
$Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s=x_t\}} = \frac{1}{\binom{n}{2}} \sum_{j \in \mathcal{X}} \binom{N_j}{2}$
where $N_j \leftarrow \sum_{t=1}^n \mathbb{1}_{\{x_t=j\}}$ .
3: if $Z \geq \tau$ then return no
4: else return yes
▷ Not uniform
▷ Uniform

# Testing in TV distance: uniformity testing

**Input:** Multiset of  $n$  i.i.d. samples  $x_1, \dots, x_n \in \mathcal{X}$ , parameters  $\varepsilon \in (0, 1]$  and  $k = |\mathcal{X}|$   
1: Set  $\tau \leftarrow \frac{1+2\varepsilon^2}{k}$   
2: Compute  $Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}}$   $= \frac{1}{\binom{n}{2}} \sum_{j \in \mathcal{X}} \binom{N_j}{2}$

where  $N_j \leftarrow \sum_{t=1}^n \mathbb{1}_{\{x_t=j\}}$ .  
3: if  $Z \geq \tau$  then return no ▷ Not uniform  
4: else return yes ▷ Uniform

# Testing in TV distance: uniformity testing

**Input:** Multiset of  $n$  i.i.d. samples  $x_1, \dots, x_n \in \mathcal{X}$ , parameters  $\varepsilon \in (0, 1]$  and  $k = |\mathcal{X}|$   
1: Set  $\tau \leftarrow \frac{1+2\varepsilon^2}{k}$   
2: Compute  $Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}}$   $\Rightarrow O(n)$  time if  $\mathcal{X}$  is known

$$Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}} = \frac{1}{\binom{n}{2}} \sum_{j \in \mathcal{X}} \binom{N_j}{2}$$

where  $N_j \leftarrow \sum_{t=1}^n \mathbb{1}_{\{x_t=j\}}$ .

- 3: if  $Z \geq \tau$  then **return no**  $\triangleright$  Not uniform  
4: **else return yes**  $\triangleright$  Uniform

# Testing in TV distance: uniformity testing

**Input:** Multiset of  $n$  i.i.d. samples  $x_1, \dots, x_n \in \mathcal{X}$ , parameters  $\varepsilon \in (0, 1]$  and  $k = |\mathcal{X}|$   
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where  $N_j \leftarrow \sum_{t=1}^n \mathbb{1}_{\{x_t=j\}}$ .

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4: **else return yes**  $\triangleright$  Uniform

# Testing in TV distance: uniformity testing

**Input:** Multiset of  $n$  i.i.d. samples  $x_1, \dots, x_n \in \mathcal{X}$ , parameters  $\varepsilon \in (0, 1]$  and  $k = |\mathcal{X}|$   
1: Set  $\tau \leftarrow \frac{1+2\varepsilon^2}{k}$   
2: Compute  $Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}}$   $\Rightarrow O(n)$  time if  $\mathcal{X}$  is known

$$Z = \frac{1}{\binom{n}{2}} \sum_{1 \leq s < t \leq n} \mathbb{1}_{\{x_s = x_t\}} = \frac{1}{\binom{n}{2}} \sum_{j \in \mathcal{X}} \binom{N_j}{2}$$

where  $N_j \leftarrow \sum_{t=1}^n \mathbb{1}_{\{x_t=j\}}$ .

- 3: if  $Z \geq \tau$  then **return no**  $\triangleright$  Not uniform  
4: **else return yes**  $\triangleright$  Uniform

## Testing in TV distance: uniformity testing, summary

**Theorem 55.** Testing uniformity of an unknown distribution  $\mathbf{p} \in \Delta(\textcolor{blue}{k})$  to total variation distance  $\varepsilon$  (with success probability  $2/3$ ) can be done with

$$\textcolor{red}{n} = O\left(\frac{\sqrt{\textcolor{blue}{k}}}{\varepsilon^2}\right)$$

i.i.d. samples, using Algorithm 21. (Moreover, this is optimal for constant success probability.)

Tight bound (other algo)

$$n = O\left(\frac{\sqrt{k \log \frac{1}{\delta}} + \log \frac{1}{\delta}}{\varepsilon^2}\right)$$

# Summary

- Learning  $p$

$\tilde{V}_1$

$$\boxed{\frac{k}{\varepsilon^2}}$$

- Testing

$$p = u_h$$

vs

$$TV(p, u_h) > \varepsilon$$

$$\boxed{\frac{\sqrt{k}}{\varepsilon^2}}$$

- Estimating TV distance

$$TV(p, u_k)$$

$$\boxed{\frac{k}{\varepsilon^2 \log k}}$$