Climate change reshapes the drivers of false spring risk across European trees

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$_{\scriptscriptstyle 16}$ Abstract

Temperate and boreal forests are at risk of late spring freezing events after budburst—also known as false springs. Research to date has generated conflicting results of whether climate change will decrease false springs, and thus reshape a fundamental factor that influences species' ranges. Conflicting results may be due to the myriad climatic and geographic factors that contribute to a plant's risk of a false spring, which—to date—no study has compared at once. Here, we assessed the effects of mean spring temperature, distance from the coast, elevation and the North Atlantic Oscillation (NAO) using PEP725 leafout data for six tree species across 11,648 sites in Europe, to determine which were the strongest predictors of false spring risk and how these predictors shifted with climate change. Across species, mean spring temperature and distance from the coast were the strongest predictors, with higher mean spring temperatures having a decreased risk in false springs (-7.64% for every 2°C increase) and sites further from the coast experiencing an increased risk in false springs (5.32% for every 150km from the coast). Elevation (2.23% for every 200m increase in elevation) and

NAO index (1.91% for every 0.3 increase) also increased false spring risk. The effects of these major drivers
have been reshaped by climate change and are influencing false spring risk in new ways, with the climatic
factors shifting considerably compared to the geographic factor hough false spring risk is decreasing across
species, the residual, unexplained effects of climate change have magnified the variation in risk across species,
with a split among early- versus late-leafout specie ur results suggest that considering multiple spatial
and climatic factors is essential for predicting false spring risk and having a better understanding of the
unexplained complexities of climate change is critical.

Introduction

Temperate tree and shrub species are at risk of damage from late spring freezing events after budburst, also known as false springs, and this risk may shift with climate change. With earlier springs due to warming (IPCC, 2015; Wolkovich et al., 2012), the growing season is lengthening across many regions in the northern 38 hemisphere (Chen et al., 2005; Kukal & Irmak, 2018; Liu et al., 2006). Longer growing seasons could translate to increased plant growth, assuming such increases are not offset by tissue losses due to false springs. Last spring freeze dates are not predicted to advance at the same rate as warming (Inouye, 2008; Labe et al., 2016; Martin et al., 2010; Wypych et al., 2016b; Sgubin et al., 2018), potentially amplifying the effects of false spring events in some regions. In Germany, for example, the last freeze date has advanced by 2.6 days per decade since 1955 (Zohner et al., 2016), but budburst has advanced roughly twice as fas lajor false spring events have been recorded in recent years but studies have variously found that spring freeze damage may increase (Augspurger, 2013; Hänninen, 1991; Labe et al., 2016), remain the same (Scheifinger et al., 2003) or even decrease (Kramer, 1994; Vitra et al., 2017) with climate change. Regardless, studies have found it can take 16-38 days for trees to refoliate after a freeze (Augspurger, 2009, 2013; Gu et al., 2008; Menzel et al., 2015), which can detrimentally affect crucial processes such as carbon uptake and nutrient cycling (Hufkens et al., 2012; Klosterman et al., 2018; Richardson et al., 2013). Spring freezes are one of the largest limiting factors to species ranges and have greatly shaped plant life history strategies (Kollas et al., 2014). Temperate plants are exposed to freezing temperatures numerous times throughout the year, however, individuals are most at risk to damage in the spring, when freeze 53 tolerance is lowest (Sakai & Larcher, 1987). Plants have adapted to these early spring risks through various mechanisms with one common strategy being avoidance (Vitasse et al., 2014). Many temperate species

minimize freeze risk and optimize growth by using a complex mix of cues to initiate budburst: low winter temperatures (i.e., chilling), warm spring temperatures (i.e., forcing), and increasing spring daylengths (i.e., photoperiod). With climate change advancing, the interaction of these cues may shift spring phenologies both across and within species and sites, making some species less—or more—vulnerable to false springs than before. Species that leafout first each spring are especially at risk of false springs, as their budburst occurs during times of year when the risk of freeze events is relatively high. To date these species also appear to advance the most with warming (Wolkovich et al., 2012), thus, if climate change increases the prevalence of late spring freezes, we would expect these species to see major increases in false spring risk. If climate change has restructured the timing and prevalence of false springs to later in the spring, then later-leafout species may experience major increases in false spring risk with climate change.

Some research suggests false spring incidence has already begun to decline in many regions (i.e. across parts of North America and Asia), however the prevalence of spring frost consistently increased across Europe since 1982 (Liu et al., 2018). Furthermore, recent studies have demonstrated site-specific effects may be more closely related to false spring risk: whether via elevation, where higher elevations appear at higher risk (Ma et al., 2018; Vitra et al., 2017), or distance from the coast, where inland areas appear at higher risk (Ma et al., 2018; Wypych et al., 2016b). Improved understanding of which regional climatic factors impact false spring risk, including which factors are most crucial for predicting risk, we may be able to determine which regions are at risk currently and which regions will be more at risk in the future.

The majority of false spring studies assess the effects of one predictor (e.g. temperature, elevation or distance from the coast) on false spring prevalence, thus failing to compare how multiple factors may together shape risk. False spring risk is influenced by multiple climatic and geographic factors, which may vary across species and time. Further, because predictors can co-vary—for example, higher elevation sites are often more distant from the coast—the best predictions of false springs should examine all predictors at once.

Here we investigate the influence of known spatial and climatic factors on false spring risk (defined here as
when fell temperatures below -2.2° between estimated budburst and leafout Schwartz, 1993). We assessed
the number of false springs that occurred across 11,648 sites across Europe using observed phenological
data (754,786 observations) for six temperate, deciduous trees and combined that with daily gridded climate
data for each site that extended from 1951-2016. We focus on the major factors shown or hypothesized to
influence false spring risk: mean spring temperature, elevation, distance from the coast, and a major climatic
oscillation that structures European climate—the North Atlantic Oscillation (NAO). The NAO is tied to

winter and spring circulation across Europe, with more positive NAO phases tending to result in higher than
average winter and spring temperatures. With climate-change induced shifts, years with higher NAO indices
have correlated to even earlier budburst dates since the late 1980s in some regions (Chmielewski & Rötzer,
2001), however little research has tested if more positive NAO phases also translates into more false springs.
We aimed to understand which factors are the strongest predictors of false spring risk, and how the major
predictors have shifted with climate change

We obtained phenological data from the Pan European Phenology network (PEP725, www.pep725.edu),

which provides open access phenology records across Europe (Templ et al., 2018). Since plants are most

$_{\scriptscriptstyle{92}}$ Methods

93 Phenological Data and Calculating Vegetative Risk

susceptible to damage from freezing temperatures between budburst and full leafout, we selected leafout data (i.e., in Meier, 2001, BBCH 11, which is defined as the point of leaf unfolding and the first visible leaf stalk) from the PEP725 dataset. The species used in the study were Aesculus hippocastanum Poir., Alnus glutinosa (L.) Gaertn., Betula pendula Roth., Fagus sylvatica Ehrh., Fraxinus excelsior L., and Quercus robur L. Selection criteria for the species were as follows: (1) to be temperate, deciduous species that were not cultivars or used as crops, (2) there were at least 90,000 observations of BBCH 11 (leafout), (3) to represent 101 over half of the total number of sites available (11,684), and (4) there were observations for at least 65 out of 102 the 66 years of the study (1951-2016) (Table S1). 103 Plants are generally the most freeze tolerant in the winter but this freeze tolerance greatly diminishes once individuals exit the dormancy phase (i.e. processes leading to budburst) through full leaf expansion (Lenz 105 et al., 2016; Vitasse et al., 2014). Thus, for most individuals that initiate budburst and have not fully leafed out before the last spring freeze are at risk of leaf tissue loss, damage to the xylem, and slowed canopy 107 development (Gu et al., 2008; Hufkens et al., 2012). To capture this 'high-risk' timeframe, we subtracted 12 108 days from the leafout date to establish a standardized estimate for day of budburst (Donnelly et al., 2017; 109 Flynn & Wolkovich, 2018; USA-NPN, 2019) since the majority of the individuals were missing budburst 110 observation We additionally considered a model that altered the timing between budburst and leafout 111 for each species. For this alternate model, we calculated budburst by subtracting 11 days from leafout 112 for Aesculus hippocastanum and Betula pendula, 12 days for Alnus glutinosa, 5 days for Fagus sylvatica,

and 7 days for both *Fraxinus excelsior* and *Quercus robur* based on growth chamber experiment data from phylogenetically related species (Buerki *et al.*, 2010; Wang *et al.*, 2016; Hipp *et al.*, 2017; Flynn & Wolkovich, 2018).

117 Climate Data

We collected daily gridded climate data from the European Climate Assessment & Dataset (ECA&D) and used the E-OBS 0.25 degree regular latitude-longitude grid from version 16. We used the daily minimum 119 temperature dataset to determine if a false spring occurred. False springs in this study were defined as temperatures at or below -2.2°C (Schwartz, 1993) between budburst to leafout. We additionally tested this 121 model by changing the definition of a freezing temperature from -2.2°C (Schwartz, 1993) to -5°C (Lenz et al., 2013; Sakai & Larcher, 1987) in an additional model. In order to assess regional climatic effects we calculated 123 the mean spring temperature by using the daily mean temperature from March 1 through May 31. We used 124 this date range to best capture temperatures likely after chilling had accumulated to compare differences in 125 spring forcing temperatures across sites (Basler & Körner, 2012; Körner et al., 2016). We collected NAO-126 index data from the KNMI Climate Explorer CPC daily NAO time series and selected the NAO indices from 127 November until April to capture the effects of NAO on budburst for each region and then took the mean 128 NAO index during these months (KNMI, 2018). Since the primary aim of the study is to predict false spring incidence in a changing climate, we split the data: before temperature trends increased (1951-1983) and after 130 trends increased (1984-2016, Kharouba et al., 2018; Stocker et al., 2013) to represent climate change and which will be referred to as the 'climate change' parameter henceforth. 132

133 Data Analysis

134 Simple regression models

We initally ran three simple regression models—following the same equation (below) but with varying responses—to assess the effects of climate change on budburst, minimum temperatures between budburst and leafout and the number of false springs across species (Equation 2).

$$\epsilon_i \sim Normal(y_i, \sigma^2) y_i \qquad = \alpha_{[i]} + \beta_{ClimateChange_{[i]}} + \beta_{Species} \beta_{ClimateChange \times Species_{[i]}} + \epsilon_{[i]}$$
 (2)

138 Main Model

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To best compare across the effects of each climatic and geographic variable, we scaled all of the predictors and used a z-score following the binary predictor approach (Gelman & Hill, 2006). To control for spatial 140 autocorrelation and to account for spatially structured processes independent from our regional predictors of false springs, we generate an additional 'space' parameter for the model. To generate our space parameter 142 we first extracted spatial eigenvectors corresponding to our analyses' units and selected the subset that 143 minimizes spatial autocorrelation of the residuals of a model including all predictors except for the space parameter (Diniz-Filho et al., 2012; Bauman et al., 2017) (see supplement 'Methods: Spatial parameter' for 145 more details). We then took the eigenvector subset determined from the minimization of Moran's the residuals (MIR approach) and regressed them against the above residuals—i.e. number of false springs vs. 147 climatic and geographical factors. Finally we used the fitted values of that regression as our space parameter, which, by definition, represents the portion of the variation in false springs that is both spatially structured 149 and independent from all other predictors in the model (e.g. average spring temperature, elevation, etc. 150 Griffith & Peres-Neto, 2006; Morales-Castilla et al., 201 151 To estimate the probability of false spring risk across species and our predictors we used a Bayesian modeling 152 approach. By including all parameters in the model, as well as species, we were able to distinguish the 153 strongest contributing factors to false spring risk. We fit a Bernoulli distribution model (also know as a logistic regression) using mean spring temperature (written as MST in the model equation), NAO, elevation, distance 155

from the coast, space, and climate change as predictors and all two-way interactions and species as two-way

interactions (Equation 1), using the brms package (Bürkner, 2017), version 2.3.1, in R (R Development Core

Team, 2017), version 3.3.1, and was written as follows:

$$y_{i} \sim Binomial(1, p) \tag{1}$$

$$logit(p) = \alpha_{[i]} + \beta_{MST_{[i]}} + \beta_{DistanceCoast_{[i]}} + \beta_{Elevation_{[i]}} + \beta_{NAO_{[i]}} + \beta_{Space_{[i]}} + \beta_{ClimateChange_{[i]}} + \beta_{Species_{[i]}}$$

$$+ \beta_{MST \times Species_{[i]}} + \beta_{DistanceCoast \times Species_{[i]}} + \beta_{Elevation \times Species_{[i]}} + \beta_{NAO \times Species_{[i]}}$$

$$+ \beta_{Space \times Species_{[i]}} + \beta_{ClimateChange \times Species_{[i]}} + \beta_{MST \times ClimateChange_{[i]}}$$

$$+ \beta_{DistanceCoast \times ClimateChange_{[i]}} + \beta_{Elevation \times ClimateChange_{[i]}}$$

$$+ \beta_{NAO \times ClimateChange_{[i]}} + \beta_{Space \times ClimateChange_{[i]}}$$

We ran four chains, each with 2,500 warm-up iterations and 4,000 sampling iterations for a total of 6,000 posterior samples for each predictor. We evaluated our model performance based on \hat{R} values that were close to one. We also evaluated effective sample size estimates, which were 1,994 or above. We additionally assessed chain convergence visually and posterior predictive checks. Due to the large number of observations in the data we used the FASRC Cannon cluster supported by the FAS Division of Science Research Computing Group at Harvard University to run the model.

Model estimates were on the logit scale (shown in all tables) and were converted to probability percentages in all figures for easier interpretation by using the 'divide by 4' rule (Gelman & Hill, 2006) and then back

in all figures for easier interpretation by using the 'divide by 4' rule (Gelman & Hill, 2006) and then back converted to the original scale by multiplying by two standard deviations. We calculated overall estimates (i.e., across species) of main effects in Figure 3, Figure S3 and Figure S4 from the average of the posteriors of each effect by species. We report all estimated values in-text as mean \pm 98% uncertainty intervals, unless otherwise noted. The combined effects of climate change with all of the climatic and geographic factors across species were determined by adding all effects in the model plus species for after climate change and subtracting this from the combined effects in the model for each species after climate change. This difference was reported as the combined change in false spring risk for each species.

174 Results

Basic shifts in budburst and number of false springs

Day of budburst varied across the six species and across geographical gradients (Figure 1). Betula pendula,

Aesculus hippocastanum, Alnus glutinosa (Figure 1A-C) generally initiated budburst earlier than Fagus sylvatica, Quercus robur, and Fraxinus excelsior (Figure 1D-F). Across all six species, higher latitude sites and sites closer to the coast tended to initiate budburst later in the season (Figure 1).

Across species, budburst dates advanced -5.22 \pm 0.15 days after 1983 (Table S3) and minimum temperatures between budburst and leafout increased by 0.62 \pm 0.3°C after climate change (Table S4). This trend in advancing day of budburst for each species corresponds closely with increasing mean spring temperatures (Figure S1). While all species initiated budburst approximately seven days earlier (Figure 2A, Table S2 and Table S3), the average minimum temperature between budburst and leafout varied across the six species with Betula pendula and Aesculus hippocastanum experiencing the lowest minimum temperatures (Figure 2B), Quercus robur and Fraxinus excelsior experiencing the highest minimum temperatures, and Fraxinus excelsior experiencing the greatest variation (Figure 2B).

A simplistic view of changes in false springs—one that does not consider changes in climatic and geographic factors or effects of spatial autocorrelation—suggests that the number of false springs increased across species by 0.03% (± 0.05%) after climate change (i.e., after 1983), but with important variation by species (Figure 2C). Early-leafout species (Aesculus hippocastanum, Alnus glutinosa and Betula pendula) showed an increased risk whereas later bursting species (Fagus sylvatica, Quercus robur and Fraxinus excelsior) showed a decrease in risk (Table S5).

The effects of climatic and geographic variation coupled with climate change on false spring risk

Climatic and geographic factors underlie variation across years and space in false springs (Figure 3 and Table S6) before recent climate change (1983). Mean spring temperature had the strongest effect on false springs, with warmer spring temperatures resulting is fewer false springs (Figure 3 and Table S6; comparable estimates come from using standardized variables, see Methods for more details). For every 2° C increase in mean spring temperature there was a -7.64% decrease in the probability of a false spring (-0.48 ± 0.03

probability of false spring/standard un Distance from the coast had the second biggest effect on false 201 spring incidence. Individuals at sites further from the coast tended to have earlier leafout dates, which corresponded to an increased risk in false springs (Figure 3 and Table S6). For every 150km away from the 203 coast there was a 5.32\% increase in risk in false springs (0.4 \pm 0.03 probability of false spring/standard unit). Sites at higher elevations also had higher risks of false spring incidence—likely due to more frequent colder 205 temperatures—with a 2.23% increase in risk for every 200m increase in elevation (0.19 \pm 0.04 probability 206 of false spring/standard unit, Figure 3 and Table S6). More positive NAO indices, which generally advance 207 leafout, slightly heightened the risk of false spring, with every 0.3 unit increase in NAO index there was a 1.91% increased risk in false spring or 0.14 ± 0.03 probability of false spring/standard unit (Figure 3 and 209 Table S6). 210

These effects varied across species (Figure 4). While there were fewer false springs for each species with increasing mean spring temperatures, *Betula pend* had the greatest risk of false springs and *Fraxinus excelsior* had the lowest risk (Figure 4A). There was an increased risk of false spring for all species at sites further from the coast (Figure 4B), with a sharp increase in risk for *Fraxinus excelsior* at sites further from the coast. With increasing elevation, all species had a greater risk of a false spring, except for *Fraxinus excelsior*, which had a slightly decreased risk at higher elevations (Figure 4C). With increasing NAO indices, the risk of false spring remained consistent for most species, except *Fagus sylvatica* experienced more with higher NAO indices (Figure 4D).

After climate change, the effects of these climatic and geographic factors on false spring risk shifted (Figure 219 3). Warmer sites still tended to have lower risks of false springs, but with climate change, increasing mean 220 spring temperatures had much less of an effect on false spring risk with -2.84% decrease in risk per 2°C 221 (or -0.06 ± 0.06 probability of false spring/standard unit versus -7.64% decrease per 2°C or -0.48 before climate change; Figure 3 and Figure S2A). There was a slightly reduced risk in false springs further from the coast after climate change (Figure 3 and Figure S2B) with 3.68% increase in risk per 150km (or 0.28 ± 0.07 224 probability of risk/standard unit versus 5.32% increase 150km or 0.4 ± 0.04 before climate change). The level of risk remained consistent before and after 1983 across elevations (Figure 3 and Figure S2C), with false 226 spring risk being higher at higher elevations. After climate change, the rate of false spring incidence largely 227 decreased with increasing NAO indices (Figure 3 and Figure S2D), now with a -9.15% decrease in risk per 0.3 228 unit increase in the NAO index (or -0.69 ± 0.06 probability of false spring/standard unit or versus 1.91% 0.3 unit increase in the NAO index or 0.14 ± 0.03 before climate change). After climate change, NAO had the 230

strongest effect on false spring risk, with higher NAO indices rendering fewer false springs. Together there
was a -14.55% decrease in risk of false springs (or a 0.58 decrease in probability or risk/standard unit) due
to the combined effects of all the climatic and geographic factors after 1983 across species.

In addition to these combined changes in climatic and geographic factors, which we estimated across all species, there is still an unexplained shift in false spring risk due to climate change for most species. This unexplained effect of climate change varied strongly by species, with an 8.86% increased risk in false springs after climate change for Aesculus hippocastanum (or 0.35 ± 0.03 probability of false spring/standard unit; Figure 3, Figure 4E and Table S6), a 10.54% increase for Alnus glutinosa, a 10.29% increase for Betula pendula, and a 0.75% for Fagus sylvatica (or a 0.4 ± 0.08 , 0.41 ± 0.08 and 0.032 ± 0.08 probability of false spring/standard unit respectively; Figure 3, Figure 4E and Table S6). Climate change decreased risk for Fraxinus excelsior by -4.27% and Quercus robur by -1.76% (or a -1.08 ± 0.1 and -0.67 ± 0.08 probability of false spring/standard unit respectively; Figure 3, Figure 4E and Table S6).

sidering the total effect of climate change on species—by combining the unexplained shifts in false spring 243 risk with climate change for each species with the effects of the climatic and geographic factors after climate change—yields an overall mean decrease in risk of false springs after climate change for all species, but effects 245 vary between early and later-leafout species. Earlier leafout species tended to see smaller declines in risk after climate change with a -5.77% decrease in risk for Aesculus hippocastanum (or -0.23 \pm 0.06 probability 247 of risk/standard unit), and a -4.27% decrease in risk for Alnus qlutinosa and Betula pendula (or -0.17 \pm 0.09 probability of risk/standard unit). Whereas the later leafout species had larger declines in risk, with a -13.8% 249 decrease in risk for Fagus sylvatica (or -0.55 ± 0.08 probability of risk/standard unit), -18.82% decrease in risk for Frazinus excelsior (or -0.75 ± 0.11 probability of risk/standard unit), and -16.06% decrease in risk 251 for Quercus robur (or -0.64 ± 0.09 probability of risk/standard unit). 252

Sensitivity of results to duration of risk and temperature thresholds

Our results remained consistent (in direction and magnitude) when we applied different rates of leafout for each species (i.e., varied the length of time between estimated budburst and leafout). Mean spring temperature (-8.08% decrease for every 2° C or -0.5 ± 0.04 probability of risk/standard unit) and distance from the coast (5.36% increase for every 150km or 0.4 ± 0.03 probability of risk/standard unit) were the strongest predictors for false spring risk (Figure S3 and Table S7). After climate change, there was a slight increase in false spring risk at higher elevations (Figure S3 and Table S7) compared to our main findings.

Results remained generally consistent also when we applied a lower temperature threshold for defining a false spring (i.e., -5°C), though there were more shifts in the magnitude of some effects, especially those of climate change. Mean spring temperature (-11.56% decrease for every 2° or -0.72 ± 0.07 probability of risk/standard unit) and elevation (7.35% increase in risk for every 200m or 0.63 ± 0.08 probability of risk/standard unit) were the strongest predictors, with a weaker effect of distance from the coast (2.75% for every 150km or 0.21 ± 0.08 probability of risk/standard unit; Figure S4 and Table S8). There was much higher risk of false springs by climate change unexplained by climatic and geographic factors included in the model (14.55% increase or 0.58 ± 0.07 probability of risk/standard unit; Figure S4 and Table S8) and this was consistent across all six species, averaging a 10.0% increase (0.4 probability of risk/standard unit).

Discussion

Integrating over 66 years of data and 11648 sites across Central Europe and major climatic and geographic 270 factors known to influence false spring risk, our results suggest climate change has decreased false spring risk for all six studied species. In line with previous work, our results support that higher elevations tend 272 to experience more false springs (Vitasse et al., 2018; Vitra et al., 2017) and sites that are generally warmer have lower risks of false springs (Wypych et al., 2016a). Individuals further from the coast typically initiated 274 leafout earlier in the season, which subsequently lead to an increase in risk and, similarly, years with higher NAO indices experienced a slight increase in risk. But many of these factors have been re-shaped by climate change, in particular the effect of climatic factors have shifted dramatically compare to shifts in geographical 277 factors. Across species, we find that NAO and mean spring temperature have shifted the most after 1983, while the effect of distance from the coast has only shifted slightly and the effect of elevation has not shifted 279 (Figure S2.) These shifts in the influence of climatic and geographic factors in turn result in different effects of climate change on species. Though there has been a consistent decrease in false spring risk for all species we studiedgiven the combined effects of all factors that contribute to false spring risk—some species (e.g. Fraxinus 283 excelsior and Quercus robur) have experienced total reases while others have experienced smaller shifts in risk (e.g., Aesculus hippocastanum, Alnus glutinosa and Betula pendula.) These species-specific effects integrate over shifts in the influence of climatic and geographic factors on false spring risk, as well as residual variation not explained by these factors, suggesting for which species we have a robust understanding of what

- drivers underlie shifts in false spring risk with climate change, versus those species where more understanding
- is most critically needed.

²⁰⁰ Climatic and geographic effects on false spring risk

Past studies using single parameters for false spring events (Liu et al., 2018; Vitasse et al., 2018; Ma et al., 291 2018; Wypych et al., 2016b) have led to contradicting predictions in future false spring risk. Some studies are finding an increased risk at higher elevations after climate change (Vitasse et al., 2018), others are finding 293 an increase in risk only in Europe but not in other regions (Liu et al., 2018), while still others found a decrease in false spring risk across Central Europe (Wypych et al., 2016b). Research to date has also found 295 variation in false spring risk after climate change across species (Ma et al., 2018). By integrating both climate gradients and geographical factors, we were able to disentangle the major predictors of false spring risk and 297 merge these with species differences to determine which factors have the strongest effects on false spring risk. Mean spring temperature, distance from the coast and climate change were the strongest predictors 299 for false spring risk, however, NAO and elevation also affected risk, emphasizing the need to incorporate 300 multiple predictors. Further, climatic and geographic factors varied in how consistent, or not, they were 301 across species. Mean spring temperature, distance from the coast and NAO effects were fairly consistent 302 across species in direction, though Frazinus excelsior experienced a much greater increase in risk at sites further from the coast and Fagus sylvatica had a heightened risk to higher NAO indices compared to the 304 other species. Elevation was the only factor that varied in direction among the species with most species having an increased risk at higher elevations except for Fraxinus excelsior, which had a decreased risk. These 306 inconsistencies may capture range differences among species, with contrasting effects of factors on individuals 307 closer to range edges (Chuine & Beaubien, 2001). Since the onset of recent major climate change, the strength of these climatic and geographic effects have changed, highlighting the need to better understand and model shifting drivers of false spring. After climate 310 change, our results show a large decrease in risk of false spring with higher NAO indices. This could be because high NAO conditions no longer lead to temperatures low enough to trigger a false spring—that is, with climate-312 change induced warming high NAO conditions may no longer produce the freezing temperatures needed for 313 false springs (Screen, 2017). Conversely, we found an increased risk with warmer mean spring temperatures after climate change, which may be driven by our studied plant species responding very strongly to increased 315 spring warming with climate change (i.e., large advances in spring phenology, Figure S1), resulting in an

increased risk of exposure to false springs at these locations. Improved mechanistic models of how warming temperatures affect budburst (Gauzere *et al.*, 2017; Chuine *et al.*, 2016) could improve our understanding of how NAO and mean spring temperatures contribute to false spring risk.

Variation in risk across species

By integrating climatic and regional factors—e.g., elevation, distance from the coast—we can unravel pheno-321 logical effects on the probability risk from the climatic and geographic factors that contribute to an individual's 322 level of false spring risk. Due to the prominent shifts in the geographic and climatic factors (i.e., mean spring temperature, elevation, distance from the coast and NAO indices) with climate-change induced warming, we 324 estimated a combined decrease in risk of false springs across all species after 1983, though residual (unexplained by climatic and geographic) effects of climate change resulted in marked differences in risk between 326 early- and late-leafout species. Before 1983, false spring risk was slightly higher for species initiating leafout 327 earlier in the spring but overall the risk was more consistent across species (Figure 4E). After climate change, 328 however, species differences in risk amplified: the early-leafout species (i.e., Aesculus hippocastanum, Alnus 329 glutinosa and Betula pendula) had an increased risk, the middle-leafout species—i.e. Fagus sylvatica—had a 330 similar level of risk as before and the later-leafout species (i.e., Frazinus excelsior and Quercus robur) had a 331 decreased risk (Figure 4E). Our combined estimates suggest all species have experienced a decline in false spring risk, in contrast with 333 simple estimates of absolute changes in number of false springs across species . These simple estimates, 334 which suggested an increase in risk for early-leafout species and a decline or no change for later-leafout 335

simple estimates of absolute changes in number of false springs across species . These simple estimates,
which suggested an increase in risk for early-leafout species and a decline or no change for later-leafout
species, correlated more closely with estimated effects of climate change on species unexplained by climatic
or geographic factors. The three early species (Betula pendula, Aesculus hippocastanum, Alnus glutinosa) had
much higher unexplained shift sees spring with climate change than the later species (Fagus sylvatica,
Quercus robur, Fraxinus excelsior), suggesting the climatic and geographic factors we examined are better
capturing variation in false spring risk for later species—and that we still fundamentally lack information on
what drives false spring risk for the early-species, which are also fundamentally the species with highest risk
While our model examines the major factors expected to influence false spring risk (Liu et al., 2018; Ma et al.,
2018; Vitasse et al., 2018; Wypych et al., 2016b), these results highlight the need to explore other climatic
factors to improve forecasting. We expect factors that affect budburst timing, such as shifts in over-winter

chilling temperature, may help explain these discrepancies, but progress will require improved models of

chilling beyond the current models, which have been mainly developed for crops (Luedeling & Brown, 2011;
Dennis, 2003).

Habitat preference and range differences among the species could also explain some of the species-specific 348 variation in the results, but would require data on more species—and species that vary strongly in their climatic and geographic ranges—for robust analyses. The overall ranges of the predictors are similar across 350 species, but Betula pendula extends to the highest elevation and latitude and spans the greatest range of 351 distances from the coast, while Quercus robur experiences the greatest range of mean spring temperatures. 352 Within our species, Betula pendula has the largest global distribution, extending the furthest north and east into Asia. The distribution of Frazinus excelsior extends the furthest south (into the northern region of 354 Iran). These range differences could potentially underlie the unexplained effect of climate change seen in our results and why the shifts in climatic and geographic factors did not explain much of the variation in false 356 spring risk across species. Fagus sylvatic as better explained by the model and this species has a smaller range, more confined to Central Europe. Future research that captures these spatial, temporal and climatic 358 differences across myriad of species could greatly enhance predictions and help us understand these residual effects of climate change.

Forecasting false springs

Our study shows how robust forecasting must integrate across major climatic and geographic factors that underlie false spring, and allow for variation in these factors across species and over time as warming continues.

Of the four climatic and geographic factors we examined, only the effect of elevation remained constant before and after climate change. This is perhaps not surprising as climate change is shifting critical spring temperatures and reshaping the temporal and spatial dynamics of how climate effect udburst, leafout and freezing temperatures, yet it highlights the complexity that robust forecasting will require. Further, the differences we found across species suggest we can forecast some species better than others—such as Fagus sylvatica—which experienced almost zero unexplained climate change effects, thus, risk was likely determined by the climatic and geographic factors already included in the model.

Moving forward, more data on more species will be critical for estimates at community or ecosystem scales (at least in species-rich ecosystems). Related to this, more research on the effects of climate change on both budburst and leafout, the timing when individuals are most at risk to spring freeze damage (Chamberlain et al., 2019; Lenz et al., 2016) and on what temperatures cause leaf damage will help better understand

differences across species. Though we found that differing rates of leafout across species had minimal effects 375 on predicting risk, we did find that the lower temperature threshold can have an impact on model estimates (and thus forecasts), with lower temperature thresholds (i.e., -5°C versus -2.2°C) predicting increased risk 377 across all six study species. Our study uses an index of false spring risk, to estimate when damage may have occurred; it does not assess the intensity or severity of the false spring events observed, nor does it 379 record the amount of damage to individuals. Other research has shown that this temperature threshold may vary importantly by species (Körner et al., 2016; Lenz et al., 2013; Zhuo et al., 2018; Bennett et al., 2018). 381 Some species or individuals may be less freeze tolerant (i.e., are damaged from higher temperatures than -2.2°C), whereas other species or individuals may be able to tolerate temperatures as low as -8.5°C (Lenz 383 et al., 2016). Further, cold tolerance can be highly influenced by fall and winter climatic dynamics that 384 influence tissue hardiness (Hofmann & Bruelheide, 2015; Vitasse et al., 2014; Charrier et al., 2011) and can also influence budburst timing (Morin et al., 2007). Thus, we expect these effects are likely integrated and that 386 useful forecasting will require far better species-specific models of budburst, leafout and hardiness—including whether budburst and hardiness may be inter-related. 388

Our results highlight how climate change complicates forecasting through multiple levels. It has shifted the influence of climatic and geographic factors, fundamentally reshaping relationships with major climatic factors such that relationships before climate change no longer hold. It has also magnified species-level variation in false spring risk. Layered onto this complexity is unexplained variation that suggests we are missing key factors that drive interspecific variation in false spring risk. Our study focuses on one region (i.e., Central Europe) with high-quality and abundant data and we hope that our results can be applied to other systems as more and more data becomes available. Integrating these findings into future models will provide more robust forecasts and help us begin to unravel the complexities of climate change effects across species.

False spring events increased w climate change, though it was more pronounced in species that initiated budburst earlier in the season. Thus we need a better understanding of the major drivers of false spring risk, how these events are changing in duration and intensity and if there are shifts in the level of damage to individuals. Our integrated approach may help direct future modelling advancements in false spring research. We show here the importance of using multiple geographic and climatic factors in predicting false spring risk and how that risk varies across species. By using phenology data to provide a better estimate for budburst and leafout, predictions for false springs will be more accurate for inter-specific risk. Additionally, we demonstrate that incorporating all regional effects is more important than simply assessing budburst

timing across species. Individuals that initiate budburst earlier in the season are not necessarily exposed to more false springs, thus, investigating site effects is essential talse spring risk in addition to day of budburst. Our results suggest there is a heightened risk of false springs with climate change for some species and that there will be complex responses to warming in the future, which could in turn, have escalating impacts on plant community dynamics and further augment climatic shifts.

10 References

- Augspurger CK (2009) Spring 2007 warmth and frost: phenology, damage and refoliation in a temperate deciduous forest. Functional Ecology, 23, 1031–1039. doi:10.1111/j.1365-2435.2009.01587.x.
- Augspurger CK (2013) Reconstructing patterns of temperature, phenology, and frost damage over 124 years:

 Spring damage risk is increasing. *Ecology*, **94**, 41–50. doi:10.1890/12-0200.1.
- Basler D, Körner C (2012) Photoperiod sensitivity of bud burst in 14 temperate forest tree species. Agricultural
 and Forest Meteorology, **165**, 73–81. doi:10.1016/j.agrformet.2012.06.001.
- Bauman D, Drouet T, Dray S, Vleminckx J (2017) Disentangling good from bad practices in the selection of
 spatial or phylogenetic eigenvectors. *Ecography*, **0**. doi:10.1111/ecog.03380.
- Bennett JM, Calosi P, Clusella-Trullas S, et al. (2018) Globtherm, a global database on thermal tolerances for aquatic and terrestrial organisms. Scientific data, 5, 180022.
- Buerki S, Lowry II P, Alvarez N, Razafimandimbison S, Kupfer P, Callmander M (2010) Phylogeny and circumscription of *Sapindaceae* revisited: Molecular sequence data, morphology and biogeography support recognition of a new family, *Xanthoceraceae*. *Plant Ecology and Evolution*, **143**, 148–159. doi:10.5091/plecevo.2010.437.
- Bürkner PC (2017) brms: An R Package for Bayesia Multilevel Models. Journal of Statistical Software, 80,

 1–28.
- ⁴²⁷ Chamberlain CJ, Cook BI, de Cortazar Atauri IG, Wolkovich EM (2019) Rethinking false spring risk. *Global*⁴²⁸ Change Biology, **25**, 2209–2220. doi:10.1111/gcb.14642.
- Charrier G, Bonhomme M, Lacointe A, Améglio T (2011) Are budburst dates, dormancy and cold acclimation in walnut trees (juglans regia l.) under mainly genotypic or environmental control? *International Journal of*

- Biometeorology, 55, 763-774. doi:10.1007/s00484-011-0470-1. URL https://doi.org/10.1007/s00484-
- o11-0470-1.
- ⁴³³ Chen X, Hu B, Yu R (2005) Spatial and temporal variation of phenological growing season and climate
- change impacts in temperate eastern China. Global Change Biology, 11, 1118–1130. doi:10.1111/j.1365-
- 2486.2005.00974.x.
- Chmielewski FM, Rötzer T (2001) Response of tree phenology to climate change across Europe. Agricultural
 and Forest Meteorology, 108, 101 112. doi:https://doi.org/10.1016/S0168-1923(01)00233-7.
- ⁴³⁸ Chuine I, Beaubien EG (2001) Phenology is a major determinant of tree species range. Ecology Letters, 4,
- 500-510. doi:10.1046/j.1461-0248.2001.00261.x. URL https://onlinelibrary.wiley.com/doi/abs/10.
- 440 1046/j.1461-0248.2001.00261.x.
- Chuine I, Bonhomme M, Legave JM, García de Cortázar-Atauri I, Charrier G, Lacointe A, Améglio T
- (2016) Can phenological models predict tree phenology accurately in the future? the unrevealed hurdle
- of endodormancy break. Global Change Biology, 22, 3444-3460. doi:10.1111/gcb.13383. URL http:
- //dx.doi.org/10.1111/gcb.13383.
- Dennis F (2003) Problems in standardizing methods for evaluating the chilling requirements for the breaking
 of dormancy in buds of woody plants. *HortScience*, **38**, 347–350.
- Diniz-Filho JAF, Bini LM, Rangel TF, Morales-Castilla I, Olalla-Tárraga MÁ, Rodríguez MÁ, Hawkins BA

 (2012) On the selection of phylogenetic eigenvectors for ecological analyses. *Ecography*, **35**, 239–249.
- Donnelly A, Yu R, Caffarra A, et al. (2017) Interspecific and interannual variation in the duration of spring

 phenophases in a northern mixed forest. Agricultural and Forest Meteorology, 243, 55–67.
- Flynn DFB, Wolkovich EM (2018) Temperature and photoperiod drive spring phenology across all species in a temperate forest community. *New Phytologist.* doi:10.1111/nph.15232.
- 453 Gauzere J, Delzon S, Davi H, Bonhomme M, Garcia de Cortazar-Atauri I, Chuine I (2017) Integrating
- interactive effects of chilling and photoperiod in phenological process-based models. A case study with two
- European tree species: Fagus sylvatica and Quercus petraea. Agricultural and Forest Meteorology, pp. 9–20.
- Gelman A, Hill J (2006) Data analysis using regression and multilevel/hierarchical models. Cambridge university press.

- Griffith DA, Peres-Neto PR (2006) Spatial modeling in ecology: the flexibility of eigenfunction spatial analyses. *Ecology*, **87**, 2603–2613.
- Gu L, Hanson PJ, Post WM, et al. (2008) The 2007 Eastern US spring freeze: Increased cold damage in a
 warming world. BioScience, 58, 253. doi:10.1641/B580311.
- Hänninen H (1991) Does climatic warming increase the risk of frost damage in northern trees? Plant, Cell

 Environment, 14, 449–454. doi:10.1111/j.1365-3040.1991.tb01514.x.
- Hipp A, S Manos P, González-Rodríguez A, et al. (2017) Sympatric parallel diversification of major oak clades
 in the Americas and the origins of Mexican species diversity. New Phytologist, 217. doi:10.1111/nph.14773.
- Hofmann M, Bruelheide H (2015) Frost hardiness of tree species is independent of phenology and macrocli matic niche. Journal of Biosciences, 40, 147–157. doi:10.1007/s12038-015-9505-9.
- Hufkens K, Friedl MA, Keenan TF, Sonnentag O, Bailey A, O'Keefe J, Richardson AD (2012) Ecological
 impacts of a widespread frost event following early spring leaf-out. Global Change Biology, 18, 2365–2377.
 doi:10.1111/j.1365-2486.2012.02712.x.
- Inouye DW (2008) Effects of climate change on phenology, frost damage, and floral abundance of montane wildflowers. *Ecology*, **89**, 353–362.
- 473 IPCC (2015) Climate change 2014: mitigation of climate change, vol. 3. Cambridge University Press.
- Kharouba HM, Ehrlén J, Gelman A, Bolmgren K, Allen JM, Travers SE, Wolkovich EM (2018) Global shifts in the phenological synchrony of species interactions over recent decades. *Proceedings of the National* Academy of Sciences, 115, 5211–5216. doi:10.1073/pnas.1714511115.
- Klosterman S, Hufkens K, Richardson AD (2018) Later springs green-up faster: the relation between onset and completion of green-up in deciduous forests of North America. *International Journal of Biometeorology*. doi:10.1007/s00484-018-1564-9.
- KNMI (2018) Daily CPC NAO data. URL https://climexp.knmi.nl/getindices.cgi?WMO=NCEPData/
 cpc_nao_daily&STATION=NAO&TYPE=i&id=someone@somewhere&NPERYEAR=366.
- Kollas C, Körner C, Randin CF (2014) Spring frost and growing season length co-control the cold range limits of broad-leaved trees. *Journal of Biogeography*, **41**, 773–783. doi:10.1111/jbi.12238.

- 484 Körner C, Basler D, Hoch G, et al. (2016) Where, why and how? Explaining the low-temperature range
- limits of temperate tree species. Journal of Ecology, 104, 1076–1088. doi:10.1111/1365-2745.12574. URL
- http://dx.doi.org/10.1111/1365-2745.12574.
- 487 Kramer K (1994) A modelling analysis of the effects of climatic warming on the probability of spring frost
- damage to tree species in The Netherlands and Germany. Plant, Cell & Environment, 17, 367–377. doi:
- 489 10.1111/j.1365-3040.1994.tb00305.x.
- 490 Kukal MS, Irmak S (2018) U.S. Agro-Climate in 20th century: growing degree days, first and last frost,
- growing season length, and impacts on crop yields. Scientific Reports, 8. doi:10.1038/s41598-018-25212-2.
- Labe Z, Ault T, Zurita-Milla R (2016) Identifying anomalously early spring onsets in the CESM large ensemble
- project. Climate Dynamics, 48, 3949–3966. doi:10.1007/s00382-016-3313-2.
- Lenz A, Hoch G, Körner C, Vitasse Y (2016) Convergence of leaf-out towards minimum risk of freezing
- damage in temperate trees. Functional Ecology, **30**, 1–11. doi:10.1111/1365-2435.12623.
- Lenz A, Hoch G, Vitasse Y, Körner C (2013) European deciduous trees exhibit similar safety margins against
- damage by spring freeze events along elevational gradients. New Phytologist, 200, 1166–1175. doi:10.1111/
- nph.12452.
- ⁴⁹⁹ Liu Q, Piao S, Janssens IA, et al. (2018) Extension of the growing season increases vegetation exposure to
- frost. Nature Communications, 9. doi:10.1038/s41467-017-02690-y.
- 501 Liu X, Yin Z, Shao X, Qin N (2006) Temporal trends and variability of daily maximum and minimum,
- extreme temperature events, and growing season length over the eastern and central Tibetan Plateau
- during 1961–2003. Journal of Geophysical Research: Atmospheres, 111. doi:10.1029/2005JD006915.
- 504 Luedeling E, Brown PH (2011) A global analysis of the comparability of winter chill models for fruit and nut
- trees. International Journal of Biometeorology, **55**, 411–421.
- Ma Q, Huang JG, Hänninen H, Berninger F (2018) Divergent trends in the risk of spring frost damage to
- trees in europe with recent warming. Global Change Biology, 0. doi:10.1111/gcb.14479.
- Martin M, Gavazov K, Körner C, Hattenschwiler S, Rixen C (2010) Reduced early growing season freezing
- resistance in alpine treeline plants under elevated atmospheric CO₂. Global Change Biology, 16, 1057–1070.
- doi:10.1111/j.1365-2486.2009.01987.x.

- ₅₁₁ Meier U (2001) Growth stages of mono-and dicotyledonous plants BBCH Monograph Edited by Uwe Meier
- Federal Biological Research Centre for Agriculture and Forestry. Agriculture, 12, 141—147 ST Geo-
- chemical study of the organic mat. doi:10.5073/bbch0515.
- Menzel A, Helm R, Zang C (2015) Patterns of late spring frost leaf damage and recovery in a European beech
- (Fagus sylvatica L.) stand in south-eastern Germany based on repeated digital photographs. Frontiers in
- ⁵¹⁶ Plant Science, **6**, 110. doi:10.3389/fpls.2015.00110.
- Morales-Castilla I, Olalla-Tarraga MA, Purvis A, Hawkins BA, Rodriguez MA (2012) The imprint of cenozoic
- migrations and evolutionary history on the biogeographic gradient of body size in new world mammals.
- The American Naturalist, 180, 246–256.
- Morin X, Améglio T, Ahas R, et al. (2007) Variation in cold hardiness and carbohydrate concentration from
- dormancy induction to bud burst among provenances of three European oak species. Tree Physiology, 27,
- 817-825. doi:10.1093/treephys/27.6.817. URL https://doi.org/10.1093/treephys/27.6.817.
- R Development Core Team (2017) R: A language and environment for statistical computing. R Foundation
- for Statistical Computing, Vienna, Austria.
- Richardson AD, Keenan TF, Migliavacca M, Ryu Y, Sonnentag O, Toomey M (2013) Climate change, phe-
- nology, and phenological control of vegetation feedbacks to the climate system. Agricultural and Forest
- 527 Meteorology, **169**, 156 173. doi:https://doi.org/10.1016/j.agrformet.2012.09.012.
- Sakai A, Larcher W (1987) Frost Survival of Plants. Springer-Verlag.
- 529 Scheifinger H, Menzel A, Koch E, Peter C (2003) Trends of spring time frost events and phenological dates
- in Central Europe. Theoretical and Applied Climatology, 74, 41-51. doi:10.1007/s00704-002-0704-6.
- 531 Schwartz MD (1993) Assessing the onset of spring: A climatological perspective. Physical Geography, 14(6),
- 536-550.
- Screen JA (2017) The missing northern european winter cooling response to arctic sea ice loss. Nature
- 534 Communications, 8, 14603. doi:10.1038/ncomms14603. URL https://doi.org/10.1038/ncomms14603.
- 555 Sgubin G, Swingedouw D, Dayon G, de Cortázar-Atauri IG, Ollat N, Pagé C, van Leeuwen C (2018) The risk
- of tardive frost damage in French vineyards in a changing climate. Agricultural and Forest Meteorology,
- 250-251, 226 242. doi:https://doi.org/10.1016/j.agrformet.2017.12.253.

- 558 Stocker TF, Qin D, Plattner GK, et al. (2013) Climate Change 2013: The Physical Science Basis. Contribution
- of Working Group I to the Fifth Assessment Report of the Intergovernmental Panel on Climate Change,
- 540 1535 pp.
- Templ B, Koch E, Bolmgren K, et al. (2018) Pan European Phenological database (PEP725): a single point
- of access for European data. International Journal of Biometeorology, 62, 1109–1113. doi:10.1007/s00484-
- 018-1512-8. URL https://doi.org/10.1007/s00484-018-1512-8.
- USA-NPN (2019) Plant and animal phenology data. USA National Phenology Network. doi:10.5066/
- F78S4N1V. URL http://doi.org/10.5066/F78S4N1V.
- ⁵⁴⁶ Vitasse Y, Lenz A, Körner C (2014) The interaction between freezing tolerance and phenology in temperate
- deciduous trees. Frontiers in Plant Science, 5. doi:10.3389/fpls.2014.00541.
- Vitasse Y, Schneider L, Rixen C, Christen D, Rebetez M (2018) Increase in the risk of exposure of forest and
- fruit trees to spring frosts at higher elevations in Switzerland over the last four decades. Agricultural and
- Forest Meteorology, **248**, 60 69. doi:https://doi.org/10.1016/j.agrformet.2017.09.005.
- Vitra A, Lenz A, Vitasse Y (2017) Frost hardening and dehardening potential in temperate trees from winter
- to budburst. New Phytologist, **216**, 113–123. doi:10.1111/nph.14698.
- Wang N, McAllister HA, Bartlett PR, Buggs RJA (2016) Molecular phylogeny and genome size evolution of
- the genus Betula (Betulaceae). Annals of Botany, 117, 1023–1035. doi:10.1093/aob/mcw048.
- Wolkovich EM, Cook BI, Allen JM, et al. (2012) Warming experiments underpredict plant phenological
- responses to climate change. Nature, 485, 18–21. doi:10.1038/nature11014.
- 557 Wypych A, Sulikowska A, Ustrnul Z, Czekierda D (2016a) Variability of growing degree days in Poland
- in response to ongoing climate changes in Europe. International Journal of Biometeorology, 61, 49-59.
- doi:10.1007/s00484-016-1190-3. URL http://dx.doi.org/10.1007/s00484-016-1190-3.
- 560 Wypych A, Ustrnul Z, Sulikowska A, Chmielewski FM, Bochenek B (2016b) Spatial and temporal variability of
- the frost-free season in Central Europe and its circulation background. *International Journal of Climatology*,
- **37**, 3340–3352. doi:10.1002/joc.4920.
- ⁵⁶³ Zhuo X, Zheng T, Zhang Z, et al. (2018) Genome-wide analysis of the NAC transcription factor gene family
- reveals differential expression patterns and cold-stress responses in the woody plant Prunus mume. Genes.
- 9. doi:10.3390/genes9100494.

Zohner CM, Benito BM, Svenning JC, Renner SS (2016) Day length unlikely to constrain climate-driven shifts in leaf-out times of northern woody plants. *Nature Climate Change*, **6**, 1120–1123. doi:10.1038/

Tables and Figures

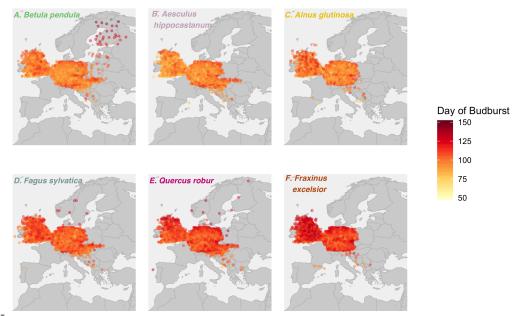


Figure 1: The average day of budburst mapped by site for each species (ordered by day of budburst starting with *Betula pendula* as the earliest budburst date to *Fraxinus excelsior*). Species names are color-coded to match figures throughout the text.

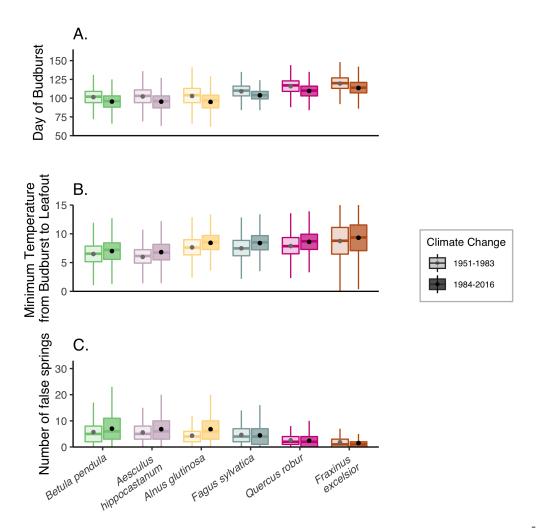


Figure 2: Day of budburst (A.), minimum temperatures between budburst and leafout (B.) and number of false springs (C.) before and after 1983 across species for all sites. Box and whisker plots show the 25th and 75th percentiles (i.e., the interquartile range) with notches indicating 95% uncertainty intervals. Dots and error bars overlaid on the box and whisker plots represent the model regression outputs (Tables S3-S5). Error bars from the model regressions indicate 98% uncertainty intervals but, given the number of sites, are quite small and thus not easily visible (see Tables S3-S5). Species are ordered by day of budburst and are color-coded to match the other figures.

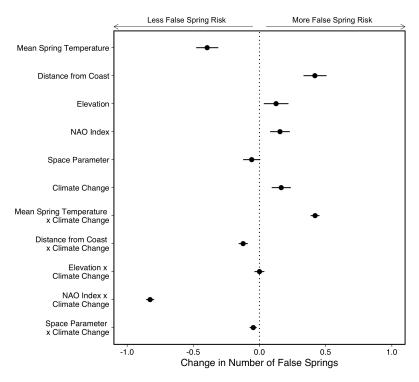


Figure 3: Effects of species, climatic and geographical predictors on false spring risk. More positive values indicate an increased probability of a false spring whereas more negative values suggest a lower probability of a false spring. Dots and lines show means and 98% uncertainty intervals. Values closer to zero have less of an effect on false springs. There were 582,211 zeros and 172,877 ones for false springs in the data. See Table S6 for full model output.

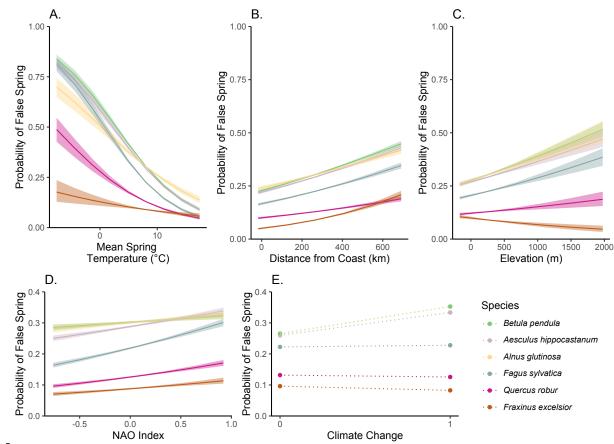


Figure 4: Species-level variation across geographic and spatial predictors (i.e., mean spring temperature (A.), distance from the coast (B.), elevation (C.), and NAO index (D.)). Lines and shading are the mean and 98% uncertainty intervals for each species. To reflect the raw data, we converted the model output back to the original scale for the x-axis in each panel. See Table S6 for full model output.