Climate change reshapes the drivers of false spring risk across European trees

- 3 Authors:
- ⁴ C. J. Chamberlain ^{1,2}, B. I. Cook ³, I. Morales-Castilla ^{4,5} & E. M. Wolkovich ^{1,2,6}
- 5 Author affiliations:
- ⁶ Arnold Arboretum of Harvard University, 1300 Centre Street, Boston, Massachusetts, USA;
- ²Organismic & Evolutionary Biology, Harvard University, 26 Oxford Street, Cambridge, Massachusetts, USA;
- ⁸ NASA Goddard Institute for Space Studies, New York, New York, USA;
- ⁹ GloCEE Global Change Ecology and Evolution Group, Department of Life Sciences, Universidad de Al-
- calá, Alcalá de Henares, 28805, Spain
- ⁵Department of Environmental Science and Policy, George Mason University, Fairfax, VA 22030;
- ¹² ⁶Forest & Conservation Sciences, Faculty of Forestry, University of British Columbia, 2424 Main Mall, Van-
- 13 couver, BC V6T 1Z4
- *Corresponding author: 248.953.0189; cchamberlain@g.harvard.edu

Abstract

Temperate and boreal forests are shaped by late spring freezing events after budburst—false springs—which
may shift with climate change. Research to date has generated conflicting results, potentially because no
study has compared the myriad climatic and geographic factors that contribute to a plant's risk of a false
spring. Here, we assessed the effects of mean spring temperature, distance from the coast, elevation and
the North Atlantic Oscillation (NAO) using PEP725 leafout data for six tree species across 11,648 sites in
Europe, to determine which were the strongest predictors of false spring risk and how these predictors shifted
with climate change. Across species before recent warming, mean spring temperature and distance from
the coast were the strongest predictors, with higher mean spring temperatures associated with decreased
risk in false springs (-7.64% per 2°C increase) and sites further from the coast experiencing an increased
risk (5.32% per 150km from the coast). Elevation (2.23% per 200m increase in elevation) and NAO index
(1.91% per 0.3 increase) also increased false spring risk. With recent warming, geographic effects (elevation

and distance from coast) remain relatively stable through time, while climatic factors have shifted in both magnitude, for mean spring temperature (down to -2.84% decrease in risk per 2°C), and direction, with positive NAO phases leading to lower risk (-9.15% decrease per 0.3). These shifts have magnified the residual effects of climate change resulting in an increased risk of false spring among early-leafout species (i.e., Aesculus hippocastanum, Alnus glutinosa and Betula pendula) versus a decline or no change in risk among late-leafout species (i.e., Fagus sylvatica, Fraxinus excelsior and Quercus robur). Our results show that climate change has reshaped the major drivers of false spring risk and highlight how considering multiple factors can yield a better understanding of the complexities of climate change.

36 Keywords: false spring, climate change, phenology, spring freeze, elevation, risk, leafout, temperate tree

Introduction

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False springs—late spring freezing events after budburst that can cause damage to temperate tree and shrub species—may shift with climate change. With earlier springs due to warming (IPCC, 2015; Wolkovich et al... 2012), the growing season is lengthening across many regions in the northern hemisphere (Chen et al., 2005; Kukal & Irmak, 2018; Liu et al., 2006). Longer growing seasons could translate to increased plant growth, assuming such increases are not offset by tissue losses due to false springs. Last spring freeze dates are not predicted to advance at the same rate as warming (Inouye, 2008; Labe et al., 2016; Martin et al., 2010; Sgubin et al., 2018; Wypych et al., 2016b), potentially amplifying the effects of false spring events in some regions. In Germany, for example, the last freeze date has advanced by 2.6 days per decade since 1955 (Zohner et al., 2016), but budburst has advanced roughly twice as fast. Major false spring events have been recorded in 47 recent years but studies have variously found that spring freeze damage may increase (Augspurger, 2013; Hänninen, 1991; Labe et al., 2016), remain the same (Scheifinger et al., 2003) or even decrease (Kramer, 1994; Vitra et al., 2017) with climate change. When damage does occur, studies have found it can take 16-38 days for trees to refoliate after a freeze (Augspurger, 2009, 2013; Gu et al., 2008; Menzel et al., 2015), which 51 can detrimentally affect crucial processes such as carbon uptake and nutrient cycling (Hufkens et al., 2012; Klosterman et al., 2018; Richardson et al., 2013).

Spring freezes are one of the largest limiting factors to species ranges and have greatly shaped plant life

times throughout the year, however, individuals are most at risk to damage in the spring, when freeze tolerance is lowest (Sakai & Larcher, 1987). Plants have adapted to these early spring risks through various mechanisms with one common strategy being avoidance (Vitasse et al., 2014). Many temperate species minimize freeze risk and optimize growth by using a complex mix of cues to initiate budburst: low winter temperatures (i.e., chilling), warm spring temperatures (i.e., forcing), and increasing spring daylengths (i.e., photoperiod). With climate change advancing, the interaction of these cues may shift spring phenologies both across and within species and sites, making some species less—or more—vulnerable to false springs than before. Species that leafout first each spring are especially at risk of false springs, as their budburst occurs during times of year when the risk of freeze events is relatively high. To date these species also appear to advance the most with warming (Wolkovich et al., 2012). Thus, if climate change increases the prevalence of late spring freezes, we would expect these species to see major increases in false spring risk. If climate change has restructured the timing and prevalence of false springs to later in the spring, then later-leafout species may experience major increases in false spring risk with climate change.

Some research suggests false spring incidence has already begun to decline in many regions (i.e. across parts of North America and Asia), however the prevalence of false springs has consistently increased across Europe since 1982 (Liu et al., 2018). Furthermore, recent studies have demonstrated site-specific effects may be more closely related to false spring risk: whether via elevation, where higher elevations appear at higher risk (Ma et al., 2018; Vitasse et al., 2018; Vitra et al., 2017), or distance from the coast, where inland areas appear at higher risk (Ma et al., 2018; Wypych et al., 2016b). Improved understanding of which geographical climatic factors impact false spring risk, including which factors are most crucial for predicting risk, we may be able to determine which regions are at risk currently and which regions will be more at risk in the future.

The majority of false spring studies assess the effects of one predictor (e.g. temperature, elevation or distance from the coast) on false spring prevalence, thus failing to compare how multiple factors may together shape risk. False spring risk is influenced by multiple climatic and geographic factors, which may vary across species and time. Further, because predictors can co-vary—for example, higher elevation sites are often more distant from the coast—the best estimates of what drives false springs should come from examining all predictors at once.

Here we investigate the influence of known climatic and geographic factors on false spring risk (defined here
as when temperatures fell below -2.2° between estimated budburst and leafout Schwartz, 1993). We assessed
the number of false springs that occurred across 11,648 sites across Europe using observed phenological data

(754,786 observations) for six temperate, deciduous trees and combined that with daily gridded climate data for each site that extended from 1951-2016. We focus on the major factors shown or hypothesized to influence false spring risk: mean spring temperature, elevation, distance from the coast, and a major climatic oscillation that structures European climate—the North Atlantic Oscillation (NAO). The NAO is tied to winter and spring circulation across Europe, with more positive NAO phases tending to result in higher than average winter and spring temperatures. With climate-change induced shifts, years with higher NAO indices have correlated to even earlier budburst dates since the late 1980s in some regions (Chmielewski & Rötzer, 2001), however little research has tested if more positive NAO phases also translate into more false springs. We aimed to understand (1) which climatic and geographic factors are the strongest predictors of false spring risk, and (2) how these major predictors have shifted with climate change across species.

96 Methods

97 Phenological Data and Calculating Vegetative Risk

We obtained phenological data from the Pan European Phenology network (PEP725, www.pep725.edu), which provides open access phenology records across Europe (Templ et al., 2018). Since plants are most susceptible to damage from freezing temperatures between budburst and full leafout, we selected leafout data (i.e., in Meier, 2001, BBCH 11, which is defined as the point of leaf unfolding and the first visible leaf stalk) 101 from the PEP725 dataset. The species used in the study were Aesculus hippocastanum Poir., Alnus glutinosa 102 (L.) Gaertn., Betula pendula Roth., Fagus sylvatica Ehrh., Fraxinus excelsior L., and Quercus robur L. Given 103 our focus on understanding how climatic and geographic factors underlie false spring risk, we selected species well-represented across space and time and not expected to be altered dominantly by human influence (i.e., 105 as crops and ornamental species often are), thus our selection criteria were as follows: (1) to be temperate, 106 deciduous species that were not cultivars or used as crops, (2) there were at least 90,000 observations of BBCH 11 (leafout), (3) to represent over half of the total number of sites available (11,684), and (4) there 108 were observations for at least 65 out of the 66 years of the study (1951-2016) (Table S1). Plants are generally the most freeze tolerant in the winter but this freeze tolerance greatly diminishes once 110 individuals exit the dormancy phase (i.e. processes leading to budburst) through full leaf expansion (Lenz 111 et al., 2016; Vitasse et al., 2014). Thus, for most individuals that initiate budburst and have not fully leafed 112 out before the last spring freeze are at risk of leaf tissue loss, damage to the xylem, and slowed canopy

development (Gu et al., 2008; Hufkens et al., 2012). To capture this 'high-risk' timeframe, we subtracted 114 12 days from the leafout date—which is the average rate of budburst across multiple studies and species (Donnelly et al., 2017; Flynn & Wolkovich, 2018; USA-NPN, 2019)—to establish a standardized estimate for 116 day of budburst since the majority of the individuals were missing budburst observations. We additionally considered a model that altered the timing between budburst and leafout for each species. For this alternate 118 model, we calculated budburst by subtracting 11 days from leafout for Aesculus hippocastanum and Betula 119 pendula, 12 days for Alnus glutinosa, 5 days for Fagus sylvatica, and 7 days for both Fraxinus excelsior and 120 Quercus robur based on growth chamber experiment data from phylogenetically related species (Buerki et al., 121 2010; Flynn & Wolkovich, 2018; Hipp et al., 2017; Wang et al., 2016). 122

123 Climate Data

We collected daily gridded climate data from the European Climate Assessment & Dataset (ECA&D) and 124 used the E-OBS 0.25 degree regular latitude-longitude grid from version 16. We used the daily minimum 125 temperature dataset to determine if a false spring occurred. False springs in this study were defined as 126 temperatures at or below -2.2°C (Schwartz, 1993) between budburst to leafout. We additionally tested this 127 model by changing the definition of a freezing temperature from -2.2°C (Schwartz, 1993) to -5°C (Lenz et al., 128 2013; Sakai & Larcher, 1987) in a separate model. In order to assess climatic effects, we calculated the mean spring temperature by using the daily mean temperature from March 1 through May 31. We used this date 130 range to best capture temperatures likely after chilling had accumulated to compare differences in spring forcing temperatures across sites (Basler & Körner, 2012; Körner et al., 2016). We collected NAO-index data 132 from the KNMI Climate Explorer CPC daily NAO time series and selected the NAO indices from November 133 until April to capture the effects of NAO on budburst for each region. We then took the mean NAO index 134 during these months (KNMI, 2018). Since the primary aim of the study is to predict false spring incidence 135 in a changing climate, we split the data: before temperature trends increased (1951-1983) and after trends 136 increased (1984-2016, Kharouba et al., 2018; Stocker et al., 2013) to represent recent climate change, which 137 will be referred to as the 'climate change' parameter henceforth.

Data Analysis

140 Simple regression models

We initally ran three simple regression models—following the same equation (below) but with varying response variables—to assess the effects of climate change on budburst, minimum temperatures between budburst and leafout and the number of false springs across species (Equation 1).

$$\epsilon_{i} \sim Normal(y_{i}, \sigma^{2})$$

$$y_{i} = \alpha_{[i]} + \beta_{ClimateChange_{[i]}} + \beta_{Species_{[i]}} + \beta_{ClimateChange \times Species_{[i]}} + \epsilon_{[i]}$$

$$(1)$$

144 Main Model

To best compare across the effects of each climatic and geographic variable, we scaled all of the predictors 145 and used a z-score following the binary predictor approach (Gelman & Hill, 2006). To control for spatial autocorrelation and to account for spatially structured processes independent from our regional predictors of 147 false springs, we generate an additional 'space' parameter for the model. To generate our space parameter we 148 first extracted spatial eigenvectors corresponding to our analyses' units and selected the subset that minimizes 149 spatial autocorrelation of the residuals of a model including all predictors except for the space parameter 150 (Bauman et al., 2017; Diniz-Filho et al., 2012) (see supplemental materials 'Methods: Spatial parameter' for 151 more details). We then took the eigenvector subset determined from the minimization of Moran's I in the 152 residuals (MIR approach) and regressed them against the above residuals—i.e. number of false springs vs. climatic and geographical factors. Finally we used the fitted values of that regression as our space parameter, 154 which, by definition, represents the portion of the variation in false springs that is both spatially structured and independent from all other predictors in the model (e.g. average spring temperature, elevation, etc. 156 Griffith & Peres-Neto, 2006; Morales-Castilla et al., 2012). A spatial predictor generated in this way has 157 three major advantages. First, it ensures that no spatial autocorrelation is left in model residuals. Second, it 158 avoids introducing collinearity issues with other predictors in the model. And third, it can be interpreted as 159 a latent variable summarizing spatial processes (e.g. local adaptation, plasticity, etc) occurring at multiple

161 scales.

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To estimate the probability of false spring risk across species and our predictors we used a Bayesian modeling approach. By including all parameters in the model, as well as species, we were able to distinguish the strongest contributing factors to false spring risk. We fit a Bernoulli distribution model (also know as a logistic regression) using mean spring temperature (written as MST in the model equation), NAO, elevation, distance from the coast (written as DistanceCoast in the model equation), space, and climate change as predictors and all two-way interactions and species as two-way interactions (Equation 2), using the brms package (Bürkner, 2017), version 2.3.1, in R (R Development Core Team, 2017), version 3.3.1, and was written as follows:

$$y_{i} \sim Binomial(1, p) \tag{2}$$

$$logit(p) = \alpha_{[i]} + \beta_{MST_{[i]}} + \beta_{DistanceCoast_{[i]}} + \beta_{Elevation_{[i]}} + \beta_{NAO_{[i]}} + \beta_{Space_{[i]}} + \beta_{ClimateChange_{[i]}} + \beta_{Species_{[i]}}$$

$$+ \beta_{MST \times Species_{[i]}} + \beta_{DistanceCoast \times Species_{[i]}} + \beta_{Elevation \times Species_{[i]}} + \beta_{NAO \times Species_{[i]}}$$

$$+ \beta_{Space \times Species_{[i]}} + \beta_{ClimateChange \times Species_{[i]}} + \beta_{MST \times ClimateChange_{[i]}}$$

$$+ \beta_{DistanceCoast \times ClimateChange_{[i]}} + \beta_{Elevation \times ClimateChange_{[i]}}$$

$$+ \beta_{NAO \times ClimateChange_{[i]}} + \beta_{Space \times ClimateChange_{[i]}}$$

We ran four chains of 4 000 iterations, each with 2 500 warm-up iterations for a total of 6 000 posterior

samples for each predictor. We evaluated our model performance based on R values that were close to one. 171 We also evaluated effective sample size estimates, which were 1 994 or above. We additionally assessed chain 172 convergence visually and posterior predictive checks. Due to the large number of observations in the data we 173 used the FASRC Cannon cluster (FAS Division of Science Research Computing Group at Harvard University) 174 to run the model. Model estimates were on the logit scale (shown in all tables) and were converted to probability percentages 176 in all figures for easier interpretation by using the 'divide by 4' rule (Gelman & Hill, 2006) and then back 177 converted to the original scale by multiplying by two standard deviations. We calculated overall estimates 178 (i.e., across species) of main effects in Figure 3, Figure S3 and Figure S4 from the average of the posteriors 179

of each effect by species. We report all estimated values in-text as mean \pm 98% uncertainty intervals, unless otherwise noted.

82 Results

Basic shifts in budburst and number of false springs

Day of budburst varied across the six species and across geographical gradients (Figure 1). Betula pendula,

Aesculus hippocastanum, Alnus glutinosa (Figure 1A-C) generally initiated budburst earlier than Fagus syl
vatica, Quercus robur, and Fraxinus excelsior (Figure 1D-F). Across all six species, higher latitude sites and

sites closer to the coast tended to initiate budburst later in the season (Figure 1).

Across species, budburst dates advanced -5.22 \pm 0.15 days after 1983 (Table S3) and minimum temperatures between budburst and leafout increased by 0.62 \pm 0.3°C after climate change (Table S4). This trend in advancing day of budburst for each species corresponds closely with increasing mean spring temperatures (Figure S1). While all species initiated budburst approximately seven days earlier (Figure 2A, Table S2 and Table S3), the average minimum temperature between budburst and leafout varied across the six species with Betula pendula and Aesculus hippocastanum experiencing the lowest minimum temperatures (Figure 2B), Quercus robur and Fraxinus excelsior experiencing the highest minimum temperatures, and Fraxinus excelsior experiencing the greatest variation (Figure 2B).

A simplistic view of changes in false springs—one that does not consider changes in climatic and geographic factors or effects of spatial autocorrelation—suggests that the number of false springs increased across species by 0.03% (± 0.05%) after climate change (i.e., after 1983), but with important variation by species (Figure 2C). Early-leafout species (Aesculus hippocastanum, Alnus glutinosa and Betula pendula) showed an increased risk whereas later bursting species (Fagus sylvatica, Quercus robur and Fraxinus excelsior) showed a decrease in risk (Table S5).

The effects of climatic and geographic variation coupled with climate change on false spring risk

Climatic and geographic factors underlie variation across years and space in false springs (Figure 3 and Table S6) before recent climate change (1983). Mean spring temperature had the strongest effect on false springs, 205 with warmer spring temperatures resulting is fewer false springs (Figure 3 and Table S6; comparable estimates come from using standardized variables—reported as 'standard units'. See Methods for more details). For 207 every 2°C increase in mean spring temperature there was a -7.64% in the probability of a false spring (-0.48 \pm 0.03 probability of false spring/standard unit). Distance from the coast had the second biggest effect on 209 false spring incidence. Individuals at sites further from the coast tended to have earlier leafout dates, which corresponded to an increased risk in false springs (Figure 3 and Table S6). For every 150km away from the 211 coast there was a 5.32\% increase in risk in false springs (0.4 \pm 0.03 probability of false spring/standard unit). 212 Sites at higher elevations also had higher risks of false spring incidence—likely due to more frequent colder 213 temperatures—with a 2.23% increase in risk for every 200m increase in elevation (0.19 \pm 0.04 probability 214 of false spring/standard unit, Figure 3 and Table S6). More positive NAO indices, which generally advance 215 leafout, slightly heightened the risk of false spring, with every 0.3 unit increase in NAO index there was a 216 1.91% increased risk in false spring or 0.14 ± 0.03 probability of false spring/standard unit (Figure 3 and Table S6). 218

These effects varied across species (Figure 4). While there were fewer false springs for each species with increasing mean spring temperatures, *Betula pendula*—an early-leafout species—had the greatest risk of false springs and *Fraxinus excelsior*—a late-leafout species—had the lowest risk (Figure 4A). There was an increased risk of false spring for all species at sites further from the coast (Figure 4B), with a sharp increase in risk for *Fraxinus excelsior* at sites further from the coast. With increasing elevation, all species had a greater risk of a false spring, except for *Fraxinus excelsior*, which had a slightly decreased risk at higher elevations (Figure 4C). With increasing NAO indices, the risk of false spring remained consistent for most species, except *Fagus sylvatica* experienced more with higher NAO indices (Figure 4D).

After climate change, the effects of these climatic and geographic factors on false spring risk shifted (Figure 3). Warmer sites still tended to have lower risks of false springs, but with climate change, increasing mean spring temperatures had much less of an effect on false spring risk with -2.84% in risk per 2° C (or -0.06 \pm 0.06 probability of false spring/standard unit versus -7.64% per 2° C or -0.48 before climate change; Figure 3 and Figure S2A). There was a slightly reduced risk in false springs further from the coast after climate change

(Figure 3 and Figure S2B) with 3.68% increase in risk per 150km (or 0.28 ± 0.07 probability of risk/standard unit versus 5.32% increase 150km or 0.4 ± 0.04 before climate change). The level of risk remained consistent before and after 1983 across elevations (Figure 3 and Figure S2C), with false spring risk being higher at higher elevations. After climate change, the rate of false spring incidence largely decreased with increasing NAO indices (Figure 3 and Figure S2D), now with a -9.15% in risk per 0.3 unit increase in the NAO index (or -0.69 ± 0.06 probability of false spring/standard unit or versus 1.91% 0.3 unit increase in the NAO index or 0.14 ± 0.03 before climate change). After climate change, NAO had the strongest effect on false spring risk, with higher NAO indices rendering fewer false springs.

Overall, there was a 4.01\% in risk of false springs across species (or a 0.16 increase in probability or 240 risk/standard unit), captured by the climate change predictor, which is leftover variability unexplained by the climatic and geographic factors after 1983. This residual effect of climate change varied strongly by species, 242 with an 8.86% increased risk in false springs after climate change for Aesculus hippocastanum (or 0.35 ± 0.03 probability of false spring/standard unit; Figure 3, Figure 4E and Table S6), a 10.54% increase for Alnus 244 glutinosa, a 10.29% increase for Betula pendula, and a 0.75% for Fagus sylvatica (or a 0.4 ± 0.08 , 0.41 ± 0.08 and 0.032 ± 0.08 probability of false spring/standard unit respectively; Figure 3, Figure 4E and Table S6). 246 Climate change decreased risk for Fraxinus excelsior by -4.27% and Quercus robur by -1.76% (or a -1.08 \pm 247 0.1 and -0.67 ± 0.08 probability of false spring/standard unit respectively; Figure 3, Figure 4E and Table S6). These residual effects of climate change on false spring risk are largely similar to the simple regression 249 model output testing the effects of climate change on the number of false springs across species (Table S5).

251 Sensitivity of results to duration of risk and temperature thresholds

Our results remained consistent (in direction and magnitude) when we applied different rates of leafout for each species (i.e., varied the length of time between estimated budburst and leafout). Mean spring temperature (-8.08% for every 2° C or -0.5 ± 0.04 probability of risk/standard unit) and distance from the coast (5.36% increase for every 150km or 0.4 ± 0.03 probability of risk/standard unit) were the strongest predictors for false spring risk (Figure S3 and Table S7). After climate change, there was a slight increase in false spring risk at higher elevations (Figure S3 and Table S7) compared to our main findings.

Results remained generally consistent also when we applied a lower temperature threshold for defining a false spring (i.e., -5°C), though there were more shifts in the magnitude of some effects, especially those of climate change. Mean spring temperature (-11.56% for every 2° or -0.72 ± 0.07 probability of risk/standard unit)

and elevation (7.35% increase in risk for every 200m or 0.63 ± 0.08 probability of risk/standard unit) were the strongest predictors, with a weaker effect of distance from the coast (2.75% for every 150km or $0.21 \pm$ 0.08 probability of risk/standard unit; Figure S4 and Table S8). There was much greater increasse in false spring risk due to the residual climate change effect across all six species (10.41% increase or 0.415 ± 0.07 probability of risk/standard unit; Figure S4 and Table S8).

Discussion

Integrating over 66 years of data, 11648 sites across Central Europe and major climatic and geographic factors, our results suggest climate change has reshaped the factors that drive false spring risk. In line with previous work, our results support that higher elevations tend to experience more false springs (Vitasse et al., 269 2018; Vitra et al., 2017) and sites that are generally warmer have lower risks of false springs (Wypych et al., 270 2016a). Individuals further from the coast typically initiated leafout earlier in the season, which subsequently 271 increased risk and, similarly, years with higher NAO indices experienced a slight increase in risk. But many of these factors have been reshaped by climate change, in particular the effect of climatic factors have shifted 273 dramatically compared to shifts in geographical factors. Across species, we find that NAO and mean spring temperature have shifted the most after 1983, while the effect of distance from the coast has only shifted 275 slightly and the effect of elevation has not at all shifted (Figure S2). These shifts in the influence of climatic and geographic factors in turn result in different effects of climate 277 change on species. The late-leafout species (e.g. Fraxinus excelsior and Quercus robur) have experienced decreases while the early-leafout species have experienced increases in risk (e.g., Aesculus hippocastanum. 279 Alnus glutinosa and Betula pendula). These species-specific effects integrate over shifts in the influence of climatic and geographic factors on false spring risk, as well as residual variation not explained by these factors. 281 This suggests that we have a robust understanding of what drivers underlie shifts in false spring for some species (i.e., Faqus sylvatica, which was not largely influenced by residual variation from climate change) versus those species where more understanding is most critically needed.

²⁸⁵ Climatic and geographic effects on false spring risk

Past studies, often considering few drivers of false spring events (Liu et al., 2018; Ma et al., 2018; Vitasse et al., 2018; Wypych et al., 2016b), have led to contradicting predictions in future false spring risk. Some

studies reported an increased risk at higher elevations after climate change (Vitasse et al., 2018), others 288 found an increase in risk only in Europe but not in other regions (Liu et al., 2018), while still others found a decrease in false spring risk across Central Europe (Wypych et al., 2016b). Research to date has also found 290 variation in false spring risk after climate change across species (Ma et al., 2018). By integrating both climate gradients and geographical factors, we were able to disentangle the major predictors of false spring risk and 292 merge these with species differences to determine which factors have the strongest effects on false spring risk. Mean spring temperature, distance from the coast and climate change were the strongest predictors 294 for false spring risk, however, NAO and elevation also affected risk, emphasizing the need to incorporate multiple predictors. Further, climatic and geographic factors varied in how consistent, or not, they were across species. Mean spring temperature, distance from the coast and NAO effects were fairly consistent 297 across species in direction, though Frazinus excelsior experienced a much greater increase in risk at sites further from the coast and Fagus sylvatica had a heightened risk to higher NAO indices compared to the 299 other species. Elevation was the only factor that varied in direction among the species with most species having an increased risk at higher elevations except for Fraxinus excelsior, which had a decreased risk. These 301 inconsistencies may capture range differences among species, with contrasting effects of factors on individuals closer to range edges (Chuine & Beaubien, 2001). 303

Since the onset of recent major climate change, the strength of these climatic and geographic effects has changed, highlighting the need to better understand and model shifting drivers of false spring. After climate 305 change, our results show a large decrease in risk of false spring with higher NAO indices. This could be because high NAO conditions no longer lead to temperatures low enough to trigger a false spring—that is, 307 with climate-change induced warming, high NAO conditions (and warmer baseline temperatures for that season) could reduce the likelihood of freezing temperatures, leading to a decreased risk of false spring 309 conditions (Screen, 2017). Conversely, we found an increased risk with warmer mean spring temperatures after climate change, which may be driven by our studied plant species responding very strongly to increased 311 spring warming with climate change (i.e., large advances in spring phenology, Figure S1), resulting in an 312 increased risk of exposure to false springs at these locations. Improved mechanistic models of how warming 313 temperatures affect budburst (Gauzere et al., 2017; Chuine et al., 2016) could improve our understanding of how NAO and mean spring temperatures contribute to false spring risk.

Variation in risk across species

By integrating climatic and geographic factors—i.e., mean spring temperature, elevation, distance from the coast and NAO indices—we can unravel phenological effects on the probability risk from these known factors 318 that contribute to an individual's level of false spring risk. Due to the prominent shifts in the climatic 319 and geographic factors with climate-change induced warming, we estimated that the residual (unexplained 320 by climatic and geographic factors) effects of climate change resulted in marked differences in risk between 321 early- and late-leafout species. Before 1983, false spring risk was slightly higher for species initiating leafout 322 earlier in the spring but overall the risk was more consistent across species (Figure 4E). After climate change, 323 however, species differences in risk amplified: the early-leafout species (i.e., Aesculus hippocastanum, Alnus glutinosa and Betula pendula) had an increased risk, the middle-leafout species—i.e. Fagus sylvatica—had a 325 similar level of risk as before and the later-leafout species (i.e., Fraxinus excelsior and Quercus robur) had a decreased risk (Figure 4E). 327

Our combined estimates are in agreement with the simple estimates of absolute changes in number of false springs across species (Figure 2C). These simple estimates, which also suggested an increase in risk for early-329 leafout species and a decline or no change for later-leafout species, correlated closely with estimated effects of climate change on species unexplained by climatic or geographic factors. The three early-leafout species 331 (Betula pendula, Aesculus hippocastanum, Alnus glutinosa) greater effects of residual climate change on false spring than the later species (Fagus sylvatica, Quercus robur, Fraxinus excelsior), suggesting the climatic and 333 geographic factors we examined are better capturing variation in false spring risk for later species—and that we still fundamentally lack information on what drives false spring risk for the early-leafout species, which are 335 also fundamentally the species with highest risk. While our model examines the major factors expected to influence false spring risk (Liu et al., 2018; Ma et al., 2018; Vitasse et al., 2018; Wypych et al., 2016b), these 337 results highlight the need to explore other climatic factors to improve forecasting. We expect factors that 338 affect budburst timing, such as shifts in over-winter chilling temperature or greater climatic stochasticity 339 earlier in the season, may help explain these discrepancies, but progress will require improved models of chilling beyond the current models, which have been mainly developed for crops (Luedeling & Brown, 2011; Dennis, 2003). 342

Our results and others (Ma et al., 2018) suggest phenological differences between species may predict their changing false spring risk with warming, but further understanding species differences will require more data and new approaches. Our focus on understanding shifting climatic and geographic factors led us to limit our study to the few species well sampled over space and time. Data on more species are available (e.g., Ma et al., 2018), but are sampled spatially and temporally much more variably. Thus, analyses of more species will need alternative datasets, or will need approaches that can detect and limit bias produced by uneven sampling of species across space and time.

Habitat preference and range differences among the species could also explain some of the species-specific 350 variation in the results, but would require data on more species—and species that vary strongly in their 351 climatic and geographic ranges—for robust analyses. The overall ranges of the predictors are similar across 352 species, but Betula pendula extends to the highest elevation and latitude and spans the greatest range of distances from the coast, while Quercus robur experiences the greatest range of mean spring temperatures. 354 Within our species, Betula pendula has the largest global distribution, extending the furthest north and east into Asia. The distribution of Frazinus excelsior extends the furthest south (into the northern region of 356 Iran). These range differences could potentially underlie the unexplained effect of climate change seen in our results and why the shifts in climatic and geographic factors did not explain much of the variation in false 358 spring risk across species. Fagus sylvatica was better explained by the model and this species has a smaller range, more confined to Central Europe. Future research that captures these spatial, temporal and climatic differences across myriad of species could greatly enhance predictions and help us understand these residual 361 effects of climate change.

Forecasting false springs

Our study shows how robust forecasting must integrate across major climatic and geographic factors that underlie false spring, and allow for variation in these factors across species and over time as warming continues. Of the four climatic and geographic factors we examined, only the effect of elevation remained constant before 366 and after climate change and there was only a slight change in the effect of distance from the coast, suggesting greater shifts in climatic factors but more stability with geographic factors. This is perhaps not surprising as 368 climate change is shifting critical spring temperatures—and ultimately the environmental drivers of phenology (Gauzere et al., 2019)—and reshaping the temporal and spatial dynamics of how climate affects budburst, 370 leafout and freezing temperatures. A general result is that climate change (e.g., warming) tends to be 371 amplified at higher elevations and further away from the coast, although how prevalent these patterns are 372 across different regions is open for discussion (Pepin et al., 2015; Rangwala & Miller, 2012; Giorgi et al., 1997). Additional studies show evidence for more consistent budburst across elevations with climate change (Vitasse et al., 2018), thus highlighting the complexity that robust forecasting will require.

Moving forward, more data on more species will be critical for estimates at community or ecosystem scales 376 (at least in species-rich ecosystems). Related to this, more research on the effects of climate change on both 377 budburst and leafout, the timing when individuals are most at risk to spring freeze damage (Chamberlain 378 et al., 2019; Lenz et al., 2016) and on what temperatures cause leaf damage will help better understand 379 differences across species. Though we found that differing rates of leafout across species had minimal effects 380 on predicting risk, we did find that the lower temperature threshold can have an impact on model estimates 381 (and thus forecasts), with lower temperature thresholds (i.e., -5°C versus -2.2°C) predicting increased risk across all six study species. Our study uses an index of false spring risk, to estimate when damage may 383 have occurred; it does not assess the intensity or severity of the false spring events observed, nor does it record the amount of damage to individuals. Other research has shown that this temperature threshold 385 may vary importantly by species (Bennett et al., 2018; Körner et al., 2016; Lenz et al., 2013; Zhuo et al., 2018). Some species or individuals may be less freeze tolerant (i.e., are damaged from higher temperatures 387 than -2.2°C), whereas other species or individuals may be able to tolerate temperatures as low as -8.5°C (Lenz et al., 2016). Further, cold tolerance can be highly influenced by fall and winter climatic dynamics that influence tissue hardiness (Charrier et al., 2011; Hofmann & Bruelheide, 2015; Vitasse et al., 2014) and 390 can also influence budburst timing (Morin et al., 2007). Thus, we expect budburst, leafout and hardiness are likely integrated and that useful forecasting will require far better species-specific models of all these 392 factors—including whether budburst and hardiness may be inter-related.

Our results highlight how climate change complicates forecasting through multiple levels. It has shifted the influence of climatic and geographic factors, fundamentally reshaping relationships with major climatic factors such that relationships before climate change no longer hold. It has also magnified species-level variation in false spring risk. Layered onto this complexity is residual effects of climate change that suggests we are missing key factors that drive interspecific variation in false spring risk. Our study focuses on one region (i.e., Central Europe) with high-quality and abundant data and we hope that our approach can be applied to other systems as more and more data becomes available. Our analysis and other analyses like ours are important for identifying not only which species will be more vulnerable to false springs, but also where in their distributions they will be at risk. Integrating these findings into future models will provide more robust forecasts and help us begin to unravel the complexities of climate change effects across species.

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References

- Augspurger CK (2009) Spring 2007 warmth and frost: phenology, damage and refoliation in a temperate deciduous forest. Functional Ecology, 23, 1031–1039. doi:10.1111/j.1365-2435.2009.01587.x.
- Augspurger CK (2013) Reconstructing patterns of temperature, phenology, and frost damage over 124 years:

 Spring damage risk is increasing. *Ecology*, **94**, 41–50. doi:10.1890/12-0200.1.
- Basler D, Körner C (2012) Photoperiod sensitivity of bud burst in 14 temperate forest tree species. *Agricultural*and Forest Meteorology, **165**, 73–81. doi:10.1016/j.agrformet.2012.06.001.
- Bauman D, Drouet T, Dray S, Vleminckx J (2017) Disentangling good from bad practices in the selection of
 spatial or phylogenetic eigenvectors. *Ecography*, **0**. doi:10.1111/ecog.03380.
- Bennett JM, Calosi P, Clusella-Trullas S, et al. (2018) Globtherm, a global database on thermal tolerances for aquatic and terrestrial organisms. Scientific data, 5, 180022.
- Buerki S, Lowry II P, Alvarez N, Razafimandimbison S, Kupfer P, Callmander M (2010) Phylogeny and circumscription of *Sapindaceae* revisited: Molecular sequence data, morphology and biogeography support recognition of a new family, *Xanthoceraceae*. *Plant Ecology and Evolution*, **143**, 148–159. doi:10.5091/plecevo.2010.437.
- Bürkner PC (2017) brms: An R Package for Bayesia Multilevel Models. Journal of Statistical Software, 80,

 1–28.
- Chamberlain CJ, Cook BI, de Cortazar Atauri IG, Wolkovich EM (2019) Rethinking false spring risk. Global
 Change Biology, 25, 2209–2220. doi:10.1111/gcb.14642.
- Charrier G, Bonhomme M, Lacointe A, Améglio T (2011) Are budburst dates, dormancy and cold acclimation in walnut trees (juglans regia l.) under mainly genotypic or environmental control? *International Journal of Biometeorology*, **55**, 763–774. doi:10.1007/s00484-011-0470-1. URL https://doi.org/10.1007/s00484-
- 429 011-0470-1.

- Chen X, Hu B, Yu R (2005) Spatial and temporal variation of phenological growing season and climate change impacts in temperate eastern China. *Global Change Biology*, **11**, 1118–1130. doi:10.1111/j.1365-
- 432 2486.2005.00974.x.
- Chmielewski FM, Rötzer T (2001) Response of tree phenology to climate change across Europe. Agricultural
 and Forest Meteorology, 108, 101 112. doi:https://doi.org/10.1016/S0168-1923(01)00233-7.
- Chuine I, Beaubien EG (2001) Phenology is a major determinant of tree species range. *Ecology Letters*, 4,
- 500-510. doi:10.1046/j.1461-0248.2001.00261.x. URL https://onlinelibrary.wiley.com/doi/abs/10.
- 1046/j.1461-0248.2001.00261.x.
- Chuine I, Bonhomme M, Legave JM, García de Cortázar-Atauri I, Charrier G, Lacointe A, Améglio T
- (2016) Can phenological models predict tree phenology accurately in the future? the unrevealed hurdle
- of endodormancy break. Global Change Biology, 22, 3444–3460. doi:10.1111/gcb.13383. URL http:
- //dx.doi.org/10.1111/gcb.13383.
- Dennis F (2003) Problems in standardizing methods for evaluating the chilling requirements for the breaking of dormancy in buds of woody plants. *HortScience*, **38**, 347–350.
- Diniz-Filho JAF, Bini LM, Rangel TF, Morales-Castilla I, Olalla-Tárraga MÁ, Rodríguez MÁ, Hawkins BA
 (2012) On the selection of phylogenetic eigenvectors for ecological analyses. *Ecography*, **35**, 239–249.
- Donnelly A, Yu R, Caffarra A, et al. (2017) Interspecific and interannual variation in the duration of spring
 phenophases in a northern mixed forest. Agricultural and Forest Meteorology, 243, 55–67.
- Flynn DFB, Wolkovich EM (2018) Temperature and photoperiod drive spring phenology across all species in a temperate forest community. *New Phytologist.* doi:10.1111/nph.15232.
- Gauzere J, Delzon S, Davi H, Bonhomme M, Garcia de Cortazar-Atauri I, Chuine I (2017) Integrating
 interactive effects of chilling and photoperiod in phenological process-based models. A case study with two
- European tree species: Fagus sylvatica and Quercus petraea. Agricultural and Forest Meteorology, pp. 9–20.
- 453 Gauzere J, Lucas C, Ronce O, Davi H, Chuine I (2019) Sensitivity analysis of tree phenology models reveals
- increasing sensitivity of their predictions to winter chilling temperature and photoperiod with warming
- climate. Ecological Modelling, 411, 108805. doi:https://doi.org/10.1016/j.ecolmodel.2019.108805. URL
- http://www.sciencedirect.com/science/article/pii/S0304380019303138.

- Gelman A, Hill J (2006) Data analysis using regression and multilevel/hierarchical models. Cambridge university press.
- Giorgi F, Hurrell JW, Marinucci MR, Beniston M (1997) Elevation dependency of the surface climate change 459 signal: a model study. Journal of Climate, 10, 288–296
- Griffith DA, Peres-Neto PR (2006) Spatial modeling in ecology: the flexibility of eigenfunction spatial anal-461 yses. Ecology, 87, 2603–2613.
- Gu L, Hanson PJ, Post WM, et al. (2008) The 2007 Eastern US spring freeze: Increased cold damage in a 463 warming world. BioScience, 58, 253. doi:10.1641/B580311. 464
- Hänninen H (1991) Does climatic warming increase the risk of frost damage in northern trees? Plant, Cell 465 & Environment, 14, 449–454. doi:10.1111/j.1365-3040.1991.tb01514.x. 466
- Hipp A, S Manos P, González-Rodríguez A, et al. (2017) Sympatric parallel diversification of major oak clades in the Americas and the origins of Mexican species diversity. New Phytologist, 217. doi:10.1111/nph.14773. 468
- Hofmann M, Bruelheide H (2015) Frost hardiness of tree species is independent of phenology and macroclimatic niche. Journal of Biosciences, 40, 147–157. doi:10.1007/s12038-015-9505-9. 470
- Hufkens K, Friedl MA, Keenan TF, Sonnentag O, Bailey A, O'Keefe J, Richardson AD (2012) Ecological impacts of a widespread frost event following early spring leaf-out. Global Change Biology, 18, 2365–2377. 472 doi:10.1111/j.1365-2486.2012.02712.x.
- Inouye DW (2008) Effects of climate change on phenology, frost damage, and floral abundance of montane 474 wildflowers. *Ecology*, **89**, 353–362. 475
- IPCC (2015) Climate change 2014: mitigation of climate change, vol. 3. Cambridge University Press.

473

- Kharouba HM, Ehrlén J, Gelman A, Bolmgren K, Allen JM, Travers SE, Wolkovich EM (2018) Global shifts 477 in the phenological synchrony of species interactions over recent decades. Proceedings of the National Academy of Sciences, 115, 5211–5216. doi:10.1073/pnas.1714511115. 479
- Klosterman S, Hufkens K, Richardson AD (2018) Later springs green-up faster: the relation between onset and completion of green-up in deciduous forests of North America. International Journal of Biometeorology. 481 doi:10.1007/s00484-018-1564-9. 482

- KNMI (2018) Daily CPC NAO data. URL https://climexp.knmi.nl/getindices.cgi?WMO=NCEPData/
 cpc_nao_daily&STATION=NAO&TYPE=i&id=someone@somewhere&NPERYEAR=366.
- Kollas C, Körner C, Randin CF (2014) Spring frost and growing season length co-control the cold range limits of broad-leaved trees. *Journal of Biogeography*, **41**, 773–783. doi:10.1111/jbi.12238.
- Körner C, Basler D, Hoch G, et al. (2016) Where, why and how? Explaining the low-temperature range limits of temperate tree species. Journal of Ecology, **104**, 1076–1088. doi:10.1111/1365-2745.12574. URL http://dx.doi.org/10.1111/1365-2745.12574.
- Kramer K (1994) A modelling analysis of the effects of climatic warming on the probability of spring frost
 damage to tree species in The Netherlands and Germany. *Plant, Cell & Environment*, **17**, 367–377. doi:
 10.1111/j.1365-3040.1994.tb00305.x.
- Kukal MS, Irmak S (2018) U.S. Agro-Climate in 20th century: growing degree days, first and last frost, growing season length, and impacts on crop yields. *Scientific Reports*, 8. doi:10.1038/s41598-018-25212-2.
- Labe Z, Ault T, Zurita-Milla R (2016) Identifying anomalously early spring onsets in the CESM large ensemble project. Climate Dynamics, 48, 3949–3966. doi:10.1007/s00382-016-3313-2.
- Lenz A, Hoch G, Körner C, Vitasse Y (2016) Convergence of leaf-out towards minimum risk of freezing

 damage in temperate trees. Functional Ecology, 30, 1–11. doi:10.1111/1365-2435.12623.
- Lenz A, Hoch G, Vitasse Y, Körner C (2013) European deciduous trees exhibit similar safety margins against
 damage by spring freeze events along elevational gradients. New Phytologist, 200, 1166–1175. doi:10.1111/
 nph.12452.
- Liu Q, Piao S, Janssens IA, et al. (2018) Extension of the growing season increases vegetation exposure to frost. Nature Communications, 9. doi:10.1038/s41467-017-02690-y.
- Liu X, Yin Z, Shao X, Qin N (2006) Temporal trends and variability of daily maximum and minimum, extreme temperature events, and growing season length over the eastern and central Tibetan Plateau during 1961–2003. Journal of Geophysical Research: Atmospheres, 111. doi:10.1029/2005JD006915.
- Luedeling E, Brown PH (2011) A global analysis of the comparability of winter chill models for fruit and nut trees. International Journal of Biometeorology, 55, 411–421.
- Ma Q, Huang JG, Hänninen H, Berninger F (2018) Divergent trends in the risk of spring frost damage to trees in europe with recent warming. *Global Change Biology*, **0**. doi:10.1111/gcb.14479.

- Martin M, Gavazov K, Körner C, Hattenschwiler S, Rixen C (2010) Reduced early growing season freezing
- resistance in alpine treeline plants under elevated atmospheric CO₂. Global Change Biology, 16, 1057–1070.
- doi:10.1111/j.1365-2486.2009.01987.x.
- Meier U (2001) Growth stages of mono-and dicotyledonous plants BBCH Monograph Edited by Uwe Meier
- Federal Biological Research Centre for Agriculture and Forestry. Agriculture, 12, 141—147 ST Geo-
- chemical study of the organic mat. doi:10.5073/bbch0515.
- Menzel A, Helm R, Zang C (2015) Patterns of late spring frost leaf damage and recovery in a European beech
- (Fagus sylvatica L.) stand in south-eastern Germany based on repeated digital photographs. Frontiers in
- Plant Science, 6, 110. doi:10.3389/fpls.2015.00110.
- 520 Morales-Castilla I, Olalla-Tarraga MA, Purvis A, Hawkins BA, Rodriguez MA (2012) The imprint of cenozoic
- migrations and evolutionary history on the biogeographic gradient of body size in new world mammals.
- The American Naturalist, 180, 246-256.
- 523 Morin X, Améglio T, Ahas R, et al. (2007) Variation in cold hardiness and carbohydrate concentration from
- dormancy induction to bud burst among provenances of three European oak species. Tree Physiology, 27,
- 817-825. doi:10.1093/treephys/27.6.817. URL https://doi.org/10.1093/treephys/27.6.817.
- Pepin N, Bradley RS, Diaz HF, et al. (2015) Elevation-dependent warming in mountain regions of the
- world. Nature Climate Change, 5, 424-430. doi:10.1038/nclimate2563. URL https://doi.org/10.1038/
- nclimate2563.
- R Development Core Team (2017) R: A language and environment for statistical computing. R Foundation
- for Statistical Computing, Vienna, Austria.
- Rangwala I, Miller JR (2012) Climate change in mountains: a review of elevation-dependent warming and
- its possible causes. Climatic Change, 114, 527-547. doi:10.1007/s10584-012-0419-3. URL https://doi.
- org/10.1007/s10584-012-0419-3.
- Richardson AD, Keenan TF, Migliavacca M, Ryu Y, Sonnentag O, Toomey M (2013) Climate change, phe-
- nology, and phenological control of vegetation feedbacks to the climate system. Agricultural and Forest
- *Meteorology*, **169**, 156 173. doi:https://doi.org/10.1016/j.agrformet.2012.09.012.
- Sakai A, Larcher W (1987) Frost Survival of Plants. Springer-Verlag.

- Scheifinger H, Menzel A, Koch E, Peter C (2003) Trends of spring time frost events and phenological dates
- in Central Europe. Theoretical and Applied Climatology, 74, 41–51. doi:10.1007/s00704-002-0704-6.
- Schwartz MD (1993) Assessing the onset of spring: A climatological perspective. *Physical Geography*, **14(6)**,
- 536-550.
- Screen JA (2017) The missing northern european winter cooling response to arctic sea ice loss. Nature
- 543 Communications, 8, 14603. doi:10.1038/ncomms14603. URL https://doi.org/10.1038/ncomms14603.
- 544 Sgubin G, Swingedouw D, Dayon G, de Cortázar-Atauri IG, Ollat N, Pagé C, van Leeuwen C (2018) The risk
- of tardive frost damage in French vineyards in a changing climate. Agricultural and Forest Meteorology,
- **250-251**, 226 242. doi:https://doi.org/10.1016/j.agrformet.2017.12.253.
- 547 Stocker TF, Qin D, Plattner GK, et al. (2013) Climate Change 2013: The Physical Science Basis. Contribution
- of Working Group I to the Fifth Assessment Report of the Intergovernmental Panel on Climate Change,
- 549 1535 pp.
- Templ B, Koch E, Bolmgren K, et al. (2018) Pan European Phenological database (PEP725): a single point
- of access for European data. International Journal of Biometeorology, 62, 1109–1113. doi:10.1007/s00484-
- 018-1512-8. URL https://doi.org/10.1007/s00484-018-1512-8.
- 553 USA-NPN (2019) Plant and animal phenology data. USA National Phenology Network. doi:10.5066/
- F78S4N1V. URL http://doi.org/10.5066/F78S4N1V.
- Vitasse Y, Lenz A, Körner C (2014) The interaction between freezing tolerance and phenology in temperate
- deciduous trees. Frontiers in Plant Science, 5. doi:10.3389/fpls.2014.00541.
- ⁵⁵⁷ Vitasse Y, Schneider L, Rixen C, Christen D, Rebetez M (2018) Increase in the risk of exposure of forest and
- fruit trees to spring frosts at higher elevations in Switzerland over the last four decades. Agricultural and
- Forest Meteorology, 248, 60 69. doi:https://doi.org/10.1016/j.agrformet.2017.09.005.
- Vitra A, Lenz A, Vitasse Y (2017) Frost hardening and dehardening potential in temperate trees from winter
- to budburst. New Phytologist, **216**, 113–123. doi:10.1111/nph.14698.
- ⁵⁶² Wang N, McAllister HA, Bartlett PR, Buggs RJA (2016) Molecular phylogeny and genome size evolution of
- the genus Betula (Betulaceae). Annals of Botany, 117, 1023–1035. doi:10.1093/aob/mcw048.
- Wolkovich EM, Cook BI, Allen JM, et al. (2012) Warming experiments underpredict plant phenological
- responses to climate change. *Nature*, **485**, 18–21. doi:10.1038/nature11014.

- ⁵⁶⁶ Wypych A, Sulikowska A, Ustrnul Z, Czekierda D (2016a) Variability of growing degree days in Poland
- in response to ongoing climate changes in Europe. International Journal of Biometeorology, 61, 49–59.
- doi:10.1007/s00484-016-1190-3. URL http://dx.doi.org/10.1007/s00484-016-1190-3.
- Wypych A, Ustrnul Z, Sulikowska A, Chmielewski FM, Bochenek B (2016b) Spatial and temporal variability of
- the frost-free season in Central Europe and its circulation background. International Journal of Climatology,
- **37**, 3340–3352. doi:10.1002/joc.4920.
- ⁵⁷² Zhuo X, Zheng T, Zhang Z, et al. (2018) Genome-wide analysis of the NAC transcription factor gene family
- reveals differential expression patterns and cold-stress responses in the woody plant Prunus mume. Genes,
- 9. doi:10.3390/genes9100494.
- ⁵⁷⁵ Zohner CM, Benito BM, Svenning JC, Renner SS (2016) Day length unlikely to constrain climate-driven
- shifts in leaf-out times of northern woody plants. Nature Climate Change, 6, 1120–1123. doi:10.1038/
- nclimate 3138.

Tables and Figures

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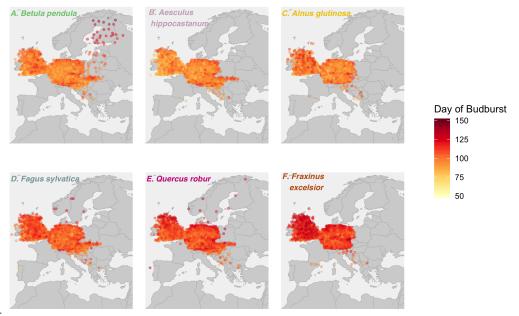


Figure 1: The average day of budburst mapped by site for each species (ordered by day of budburst starting with $Betula\ pendula$ as the earliest budburst date to $Fraxinus\ excelsior$). Species names are color-coded to match figures throughout the text.

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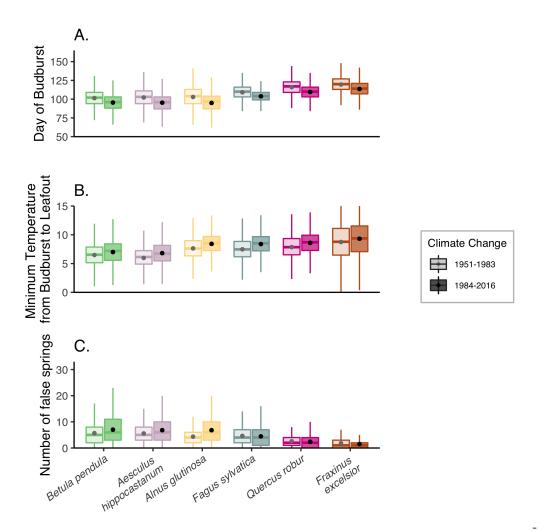


Figure 2: Day of budburst (A.), minimum temperatures between budburst and leafout (B.) and number of false springs (C.) before and after 1983 across species for all sites. Box and whisker plots show the 25th and 75th percentiles (i.e., the interquartile range) with notches indicating 95% uncertainty intervals. Dots and error bars overlaid on the box and whisker plots represent the model regression outputs (Tables S3-S5). Error bars from the model regressions indicate 98% uncertainty intervals but, given the number of sites, are quite small and thus not easily visible (see Tables S3-S5). Species are ordered by day of budburst and are color-coded to match the other figures.

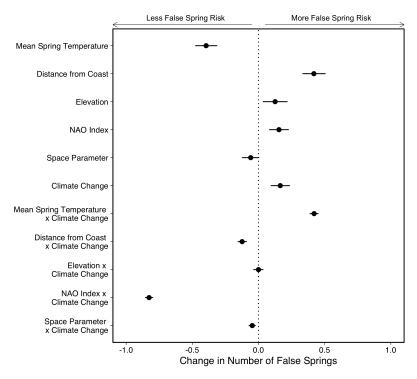


Figure 3: Effects of species, climatic and geographical predictors on false spring risk. More positive values indicate an increased probability of a false spring whereas more negative values suggest a lower probability of a false spring. Dots and lines show means and 98% uncertainty intervals. There were 582,211 zeros and 172,877 ones for false springs in the data. See Table S6 for full model output.

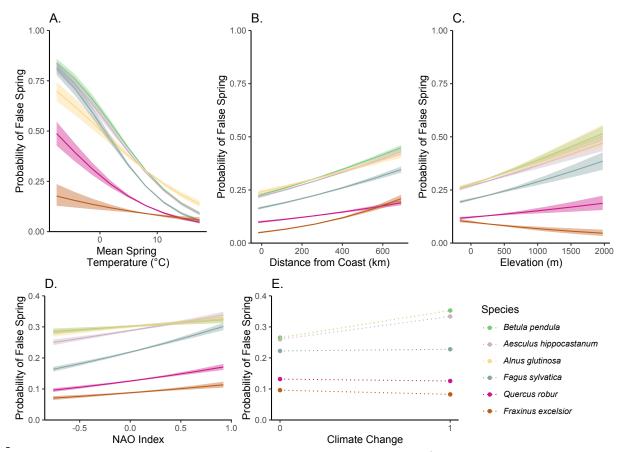


Figure 4: Species-level variation across geographic and spatial predictors (i.e., mean spring temperature (A.), distance from the coast (B.), elevation (C.), and NAO index (D.)). Lines and shading are the mean and 98% uncertainty intervals for each species. To reflect the raw data, we converted the model output back to the original scale for the x-axis in each panel. See Table S6 for full model output.