

# Class 5: Linear Model Selection and Regularization

MSA 8150: Machine Learning for Analytics

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Alireza Aghasi

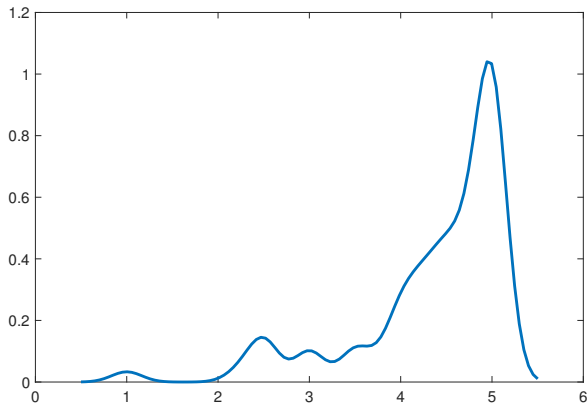
Institute for Insight, Georgia State University

# Quiz 1 Results

mean: 4.3542

std: 0.7267

Histogram:

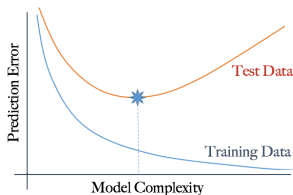


# **Brief Overview of Cross-Validation**

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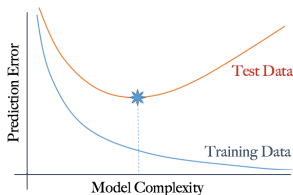
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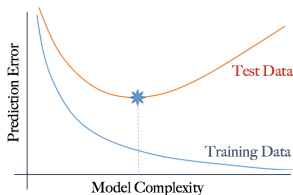
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- Instead through a process of splitting the data into training and validations sets, we were able to use LOOCV or K-Fold CV as estimates of the test error
- We discussed why K-Fold CV is a more desirable estimate, computationally and statistically

## **Adjusting the Training Statistics for Test Error Approximation**

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- We introduce few other ways of adjusting the training error **to make it a better representative of the test error**
- These adjustments are **not as reliable as Cross validation**, but they are easier to **calculate**
- These quantities were **more widely used before** the widespread use of computers for regression and machine learning
- Now that computers can help performing multiple fits computationally fast enough, often K-Fold CV is considered as the desirable test error approximation

# List of Other Techniques

Methods to adjust the training error for the number of variables to estimate the test MSE:

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- $C_p$  statistic
- Akaike information criterion (AIC)
- Bayesian information criterion (BIC)
- Adjusted  $R^2$

- For a fitted least squares model with  $d$  predictors

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- Becomes a better estimate of the test error as the sample size,  $n$ , increases

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- When  $n < 7$ , BIC imposes a smaller penalty on the number of variables, but for  $n > 7$  that  $\log n > 2$  the penalty is larger
- In other words in standard observation regimes where  $n$  is sufficiently large, BIC tends to pick smaller models than AIC or  $C_p$

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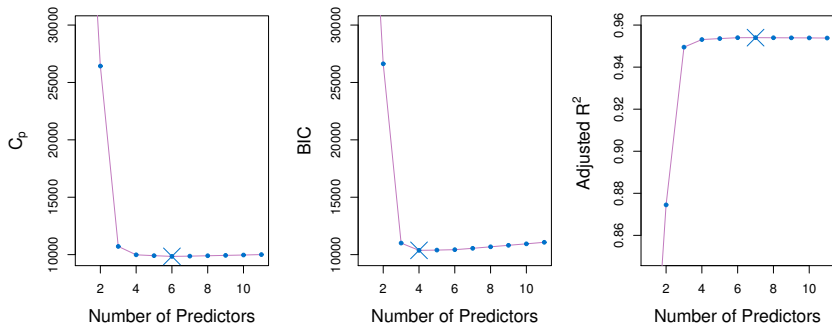
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- Unlike the other three statistics that being small indicates a better model, for adjusted  $R^2$  we are interested in models that tend to generate values closer to 1
- The use of  $C_p$ , AIC and BIC is more motivated in statistical learning theory than the adjusted  $R^2$

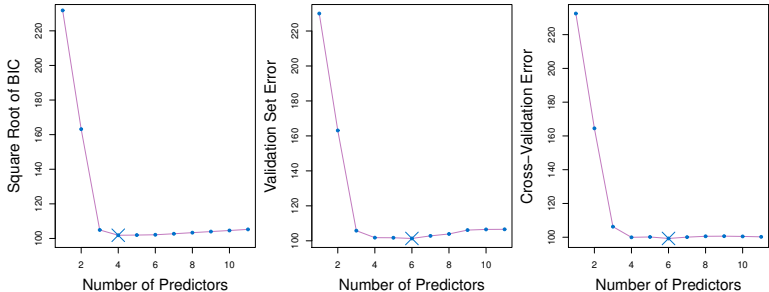
## Example Comparing the Performances



$C_p$ , BIC, and adjusted  $R^2$  for the best models of each size for the Credit data set



# Comparison Against CV Techniques



- The results are not much different
- Note that nowadays CV methods are computationally fast to implement and regardless of the model can always be used as a reliable selection tool

# How to Use These Statistics in Model Selection

- **Best subset selection** formal procedure (NP-hard and computationally not possible for large  $p$ )

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**Algorithm 6.1** *Best subset selection*

---

1. Let  $\mathcal{M}_0$  denote the *null model*, which contains no predictors. This model simply predicts the sample mean for each observation.
  2. For  $k = 1, 2, \dots, p$ :
    - (a) Fit all  $\binom{p}{k}$  models that contain exactly  $k$  predictors.
    - (b) Pick the best among these  $\binom{p}{k}$  models, and call it  $\mathcal{M}_k$ . Here *best* is defined as having the smallest RSS, or equivalently largest  $R^2$ .
  3. Select a single best model from among  $\mathcal{M}_0, \dots, \mathcal{M}_p$  using cross-validated prediction error,  $C_p$  (AIC), BIC, or adjusted  $R^2$ .
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**Algorithm 6.2** *Forward stepwise selection*

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1. Let  $\mathcal{M}_0$  denote the *null* model, which contains no predictors.
  2. For  $k = 0, \dots, p - 1$ :
    - (a) Consider all  $p - k$  models that augment the predictors in  $\mathcal{M}_k$  with one additional predictor.
    - (b) Choose the *best* among these  $p - k$  models, and call it  $\mathcal{M}_{k+1}$ . Here *best* is defined as having smallest RSS or highest  $R^2$ .
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- Forward selection can even be used when  $n < p$

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  2. For  $k = p, p - 1, \dots, 1$ :
    - (a) Consider all  $k$  models that contain all but one of the predictors in  $\mathcal{M}_k$ , for a total of  $k - 1$  predictors.
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- Backward selection requires  $p < n$  (to allow the full model to be fit)



# What are Shrinkage Methods and Why Useful?

You would probably hear **Ridge Regression** and **LASSO** quite often

- The subset selection methods use least squares to fit a linear model that contains a subset of the predictors
- As an alternative, we can fit a model containing all  $p$  predictors using a technique that constrains or regularizes the coefficient estimates, or equivalently, that shrinks the coefficient estimates towards zero
- It may not be immediately obvious why such a constraint should improve the fit, but it turns out that shrinking the coefficient estimates can significantly **reduce the model variance**

# Ridge Regression

- Recall that the least squares fitting procedure estimates  $\beta_0, \beta_1, \dots, \beta_p$  using the values that minimize

$$RSS = \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2$$

- In contrast, the ridge regression coefficient estimates  $\hat{\beta}^R$  are the values that minimize

$$RSS_{Ridge} = \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p \beta_j^2 = RSS + \lambda \sum_{j=1}^p \beta_j^2$$

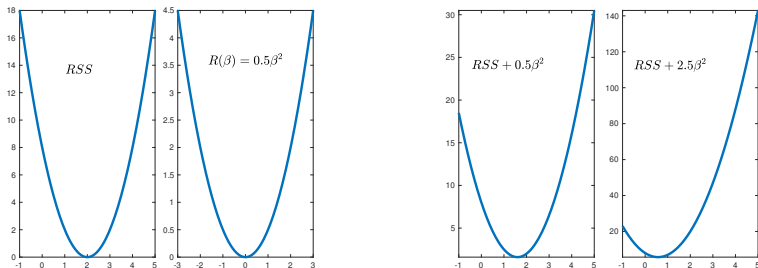
- Here,  $\lambda$  is a tuning parameter

# Ridge Regression

- As with least squares, ridge regression seeks coefficient estimates that fit the data well, by making the RSS small
- However, the second term,  $\lambda \sum_{j=1}^p \beta_j^2$ , called a shrinkage penalty, encourages solutions that are close to zero, and so it has the effect of shrinking the estimates of  $\beta_j$  towards zero
- The tuning parameter  $\lambda$  serves to control the relative impact of these two terms on the regression coefficient estimates (trade off between bias and variance)
- Selecting a good value for  $\lambda$  is critical; often cross-validation is used for this

# Effect of Increasing $\lambda$ on the $\beta$

- The figure below shows how increasing the Ridge penalty pushes the minimizers of the mixed RSS objective to zero



## Shrinkage Example

- Previously from the homework assignments you remember that the least squares solution to fit data point  $(x_1, y_1), \dots, (x_n, y_n)$  was obtained via the minimization:

$$\min_{\beta} \sum_{i=1}^N (y_i - \beta x_i)^2 \quad \therefore \quad \hat{\beta} = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2}$$

- We can show that if we run the Ridge regression problem

$$\min_{\beta} \sum_{i=1}^N (y_i - \beta x_i)^2 + \lambda \beta^2$$

the new estimate becomes

$$\hat{\beta}^R = \frac{\sum_{i=1}^n x_i y_i}{\lambda + \sum_{i=1}^n x_i^2}$$

- Note how increasing  $\lambda$  pushes  $\hat{\beta}^R$  towards zero

## In Class Exercise

- For the simple regression problem of fitting  $(x_1, y_1), \dots, (x_n, y_n)$ , to the model  $y = \beta_0 + \beta_1 x$  show that the least-squares estimates for the Ridge regularized objective

$$\sum_{i=1}^N (y_i - \beta_0 - \beta_1 x_i)^2 + \lambda(\beta_0^2 + \beta_1^2)$$

are

$$\hat{\beta}_1^R = \frac{\sum_{i=1}^n x_i y_i - \frac{n^2}{n+\lambda} \bar{x} \bar{y}}{\lambda + \sum_{i=1}^n x_i^2 + \frac{n^2}{n+\lambda} \bar{x}^2}, \quad \hat{\beta}_0^R = \frac{1}{n + \lambda} \left( \sum_{i=1}^n y_i - \hat{\beta}_1^R \sum_{i=1}^n x_i \right)$$

# What Happens in Multiple Regression?

- In this case we previously had

$$RSS = (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^\top (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$$

which led to

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$$

- In the case of regularized problem ( $\|\cdot\|$  denotes the L-2 norm)

$$(\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^\top (\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) + \lambda \|\boldsymbol{\beta}\|^2$$

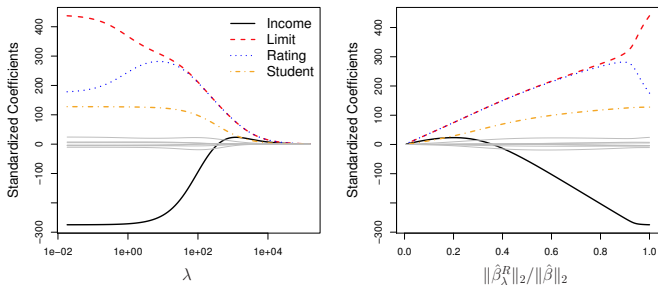
we will have

$$\hat{\boldsymbol{\beta}}^R = (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^\top \mathbf{y}$$

where  $\mathbf{I}$  is the identity matrix

# Credit Data Example

- Left: each curve corresponds to the ridge regression coefficient estimate for one of the ten variables, plotted as a function of  $\lambda$
- The right-hand panel displays the same ridge coefficient estimates as the left-hand panel, but instead of displaying  $\lambda$  on the x-axis, we display  $\|\hat{\beta}^R\|/\|\hat{\beta}\|$  (how much **shrinkage** happens by increasing  $\lambda$ )





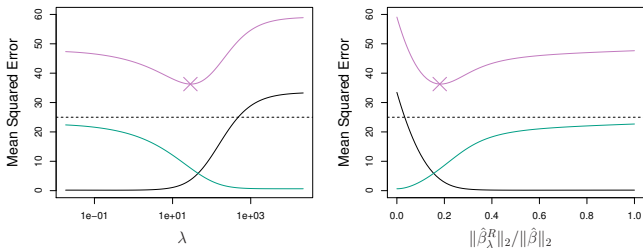
# Scaling of the Predictors

- In the standard least-squares if we scale a feature value by  $c$ , the corresponding coefficient scales by  $c^{-1}$
- However when we have the Ridge regularized objective, this is no more the case
- To see a consistent behavior, for the Ridge regularized problem we often work with standardized features:

$$\tilde{x}_{i,j} = \frac{x_{ij}}{\sqrt{\frac{1}{n} \sum_{i=1}^n (x_{ij} - \bar{x}_j)^2}}$$

# Bias-Variance Trade-Off

- A toy example: squared bias (black), variance (green), and test mean squared error (purple) for the ridge regression predictions on a simulated data set, as a function of  $\lambda$  and  $\|\hat{\beta}^R\|/\|\hat{\beta}\|$ . The horizontal dashed lines indicate the minimum possible MSE (**the standard least squares,  $\lambda = 0$  in nowhere close**). The purple crosses indicate smallest ridge regression model MSE values



- Remember (test error = bias + variance + noise variance)

**Questions?**

# References



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