# MSA 8200 Predictive Analytics Week5: SARIMA and Case Study

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#### Outline

- Integrated models for nonstationary data (ARIMA).
- Cointegration.
- Multiplicative seasonal ARIMA models.
- Walmart sales prediction example.
- Regression with correlated errors.

#### Random walk:

$$x_t = x_{t-1} + w_t$$

- $\nabla x_t = w_t$  is stationary.
- First order difference.

A nonstationary trend + a zero mean stationary component:

$$x_t = \mu_t + y_t$$

1.  $\mu_t = \beta_0 + \beta_1 t$ 

$$\nabla x_t = x_t - x_{t-1} = \beta_1 + y_t - y_{t-1} = \beta_1 + \nabla y_t$$

2.  $\mu_t = \mu_{t-1} + \gamma_t$ ,  $\gamma_t$  stationary.

$$\nabla x_t = \gamma_t + \nabla y_t$$

3.  $\mu_t = \sum_{j=0}^k \beta_j t^j \Rightarrow \nabla^k x_t$  stationary.

#### Example:

  
 • 
$$x_t = \mu_t + y_t$$
 
$$\mu_t = \mu_{t-1} + \gamma_t \text{ and } \gamma_t = \gamma_{t-1} + e_t$$

- $e_t$  stationary.
- $\nabla x_t = \gamma_t + \nabla y_t$ : not stationary.
- $\nabla^2 x_t = e_t + \nabla^2 y_t$ : stationary.

Such process is called integrated ARMA model or ARIMA model.

Definition 3.11 A process  $x_t$  is said to be ARIMA(p, d, q) if

- $\nabla^d x_t = (1 B)^d x_t$  is ARMA(p, q).
- In general, it can be written as  $\phi(B)(1-B)^d x_t = \theta(B) w_t$ .

## Integrated models - Unit Root test

Unit Root test.

## Integrated models - Cointegration

What is cointegration? How to use it in practice?

Trading strategy.
 https://quant.stackexchange.com/questions/3270/what-are-the-applications-of-cointegration

- Dependence occures most strongly at multiples of seasonal lag s.
- For example, monthly economic data, s=12.
- Quarterly data, s = 4.

Pure seasonal autoregressive moving average model: ARMA $(P,Q)_s$ 

$$\Phi_P(B^s)x_t = \Theta_Q(B^s)w_t$$

- Seasonal AR operator:  $\Phi_P(B^s) = 1 \Phi_1 B^s \Phi_2 B^{2s} \ldots \Phi_P B^{Ps}$ .
- Seasonal MA operator:  $\Theta_Q(B^s) = 1 + \Theta_1 B^s + \Theta_2 B^{2s} + \ldots + \Theta_Q B^{Qs}$ .

Example 3.46 Seasonal AR(1) series

$$(1 - \Phi B^{12})x_t = w_t$$

- $x_t = \Phi x_{t-12} + w_t$
- Estimation and forcasting is similar to AR(1) model.
- $|\Phi| < 1$

Simulated 3 years of data with  $\Phi = 0.9$ :

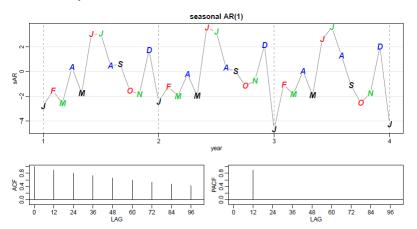


Fig. 3.20. Data generated from a seasonal (s = 12) AR(1), and the true ACF and PACF of the model  $x_t = .9x_{t-12} + w_t$ .

First order seasonal (d=12) MA model

$$x_t = w_t + \Theta w_{t-12}$$

- $\gamma(0) = (1 + \Theta^2)\sigma^2$
- $\gamma(\pm 12) = \Theta \sigma^2$
- $\gamma(h) = 0$  otherwise.
- The only non-zero correlation is  $\rho(\pm 12) = \Theta/(1+\Theta^2)$ .

First order seasonal (d = 12) AR model

$$x_t = \Phi x_{t-12} + w_t$$

- $\gamma(0) = \sigma^2/(1 \Phi^2)$
- $\gamma(\pm 12k) = \sigma^2 \Phi^k / (1 \Phi^2)$
- $\gamma(h) = 0$  otherwise.
- The only non-zero correlations are  $\rho(\pm 12k) = \Theta^k$ ,  $k = 0, 1, 2, \ldots$

Table 3.3. Behavior of the ACF and PACF for Pure SARMA Models

	$AR(P)_s$	$MA(Q)_s$	$ARMA(P,Q)_s$
ACF*	Tails off at lags $ks$ , $k = 1, 2, \dots$ ,	Cuts off after $\log Qs$	Tails off at lags $ks$
PACF*	Cuts off after $\log Ps$	Tails off at lags $ks$ $k = 1, 2, \dots,$	Tails off at lags $ks$

<sup>\*</sup>The values at nonseasonal lags  $h \neq ks$ , for  $k = 1, 2, \ldots$ , are zero.

General multiplicative seasonal ARMA, ARMA $(p,q) \times (P,Q)_s$ 

$$\Phi_P(B^s)\phi(B)x_t = \Theta_Q(B^s)\theta(B)w_t$$

Example 3.47: A mixed seasonal model

ARMA $(0,1) \times (1,0)_{12}$  model:

$$x_t = \Phi x_{t-12} + w_t + \theta w_{t-1}$$

 $ARMA(0,1) \times (1,0)_{12}$  model:

$$x_t = \Phi x_{t-12} + w_t + \theta w_{t-1}$$

ullet Take the variance of both side:  $\gamma(0)=\Phi^2\gamma(0)+\sigma_w^2+ heta^2\sigma_w^2$ 

$$\gamma(0) = \frac{1+\theta^2}{1-\Phi^2}\sigma_w^2$$

- Multiply both side by  $x_{t-1}$  and take exp:  $\gamma(1) = \Phi \gamma(11) + \theta \sigma_w^2$ .
- Multiply both side by  $x_{t-h}$   $(h \ge 2)$  and take exp:  $\gamma(h) = \Phi \gamma(h-12)$ .
- $\rho(12h) = \Phi^h$
- $\rho(12h-1) = \rho(12h+1) = \frac{\theta}{1+\theta^2}\Phi^h$ .
- $\rho(h) = 0$  otherwise.

ARMA $(0,1) \times (1,0)_{12}$  model:

$$x_t = \Phi x_{t-12} + w_t + \theta w_{t-1}$$

• Setting:  $\Phi = 0.8$  and  $\theta = -0.5$ .

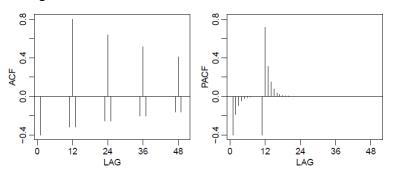


Fig. 3.21. ACF and PACF of the mixed seasonal ARMA model  $x_t = .8x_{t-12} + w_t - .5w_{t-1}$ .

#### Seasonal persistence: when the process is nearly periodic.

- For example, the average monthly temperature over the years.
- Each January would be approximately the same. Each Febuary ...
- Model average monthly temperature  $(x_t)$ :

$$x_t = S_t + w_t, w_t \sim \text{ white noise.}$$

•  $S_t$  the seasonal component, varies a little:

$$S_t = S_{t-12} + \nu_t, \nu_t \sim \text{ white noise.}$$

• ACF decays very slowly at lags h = 12k.

Seasonal persistence: when the process is nearly periodic.

•  $S_t$  the seasonal component, varies a little:

$$S_t = S_{t-12} + \nu_t, \nu_t \sim \text{ white noise.}$$

• Subtract the effect of successive years from each other (a seasonal difference of order 1):

$$(1 - B^{12})x_t = x_t - x_{t-12} = \nu + w_t - w_{t-12}$$

•  $MA(1)_{12}$ , ACF has a peak only at lag 12.

<u>Definition 3.12</u> Multiplicative seasonal autoregressive integrated moving average (SARIMA)

$$\Phi_P(B^s)\phi(B)\nabla_s^D\nabla^d x_t = \delta + \Theta_Q(B^s)\theta(B)w_t$$

• ARIMA $(p, d, q) \times (P, D, Q)_s$ .

#### Example 3.49 Air passengers.

• Monthly totals of international airline passengers, 1949 to 1960.

#### Case Study: Walmart Sales Prediction

Walmart recruiting competition:

• https://www.kaggle.com/c/walmart-recruiting-store-sales-forecasting https://rpubs.com/spillai/walmart store sales forecast

$$y_t = \sum_{j=1}^r \beta_j z_{tj} + x_t$$

• If  $x_t$  is white Gaussian noise such that

$$\gamma_{\mathsf{x}}(\mathsf{s},t) = \mathsf{0} \,\, \mathsf{for} \,\, \mathsf{s} 
eq t \,\, \mathsf{and} \,\, \gamma_{\mathsf{x}}(t,t) = \sigma^2$$

OLS will provide the efficient estimators.

• Otherwise, weighted least squares.

$$\underline{y} = Z\underline{\beta} + \underline{x}$$

•  $\underline{x} \sim \mathcal{N}(0, \Gamma)$ ,  $\Gamma = \{\gamma_x(s, t)\}$ .

$$\Gamma^{-1/2}\underline{y} = \Gamma^{-1/2}Z\underline{\beta} + \Gamma^{-1/2}\underline{x}$$
$$\underline{y}^* = Z^*\underline{\beta} + \underline{\delta}$$

- $\underline{\delta} = \Gamma^{-1/2} \underline{x} \sim \mathcal{N}(0, I)$ .
- The weighted least square estimator is

$$\hat{\underline{\beta}}_{w} = ((Z^{*})'Z^{*})^{-1}(Z^{*})'\underline{y}* = (Z'\Gamma^{-1}Z)^{-1}(Z'\Gamma^{-1}\underline{y})$$

In the time series setting,

$$y_t = \sum_{j=1}^r \beta_j z_{tj} + x_t$$

• If we know the model for  $x_t$ :

$$\phi(B)x_t = w_t, \ (\phi(B) = 1 - \phi_1 B - \dots - \phi_p B^p)$$

$$\Rightarrow \phi(B)y_t = \sum_{j=1}^r \beta_j \phi(B) z_{tj} + \phi(B) x_t$$

$$y_t^* = \sum_{j=1}^r \beta_j z_{tj}^* + w_t$$

• Least square estimator minimizes:

$$S(\phi, B) = \sum_{t=1}^{n} w_t^2 = \sum_{t=1}^{n} [\phi(B)y_t - \sum_{j=1}^{r} \beta_j \phi(B)z_{tj}]^2$$

#### The general process:

- 1. First run an ordinary regression (as if the errors are uncorrelated.) Obtain the residuals,  $\hat{x}_t = y_t \sum_{i=1}^r \hat{\beta}_i z_{tj}$ .
- 2. Identify the ARMA model for the residuals  $\hat{x}_t$ .
- 3. Run weighted least squares (or MLE) on the regression model with autocorrelated error structure.
- 4. Model diagnostics.

Example 3.45 Regression with two time series.

$$R_t = \beta_0 + \beta_1 S_{t-6} + \beta_2 D_{t-6} + \beta_3 D_{t-6} S_{t-6} + w_t$$

- R<sub>t</sub> recruitment series.
- $S_t$  the SOI index series.
- $\bullet$   $D_t$  adummy variable.

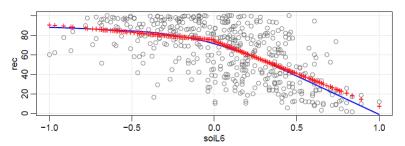


Fig. 2.10. Display for Example 2.9: Plot of Recruitment  $(R_t)$  vs SOI lagged 6 months  $(S_{t-6})$  with the fitted values of the regression as points (+) and a lowess fit (-).

<u>Example</u> Stock Prediction using dividend. https://www.econometrics-with-r.org/14-9-can-you-beat-the-market-part-ii.html

## Reading Materials

- Ch 3.6 (Integrated model)
- Ch 3.9 (SARIMA)
- Ch 3.8 p.142-145 (Regression with Autocorrelated Errors)
- https://medium.com/@arneeshaima/walmart-sales-data-analysis-sales-prediction-using-multiple-linear-regression-in-r-programming-adb14afd56fb