

MSA 8200 Predictive Analytics

Week 4: ARMA model and alternatives

Yichen Cheng

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- ARMA model
- Forecasting
- Estimation
- Model Diagnostics
- Alternatives?

ARMA model – Definition

A time series $\{x_t; t = 0, \pm 1, \dots\}$ is ARMA(p, q) if:

$$x_t = \phi_1 x_{t-1} + \dots + \phi_p x_{t-p} + w_t + \theta_1 w_{t-1} + \dots + \theta_q w_{t-q}$$

- $\phi_p \neq 0, \theta_q \neq 0, \sigma_w^2 > 0$.
- Stationary.
- p is the AR order and q is the MA order.

ARMA model – Definition

Definition: AR and MA polynomials:

- **AR polynomials:** $\Phi(z) = 1 - \phi_1 z - \dots - \phi_p z^p$, $\phi_p \neq 0$.
- **MA polynomials:** $\Theta(z) = 1 + \theta_1 z + \dots + \theta_q z^q$, $\theta_q \neq 0$.
- z is a complex number.

ARMA model – Redundancy

Parameter redundancy:

$$\begin{aligned}\Phi(B)x_t &= \Theta(B)w_t \\ \Rightarrow \eta(B)\Phi(B)x_t &= \eta(B)\Theta(B)w_t\end{aligned}$$

Example:

- True model: $x_t = w_t$.
- Let $\eta(B) = 1 - 0.5B$, then we have

$$\begin{aligned}(1 - 0.5B)x_t &= (1 - 0.5B)w_t \\ \Rightarrow x_t &= 0.5x_{t-1} - 0.5w_{t-1} + w_t\end{aligned}$$

- It looks like an ARMA model.

ARMA model – Redundancy

We use $\text{ARMA}(p,q)$ model to refer to its simplest form.

- That is $\Phi(z)$ and $\theta(z)$ have no common factors.

$$x_t = 0.5x_{t-1} - 0.5w_{t-1} + w_t.$$

- $\Phi(z) = 1 - 0.5z$.
- $\Theta(z) = 1 - 0.5z$.
- $\Phi(z)$ and $\Theta(z)$ have common factor.
- $\Phi(z) = 1$ and $\Theta(z) = 1$.
- It is a $\text{ARMA}(0,0)$ process.

Definition 3.7: Concept of causality

- An ARMA(p, q) model for x_t is said to be causal if x_t can be written as a one-sided linear process:

$$x_t = \sum_{j=0}^{\infty} \psi_j w_{t-j} = \Psi(B)w_t$$

- where $\Psi(B) = \sum_{j=0}^{\infty} \psi_j B^j$ and $\sum_{j=0}^{\infty} |\psi_j| < \infty$.
- $\psi_0 = 1$.
- It is also called the MA representation.

An ARMA(p, q) model is causal if and only if:

- $\Phi(z) \neq 0$ for $|z| \leq 1$.
- Then $x_t = \sum_{j=0}^{\infty} \psi_j w_{t-j}$.
- $\Psi(z) = \sum_{j=0}^{\infty} \psi_j z^j = \Theta(z)/\Phi(z)$, $|z| \leq 1$.
- Equivalently $\Phi(z) = 0$ only when $|z| > 1$.

Definition 3.8: An ARMA(p,q) model is invertible, if the time series $\{x_t : t = 0, \pm 1, \pm 2, \dots\}$ can be written as

$$\Pi(B)x_t = \sum_{j=0}^{\infty} \pi_j x_{t-j} = w_t$$

- $\Pi(B) = \sum_{j=0}^{\infty} \pi_j B^j$;
- $\sum_{j=0}^{\infty} |\pi_j| < \infty$;
- $\pi_0 = 1$;
- It is called the **invertible representation**.

Prop 3.2: Invertibility of an ARMA(p,q) process.

An ARMA(p,q) model is invertible if and only if

- $\Theta(z) \neq 0$ for $|z| \leq 1$.
- Then, $\pi(z) = \sum_{j=0}^{\infty} \pi_j z^j = \Phi(z)/\Theta(z)$, $|z| \leq 1$.
- In other words, roots of $\Theta(z)$ lie outside the unit circle.
- $\Theta(z) = 0$ only when $|z| > 1$.

ARMA model – Example

Example 3.7: Parameter redundancy, causality and invertibility

$$x_t = 0.4x_{t-1} + 0.45x_{t-2} + w_t + w_{t-1} + 0.25w_{t-2}$$

- $(1 - 0.4B - 0.45B^2)x_t = (1 + B + 0.25B^2)w_t$.
- It appears to be an ARMA(2,2) process.
- $\Phi(B) = (1 - 0.4B - 0.45B^2) = (1 - 0.9B)(1 + 0.5B)$.
- $\Theta(B) = (1 + B + 0.25B^2) = (1 + 0.5B)^2$.
- Common factor: $(1 + 0.5B)$. ARMA(1,1)!

Exemple 3.7 continued:

$$(1 - 0.9B)x_t = (1 + 0.5B)w_t$$

- $x_t = 0.9x_t + 0.5w_{t-1} + w_t$.
- Causal: $\Phi(z) = (1 - 0.9z) = 0 \Rightarrow z = 10/9$.
- Invertible: $\Theta(z) = (1 + 0.5z) = 0 \Rightarrow z = -2$.

ARMA model – Example

Exemple 3.7 continued:

$$(1 - 0.9B)x_t = (1 + 0.5B)w_t$$

- The MA representation:

$$x_t = (1 + \psi_1 B + \psi_2 B^2 + \dots)w_t.$$

- $(1 - 0.9z)(1 + \psi_1 z + \psi_2 z^2 + \dots) = 1 + 0.5z$
- $1 + (\psi_1 - 0.9)z + (\psi_2 - 0.9\psi_1)z^2 + \dots + (\psi_j - 0.9\psi_{j-1})z^j + \dots = 1 + 0.5z.$
- $\psi_1 - 0.9 = 0.5 \Rightarrow \psi_1 = 1.4.$
- $\psi_j - 0.9\psi_{j-1} = 0 \Rightarrow \psi_j = 1.4 \times 0.9^{(j-1)}$

How about the invertible representation?

Goal for forecasting:

- Use current and past observations to predict future.

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- Use current and past observations to predict future.
- Predict x_{n+1} based on $X_{1:n} = \{x_1, x_2, \dots, x_n\}$.
- Predict x_{n+2} based on $X_{1:n} = \{x_1, x_2, \dots, x_n\}$.
- ...

Goal: predict x_{n+m} , $m = 1, 2, \dots$ based on $X_{1:n} = \{x_1, x_2, \dots, x_n\}$

$$x_{n+m}^n = \alpha_0 + \sum_{k=1}^n \alpha_k x_k$$

- x_{n+m}^n minimize $E[x_{n+m} - x_{n+m}^n]^2$,
- It is called the best linear predictor. (BLP)

Property 3.3: Best linear prediction for stationary process

- The parameters $\alpha_0, \alpha_1, \dots, \alpha_n$ in $x_{n+m}^n = \alpha_0 + \sum_{k=1}^n \alpha_k x_k$ is found by solving

$$E[(x_{n+m} - x_{n+m}^n)x_k] = 0, k = 0, 1, \dots, n,$$

for $\alpha_0, \alpha_1, \dots, \alpha_n$, where $x_0 = 1$.

One-step-ahead prediction:

$$x_{n+1}^n = \phi_{n1}x_n + \phi_{n2}x_{n-1} + \dots + \phi_{nn}x_1 = \underline{\phi}_n' \underline{x} \quad (1)$$

- $E[(x_{n+1} - \sum_{j=1}^n \phi_{nj}x_{n+1-j})x_{n+1-k}] = 0, k = 1, \dots, n.$
 $\Rightarrow \sum_{j=1}^n \phi_{nj}\gamma(k-j) = \gamma(k)$ or $\Gamma_n \underline{\phi}_n = \underline{\gamma}_n.$

$$\underline{\phi}_n = \Gamma_n^{-1} \underline{\gamma}_n. \quad (2)$$

- $\Gamma_n = \{\gamma(k-j)\}, j = 1, \dots, n, k = 1, \dots, n$
- $\underline{\gamma}_n = (\gamma(1), \dots, \gamma(n))'$

Mean squared one-step-ahead prediction error:

$$P_{n+1}^n = E(x_{n+1} - x_{n+1}^n)^2 = \gamma(0) - \underline{\gamma}_n' \Gamma_n^{-1} \underline{\gamma}_n.$$

- $\Gamma_n = \{\gamma(k-j)\}, j = 1, \dots, n, k = 1, \dots, n$
- $\underline{\gamma}_n = (\gamma(1), \dots, \gamma(n))'$

Example Prediction for an AR(1) model

$$x_t = 0.5x_{t-1} + w_t$$

One-step-ahead prediction of x_2 based on x_1 is:

- $x_2^1 = \phi_{11}x_1$
- $\phi_{11} = \Gamma_1^{-1}\gamma_1$
- $\Gamma_1 = \gamma(0) = \sigma_w^2/(1 - \phi^2)$
- $\gamma_1 = \gamma(1) = \sigma_w^2\phi/(1 - \phi^2)$

$$\Rightarrow x_2^1 = 0.5x_1$$

- The mean squared prediction error (MSPE):

$$P_2^1 = \gamma(0) - \gamma_n' \Gamma_n^{-1} \gamma_n = \gamma(0) - \gamma_1^2 / \gamma(0) = \sigma_w^2$$

Example 3.19 Prediction for an AR(2) model

$$x_t = \phi_1 x_{t-1} + \phi_2 x_{t-2} + w_t$$

One-step-ahead prediction of x_2 based on x_1 :

$$x_2^1 = \phi_{11} x_1 = \frac{\gamma(1)}{\gamma(0)} x_1 = \rho(1) x_1$$

One-step-ahead prediction of x_3 based on x_2 is:

$$x_3^2 = \phi_{21} x_2 + \phi_{22} x_1$$

- $\phi_{21} = \phi_1, \phi_{22} = \phi_2$.

Forecasting – Durbin-Levinson Algorithm

- For general ARMA models, the prediction equations not as simple.
- Also, for n large, using Equation (2) is hard (matrix inverse is expensive).

Solve the prediction iteratively through **Durbin-Levinson Algorithm**:

- $\phi_{00} = 0, P_1^0 = \gamma(0)$
- for $n \geq 1$,
 - $\phi_{nn} = \frac{\rho(n) - \sum_{k=1}^{n-1} \phi_{n-1,k} \rho(n-k)}{1 - \sum_{k=1}^{n-1} \phi_{n-1,k} \rho(k)}$
 - $P_{n+1}^n = P_n^{n-1} (1 - \phi_{nn}^2)$
- for $n \geq 2$: $\phi_{nk} = \phi_{n-1,k} - \phi_{nn} \phi_{n-1,n-k}, k = 1, 2, \dots, n-1$

Note: the Durbin-Levinson Algorithm can be used to find the PACF.
(Example 3.21 in textbook.)

Example 3.25 Forecasting the Recruitment Series.

- Estimation
- Forecasting
- Mean square prediction error (MSPE)

Example 3.29 Method of moments estimation for an MA(1)

$$x_t = w_t + \theta w_{t-1}$$

with $|\theta| < 1$.

Use the same idea as for Yule-Walker Equation:

- $\gamma(0) = \sigma_w^2(1 + \theta^2)$
- $\gamma(1) = \sigma_w^2\theta$

$\hat{\theta}$ can be found by solving

$$\hat{\rho}(1) = \frac{\hat{\gamma}(1)}{\hat{\gamma}(0)} = \frac{\hat{\theta}}{1 + \hat{\theta}^2} \quad (3)$$

Example 3.29 Method of moments estimation for an MA(1)

$$x_t = w_t + \theta w_{t-1}$$

with $|\theta| < 1$.

$\hat{\theta}$ can be found by solving

$$\hat{\rho}(1) = \frac{\hat{\gamma}(0)}{\hat{\gamma}(0)} = \frac{\hat{\theta}}{1 + \hat{\theta}^2} \quad (4)$$

- Two solutions exist, pick the invertible one.
- If $|\hat{\rho}(1)| \leq 1/2$, the solutions are real. Otherwise, no real solution exists.
- Example: $\hat{\rho}(1) = .507$ while $\rho(1) = .9/(1 + .9^2) = .497$.

Estimation – Method of moments

Example 3.29 Method of moments estimation for an MA(1)

- When $|\hat{\rho}(1)| \leq 1/2$, the invertible estimate is

$$\hat{\theta} = \frac{1 - \sqrt{4\hat{\rho}(1)^2}}{2\hat{\rho}(1)}$$

Furthermore, it can be shown that

$$\hat{\theta} \sim AN\left(\theta, \frac{1 + \theta^2 + 4\theta^4 + \theta^6 + \theta^8}{n(1 - \theta^2)^2}\right)$$

AN means asymptotically normal.

- In contrast, the maximum likelihood estimator has an asymptotic variance $(1 - \theta^2)/n$.
- When $\theta = 0.5$, the ratio is about 3.5.

Estimation – Maximum Likelihood Estimation

AR(1): $x_t = \phi x_{t-1} + w_t$ with $|\phi| < 1$ and $w_t \sim iid \mathcal{N}(0, \sigma_w^2)$

Likelihood:

$$L(\phi, \sigma_w^2) = f(x_1)f(x_2|x_1) \dots f(x_n|x_{n-1})$$

- $x_t|x_{t-1} \sim \mathcal{N}(\phi x_{t-1}, \sigma_w^2)$
 $f(x_t|x_{t-1}) = f_w(x_t - \phi x_{t-1})$, f_w is the density of w_t
- $x_1 = w_1 + \phi w_0 + \phi^2 w_{-1} + \dots = \sum_{j=0}^{\infty} \phi^j w_{1-j}$
 $x_1 \sim \mathcal{N}(0, \sigma_w^2/(1 - \phi^2))$

Likelihood:

$$L(\phi, \sigma_w^2) = (2\phi\sigma_w^2)^{-n/2}(1 - \phi^2)^{1/2} \exp \left\{ -\frac{S(\phi)}{2\sigma_w^2} \right\}$$

- $S(\phi) = (1 - \phi^2)x_1^2 + \sum_{t=2}^n (x_t - \phi x_{t-1})^2$

Estimation – Maximum Likelihood Estimation

For a general ARMA model:

- It is difficult to write the likelihood as an explicit function of the parameters.
- Instead, it is advantageous to write it in terms of the **innovations**, one-step-ahead prediction errors $x_t - x_t^{t-1}$

Let $\beta = (\phi_1, \dots, \phi_p, \theta_1, \dots, \theta_q)$:

$$L(\beta, \sigma_w^2) = \prod_{t=1}^n f(x_t | x_{t-1}, \dots, x_1)$$

We have $x_t | x_{t-1}, \dots, x_1 \sim \mathcal{N}(x_t^{t-1}, P_t^{t-1})$, where

- $P_t^{t-1} = \gamma(0) \prod_{j=1}^{t-1} (1 - \phi_{jj}^2)$
- $\gamma(0) = \sigma_w^2 \prod_{j=0}^{\infty} \psi_j^2$

Estimation – Maximum Likelihood Estimation

The Newton-Raphson algorithm:

- Start with an initial guess β^0 .
- At iteration t : update the estimate by':

$$\beta^t = \beta^{t-1} + \{l''(\beta^{t-1})\}^{-1} l'(\beta^{t-1}) \quad (5)$$

- Stop until convergence.

Large Sample Property:

$$\sqrt{n}(\hat{\beta} - \beta) \xrightarrow{d} \mathcal{N}(0, \sigma_w^2 \Gamma_{p,q}^{-1}) \quad (6)$$

Estimation – Overfitting Caveat

Example 3.35 Overfitting Caveat.

Does it matter?

- If the time series follows an AR(1) process.
- We decide to fit an AR(2) model to the data.

By overfitting:

- We lose efficiency:
- AR(1): $\text{var}(\hat{\phi}_1) \approx (1 - \phi_1^2)/n$
- fit an AR(2): $\text{var}(\hat{\phi}_1) \approx (1 - \phi_2^2)/n = 1/n > (1 - \phi_1^2)/n$

Side note: it can be used as a diagnostic tool.

Example 3.33: Fitting the glacial varve series

- Paleoclimatic glacial varve thicknesses from Massachusetts for $n = 634$ years.
- Melting glaciers deposit yearly layers of sand and silt during the spring melting seasons.
- Beginning 11,834 years ago.

To do:

- Estimation.

Looking at the residual:

- The residuals should be i.i.d $\mathcal{N}(0, \sigma^2)$

This suggests we should check whether the residuals:

- follow a normal distribution. [Q-Q plot](#)
- are independent: no auto correlation. [Ljung-Box-Pierce Q-statistics](#)

Standardized Residuals::

$$e_t = (x_t - x_t^{t-1}) / \sqrt{P_t^{t-1}}$$

Ljung-Box-Pierce Q-statistics:

$$Q = n(n+2) \sum_{h=1}^H \frac{\hat{\rho}_e^2(h)}{n-h}$$

- $\hat{\rho}_e^2(h)$: lag h auto correlation for $\{e_1, \dots, e_n\}$
- $Q \sim \chi_{H-p-q}^2$.

Model diagnostics – Examples

Example 3.40: Diagnostics for GNP growth rate example

Example 3.41: Diagnostics for the glacial varve series

Basic steps for fitting an ARMA model

Basic steps for fitting ARMA models to time series data:

- plotting the data
- possibly transforming the data
- identifying the dependence orders of the model.
- parameter estimation
- diagnostics
- model choice

Go back to the stock prediction example

Alternatives? <https://www.analyticsvidhya.com/blog/2018/10/predicting-stock-price-machine-learningnd-deep-learning-techniques-python/>

- Moving Average
- Linear Regression
- k-Nearest Neighbors (KNN)
- Auto ARIMA
- Prophet
- Long Short Term Memory (LSTM)

What was done right/wrong?

- Ch 3.3 p.83-88 (ARMA model, Causality, Invertibility)
- Ch 3.4 (Forecasting)
- Ch 3.5 (Estimation)
- Ch 3.7 p.135-142 (Building ARIMA models)
- A short YouTube video on Ljung Box test:
<https://www.youtube.com/watch?v=FETjCDysdWY>