

PREDICTIVE ANALYTICS

HW2

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02/15/2020

Question 1

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Time Series Analysis HW2

- 1) Let $w_t: t = 0, 1, \dots$ be a white noise process with variance σ_w^2 and let $|\phi| < 1$ be a constant. Consider the process $x_0 = w_0$ and

$$x_t = \phi x_{t-1} + w_t, \quad t = 1, 2, \dots$$

- a) Show that $x_t = \sum_{j=0}^{t-1} \phi^j w_{t-j}$ for any $t = 1$.

$$x_t = \phi (\phi x_{t-2} + w_{t-1}) + w_t$$

$$x_t = \phi^2 x_{t-2} + \phi w_{t-1} + w_t$$

$$x_t = \phi^k x_{t-k} + \sum_{j=0}^{k-1} \phi^j w_{t-j}$$

assuming $|\phi| < 1$ as $j \rightarrow \infty$ we can represent the AR(1) process as a linear one given by

$$\sum_{j=0}^{\infty} \phi^j w_{t-j}$$

- b) Find $E(x_t)$

$$E(x_t) = \sum_{j=0}^{\infty} \phi^j E(w_{t-j}) = 0$$

as $j \rightarrow \infty$ the value of $x_t \rightarrow w_t$ and $E(w_t) = 0$.

c) Show that for $t = 0, 1, \dots$

$$\text{var}(x_t) = \frac{\sigma_w^2 (1 - \phi^{2(t+1)})}{1 - \phi^2}$$

$$x_t = \phi x_{t-1} + w_t$$

$$x_t = \sum_{j=0}^{\infty} \phi^j w_{t-j}$$

$$\text{var}(x_t) = \sum_{j=0}^{\infty} \phi^{2j} \cdot \text{var}(w_{t-j})$$

$$\downarrow$$

$$= \sigma^2$$

assumes $a \perp b$ so we can just add

assumes $w_t \sim N(0, \sigma^2)$

$$\text{var}(x_t) = \sum_{j=0}^{\infty} \phi^{2j} \cdot \sigma^2$$

product across an infinite series.

$$\text{var}(x_t) = \sigma^2 (1 + \phi^2 + \phi^4 + \phi^6 + \dots)$$

$$= \frac{\sigma_w^2}{1 - \phi^2}$$

d.) Show that for $h \geq 0$ $\text{cov}(x_t, x_{t+h}) = \phi^h \text{var}(x_t)$

Question: why do we use expectation w/ covariance? because we est m?

$$E [E(x_t - \hat{x}) \cdot E(x_{t-1} - \hat{x}_{t-1})]$$

TBD

HW 2. Question 1

e.) Is x_t stationary?

x_t stationary | $|\phi| < 1$.

$$AR(1) = x_t = \phi x_{t-1} + w_t$$

$$\text{if } \phi = 1 \Rightarrow x_t = 1 + x_{t-1} + E(w_t)$$

$$x_t = x_{t-1} + 0$$

$\therefore x_t$ is only dependent on the prior observation & w/
assume a random walk.

Furthermore, $\text{Var}(x_t) = t \times \sigma^2_w$ meaning that variance
will increase unchecked as $t \rightarrow \infty$

if $\phi > 1$ then the past observations of x_t grow as t increases, meaning
 x is progressively more dep on older observations. Logically, this does not
make sense, and as a geometric series does not allow for a finite sum.

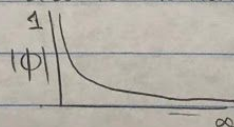
x_t is stationary | $|\phi| < 1$

Allows for an infinite geometric series.

As $t \rightarrow \infty$ $|\phi| \rightarrow 0$.

f.) Argue that as $t \rightarrow \infty$ the process becomes stationary.

• $|\phi| < 1$ leads to a decaying function



• As $t \rightarrow \infty$ $|\phi|$ will approach but

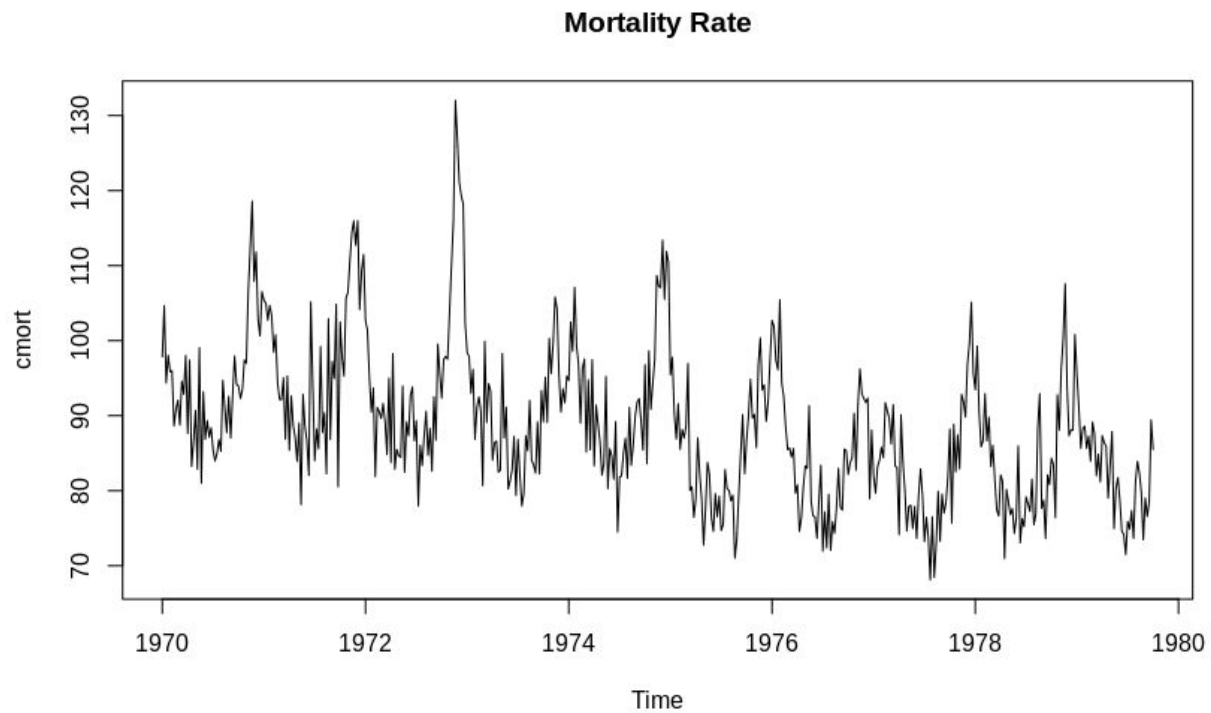
never reach zero. Therefore, past observations of x_t w/ decrease in importance.

and given $x_t = \sum_{i=0}^{\infty} \phi^i w_{t-i}$; $x_t \approx w$ which we know is stationary
when $w = N(0, \sigma^2)$ is iid.

Question 2

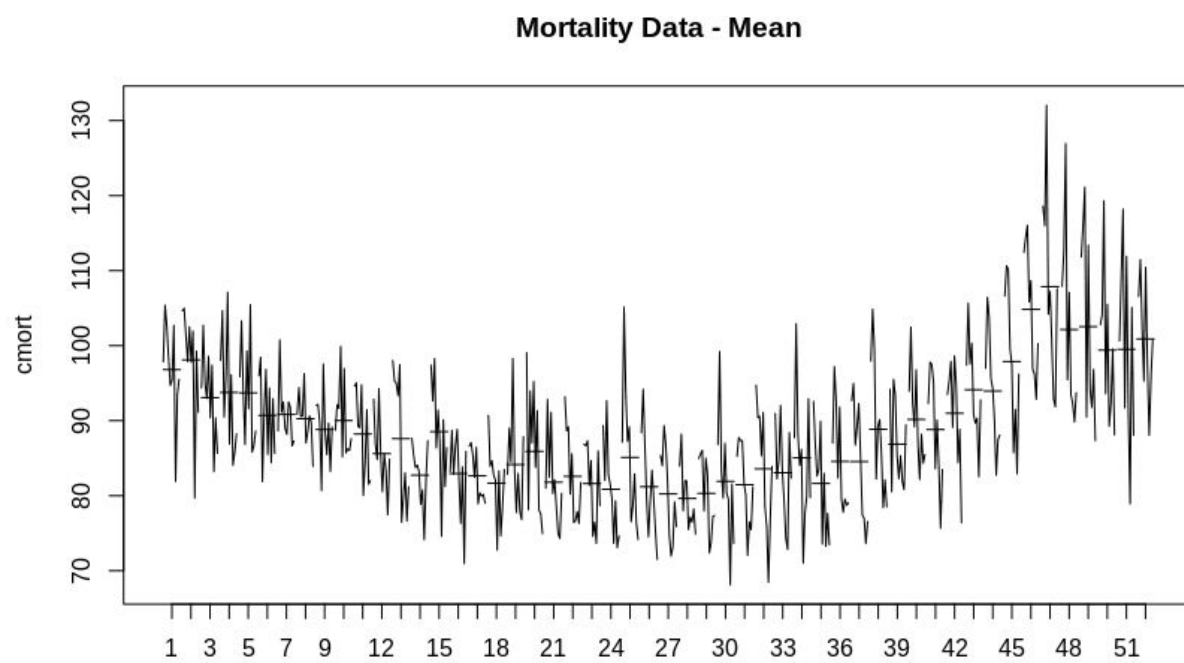
a.) Plot the data and ACF, PACF. Comment on the stationarity and discuss which model you would use.

```
plot(cmort)
```

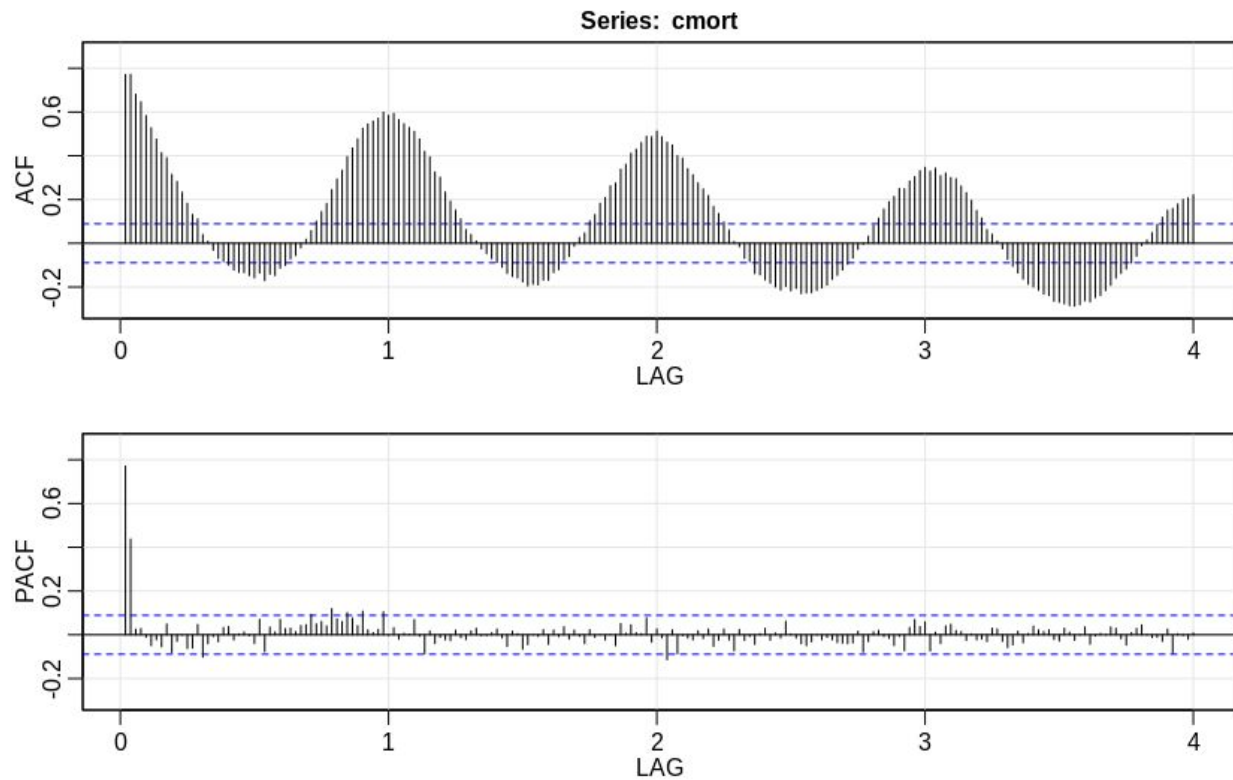


The data does not appear to be stationary. The mean appears to decline over time "trend" and there appears to be yearly seasonality.

```
cmort1 = cmort[0: 100]
cmort2 = cmort[408 : 508]
cmort1_mu = mean(cmort1)
cmort2_mu = mean(cmort2)
mu1_2 = c(cmort1_mu, cmort2_mu)
barplot(mu1_2, main='Barplot Mu1 and Mu2')
monthplot(cmort, main='Mortality Data - Mean')
```



```
'Plot ACF & PACF'  
par(mfrow=c(2,1))  
acf(cmort)  
pacf(cmort)
```



Observations:

- The ACF and PACF plots indicate that this is an AR model with some seasonality at increments of 1. In the case of the ACF, the phi value progressively decreases from 1 as the lag increases. If this were an MA model we would see a single bar and then nothing. In the case of the PACF plot, we see two bars and then minimal ones thereafter. If this were an MA model we would see progressively declining bars as is the case with the ACF plot.

Question 2(b)

Fit an AR(2) model to xt using the method of moments Yule Walker equation

```
cmort.diff <- diff(log(cmort))  
ar.yw <- ar.yw(cmort.diff, order=2)  
ar.yw
```

Call:

```
ar.yw.default(x = cmort.diff, order.max = 2)
```

Coefficients:

```
      1      2  
-0.5678 -0.0945
```

Order selected 2 sigma^2 estimated as 0.004183

Matrix

```
      [,1] [,2]  
[1,] 0.001966395 0.001020007  
[2,] 0.001020007 0.001966395
```

```
out = acf(cmort.diff)  
rho = out$acf[1:5]  
G2 = matrix(c(rho[1],rho[2],rho[2],rho[1]),2)  
g2 = c(rho[2],rho[3])  
solve(G2)%*%g2  
[1,] -0.56775711  
[2,] -0.09453654
```


Question 3

a.) $x_t = 0.8x_{t-1} - 0.15x_{t-2} + w_t - 0.3w_{t-1}$

$$x_t - 0.8x_{t-1} + 0.15x_{t-2} = w_t - 0.3w_{t-1}$$

$$(1 - 0.8B)x_t + (1 - 0.15B^2)x_t = (1 - 0.3B)w_t$$

$$(1 - 0.8B + 0.15B^2)x_t = (1 - 0.3B)w_t$$

ARMA(2,1)

$$\phi(B) = (1 - 0.8B + 0.15B^2) = (1 - 0.5B)(1 - 0.3B)$$

$$\begin{aligned} 1 - 0.3B - 0.5B + 0.15B^2 \\ 1 - 0.8B + 0.15B^2 \end{aligned}$$

• common factor = $(1 - 0.3B)$

• Therefore ARMA(1,1) model.

b.) $x_t = x_{t-1} - 0.5x_{t-2} + w_t - w_{t-1}$

$$x_t - x_{t-1} + 0.5x_{t-2} = w_t - w_{t-1}$$

$$(1 - B + 0.5B^2)x_t = (1 - B)w_t$$

• common factor = $(1 - B)$

• Therefore AR(1,1) model.

Causality: $\phi(z) \neq 0$ for $|z| \leq 1$

$\phi \neq 0 \neq 0$

Both are causal & invertible

Question 4

```
# Clear Namespace
```

```
rm(list=ls())
```

```
dev.off(dev.list())
```

```
# Load Libraries
```

```
library(readxl)
```

```
library(plyr)
```

```
library(ggplot2)
```

```
library(astsa)
```

```
library(tseries)
```

```
## DATA PREPARATION -----
```

```
# Load Data & Get Monthly Sales
```

```
sales <- read_excel("Desktop/repositories/Time_Series/hw2/sales.xls")
```

```
sales$yr <- format(sales$`Order Date`, format='%Y')
```

```
sales$yr_month <- format(sales$`Order Date`, format='%Y-%m')
```

```
sales[c('Sales', 'yr_month')]
```

```
df.sum <- ddply(sales, c('yr_month'), summarize, Sales = sum(Sales))
```

```
> df.sum
```

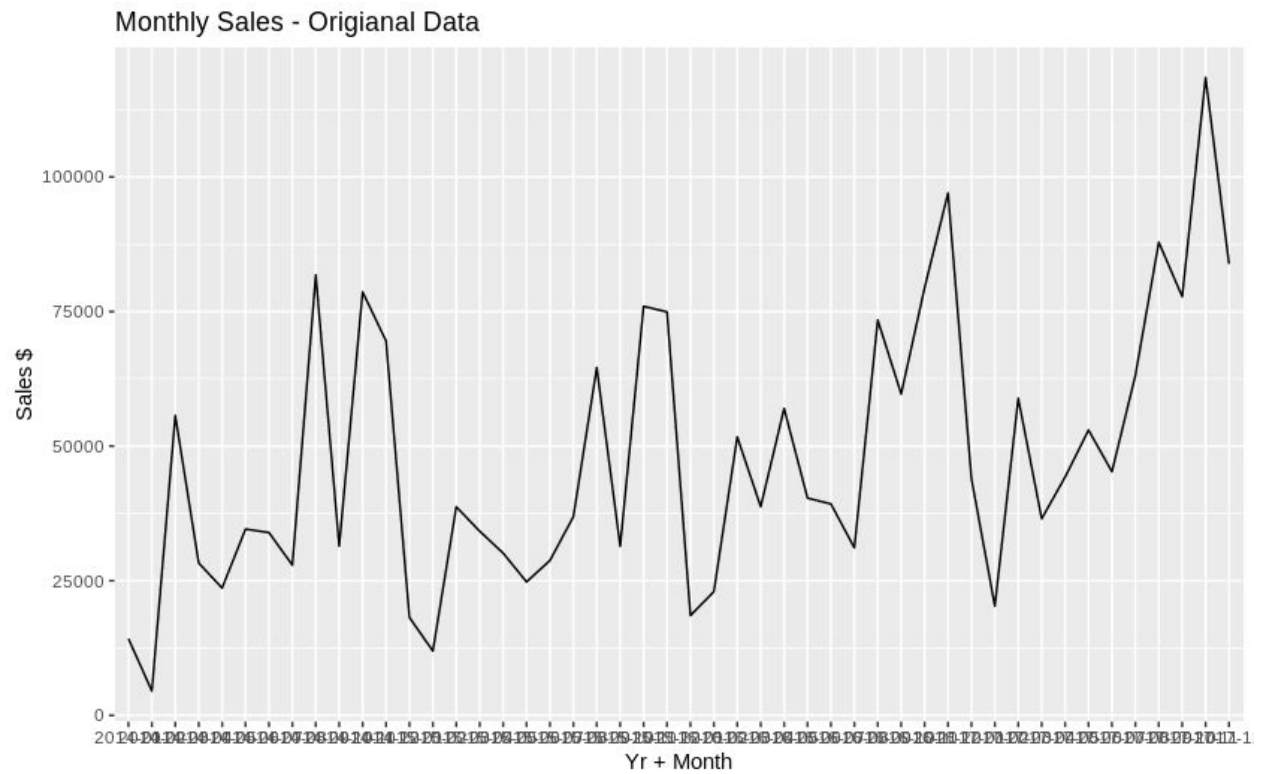
	yr_month	Sales
1	2014-01	14236.895
2	2014-02	4519.892
3	2014-03	55691.009
4	2014-04	28295.345
5	2014-05	23648.287
6	2014-06	34595.128
7	2014-07	33946.393
8	2014-08	27909.468
9	2014-09	81777.351
10	2014-10	31453.393
11	2014-11	78628.717
12	2014-12	69545.621
13	2015-01	18174.076

```
# Plot Monthly Sales
```

```
' Mean and variance do not appear to be constant'
```

```
ggplot(df.sum, aes(x=df.sum$yr_month, y=df.sum$Sales, group=1)) + ggtitle('Monthly Sales -  
Original Data') +
```

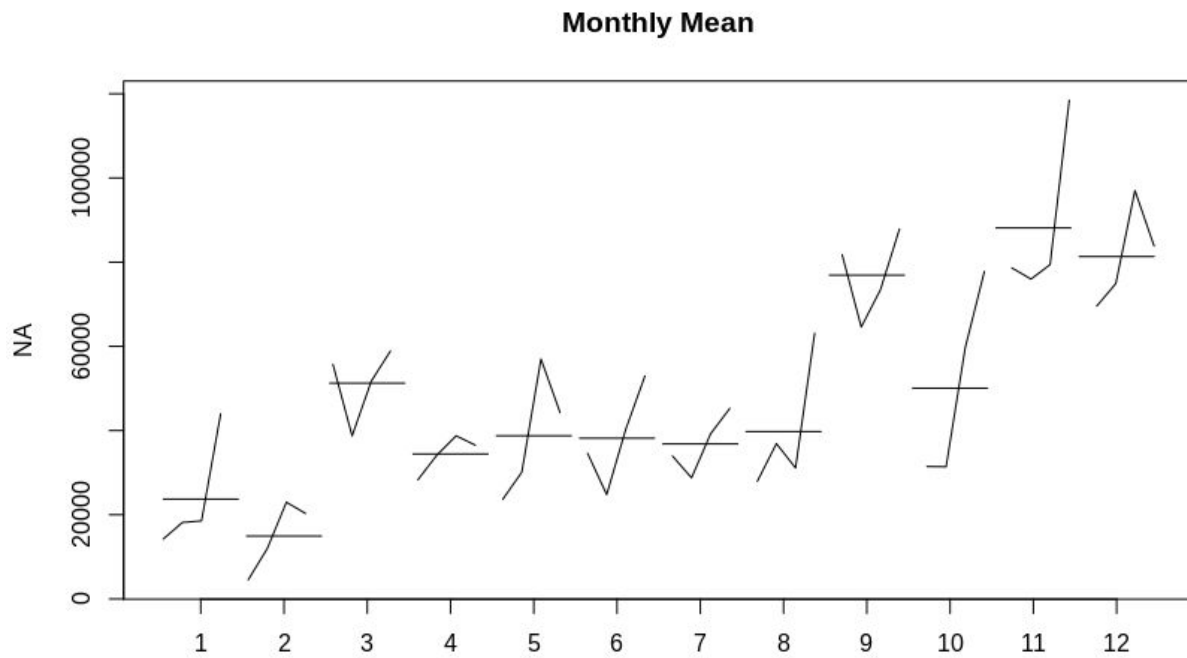
```
  xlab('Yr + Month') + ylab('Sales $') + geom_line()
```



```
# Check If Mean Constant
```

```
' Mean is not constant and therefore, the timeseries is not stationary'
```

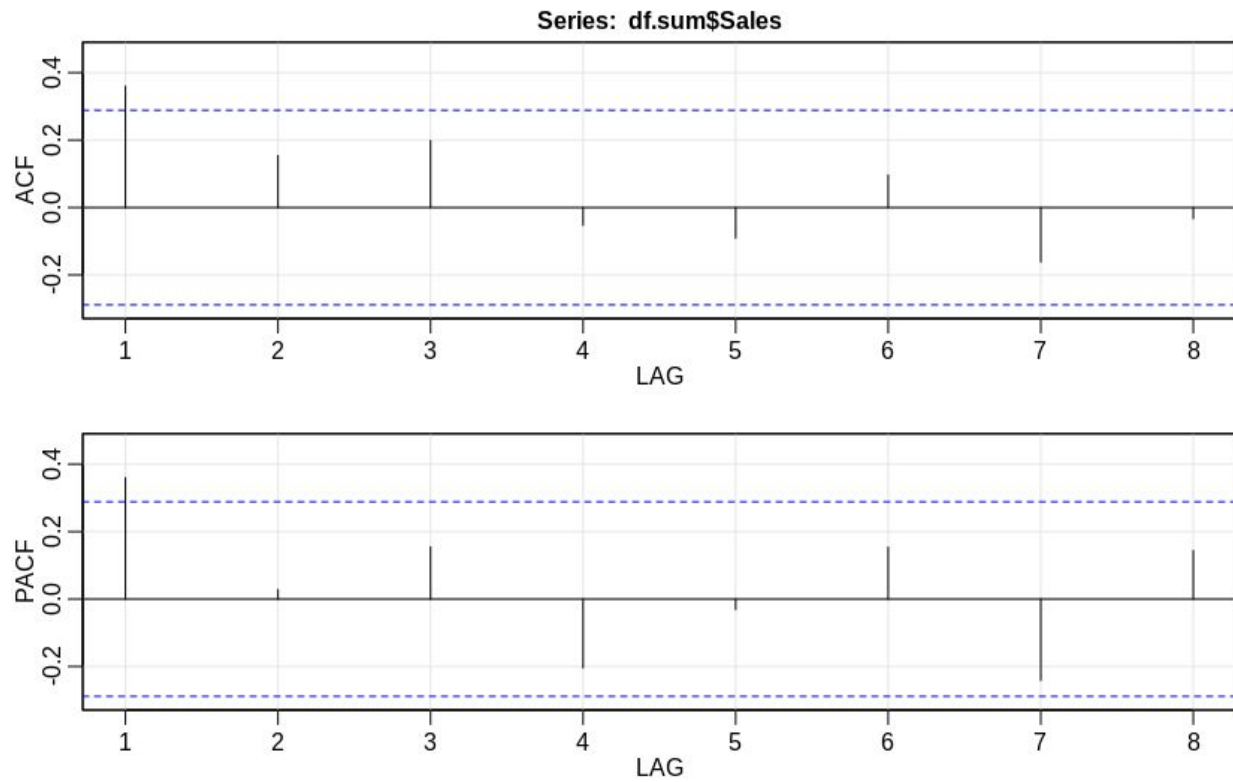
```
monthplot(df.sum$Sales, main='Monthly Mean')
```



ACF & PACF - Original Data

' ACF Looks like a slow decaying function up to lag=4 & the PACF drops from 0.36 to 0.03 or close to zero. This looks like an AR(1) process'

acf2(df.sum\$Sales)

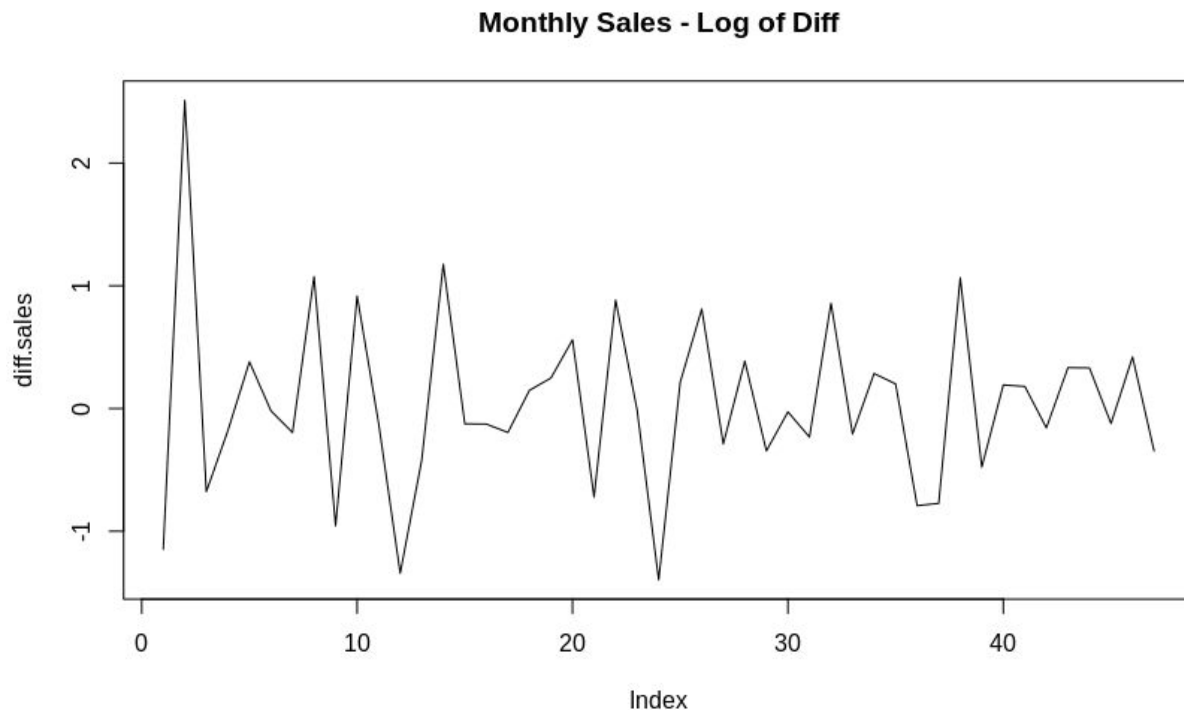


```
# Take diff and log of data
```

```
' Seems to e centered about the mean now. Some change in variance'
```

```
diff.sales <- diff(log(df.sum$Sales))
```

```
plot(diff.sales, type='l', main='Monthly Sales - Log of Diff')
```

Take ACF & PACF - Differenced Data

' Interpretation: This looks like an AR model based on the ACF and order of 2 based on the PACF'

```
acf2(diff.sales, max.lag=12)
```

FIT AR MODELS -----

YW Fit

' Note: $|\phi| < 1$ '

```
sales.yw <- ar.yw(diff.sales, order=2)
```

```
sales.yw$x.mean
```

```
sales.yw.mse <- sum(sales.yw$resid[3: length(sales.yw$resid)]^2) / length(sales.yw$resid[3 :  
length(sales.yw$resid)])
```

```
plot(sales.yw$resid, type='l')
```

```
sales.yw$ar # coefficients
```

```
sales.yw$asy.var.coef # covariance matrix
```

MLE Fit

```
sales.mle <- ar.mle(diff.sales, order=2)
```

```
sales.mle$x.mean
```

```
sales.mle.mse <- sum((sales.mle$resid[3: length(sales.mle$resid)])^2) / length(sales.mle$resid)
```

```
# YW Prediction
sales.pr <- predict(sales.yw, n.ahead=10)
comb <- c(diff.sales, sales.pr$pred)
ts.plot(comb, type='l', main = 'Plot YW Prediction') # Plot w/ prediction
lines(sales.pr$pred + sales.pr$se, lty=2) # conf intv
lines(sales.pr$pred - sales.pr$se, lty=2) # conf intv
```

```
# MLE Prediction
sales.pr <- predict(sales.mle, n.ahead=10)
comb <- c(diff.sales, sales.pr$pred)
ts.plot(comb, type='l', main='Plot MLE Prediction') # Plot w/ prediction
lines(sales.pr$pred + sales.pr$se, lty=2) # conf intv
lines(sales.pr$pred - sales.pr$se, lty=2) # conf intv
```

```
# Compare MSE
' MLE has the lower MSE'
paste('MLE MSE => ', sales.mle.mse)
paste('YW MSE => ', sales.yw.mse)
```

FIT ARIMA MODELS -----

```
# Arima Model
sales.mle.ar1 <- arima(diff.sales, c(1,0,0))
sales.mle.ar2 <- arima(diff.sales, c(2,0,0))
```

```
# MA Model
sales.mle.ma1 <- arima(diff.sales, c(0,0,1))
sales.mle.ma2 <- arima(diff.sales, c(0,0,2))
```

```
# AR(1) AR(2) Predictions
ar1.pr <- predict(sales.mle.ar1, n.ahead=10)
ar1.mse <- sum(ar1.pr$se^2) / length(ar1.pr$se)
ar2.pr <- predict(sales.mle.ar2, n.ahead=10)
ar2.mse <- sum(ar2.pr$se^2) / length(ar2.pr$se)
```

```
# MA(1) MA(2) Predictions
ma1.pr <- predict(sales.mle.ma1, n.ahead=10)
ma1.mse <- sum(ma1.pr$se^2) / length(ma1.pr$se)
dma2.pr <- predict(sales.mle.ma2, n.ahead=10)
ma2.mse <- sum(ma2.pr$se^2) / length(ma2.pr$se)
```

```
# Compare Results
mse.results <- c(ar1.mse, ar2.mse, ma1.mse, ma2.mse)
mse.labels <- c('AR1', 'AR2', 'MA1', 'MA2')
barplot(mse.results, names.arg = mse.labels, main= 'MSE RESULTS')
```